

or commonly know as:

"a bit more transparent way to state
your research assumptions and questions"

Jose Rivera

May 5, 2022

What are we going to talk about? I

- 1 About research
 - A typical scientific lab
 - Research hypothesis production
- 2 DAGs and PP
- 3 Example cases
 - Experimental design: the panacea
 - Simulation conventions
 - Fork bias: spurious relationships
 - Fork bias: masked relationships (a)
 - Fork bias: masked relationships (b)
 - Fork bias: multicollinearity



What are we going to talk about? II

- No more fork bias: neutral control
- Pipe bias: precision parasite
- Pipe bias: post-treatment
- Pipe bias: masked relationships
- Pipe/Fork bias: bias amplification
- Collider bias: M-bias
- Descendant bias: case control
- 4 Concluding remarks
- 5 Do you wanna know more???



1. About research

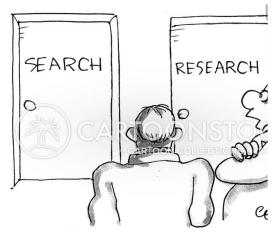
A typical scientific lab



A typical scientific lab¹

What is needed?

- 1. Quality of theory
- 2. Quality of data
- 3. Reliable procedures and code
- 4. Quality of data analysis
- 5. Documentation
- 6. Reporting



"DO WE NEED BOTH ?"

¹McElreath [12], lecture 20 and McElreath [13], chapter 17



A typical scientific lab

What we "normally" focus on?

- 1. Quality of theory
- 2. Quality of data
- 3. Reliable procedures and code
- 4. Quality of data analysis
- 5. Documentation
- 6. Reporting





A typical scientific lab

What can be improved? (with DAGs and PP)

- 1. Quality of theory
- 2. Quality of data
- 3. Reliable procedures and code
- 4. Quality of data analysis
- 5. Documentation
- 6. Reporting





1. About research

Research hypothesis production



Research hypothesis production

Well known challenges^a

- Insufficient data
- Wrong population
- Measurement error
- Selection bias
- Confounding



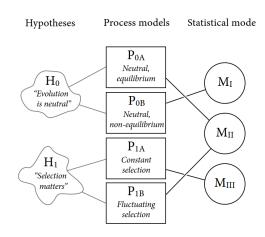


^aHernán [8], lesson 4

Research hypothesis production

but we should not forget^a

- No one-to-one relationship exists between our process models and statistical models,
- Nor between our hypothesis and a process models





^aFigure 1.2 reproduced from chapter 1 McElreath [13]

Research hypothesis production

and also

statistical models are just
 "machines to find association", not
 a reliable reflection of the theory
 (I can prove it!!).



Research hypothesis schematics²

- a. Estimand and process model
- b. Synthetic data generation
- c. Statistical model design and testing
- d. Apply statistical model to data



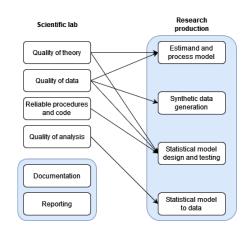
²McElreath [13], lecture 20, Pearl [16]. Follow Fogarty et al. [6] on item (c).



Research hypothesis schematic

Where does it match with the previous?

- a. Estimand and process model maps 1 (theory) and 2 (data) to a heuristic model.
- b. Synthetic data generation maps 2 (data) to an idealized data.
- c. Statistical model design and testing maps 1 (theory), 2 (data), and 3 (reliable code) to an statistical model.
- d. Apply statistical model to data maps 4 (analysis) onto a result.

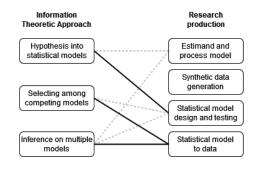




Where does the ITA fit?

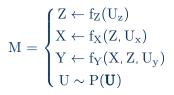
Information Theoretic Approach (ITA) is framework to select among competing models [1, 3]:

- 1. Hypothesis into statistical models, (how about a process model?)
- 2. Select among competing models, (do the code works as intended?)
- Make inferences based on one or multiple models. (do the code works as intended?, are there variables that can bias our results?)

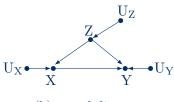




- Directed acyclic graphs (DAGs), are a type of structural causal model (SCM) [15, 4]
- DAGs can be represented by a structural model, and its associated causal diagram^a.
- we put distributional assumptions to the structural model through probabilistic programming (PP) [10]. (more in part 3)



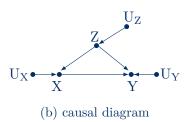
(a) structural model



^areproduced from Cinelli et al. [4].

- $V = \{Z, X, Y\}$ are called endogenous variables.
- **U** = {U_Z, U_X, U_Y} are called exogenous variables.
 (drawn when strictly required)
- $\mathbf{F} = \{f_Z, f_X, f_Y\}$ are called structural equations.

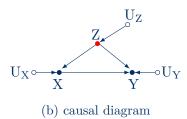
$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\textbf{U}) \end{cases}$$



Causal diagram conventions [4],

- black nodes are observed variables.
- white nodes are unobserved variables.
- red nodes are variables for which we will decide its inclusion or not.

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\textbf{U}) \end{cases}$$

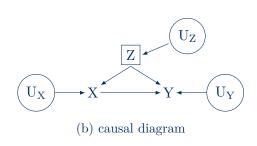




Other causal diagram conventions,

- no circle nodes are observed variables.
- circled nodes are unobserved variables.
- squared nodes are variables for which we will decide its inclusion or not.

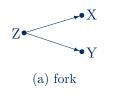
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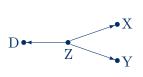




The benign case of DAG elementals

For everything can be depicted with them



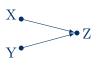


(d) descendant on fork

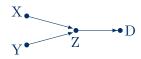




(e) descendant on pipe



(c) collider



(f) descendant on collider

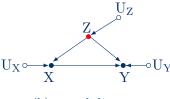
About D-separation

Causal graph theory [14, 15, 17, 18, 19],

- descendant (child, grandchild), parent (grandparent).
 (path specific)
- 2. paths (directional, non-directional).
- 3. paths are blocked or open according to the D-separation rules.
 (also path specific)
- 4. there are only four (4) D-separation rules.

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



About D-separation

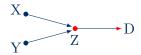
The D-separation (Directional) rules [8],

- If no variables being conditioned on, a path is blocked if and only if, two arrowheads on the path collide at some variable on the path.
- 2. Any path that contains a noncollider that has been conditioned on, is blocked (backdoor path)^a.
- 3. A collider that has been conditioned on does not block a path.
- 4. A collider that has a descendant that has been conditioned on does not block a path.











^athere is also a front-door path (if you wonder).

About D-separation

The D-separation rules implications, (independent of distributional assumptions)

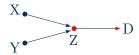
1.
$$X \perp \!\!\!\perp Y \Longrightarrow P(X, Y) = P(X) \cdot P(Y)$$

- 2. $X \perp \!\!\!\perp Y \mid Z \Longrightarrow$ $P(X, Y \mid Z) = P(X \mid Z) \cdot P(Y \mid Z)$ (same for fork or pipe)
- 3. $X \not\perp \!\!\! \perp Y | Z \Longrightarrow$ $P(X, Y|Z) \neq P(X|Z) \cdot P(Y|Z)$
- 4. $X \not\perp \!\!\! \perp Y \mid D \Longrightarrow P(X, Y \mid D) \neq P(X \mid D) \cdot P(Y \mid D)$



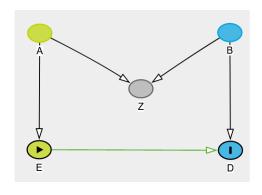






Oh DAGitty!! mijn vriendin

- browser (R package) environment for creating, editing, and analyzing causal diagrams [20].
- available online: http://dagitty.net
- But there are more fish in the sea: http://www.causalfusion.net [2] (b**** better have my \$\$\$)





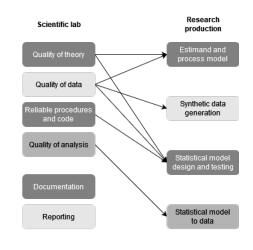
Where do DAGs and PP fit?

starts with:

- A clear definition of the estimand and process model (assumptions).
- An improved the reliability of your procedures.
- As a documentation procedure.

and leads to:

- A sound analysis, and result (even when we cannot have an answer to our question)
- An improved planning to get data.



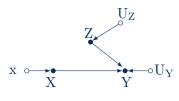




Experimental design³

- Purpose: to control all factors responsible for the outcome's variation.
 (understand the system)
- It is modeled by modifying the structural model (and causal diagram).

(a) structural model



³Cinelli et al. [4], appendix A (p. 15)



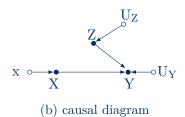
 $M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(x) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$

Experimental design

- intervention on X can be written in do-calculus^a as: $P(\mathbf{V} \mid do(X = x))$.
- remember:

$$\begin{aligned} & \boldsymbol{V} = \{Z, X, Y\}, \\ & \boldsymbol{U} = \{U_Z, U_X, U_Y\}, \text{ and } \\ & \boldsymbol{F} = \{f_Z, f_X, f_Y\}. \end{aligned}$$

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(x) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\textbf{U}) \end{cases}$$



^aan appropriate treatment can be found with the usual suspects [14, 15, 17, 18])

Effects of interest

two types of effects,

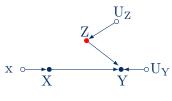
- 1. Average causal effect: ACE(x) = E[Y|do(x+1)] E[Y|do(x)]
- 2. Controlled direct effect: $CDE(x,z) = E[Y|do(x+1),do(z)] \\ E[Y|do(x),do(z)]$

points to consider:

- CDE takes a particular relevance with observational data.
- There is also a distinction between total effect and direct effect.

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_x(x) \\ Y \leftarrow f_Y(X, Z, U_Y) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



3. Example cases

Simulation conventions



Simulation conventions

one way to defined it,

$$Z = U_Z$$
 ; $U_Z \sim N(0, \sigma_Z)$

$$X = \beta_Z Z + U_X \hspace{1cm} ; U_X \sim N(0, \sigma_X)$$

$$Y = \beta_Z Z + \beta_X X + U_Y$$
; $U_Y \sim N(0, \sigma_Y)$

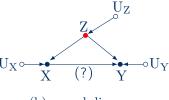
a more succinct way,

$$Z \sim N(0, \sigma_Z)$$

 $X \sim N(\beta_Z Z, \sigma_X)$
 $Y \sim N(\beta_Z Z + \beta_X X, \sigma_Y)$

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(Z, X, U_Y) \\ U \sim P(U) \end{cases}$$

(a) structural model





Spurious relationships⁴

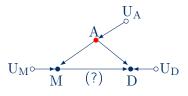
also known as,

- \blacksquare spurious association
- confounder
- \blacksquare an instance of fork bias

research question,

- Does M has a (direct) effect on D? variables,
 - A, median age at marriage
 - M, marriage rate
 - D, divorce rate

(a) structural model



⁴McElreath [12], chapter 05 (p. 125)



 $M = \begin{cases} A \leftarrow f_A(U_A) \\ M \leftarrow f_M(A, U_M) \\ D \leftarrow f_D(A, M, U_D) \\ U \sim P(\boldsymbol{U}) \end{cases}$

Simulation setting

```
# sim
A = rnorm( 100 )
M = rnorm( 100 , mean=-1*A )
D = rnorm( 100 , mean=-1*A + 0*M )
d = data.frame(A=A, M=M, D=D)
```

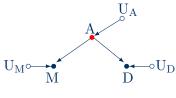
(c) R code

Implications,

- M #D
- M ⊥LD | A

$$M = \begin{cases} A \leftarrow f_A(U_A) \\ M \leftarrow f_M(A, U_M) \\ D \leftarrow f_D(A, U_D) \\ U \sim P(\textbf{U}) \end{cases}$$

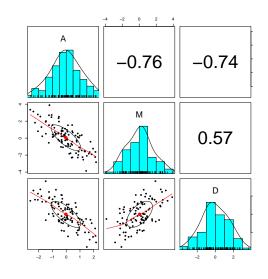
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(A, D) < 0 and cor(M, D) > 0 goes in line of our "rudimentary" understanding of the data.
- why there is cor(M, D) > 0? (hint: univariate correlation)
- we include M as a covariate in our statistical model
 (is our research hypothesis)





Regression, regression!!

based on statistical analysis,

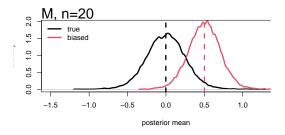
• we have two different stories, (which one is the "truth"?)

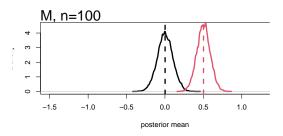
```
summary(lm(D ~ M, data=d)) # spurious relation
lm(formula = D \sim M, data = d)
Residuals:
               10 Median
-2.80012 -0.90447 -0.03866 0.80220 2.82970
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
                        0.12412 -1.877
                        0.08986 4.477 2.04e-05 ***
             0.40233
> summary(lm(D \sim A + M, data=d)) # controlled relation
lm(formula = D \sim A + M, data = d)
Residuals:
              10 Median
                                         Max
-2.27295 -0.68174 0.03781 0.78885 2.95320
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.18854
            -1.03121
            -0.06134
                        0.09362
                                -0.655
```

I'll get more data!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the more certain you are about your biased estimates (the winner's curse)







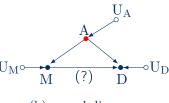
The dream team!!

based on DAG and statistical model,

- the 2nd D-separation rule requires you to control any noncollider to block the backdoor path,
 i.e. M ⊥LD | A
- conditioning on A we can find, E[D|do(m)] = E[E[D|M = m, A]] (law of total expectation)
- then we can find the ACE(m) = E[D|do(m+1)] E[D|do(m)] (Frisch-Waugh-Lovell theorem)

$$M = \begin{cases} A \leftarrow f_A(U_A) \\ M \leftarrow f_M(A, U_M) \\ D \leftarrow f_D(A, M, U_D) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



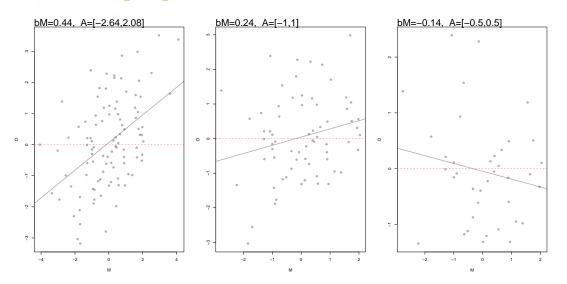
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based on DAG and statistical analysis,

■ the less biased model is the second, (assuming our DAG is true)



So, what is going on?







Masked relationships $(a)^5$

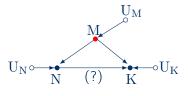
also known as,

- omitted variable bias
- an instance of fork bias

research question,

- Does N has a (direct) effect on K? variables.
 - M, mammal mass in kg.
 - N, ratio neocortex over total brain mass
 - K, Kcal. per gram of milk

(a) structural model



⁵McElreath [12], chapter 05 (p. 144)



 $M = \begin{cases} M \leftarrow f_M(U_M) \\ N \leftarrow f_N(M, U_N) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$

Simulation setting

```
# sim
M = rnorm( 100 )
N = rnorm( 100 , 1*M )
K = rnorm( 100 , 1*N + -1*M )
d = data.frame(N=N, M=M, K=K)
```

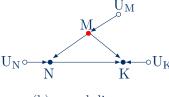
(c) R code

Implications,

- N #K
- N #K | M

$$M = \begin{cases} M \leftarrow f_M(U_M) \\ N \leftarrow f_N(M, U_N) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$$

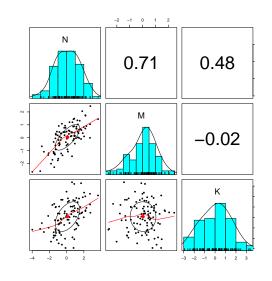
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"Eyeballing" analysis

based on correlation analysis,

- cor(N, K) > 0 goes in line of our "rudimentary" understanding of the data.
- but why there is $cor(M, k) \approx 0$? (hint: univariate correlation)
- we might not include M as a covariate in our statistical model





Regression, regression!!

based on statistical analysis,

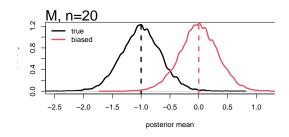
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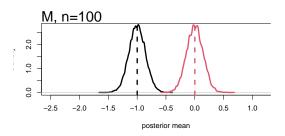
```
summarv(lm(K \sim N. data=d)) # biased estimate
Call:
lm(formula = K \sim N, data = d)
Residuals:
    Min
             10 Median
-2.8355 -0.8110 0.0188 0.7897 3.4276
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.01401
                        0.09332
                                 5.680 1.38e-07 ***
             0.53002
 summary(lm(K \sim N + M, data=d)) # less biased estima
Call:
lm(formula = K \sim N + M, data = d)
Residuals:
              10 Median
                                         Max
-2.50873 -0.72626 -0.01968 0.69016 2.93000
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.22096
                        0.09845
             0.95510
                        0.10089
                                  9.466 1.91e-15
            -1.06246
                        0.15462 -6.871 6.14e-10 ***
```

I'll get more data!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the more certain you are about your biased estimates







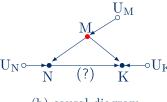
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based on DAG and statistical model,

- the 2nd D-separation rule requires you to control any noncollider to block the backdoor path, i.e. N 业K | M
- conditioning on M we can find, E[K|do(n)] = E[E[K|N = n, M]] (law of total expectation)
- then we can find the ACE(n) = E[D|do(n+1)] E[D|do(n)] (Frisch-Waugh-Lovell theorem)

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(a) structural model



the dream team!!

based on DAG and statistical analysis,

■ the less biased model is the second, (assuming our DAG is true)



So, what is going on? $fork2_triptych.pdf$

3. Example cases

Fork bias: masked relationships (b



Masked relationships $(b)^6$

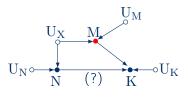
also known as,

- (unobserved) omitted variable bias
- an instance of fork bias

research question,

- Does N has a (direct) effect on K? variables.
 - U_X, unobservable (e.g. genetics)
 - M, mammal mass in kg.
 - N, neocortex over total brain mass
 - K, Kcal. per gram of milk

(a) structural model



⁶McElreath [12], chapter 05 (p. 144)



 $M = \begin{cases} N \leftarrow f_N(U_N, U_X) \\ M \leftarrow f_M(U_M, U_X) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$

Simulation setting

```
# sim
U = rnorm( 100 )
N = rnorm( 100 , 1*U )
M = rnorm( 100 , 1*U )
K = rnorm( 100 , 1*N + -1*M )
d = data.frame(U=U,N=N,M=M,K=K)
```

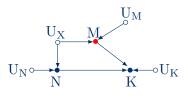
(c) R code

Implications,

- N #K
- N #K | M

$$M = \begin{cases} N \leftarrow f_N(U_N, U_X) \\ M \leftarrow f_M(U_M, U_X) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$$

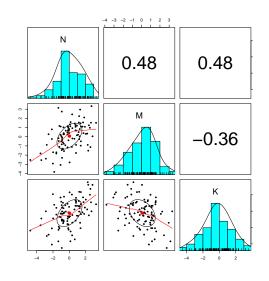
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(N, K) > 0 goes in line of our "rudimentary" understanding of the data.
- cor(M, K) < 0 does NOT goes in line of our "rudimentary" understanding of the data. (hint: univariate correlation)
- we include M as a covariate in our statistical model (by chance?)



Regression, regression!!

based on statistical analysis,

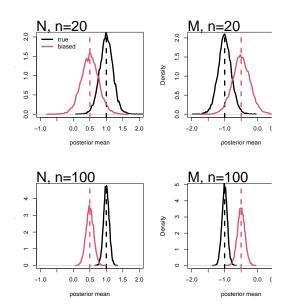
• we have two different stories, (which one is the "truth"?)

```
summarv(lm(K \sim N. data=d)) # unobserved path still
Call:
lm(formula = K \sim N. data = d)
Residuals:
             10 Median
 -3.7763 -0.8480 0.1497 0.9874 3.3530
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.24867
                        0.14573 -1.706 0.0911
                        0.09502 5.410 4.46e-07 ***
             0.51406
 summary(lm(K \sim N + M, data=d)) # unobserved path c
Call:
lm(formula = K \sim N + M, data = d)
Residuals:
     Min
               1Q Median
                                          Max
-2.58218 -0.58434 -0.00579 0.72016 1.78724
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.19978
             0.90893
                        0.06958 13.064
                                           <2e-16 ***
            -0.89676
                        0.07572 -11.843
                                           <2e-16 ***
```

I'll get more data!!

imagine we can continue sampling,

- \bullet top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the more certain you are about your biased estimates



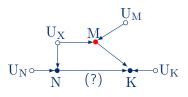
The dream team!!

based on DAG and statistical model,

- the 2nd D-separation rule requires control on any noncollider to block the backdoor path,
 i.e. N ⊥K | U_X
 (but it is unobservable)
- still we use the 2nd D-separation rule by controlling for M,
 i.e. N ⊥K | M
- conditioning on M we can still find, E[K|do(n)] = E[E[K|N = n, M]](law of total expectation)
- then we can find the ACE(n) = E[D|do(n+1)] E[D|do(n)] (Frisch-Waugh-Lovell theorem??)

$$M = \begin{cases} N \leftarrow f_N(U_N, U_X) \\ M \leftarrow f_M(U_M, U_X) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



the dream team!!

based on DAG and statistical analysis,

■ the less biased model is the second, (assuming our DAG is true)

Similar scenario⁷

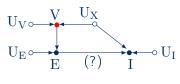
research question,

■ Does E has a (direct) effect on I? variables,

- U_X, unobservable (e.g. family context)
- V, personal values
- E, education
- I, income

$$M = \begin{cases} V \leftarrow f_M(U_V, U_X) \\ E \leftarrow f_E(V, U_E) \\ I \leftarrow f_I(E, U_X, U_I) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



⁷Cinelli et al. [4] (p. 3), McElreath [13], lecture 6



So, what is going on? $fork3_triptych.pdf$

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Multicollinearity⁸

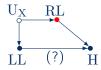
also known as,

- extreme case of masked relationships
- an instance of fork bias

research question,

- Should we include RL in our model? variables,
 - U_X, unobservable (e.g. genetics and context)
 - LL, individual's left leg
 - RL, individual's right leg
 - H, individual's height

(a) structural model



⁸McElreath [12], chapter 06 (p. 163)



 $M = \begin{cases} LL \leftarrow f_L(U_X) \\ RL \leftarrow f_L(U_X) \\ H \leftarrow f_K(RL, LL) \\ U \sim P(\textbf{U}) \end{cases}$

Simulation setting

```
# backward simulation
H = round( rnorm( 100 , 170, 2), 1)
Lp = runif( 100 , 0.5-0.05, 0.5+0.05)
LL = round( Lp*H + rnorm( 100 , 0, 1 ), 1)
RL = round( Lp*H + rnorm( 100 , 0, 1 ), 1)
d = data.frame(LL,RL,H)
```

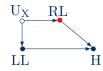
(c) R code

Implications,

■ LL #RL

$$M = \begin{cases} LL \leftarrow f_L(U_X) \\ RL \leftarrow f_L(U_X) \\ H \leftarrow f_K(RL, LL) \\ U \sim P(\textbf{U}) \end{cases}$$

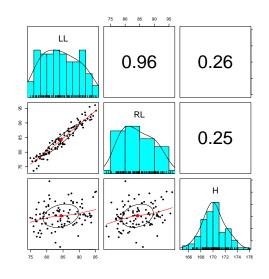
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(LL, H) > 0, cor(RL, H) > 0 and cor(LL, RL) > 0 goes in line of our understanding of the data.
- we might not include RL as a covariate in our statistical model (based on univariate correlation)





Regression, regression!!

based on statistical analysis,

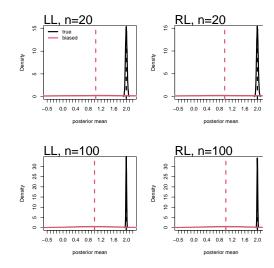
- the second regression show a smaller effect of RL,
- the second regression show larger SE values

```
summary(lm(H \sim -1 + RL, data=d)) # unbiased
lm(formula = H \sim -1 + RL, data = d)
Residuals:
                    Median
                                 3Q
     Min
-18.1223 -7.6047
                    0.8014
                             9.3119 17.3240
Coefficients:
   Estimate Std. Error t value Pr(>|t|)
RL 2.00234
               0.01198 167.2 <2e-16 ***
> summary(\lim(H \sim -1 + LL + RL, data=d)) # ine
Call:
lm(formula = H \sim -1 + LL + RL, data = d)
Residuals:
    Min
             10 Median
                                      Max
-20.357 -8.608
                 1.473 10.207 19.770
Coefficients:
   Estimate Std. Error t value Pr(>|t|)
                                  0.4856
     0.5090
                0.7271
                          0.700
     1.5017
                0.7284
                          2.062
                                  0.0419 *
```

I'm sure data won't help

imagine we can continue sampling,

- \bullet top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the less certain are your biased estimates





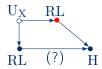
Not so great now??

based on DAG and statistical model,

- the 2nd D-separation rule requires control on any noncollider to block the backdoor path,
 i.e. LL ⊥H | U_X
 (but it is unobservable)
- we still use the 2nd D-separation rule by controlling for RL, but still we have LL ⊥H | RL
- issue goes beyond the backdoor path RL and LL provide the same information, i.e. singular matrix, (is like having a causal model like b)

$$M = \begin{cases} LL \leftarrow f_L(U_X) \\ RL \leftarrow f_L(U_X) \\ H \leftarrow f_K(RL) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



Not so great now??

based on DAG and statistical analysis,

■ the less biased model is the first, (assuming our DAG is true)

```
summary(lm(H \sim -1 + RL, data=d)) # unbiased
lm(formula = H \sim -1 + RL, data = d)
Residuals:
     Min
               10
                    Median
                                  30
                                           Max
-18.1223 -7.6047
                    0.8014
                              9.3119 17.3240
Coefficients:
   Estimate Std. Error t value Pr(>|t|)
RL 2.00234
               0.01198
                          167.2
```

So, what is going on? $fork4_triptych.pdf$

3. Example cases

No more fork bias: neutral contro



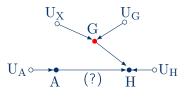
Neutral control⁹

also known as,

- precision "booster"
- similar to experimental design research question,
- Should we include G on our model? variables,
 - A, "hearing" age
 - G, gender
 - \blacksquare U_X, unobservable (e.g. no idea yet)
 - H, inverse logit of entropy (approximate of speech intelligibility)

$$M = \begin{cases} G \leftarrow f_G(U_G, U_X) \\ A \leftarrow f_A(U_A) \\ H \leftarrow f_H(A, G, U_H) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model





⁹Cinelli et al. [4] (p. 4)

Simulation setting

```
# sim
G = sample( 0:1, 100 , replace=T )
A = rnorm( 100 )
H = rnorm( 100 , -1*A + -1*G )
d = data.frame(G=G,A=A,SI=SI)
```

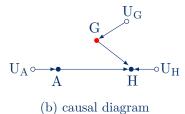
(c) R code

Implications,

- A ⊥LG
- A #H
- G #H

$$M = \begin{cases} G \leftarrow f_G(U_G) \\ A \leftarrow f_A(U_A) \\ H \leftarrow f_H(A, G, U_H) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

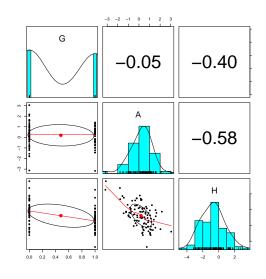
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(G, H) < 0, $cor(G, A) \approx 0$ and cor(A, H) < 0 goes in line of our "rudimentary" understanding of the data.
- we include both as a covariate in our statistical model





Regression, regression!!

based on statistical analysis,

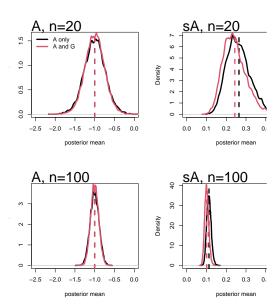
- almost no change on our estimates,
- lower standard errors for A when G is included

```
summary(lm(H \sim A, data=d)) # correct estimate
Call:
lm(formula = H \sim A, data = d)
Residuals:
             10 Median
-3.4714 -0.8797 -0.0633 0.8963 2.4346
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.5770
                         0.1216 -4.746 7.07e-06 ***
             -0 8410
                         0.1183 -7.108 1.92e-10 ***
Call:
lm(formula = H \sim A + G, data = d)
Residuals:
    Min
             10 Median
-2.7994 -0.6914 0.0579 0.7796 1.8274
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.03317
            -0.87360
                        0.10090 -8.658 1.05e-13
            -1.25786
                        0.20371 -6.175 1.55e-08
```

I'll get more data!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the second model, the larger the sample size,
 - the more certain you are about your non-biased estimates (under the any model)



So, what is going on? $fork5_triptych.pdf$

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3. Example cases

Pipe bias: precision parasite



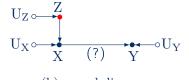
Precision parasite¹⁰

research question,

- Should we include Z in the model?
- variables,
 - **■** Z,
 - X,
 - Y,

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



¹⁰McElreath [13], lecture 6, Cinelli et al. [4] (p. 5)



Simulation setting

```
# Sim
Z = rnorm( 100 )
X = rnorm( 100 , 1*Z )
Y = rnorm( 100 , 1*X )
d = data.frame(Z,X,Y)
```

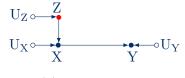
(c) R code

Implications,

- X #Y
- Z #Y | X

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

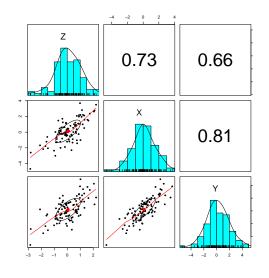
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(Z, X) > 0 is not large enough to discard it as multicollinearity.
- cor(Z, Y) > 0 and cor(X, Y) > 0 indicate both should be in our model
 (it might be our research hypothesis)





Regression, regression!!

based on statistical analysis,

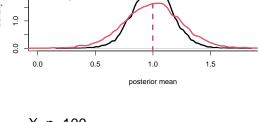
- no bias in parameter if Z is in,
- but we loose precision on X

```
> summary(lm(Y ~ X, data=d)) # unbiased effect, more
lm(formula = Y \sim X, data = d)
Residuals:
     Min
               10 Median
                                          Max
-2.41746 -0.73659 -0.09384 0.63812 2.10338
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.03433
             1 16908
                        0.06717 17.405
                                           <2e-16 ***
 summary(lm(Y \sim X + Z, data=d)) # unbiased effects,
Call:
lm(formula = Y \sim X + Z, data = d)
Residuals:
               10 Median
                                          Max
-2.27019 -0.74072 -0.06355 0.66643 2.20770
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.04722
                                            0.649
             1.08881
                                 10.512
                                           <2e-16 ***
             0.15431
                        0.15159
                                  1.018
                                            0.311
```

With more data??

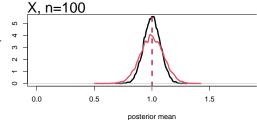
imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the second model, the larger the sample size,
 - still less precise estimates
 - more difficult to test hypothesis



X, n=20

less precise





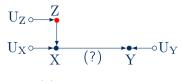
Now, what is going on here??

based on DAG and statistical model,

conditioning on Z reduces variation on X

$$M = \begin{cases} Z \leftarrow f_Z(U_Z) \\ X \leftarrow f_X(Z, U_X) \\ Y \leftarrow f_Y(X, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



Now, what is going on here??

based on DAG and statistical analysis,

■ the more appropriate model (for inference) is the first,
(assuming our DAG is true)

So, what is going on? $pipe1_triptych.pdf$

3. Example cases

Pipe bias: post-treatment



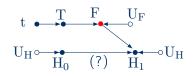
Post-treatment bias¹¹

case of,

- full mediation
- research question,
- Does the treatment T works? variables,
 - H_0 , height of plant at t=0
 - T, antifungal treatment
 - F, presence of fungus
 - H_1 , height of plant at t = 1

$$\mathbf{M} = \begin{cases} \mathbf{H}_0 \leftarrow \mathbf{f}_{H}(\mathbf{U}_{H}) \\ \mathbf{T} \leftarrow \mathbf{f}_{T}(\mathbf{t}) \\ \mathbf{F} \leftarrow \mathbf{f}_{F}(\mathbf{T}, \mathbf{U}_{F}) \\ \mathbf{H}_1 \leftarrow \mathbf{f}_{H}(\mathbf{F}, \mathbf{H}_0, \mathbf{U}_{H}) \\ \mathbf{U} \sim \mathbf{P}(\mathbf{U}) \end{cases}$$

(a) structural model



¹¹McElreath [12], chapter 6 (p. 170)



Simulation setting

```
# sim
h0 = rnorm( 100 , 10, 2)
Tr = rep( 0:1 , each=100/2 )
Fu = rbinom( n , size=1 , prob=0.5 + -0.4*Tr )
h1 = h0 + rnorm( n , 5 + -3*Fu)
d = data.frame( h0=h0, h1=h1, Tr=Tr, Fu=Fu )
```

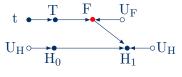
(c) R code

Implications,

- T ⊥⊥H₀
- \blacksquare T $\not\perp H_1$
- \blacksquare T $\not\perp$ H₁ | F

$$M = \begin{cases} H_0 \leftarrow f_H(U_H) \\ T \leftarrow f_T(t) \\ F \leftarrow f_F(T, U_F) \\ H_1 \leftarrow f_H(F, H_0, U_H) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



Descriptive analysis

based on descriptive analysis,

- positive change in height with treatment.
- negative change in height with fungus.
- diluted relationship for T when both are in the model
 (hint: blocking path of information)

```
mean
                sd
                              se
                       50 0.248
                       50 0.205
Fu
       mean
       4.93 0.953
                      62 0.121
       1.81 0.902
      Fu
              mean
                                   se
                             19 0.186
              1.76 0.861
                             31 0.155
              4.93 1.02
                             43 0.155
              2.011.12
                              7 0.423
```



Again regression!!

based on statistical analysis we have two different stories (but not quite),

- treatment has a significant effect,
- but gets completely diluted when fungus is considered in the model

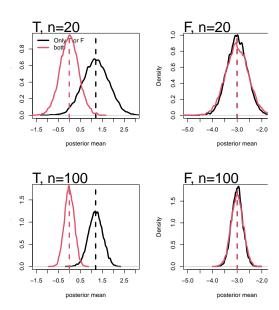
```
summary(lm(h1-h0 \sim Tr, data=d)) # only treatment
lm(formula = h1 - h0 \sim Tr, data = d)
Residuals:
             10 Median
-4.1166 -1.0929 0.1755 1.2621 3.3990
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)
              1 5656
                         0.3215
> summary(lm(h1-h0 ~ Tr + Fu, data=d)) # only fungus
lm(formula = h1 - h0 \sim Tr + Fu. data = d)
Residuals:
    Min
             10 Median
                                     Max
-2.1691 -0.4823 0.0963 0.5315 2.0357
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 4.86397
                                           <2e-16 ***
             0.09138
                        0.21579 0.423
                                            0.673
            -3.07122
                        0.22229 -13.816
                                           <2e-16 ***
```



I can guess what happens with more data!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the "incorrect" model, the larger the sample size,
 - the more certain you are about your biased T estimates (not F) (this result is not wrong!)





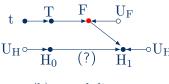
The dream team!!

based on DAG and statistical model,

- the 2nd D-separation states that if you to control any noncollider you block the backdoor path,
 i.e. T ⊥⊥H₁ | F
- therefore if we want to find if T = 1 works, we should not stratify by F

$$M = \begin{cases} H_0 \leftarrow f_H(U_H) \\ T \leftarrow f_T(t) \\ F \leftarrow f_F(T, U_F) \\ H_1 \leftarrow f_H(F, H_0, U_H) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model





the dream team!!

based on DAG and statistical analysis,

■ the model that answers our research question is the first one, (assuming our DAG is true)

So, what is going on? pipe2_triptych.pdf

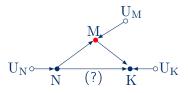


Masked relationships¹²

also known as,

- mediation
- Simpson's paradox
- an instance of pipe bias research question,
- Does N has a (direct) effect on K? variables,
 - M, mammal mass in kg.
 - N, neocortex over total brain mass
 - K, Kcal. per gram of milk

(a) structural model



¹²McElreath [12], chapter 05 (p. 144)



 $M = \begin{cases} N \leftarrow f_N(U_N) \\ M \leftarrow f_M(N, U_M) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$

Simulation setting

```
# sim
N = rnorm( 100 )
M = rnorm( 100 , 1*N )
K = rnorm( 100 , 1*N + -1*M )
d = data.frame(N=N,M=M,K=K)
```

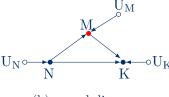
(c) R code

Implications,

- N #K
- N #K | M

$$M = \begin{cases} N \leftarrow f_N(U_N) \\ M \leftarrow f_M(M, U_M) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$$

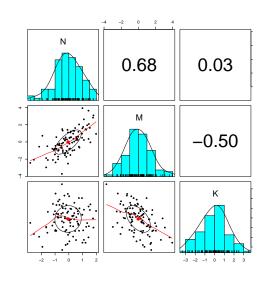
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(M, K) < 0 does NOT goes in line of our "rudimentary" understanding of the data.
- and why there is $cor(N, K) \approx 0$? (hint: univariate correlation)
- we include N as a covariate in our statistical model
 (is our research hypothesis)





Regression, regression!!

based on statistical analysis,

■ two regressions with two different results, which model is the "true"?

```
summary(lm(K \sim N, data=d)) # biased estimate
Call:
lm(formula = K \sim N, data = d)
Residuals:
             10 Median
                                     Max
-3.1751 -0.9009 0.1519 0.8574 3.6041
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.10412
                         0.13808
             0.05005
                        0.14487 0.345
                                            0.730
> summary(lm(K ~ N + M, data=d)) # less biased estimate
Call:
lm(formula = K \sim N + M, data = d)
Residuals:
              1Q Median
-2.58484 -0.59175 0.04378 0.61175 2.43360
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.06181
                        0.09825
            0.98297
                        0.13994
                                7.024 2.98e-10 ***
            -0.93107
                        0.09457 -9.846 2.89e-16 ***
```

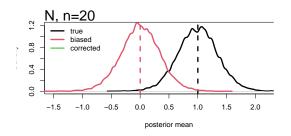
I'll get more data!!

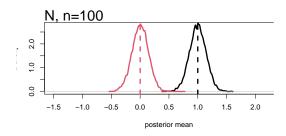
imagine we can continue sampling,

- \blacksquare top: 10,000 samples n = 20
- \blacksquare bottom: 10,000 samples n = 100

the larger the sample size,

- the more certain you are about your estimates
- the more mistaken you are about your research question (under the "incorrect" model)







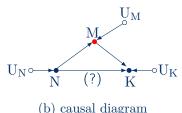
The dream team!!

based on DAG and statistical model,

- the 2nd D-separation rule requires you to control any noncollider to block the backdoor path,
 i.e. N ⊥K | M
- conditioning on M we can find, E[K|do(n)] = E[E[K|N = n, M]] (law of total expectation)
- then we can find the ACE(n) = E[D|do(n+1)] E[D|do(n)] (Frisch-Waugh-Lovell theorem)

$$M = \begin{cases} N \leftarrow f_N(U_N) \\ M \leftarrow f_M(M, U_M) \\ K \leftarrow f_K(M, N, U_K) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model





the dream team!!

based on DAG and statistical analysis,

■ the less biased model is the second, (assuming our DAG is true)

So, what is going on? $pipe3_triptych.pdf$

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Bias amplification¹³

also known as,

- (unobserved) omitted variable bias
- related to instrumental variables
- an instance of fork bias

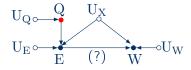
research question,

■ Do E has a (direct) effect on W? variables,

- Q, instrumental variable (e.g. quarter of the year)
- E, educational level
- \blacksquare U_X, unobservables (e.g. ability)
- W, future wages

$$M = \begin{cases} Q \leftarrow f_Q(U_Q) \\ E \leftarrow f_E(Q, U_X, U_E) \\ W \leftarrow f_W(E, U_X, U_W) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model



¹³McElreath [12], chapter 14 (p. 455), Cinelli et al. [4] (p. 5)



Simulation setting

```
# sim
U = rnorm( 100 )
Q = sample( 1:4, 100, replace=T )
E = rnorm( 100 , 1*Q + 1*U )
W = rnorm( 100 , 0*E + 1*U )
d = data.frame(U=U,Q=Q,E=E,W=W)
```

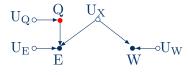
(c) R code

Implications,

- **■** E #W
- \blacksquare E $\bot\!\!\!\bot W \mid U_X \text{ (impossible)}$
- \blacksquare Q $\bot\!\!\!\bot U_X$ (cannot be tested)
- Q ⊭E
- Q ⊥LW | E (cannot be tested) (exclusion restriction)

$$M = \begin{cases} Q \leftarrow f_Q(U_Q) \\ E \leftarrow f_E(Q, U_X, U_E) \\ W \leftarrow f_W(U_X, U_W) \\ U \sim P(\textbf{U}) \end{cases}$$

(a) structural model

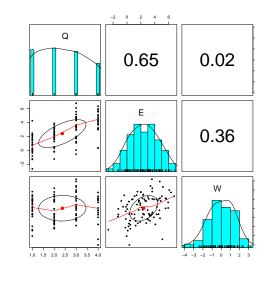


"Eyeballing" analysis

based on correlation analysis,

- cor(Q, E) > 0 and cor(E, W) > 0 goes in line of our "rudimentary" understanding of the data.
- cor(Q, W) > 0 tells you about the exclusion restriction?

 (hint: No)
- we might NOT include Q as a covariate in our statistical model (but is the instrumental variable!!!)





Regression, regression!!

based on statistical analysis,

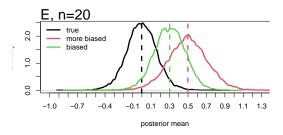
- two different stories (which model is the "truth"?)
- one is "worse"/"better" than the other?
- are both wrong?

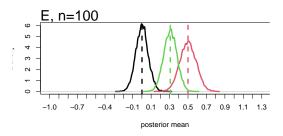
```
summary(lm(W ~ E, data=d)) # biased
lm(formula = W \sim E, data = d)
Residuals:
    Min
             10 Median
-4.0726 -0.9674 0.1771 0.9234 2.8787
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.60816
                       0.20506 -2.966 0.003793 **
             0.25408
                       0.06559 3.873 0.000194 ***
> summary(lm(W ~ E + O, data=d)) # more biased
call.
lm(formula = W \sim E + 0. data = d)
Residuals:
             10 Median
-3.7405 -0.9774 0.0879 0.9162 2.9825
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.12229
                        0.30650
             0.42054
                        0.08262
                                 5.090 1.75e-06 ***
            -0.47716
                        0.15361 -3.106 0.00249 **
```

I'll get more data!!

imagine we can continue sampling,

- \bullet top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the more certain you are about your biased estimates (under the any model!!)







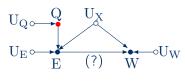
Yo, what is going on??

based on DAG and statistical model,

- the 2nd D-separation rule requires control on any noncollider to block the backdoor path,
 i.e. E ⊥⊥W | U_X
 (but U_X is unobservable)
- if we use Q in the model, the 3rd D-separation rule kicks in:
 "A collider that has been conditioned on does not block a path."
 i.e. Q ⊥ U_X | E
 (e.g. switch, electricity, and light bulb)

$$M = \begin{cases} Q \leftarrow f_Q(U_Q) \\ E \leftarrow f_E(Q, U_X, U_E) \\ W \leftarrow f_W(E, U_X, U_W) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



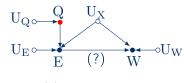
Yo, what is going on??

open paths?:

- \blacksquare E \rightarrow W
- \blacksquare E \rightarrow U_x \rightarrow W
- $\blacksquare \ E \to U_x \to Q \to E \to W$
- \blacksquare E \to U_x \to Q \to E \to U_X \to W

$$M = \begin{cases} Q \leftarrow f_Q(U_Q) \\ E \leftarrow f_E(Q, U_X, U_E) \\ W \leftarrow f_W(E, U_X, U_W) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



What should I do then??

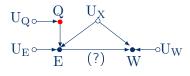
$$egin{aligned} \begin{pmatrix} \mathbf{W} \\ \mathbf{E} \end{pmatrix} &\sim \mathbf{MVN} \left[\begin{pmatrix} \mu_{\mathbf{W}} \\ \mu_{\mathbf{E}} \end{pmatrix}, \mathbf{\Sigma} \right] \\ \mu_{\mathbf{W}} &= \alpha_{\mathbf{W}} + \beta_{\mathbf{EW}} \mathbf{E} \\ \mu_{\mathbf{E}} &= \alpha_{\mathbf{E}} + \beta_{\mathbf{QE}} \mathbf{Q} \end{aligned}$$
(c) probabilistic model

based on DAG and statistical model, use the knowledge of the system

- \blacksquare one model for $Q \to E$
- \blacksquare one model for $E \to W$
- use the knowledge that cov(E, W) > 0 due to unobserved confounder U_X, (i.e. cov(E, W) = Σ = SRS)

$$M = \begin{cases} Q \leftarrow f_Q(U_Q) \\ E \leftarrow f_E(Q, U_X, U_E) \\ W \leftarrow f_W(E, U_X, U_W) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



did it worked???

based on DAG and bayesian statistical analysis,

- appropriate value estimated, (assuming our DAG is true)
- it picks up some of the unobserved correlation R[1, 2]

```
mean
        0.02 \ 0.18 \ -0.26
       -0.14 0.16 -0.40
        1.00 0.07
boe
bEW
        0.05 0.07 -0.06
R[1,1]
                   1.00
R[1,2]
                    0.15
                          0.50
R[2,1]
                    0.15
                          0.50
        1.25 0.10
                    1.11
        1.39 0.10
                   1.24
```



did it worked???

frequentists guys apply Two Stage Least Squares (2SLS)^a:

- \blacksquare regress $E \leftarrow Q$,
- predict Ê,
- \blacksquare regress W \leftarrow \hat{E}

```
s1 = lm( E ~ Q, data=d)
Ehat = s1$fitted.values
s2 = lm( W ~ Ehat, data=d)
# se not corrected
require(AER)
tsls = ivreg( W ~ E | Q, data=d)
# se corrected
```

```
summary(s2)
lm(formula = W \sim Ehat. data = d)
Residuals:
             10 Median
-4.1911 -1.0670 -0.0643 1.3802 4.6813
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.5992
              0.2093
                        0.1263 1.658
Fhat
                                           0.101
> summary(tsls)
ivreg(formula = W \sim E \mid Q, data = d)
Residuals:
             10 Median
-3.6571 -1.1852 -0.1819 1.0945 4.5328
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
                         0.3325 -1.802
(Intercept) -0.5992
              0.2093
                         0.1145
                                 1.829
                                          0.0705
```

^aHanck et al. [7], section 12.1, See McElreath [12] chapter 14 (p. 460) for a discussion on the method.

So, what is going on? pipefork1_triptych.pdf

Similar case, contextual confounds¹⁴ research question,

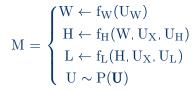
- Does W has an effect on L?
- should we include H in our model?

variables,

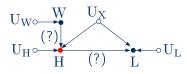
- W, win the lottery
- H, happiness
- U_X, contextual confound
- L, lifespan

Short answer;

- for total effects: No (two question marks together)
- for direct effect of $H \to L$: will be always counfounded



(a) structural model





¹⁴McElreath [13], lecture 6

3. Example cases

Collider bias: M-bias



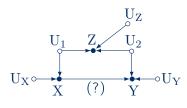
M-bias¹⁵

case of,

- bias on pre-treatment variable research question,
- Should we include Z in our model? variables,
 - Z, "health" quality of friends (defined as a continuum)
 - X, health of individual 1
 - \blacksquare U₁, hobbies of individual 1
 - Y, health of individual 2
 - \blacksquare U₂, hobbies of individual 2

$$M = \begin{cases} X \leftarrow f_X(U_1, U_X) \\ Z \leftarrow f_Z(U_1, U_2, U_Z) \\ Y \leftarrow f_Y(X, U_2, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model



¹⁵McElreath [13], lecture 6; Cinelli et al. [4] (p. 4)



Simulation setting

```
# sim
U1 = sample(1:5, 100 , replace=T)
U2 = sample(1:5, 100 , replace=T)
Z = rnorm( 100 , 0.5*U1 + 0.5*U2)
X = rnorm( 100 , 1*U1 + 0*Z)
Y = rnorm( 100 , 1*U2 + 0*X + 0*Z)
d = data.frame(U1,U2,Z,X,Y)
```

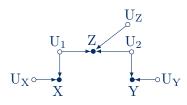
(c) R code

Implications,

- X ⊥1Y
- X #Y | Z

$$M = \begin{cases} X \leftarrow f_X(U_1, U_X) \\ Z \leftarrow f_Z(U_1, U_2, U_Z) \\ Y \leftarrow f_Y(X, U_2, U_Y) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

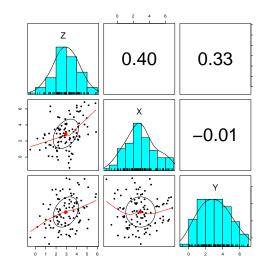
(a) structural model



Descriptive analysis

based on descriptive analysis,

- \blacksquare Cor(X, Y) \approx 0, quite low
- larger Cor(Z, Y), while Cor(Z, X) is not high enough to discard it as a cause of multicollinearity,
- we might include Z rather than X (but the effect of X is our interest!!)
- then we include Z and X





Again regression!!

based on statistical analysis,

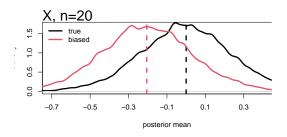
- X does not have an effect of Y, (in the first nor the second model)
- but X has an (non-negligible)
 effect when Z is in the model
 (but we do not reject the null)
- The increase of the X effect might lead you to think that with more data, we can reject the null (and you would be right!!)
- But is it correct to include Z?

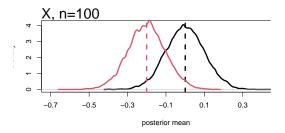
```
summary(lm(Y ~ X, data=d)) # unbiased effects (effi
lm(formula = Y \sim X. data = d)
Residuals:
             10 Median
-3.7452 -1.3300 -0.0253 1.2414 3.9479
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.067474
> summary(lm(Y \sim X + Z, data=d)) # biased effects (e
lm(formula = Y \sim X + Z, data = d)
Residuals:
    Min
             10 Median
                                      Max
 -3.3756 -1.0898 0.0351 1.0540
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.83324
             -0.15765
             0.56590
                         0.14974
                                    3 779 0 000272
```

Ok, I get it!!, more data, more wrong!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the "incorrect" model, the larger the sample size,
 - the more certain you are about your biased estimates (with enough you could reject the null)







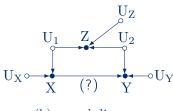
The dream team strikes back!!

based on DAG and statistical model,

- the 3rd D-separation states that a collider that has been conditioned on does not block a path, in this case: $X \to U_1 \to Z \to U_2 \to Y$ i.e. $X \not\perp \!\!\! \perp \!\!\! \perp Y \mid Z$
- therefore if we want to find the direct effect of $X \to Y$, we should not stratify by Z

$$\mathbf{M} = \begin{cases} \mathbf{X} \leftarrow f_{\mathbf{X}}(\mathbf{U}_1, \mathbf{U}_{\mathbf{X}}) \\ \mathbf{Z} \leftarrow f_{\mathbf{Z}}(\mathbf{U}_1, \mathbf{U}_2, \mathbf{U}_{\mathbf{Z}}) \\ \mathbf{Y} \leftarrow f_{\mathbf{Y}}(\mathbf{X}, \mathbf{U}_2, \mathbf{U}_{\mathbf{Y}}) \\ \mathbf{U} \sim P(\mathbf{U}) \end{cases}$$

(a) structural model





The dream team strikes back!!

based on DAG and statistical analysis,

■ the model that answers our research question is the first one, (assuming our DAG is true)

So, what is going on? $\operatorname{collider4_triptych.pdf}$

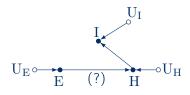


Case control¹⁶

also,

- an instance of descendant bias research question,
 - Does E has a (direct) effect on H?
- Should we include I on our model? variables,
- E, education
 - H, hours in occupation (standardized)
 - I, income

(a) structural model



¹⁶McElreath [13], lecture 06; Cinelli et al. [4] (p. 8)



 $M = \begin{cases} E \leftarrow f_E(U_E) \\ H \leftarrow f_H(E, U_H) \\ I \leftarrow f_I(H, U_I) \\ U \sim P(\textbf{U}) \end{cases}$

Simulation setting

```
# sim
E = rnorm( 100 )
H = rnorm( 100 , -1*E )
I = rnorm( 100 , -1*H )
d = data.frame(E,H,I)
```

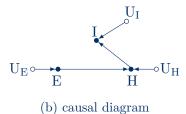
(c) R code

Implications,

- E #H
- E #I | H

$$M = \begin{cases} E \leftarrow f_E(U_E) \\ H \leftarrow f_H(E, U_H) \\ I \leftarrow f_I(H, U_I) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

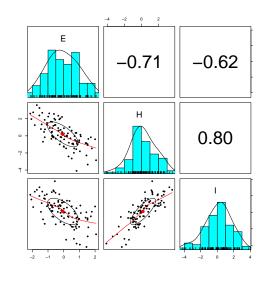
(a) structural model



"Eyeballing" analysis

based on correlation analysis,

- cor(E, I) < 0 does NOT goes in line of our "rudimentary" understanding of the data.
- while cor(H, I) > 0 indicate the more you work the more you gain (but is it the only way?)
- since cor(H, I) is high we might include it as a covariate in our statistical model (to improve the precision?)





Regression, regression!!

based on statistical analysis,

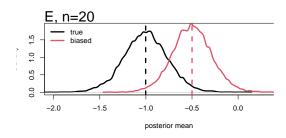
- we now have two models with two different "levels" of effects
- which one is the "truth"?

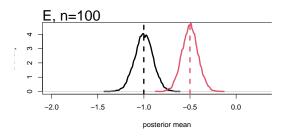
```
summarv(lm(H \sim E. data=d)) # unbiased effects
Call:
lm(formula = H \sim E. data = d)
Residuals:
               10 Median
 -2.19319 -0.60621 -0.06694 0.55674 2.77776
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.04689
                        0.09534 0.492
            -0.90223
                        0.08930 -10.104
 summary(lm(H ~ E + I. data=d)) # biased effects
Call.
lm(formula = H \sim E + I, data = d)
Residuals:
                    Median
                                         Max
-1 22991 -0 54882
                   0.01153 0.46886 1.79879
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.05036
            -0.45847
                        0.07915
                                 -5.793 8.57e-08
             0.55017
                        0.05728
                                  9.604 9.61e-16
```

The data tell-tell story!!

imagine we can continue sampling,

- top: 10,000 samples n = 20
- bottom: 10,000 samples n = 100 under the incorrect model, the larger the sample size,
 - the more certain you are about your biased estimates







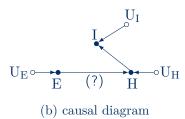
The dream team!!

based on DAG and statistical model,

- stratifying (controlling) on I, we are already explaining variability in H
 (I is a descendant of H)
- the remaining is explained by E (big chunk is already explained)

$$M = \begin{cases} E \leftarrow f_E(U_E) \\ H \leftarrow f_H(E, U_H) \\ I \leftarrow f_I(H, U_I) \\ U \sim P(\boldsymbol{U}) \end{cases}$$

(a) structural model





the dream team!!

based on DAG and statistical analysis,

■ the less biased model is the first, (assuming our DAG is true)



So, what is going on? $descendant 1_trip tych.pdf$

4. Concluding remarks



Concluding remarks

- Research is filled with challenges (you: duh!!)
- Statistical models are not theory (you: so obvious again!!)
- Don't trust your statistical model when no DAG is involved (me: how about that?!)
- For explanation, without a DAG the (sample) size does not matter (me: sorry?!)
- For prediction, sometimes a DAG can help (me: did you expect this one?!)



Concluding remarks

Additionally, simulation can serve,

- a reflection of a hypothesis, and its research complexities (me: DAGs baby!!)
- a place where you can reflect the status of a population (test what happens if??, post-stratification??)
- as a data where you can test your statistical model on various purposes
 (parameter recovery, power?)

think2.jpg



5. Do you wanna know more???



5. Do you wanna know more????



- [1] Anderson, D. [2008]. Model Based Inference in the Life Sciences: A Primer on Evidence, Springer.
- [2] Bareinboim, E. and Pearl, J. [2016]. Causal inference and the data-fusion problem, Proceedings of the National Academy of Sciences 113(27): 7345–7352. doi: https://doi.org/10.1073/pnas.1510507113.
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