

# Objective rating method: Entropy

Speech intelligibility estimation

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# What are we going to talk about?

- 1 Preliminars
  - Research question
  - Research hypothesis production
- 2 Research hypothesis procedure
  - Estimand and process model
  - Synthetic data generation
  - Statistical model design and testing
  - Apply statistical model to data
- 3 References



#### 1. Preliminars

Research question



# Research question

#### On two fronts:

1. Can comparative judgement (CJ) methods be used to assess speech intelligibility (SI)?,

To investigate this wee need:

- an objective measure of SI
- 2. where CJ stands versus absolute holistic judgement (HJ) methods?, In terms of:
  - validity
  - $\blacksquare$  reliability
  - statistical efficiency
  - time efficiency

# Objective measure of SI

the most objective (we know of) measure of SI comes from a transcription task:

- 1. transcribe children's utterances (made by multiple judges),
- 2. align transcriptions at the utterance level,
- 3. calculate an entropy measure (H), per utterance per child:

$$H = H(\mathbf{p}) = \frac{-\sum_{i=1}^{n} p_i \cdot \log_2(p_i)}{\log_2(N)}$$

- 4. characteristics of H [3, 7]
  - $\blacksquare$  bounded in [0,1] space,
  - $\blacksquare$  utterances with more agreement are more intelligible, and therefore H  $\rightarrow$  0,
  - $\blacksquare$  utterances with low agreement are less intelligible, and therefore H  $\rightarrow$  1.

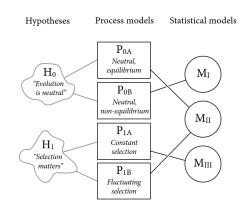
### 1. Preliminars



# A typical scientific lab

#### What is needed / deal with<sup>a</sup>

- 1. Quality of theory
- 2. Quality of data
- 3. Reliable procedures and code
- 4. Quality of data analysis
- 5. Documentation
- 6. Reporting



<sup>&</sup>lt;sup>a</sup>Figure extracted from McElreath [11]

# Research hypothesis production<sup>1</sup>

#### Well known challenges

- Insufficient data
- Wrong population
- Measurement error
- Selection bias
- Confounding

#### Known challenges in our research;

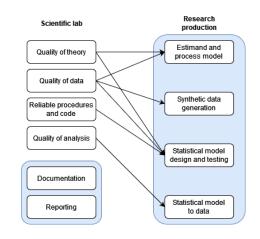
- Insufficient data (possibly)
- Wrong population
- Measurement error
- Selection bias
- Confounding

<sup>&</sup>lt;sup>1</sup>Hernán [9], lesson 4



# Research hypothesis schematics<sup>2</sup>

- a. Estimand and process model
- b. Synthetic data generation
- c. Statistical model design and testing
- d. Apply statistical model to data



<sup>&</sup>lt;sup>2</sup>McElreath [12], lecture 20, Pearl [13]. Follow Fogarty et al. [8] on item (c).



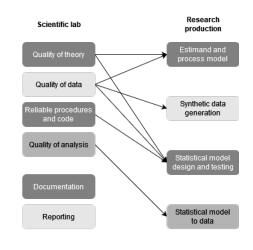
# Why do we need to follow this?

#### Because the improvement of:

- A clear definition of the estimand and process model (assumptions).
- An improved the reliability of your procedures.
- As a documentation procedure.

#### leads to:

- A sound analysis and results (even when we cannot answer our question).
- An improved planning to get data.





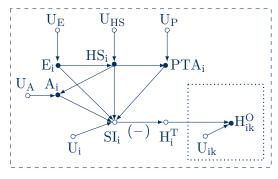
2. Research hypothesis procedure

Estimand and process mode



# The theory behind our research

- $\blacksquare$  H<sub>ik</sub> = (observed) entropy replicates
- $\blacksquare$  H<sub>i</sub> = (latent) "true" entropy
- $SI_i = (latent) SI score$ (inversely related to  $H_i^T$ )
- $\blacksquare$  A<sub>i</sub> = "hearing" age (minimum)
- $\blacksquare$   $E_i = etiology of disease$
- $\blacksquare$  HS<sub>i</sub> = hearing status
- PTA<sub>i</sub> = pure tone average (standardized)
- variables assumed independent, beyond the described relationships,



General structural diagram

$$P(U) = P(U_{ik}, U_i, U_A, U_E, U_{HS}, U_P)$$
  
=  $P(U_{ik})P(U_i)P(U_A)P(U_E)P(U_{HS})P(U_P)$ 

$$\begin{split} & H_{ik}^O \leftarrow f(H_i^T, U_{ik}) \\ & H_i^T \leftarrow f(SI_i) \\ & SI_i \leftarrow f(HS_i, E_i, A_i, PTA_i, U_i) \\ \\ & HS_i \leftarrow f(U_{HS}) \\ & A_i \leftarrow f(U_A) \\ & E_i \leftarrow f(U_E) \\ \\ & PTA_i \leftarrow f(U_P) \\ & U \sim P(\textbf{U}) \\ & (a) \text{ general structural model} \end{split}$$

$$\begin{split} H_{ik}^{O} &\sim \, \mathrm{BetaProp}(H_{i}^{\mathrm{T}}, M_{ik}) \\ H_{i}^{\mathrm{T}} &= \, \mathrm{inv\_logit}(-\mathrm{SI}_{i}) \\ \mathrm{SI}_{i} &\sim \, \mathrm{Normal}(\mu_{\mathrm{SI}}, \sigma_{\mathrm{Ui}} +) \\ \mu_{\mathrm{SI}} &= \alpha + \alpha_{\mathrm{E[i]}, \mathrm{HS[i]}} + \beta_{\mathrm{P}} \mathrm{PTA}_{i} \\ &+ \beta_{\mathrm{A}, \mathrm{HS[i]}}(A_{i} - \bar{A}) \\ \mathrm{HS}_{i} &\sim \, \mathrm{data} \\ A_{i} &\sim \, \mathrm{data} \\ E_{i} &\sim \, \mathrm{data} \\ \mathrm{PTA}_{i} &\sim \, \mathrm{data} \\ \mathrm{U} &\sim \, \mathrm{unobservable} \\ \mathrm{(a) \, general \, probabilistic \, model} \end{split}$$

First form

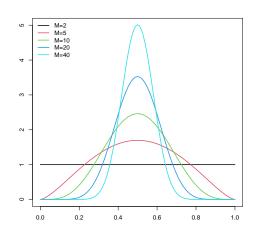
#### Notice,

- $\alpha$ ,  $\alpha_{\text{HS[i]}}$ ,  $\alpha_{\text{E[i]}}$ ,  $\beta_{\text{A,HS[i]}}$ ,  $\beta_{\text{P}}$  are structural parameters (as in SEM)
- $U_{ik}$  = replicates measurement error  $U_i$  = between child SI variability
- variability of  $U_{ik}$  is modeled by  $M_{ik}$  variability of  $U_i$  is modeled by  $\sigma_{Ui}$

```
H_{ik}^{O} \sim BetaProp(H_{i}^{T}, M_{ik})
   H_i^T = inv logit(-SI_i)
    SI_i \sim Normal(\mu_{SI}, \sigma_{IIi})
        \mu_{\rm SI} = \alpha + \alpha_{\rm E[i].HS[i]} + \beta_{\rm P}PTA_{\rm i}
            +\beta_{A \text{ HS[i]}}(A_i - \bar{A})
  HS_i \sim data
    A_i \sim data
    E_i \sim data
PTA_i \sim data
     U \sim unobservable
      (a) general probabilistic model
```

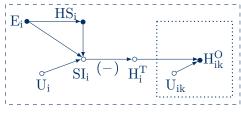
#### Express variability in BetaProp

$$\begin{split} & \boldsymbol{H}_{ik}^{O} \sim \ \operatorname{BetaProp}(\boldsymbol{H}_{i}^{T}, \boldsymbol{M}_{ik}) \\ & \boldsymbol{H}_{i}^{T} = \ \alpha/(\alpha + \beta) \\ & \boldsymbol{M}_{ik} = \ \alpha + \beta \\ \\ & \alpha = \ \boldsymbol{H}_{i}^{T} \cdot \boldsymbol{M}_{ik} \\ & \beta = \ (1 - \boldsymbol{H}_{i}^{T}) \cdot \boldsymbol{M}_{ik} \\ \\ & \alpha = \ 0.5 \cdot 2 \\ & \beta = \ (1 - 0.5) \cdot 2 \end{split}$$

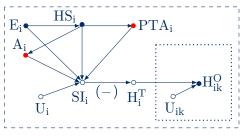


#### Interested in two effects

- 1. total effects model inherits:
  - children's characteristics that lead to the fitting of specific apparatus,
  - the (convenience of) sample selection (fixed with post-stratification)
- 2. to do the last, we stratify for all variables that explain variability, ergo, use a direct effects model
- 3. two levels: replicates (k), children (i), denoted by discontinuous squares



(b) total effects



$$\begin{split} & H_{ik}^O \leftarrow f(H_i^T, U_{ik}) \\ & H_i^T \leftarrow f(SI_i) \\ & SI_i \leftarrow f(HS_i, E_i, \textbf{A}_i, \textbf{PTA}_i, U_i) \end{split}$$

$$& HS_i \leftarrow f(U_{HS}) \\ & A_i \leftarrow f(U_A) \\ & E_i \leftarrow f(U_E) \\ & \textbf{PTA}_i \leftarrow f(U_P) \\ & U \sim P(\textbf{U}) \end{split}$$
(a) general structural model

$$\begin{split} H_{ik}^{O} &\sim \, \mathrm{BetaProp}(H_{i}^{T}, M_{ik}) \\ H_{i}^{T} &= \, \mathrm{inv\_logit}(-\mathrm{SI}_{i}) \\ \mathrm{SI}_{i} &= \, \mathrm{a}_{i} + \alpha + \alpha_{\mathrm{E[i],HS[i]}} + \beta_{\mathrm{P}} \mathrm{PTA}_{i} \\ &\quad + \beta_{\mathrm{A,HS[i]}} (A_{i} - \bar{A}) \\ \mathrm{a}_{i} &\sim \, \mathrm{Normal}(0, \sigma_{\mathrm{Ui}}) \\ \mathrm{HS}_{i} &\sim \, \mathrm{data} \\ A_{i} &\sim \, \mathrm{data} \\ \mathrm{E}_{i} &\sim \, \mathrm{data} \\ \mathrm{PTA}_{i} &\sim \, \mathrm{data} \\ \mathrm{U} &\sim \, \mathrm{unobservable} \end{split}$$

## 2. Research hypothesis procedure

Synthetic data generation



## Idealized data<sup>3</sup>

Simulation data can serve as [10, 11],

- 1. A place where to test your model, on multiple purposes,
  - parameter recovery
  - power
- 2. A reflection of a population,
  - De Raeve [6]: 70 HI/CI, 130 HI/HA
  - Our idealized data:150 NH, 70 HI/CI, 130 HI/HA
- 3. A reflection of a hypothesis,
  - size of effects

<sup>&</sup>lt;sup>3</sup>more details in file: 1 2 E sim fun.R



<sup>(</sup>sim\_name=NULL, # file\_name need to include sim\_save=NULL, # file\_save need to include seed=NULL, # seed I=350, # experimental units (children) K=10, # replicates (utterances) p=c(0.50, 0.175, 0.325), # children prop. par=list( m\_i=0, s\_i=0.5, # hyperprior ch m\_M=10, s\_M=NULL, # generation a=0. bP = -0.1.aHS=c(0.4,0,-0.4), bA=0.15. bAHS=rep(0,3), aE=rep(0,4), aEHS=matrix( c(rep(0,4), #NH)c(0, seq(0.1, -0)c(0, seq(0.1, -0))ncol=3, byrow=F)

# About the size of the effects (in logits, no previous info)

- 1. E has a full interaction with HS
- 2.  $aHS[j] aHS[i] \approx -0.4$ , NH vs HI/CI (depends on E)
- 3. bP = -0.1, per PTA unit  $(+10 \text{ PTA units} \Rightarrow -1 \text{ logit})$ ,
- 4.  $bA \approx 0.15$ , per A unit, above the minimum (depends on HS)  $(+10 \text{ A units} \Rightarrow +1.5 \text{ logits})$

```
(sim_name=NULL, # file_name need to include
sim_save=NULL, # file_save need to include
seed=NULL, # seed
I=350, # experimental units (children)
p=c(0.50, 0.175, 0.325), # children prop.
par=list( m_i=0, s_i=0.5, # hyperprior ch
          m_M=10, s_M=NULL, # generation
          a=0.
          bP = -0.1.
          aHS=c(0.4,0,-0.4),
          bA=0.15.
          bAHS=rep(0,3),
          aE=rep(0,4),
          aEHS=matrix( c(rep(0,4), #NH)
                           c(0, seq(0.1, -0))
                           c(0, seq(0.1, -0))
                        ncol=3, byrow=F)
```

- 1. variables are generated in a random fashion
- 2. between children SI variability are defined by the random effects

```
# 1. true data ####
dT = data.frame(matrix(NA, nrow=I, ncol=1))
names(dT) = c('child_id')
dTschild id = 1:I
n = round(p*I)
if( sum(n) != I ){
   n[3] = I - sum(n[c(1,3)]) # to sum the right amount
if(!is.null(seed)){
 set.seed(seed+1)
dT$HS = c(rep(1, n[1]), rep(2, n[2]), rep(3, n[3]))
dT$A = round(rnorm(sum(n), 5, 1))
dTSA = with(dT. ifelse(A>7. 7. A))
dTSE = c(rep(1, n[1]), # no way to know true effects
          sample(2:3, size=n[2], replace=T),
         sample(3:4, size=n[3], replace=T))
dTPTA = c(round(rnorm(n[1], 60, 15)), # first 12 NH
           round(rnorm(n[2], 90, 15)), # next 10
            round(rnorm(n[3], 110, 15))) # last 10
if(!is.null(seed)){
 set.seed(seed-1)
par$re_i = rnorm(I, par$m_i, par$s_i)
dT$re i = par$re i # children's random effects (between SI
```

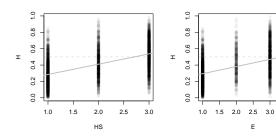
- 1. we use second form of the probabilistic model
- 2. "true' entropy (Ht) is inversely related to SI
- 3. we simulate measurement error through M from BetaProp() distribution.

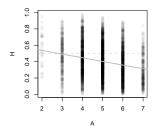
```
dTSSI = NA
A bar = min(dT$A)
sPTA = standardize( dT$PTA )
for(i in 1:I){
 dT$SI[i] = with(dT, par$re_i[i] +
                    par$a +
                    par$bP*sPTA[i] +
                    par$aHS[ HS[i] ] +
                    par$aE[ E[i] ] +
                    par$aEHS[ E[i], HS[i] ] +
                    par$bA*( A[i] - A_bar ) +
                    par$bAHS[ HS[i] ]*( A[i] - A bar ) )
dT$Ht = inv_logit(-dT$SI) # true entropy (SI -> Ht: negat
# variability of H
if(!is.null(seed)){
 set.seed(seed+2)
if( is.numeric(par$m_M) & !is.numeric(par$s_M) ){
 par$M = rep(par$m_M, I)
} else{
 par$M = round( rlnorm(I, meanlog=par$m_M, sdlog= par$s_
dT$M = par$M # same df for all children (not same shape!
dT[,6:ncol(dT)] = round(dT[,6:ncol(dT)], 5)
```

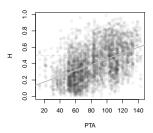
- 1. we simulate replicate measures of entropy (H)
- 2. we storage all relevant parameters and data

```
# 2. observed data ####
N = I*K
d0 = data.frame(matrix(NA, nrow=N, ncol=3))
names(d0) = c('child_id', 'utt_id', 'H')
dO$child_id = rep(1:I, each=K)
d0$utt_id = rep(1:K, I)
# generating observed H
if(!is.null(seed)){
  set.seed(seed-2)
for(i in 1:I){
  # identify data
  idx = d0$child_id == i
  # linear predictor
  dO$H[idx] = rbeta2(n=K, prob=dT$Ht[i], theta=dT$M[
dO$H = round(dO$H, 5)
```

# Example







3.5 4.0



2. Research hypothesis procedure

Statistical model design and testing



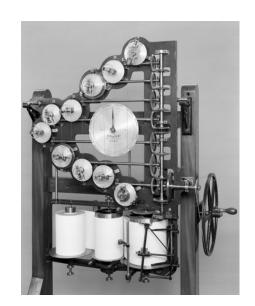
# Model design and test<sup>4</sup>

#### Purpose:

- to have reliable procedures,
- to maintain a clear documentation,
- to have a sound analysis

#### Procedure:

- step by step, instantiating one difficulty at the time
- use probabilistic assumptions defined in estimand and process model
- is like running synthetic data generation backwards



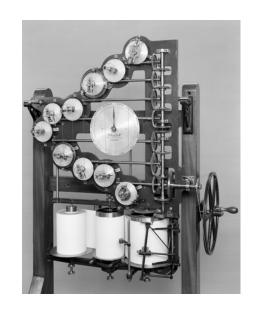


<sup>&</sup>lt;sup>4</sup>Following Fogarty et al. [8]

# Model design and test

#### We evaluate,

- the probabilistic model implementation [11, 2] (centered and non-centered versions)
- prior predictive
- "health" of MCMC chains
- parameter recovery
- posterior predictive
- power



#### starts with,

- the simplest model
- the simplest data generating procedure

```
transformed parameters{
    vector[I] SI;
                            // SI index (per child)
    vector[I] Ht;
                            // true entropy (per chi
    SI = a + re_i;  // linear predictor
Ht = inv_logit(-SI);  // average entropy (SI -
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    // priors
    a \sim normal(0, 0.2);
    re_i \sim normal(m_i . s_i):
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , 10 );
```

#### try,

• the centered and non-centered parametrization

```
transformed parameters{
    vector[I] re_i;
                           // random intercepts (per
    vector[I] SI; // SI index
    vector[I] Ht:
                          // true entropy (per child
    re_i = m_i + s_i*z_re;// non-centered RE
   SI = a + re_i;  // linear predictor
Ht = inv_logit(-SI);  // average entropy (SI ->
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    s_i \sim exponential(1):
    // priors
    z_re ~ std_normal();
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , 10 );
```

#### escalate,

- complexity of model
- traits of the data

```
transformed parameters{
    vector[I] SI:
                          // SI index (per child)
    vector[I] Ht;
                          // true entropy (per chi
    // linear predictor
    for(i in 1:I){
      SI[i] = re_i[i] + a + aHS[HS[i]] +
        bA*Am[i] + bP*sPTA[i]:
      // no multicollinearity between E and HS
    // average entropy (SI -> Ht: negative)
    Ht = inv logit(-SI): // average entropy (SI -
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    s_i \sim exponential(1):
    // priors
    a \sim normal(0, 0.2);
    re_i \sim normal( m_i , s_i );
    aHS \sim normal(0, 0.5);
    bP \sim normal(0, 0.3);
    bA \sim normal(0, 0.3);
    m_M \sim lognormal(1.5, 0.5);
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , m_M );
```

We tested 5 random effects models: (from 5 synthetic data types) (centered, and non-centered)

- only intercept, M = 10,
- $\blacksquare$  causal model, M = 10,
- causal model, M per individual,
- no known process,
- causal model with interactions,
   M per individual,

```
transformed parameters{
   vector[I] SI:
                          // SI index (per child)
   vector[I] Ht;
                          // true entropy (per child
   // linear predictor
    for(i in 1:I){
      SI[i] = re_i[i] + a + aHS[HS[i]] +
        bA*Am[i] + bP*sPTA[i]:
      // no multicollinearity between E and HS
   // average entropy (SI -> Ht: negative)
   Ht = inv_logit(-SI): // average entropy (SI ->
model{
   // hyperpriors
   m_i \sim normal(0.0.2):
   s_i \sim exponential(1);
   // priors
    a \sim normal(0, 0.2);
    re_i \sim normal(m_i . s_i):
    //aE \sim normal(0.0.5):
   aHS \sim normal(0, 0.5);
   bP \sim normal(0, 0.3);
   bA \sim normal(0, 0.3);
   m M ~ lognormal( 1.5 . 0.5 ):
   // likelihood
   for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , m_M );
```

# Prior predictive

#### Priors and hyper-priors

- In the probabilistic (causal) model there were no priors for our parameters,
- To decide our priors we follow McElreath [11]: "priors are part of the assumptions, and should be inspected as such",
- We will evaluate the implications of our priors on the outcome scale.
   We have three outcomes scales: SI<sub>i</sub>, H<sub>i</sub><sup>T</sup>, and H<sub>ik</sub><sup>O</sup>

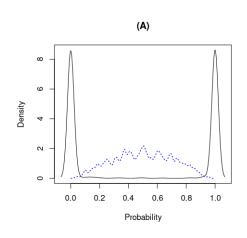
```
Priors
                  a_i \sim Normal(\mu_a, \sigma_a)
                M_i \sim LogNormal(\mu_M, \sigma_M)
                  \alpha \sim \text{Normal}(0.0.2)
           \alpha_{\rm HS[i]} \sim {\rm Normal}(0, 0.3)
             \alpha_{\rm E[i]} \sim {\rm Normal}(0, 0.3)
        \beta_{\text{A.HS[i]}} \sim \text{Normal}(0, 0.3)
                 \beta_{\rm P} \sim \text{Normal}(0, 0.3)
Hyper-priors
                 \mu_a \sim \text{Normal}(0, 0.2)
                 \sigma_a \sim \text{Exp}(1)
                \mu_{\rm M} \sim {\rm Normal}(0, 0.5)
```

 $\sigma_{\rm M} \sim {\rm Exp}(1)$ 

# Prior predictive

Undesired assumptions can easily creep in non-linear models<sup>a</sup> Example:

(black line) 
$$\theta \sim N(0, 100)$$
 
$$logit(p) = \theta$$
 (blue line) 
$$\theta \sim N(0, 1)$$
 
$$logit(p) = \theta$$



<sup>&</sup>lt;sup>a</sup>Figure extracted from Rivera [14].

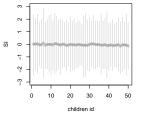
# Prior predictive

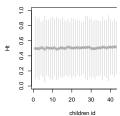
What our priors imply?

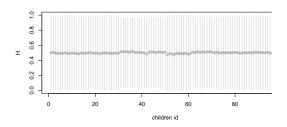
NO undesired assumption has crept in:

- the SI<sub>i</sub> scale,
- the H<sub>i</sub><sup>T</sup> scale,
- $\blacksquare$  the  $H_{ik}^{O}$  scale

i.e. the scales' full space can be reached by (a combination of) the parameters







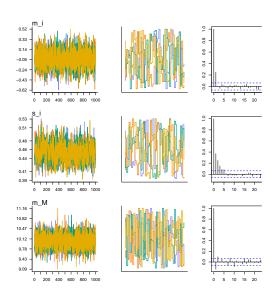


#### "Health" of MCMC chains

The MCMC chains achieve,

- good convergence
- good mixing
- lack of autocorrelation

same results on the n\_eff and RHat statistics.



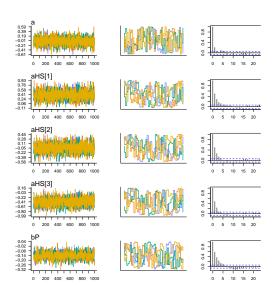


#### "Health" of MCMC chains

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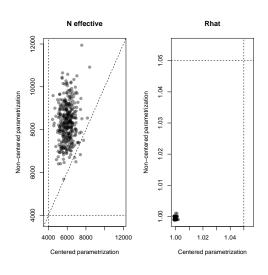




### "Health" of MCMC chains

The non-centered parametrization has,

- better n\_eff (denoting lack of autocorrelation)
- better Rhat (denoting good convergence)
- better mixing (inspected visually, not shown)



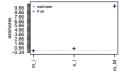


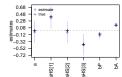
# Parameter recovery

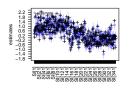
#### On idealized data

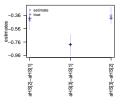
# The model, with less working assumptions<sup>a</sup>

- recovers the parameters in the right scale,
- most of the "true" parameters are inside of the compatibility intervals (CI)
- contrast are approximately correct









 $<sup>^{\</sup>rm a}$ random effects causal model, with M = 10

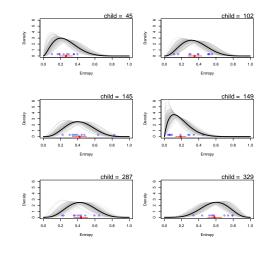
# Posterior predictive

But how well reproduces the data?, with less working assumptions<sup>a</sup>

- captures the variability of the replicates,
- provides a "true" H and SI

More complex models,

- captures even better the data, but might overfit,
- avoid overfit when the model is fit to the data (ITA [1, 4])





 $<sup>^{\</sup>rm a}$ random effects causal model, with M = 10

Equivalent prior sampling method [18, 10]

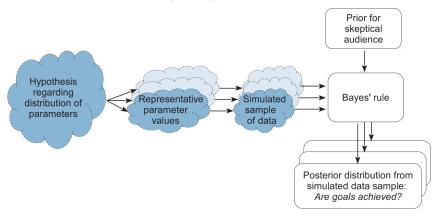
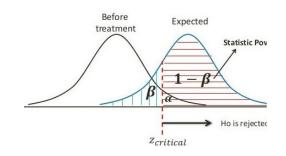


Figure: Flow of information in a power analysis, in which simulated data come from random hypothetical parameters. Extracted from Kruschke [10].

#### Equivalent prior sampling method

- 1. generate idealized data
- 2. generate parameters' distribution (with Bayes theorem)
- 3. simulate data sample (variability in data and parameters)
- 4. apply models to simulated data (include priors)
- 5. evaluate desire goals
  - lacktriangledown reject the null hypothesis
  - affirm predicted value
  - $\blacksquare$  achieve precision in estimate
- 6. repeat procedure (to approximate power)





#### Equivalent prior sampling method

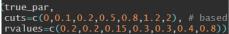
- 1. generate idealized data
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Epower( power\_save=file.palth(getwd(), 'sim\_chain'), # po
 sim\_name='Hbeta\_sim2\_power.RData', # file\_save n
 sim\_save=file.path(getwd(), 'sim\_data'), # file\_
 model\_name='Hbeta\_NC\_sim2', # model for which we
 model\_in=file.path(getwd(), 'sim\_nodels'), # loc
 model\_out=file.path(getwd(), 'sim\_chain'), # loc
 Nsim=100, # number of simulation for power
 I\_grid = c(48, 60), # experimental units (childr
 K\_grid = c(10, 20), # replicates (utterances)
 p=c(0.34, 0.33, 0.33), # children prop. on each
 par\_int=c('aHS', 'bP', 'bA', 'm\_i', 's.i', "m\_M', 'SI'
 par\_cont=c('aHS', 'SI')) # parameters to contras



#### Equivalent prior sampling method

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2,0.2,0.13,0.3,		
	Effect size	d
	Very small	0.01
	Small	0.20
	Medium	0.50
	Large	0.80
	Very large	1.20
	Huge	2.0



<sup>&</sup>lt;sup>a</sup>use ROPE [10], and effect sizes [5, 15]

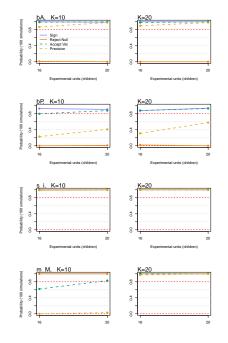
# Power results

#### For this parameter set,

- reject the null hypothesis, never achieved in structural parameters,
- affirm predicted value, in all shown parameters
- achieve precision in estimate, in some parameters is reached

#### Notice,

- between child SI variability,
- replicates measurement error (M),
- not much difference with 2x comparisons (except for M)



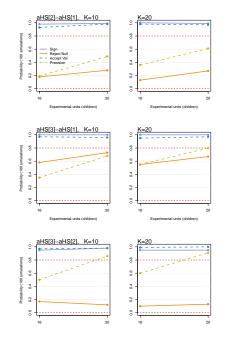
# Power results

#### For group contrasts,

- reject the null hypothesis,
   closest for the largest contrast,
- affirm predicted value, in all shown contrasts,
- achieve precision in estimate, grows with more children

#### Notice,

- not much difference with 2x comparisons
- group contrasts are easier to identify (larger sample size per comparison)



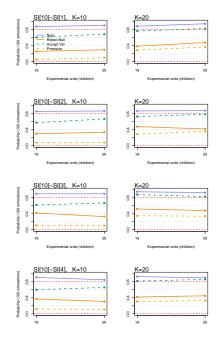
# Power results

#### For individual contrasts,

- reject the null hypothesis, depends on the contrast of interest,
- affirm predicted value, for some is achieved,
- achieve precision in estimate, clearly requires more sample size

#### Notice,

■ we see a clear difference with 2x comparisons (more sample at the appropriate level)



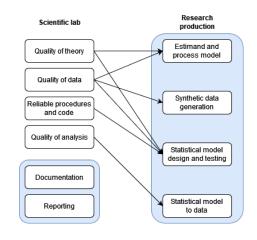
# 2. Research hypothesis procedure

Apply statistical model to data



# What we have so far

- 1. measure: replicated entropies  $H_{ik}^{O}$
- 2. estimand:
  SI index, structural parameters,
  contrasts
- 3. structural models: total and direct effects
- 4. probabilistic models: three possible fitting models
- 5. statistical models: works as intended
- 6. power: enough



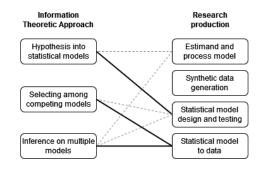
# Information Theoretic Approach [1, 4]

The last step would be select the most fitting model using the ITA,

- 1. hypothesis into statistical models,
- 2. select among competing models,
- 3. make inferences based on one or multiple models.

the most fitting model based on information criteria,

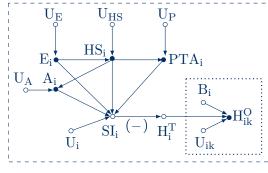
- WAIC [17]
- PSIS [16]





# Revisiting our theory

- $\blacksquare$  H<sub>ik</sub> = (observed) entropy replicates
- $\blacksquare$  H<sub>i</sub> = (latent) "true" entropy
- $SI_i = (latent) SI score$ (inversely related to  $H_i^T$ )
- $\blacksquare$  A<sub>i</sub> = "hearing" age (minimum)
- $\blacksquare$   $E_i = etiology of disease$
- $\blacksquare$  HS<sub>i</sub> = hearing status
- PTA<sub>i</sub> = pure tone average (standardized)
- $\blacksquare$   $B_i = block$
- variables assumed independent, beyond the described relationships,



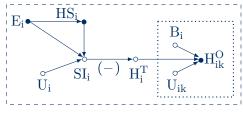
General structural diagram

$$\begin{split} P(\boldsymbol{U}) &= P(U_{ik}, U_i, U_A, U_E, U_{HS}, U_P, U_B) \\ &= P(U_{ik})P(U_i)P(U_A)P(U_E)P(U_{HS})P(U_P)P(U_B) \end{split}$$

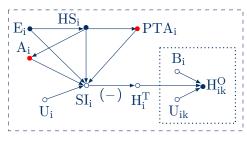


### Interested in two effects

- 1. total effects model inherits:
  - children's characteristics that lead to the fitting of specific apparatus,
  - the (convenience of) sample selection (fixed with post-stratification)
- 2. to do the last, we stratify for all variables that explain variability, ergo, use a direct effects model
- 3. two levels: replicates (k), children (i), denoted by discontinuous squares



(b) total effects



# Competing models

models\_table.png

3. References

# 3. References



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