

Objective rating method: Entropy

Speech intelligibility estimation

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What are we going to talk about?

- 1 Preliminars
 - Research question
 - Research hypothesis production
- 2 Research hypothesis procedure
 - Estimand and process model
 - Synthetic data generation
 - Statistical model design and testing
 - Apply statistical model to data
- 3 References



1. Preliminars

Research question



Research question

On two fronts:

1. Can comparative judgement (CJ) methods be used to assess speech intelligibility (SI)?,

To investigate this wee need:

- an objective measure of SI
- 2. where CJ stands versus absolute holistic judgement (HJ) methods?, In terms of:
 - validity
 - \blacksquare reliability
 - statistical efficiency
 - time efficiency



Objective measure of SI

the most objective (we know of) measure of SI comes from a transcription task:

- 1. transcribe children's utterances (made by multiple judges),
- 2. align transcriptions at the utterance level,
- 3. calculate an entropy measure (H), per utterance per child:

$$H = H(\mathbf{p}) = \frac{-\sum_{i=1}^{n} p_i \cdot \log_2(p_i)}{\log_2(N)}$$

- 4. characteristics of H [3, 7]
 - \blacksquare bounded in [0,1] space,
 - \blacksquare utterances with more agreement are more intelligible, and therefore $H \to 0$,
 - \blacksquare utterances with low agreement are less intelligible, and therefore H \rightarrow 1.

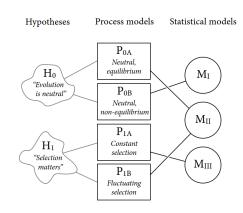
1. Preliminars



A typical scientific lab

What is needed / deal with^a

- 1. Quality of theory
- 2. Quality of data
- 3. Reliable procedures and code
- 4. Quality of data analysis
- 5. Documentation
- 6. Reporting



^aFigure extracted from McElreath [11]

Research hypothesis production¹

Well known challenges

- Insufficient data
- Wrong population
- Measurement error
- Selection bias
- Confounding

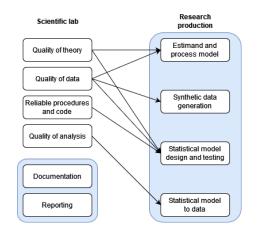
Known challenges in our research;

- Insufficient data (possibly)
- Wrong population
- Measurement error
- Selection bias
- Confounding

¹Hernán [9], lesson 4

Research hypothesis schematics²

- a. Estimand and process model
- b. Synthetic data generation
- c. Statistical model design and testing
- d. Apply statistical model to data



²McElreath [12], lecture 20, Pearl [13]. Follow Fogarty et al. [8] on item (c).



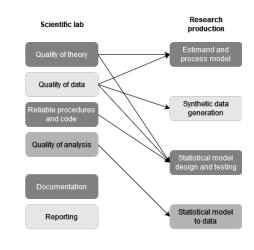
Why do we need to follow this?

Because the improvement of:

- A clear definition of the estimand and process model (assumptions).
- An improved the reliability of your procedures.
- As a documentation procedure.

leads to:

- A sound analysis and results (even when we cannot answer our question).
- An improved planning to get data.





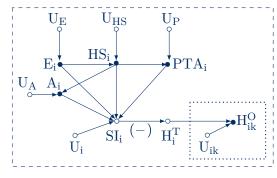
2. Research hypothesis procedure

Estimand and process mode



The theory behind our research

- \blacksquare H_{ik} = (observed) entropy replicates
- \blacksquare H_i = (latent) "true" entropy
- $SI_i = (latent) SI score$ (inversely related to H_i^T)
- A_i = "hearing" age (minimum)
- \blacksquare $E_i = etiology of disease$
- \blacksquare HS_i = hearing status
- PTA_i = pure tone average (standardized)
- variables assumed independent, beyond the described relationships,



General causal diagram

$$\begin{split} P(\mathbf{U}) &= P(U_{ik}, U_i, U_A, U_E, U_{HS}, U_P) \\ &= P(U_{ik})P(U_i)P(U_A)P(U_E)P(U_{HS})P(U_P) \end{split}$$

First form

$$\begin{split} &H_{ik}^{O} \leftarrow f(H_{i}^{T}, U_{ik}) \\ &H_{i}^{T} \leftarrow f(SI_{i}) \\ &SI_{i} \leftarrow f(HS_{i}, A_{i}, E_{i}, PTA_{i}, U_{i}) \end{split}$$

$$&HS_{i} \leftarrow f(U_{HS}) \\ &A_{i} \leftarrow f(U_{A}) \\ &E_{i} \leftarrow f(U_{E}) \\ &PTA_{i} \leftarrow f(U_{P}) \\ &U \sim P(\mathbf{U}) \end{split}$$

$$&(a) \text{ general structural model}$$

 $H_{i\nu}^{O} \sim \text{BetaProp}(H_{i}^{T}, M_{ik})$ $H_i^T = inv logit(-SI_i)$ $SI_i \sim Normal(\mu_{SI}, \sigma_{II_i})$ $\mu_{\rm SI} = \alpha + \alpha_{\rm HS[i]} + \alpha_{\rm E[i]}$ $+ \beta_{A \text{ HSfil}}(A_i - \bar{A}) + \beta_P PTA_i$ $HS_i \sim data$ $A_i \sim data$ $E_i \sim data$ $PTA_i \sim data$ $U \sim unobservable$ (a) general probabilistic model

First form

Notice,

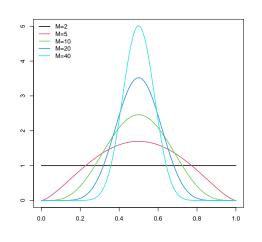
- α , $\alpha_{\text{HS[i]}}$, $\alpha_{\text{E[i]}}$, $\beta_{\text{A,HS[i]}}$, β_{P} are structural parameters (as in SEM)
- U_{ik} = replicates measurement error U_i = between child SI variability
- variability of U_{ik} is modeled by M_{ik} variability of U_i is modeled by σ_{Ui}

```
H_{ik}^{O} \sim BetaProp(H_{i}^{T}, M_{ik})
   H_i^T = inv_logit(-SI_i)
    SI_i \sim Normal(\mu_{SI}, \sigma_{IJi})
        \mu_{\rm SI} = \alpha + \alpha_{\rm HS[i]} + \alpha_{\rm E[i]}
            + \beta_{A,HS[i]}(A_i - \bar{A}) + \beta_P PTA_i
  HS_i \sim data
    A_i \sim data
     E_i \sim data
PTA_i \sim data
      U \sim unobservable
```

(a) general probabilistic model

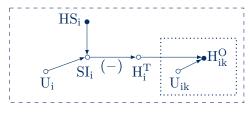
Express variability in BetaProp

$$\begin{split} & \boldsymbol{H}_{ik}^{O} \sim \ \operatorname{BetaProp}(\boldsymbol{H}_{i}^{T}, \boldsymbol{M}_{ik}) \\ & \boldsymbol{H}_{i}^{T} = \ \alpha/(\alpha + \beta) \\ & \boldsymbol{M}_{ik} = \ \alpha + \beta \\ \\ & \alpha = \ \boldsymbol{H}_{i}^{T} \cdot \boldsymbol{M}_{ik} \\ & \beta = \ (1 - \boldsymbol{H}_{i}^{T}) \cdot \boldsymbol{M}_{ik} \\ \\ & \alpha = \ 0.5 \cdot 2 \\ & \beta = \ (1 - 0.5) \cdot 2 \end{split}$$

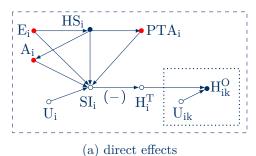


Interested in two effects

- 1. total effects model inherits:
 - children's characteristics that lead to the fitting of specific apparatus,
 - the (convenience of) sample selection (fixed with post-stratification)
- 2. to do the last, we stratify for all variables that explain variability, ergo, use a direct effects model
- 3. two levels: replicates (k), children (i), denoted by discontinuous squares



(b) total effects





Second form

$$\begin{split} H_{ik}^O \leftarrow f(H_i^T, U_{ik}) & H_{ik}^O \sim \\ H_i^T \leftarrow f(SI_i) & H_i^T = \\ SI_i \leftarrow f(HS_i, \textbf{A}_i, \textbf{E}_i, \textbf{PTA}_i, U_i) & SI_i = \\ \\ HS_i \leftarrow f(U_{HS}) & HS_i \sim \\ \\ A_i \leftarrow f(U_A) & A_i \sim \\ \\ E_i \leftarrow f(U_E) & E_i \sim \\ \\ \textbf{PTA}_i \leftarrow f(U_P) & \textbf{PTA}_i \sim \\ \\ U \sim P(\textbf{U}) & U \sim \\ \\ \textbf{(a)} & \text{general structural model} & \textbf{(a)} \end{split}$$

$$\begin{split} H_{ik}^{O} &\sim \ \mathrm{BetaProp}(H_{i}^{T}, M_{ik}) \\ H_{i}^{T} &= \ \mathrm{inv_logit}(-\mathrm{SI}_{i}) \\ \mathrm{SI}_{i} &= \ \mathrm{a}_{i} + \alpha + \alpha_{\mathrm{HS}[i]} + \alpha_{\mathrm{E}[i]} \\ &+ \beta_{\mathrm{A,HS}[i]}(\mathrm{A}_{i} - \bar{\mathrm{A}}) + \beta_{\mathrm{P}}\mathrm{PTA}_{i} \\ \mathrm{a}_{i} &\sim \ \mathrm{Normal}(0, \sigma_{\mathrm{U}i}) \\ \mathrm{HS}_{i} &\sim \ \mathrm{data} \\ \mathrm{A}_{i} &\sim \ \mathrm{data} \\ \mathrm{E}_{i} &\sim \ \mathrm{data} \\ \mathrm{PTA}_{i} &\sim \ \mathrm{data} \\ \mathrm{U} &\sim \ \mathrm{unobservable} \end{split}$$

2. Research hypothesis procedure

Synthetic data generation



Idealized data³

Simulation data can serve as [10, 11],

- 1. A place where to test your model, on multiple purposes,
 - parameter recovery
 - power
- 2. A reflection of a population,
 - De Raeve [6]: 70 HI/CI, 130 HI/HA
 - Our idealized data:
 150 NH, 70 HI/CI, 130 HI/HA
- 3. A reflection of a hypothesis,
 - size of effects

³more details in file: 1 2 E sim fun.R



About the size of the effects (in logits, no previous info)

- 1. aE = 0, ordered by severity $(cor(E, HS) \approx 1$, ordering might not be possible),
- 2. aHS = -0.4, difference between NH children and HI/CI,
- 3. bP = -0.1, per PTA unit $(+10 \text{ PTA units} \Rightarrow -1 \text{ logit})$,
- 4. bA = -0.15, per A unit, above the minimum
 (+10 A units ⇒ +1.5 logits),

- 1. variables are generated in a random fashion
- 2. between children SI variability are defined by the random effects

```
# 1. true data ####
dT = data.frame(matrix(NA, nrow=I, ncol=1))
names(dT) = c('child_id')
dTschild id = 1:I
n = round(p*I)
if( sum(n) != I ){
   n[3] = I - sum(n[c(1,3)]) # to sum the right amount
if(!is.null(seed)){
 set.seed(seed+1)
dT$HS = c(rep(1, n[1]), rep(2, n[2]), rep(3, n[3]))
dT$A = round(rnorm(sum(n), 5, 1))
dTSA = with(dT. ifelse(A>7. 7. A))
dTSE = c(rep(1, n[1]), # no way to know true effects
          sample(2:3, size=n[2], replace=T),
         sample(3:4, size=n[3], replace=T))
dTPTA = c(round(rnorm(n[1], 60, 15)), # first 12 NH
           round(rnorm(n[2], 90, 15)), # next 10
            round(rnorm(n[3], 110, 15))) # last 10
if(!is.null(seed)){
 set.seed(seed-1)
par$re_i = rnorm(I, par$m_i, par$s_i)
dT$re i = par$re i # children's random effects (between SI
```

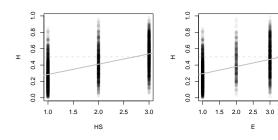
- 1. we use second form of the probabilistic model
- 2. "true' entropy (Ht) is inversely related to SI
- 3. we simulate measurement error through M from BetaProp() distribution.

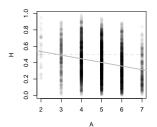
```
dT$SI = with(dT, re_i + par$a + par$aE[E] + par$aHS[HS] +
               par$bA*(A - min(A)) +
               par$bAHS[HS]*(A - min(A)) +
               par$bP*c( standardize(PTA) ) )
# true entropy
dT$Ht = inv_logit(-dT$SI) # true entropy (SI -> Ht: negative
if(!is.null(seed)){
  set.seed(seed+2)
if( is.numeric(par$m_M) & !is.numeric(par$s_M) ){
  par$M = rep(par$m_M, I)
  par$M = round( rlnorm(I, meanlog=par$m_M, sdlog= par$s_M)
dT$M = par$M # same df for all children (not same shape!!)
dT[,6:ncol(dT)] = round(dT[,6:ncol(dT)], 5)
N = I*K
d0 = data.frame(matrix(NA, nrow=N, ncol=3))
names(d0) = c('child_id', 'utt_id', 'H')
d0$child_id = rep(1:I, each=K)
dO$utt_id = rep(1:K, I)
if(!is.null(seed)){
  set.seed(seed-2)
```

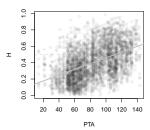
- 1. we simulate replicate measures of entropy (H)
- 2. we storage all relevant parameters and data

```
idx = d0$child_id == i
 dO$H[idx] = rbeta2(n=K, prob=dT$Ht[i], theta=dT$M[i])
dO$H = round(dO$H, 5)
 N = nrow(dO), # observations
 I = max(d0$child_id), # children
 K = max(dO$utt_id), # utterances
 cHS = max(dT$HS),
 cE = max(dTSE).
 sPTA = c( standardize( dT$PTA ) ),
 H = with(d0, ifelse(H==0, 0.0001, ifelse(H==1, 0.9999, H))),
 cid = d0$child_id,
 uid = dO$utt_id
nom = list(dS=list( dT=dT, dO=dO, par=par), dL=dL)
```

Example







3.5 4.0



2. Research hypothesis procedure

Statistical model design and testing



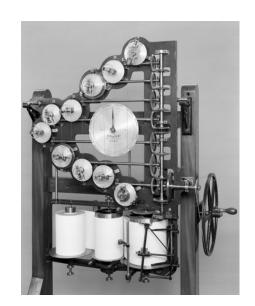
Model design and test⁴

Purpose:

- to have reliable procedures,
- to maintain a clear documentation,
- to have a sound analysis

Procedure:

- step by step, instantiating one difficulty at the time
- use probabilistic assumptions defined in estimand and process model
- is like running synthetic data generation backwards



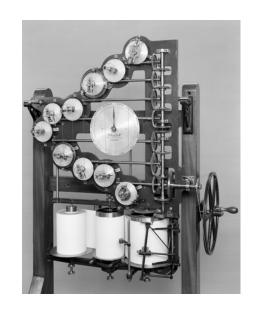


⁴Following Fogarty et al. [8]

Model design and test

We evaluate,

- the probabilistic model implementation [11, 2] (centered and non-centered versions)
- prior predictive
- "health" of MCMC chains
- parameter recovery
- posterior predictive
- power



starts with,

- the simplest model
- the simplest data generating procedure

```
transformed parameters{
    vector[I] SI;
                            // SI index (per child)
    vector[I] Ht;
                            // true entropy (per chi
    SI = a + re_i;  // linear predictor
Ht = inv_logit(-SI);  // average entropy (SI -
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    // priors
    a \sim normal(0, 0.2);
    re_i \sim normal(m_i . s_i):
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , 10 );
```

try,

• the centered and non-centered parametrization

```
transformed parameters{
    vector[I] re_i;
                           // random intercepts (per
    vector[I] SI; // SI index
    vector[I] Ht:
                          // true entropy (per child
    re_i = m_i + s_i*z_re;// non-centered RE
   SI = a + re_i;  // linear predictor
Ht = inv_logit(-SI);  // average entropy (SI ->
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    s_i \sim exponential(1):
    // priors
    z_re ~ std_normal();
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , 10 );
```

escalate,

- complexity of model
- traits of the data

```
transformed parameters{
    vector[I] SI:
                          // SI index (per child)
    vector[I] Ht;
                          // true entropy (per chi
    // linear predictor
    for(i in 1:I){
      SI[i] = re_i[i] + a + aHS[HS[i]] +
        bA*Am[i] + bP*sPTA[i]:
      // no multicollinearity between E and HS
    // average entropy (SI -> Ht: negative)
    Ht = inv logit(-SI): // average entropy (SI -
model{
    // hyperpriors
    m_i \sim normal(0, 0.2);
    s_i \sim exponential(1):
    // priors
    a \sim normal(0, 0.2);
    re_i \sim normal( m_i , s_i );
    aHS \sim normal(0, 0.5);
    bP \sim normal(0, 0.3);
    bA \sim normal(0, 0.3);
    m_M \sim lognormal(1.5, 0.5);
    // likelihood
    for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , m_M );
```

We tested 5 random effects models: (from 5 synthetic data types) (centered, and non-centered)

- \bullet only intercept, M = 10,
- \blacksquare causal model, M = 10,
- causal model, M per individual,
- no known process,
- causal model with interaction, M per individual,

```
transformed parameters{
   vector[I] SI:
                          // SI index (per child)
   vector[I] Ht;
                          // true entropy (per child
   // linear predictor
    for(i in 1:I){
      SI[i] = re_i[i] + a + aHS[HS[i]] +
        bA*Am[i] + bP*sPTA[i]:
      // no multicollinearity between E and HS
   // average entropy (SI -> Ht: negative)
   Ht = inv_logit(-SI): // average entropy (SI ->
model{
   // hyperpriors
   m_i \sim normal(0.0.2):
   s_i \sim exponential(1);
   // priors
    a \sim normal(0, 0.2);
    re_i \sim normal(m_i . s_i):
    //aE \sim normal(0.0.5):
   aHS \sim normal(0, 0.5);
   bP \sim normal(0, 0.3);
   bA \sim normal(0, 0.3);
   m M ~ lognormal( 1.5 . 0.5 ):
   // likelihood
   for(n in 1:N){
      H[n] ~ beta_proportion( Ht[cid[n]] , m_M );
```

Prior predictive

Priors and hyper-priors

- In the probabilistic (causal) model there were no priors for our parameters,
- To decide our priors we follow McElreath [11]: "priors are part of the assumptions, and should be inspected as such",
- We will evaluate the implications of our priors on the outcome scale.
 We have three outcomes scales: SI_i, H_i^T, and H_{ik}^O

```
Priors
                  a_i \sim Normal(\mu_a, \sigma_a)
                M_i \sim LogNormal(\mu_M, \sigma_M)
                  \alpha \sim \text{Normal}(0.0.2)
           \alpha_{\rm HS[i]} \sim {\rm Normal}(0, 0.3)
             \alpha_{\rm E[i]} \sim {\rm Normal}(0, 0.3)
        \beta_{\text{A.HS[i]}} \sim \text{Normal}(0, 0.3)
                 \beta_{\rm P} \sim \text{Normal}(0, 0.3)
Hyper-priors
                 \mu_a \sim \text{Normal}(0, 0.2)
                 \sigma_a \sim \text{Exp}(1)
                \mu_{\rm M} \sim {\rm Normal}(0, 0.5)
```

 $\sigma_{\rm M} \sim {\rm Exp}(1)$

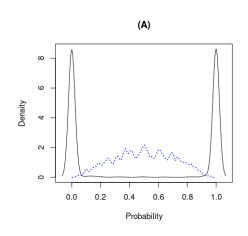
Prior predictive

Undesired assumptions can easily creep in non-linear models^a Example:

(black line)
$$\theta \sim N(0, 100)$$

$$logit(p) = \theta$$
 (blue line)
$$\theta \sim N(0, 1)$$

$$logit(p) = \theta$$



^aFigure extracted from Rivera [14].

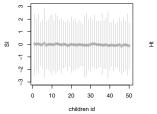
Prior predictive

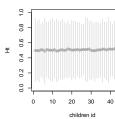
What our priors imply?

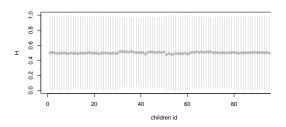
NO undesired assumption has crept in:

- the SI_i scale,
- the H_i^T scale,
- \blacksquare the H_{ik}^{O} scale

i.e. the scales' full space can be reached by (a combination of) the parameters







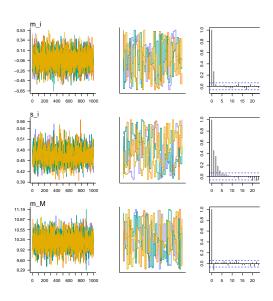


"Health" of MCMC chains

The MCMC chains achieve,

- good convergence
- good mixing
- lack of autocorrelation

same results on the n_eff and RHat statistics.



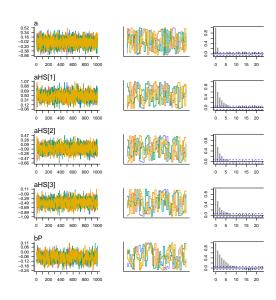


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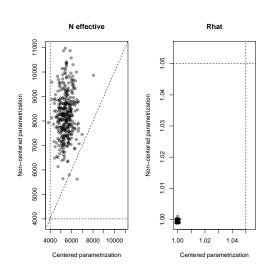




"Health" of MCMC chains

The non-centered parametrization has,

- better n_eff (denoting lack of autocorrelation)
- better Rhat (denoting good convergence)
- better mixing (inspected visually, not shown)



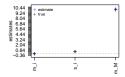


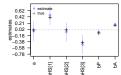
Parameter recovery

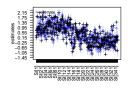
On idealized data

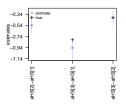
The model,

- recovers the parameters in the right scale,
- most of the "true" parameters are inside of the compatibility intervals (CI)
- contrast are approximately correct









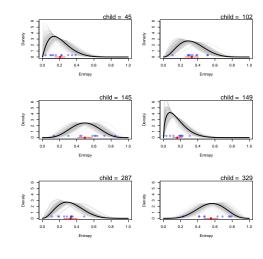
Posterior predictive

But how well reproduces the data?, The model with less working assumptions^a,

- captures the variability of the replicates,
- provides a "true" H and SI

More complex models,

- captures even better the data, but might overfit,
- avoid overfit when the model is fit to the data (ITA [1, 4])



 $^{^{\}rm a}$ random effects causal model, with M = 10

Equivalent prior sampling method [18, 10]

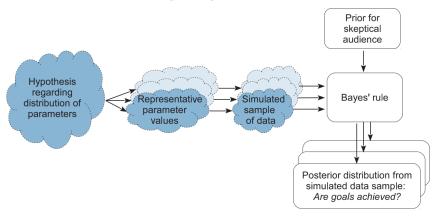
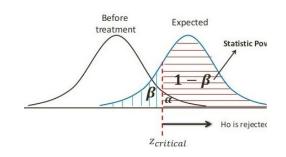


Figure: Flow of information in a power analysis, in which simulated data come from random hypothetical parameters. Extracted from Kruschke [10].

Equivalent prior sampling method

- 1. generate idealized data
- 2. generate parameters' distribution (with Bayes theorem)
- 3. simulate data sample (variability in data and parameters)
- 4. apply models to simulated data (include priors)
- 5. evaluate desire goals
 - lacktriangledown reject the null hypothesis
 - affirm predicted value
 - achieve precision in estimate
- 6. repeat procedure (to approximate power)





Equivalent prior sampling method

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 - achieve precision in estimate
- 6. repeat procedure (to approximate power)

Epower(power_save=file.palth(getwd(), 'sim_chain'), # po
sim_name='Hbeta_sim2_power.RData', # file_save n
sim_save=file.path(getwd(), 'sim_data'), # file_
model_name='Hbeta_NC_sim2', # model for which we
model_in=file.path(getwd(), 'sim_models'), # loc
model_out=file.path(getwd(), 'sim_chain'), # loc
Nsim=100, # number of simulation for power
L_grid = c(48, 60), # experimental units (childr
K_grid = c(10, 20), # replicates (utterances)
p=c(0.34, 0.33, 0.33), # children prop. on each
par_int=c('aHs', 'bP', 'bA', 'mi', 's.i', 'mi', 'SI'
par_cont=c('aHs', 'SI')) # parameters to contras



Equivalent prior sampling method

- 1. generate idealized data
- 2. generate parameters' distribution (with Bayes theorem)
- 3. simulate data sample (variability in data and parameters)
- 4. apply models to simulated data (include priors)
- 5. evaluate desire goals^a
 - reject the null hypothesis
 - affirm predicted value
 - achieve precision in estimate
- 6. repeat procedure (to approximate power)



2,012,0123,013,		
	Effect size	d
	Very small	0.01
	Small	0.20
	Medium	0.50
	Large	0.80
	Very large	1.20
	Huge	2.0



^ause ROPE [10], and effect sizes [5, 15]

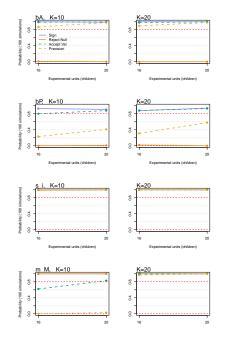
Power results

For this parameter set,

- reject the null hypothesis, never achieved in structural parameters,
- affirm predicted value, in all shown parameters
- achieve precision in estimate, in some parameters is reached

Notice,

- between child SI variability,
- replicates measurement error (M),
- not much difference with 2x comparisons (except for M)





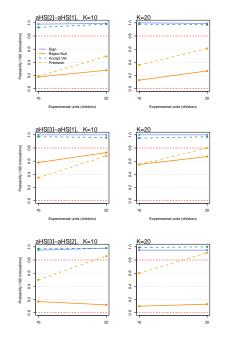
Power results

For group contrasts,

- reject the null hypothesis,
 closest for the largest contrast,
- affirm predicted value, in all shown contrasts,
- achieve precision in estimate, grows with more children

Notice,

- not much difference with 2x comparisons
- group contrasts are easier to identify (larger sample size per comparison)



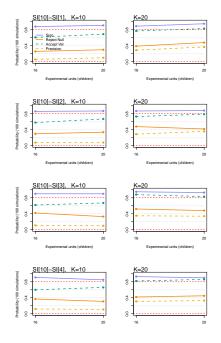
Power results

For individual contrasts,

- reject the null hypothesis,
 depends on the contrast of interest,
- affirm predicted value, for some is achieved,
- achieve precision in estimate, clearly requires more sample size

Notice,

■ we see a clear difference with 2x comparisons (more sample at the appropriate level)



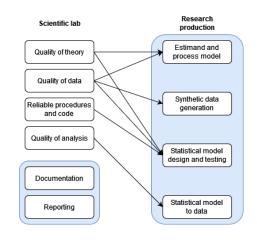
2. Research hypothesis procedure

Apply statistical model to data



What we have so far

- 1. measure: replicated entropies H_{ik}^{O}
- 2. estimand:
 SI index, structural parameters,
 contrasts
- 3. structural models: total and direct effects
- 4. probabilistic models: three possible fitting models
- 5. statistical models: works as intended
- 6. power: enough



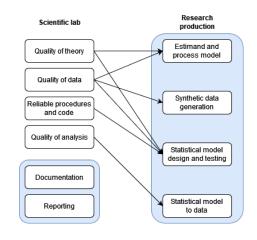
Information Theoretic Approach [1, 4]

The last step would be to select the most fitting model, and we do it with the ITA,

- 1. hypothesis into statistical models,
- 2. select among competing models,
- 3. make inferences based on one or multiple models.

Select models with information criteria,

- 1. WAIC [17]
- 2. PSIS [16]



Competing models

models_table.png

3. References



3. References



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