Causes and effects in Dichotomous Comparative Judgments: an information-theoretical system with plausible mechanism

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Abstract

Dichotomous Comparative Judgment (DCJ, Pollitt, 2012a, @Pollitt 2012b) requires judges to evaluate the relative manifestation of traits between pairs of stimuli, resulting in a dichotomous outcome indicating which stimulus exhibits the trait more strongly. Research has demonstrated DCJ's effectiveness and reliability in various domains (Pollitt, 2012b; Bartholomew et al., 2018; van Daal et al., 2019; Lesterhuis, 2018; Bartholomew and Williams, 2020; Boonen et al., 2020). Nevertheless, despite the method's widespread use, the literature lacks a transparent depiction of the DCJ system and the plausible mechanisms that generate the DCJ data. Particularly, there is no detailed explanation of how different assessment factors can potentially influence the observed DCJ data. This study aims to fill this gap by applying the framework of causal analysis and Directed Acyclic Graphs [DAG; Pearl (2009)]. Using this framework, the study will construct a scientific model to elucidate the causal assumptions and mechanisms inherent the system. This model will enable researchers to draw inferences about causal relationships from DCJ data. Subsequently, the study will translate this model into a probabilistic statistical model, aiming to derive statistical estimands for different targets of inference. The outcomes of this study will inform the planning of DCJ experiments and hold significance for researchers or analysts involved in education and assessment procedures who implement the DCJ methodology.

 $\it Keywords:$ comparative judgement, directed acycilc graph, causal analysis, probabilistic statistics

1. Introduction

In contemporary contexts, Thurstone's law of comparative judgment (1927) primarily refers to the method of *dichotomous* comparative judgment (DCJ, Pollitt, 2012a,b).

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In DCJ, a judge assesses the relative manifestation of a *trait* within a pair of stimuli. This assessment results in a dichotomous value indicating which stimulus possesses a higher degree of the trait. After different judges perform multiple rounds of pairwise comparisons, an outcome vector is produced. This vector is modeled using the Bradley-Terry-Luce model (BTL, Bradley and Terry, 1952; Luce, 1959), which creates a score that corresponds with the trait of interest. This score is then used to rank the stimuli from lowest to highest or to evaluate the influence of certain variables on the stimuli's positions in the ranking.

DCJ has proven effective in assessing competencies and traits predominantly within the educational realm, as demonstrated by Pollitt (2012b), Jones (2015), van Daal et al. (2019), Bartholomew et al. (2018), Lesterhuis (2018), Bartholomew and Williams (2020), and Marshall et al. (2020). However, its application transcends education, as exemplified by Boonen et al. (2020). The methodology has also evolved to include multiple, as opposed to pairwise comparisons (Luce, 1959; Plackett, 1975), and to accommodate comparisons with ordinal outcomes (Tutz, 1986; Agresti, 1992). Overall, research suggests that DCJ offers an alternative and efficient approach to measurement and evaluation, characterized by its reliability and validity (Lesterhuis, 2018; van Daal, 2020; Marshall et al., 2020). Nevertheless, despite the method's widespread use, the literature does not offer a clear representation of the plausible mechanisms that generate DCJ data. Particularly, there is no depiction of the complexity and the underlying assumptions of the DCJ system, nor how different assessment factors can potentially influence the observed DCJ outcome.

According to Verhavert et al. (2019) and van Daal (2020), several assessment factors interact and influence the method's outcome. These factors include the number and characteristics of the stimuli, their proximity in terms of the assessed trait, the number of comparison per stimulus, and the pairing algorithm used. Furthermore, since the method relies on judges' assessments, the number and characteristics of judges, their discrimination abilities, and the number of comparisons per judge also play pivotal roles. Moreover, when the stimuli represent sub-units of higher-levels units, factors such as the number and characteristics of these units, along with their proximity in terms of the assessed trait, can significantly influence the outcome. For example, van Daal et al. (2019) assessed the academic writing skills of university students (units) using multiple argumentative essays (sub-units).

Although several studies have examined the individual impact of these factors on the method's reliability (Bramley, 2015; Pollitt, 2012b; Bramley and Vitello, 2019; Verhavert et al., 2019; Crompvoets et al., 2022; van Daal et al., 2017; Gijsen et al., 2021), none, to the best of the authors' knowledge, have provided a transparent depiction of the DCJ system and the mechanisms generating the DCJ outcome. This study aims to fill this gap by representing DCJ within the causal inference framework. Specifically, using the *structural approach* to causal inference (Wright, 1927; Pearl, 2009; Pearl et al., 2016), the study aims to construct a scientific model. This model will elucidate the underlying assumptions of the DCJ system, providing plausible mechanisms for how the DCJ outcome is generated. Next, using a minimal set of assumptions, the study will translate the scientific model into a probabilistic statistical model. This model will produce statistical estimates to draw inferences about plausible causal relationships within the DCJ system.

Ultimately, this research aims to extend the law of comparative judgment initially proposed by Thurstone (1927) and provide a sound probabilistic base for the statistical analysis of DCJ data. Consequently, this research holds significance for researchers and analysts involved in education and assessment procedures who implement or design DCJ experiments.

2. Preliminaries

2.1. The structural approach to causal inference

In statistics, causal inference refers to the process of identifying the causes of a phenomenon and estimating their effects using data (Shaughnessy et al., 2010; Neal, 2020). Unlike classical statistical modeling, which focuses solely on summarizing data and inferring associations, causal inference provides a coherent mathematical representation for analyzing causes and counterfactuals (Pearl, 2009). Counterfactuals reflect scenarios contrary to fact, where alternative potential outcomes have not been observed and cannot be observed (Neal, 2020; Counterfactual, 2024). Using counterfactuals, researchers can construct a theory of the world that explains why specific causes have specific effects, and what happens in the absence of those causes. Consequently, counterfactuals represent the highest level of abstraction in causal inference (Pearl and Mackenzie, 2018).

Several approaches to constructing counterfactuals and applying causal inference exist, but two are particularly prominent: the potential outcomes approach and the structural approach. The potential outcomes approach, also known as the Neyman-Rubin causal model, was developed by Neyman et al. (1990) and Rubin (1974). In contrast, the structural approach was pioneered by Wright (1927), and later extended by Pearl (Pearl, 2009; Pearl et al., 2016). Both approaches employ rigorous mathematical notation to characterize causal inference, but they do so in different ways (Neal, 2020). The potential outcomes approach relies on counterfactual notation, whereas the structural approach utilizes the do-operator and Structural Causal Models (SCM, Pearl, 2009; Pearl et al., 2016). Despite these differences, both approaches provide methods for using experimental and observational data to estimate causal effects (Pearl, 2010).

However, the structural approach offers a significant advantage over the potential outcomes approach by enabling the graphical representation of systems through directed acyclic graphs (DAG, Gross et al., 2018; Neal, 2020). These graphical representations provides a transparent depiction of a system's complexity, revealing its underlying assumptions and the plausible mechanisms generating the system's outcome. This capability aligns well with the aims of this study. Consequently, the structural approach will be the primary method used to address the study's research goal.

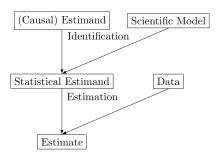


Figure 1: Identification-Estimation flowchart. Extracted from Neal (2020, 32)

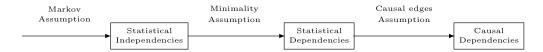


Figure 2: The flow of association and causation in graphs. Extracted from Neal $(2020,\,31)$

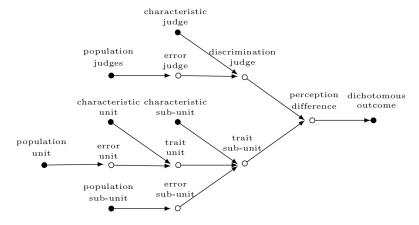


Figure 3: DCJ causal diagram, simplified description

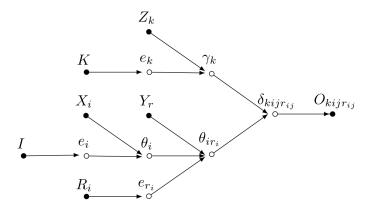


Figure 4: DCJ causal diagram, simplified mathematical description

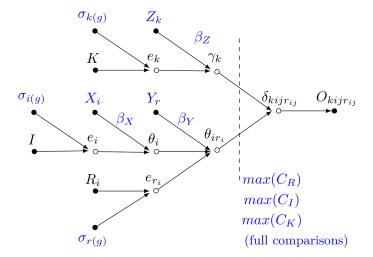


Figure 5: DCJ causal diagram, population mathematical description $% \left(1\right) =\left(1\right) \left(1\right) \left$

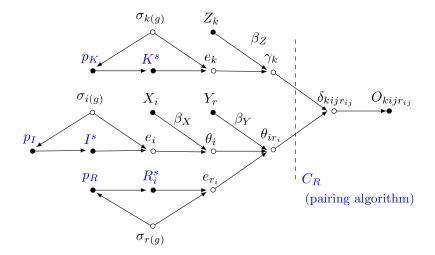


Figure 6: DCJ causal diagram, sample with comparisons mathematical description

- 2.2. Graphs, DAGs and SCMs
- 2.3. The flow of association and causation

3. Theoretical framework

- 3.1. A scientific model for the DCJ
- $\it 3.2.\ Probabilities\ assumptions\ of\ the\ scientific\ model$

$$\begin{split} O_{kijr_{ij}} &:= f_O(\delta_{kijr_{ij}}) \\ \delta_{kijr_{ij}} &:= f_D(\gamma_k, \theta_{ir_i}) \\ \gamma_k &:= f_G(Z_k, e_k) \\ \theta_{ir_i} &:= f_R(\theta_i, Y_r, e_{r_i}) \\ \theta_i &:= f_T(X_i, e_i) \\ e_k & \!\!\!\perp \!\!\!\perp e_i \\ e_k & \!\!\!\perp \!\!\!\perp e_{r_i} \\ e_i & \!\!\!\perp \!\!\!\perp e_{r_i} \end{split}$$

3.3. From the scientific to statistical model

$$\begin{split} O_{kijr_{ij}} &\sim \text{Bernoulli} \left[\ logit^{-1} \left(\delta_{kijr_{ij}} \right) \ \right] \\ \delta_{kijr_{ij}} &= \gamma_k (\theta_{ir_i} - \theta_{jr_j}) \\ \gamma_k &= logit^{-1} \left[\beta_Z Z_k + e_k \right] \\ \theta_{ir_i} &= \theta_i + \beta_Y Y_r + e_{r_i} \\ \theta_i &= \beta_X X_i + e_i \\ e_k &\sim \text{Normal}(0, \sigma_{k(g)}) \\ e_i &\sim \text{Normal}(0, \sigma_{i(g)}) \\ e_i &\sim \text{Normal}(0, \sigma_{r(g)}) \end{split}$$
 (2)

for identification purposes we can set $\frac{1}{G}\sum_{g=1}^G\sigma_{k(g)}=0.02,\;\frac{1}{G}\sum_{g=1}^G\sigma_{i(g)}=1,\;$ and $\frac{1}{G}\sum_{g=1}^{G}\sigma_{r(g)}=1$. A special case of this would be to assume that the data comes from the same population, in that case, $\sigma_{k(g)} = \sigma_k = 0.02, \, \sigma_{i(g)} = \sigma_i = 1$

3.4. Let's talk about Thurstone

Thurstone's comparative judgment Thurstone (1927) is based on the formula:

$$X_{AB} = \frac{S_A - S_B}{\sigma_{AB}}$$

where X_{AB} defines the comparative judgment outcome, S_A and S_B are the modal discriminal processes, $\sigma_{AB} = \sqrt{\sigma_A^2 + \sigma_B^2 + 2\rho\sigma_A\sigma_B}$, with σ_A and σ_B being the dispersion of discriminal processes A and B, respectively, and ρ the correlation between discriminal processes.

The theory identifies five cases:

- Case 1: only constant ρ (not ρ_{ij}) Case 2: X_{ij} becomes X_{kij} with $k=1,\ldots,K$ judges (replication, not duplication)
- Case 3: $\rho = 0$, then $\sigma_{AB} = \sqrt{\sigma_A^2 + \sigma_B^2}$
- Case 4: $\sigma_B=\sigma_A+d,$ then $\lim_{d\leq 0.1\sigma_A}\sigma_{AB}=(\sigma_A+\sigma_B)/\sqrt{2}$
- Case 5: $\sigma_B = \sigma_A$, then $\sigma_{AB} = \sqrt{2}\sigma$

Now using the DAG and statistical notation

$$\begin{split} O_{kijr_{ij}} &:= f_O(\delta_{kijr_{ij}}) \\ \delta_{kijr_{ij}} &= \gamma_k(\theta_{ir_i} - \theta_{jr_j}) \\ \gamma_k &= f_G(Z_k, e_k) \\ \theta_{ir_i} &= \theta_i + \beta_Y Y_r + e_{r_i} \\ \theta_i &= \beta_X X_i + e_i \\ e_k &\sim \text{Normal}(0, \sigma_{k(g)}) \\ e_i &\sim \text{Normal}(0, \sigma_{i(g)}) \\ e_{r_i} &\sim \text{Normal}(0, \sigma_{r(g)}) \end{split}$$

The theory identifies five cases:

- Case 1: only constant $\rho \approx \sigma_i$
- Case 2: now judges are separated by using γ_k
- Case 3: $\rho \approx \sigma_{e_i} = 0$ (no nesting of texts on students), then $\sigma_{AB} = \sqrt{\sigma_A^2 + \sigma_B^2}$
- Case 4: $\sigma_B=\sigma_A+d,$ then $\lim_{d\leq \underline{0}.1\sigma_A}\sigma_{AB}=(\sigma_A+\sigma_B)/\sqrt{2}$
- Case 5: $\sigma_B = \sigma_A$, then $\sigma_{AB} = \sqrt{2}\sigma$

But now can we see other scenarios than just those 5 cases?

- consider different $\rho \approx \sum_{p=1}^P \sigma_p$, depending on P nesting structures we can now investigate γ_k
- we can assume $\sigma_B \neq \sigma_A$, no need for results on the limit

4. Discussion

- 4.1. Findings
- 4.2. Limitations and further research
- 5. Conclusion

Declarations

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Code availability: All the code utilized in this research is available in the digital document located at: https://jriveraespejo.github.io/paper2_manuscript/.

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6. Appendix

6.1. Additional definitions

${\bf Counterfactual:}$

6.2. Why do we need to estimate judges' abilities?

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