

# Generalized Linear Latent and Mixed Models:

method, estimation procedures, advantages, and applications to educational policy.

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# Dedication

To Manuel, for being my friend and father.  
To Margarita, Susan, and Marysu, for their relentless encouragement.  
To Ana, for showing me the value of family, here in this moorland.  
To both of you, as you are always in my mind.  
And to all that knowingly or not, help me to get here.  
I am lucky due to all of you.  
I hope I make you all proud.

A Manuel, por ser mi amigo y mi padre.  
A Margarita, Susan y Marysu, por su incansable aliento.  
A Ana, por mostrarme el valor de la familia, aquí en este páramo.  
A ustedes dos, que siempre las tengo en mente.  
Y a todos los que sabiendolo o no, me ayudaron a llegar aquí.  
Soy un suertudo gracias todos ustedes.  
Espero llenarlos de orgullo.

# Acknowledgment

(work in progress)

# Abstract

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**Keywords:**

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# Abbreviations

CFA	Confirmatory Factor Analysis.
DDM	Dual Dependence Models.
EFA	Exploratory Factor Analysis.
GLLAMM	Generalized Linear Latent and Mixed Model.
GLM	Generalized Linear Model.
GLMM	Generalized Linear Mixed Model.
HMC	Hamiltonian Monte Carlo.
IRT	Item Response Theory models.
MCMC	Markov Chain Monte Carlo.
ML	Maximum Likelihood.
MSEM	Multilevel Structural Equation Model.
SEM	Structural Equation Model.



# Chapter 1

## Introduction

### 1.1 Preliminar considerations

Local independence is one of the key assumptions of Item Response Theory (IRT) models, and it is comprised of two parts: (i) local item independence and (ii) local individual independence [3, 20]. In the former case, the assumption entails that the individual's response to an item does not affect the probability of endorsing another item, after conditioning on the individual's ability. While in the case of the latter, the assumption considers that an individual's response to an item is independent of another person's response to that same item [43].

The literature has shown that IRT models are not robust to the violation of local independence. The transgression of the assumption affects model parameter estimates, inflates measurement reliabilities and test information, and underestimates standard errors (see Yen [54], Chen and Thissen [8], and Jiao et al. [23]).

However, item response data arising from educational assessments often display several types of dependencies, e.g. testlets, where items are constructed around a common stimulus [52]; the measurement of multiple latent traits within individuals [43]; cluster effects, where correlation among individuals results from the sampling and measurement mechanism used to gather the data [42]; among others. A good motivating example, that will permeate this research, is the reading comprehension sub-test, from the Peruvian public teaching career national assessment. The test is designed to measure three hierarchically nested sub-dimensions of reading comprehension: literal, inferential, and reflective abilities. Furthermore, the items are bundled together in testlets related to a common text or passage. Finally, multiple cluster effects are present, e.g. at the region, and district level, just to mention a few.

Recent studies have proposed IRT Dual Dependency Models (DDM) to deal with the testlets and individual clustering dependencies observed in the data [14, 13, 12, 23, 10, 11, 43, 6]. The majority of these representations have been developed under the Bayesian framework, and they are similar in parametrization to multilevel models. On the other hand, an almost independent line of research, the Generalized Linear Latent and Mixed Models (GLLAMM) [37, 39, 47, 40], have extended the capabilities of hierarchical models on the estimation of multiple latent traits at different hierarchical levels. These developments have been motivated mostly under the frequentist framework, and they are similar in parametrization to a Multilevel Structural Equation Model (MSEM).

While the initial sense is that both developments are independent of each other, follow-

ing their literature, one can easily notice that they share more than a resemblance. Both follow a multilevel/hierarchical multidimensional approach to account for the clustering of persons within the samples and the item bundles (DDM), or the latent structures within the individuals (GLLMM). However, it is important to point out that in some cases the model parametrization between the two developments differs in a way, that some of them appear to be useful only under their specific contexts. Fortunately, their integration under the Bayesian framework is not only trivial but can be motivated under either type of model.

The benefits of the integration revolve around two facts: (i) the educational data often presents all of the aforementioned dependencies and more, as in the motivating example; and (ii) as it was hinted in the second paragraph, to reach appropriate conclusions from the parameter estimates, IRT models need to account for all of these dependencies. The latter is particularly important as, more often than not, a researcher is interested in producing inferences at the structural level of the model, i.e. how a different set of manifest variables explain the variability in the latent variables, or how the latent variables explain other manifest or latent variables, at different levels. As an example, one might be interested in finding evidence if the latent “abilities” of the teachers are explained by their initial educational conditions, i.e. if they were educated in an institute, university, or both. The main purpose of this would be to identify the type of teacher that might benefit more from the in-service training<sup>1</sup>, offered by the national educational authorities, making the intervention cost-effective.

From the previous description, one can infer that the proposed IRT representation would be complex and highly dimensional. Moreover, as educational assessments are usually scored in a binary way (the individual either endorse or did not the item), and because not all individuals are assessed by all the items, the model will be trained on sparse data. From the modeling perspective, neither of the previous points presents a challenge for the bayesian framework. However, it has long been recognized that complex parametrizations, that allow this powerful modeling schemes, introduce pathologies that make Markov Chain Monte Carlo methods (MCMC) face performance challenges [15, 16, 32, 33, 4]. This is highly relevant because, in order to make inferences about the posterior distribution of the parameters, the chains need to achieve three requirements, highly related to the performance of the method: stationarity, convergence, and good mixing [29].

Throughout the bayesian IRT literature, one often finds that four solutions are offered to ensure the fulfillment of the previous requirements, and they can be classified into two broad groups: (i) solutions that involve changing the settings of the MCMC methods, and (ii) solutions that involve readjusting the Bayesian model.

In the first category, we find two proposals: (a) increasing the number of iterations per chain, with large burn-in and thinning processes, and (b) designing model-specific MCMC algorithms. The easiest to implement and more prevalent in the literature is the former, e.g. Fujimoto [13] used chains with 60,000 iterations, where 15,000 were discarded and the remaining were thinned in jumps of 3; while Fujimoto [12] used 225,000 iterations, with burn-in of 30,000 and thinning with jumps of 15. Among the drawbacks of this solution are the large computational times; the user involvement in deciding the specific setting for the process, which could be different for different parameters in the same model;

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<sup>1</sup>Intervention designed with the purpose of potentiating specific abilities in teachers that are currently part of the public teaching career.

and finally, the lack of confidence that larger chain iterations actually produce a proper posterior investigation. On the other hand, several authors have developed high-tech MCMC algorithms that aim to optimize their performance within a particular class of models [33]. In these cases, the developers re-evaluate not only the use of the programming language, with the purpose of speeding and improving performance (e.g. Fujimoto [13]); but also the inclusion of ad-hoc model assumptions, like the ones used in staple software developments like Mplus [30] or Stata [38]. It is clear from the previous that this solution is not accessible to all researchers, either because of the lack of programming skills, or the restrictive cost of access involved in acquiring the software. But more importantly, these solutions are not always applicable to a wider framework of similar models [33].

In the second category, we also find two proposed solutions: (a) re-write the model in an alternative parametrization, and (b) encode prior information through the prior distributions. On both solutions, the purpose is to ensure the identification of the parameters within the model, which helps to stabilize the MCMC procedure [17]. An example of the former is Fujimoto [13], which decomposed the items' discriminatory parameters into overall and specific item discriminations. For the latter, Fujimoto [14] used informative priors also for the items' discrimination parameters.

More often than not, researchers use two or more of the aforementioned solutions to reach an acceptable performance in the chains. However, as point out by Betancourt and Girolami [4], even the most simple hierarchical models present formidable pathologies, that no simple correction can be performed to visit the posterior distribution properly. This is true no matter the rotation/rescaling of the parameter, or the amount of data. In this context, several authors [15, 16, 32, 33, 4] showed that prior information can be included in the model, not only through the prior distributions but also by encoding it in the model itself, changing the posterior sampling geometries and favoring the performance of MCMC chains.

Given all of the above, the present research will focus its attention on showing how easy is to account for all of the dependencies that educational data often display, under the GLLAMM framework. Furthermore, given that only the literature related to gaussian hierarchical models have shown the benefits of changing the posterior sampling geometries, through the use of the non-centered parameterization [15, 16, 32, 33, 4], it seems sensible to provide a similar assessment for nonlinear hierarchical models, and in particular, the ones with latent stochastic processes like IRT [33]. Finally, the research will apply the newfound knowledge to data coming from a large Teacher's standardized educational assessments from Peru.

## 1.2 Objectives

As mentioned in the previous paragraph, the present research has a three-fold purpose:

1. Motivate the Bayesian GLLAMM for binary outcomes [37, 39, 47, 40]. The representation will emphasize the modeling of multiple hierarchical latent structures and testlets. This, in turn, will effectively blur the division between the GLLAMM framework and IRT models.
2. Empirically evaluate the benefits of changing the posterior sampling geometry. The emphasis here will be on comparing the centered and non-centered parametrizations

[15, 16] in terms of performance of the chains, the parameters' recovery, and the ability of the model to produce appropriate inferences.

3. Apply the model and its parametrization to a real data setting. Here the emphasis will be to assess the conclusions arrived from the application of the model and what they could imply for the educational authorities.

Given the aforementioned goals, the researcher believes the master's thesis contributes to the literature in two aspects:

1. In a theoretical and methodological sense, as the research is focused on describing a model that effectively controls for the multiple dependencies usually observed in educational datasets; and
2. In a more practical sense, as the study will provide empirical evidence if changing the sampling posterior geometries could (could not) benefit the performance of MCMC methods and therefore the inferences, under IRT models.

Finally, it is important to mention, that the computational implementation of the method will be developed in `Stan` [50] and `R` [34, 49].

## 1.3 Organization

Chapter 2 will motivate the GLLAMM for dichotomous outcomes, and define its components. Chapter 3 will describe the bayesian framework, its benefits and shortcomings. Moreover, it will outline the evidence behind the change in posterior sampling geometries, and the computational implementation of the model. Chapter 4 will show the results of an empirical simulation study designed to assess the benefits of the re-parametrization, proposed in the previous chapter. Chapter 5 will describe the instruments, the data collection process, and scales under analysis, for a large standardized educational assessment. Furthermore, the chapter will show the conclusions achieved by the application of the model in the said data. Finally, Chapter 6, will discuss the conclusion for the research, and it will outline the path of future research topics that can be derived from the present effort.

# Chapter 2

## The GLLAMM for dichotomous outcomes

The Generalized Linear Latent and Mixed Model (GLLAMM) is a framework that unifies a wide range of latent variable models. Developed by Rabe-Hesketh and colleagues [37, 39, 38, 47, 40], the method was motivated by the need of a Multilevel Structural Equation Model (MSEM) that accommodates for unbalanced data, noncontinuous responses and cross-level effects among latent variables. The authors focused its development mainly from the frequentist perspective, however, they offered a general guidance on implementing the model under the bayesian framework (see Skrondal and Rabe-Hesketh [47]).

### 2.1 Model motivation

Consider a large standardized assessment composed of three sub-test designed to evaluate the reading comprehension, mathematical reasoning, and pedagogical knowledge of teachers; where each sub-test has several dichotomously scored items.

Focusing on the first sub-test, the items were designed to measure only one of the three hierarchically nested sub-dimensions of reading comprehension: literal, inferential, and reflective abilities. Furthermore, its is assumed the three sub-dimensions are all that is needed to measure the reading comprehension ability, effectively making this scale the highest level latent variable in the model, similar to a hierarchical Confirmatory Factor Analysis (CFA). Finally, the items were bundled in groups of five to a common text or passage, i.e. testlets, that provided the stimulus over which the individual is assessed. Figure 2.1 shows the path diagram of the hypothesized dimensional structure, for the hierarchical cross-classified IRT model corresponding with the instrument.

With the purpose of providing an easier motivation of the model, we will not consider yet the cluster effects; however, later in the presentation we will show how easy is to introduce them in the model. Just for future reference, under this example, one expects to observe clustering effects, because individuals from different regions did not have the same educational opportunities, effectively causing differences among them at a regional level.

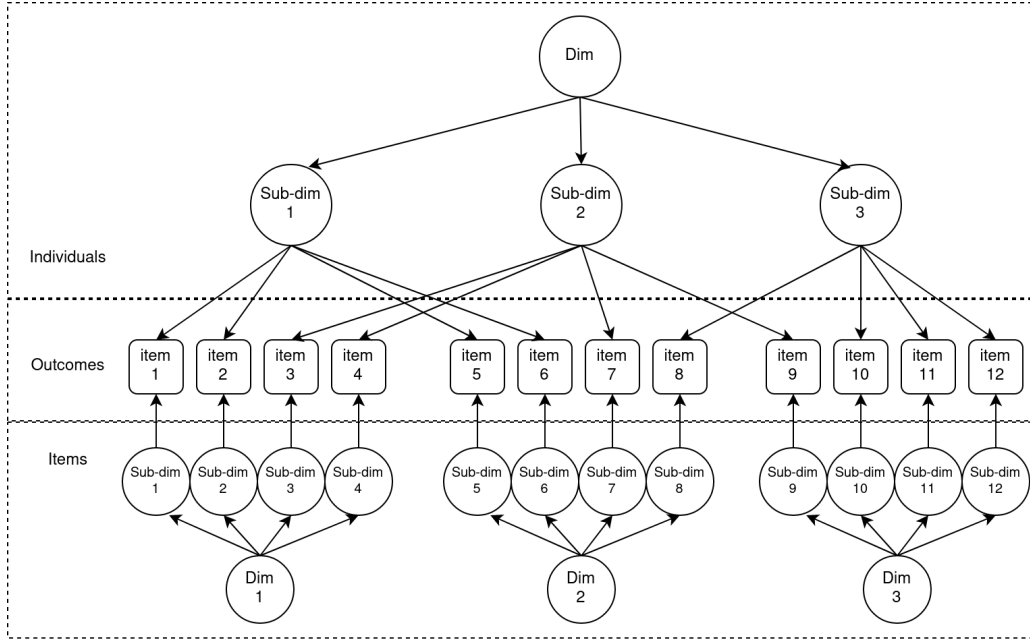


Figure 2.1: Path diagram of the dimensional structure for a hierarchical cross-classified IRT model. Squares represent dichotomous manifest variables, and circles represent latent variables. The figure is based on a reduced set of items, while the errors and scales of the latent variables are not represented. Different sub-dimensions at the individuals block represent the literal, inferential and reflective abilities, while at the items blocks represent the items' difficulties. The dimensions at the individuals block represent the reading comprehension ability, while at the items block represent the multiple testlets.

## 2.2 Model definition

Following Rabe-Hesketh et al. [37, 39], we continue defining the GLLAMM in two parts: (i) the response model, and (ii) the latent structure.

In case the reader is interested in outcomes different than the dichotomous case, refer to Appendix A.1.

### 2.2.1 Response model

Conditional on the regression parameters, explanatory and latent variables  $(\mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\Lambda}, \boldsymbol{\Theta})$ , the response model can be represented by a Generalized Linear Model (GLM) [31, 28] with a systematic and distributional part. The former composed of a linear predictor and a link function.

For the distributional part, the dichotomous items  $y_{jkd}$  are modeled at the first level by a Bernoulli probability mass function  $f(\cdot)$ , in the following form:

$$f(y_{jkd} = 1 \mid \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\Lambda}, \boldsymbol{\Theta}) = p_{jkd}^n (1 - p_{jkd})^{1-n} \quad (2.1)$$

where  $n$  denotes the endorsement of the item in the Bernoulli trial. On the other hand, in the systematic part, the probability of endorsing the item  $p_{jkd}$  is linked to a linear predictor  $v_{jkd}$  through an inverse-link function  $h(\cdot)$ , in the following form:

$$P[y_{jkd} = 1 \mid \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\Lambda}, \boldsymbol{\Theta}] = p_{jkd} = h(\tau_k + v_{jkd}) \quad (2.2)$$

where  $\tau_k$  is  $k$ 'th item threshold (usually assumed at zero). Furthermore, the inverse-link function can be defined in three ways:

$$h(x) = \begin{cases} \exp(x)[1 + \exp(x)]^{-1} \\ \Phi(x) \\ \exp(-\exp(x)) \end{cases} \quad (2.3)$$

corresponding to the logistic, standard normal  $\Phi(x)$ , and Gumbel (extreme value type I) cumulative distributions, respectively. It is usual to report the last in terms of link functions  $g(\cdot) = h^{-1}(\cdot)$ . In that case, the link functions corresponds to the well known logit, probit and complementary log-log, respectively. Finally, the linear predictor is defined by:

$$v_{jkd} = \sum_{q=2}^Q x_{jq}\beta_q + \sum_{m=2}^{M+1} \sum_{k=1}^{K(m)} \eta_k^{(m)} \alpha_k^{(m)} + \sum_{l=2}^{L+1} \sum_{d=1}^{D(l)} \theta_{jd}^{(l)} \lambda_d^{(l)} \quad (2.4)$$

where individuals are indexed by  $j = 1, \dots, J$ , and  $J$  represents the total number of individuals in the sample.  $\beta_q$  denotes regression parameter for the  $x_{jq}$  explanatory variable with  $q = 1, \dots, Q$ , and  $Q$  denoting the total number of explanatory variables.  $\eta_k^{(m)}$  is the  $k$ th item latent dimension at level  $m$  with loading  $\alpha_k^{(m)}$ , where  $k = 1, \dots, K(m)$ ,  $K(m)$  denotes the number of dimensions at level  $m = 2, \dots, M + 1$ , and  $M$  represents the number of levels in the items block.  $\theta_{jd}^{(l)}$  is the individual's  $d$ th latent dimension at level  $l$  with loading  $\lambda_d^{(l)}$ , where  $d = 1, \dots, D(l)$ ,  $D(l)$  represents the number of dimensions at level  $l = 2, \dots, L + 1$ , and  $L$  denotes the number of levels observed in the individuals block. Using figure 2.1 as a reference, we have  $Q = 0$ , as we do not have explanatory variables;  $M = 2$  levels in the items block, where  $K_2 = 12$  corresponding to the items difficulties, and  $K_3 = 3$  corresponding to the testlets; and  $L = 2$  levels in the individuals block, with  $D_2 = 3$  corresponding to the literal, inferential and reflective abilities, and  $D_3 = 1$  corresponding to the reading comprehension latent variable.

Equation 2.4 can be re-written in matrix form in the following way:

$$v_{jkd} = \mathbf{X}_j \boldsymbol{\beta} + \sum_{m=2}^{M+1} \boldsymbol{\eta}^{(m)} \boldsymbol{\alpha}^{(m)} \mathbf{A}_j^{(m)} + \sum_{l=2}^{L+1} \boldsymbol{\theta}_j^{(l)} \boldsymbol{\lambda}^{(l)} \mathbf{B}_j^{(l)} \quad (2.5)$$

where  $\mathbf{X}_j$  represents the individual's design matrix of explanatory variables that maps the parameter vector  $\boldsymbol{\beta}$  to the linear predictor. Moreover,  $\boldsymbol{\eta}^{(m)} = [\eta_1^{(m)}, \dots, \eta_{K(m)}^{(m)}]^T$ , and  $\boldsymbol{\alpha}^{(m)} = [\alpha_1^{(m)}, \dots, \alpha_{K(m)}^{(m)}]^T$  are the vectors of the item's latent dimensions with corresponding loadings, mapped by a block matrix  $\mathbf{A}_j^{(m)}$ . Similarly,  $\boldsymbol{\theta}_j^{(l)} = [\theta_{j1}^{(l)}, \dots, \theta_{jD(l)}^{(l)}]^T$ , and  $\boldsymbol{\lambda}^{(l)} = [\lambda_1^{(l)}, \dots, \lambda_{D(l)}^{(l)}]^T$  are the vectors of the individuals's latent dimensions with corresponding loadings, mapped by a block matrix  $\mathbf{B}_j^{(l)}$ .

Notice the model departs from the traditional multivariate framework for formulating structural models, i.e. a wide data format, where the subject's repeated outcomes are stored in a single row, with multiple response vectors, and explanatory variables appended column-wise to the outcome data; and adopts a univariate approach, i.e. a long data format, where the subject's repeated outcomes are stored in a single "stacked" response

vector with as many rows as there are repeated measurements, and explanatory variables appended column-wise to the outcome data, distinguished from each other, by a design matrix.

Finally, as it was declared at the beginning of the section, we will use  $\mathbf{X} = [\mathbf{X}_1^T, \dots, \mathbf{X}_J^T]^T$  to represent the “stacked” design matrix of explanatory variables for all individuals.  $\mathbf{\Lambda} = [\boldsymbol{\alpha}^{(1)T}, \dots, \boldsymbol{\alpha}^{(M+1)T}, \boldsymbol{\lambda}^{(1)T}, \dots, \boldsymbol{\lambda}^{(L+1)T}]^T$  as the “stacked” vector of loadings, and  $\mathbf{\Theta} = [\boldsymbol{\eta}^{(1)T}, \dots, \boldsymbol{\eta}^{(M+1)T}, \boldsymbol{\theta}^{(1)T}, \dots, \boldsymbol{\theta}^{(L+1)T}]^T$  as the “stacked” vectors of dimensions and sub-dimensions.

### Cluster effects

Considering the previous, modeling clustering is as easier as to add more random effects to the linear predictor defined in equation 2.4, in the following form:

$$\begin{aligned} v_{jkd} &= v_{jkd} + \sum_{c=1}^C \delta_c \\ &= v_{jkd} + \boldsymbol{\delta} \mathbf{Z}_j \end{aligned} \quad (2.6)$$

where  $c = 1, \dots, C$ , which denotes the number of clusters,  $v_{jkd}$  is defined as in equation 2.4, and  $\mathbf{Z}_j$  is a block matrix.

### 2.2.2 Latent structure

The structural model for the latent variables has the form:

$$\underset{(M \times M)(M \times 1)}{\boldsymbol{\eta}} = \underset{(M \times M)(M \times 1)}{\mathbf{B}} \underset{(M \times 1)}{\boldsymbol{\eta}} + \underset{(M \times Q)(Q \times 1)}{\mathbf{\Gamma}} \underset{(M \times 1)}{\mathbf{W}} + \underset{(M \times 1)}{\boldsymbol{\zeta}} \quad (2.7)$$

where  $\mathbf{B}$  and  $\mathbf{\Gamma}$  are parameter matrices that maps the relationship between the latent variables  $\boldsymbol{\eta}$ , and the vector of “stacked” covariates  $\mathbf{W}$ , respectively;  $\boldsymbol{\zeta}$  is a vector of errors or disturbances, and  $M = \sum_l M_l$ . Notice that while equation (2.7) resembles to single-level structural equation models, the main difference lies in the fact that the latent variables may vary at different levels. Additionally, considering that  $\boldsymbol{\eta}$  has no feedback effects, and it is permuted and sorted according to the levels,  $\mathbf{B}$  is defined as a strictly upper triangular matrix. In this regard, it is important to mention that,

1. The absence of feedback loops implies that the method deals with non-recursive models, i.e. none of the latent variables are specified as both causes and effects of each other [26]; **this in turn allows the easy estimation of the model parameters.**
2. The strictly upper triangular structure reveals that the framework does not allow latent variables to be regressed on lower level latent or observed variables, as such specification is more related to the use of formative, rather than reflective, latent variables. For a detail explanation on the topic refer to Edwards and Bagozzi [9].

Notice, however, the previous restrictions does not hinder the ability of the method to model contextual effects, after controlling the lower level compositional effects. For examples of such refer to Appendix A.



# Chapter 3

## The Bayesian estimation

The practical use of GLLAMM requires the estimation of the parameters associated with the items and the individuals' latent abilities. These can be obtained within two frameworks: the classical (frequentist), and the bayesian. The current chapter center its attention on describing the bayesian framework using the Markov Chain Monte Carlo method (MCMC). For a full development of GLLAMM under the frequentist estimation framework refer to Rabe-Hesketh et al. [37, 39], Skrondal and Rabe-Hesketh [47], Rabe-Hesketh et al. [40].

### 3.1 Benefits and shortcomings

The reasons on why bayesian statistics is attractive to perform the estimation of the parameters of any model, and especially under the GLLAMM framework, are:

1. The bayesian estimates are at least as good as the frequentist estimates [2, 53, 22].
2. It is built on a simulation-based estimation method, therefore, it can handle all kinds of priors and data-generating processes [11]. This is especially useful with highly complex and over-parameterized models, where other methods are unfeasible or work poorly [2, 25].
3. The model definitions, i.e. the likelihood for the data and priors for the parameters, are used to estimate the corresponding posterior distributions. However, the definitions can also be used in a generative way, i.e. simulate observations, allowing us to test the ability of the method/data to recover the parameters of interest [29].
4. It allow us to integrate prior beliefs or knowledge about the parameters beyond the observed responses [11, 47]. This is especially useful when we have issues of non-convergence or improper estimation of the parameters under the Maximum Likelihood methods (ML). Examples of these cases are:
  - (a) Estimating abilities when individuals have null scores or aberrant response patterns, i.e. examinees that answered some relatively difficult and discriminating items correctly, while answering some of the easiest incorrectly. [20, 1].

- (b) Estimating parameters that need to be confined to a permitted parameter space, e.g. the estimation of positive unique factors variances, where the opposite is known as ‘Heywood cases’ [27]
- (c) Estimating parameters under a sparse data structure, where the asymptotic theory is unlikely to hold [11];

Finally, in terms of shortcomings, the bayesian framework has the following inconveniences:

1. It exposes the user to arbitrary” decisions about the running of the chains, e.g. how many iterates does the chain need to achieve precise estimates?, what is the right size for the burn-in and warm-up phases?, how should the thinning procedure be performed? [47].
2. The user has many options to assess if the chain achieves stationarity, convergence or good mixing, and most of them are visual. This makes it hard to assess if the chain converges to a proper distribution [18].
3. The procedure makes it hard to discover parameters’ lack of identification [47]. Inadequate mixing of the chain could lead us to think unidentified parameters have been estimated with precision, when in fact they have a ‘flat’ posterior [24].
4. Sometimes the geometry of the model makes it hard to find proper solutions for the parameter space. This is especially true in hierarchical models. Under this circumstances, the scientist needs to re-parameterize the model to a non-centered form, i.e. remove the dependence of the parameters on other sampled parameters [19]. In those cases, the complexity of the transformation limit the ability of the scientist to communicate/share the implementation [29].
5. The procedure requires more time to achieve a proper solution, compared to the classical methods. This is especially true in models with high complexity [51, 44].

Although some of the shortcomings has made the use of bayesian methods a ”controversial” issue, most of these already have an acceptable solution.

For the first point, a popular approach to solve the issues is to use a large number of iterates, or multiple chains with different initial states. This is mostly applicable under the Metropolis-Hastings and Gibbs sampling methods. However, as we will see in section 3.1, the Hamiltonian Monte Carlo method (HMC) [4] implements a different sampling mechanism that is less reliant on these decisions.

About the second shortcoming, it is well accepted that the visual assessment of stationarity and convergence is easier, and this procedure usually has additional support from statistics like  $\hat{R}$  [17]. On the contrary, a visual evaluation of ‘good’ mixing remains as a hard task. A popular approach to increase the possibility of a well mixed chain is to change the geometry of the model [29]. However, the implementation of the approach does not necessarily ensure the required property.

On the third point, the most common solution is to use regularizing priors, i.e. priors that are more ‘skeptical’ of wider parameter spaces [29]. However, it is important to mention, there are scenarios where one can achieve poor parameter estimates, even in

the presence of ‘enough’ data and regularizing priors, e.g. the estimation of the variance parameters in random effects models [47], but this is also applicable to the classical estimation procedures.

Finally, the fourth and fifth points can be considered as the ‘price’ a scientist has to pay to be able to fit complex models, that are in more accordance with the observed data generating processes.

## 3.2 The Bayesian GLLMM for dichotomous outcomes

### 3.2.1 Prior distribution

### 3.2.2 Initial start

### 3.2.3 Likelihood

### 3.2.4 Posterior distribution

### 3.2.5 Model identification

(work in progress)

The structure of the latent variables is specified by the number of levels  $L$  and the number of latent variables  $M_l$  at each level. A particular level may coincide with a level of clustering in the hierarchical dataset. However, there will often not be a direct correspondence between the levels of the model and the levels of the data hierarchy.

### 3.2.6 Distribution of the latent variables

Finally, to fully specify the framework, and provide a scale for the latent variables, we have to make assumptions for either the distribution of the disturbances  $\zeta$  or the latent variables  $\eta$ . If our research interest lies in the structural equation model, it is more convenient to make assumptions for the distribution of the disturbances; otherwise, we make assumptions for the distributions for the latent variables.

Furthermore, as in the hierarchical framework, it is assumed the latent variables at different levels are independent, whereas latent variables at the same level may present dependency. In that sense, we presume all latent variables at level  $l$  to have a multivariate normal distribution with zero mean and covariance matrix  $\Sigma_l$ , i.e.  $\eta^{(l)} \sim MVN(\mathbf{0}, \Sigma_l)$ . It is important to emphasize that, while the multivariate normal distribution is widely used in these settings, it is not the only distribution that can be assumed. Rabe-Hesketh, Skrondal and Pickles [36] have provided evidence that it can be even left unspecified, by using non-parametric maximum likelihood estimation.

## 3.3 Computational implementation

(work in progress)

see

- Gelman et al (2011) - Handbook of Markov Chain Monte Carlo
- McElreath (2020) - Statistical Rethinking

Rethinking: Warmup is not burn-in. Other MCMC algorithms and software often discuss burn-in. With a sampling strategy like ordinary Metropolis, it is conventional and useful to trim off the front of the chain, the “burn-in” phase. This is done because it is unlikely that the chain has reached stationarity within the first few samples. Trimming off the front of the chain hopefully removes any influence of which starting value you chose for a parameter. 156 But Stan’s sampling algorithms use a different approach. What Stan does during warmup is quite different from what it does after warmup. The warmup samples are used to adapt sampling, to find good values for the step size and the number of steps. Warmup samples are not representative of the target posterior distribution, no matter how long warmup continues. They are not burning in, but rather more like cycling the motor to heat things up and get ready for sampling. When real sampling begins, the samples will be immediately from the target distribution, assuming adaptation was successful.

The procedure will be with the aid of Stan [50] and R [34, 49] to retrieve .

## Chapter 4

### Simulation Study

# Chapter 5

## Application

### 5.1 Instruments

### 5.2 Data

#### 5.2.1 Collection

#### 5.2.2 Sample scheme

### 5.3 Results

# Chapter 6

## Conclusion and Discussion

6.1 Discussion

6.2 Conclusions

6.3 Future development

# Appendix A

## Additional Theory

### A.1 Other links and distributions

As stated in chapter 2, the GLLAMM is a unifying framework for a wide range of latent models. This is possible thanks to the flexibility in the modeling of the outcomes, coming from defining the response model as a GLM [28] (see section 2.2).

In the current section the author will try to briefly describe the most important link functions and outcomes distributions, that can be accommodated within the framework. For a more detailed approach on either of these, refer to Rabe-Hesketh et al. [37, 39, 38], Skrondal and Rabe-Hesketh [47], and Rabe-Hesketh et al. [40].

#### 1. Continuous:

It results from selecting an identity link function for the scaled mean response,

$$\begin{aligned}\mu^* &= E[y^*|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\ &= v\end{aligned}\tag{A.1}$$

where  $\mu^* = \mu\sigma^{-1}$ ,  $y^* = y\sigma^{-1}$ , and  $\sigma$  denotes the standard deviation of the errors.

On the other hand, the distributional part is defined by a Standard Normal distribution  $\phi(x) = (2\pi)^{-1/2}\exp(-x^2/2)$ ,

$$\begin{aligned}f(y^*|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}) &= \phi(\mu^*)\sigma^{-1} \\ &= \phi(v)\sigma^{-1}\end{aligned}\tag{A.2}$$

Notice that the same parametrization can be achieved considering  $y^* = v + \epsilon^*$ , and  $\epsilon^* \sim N(0, 1)$ . Additionally, the decision to standardize the response variables has been made with the purpose of making the estimation process easier, as such distribution is free of unknown parameters.

#### 2. Polytomous:

It results from selecting a generalized logistic inverse-link function [5] for the expected value of the response, which in this case, describe the probability of endorsing



one of the  $S$  unordered available categories,

$$\begin{aligned}
 \mu_s &= E[y = y_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= P[y = y_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= \pi_s \\
 &= h(v_s)
 \end{aligned} \tag{A.3}$$

where  $v_s$  is the linear predictor for category  $s$  ( $s = 1, \dots, S$ ), and  $h(\cdot)$  is defined as:

$$h(x) = \exp(x) \cdot \left[ \sum_{s=1}^S \exp(x) \right]^{-1} \tag{A.4}$$

It is important to note that, as in the dichotomous case, the same parametrization can be achieved using the concept of underlying continuous responses in the form  $y_s^* = v_s + \epsilon_s$ , where  $y = s$  if  $y_s^* > y_k^* \forall s, s \neq k$ ,  $\epsilon_s$  have a Gumbel (extreme value type I) distribution, as the one defined in equation (2.3), and  $y_s$  denotes the random utility for the  $s$  category.

Finally, the distributional part is defined by a Multinomial distribution,

$$\begin{aligned}
 f[y = \{y_1, \dots, y_S\} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] &= \frac{n!}{y_1! \dots y_S!} \prod_{s=1}^S \mu_s^{y_s} \\
 &= \frac{n!}{y_1! \dots y_S!} \prod_{s=1}^S \pi_s^{y_s}
 \end{aligned} \tag{A.5}$$

where  $y_s$  denotes the number of "successes" in category  $s$ .

### 3. Ordinal and discrete time duration:

For the ordinal case, the linear predictor is "linked" to the probability of endorsing category  $s$ , against all previous categories, in the following form:

$$\begin{aligned}
 \mu_s &= E[y = y_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= P[y \leq y_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] - P[y \leq y_{s-1} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= h(\kappa_s - v_s) - h(\kappa_{s-1} - v_{s-1})
 \end{aligned} \tag{A.6}$$

where  $\kappa_s$  denotes the thresholds for category  $s$ . For discrete time duration, the linear predictor is "linked" to the probability of survival, in the  $s$ th time interval, as follows:

$$\begin{aligned}
 \mu_s &= E[t_{s-1} \leq T \leq t_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= P[T \leq t_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] - P[T \leq t_{s-1} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\
 &= h(v_s + t_s) - h(v_{s-1} + t_{s-1})
 \end{aligned} \tag{A.7}$$

where  $T$  is the unobserved continuous time, and  $t_s$  its observed discrete realization. Additionally, for both type of responses,  $h(\cdot)$  can be defined as the logistic, standard normal, and Gumbel (extreme value type I) *cumulative distributions*, as in equation (2.3).

Similar to the dichotomous and polytomous case, the same parametrization can be achieved using the concept of underlying latent variables with  $y_s^* = v_s + \epsilon_s$ , where  $y = s$  if  $\kappa_{s-1} < y_s^* \leq \kappa_s$ ,  $\kappa_0 = -\infty$ ,  $\kappa_1 = 0$ ,  $\kappa_S = +\infty$ ,  $\epsilon_s$  has one of the distributions in equation (2.3), and  $y_s$  denotes the random utility for the  $s$  category.

It is important to note, for discrete time duration responses, the logit link corresponds to a *Proportional-Odds model*, while the complementary log-log link to a *Discrete Time Hazards model* [41]. Other models for ordinal responses, such as the *Baseline Category Logit* or the *Adjacent Category Logit* models can be specified as special cases of the generalized logistic response function, defined in equation (A.4).

Finally, the distributional part is defined by a Multinomial distribution, as the one defined in equation (A.5).

#### 4. Counts and continuous time duration:

It results from selecting an exponential inverse-link function (log link) for the expected value of the response,

$$\begin{aligned}\mu &= E[y|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\ &= \lambda \\ &= \exp(v)\end{aligned}\tag{A.8}$$

and a Poisson conditional distribution for the counts,

$$\begin{aligned}f[y|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] &= \exp(-\mu)\mu^y(y!)^{-1} \\ &= \exp(-\lambda)\lambda^y(y!)^{-1}\end{aligned}\tag{A.9}$$

It is important to mention that unlike the models for dichotomous, polytomous and ordinal responses, model for counts cannot be written under the random utility framework.

#### 5. Rankings and pairwise comparisons:

Following Skrandal and Rabe-Hesketh [45], the parametrization for polytomous responses can serve as the building block for the conditional distribution of rankings. Selecting a "exploded logit" inverse-link function [7] for the expected value of the response, which describes the probability of the full rankings of category  $s$ ,

(work in progress)

$$\begin{aligned}\mu_s &= P[\mathbf{R}_s = \{r_s^1, \dots, r_s^1\}|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\ &= \pi_s \\ &= h(v_s)\end{aligned}\tag{A.10}$$

where  $v_s$  is the linear predictor for category  $s$  ( $s = 1, \dots, S$ ), and  $h(\cdot)$  is defined as:

$$h(x) = \prod_{s=1}^S \exp(x^s) \left[ \sum_{s=1}^S \exp(x^s) \right]^{-1}\tag{A.11}$$

Again, as in specific previous cases, the same parametrization can be achieved using the concept of underlying latent variables.

Finally, the distributional part is defined by a Multinomial distribution,

$$\begin{aligned} f[y = \{y_1, \dots, y_S\} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] &= \frac{n!}{y_1! \dots y_S!} \prod_{s=1}^S \mu_s^{y_s} \\ &= \frac{n!}{y_1! \dots y_S!} \prod_{s=1}^S \pi_s^{y_s} \end{aligned} \quad (\text{A.12})$$

where  $y_s$  denotes the number of "success cases" in category  $s$ .

## 6. Mixtures:

Given the previous definitions, the framework easily lends itself to model five additional settings:

- (a) **Different links and distributions for different latent variables.** This can be easily achieved by setting different links and distributions for each of the  $M_2$  latent variables located at level 2.
- (b) **Left- or right-censored continuous responses.** Common in selection models (e.g. Heckman [21]), they can be achieved by specifying an identity link and Normal distribution for the uncensored scaled responses, as in equations (A.1) and (A.2); and a scaled probit link and Binomial distribution otherwise, as in equations (2.3) and (??).
- (c) **zero-inflated count responses.** where a log link and a Poisson distribution is set for the counts, as in equations (A.8) and (A.9); and a logit link and Binomial distribution is specified to model the zero center of mass, as in equations (??) and (??).
- (d) **Measurement error in covariates.** this setting occurs when standard models use variables, with measurement error, as covariates, e.g. a logistic regression with a continuous covariate that presents measurement error. For more details on this type of setting see Rabe-Hesketh, Skrondal and Pickles [36], Rabe-Hesketh, Pickles and Skrondal [35], and Skrondal and Rabe-Hesketh [46].
- (e) **Composite links.** Useful for specifying proportional odds models for right-censored responses, for handling missing categorical covariates and many other model types. For more details on this type of settings see Skrondal and Rabe-Hesketh [48].

## Heteroscedasticity and over-dispersion in the response

Much like the Generalized Linear Mixed Model framework (GLMM), the GLLAMM allows to model heteroscedasticity, and over- or under-dispersion by adding random effects to the linear predictor, at level 1. The types of responses, in which such characteristics can be modeled, are the following:

### 1. Continuous:

We model **heteroscedasticity** in the following form:

$$\sigma = \exp(\boldsymbol{\alpha}^T \mathbf{Z}^{(1)}) \quad (\text{A.13})$$

Notice that the previous formula implies that equation (A.2) can be re-written in the following form:

$$f(y^*|\mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}) = \phi(v + \boldsymbol{\alpha}^T \mathbf{Z}^{(1)}) \quad (\text{A.14})$$

where  $\mathbf{Z}^{(1)}$  is the design matrix that maps the random effects  $\boldsymbol{\alpha}$ . Notice that equation (A.14) effectively corresponds to a model that includes random intercepts at level 1.

## 2. Ordinal, and discrete time duration:

Similar to the dichotomous case, by including random intercepts at level 1 in equation (A.6), we can model over- or under-dispersion:

$$\begin{aligned} \mu_s &= P[y \leq y_s | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] - P[y \leq y_{s-1} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\ &= h(\kappa_s - v_s + \boldsymbol{\alpha}^T \mathbf{Z}^{(1)}) - h(\kappa_{s-1} - v_{s-1} + \boldsymbol{\alpha}^T \mathbf{Z}^{(1)}) \end{aligned} \quad (\text{A.15})$$

A similar parametrization can be used for discrete time duration.

## 3. Counts, and continuous time duration:

Finally, modifying equation (A.8) allow us to model over- or under-dispersion under a counts model:

$$\begin{aligned} \mu &= E[y | \mathbf{X}, \mathbf{Z}, \boldsymbol{\eta}] \\ &= \lambda \\ &= \exp(v + \boldsymbol{\alpha}^T \mathbf{Z}^{(1)}) \end{aligned} \quad (\text{A.16})$$

## A.2 Sampling scheme

# Appendix B

## Code

# Bibliography

- [1] Azevedo, C. [2003]. *Métodos de estimação na teoria de resposta ao item*, Master's thesis, Universidade de São Paulo (USP).  
**url:** <https://teses.usp.br/teses/disponiveis/45/45133/tde-05102004-163906/pt-br.php>.
- [2] Baker, F. [1998]. An investigation of the item parameter recovery characteristics of a gibbs sampling procedure, *Applied Psychological Measurement* **22**(22): 153–169.  
**doi:** <https://doi.org/10.1177/01466216980222005>.
- [3] Baker, F. [2001]. The basic of item response theory, *Technical report*, ERIC Clearinghouse on Assessment and Evaluation.
- [4] Betancourt, M. and Girolami, M. [2012]. Hamiltonian monte carlo for hierarchical models.  
**url:** [arxiv.org/abs/1312.0906v1](http://arxiv.org/abs/1312.0906v1).
- [5] Bock, R. [1972]. Estimating item parameters and latent ability when responses are scored in two or more nominal categories, *Psychometrika* **37**(1).  
**doi:** <https://doi.org/10.1007/BF02291411>.
- [6] Bradlow, E., Wainer, H. and Wang, X. [1999]. A bayesian random effects model for testlets, *Psychometrika* **64**(2): 153–168.  
**doi:** <https://doi.org/10.1007/BF02294533>.
- [7] Chapaaan, R. and Staelin, R. [1982]. Exploiting rank ordered choice set data within the stochastic utility model, *Journal of Marketing Research* **19**(3): 288–301.  
**doi:** <https://www.doi.org/10.1177/002224378201900302>.
- [8] Chen, W. and Thissen, D. [1997]. Local dependence indexes for item pairs using item response theory, *Journal of Educational and Behavioral Statistics* **22**(3): 265–289.  
**doi:** <https://doi.org/10.3102/10769986022003265>.
- [9] Edwards, J. and Bagozzi, R. [2000]. On the nature and direction of relationships between constructs and measures, *Psychological Methods* **5**(2): 155–174.  
**doi:** <https://www.doi.org/10.1037/1082-989X.5.2.155>.
- [10] Flores, S. [2012]. *Modelos testlet logísticos y logísticos de exponente positivo para pruebas de comprensión de textos*, Master's thesis, Pontificia Universidad Católica del Perú.

- [11] Fox, J. [2010]. *Bayesian Item Response Modeling, Theory and Applications*, Statistics for Social and Behavioral Sciences, fienberg, s. and van der linden, w. edn, Springer Science+Business Media, LLC.
- [12] Fujimoto, K. [2018a]. The bayesian multilevel trifactor item response theory model, *Educational and Psychological Measurement* **79**(3): 462–494.  
**doi:** <https://doi.org/10.1177/0013164418806694>.
- [13] Fujimoto, K. [2018b]. A general bayesian multilevel multidimensional irt model for locally dependent data, *Br J Math Stat Psychol* **71**(3): 536–560.  
**doi:** <https://doi.org/10.1111/bmsp.12133>.
- [14] Fujimoto, K. [2020]. A more flexible bayesian multilevel bifactor item response theory model, *Journal of Educational Measurement* **57**(2): 255–285.  
**doi:** <https://doi.org/10.1111/jedm.12249>.
- [15] Gelfand, A., Sahu, S. and Carlin, B. [1995]. Efficient parametrisations for normal linear mixed models, *Biometrika* **82**(3): 479–488.  
**doi:** <https://doi.org/10.1093/biomet/82.3.479>.
- [16] Gelfand, A., Sahu, S. and Carlin, B. [1996]. Efficient parameterizations for generalised linear models (with discussion), in J. Bernardo, J. Berger, A. Dawid and a. Smith (eds), *Bayesian Statistics*, Vol. 5, pp. 165–180.
- [17] Gelman, A., Carlin, J., Stern, H., Dunson, D., Vehtari, A. and Rubin, D. [2014]. *Bayesian Data Analysis*, Texts in Statistical Science, third edn, Chapman and Hall/CRC.
- [18] Gelman, A. and Rubin, D. [1996]. Markov chain monte carlo methods in biostatistics, *Statistical Methods in Medical Research* **5**(4): 339–355.  
**doi:** <https://doi.org/10.1177/096228029600500402>.
- [19] Gorinova, M., Moore, D. and Hoffman, M. [2019]. Automatic reparameterisation of probabilistic programs.  
**url:** <https://arxiv.org/abs/1906.03028>.
- [20] Hambleton, R., Swaminathan, H. and Rogers, H. [1991]. *Fundamentals of Item Response Theory*, SAGE Publications Inc.
- [21] Heckman, J. [1979]. Sample selection bias as a specification error, **47**(1): 153–161.  
**doi:** <https://www.doi.org/10.2307/1912352>.  
**url:** <https://www.jstor.org/stable/1912352>.
- [22] Hsieh, M., Proctor, T., Hou, J. and Teo, K. [2010]. A comparison of bayesian mcmc and marginal maximum likelihood methods in estimating the item parameters for the 2pl irt model, *International Journal of Innovative Management, Information and Production* **1**(1): 81–89.  
**url:** <http://ismeip.org/IJIMIP/contents/imip1011/10IN15T.pdf>.
- [23] Jiao, H., Kamata, A., Wang, S. and Jin, Y. [2012]. A multilevel testlet model for dual local dependence, *Journal of Educational Measurement* **49**: 82–100.  
**doi:** <https://doi.org/10.1111/j.1745-3984.2011.00161.x>.

- [24] Keane, M. [1992]. A note on identification in the multinomial probit model, *Journal of Business and Economic Statistics* **10**(2): 193–200.  
**doi:** <https://doi.org/10.2307/1391677>.  
**url:** <https://www.jstor.org/stable/1391677>.
- [25] Kim, S. and Cohen, A. [1999]. Accuracy of parameter estimation in gibbs sampling under the two-parameter logistic model, *Annual Meeting of the American Educational Research Association*, American Educational Research Association.  
**url:** <https://eric.ed.gov/?id=ED430012>.
- [26] Kline, R. [2012]. Assumptions in structural equation modeling, in R. Hoyle (ed.), *Handbook of Structural Equation Modeling*, The Guilford Press, chapter 7, pp. 111–125.
- [27] Martin, J. and McDonald, R. [1975]. Bayesian estimation in unrestricted factor analysis: A treatment for heywood cases, *Psychometrika* (40): 505–517.  
**doi:** <https://doi.org/10.1007/BF02291552>.
- [28] McCullagh, P. and Nelder, J. [1989]. *Generalized Linear Models*, Monographs on Statistics Applied Probability, Chapman Hall/CRC Press.
- [29] McElreath, R. [2020]. *Statistical Rethinking: A Bayesian Course with Examples in R and Stan*, Texts in Statistical Science, 2 edn, Chapman and Hall/CRC.  
**doi:** <https://doi.org/10.1201/9780429029608>.
- [30] Muthén, L. and Muthén, B. [1998-2011]. *Mplus User's Guide*, CA: Muthén Muthén.
- [31] Nelder, J. and Wedderburn, W. [1972]. Generalized linear models, *Royal Statistical Society* **135**(3): 370–384.  
**doi:** <https://doi.org/10.2307/2344614>.  
**url:** <https://www.jstor.org/stable/2344614>.
- [32] Papaspiliopoulos, O., Roberts, G. and Skold, M. [2003]. Non-centered parameterisations for hierarchical models and data augmentation, *Bayesian Statistics* **7**: 307–326.  
**url:** <http://econ.upf.edu/~omiros/papers/val7.pdf>.
- [33] Papaspiliopoulos, O., Roberts, G. and Skold, M. [2007]. A general framework for the parametrization of hierarchical models, *Statistical Science* **22**(1): 59–73.  
**doi:** <https://www.doi.org/10.1214/088342307000000014>.
- [34] R Core Team [2015]. *R: A Language and Environment for Statistical Computing*, R Foundation for Statistical Computing, Vienna, Austria.  
**url:** <http://www.R-project.org/>.
- [35] Rabe-Hesketh, S., Pickles, A. and Skrondal, A. [2003]. Correcting for covariate measurement error in logistic regression using nonparametric maximum likelihood estimation, *Statistical Modelling* **3**(3): 215–232.  
**doi:** <https://www.doi.org/10.1191/1471082X03st0560a>.



- [36] Rabe-Hesketh, S., Skrondal, A. and Pickles, A. [2003]. Maximum likelihood estimation of generalized linear models with covariate measurement error, *The Stata Journal* **3**(4): 386–411.  
**doi:** <https://www.doi.org/10.1177/1536867X0400300408>.  
**url:** <https://journals.sagepub.com/doi/pdf/10.1177/1536867X0400300408>.
- [37] Rabe-Hesketh, S., Skrondal, A. and Pickles, A. [2004a]. Generalized multilevel structural equation modeling, *Psychometrika* **69**(2): 167–190.  
**doi:** <https://www.doi.org/10.1007/BF02295939>.
- [38] Rabe-Hesketh, S., Skrondal, A. and Pickles, A. [2004b]. *GLLAMM Manual*, UC Berkeley Division of Biostatistics.  
**url:** <http://www.biostat.jhsph.edu/~fdominic/teaching/bio656/software-gllamm.manual.pdf>.
- [39] Rabe-Hesketh, S., Skrondal, A. and Pickles, A. [2004c]. Maximum likelihood estimation of limited and discrete dependent variable models with nested random effects, *Journal of Econometrics* **128**(2): 301–323.  
**doi:** <https://www.doi.org/10.1016/j.jeconom.2004.08.017>.  
**url:** <http://www.sciencedirect.com/science/article/pii/S0304407604001599>.
- [40] Rabe-Hesketh, S., Skrondal, A. and Zheng, X. [2012]. Multilevel structural equation modeling, in R. Hoyle (ed.), *Handbook of Structural Equation Modeling*, The Guilford Press, chapter 30, pp. 512–531.
- [41] Rabe-Hesketh, S., Yang, S. and Pickles, A. [2001]. Multilevel models for censored and latent responses, *Statistical Methods in Medical Research* **10**(6): 409–427.  
**doi:** <https://www.doi.org/10.1177/096228020101000604>.
- [42] Raudenbush, S. and Bryk, A. [2002]. *Hierarchical linear models: Applications and data analysis methods (Vol. 1)*, Advanced Quantitative Techniques in the Social Sciences, SAGE Publications Inc.
- [43] Reckase, M. [2009]. *Multidimensional Item Response Theory*, Statistics for Social and Behavioral Sciences, Springer Science+Business Media, LLC.
- [44] Rivera, J. [2019]. *El modelo de respuesta nominal: Aplicación a datos educacionales*, Master’s thesis, Pontificia Universidad Católica del Peru.  
**url:** <http://hdl.handle.net/20.500.12404/14600>.
- [45] Skrondal, A. and Rabe-Hesketh, S. [2003a]. Multilevel logistic regression for polytomous data and rankings, *Psychometrika* **68**: 267–287.  
**doi:** <https://www.doi.org/10.1007/BF02294801>.
- [46] Skrondal, A. and Rabe-Hesketh, S. [2003b]. Some applications of generalized linear latent and mixed models in epidemiology: Repeated measures, measurement error and multilevel modeling, *Norsk Epidemiologi* **13**(2): 265–278.
- [47] Skrondal, A. and Rabe-Hesketh, S. [2004a]. *Generalized Latent Variable Modeling: Multilevel, Longitudinal, and Structural Equation Models*, Interdisciplinary Statistics, Chapman Hall/CRC Press.

- [48] Skrondal, A. and Rabe-Hesketh, S. [2004b]. Generalized linear latent and mixed models with composite links and exploded likelihoods, *in* Biggeri A., E. Dreassi, C. Lagazio and M. Marchi (eds), *Proceedings of the 19th International Workshop on Statistical Modeling*, Firenze University Press, Florence, Italy, pp. 27–39.  
**url:** [http://www.gllamm.org/composite\\_conf.pdf](http://www.gllamm.org/composite_conf.pdf).
- [49] Stan Development Team [2020a]. RStan: the R interface to Stan. R package version 2.21.2.  
**url:** <http://mc-stan.org/>.
- [50] Stan Development Team. [2020b]. *Stan Modeling Language Users Guide and Reference Manual, version 2.26*, Vienna, Austria.  
**url:** <https://mc-stan.org>.
- [51] Tarazona, E. [2013]. *Modelos alternativos de respuesta graduada con aplicaciones en la calidad de servicios*, Master’s thesis, Pontificia Universidad Católica del Perú (PUCP).  
**url:** <http://hdl.handle.net/20.500.12404/6175>.
- [52] Wainer, H., Bradlow, E. and Wang, X. [2007]. *Testlet response theory and its applications*, Cambridge University Press.
- [53] Wollack, J. A., Bolt, D. M., Cohen, A. S. and Lee, Y.-S. [2002]. Recovery of item parameters in the nominal response model: A comparison of marginal maximum likelihood estimation and markov chain monte carlo estimation, *Applied Psychological Measurement* **26**(3): 339–352.  
**doi:** <https://www.doi.org/10.1177/0146621602026003007>.
- [54] Yen, W. [1984]. Effects of local item dependence on the fit and equating performance of the three-parameter logistic model, *Applied Psychological Measurement* **8**(2): 125–145.  
**doi:** <https://doi.org/10.1177/014662168400800201>.

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