

- Chp 4 Exercise 4.3.12: [Note  $M_n = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ . Use R to generate the random variable w/resp.] Generate  $X_1, \dots, X_n$  i.i.d.  $\text{Exp}(5)$  and compute  $M_n$  when  $n=20$ . Repeat this  $N=10^5$  times, and compute the proportion of values  $M_n$  that lie between  $0.19 \leq 0.21$ . Repeat this w/  $n=50$ . What property of convergence in probability do your results illustrate?
- $P(0.19 \leq M_{20} \leq 0.21) \approx 0.1789$ ,  $P(0.19 \leq M_{50} \leq 0.21) \approx 0.27721$
  - This result illustrates the weak LLN.

Exercise 4.4.4: let  $W_n$  have density  $\frac{1+x/n}{1+1/2n}$  for  $0 < x < 1$ ; 0 o.w.  
let  $W \sim \text{Unif}(0,1)$ . Prove  $W_n$  converges in distribution to  $W$ .

DEF 4.4.1: let  $X_1, X_2, \dots, X_n$  be random variables. Then we say the sequence  $\{X_n\}$  converges in distribution to  $X$ , if  $\forall x \in \mathbb{R}$  s.t.  $P(X=x)=0$  we have  $\lim_{n \rightarrow \infty} P(X_n \leq x) = P(X \leq x)$  and we write  $X_n \xrightarrow{D} X$

$$\lim_{n \rightarrow \infty} W_n = \lim_{n \rightarrow \infty} \frac{1+x/n}{1+1/2n} = 1 = W$$

- Chp 4 Exercise 4.4.12: Add (d-f): Determine the exact distribution of the average time to service for the first  $n$  customers when  $n=16, 36, 100$ . Then use pgamma function in R to find the exact probability and compare to the normal approx.
- If the service time, in minutes at a bank has  $\text{Exp}(1/2)$  distribution. Use the cdf to estimate the probability that the service time of the 1st  $n$  customers is less than 2.5 minutes when: (NOTE: let  $X$  be the variable of interest)

(a)  $n=16$ .

$$P(\bar{X}_{16} \leq 2.5) \approx 0.84265$$

(b)  $n=36$

$$P(\bar{X}_{36} \leq 2.5) \approx 0.92641$$

(c)  $n=100$

$$P(\bar{X}_{100} \leq 2.5) \approx 0.99037$$

3.) (Contd)

NOTE: IF  $Y = \sum_{i=1}^n X_i$  where  $X_i$  iid  $\text{Exp}(\lambda)$

$$\text{Then } M_Y(s) = M_{X_1 + \dots + X_n}(s) = E[e^{s(X_1 + \dots + X_n)}] = \prod_{i=1}^n E[e^{sX_i}]$$

$$= (\lambda/(\lambda - s))^n$$

$$\Rightarrow Y \sim \text{Gamma}(n, \lambda)$$

$$\textcircled{1} \bar{X} = \frac{1}{n} Y \Rightarrow Y = n\bar{X} = n\bar{x}; \quad \frac{d}{dx}[n^{-1}(x)] = n$$

$$f_{\bar{X}}(\bar{x}) = \left( \frac{\lambda^n}{\Gamma(n)} \right) (n\bar{x})^{n-1} e^{-\lambda(n\bar{x})} \quad (n)$$

$$= \left( \frac{\lambda^n}{\Gamma(n)} \right) n^n \bar{x}^{n-1} e^{-(\lambda n)\bar{x}}$$

$$f_{\bar{X}}(\bar{x}) = \left( \frac{(\lambda n)^n}{\Gamma(n)} \right) \bar{x}^{n-1} e^{-(\lambda n)\bar{x}}$$

$$\Rightarrow \bar{X} \sim \text{Gamma}(n, \lambda n)$$

$$(d) n=16 \Rightarrow \bar{X} \sim \text{Gamma}(16, 8)$$

$$P(\bar{X} \leq 2.5) = 0.8434869 \text{ which is close to our approximation using the CLT of } 0.84265$$

$$(e) n=36 \Rightarrow \bar{X} \sim \text{Gamma}(36, 18)$$

$$P(\bar{X} \leq 2.5) = 0.9257825 \text{ which is close to our approximation using the CLT of } 0.92641$$

$$(f) n=100 \Rightarrow \bar{X} \sim \text{Gamma}(100, 50)$$

$$P(\bar{X} \leq 2.5) = 0.9906209 \text{ which is close to our approximation using the CLT of } 0.99037$$

4.) Chp 4 Exercise 4.4.16:

Generate  $N = 10^4$  samples of  $X_1, X_2, \dots, X_{30}$  Unif $[-20, 10]$ .

Use these samples to estimate the probability  $P(M_{30} \leq -5)$ . How does your answer compare to what the CLT gives as an approximation?

•  $P(\bar{X} \leq -5) \approx 0.505$  according to our simulation. This is close to the CLT approximation of 0.50.

5.) Chp 4 Exercises 4.6.1, 4.6.2, 4.6.7

Chp 4 Exercise 4.6.1:

Let  $X_1 \sim N(3, 2^2)$  and  $X_2 \sim N(-8, 5^2)$  be independent. Let

$U = X_1 - 5X_2$  and  $V = -6X_1 + CX_2$  where  $C$  is a constant.

(a) What are the distributions of  $U, V$ ? (H.O.4 pg 30)\*

$$U \sim N(43, 629)$$

$$V \sim N(-14 - 8C, 144 + 25C^2)$$

(b) What is the value of  $C$  that makes  $U, V$  independent? (H.O.4 pg 31)\*

NOTE Exercise 4.6.2: If  $X_i \sim N(\mu_i, \sigma_i^2)$  for  $i = 1, 2, \dots, n$  and also that the  $\{X_i\}$  are independent. Let  $U = \sum_{i=1}^n a_i X_i$  and  $V = \sum_{i=1}^n b_i X_i$  for some constants  $\{a_i\}$  &  $\{b_i\}$ . Then  $\text{Cov}(U, V) = \sum_{i=1}^n a_i b_i \sigma_i^2$ . Furthermore,  $\text{Cov}(U, V) = 0 \Leftrightarrow U, V$  are independent

$$\text{Cov}(U, V) = 1(-6)(2^2) + (-5)(C)(5^2) = 0$$

$$= -24 - 125C = 0$$

(Contd.)

Chp 4 Exercise 4.6.2: Let  $X \sim N(3, 5)$ ,  $Y \sim N(-7, 2)$  be independent

(a) write the distribution of  $Z = 4X - 13Y$  \* (H.O.4 pg 30)

$$Z \sim N(4(3) + (-13)(-7), (4^2)(5) + (-13)^2(2))$$

$$Z \sim N(43/3, 722/3)$$

(b) What is  $\text{cov}(X, Z)$ ? (see H.O.3 slide 34) (also Thm 3.3.2)

Theorem 3.3.2 (Linearity of Covariance)

Let  $X, Y, Z$  be three random variables. Let  $a, b \in \mathbb{R}$ . Then

$$\text{Cov}(aX + bY, Z) = a \text{Cov}(X, Z) + b \text{Cov}(Y, Z)$$

$$\text{Cov}(X, Z) = \text{Cov}(X, 4X - 13Y) = 4 \text{Cov}(X, X) - 13 \text{Cov}(X, Y)$$

$$* X, Y \text{ are ind} \Rightarrow \text{Cov}(X, Y) = 0$$

$$= 4 \text{Var}(X)$$

$$= 4(5) = \boxed{20 = \text{Cov}(X, Z)}$$

Chp 4 Exercise 4.6.7: Let  $X_1, \dots, X_n$  be iid  $N(0, 1)$ . Find a value  $C$  s.t.

$$CX_1 / \sqrt{X_2^2 + \dots + X_n^2} \sim t(n) \quad * (\text{see H.O.4 pg 35})$$

From H.O.4 pg 35: If  $Z \sim N(0, 1)$ ,  $U \sim \chi^2(n)$  then

$$T = \frac{Z}{(U/n)^{1/2}} \sim t(n).$$

$$T = \frac{\sqrt{n} Z}{\sqrt{U}} \Rightarrow \boxed{C = \sqrt{n}}$$

Chp 4 Exercise 4.6.10: Let  $X_1, \dots, X_{100}$  be iid  $N(0, 1)$ . Compute the dist of the following:

$$(a) X_1^2 \sim \chi^2(1)$$

$$(b) X_3^2 + X_5^2 \sim \chi^2(2)$$

$$(c) X_{10} / \sqrt{(X_{10}^2 + X_{30}^2 + X_{40}^2)/3} \sim t(3)$$

$$(d) 3X_{10}^2 / [X_{10}^2 + X_{30}^2 + X_{40}^2] \sim F(1, 3)$$

$$(e) 30(X_1^2 + \dots + X_{70}^2) / 70(X_{71}^2 + \dots + X_{100}^2) \sim F(70, 30)$$



7.)  $\beta$   $X_1, \dots, X_n$  are i.i.d. R.Vs from the  $\text{Exp}(\lambda)$  dist. Write down the joint pdf for the random vector  $(X_1, \dots, X_n)$ .

$$f_{X_i}(x_i) = \lambda e^{-\lambda x_i} \Rightarrow f_{X_1, \dots, X_n}(x_1, \dots, x_n) = \prod_{i=1}^n \lambda e^{-\lambda x_i}$$

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n) = \lambda^n e^{-\lambda \sum_{i=1}^n x_i}$$

8.)  $\beta$   $T_1, \dots, T_n$  are i.i.d. binomial  $(4, \theta)$ . Write down the pmf for the random vector  $(T_1, \dots, T_n)$ .

$$f_{T_i}(t_i) = \binom{4}{t_i} \theta^{t_i} (1-\theta)^{4-t_i} = \binom{4}{t_i} \theta^{t_i} (1-\theta)^4 (1-\theta)^{-t_i}$$

$$f_{T_1, \dots, T_n}(t_1, \dots, t_n) = \prod_{i=1}^n \binom{4}{t_i} \theta^{t_i} (1-\theta)^4 (1-\theta)^{-t_i}$$

$$= \left[ \binom{4}{t_1} \dots \binom{4}{t_n} \right] (1-\theta)^{4n} (1-\theta)^{-\sum_{i=1}^n t_i} \theta^{\sum_{i=1}^n t_i}$$

$$f_{T_1, \dots, T_n}(t_1, \dots, t_n) = (1-\theta)^{4n} \left[ \prod_{i=1}^n \binom{4}{t_i} \right] \left[ \left( \frac{\theta}{1-\theta} \right)^{\sum_{i=1}^n t_i} \right]$$

9.) Let  $X_1, \dots, X_n$  be i.i.d w/  $f_{X_i}(x_i) = \frac{1}{2\beta} e^{-|x_i - \mu|/\beta}$ . Write down the joint pdf of the random vector  $(X_1, \dots, X_n)$ . Simplify as possible.

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n) = \prod_{i=1}^n \frac{1}{2\beta} e^{-|x_i - \mu|/\beta}$$

$$= \left( \frac{1}{2\beta} \right)^n e^{-\left( \frac{1}{\beta} \right) \sum_{i=1}^n |x_i - \mu|}$$

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n) = \left( \frac{1}{2\beta} \right)^n \sum_{k=0}^{\infty} \frac{-\left( \frac{1}{\beta} \right)^k \left( \sum_{i=1}^n |x_i - \mu| \right)^k}{k!}$$

10.) Use R to simulate  $N = 10^4$  random samples  $(Z_1, \dots, Z_n)$  from a  $N(0, 1)$ , and compute  $T = \max(Z_1, \dots, Z_n)$  for each sample. Use  $n = 20$ . Plot a histogram and comment on its shape and symmetry.

- The histogram of  $T$  looks close to a normal distribution for  $n = 20$ . However it is skewed slightly to the right. As  $n$  gets larger the right skew becomes more prominent.