

Statistics 630 – Exam I
Wednesday, 30 September 2020

Printed Name: _____ Email: _____

INSTRUCTIONS FOR THE STUDENT:

1. You have 50 minutes to complete the exam (after taking a moment to read these instructions). Please indicate your start time: _____, and your end time: _____
2. There are 6 pages including this cover sheet and the formula sheets.
3. Questions 1–4 are multiple choice and worth 5 points each.
4. Questions 5–8 require solutions to be worked out and are 10 points each. Please write out your answers in *the spaces provided*, explaining your steps. You may refer to theorems by name/description rather than by its number in the book.
5. If you *cannot* print out the exam, please write your answers on blank sheet of paper – in order.
6. You may use the *attached formula sheets*. No other resources are allowed. Do not use the textbook, the class notes, homework or formula sheets that were posted online.
7. *Do not use a calculator.* You may leave answers in forms that can easily be put into a calculator such as $\frac{12}{19}$, $\binom{40}{5}$, e^{-3} , $\Phi(1.5)$, etc.
8. Do not discuss or provide any information to anyone concerning any of the questions on this exam until your solutions are returned or I post my solutions.

I attest that I spent no more than 50 minutes to complete the exam. I used only the materials allowed above. I did not receive assistance from or provide assistance to anyone either before or while taking this exam.

Student's Signature _____

Questions 1–4 are multiple choice: circle the single correct answer. No partial credit!

1. (5 points) Consider the function $F(t) = \begin{cases} 0 & \text{if } t \leq 0, \\ t^2 & \text{if } 0 < t < \frac{1}{2}, \\ 1 - t & \text{if } \frac{1}{2} \leq t < 1, \\ 1 & \text{if } t \geq 1. \end{cases}$

$F(t)$ is not a cdf for some random variable X because

- (a) it is not continuous.
 - (b) it is not nondecreasing.
 - (c) it does not integrate to 1.
 - (d) it is not a function of x .
 - (e) all of the above.
2. (5 points) Suppose $W \sim \text{Poisson}(\lambda)$. We know that $E(W) = \lambda$ and $\text{Var}(W) = \lambda$. What is $E(W(W-1))$?
- (a) $\lambda(\lambda-1)$.
 - (b) λ .
 - (c) λ^2 .
 - (d) 2λ .
 - (e) 0.

3. (5 points) A random vector (X, Y) has joint pdf $f(x, y) = \begin{cases} 5x^2y & \text{if } 0 \leq y \leq x \leq 1, \\ 5xy^2 & \text{if } 0 \leq x \leq y \leq 1, \\ 0 & \text{otherwise.} \end{cases}$

Without any computation, we can say that

- (a) X and Y are independent because $f(x, y) = g(x)h(y)$ for some functions $g(x), h(y)$.
 - (b) X and Y are not independent because $f(x, y) = g(x)h(y)$ for some functions $g(x), h(y)$.
 - (c) X and Y are independent because $f(x, y) \neq g(x)h(y)$ for some functions $g(x), h(y)$.
 - (d) X and Y are not independent because $f(x, y) \neq g(x)h(y)$ for some functions $g(x), h(y)$.
 - (e) none of the above.
4. (5 points) Randy has 17 blue socks in his drawer, 11 of which have no holes, and 13 red socks, 5 of which have no holes. Randy pulls out two socks at random and sees they are both blue. What is the chance neither has a hole?
- (a) $\frac{11 \times 5}{17 \times 13}$.
 - (b) $\frac{11 \times 10}{17 \times 16}$.
 - (c) $\frac{\binom{11}{2}}{\binom{30}{2}}$.
 - (d) $1 - \frac{\binom{6}{2}}{\binom{17}{2}}$.
 - (e) $1 - \frac{\binom{19}{2}}{\binom{30}{2}}$.

Provide solutions to Questions 5–8, to the point of a calculable expression.

5. (10 points) The number of connections between a computer server and customer A is a Poisson(1.3) random variable and, independently, the number of connections with customer B is a Poisson(0.7) random variable. What is the chance the server has at least two connections with each of the two customers? (Show a computable expression.)

6. (10 points) At a given point in time, 2% of individuals in a community have COVID-19. Of those individuals, 0.8% are hospitalized in ICU with the disease. Of individuals without COVID-19, 0.01% are in ICU. If a patient in ICU is randomly selected, what is the chance they have COVID-19? (Show a computable expression.)

7. (10 points) Suppose Z has beta(2,2) pdf $f_Z(z) = 6z(1 - z)$ for $0 \leq z \leq 1$. What *equation must be solved* to get the .25-quantile (25th percentile) for Z ? Simplify the equation as much as possible, but you do not need to solve it.

8. (10 points) $Z \sim \text{gamma}(3, \lambda)$ and the conditional pdf for Y , given $Z = z$, is $f_{Y|Z}(y|z) = \frac{2y}{z^2}$, $0 \leq y \leq z$. Recall that the joint pdf for (Y, Z) is $f(y, z) = f_Z(z)f_{Y|Z}(y|z)$. Use this to find the marginal pdf for Y .

Formulas for Exam I

permutations $P_{n,k} = \frac{n!}{(n-k)!} = n(n-1) \cdots (n-k+1)$.

combinations $C_{n,k} = \binom{n}{k} = \frac{n!}{k!(n-k)!}$.

complement and union $P(A^c) = 1 - P(A)$; $P(A \cup B) = P(A) + P(B) - P(A \cap B)$.

conditional probability $P(A | B) = \frac{P(A \cap B)}{P(B)}$.

independent events $P(A \cap B) = P(A)P(B)$.

total probability $P(A) = \sum_{k=1}^n P(A | B_k)P(B_k)$ if B_1, \dots, B_n are disjoint, $\bigcup_{k=1}^n B_k = \Omega$.

Bayes' rule $P(B_j | A) = \frac{P(A|B_j)P(B_j)}{\sum_{k=1}^n P(A|B_k)P(B_k)}$ if B_1, \dots, B_n are disjoint and $\bigcup_{k=1}^n B_k = \Omega$.

cdf of random variable $F_X(x) = P(X \leq x) = \sum_{y \leq x} p_X(y)$ if X is discrete;

$F_X(x) = P(X \leq x) = \int_{-\infty}^x f_X(y) dy$ if X is continuous.

$P(a < X \leq b) = F_X(b) - F_X(a)$.

pmf of random variable $p_X(x) = P(X = x)$ if X is discrete.

pdf of random variable $f_X(x) = \frac{d}{dx} F_X(x)$ if X is continuous.

quantile function $Q_X(p)$ such that $F(Q_X(p)) = p$ if X is continuous. Otherwise, $Q_X(p)$ satisfies $F_X(x) \leq p \leq F(Q_X(p))$ for $x < Q_X(p)$.

distribution of a function of X $F_Y(y) = P(h(X) \leq y)$ for $Y = h(X)$.

If X is a discrete rv or $h(x)$ takes only countably many values then Y has pmf $p_Y(y) = P(h(X) = y)$.

If X is a continuous rv and $h(x)$ is a continuous function then Y has pdf $f_Y(y) = \frac{d}{dy} P(h(X) \leq y)$.

binomial theorem $\sum_{k=0}^n \binom{n}{k} a^k b^{n-k} = (a+b)^n$.

geometric sum $\sum_{k=n}^{\infty} a^k = \frac{a^n}{1-a}$ if $-1 < a < 1$.

exponential expansion $\sum_{k=0}^{\infty} \frac{a^k}{k!} = e^a$.

integral of a power function $\int_u^v x^a dx = \frac{v^{a+1} - u^{a+1}}{a+1}$ if $a \neq -1$, and $\int_u^v x^{-1} dx = \log_e(v/u)$.

integral of an exponential function $\int_u^v e^{ax} dx = \frac{1}{a} (e^{av} - e^{au})$.

gamma integral $\int_0^{\infty} x^{a-1} e^{-x} dx = \Gamma(a) = (a-1)!$ for $a > 0$.

integral of exponential of a quadratic $\int_{-\infty}^{\infty} e^{a+bx-cx^2} dx = \sqrt{\frac{\pi}{c}} e^{b^2/(4c)+a}$ for $c > 0$.

discrete uniform(N) pmf $p(x) = \frac{1}{N}$ for $x = 1, 2, \dots, N$.

hypergeometric(N, M, n) pmf $p(x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$ for $x = 0, 1, \dots, n$, $M \leq N$.

binomial(n, θ) pmf $p(x) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$ for $x = 0, 1, \dots, n$, $0 < \theta < 1$.

geometric(θ) pmf $p(x) = \theta(1-\theta)^x$ for $x = 0, 1, 2, \dots$, $0 < \theta < 1$.

negative binomial(r, θ) **pmf** $p(x) = \binom{r-1+x}{r-1} \theta^r (1-\theta)^x$ for $x = 0, 1, 2, \dots$, $0 < \theta < 1$.

Poisson(λ) **pmf** $p(x) = \frac{\lambda^x}{x!} e^{-\lambda}$ for $x = 0, 1, 2, \dots$, $\lambda > 0$.

uniform(a, b) **pdf** $f(x) = \frac{1}{b-a}$ for $a < x < b$.

exponential(λ) **pdf** $f(x) = \lambda e^{-\lambda x}$ for $x > 0$, $\lambda > 0$.

gamma(α, λ) **pdf** $f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$ for $x > 0$, $\lambda > 0$, $\alpha > 0$.

normal(μ, σ^2) **pdf** $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/(2\sigma^2)}$ for $-\infty < x < \infty$, $\sigma^2 > 0$.

Weibull(α, β) **pdf** $f(x) = \frac{\alpha}{\beta} (x/\beta)^{\alpha-1} e^{-(x/\beta)^\alpha} I_{(0,\infty)}(x)$ for $\alpha > 0$, $\beta > 0$. $E(X^k) = \beta^k \Gamma(1 + \frac{k}{\alpha})$.

beta(a, b) **pdf** $f(x) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1} (1-x)^{b-1}$ for $0 < x < 1$, $a > 0$, $b > 0$.

joint cdf $F_{X,Y}(x, y) = P(\{X \leq x\} \cap \{Y \leq y\})$.

joint pmf $p_{X,Y}(x, y) = P(\{X = x\} \cap \{Y = y\})$, $F_{X,Y}(x, y) = \sum_{u \leq x} \sum_{v \leq y} p_{X,Y}(u, v)$.

joint pdf $f_{X,Y}(x, y) = \frac{\partial^2}{\partial x \partial y} F_{X,Y}(x, y)$, $F_{X,Y}(x, y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(u, v) dv du$.

marginal pmf/pdf $p_X(x) = \sum_y p_{X,Y}(x, y)$, $p_Y(y) = \sum_x p_{X,Y}(x, y)$;
 $f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy$, $f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx$.

conditional pmf/pdf $p_{X|Y}(x|y) = \frac{p_{X,Y}(x, y)}{p_Y(y)}$; $f_{X|Y}(x|y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$.

independent random variables $p(x, y) = p_X(x)p_Y(y)$ if (X, Y) is discrete;
 $f(x, y) = f_X(x)f_Y(y)$ if (X, Y) is continuous.

discrete convolution $p_{X+Y}(z) = \sum_x p_X(x)p_Y(z-x)$ for independent X, Y .

continuous convolution $f_{X+Y}(z) = \int_{-\infty}^{\infty} f_X(x)f_Y(z-x) dx$ for independent X, Y .

cdf of minimum $F_{\min(X_1, \dots, X_n)}(u) = 1 - (1 - F_{X_1}(u)) \times \dots \times (1 - F_{X_n}(u))$ for independent X_1, \dots, X_n .

cdf of maximum $F_{\max(X_1, \dots, X_n)}(u) = F_{X_1}(u) \times \dots \times F_{X_n}(u)$ for independent X_1, \dots, X_n .

expectation for a discrete rv $E(h(X)) = \sum_x h(x)p_X(x)$.

expectation for a continuous rv $E(h(X)) = \int_{-\infty}^{\infty} h(x)f_X(x) dx$.

mean and variance $\mu_X = E(X)$; $\sigma_X^2 = \text{Var}(X) = E((X - \mu_X)^2) = E(X^2) - \mu_X^2$.

standard deviation $\sigma_X = \sqrt{\text{Var}(X)}$.

covariance and correlation $\text{Cov}(X, Y) = E((X - \mu_X)(Y - \mu_Y)) = E(XY) - \mu_X \mu_Y$; $\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$.

For independent X and Y , $\text{Cov}(X, Y) = \text{Corr}(X, Y) = 0$.

expectation of a sum $E(a_1 X_1 + \dots + a_n X_n) = a_1 E(X_1) + \dots + a_n E(X_n)$.

expectation of a product If X_1, \dots, X_n are independent, $E\left(\prod_{i=1}^n h_i(X_i)\right) = \prod_{i=1}^n E(h_i(X_i))$.

variance of a sum $\text{Var}(aX + bY) = a^2 \text{Var}(X) + 2ab \text{Cov}(X, Y) + b^2 \text{Var}(Y)$.

variance of a sum of independent rvs $\text{Var}(a_1 X_1 + \dots + a_n X_n) = a_1^2 \text{Var}(X_1) + \dots + a_n^2 \text{Var}(X_n)$.

moments k -th moment is $E(X^k)$, $k = 1, 2, \dots$