

STAT 608, Spring 2022 - Assignment 6
SOLUTIONS

1. For logistic regression with one predictor, we use the model

$$\log \left(\frac{\theta(x)}{1 - \theta(x)} \right) = \beta_0 + \beta_1 x$$

- (a) Show that solving for the probability of success for a given value of the predictor, $\theta(x)$, gives

$$\theta(x) = \frac{\exp(\beta_0 + \beta_1 x)}{1 + \exp(\beta_0 + \beta_1 x)}$$

$$\log \left(\frac{\theta(x)}{1 - \theta(x)} \right) = \beta_0 + \beta_1 x$$

$$\frac{\theta(x)}{1 - \theta(x)} = \exp(\beta_0 + \beta_1 x)$$

$$\theta(x) = \exp(\beta_0 + \beta_1 x) - \theta(x)\exp(\beta_0 + \beta_1 x)$$

$$\theta(x) + \theta(x)\exp(\beta_0 + \beta_1 x) = \exp(\beta_0 + \beta_1 x)$$

$$\theta(x) = \frac{\exp(\beta_0 + \beta_1 x)}{1 + \exp(\beta_0 + \beta_1 x)}$$

- (b) and

$$\theta(x) = \frac{1}{1 + \exp(-\{\beta_0 + \beta_1 x\})}$$

$$\begin{aligned} \theta(x) &= \frac{\exp(\beta_0 + \beta_1 x)}{1 + \exp(\beta_0 + \beta_1 x)} \\ &= \frac{\exp(\beta_0 + \beta_1 x)}{1 + \exp(\beta_0 + \beta_1 x)} \left[\frac{\exp(-(\beta_0 + \beta_1 x))}{\exp(-(\beta_0 + \beta_1 x))} \right] \\ &= \frac{1}{\exp(-(\beta_0 + \beta_1 x)) + 1} \\ &= \frac{1}{1 + \exp(-(\beta_0 + \beta_1 x))} \end{aligned}$$

2. On page 285 of the text, it says “When X is a dummy variable, it can be shown that the log odds are also a linear function of x .” Suppose that X is a dummy variable, taking the value 1 with probability π_j , $j = 0, 1$, conditional on $Y = 0, 1$.

(a) Show that the log odds are a linear function of x .

FIRST, DEFINE THE BERNOULLI PROBABILITY $P(X|Y = j) = \pi_j^x(1 - \pi_j)^{1-x}$. THEN NOTE THAT

$$\frac{\theta(x)}{1 - \theta(x)} = \frac{P(Y = 1|X)}{P(Y = 0|X)} = \frac{P(Y = 1, X = x)P(X = x)}{P(Y = 0, X = x)P(X = x)} = \frac{P(Y = 1)P(X = x|Y = 1)}{P(Y = 0)P(X = x|Y = 0)}$$

FINALLY, WE HAVE

$$\begin{aligned} \log\left(\frac{\theta(x)}{1 - \theta(x)}\right) &= \log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + \log\left(\frac{P(X = x|Y = 1)}{P(X = x|Y = 0)}\right) \\ &= \log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + \log\left(\frac{\pi_1^x(1 - \pi_1)^{1-x}}{\pi_0^x(1 - \pi_0)^{1-x}}\right) \\ &= \log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + x \log\left(\frac{\pi_1}{\pi_0}\right) + (1 - x) \log\left(\frac{1 - \pi_1}{1 - \pi_0}\right) \\ &= \log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + \log\left(\frac{1 - \pi_1}{1 - \pi_0}\right) + x \left(\log\left(\frac{\pi_1}{\pi_0}\right) - \log\left(\frac{1 - \pi_1}{1 - \pi_0}\right)\right) \\ &= \log\left(\frac{P(Y = 1)P(X = 0|Y = 1)}{P(Y = 0)P(X = 0|Y = 0)}\right) + x \log\left(\frac{\pi_1/(1 - \pi_1)}{\pi_0/(1 - \pi_0)}\right) \\ &= a + bx \end{aligned}$$

(b) Define the slope and intercept for the linear function.

THE INTERCEPT IS

$$\log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + \log\left(\frac{1 - \pi_1}{1 - \pi_0}\right)$$

AND THE SLOPE IS

$$\log\left(\frac{\pi_1/(1 - \pi_1)}{\pi_0/(1 - \pi_0)}\right)$$

3. On page 284 of the text, the author quotes Cook and Weisberg: “When conducting a binary regression with a skewed predictor, it is often easiest to assess the need for x and $\log(x)$ by including them both in the model so that their relative contributions can be assessed directly.” Show that indeed the log odds are a function of x and $\log(x)$ for the gamma distribution.

THE TEXT GIVES THE FOLLOWING, SO WE HANDLE ONLY THE SECOND TERM IN THE SUM, AS THE FIRST TERM IS A CONSTANT WITH RESPECT TO x :

$$\log\left(\frac{\theta(x)}{1 - \theta(x)}\right) = \log\left(\frac{P(Y = 1)}{P(Y = 0)}\right) + \log\left(\frac{f(x|Y = 1)}{f(x|Y = 0)}\right)$$

WE’LL USE THE PARAMETRIZATION OF THE GAMMA DISTRIBUTION AS FOLLOWS:

$$f(x|\alpha, \beta) = \frac{x^{\alpha-1}e^{-x/\beta}}{\Gamma(\alpha)\beta^\alpha}$$

THEN WE CAN START REARRANGING:

$$\begin{aligned}\log\left(\frac{f(x|Y=1)}{f(x|Y=0)}\right) &= \log\left(\frac{x^{\alpha_1-1}e^{-x/\beta_1}\Gamma(\alpha_0)\beta_0^{\alpha_0}}{x^{\alpha_0-1}e^{-x/\beta_0}\Gamma(\alpha_1)\beta_1^{\alpha_1}}\right) \\ &= \log\left(\frac{\Gamma(\alpha_0)\beta_0^{\alpha_0}}{\Gamma(\alpha_1)\beta_1^{\alpha_1}}\right) + (\alpha_1 - \alpha_0)\log(x) + \left(\frac{1}{\beta_0} + \frac{1}{\beta_1}\right)x\end{aligned}$$

FINALLY, WE SEE THAT THE LOG ODDS IS A FUNCTION OF x AND $\log(x)$ WHEN x HAS THE GAMMA DISTRIBUTION.

4. Chapter 8, Question 4

- (a) No, MODEL (8.6) IS NOT A VALID MODEL: WE SEE FROM THE MARGINAL MODEL PLOTS THAT OUR MODEL DOES NOT MATCH A NONPARAMETRIC FIT TO THE DATA.
- (b) THE REAL ANSWER HERE IS THAT THE PREDICTORS MAY VERY WELL HAVE NONLINEAR ASSOCIATIONS AND WE MAY BE MISSING SOME IMPORTANT PREDICTORS LIKE DIET AND EXERCISE; HOWEVER, BASED ON THE KERNEL DENSITY ESTIMATES, WE FIRST OBSERVE THAT x_1 AND x_4 ARE BOTH RIGHT-SKEWED. AS SHOWN ABOVE, THIS MAY INDICATE THAT ADDING THE LOG TRANSFORMATIONS OF BOTH OF THESE VARIABLES MAY BE IMPORTANT. TO A SMALLER DEGREE, IT APPEARS VARIABLE x_1 HAS UNEQUAL VARIANCES WHEN HEART DISEASE = YES AND NO, WHICH MAY INDICATE ADDING A QUADRATIC TERM IF WE THINK x_1 IS APPROXIMATELY NORMALLY DISTRIBUTED. THE FACT THAT THE PLOT IS MUCH WIDER THAN IT IS NARROW MAY BE EXAGGERATING THE SKEWNESS TO OUR EYES.
- (c) THIS IS A MUCH-IMPROVED MODEL. WE SEEM TO STILL BE HAVING TROUBLE MODELING THE RELATIONSHIP BETWEEN x_1 AND THE ODDS OF A HEART ATTACK AT VERY SMALL VALUES OF x_1 , BUT OTHERWISE, THE MODEL SEEMS TO MATCH A NONPARAMETRIC FIT PRETTY WELL. WE DON'T HAVE THE ABILITY TO CONDUCT A GOODNESS OF FIT TEST FOR THESE BINARY DATA.
- (d) HOLDING SYSTOLIC BLOOD PRESSURE, CHOLESTEROL, OBESITY, AND AGE CONSTANT, WHEN A PATIENT HAS A FAMILY HISTORY OF HEART DISEASE, OUR MODEL PREDICTS THEIR ODDS OF A HEART ATTACK TO BE $e^{0.941056} = 2.56$ TIMES AS LARGE AS WHEN A PATIENT DOES NOT HAVE A FAMILY HISTORY OF HEART DISEASE.