

Regression 2

StatML

20.02.2014

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What happens now?

- The TAs have graded your assignments
- General and individual feedback at TA sessions
- **Optional lecture on the Perceptron**
by Christian Friday 13.30-14.15 in Aud 3 (HCØ)
- **Math Q&A / help session Friday afternoon**
14.15 - ca 16.00, A103, A104 and A105 at HCØ

About Assignment 1

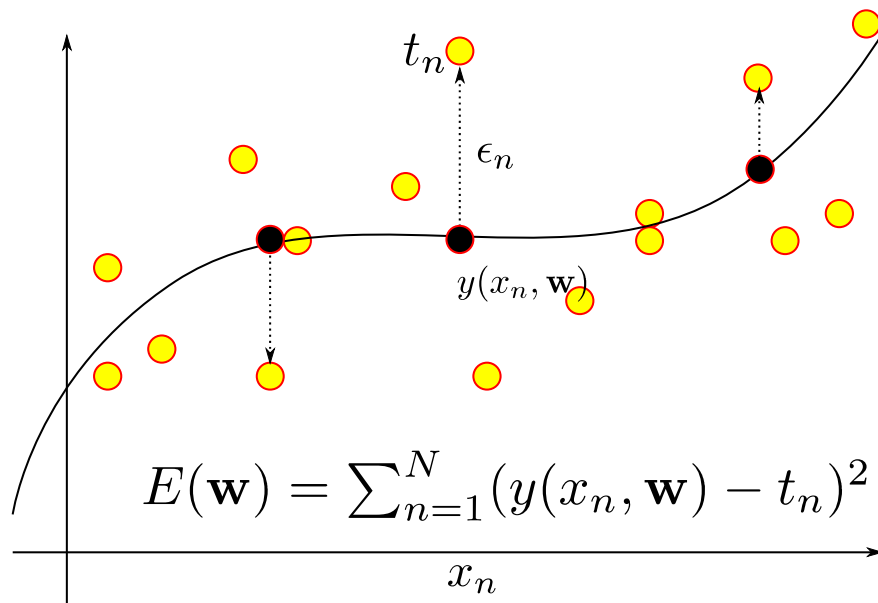
- Deadline for resubmission is Tuesday 25.02.
- There will only be one resubmission round.
- If you are asked to resubmit Exercise 1.3, you may instead choose to resubmit the make-up assignment posted in Absalon.
- This is a one-time only exception.

After today's lecture you should:

- Be able to produce a regularized maximum likelihood solution to a linear regression model
- Be able to produce a maximum a posteriori solution to a linear regression model
- Understand the relation between maximum a posteriori solutions and regularized maximum likelihood solutions
- Be familiar with different choices of regularization of why you would want to use them
- Understand the curse of dimensionality and its impact on solving regression problems
- Understand the effect of choice of prior in MAP estimates for different problems
- Be able to recognize and pose practical regression problems

Last time: Geometric and Probabilistic approaches to Regression

Geometric approach



Maximum Likelihood approach

Assume: Gaussian noise model

$$\mathcal{N}(t|y(x, \mathbf{w}), \beta^{-1})$$

Likelihood of data \mathbf{t} under model fixed by \mathbf{w}, \mathbf{x}

$$\begin{aligned} p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) \\ = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1}) \end{aligned}$$

Maximizing the likelihood is equivalent to minimizing $\sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2$

Last time: Analytic solution to Maximum Likelihood a.k.a. Geometric Least Square regression

Minimizing $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ when $y(\mathbf{x}, \mathbf{w}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$.

$$\begin{aligned} & \frac{\partial}{\partial w_i} [\sum_{n=1}^N (\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) - t_n)^2] \\ &= \sum_{n=1}^N \frac{\partial}{\partial w_i} [(\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) - t_n)^2] \\ &= \sum_{n=1}^N 2(\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) - t_n) \cdot \frac{\partial}{\partial w_i} (\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) - t_n) \\ &= 2 \sum_{n=1}^N (\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) - t_n) \cdot \phi_i(\mathbf{x}_n) - t_n = 0 \quad \text{for all } i \end{aligned}$$

Since $\boldsymbol{\phi}(\bar{x})^T = (\phi_0(\mathbf{x}), \phi_1(\mathbf{x}), \dots, \phi_{M-1}(\mathbf{x}))$, we get

$$\sum_{n=1}^N \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) \boldsymbol{\phi}(\mathbf{x}_n)^T - \sum_{n=1}^N t_n \boldsymbol{\phi}(\mathbf{x}_n)^T = 0,$$

or

$$0 = \mathbf{w}^T \sum_{n=1}^N \boldsymbol{\phi}(\mathbf{x}_n) \boldsymbol{\phi}(\mathbf{x}_n)^T - \sum_{n=1}^N t_n \boldsymbol{\phi}(\mathbf{x}_n)^T \quad (*)$$

$$\text{Setting } \Phi = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \dots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \dots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & & & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \dots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix}$$

we rewrite (*) as $0 = \mathbf{w}^T (\Phi^T \Phi) - \mathbf{t}^T \Phi$

$$\Rightarrow \mathbf{w}^T (\Phi^T \Phi) = \mathbf{t}^T \Phi$$

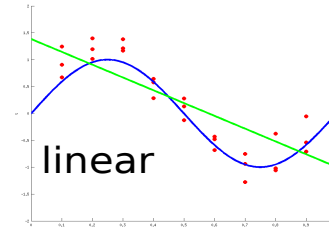
$$\Rightarrow (\Phi^T \Phi)^T \mathbf{w} = (\Phi^T \Phi) \mathbf{w} = \Phi^T \mathbf{t} \quad (\text{transpose})$$

$$\Rightarrow \mathbf{w} = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t}$$

Different basis functions

$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$ where $\{\phi_j(\mathbf{x})\}$ are **basis functions**

$$\mathbf{w} = \begin{pmatrix} w_0 \\ w_1 \\ \vdots \\ w_{M-1} \end{pmatrix} \text{ and } \boldsymbol{\phi} = \begin{pmatrix} \phi_0 \\ \phi_1 \\ \vdots \\ \phi_{M-1} \end{pmatrix}$$

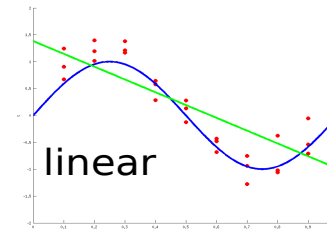


$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 x_1 + w_2 x_2 + \dots + w_D x_D$$

Different basis functions

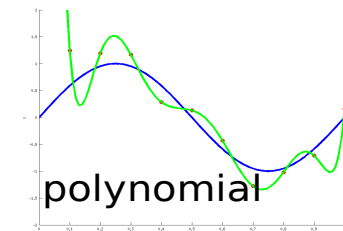
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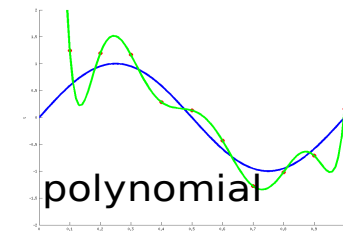
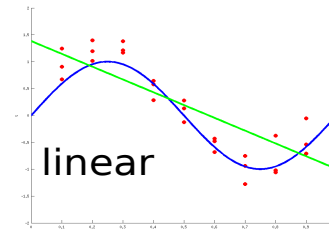
$$y(x, \mathbf{w}) = w_0 + w_1 x^1 + w_2 x^2 + \dots + w_{M-1} x^{M-1}$$



Different basis functions

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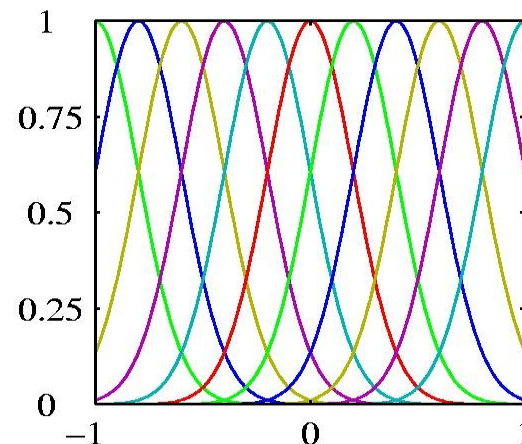


$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 x_1 + w_2 x_2 + \dots + w_D x_D$$

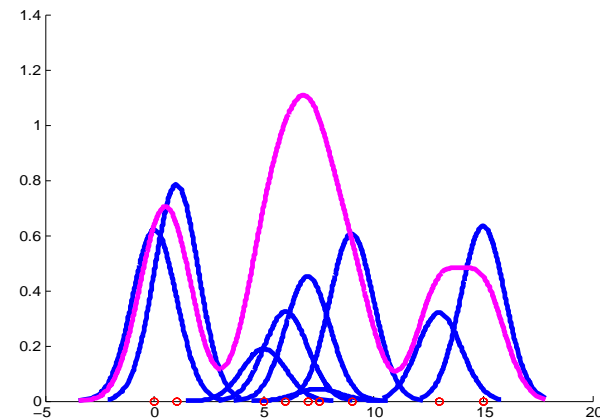
$$y(x, \mathbf{w}) = w_0 + w_1 x^1 + w_2 x^2 + \dots + w_{M-1} x^{M-1}$$

$$y(x, \mathbf{w}) = w_0 + w_1 e^{-\frac{1}{2s^2}(x-x_1)^2} + \dots + w_{M-1} e^{-\frac{1}{2s^2}(x-x_{M-1})^2}$$

Radial basis functions
(Gaussians)



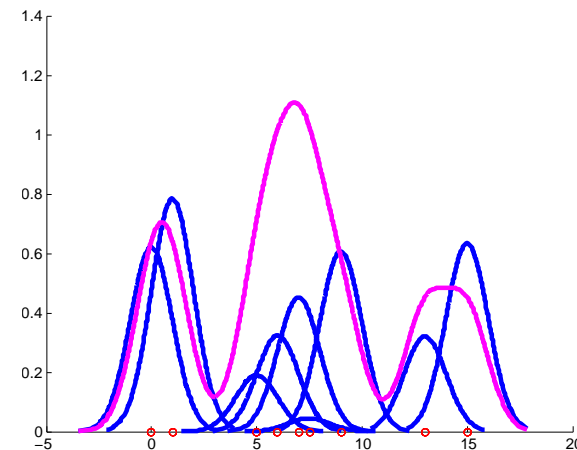
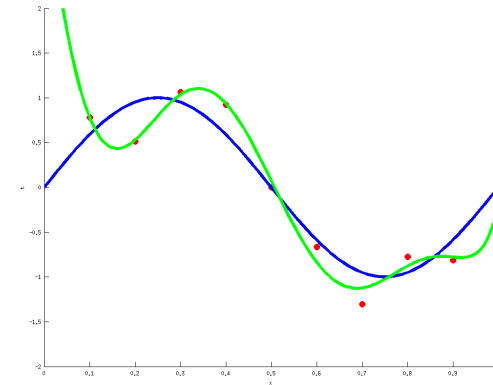
From Bishop



Basis functions:

Global versus local effect

- **Polynomials** fits data globally: Change a parameter and it has effect globally by changing the whole curve.
- The **radial basis functions** fits data locally: Changing a parameter changes the basis weight locally and only changes the curve locally. Have infinite support (will cause very small changes far away).
- Splines (piecewise polynomials) fit data locally: Changing a parameter only affects the curve locally (in the region of the local polynomial).



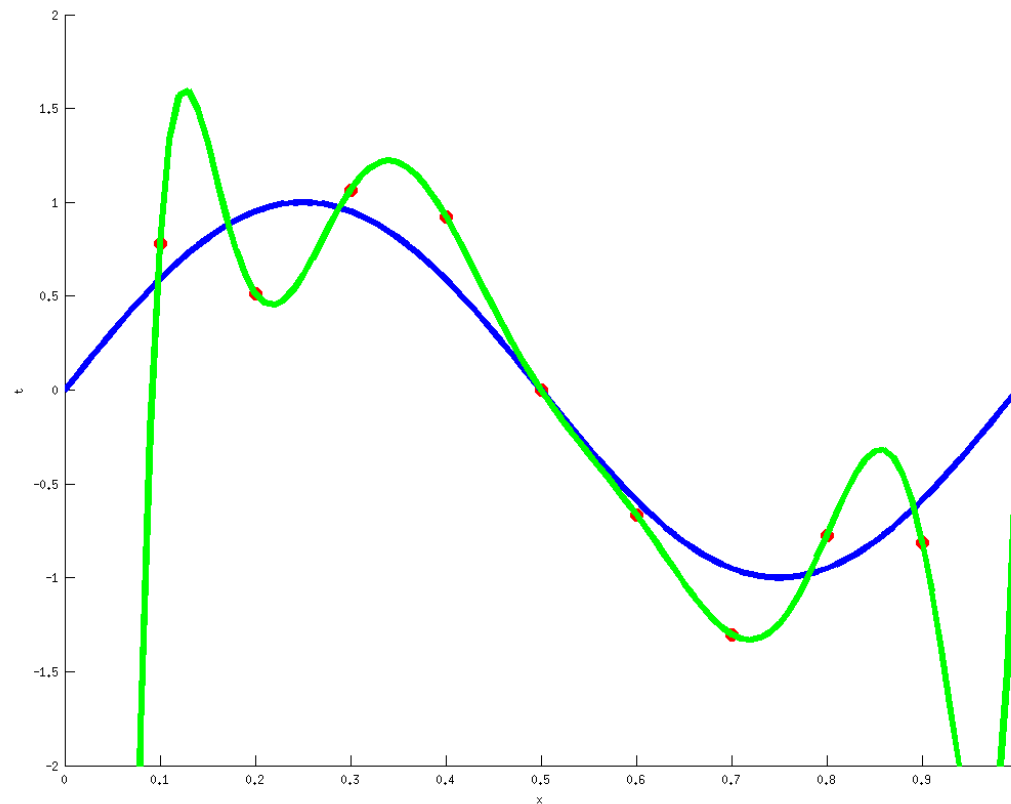
Curse of Dimensionality

- D-dimensional polynomial curve fitting, $M = 3$:

$$y(\mathbf{x}, \mathbf{w}) = w_0 + \sum_{i=1}^D w_i x_i + \sum_{i=1}^D \sum_{j=1}^D w_{ij} x_i x_j + \sum_{i=1}^D \sum_{j=1}^D \sum_{k=1}^D w_{ijk} x_i x_j x_k$$

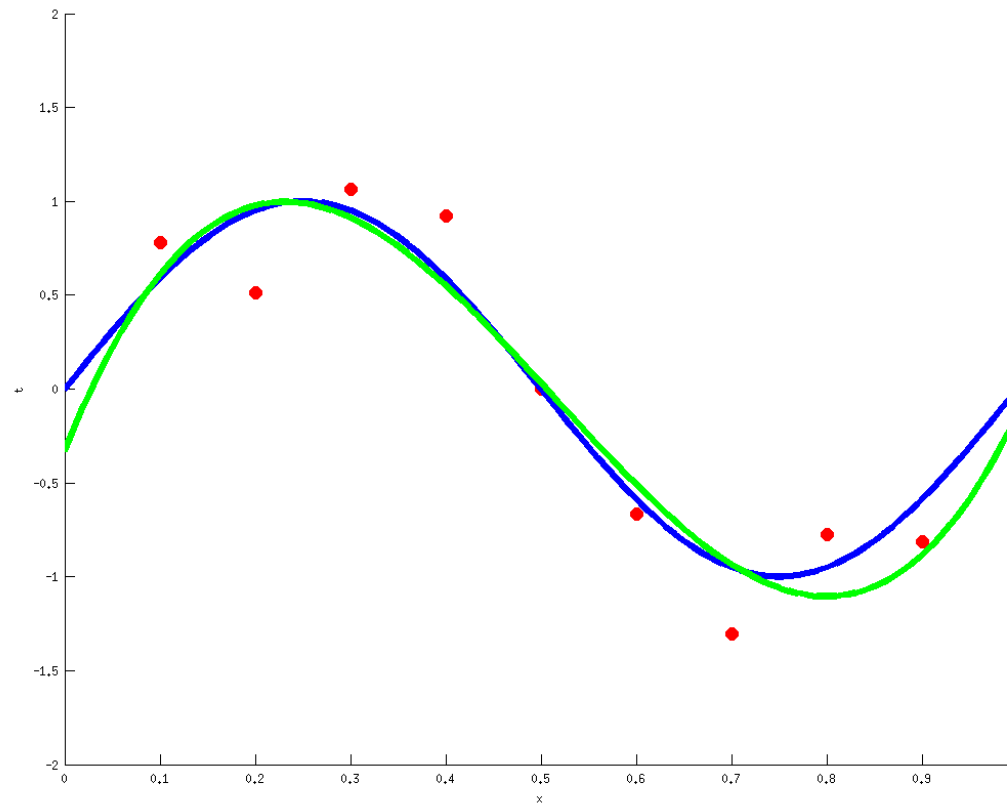
- **In general:**
 - Number of free model parameters grows polynomially in D^M with the dimensionality D .
 - The data set size N should grow polynomially to keep same precision on parameter estimates.

Example from last time:



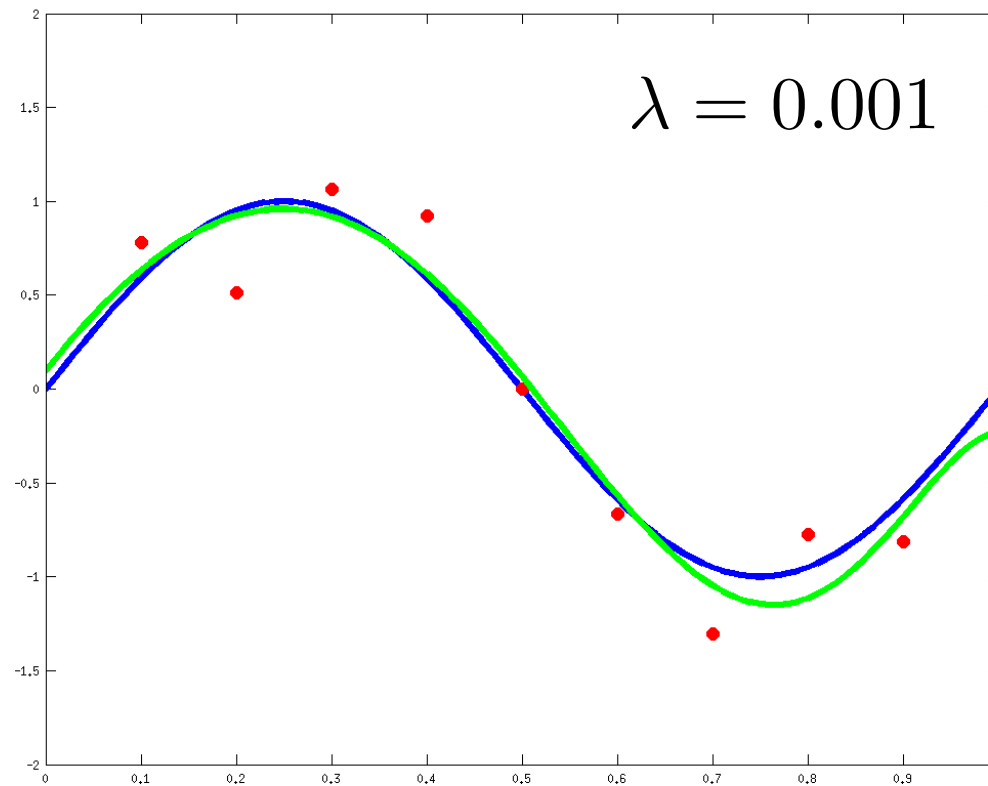
$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + \dots + w_9x^9$$

Example from last time:



$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + w_3x^3$$

Example from last time: Regularization



$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + \dots + w_9x^9$$

$$E(\mathbf{w}) = \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 + \lambda \|\mathbf{w}\|^2$$

Solving the regularized regression problem

Minimizing $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \lambda \|\mathbf{w}\|^2$ when $y(\mathbf{x}, \mathbf{w}) = \mathbf{w}^T \phi(\mathbf{x})$.

$$\begin{aligned} & \frac{\partial}{\partial w_i} [\sum_{n=1}^N (\mathbf{w}^T \phi(\mathbf{x}_n) - t_n)^2 + \lambda \mathbf{w}^T \mathbf{w}] \\ &= \sum_{n=1}^N \frac{\partial}{\partial w_i} [(\mathbf{w}^T \phi(\mathbf{x}_n) - t_n)^2 + \lambda \mathbf{w}^T \mathbf{w}] \\ &= \sum_{n=1}^N 2(\mathbf{w}^T \phi(\mathbf{x}_n) - t_n) \cdot \frac{\partial}{\partial w_i} (\mathbf{w}^T \phi(\mathbf{x}_n) - t_n) + 2\lambda w_i \\ &= 2 \sum_{n=1}^N (\mathbf{w}^T \phi(\mathbf{x}_n) - t_n) \cdot \phi_i(\mathbf{x}_n) + 2\lambda w_i = 0 \quad \text{for all } i \end{aligned}$$

Since $\phi(\bar{x})^T = (\phi_0(\mathbf{x}), \phi_1(\mathbf{x}), \dots, \phi_{M-1}(\mathbf{x}))$, we get

$$\sum_{n=1}^N \mathbf{w}^T \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T - \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T + \lambda \mathbf{w}^T = 0,$$

or

$$0 = \mathbf{w}^T \sum_{n=1}^N \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T - \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T + \lambda \mathbf{w}^T (*)$$

$$\text{Setting } \Phi = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \dots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \dots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \dots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix}$$

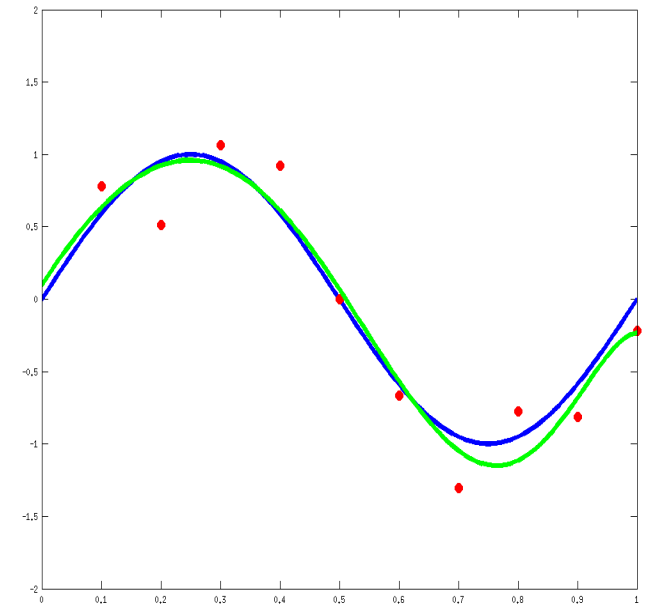
we rewrite (*) as $0 = \mathbf{w}^T (\Phi^T \Phi) - \mathbf{t}^T \Phi + \lambda \mathbf{w}^T$

$$\Rightarrow \mathbf{w}^T (\Phi^T \Phi + \lambda \mathbf{I}) = \mathbf{t}^T \Phi$$

$$\Rightarrow (\Phi^T \Phi + \lambda \mathbf{I})^T \mathbf{w} = \Phi^T \mathbf{t}$$

$$\Rightarrow (\Phi^T \Phi + \lambda \mathbf{I}) \mathbf{w} = \Phi^T \mathbf{t}$$

$$\Rightarrow \mathbf{w} = (\Phi^T \Phi + \lambda \mathbf{I})^{-1} \Phi^T \mathbf{t}$$



$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 x + w_2 x^2 + \dots + w_9 x^9$$

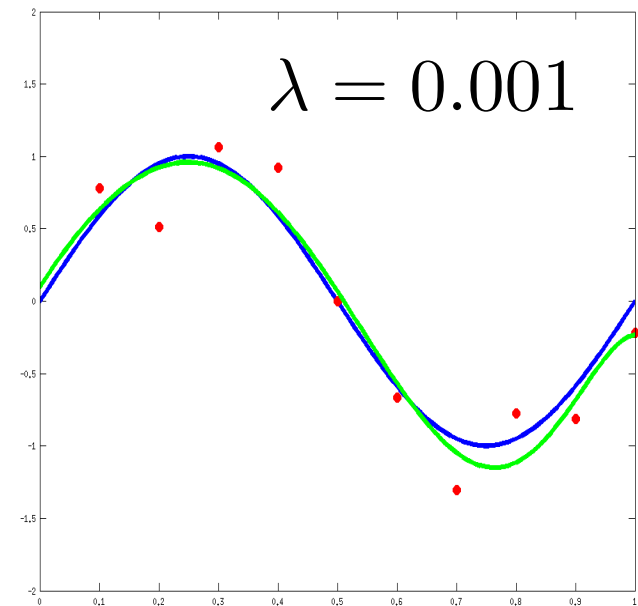
$$E(\mathbf{w}) = \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 + \lambda \|\mathbf{w}\|^2$$

OBS! Also stabilizes the matrix inversion... 15

Regularization

- Adding an L2 punishment of the weight vector is referred to as *ridge regression*
- Drives weights towards small norm
- **Interpretation of weights:**

Tell you about the importance of each basis function for describing the data (*this interpretation is a heuristic!*)



$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1\phi_1(x) + w_2\phi_2(x) + \dots + w_9\phi_9(x)$$

$$E(\mathbf{w}) = \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 + \lambda \|\mathbf{w}\|_{10}^2$$

Regularization

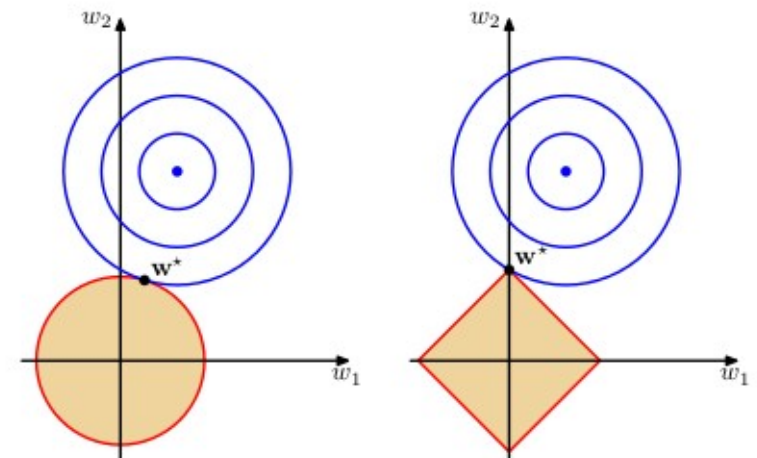
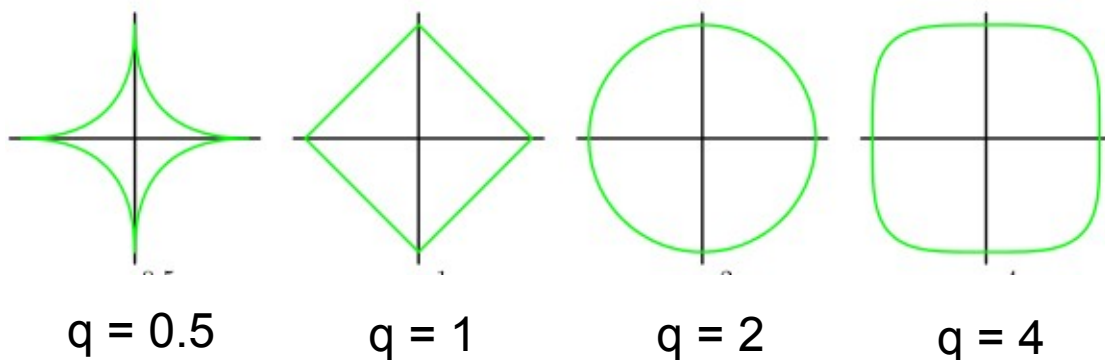
- More generally, regularization can be done by adding a term of degree q

$$E(\mathbf{w}) = \sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 + \lambda \|\mathbf{w}\|^q$$

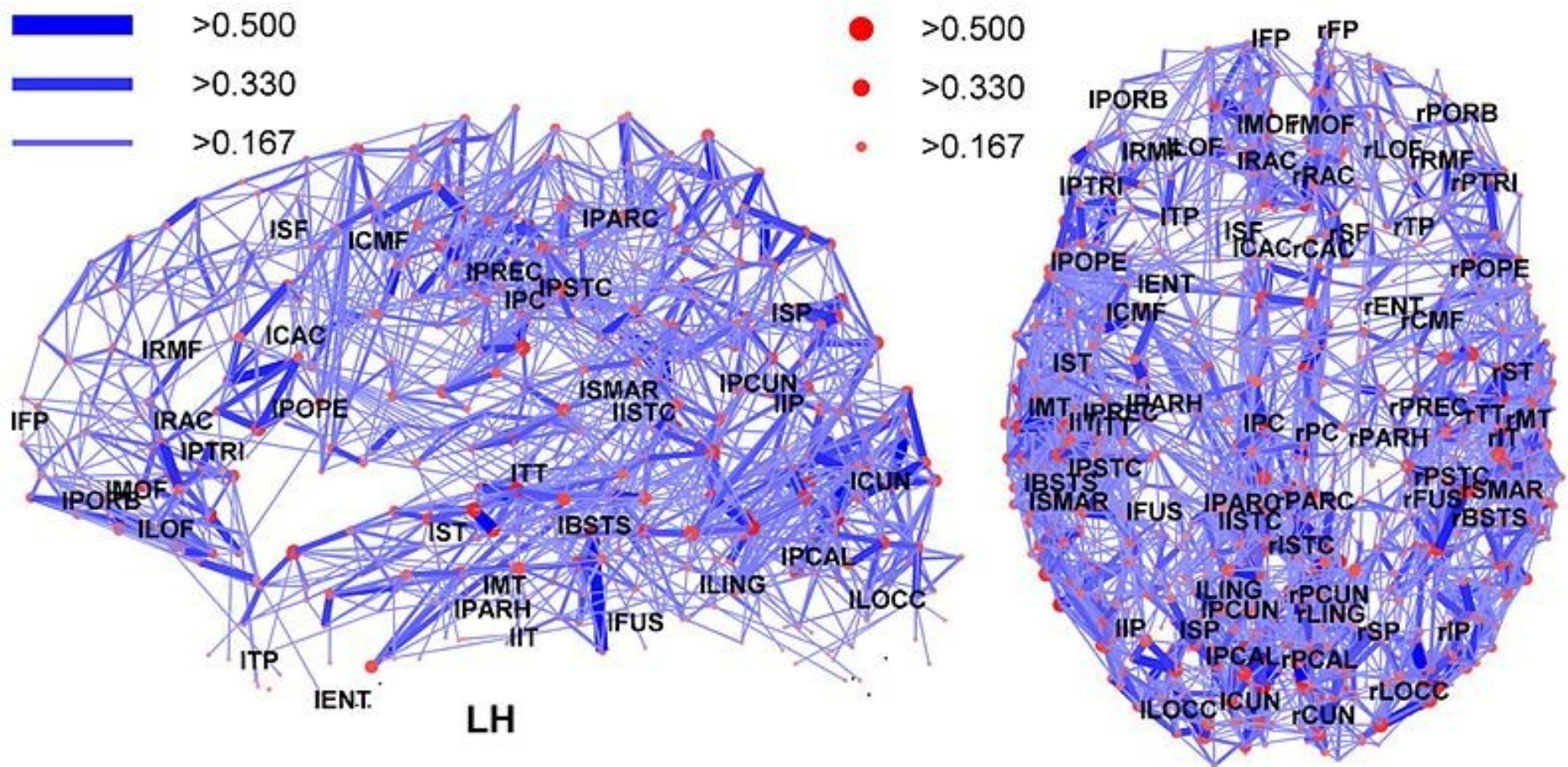
- When $q = 1$, this is called the *lasso*.
- For $q = 1$ or smaller, minimization will prefer weights = 0
- Interpretation of weights:

Curse of dim

- Supervised **dimensionality reduction** / **feature selection**
- Importance of different basis functions

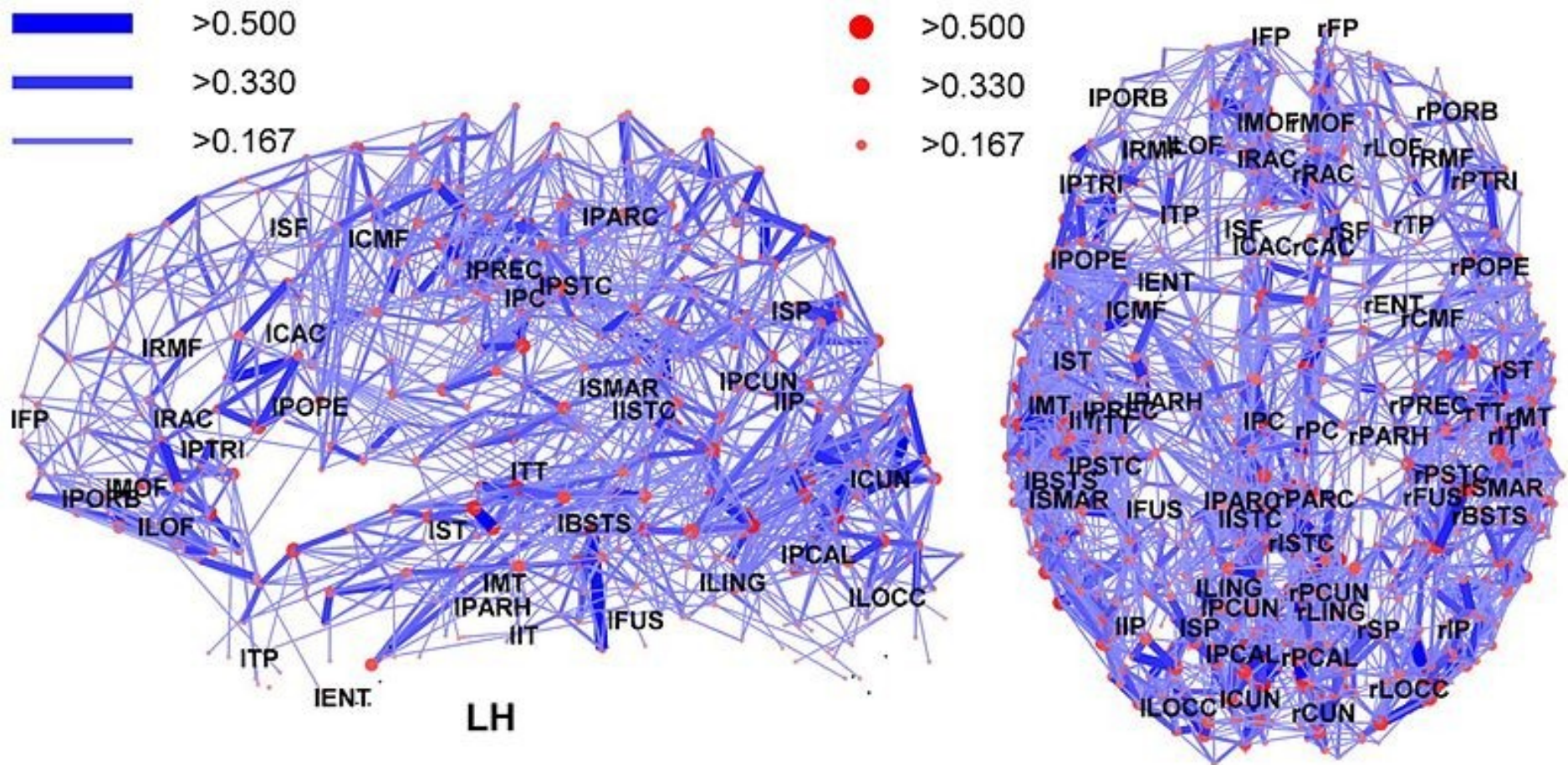


Example: Sparse regression for brain connectivity



$$E(\mathbf{w}) = \sum_{n=1}^N (\sum_{e \in E} w_e \phi_e - t_n)^2 + \lambda \|\mathbf{w}\|$$

Example: Structured sparse regression for brain connectivity



$$E(\mathbf{w}) = \sum_{n=1}^N (\sum_{e \in E} w_e \phi_e - t_n)^2 + \lambda \|\mathbf{w}\| + \sum_{G \subset E} \lambda_2 \|\mathbf{w}_G\|^2$$

Recall from Lecture 2:

- Maximum Likelihood estimates

Find the model parameters

\mathbf{w}

that maximize the joint probability

$$p(D \mid \mathbf{w})$$

of observing the data given the model

- Maximum a posteriori estimates

Find the most likely model parameters given the data, that is find the model parameters

$$p(\mathbf{w} \mid D) \propto p(D \mid \mathbf{w})p(\mathbf{w})$$

Bayesian regression: Maximum a Posteriori solution

N input variables

$$\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_N \end{pmatrix}$$

with target variables

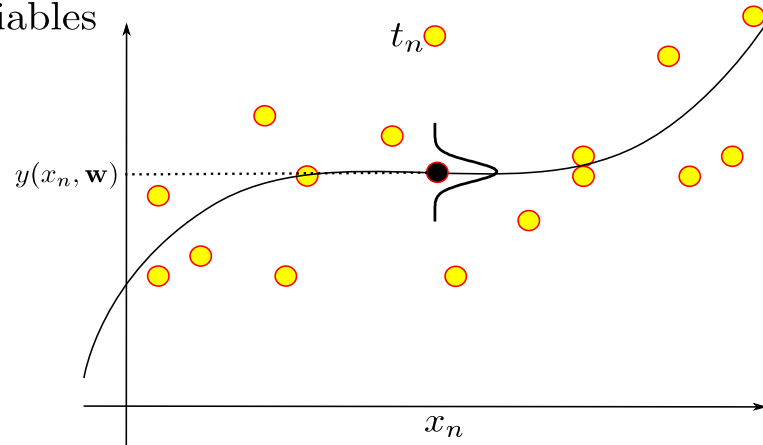
$$\mathbf{t} = \begin{pmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{pmatrix}$$

Assume: Gaussian noise model

$$\mathcal{N}(t|y(x, \mathbf{w}), \beta^{-1})$$

Likelihood of data \mathbf{t} under model fixed by \mathbf{w}, \mathbf{x}

$$p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1})$$



Bayesian regression: Maximum a Posteriori solution

N input variables

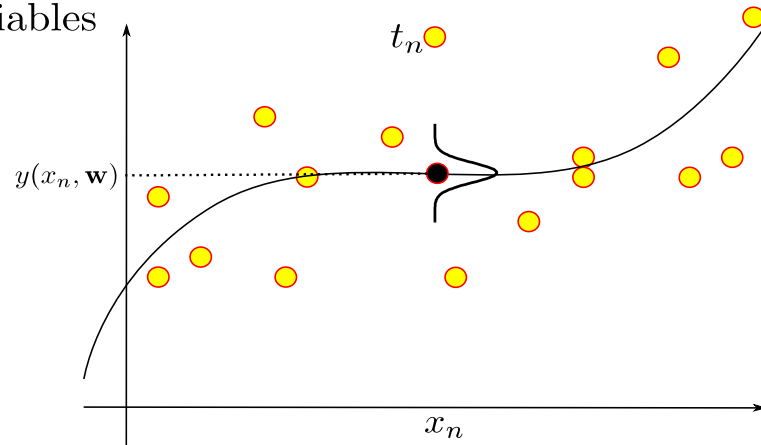
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$$p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1})$$

Conjugate prior:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{m}_0, \mathbf{S}_0)$$

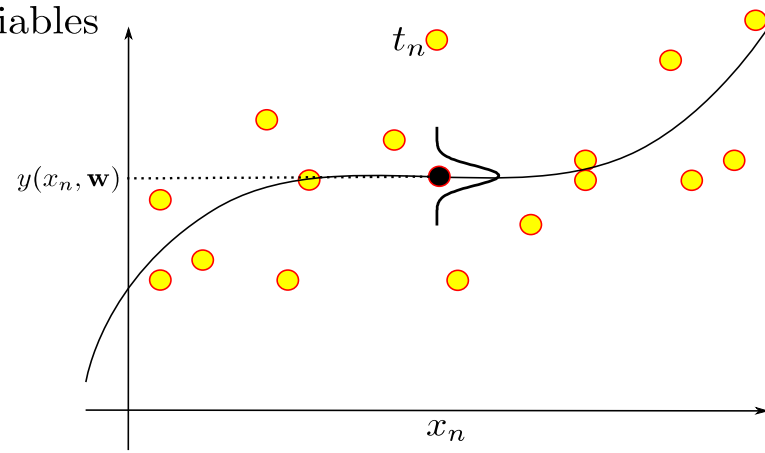
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Posterior distribution:

$$p(\mathbf{w}|\mathbf{t}) = p(\mathbf{t}|\mathbf{w})p(\mathbf{w}) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1}) \cdot \mathcal{N}(\mathbf{w}|\mathbf{m}_0, \mathbf{S}_0)$$

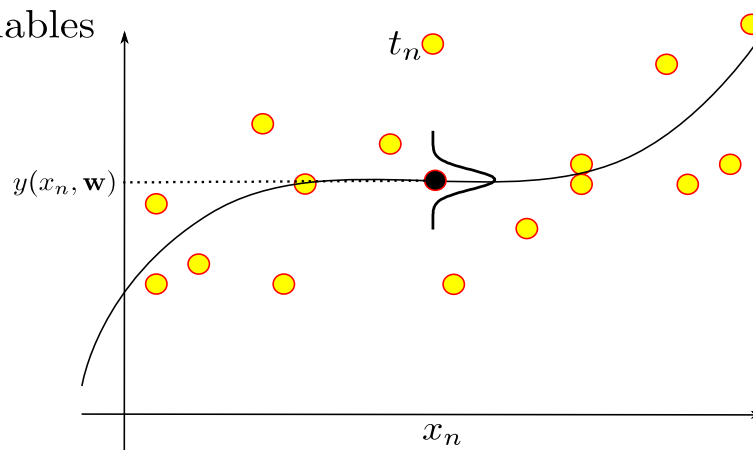
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where

$$\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\mathbf{\Phi}^T\mathbf{t}) \quad \mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta\mathbf{\Phi}^T\mathbf{\Phi} \quad (\text{See CB for proof})$$

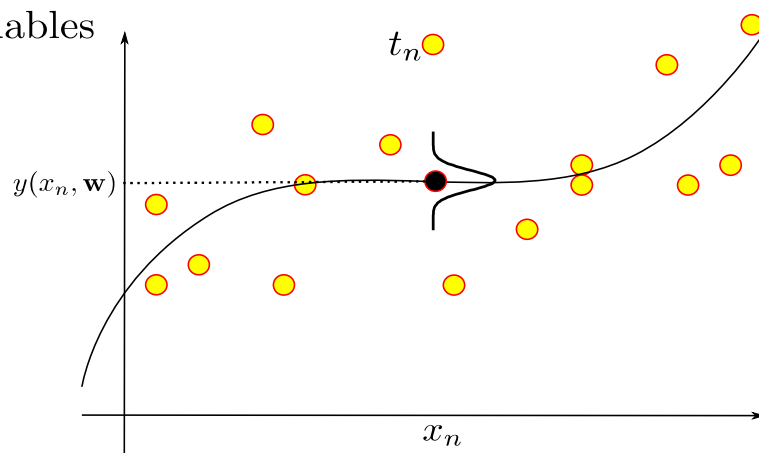
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$\mathbf{w}_{MAP} = \mathbf{m}_N$ since $p(\mathbf{w}|\mathbf{t})$ is a (unimodal) Gaussian

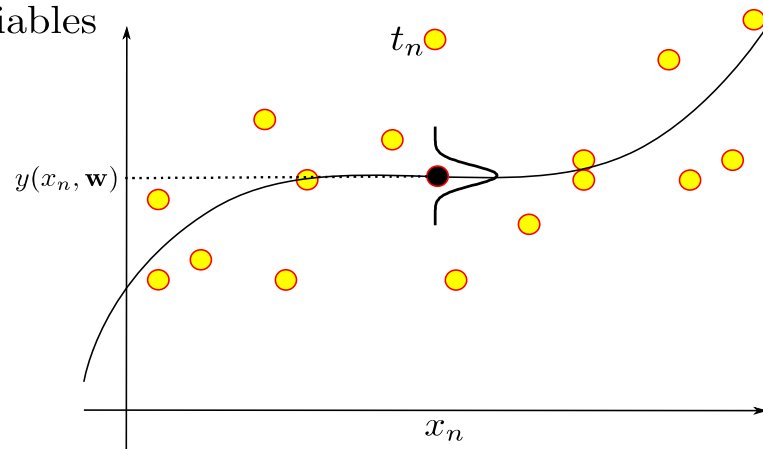
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Conjugate prior:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{m}_0, \mathbf{S}_0)$$

$$\Phi = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \dots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \dots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & & & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \dots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix}$$

Posterior distribution:

$$\begin{aligned} p(\mathbf{w}|\mathbf{t}) &= p(\mathbf{t}|\mathbf{w})p(\mathbf{w}) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1}) \cdot \mathcal{N}(\mathbf{w}|\mathbf{m}_0, \mathbf{S}_0) \\ &= \mathcal{N}(\mathbf{w}|\mathbf{m}_N, \mathbf{S}_N) \quad (\text{product of Gaussians}) \end{aligned}$$

where

$$\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\Phi^T\mathbf{t}) \quad \mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta\Phi^T\Phi$$

(See CB for proof)

Design matrix

$\mathbf{w}_{MAP} = \mathbf{m}_N$ since $p(\mathbf{w}|\mathbf{t})$ is a (unimodal) Gaussian

Analytic
solution

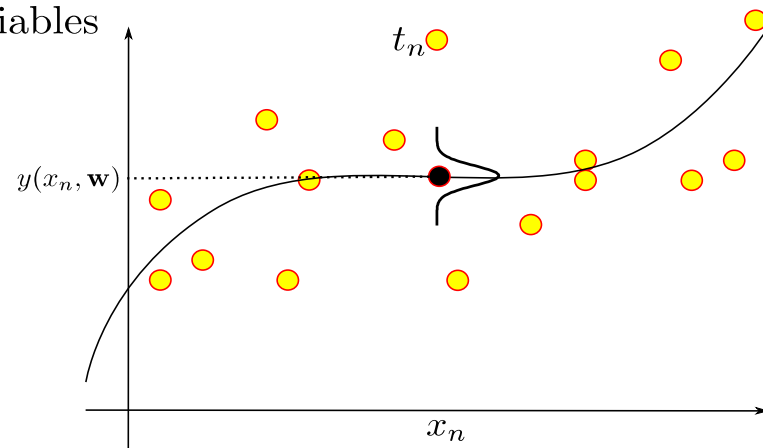
Effect of the prior

N input variables

$$\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_N \end{pmatrix}$$

with target variables

$$\mathbf{t} = \begin{pmatrix} t_1 \\ t_2 \\ \vdots \\ t_N \end{pmatrix}$$



Assume: Gaussian noise model

$$\mathcal{N}(t|y(x, \mathbf{w}), \beta^{-1})$$

Likelihood of data \mathbf{t} under model fixed by \mathbf{w}, \mathbf{x}

$$p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1})$$

Conjugate prior:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{m}_0, \mathbf{S}_0)$$

Posterior distribution: $\mathcal{N}(\mathbf{w}|\mathbf{m}_N, \mathbf{S}_N)$ where

$$\mathbf{m}_N = \mathbf{S}_N(\mathbf{S}_0^{-1}\mathbf{m}_0 + \beta\Phi^T\mathbf{t}) \quad \mathbf{S}_N^{-1} = (\mathbf{S}_0^{-1} + \beta\Phi^T\Phi)$$

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Effect of prior:

If $\mathbf{S}_0 = \alpha^{-1}I$ with $\alpha \rightarrow 0$, then $\mathbf{m}_N \rightarrow \mathbf{w}_{ML} = (\Phi^T\Phi)^{-1}\Phi^T\mathbf{t}$

If $N = 0$, then $\mathbf{m}_N = \mathbf{m}_0$

Design matrix

Analytic
solution

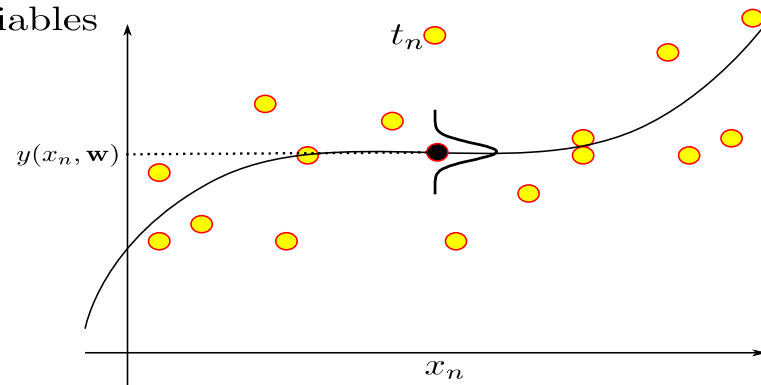
Relation to Maximum Likelihood

N input variables

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$$\mathcal{N}(t|y(x, \mathbf{w}), \beta^{-1})$$

Likelihood of data \mathbf{t} under model fixed by \mathbf{w}, \mathbf{x}

$$p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|y(x_n, \mathbf{w}), \beta^{-1})$$

Conjugate prior:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha\mathbf{I})$$

Maximize the posterior:

$$\begin{aligned} & \underset{\mathbf{w}}{\operatorname{argmax}} p(\mathbf{w}|\mathbf{x}, \mathbf{t}, \alpha, \beta) \\ &= \underset{\mathbf{w}}{\operatorname{argmin}} -\ln(p(\mathbf{w}|\mathbf{x}, \mathbf{t}, \alpha, \beta)) \\ &= \underset{\mathbf{w}}{\operatorname{argmin}} -\ln(p(\mathbf{t}|\mathbf{w}, \mathbf{x}, \alpha, \beta)p(\mathbf{w})) \\ &= \underset{\mathbf{w}}{\operatorname{argmin}} [-\ln(p(\mathbf{t}|\mathbf{w}, \mathbf{x}, \alpha, \beta)) - \ln(p(\mathbf{w}))] \\ &= \underset{\mathbf{w}}{\operatorname{argmin}} \left[-\sum_{n=1}^N (y(x_n, \mathbf{w}) - t_n)^2 + \frac{\alpha}{2} \|\mathbf{w}\|^2 + \text{const} \right] \end{aligned}$$

So adding a prior in the MAP estimate is equivalent to adding a regularizer in the ML estimate

Sequential learning

Assume: measurements are arriving sequentially

$$(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_N, t_N)$$

Want to learn "on the go" – e.g. tracking
personalized models
etc

Sequential learning

Assume: measurements are arriving sequentially

$$(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_N, t_N)$$

Want to learn "on the go" – e.g. tracking
personalized models
etc

$$p(\mathbf{w}|\mathbf{t}, \mathbf{x}, \beta) \propto \underbrace{p(\mathbf{t}|\mathbf{x}, \mathbf{w}, \beta)}_{\text{likelihood}} \underbrace{p(\mathbf{w})}_{\text{prior}} \quad (\text{Bayes})$$

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$$= \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1}) \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$$

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$$= \mathcal{N}(t_N | \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_N), \beta^{-1}) \prod_{n=1}^{N-1} \mathcal{N}(t_n | \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1}) \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0)$$

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$$= \underbrace{p(t_N | x_N, \mathbf{w}, \beta)}_{\text{likelihood for } t_N} \underbrace{p(\mathbf{w})}_{\text{original prior}} \underbrace{\prod_{n=1}^{N-1} p(t_n | \mathbf{x}_n, \mathbf{w}, \beta)}_{\text{likelihood of observing } t_1, \dots, t_{N-1}}$$

Sequential learning

Assume: measurements are arriving sequentially

$$(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_N, t_N)$$

Want to learn "on the go" – e.g. tracking
personalized models
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posterior for $N - 1 = \text{prior for } N$

Sequential learning

Assume: measurements are arriving sequentially

$$(\mathbf{x}_1, t_1), (\mathbf{x}_2, t_2), \dots, (\mathbf{x}_N, t_N)$$

Want to learn "on the go" – e.g. tracking
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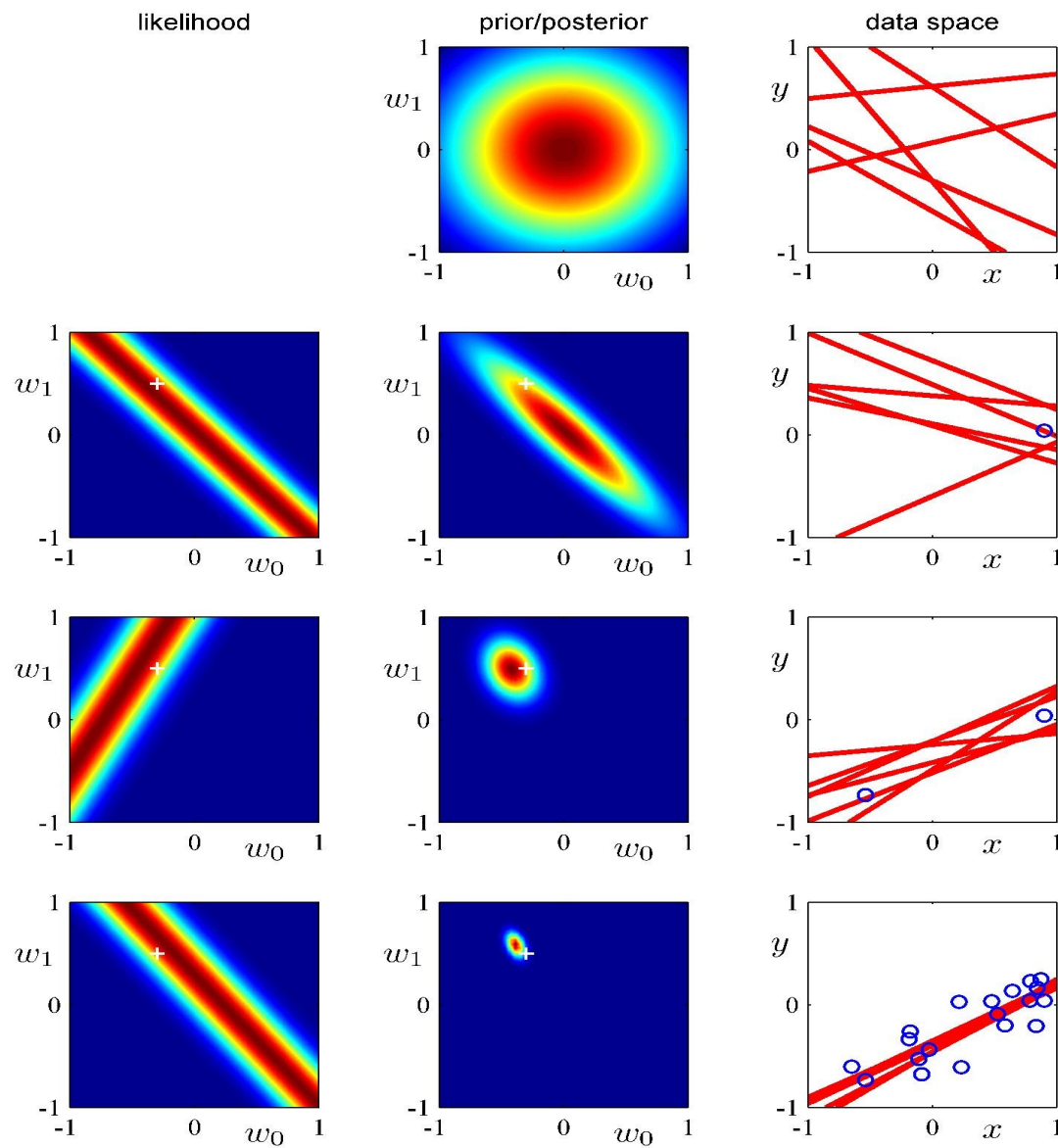
$$= \underbrace{p(t_N | x_N, \mathbf{w}, \beta)}_{\text{likelihood for } t_N} \underbrace{p(\mathbf{w})}_{\text{original prior}} \underbrace{\prod_{n=1}^{N-1} p(t_n | \mathbf{x}_n, \mathbf{w}, \beta)}_{\text{likelihood of observing } t_1, \dots, t_{N-1}}$$

posterior for $N - 1$ = prior for N

Posterior for $N - 1$ acts as prior for N

Sequential learning

Example from CB



Summary

Three views on regression

Geometric least squares

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$

Summary

Three views on regression

Geometric least squares

Maximum Likelihood

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$
 $= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$

Tend to overfit \rightsquigarrow regularization $= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^2$

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$
 $= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Tend to overfit \rightsquigarrow regularization

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^2$



MAP \Leftrightarrow ML + L_2 regularizer

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$
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MAP \Leftrightarrow ML + L_2 regularizer

Alternative regularizers:

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^q$

Built-in dimensionality reduction / feature selection

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$
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Built-in dimensionality reduction / feature selection

No trivial Bayesian counterpart

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$
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Frequentist

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$

Tend to overfit \rightsquigarrow regularization

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^2$

Alternative regularizers:

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^q$

Built-in dimensionality reduction / feature selection

No trivial Bayesian counterpart

Frequentist

$= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Bayesian

MAP \Leftrightarrow ML + L_2 regularizer

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood \Leftrightarrow Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$

Tend to overfit \rightsquigarrow regularization

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Alternative regularizers:

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Built-in dimensionality reduction / feature selection

No trivial Bayesian counterpart

Frequentist

$= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Bayesian

MAP \Leftrightarrow ML + L_2 regularizer

**For solving real problems,
ask yourself questions like:**

- Do I have a good prior (prior knowledge can be better than standard regularizer)?
- Is my data normally (or similarly nicely) distributed?
- Do I need a sparse regularizer?

Summary

Three views on regression

Geometric least squares \Leftrightarrow Maximum Likelihood Maximum a Posteriori

Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2$ Find \mathbf{w} that maximizes $p(\mathbf{t}|\mathbf{x}, \mathbf{w})$ Find \mathbf{w} that maximizes $p(\mathbf{w}|\mathbf{x}, \mathbf{t})$

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Minimize $\sum_{n=1}^N (y(\mathbf{x}_n, \mathbf{w}) - t_n)^2 + \|\mathbf{w}\|^q$

Built-in dimensionality reduction / feature selection

No trivial Bayesian counterpart

Frequentist

$= p(\mathbf{t}|\mathbf{w}, \mathbf{x})p(\mathbf{w})$

Bayesian

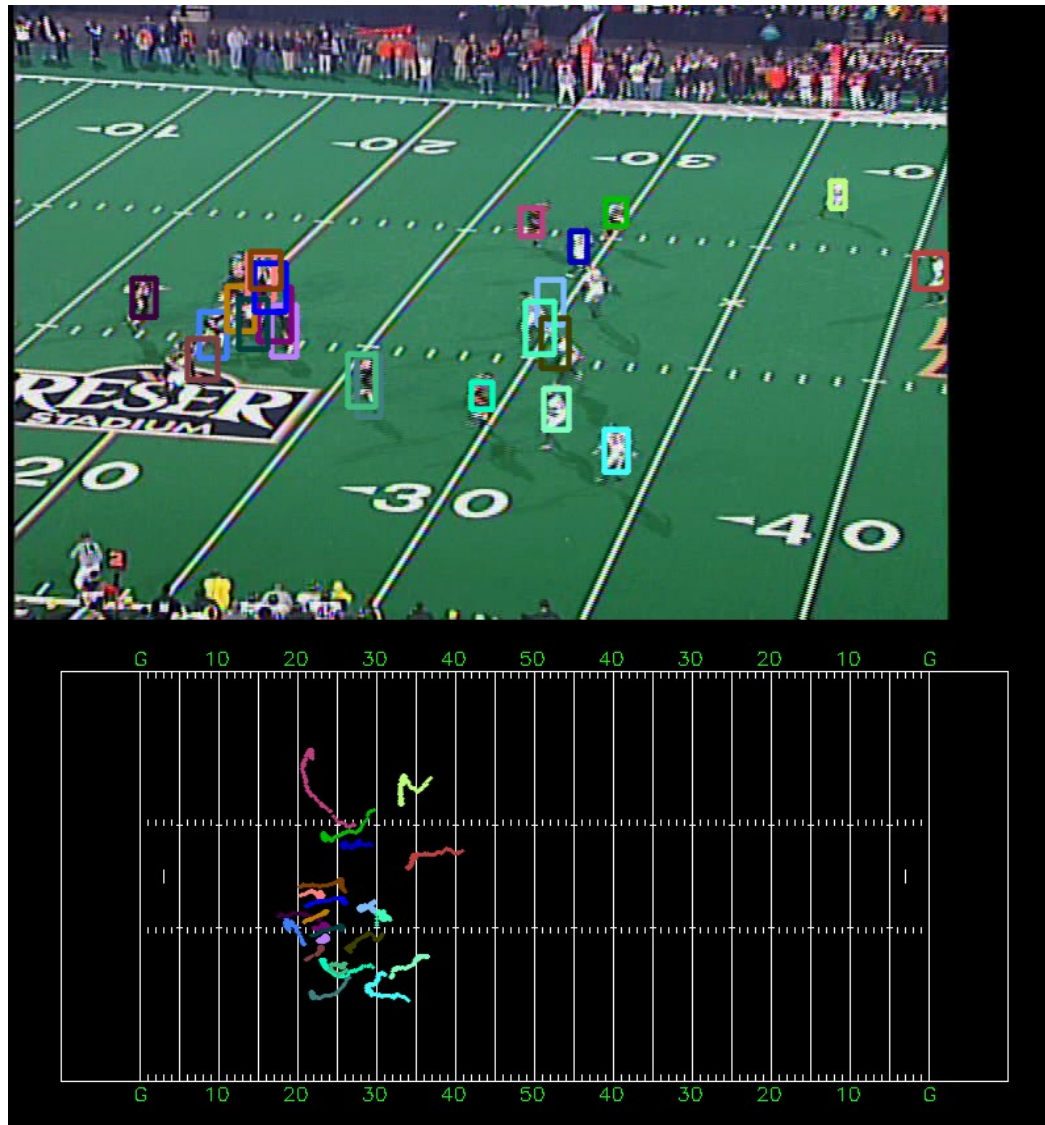
MAP \Leftrightarrow ML + L_2 regularizer

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- Is my data normally (or similarly nicely) distributed?
- Do I need a sparse regularizer?

Scientist

Case: Tracking humans in video



After today's lecture you should:

- Be able to produce a regularized maximum likelihood solution to a linear regression model
- Be able to produce a maximum a posteriori solution to a linear regression model
- Understand the relation between maximum a posteriori solutions and regularized maximum likelihood solutions
- Be familiar with different choices of regularization of why you would want to use them
- Understand the curse of dimensionality and its impact on solving regression problems
- Understand the effect of choice of prior in MAP estimates for different problems
- Be able to recognize and pose practical regression problems
- Reading material: CB 138-147, 152-156.

Next time!

- Neural networks (Christian)
- CB sections 5.1 - 5.3.3