

# A Few Variables Capture over 70% of State-Level Variance in U.S. Covid Deaths

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## **Abstract**

A U.S. state-level analysis of factors associated with covid-19 deaths through early June 2021 reveals inequality (as defined by the Gini coefficient) to be far and away the strongest single-variable predictor, capturing 40% of variance in covid deaths and 49% of variance in all-cause excess deaths since the start of the pandemic. A linear regression model with 5 independent variables accounts for over 70% of variation in covid deaths, as does a 4-variable linear regression model for all-cause excess deaths. Similar models for covid and all-cause excess deaths since October 1 achieve similar results. Coefficients are highly significant ( $p < 0.01$ ) in almost all cases. A consistent finding across all 4 models is that a state's relative humidity is strongly associated with fewer deaths after controlling for other factors. Lockdown stringency (as measured by the Oxford stringency index) is also strongly associated with fewer deaths for all models. Other significant factors for some models include population density, nursing home resident density, voting patterns shifting towards Donald Trump vs. previous Republican candidates, and share of population under 18 years old. The models pass various robustness checks. The results are reproducible via an open-access data repository and Python notebook made available online.

# 1 Introduction

Covid-19 death rates vary widely across U.S. states, ranging from 297 per 100K population in New Jersey to below 50 per 100K in Alaska, Vermont and Hawaii as of 6/16/2021 [6]. The reasons for this variation remain hotly debated within the public sphere and many observers claim that the variation seems largely random. A recent article in a public affairs online magazine acknowledges some evidence for the effectiveness of lockdown measures but also observes that looking “at a list of states by their number of Covid-19 deaths per capita, it's hard to discern much of a pattern ” [5]. A leading covid modeller, when looking at state-level data, said the data supports the theory that the virus is “ inherently unpredictable ” [12] and “ the lack of correlations here suggest luck + randomness may play a large role ” [13]. The 45th president of the United States recently issued a statement making the claim “blue state lockdowns didnt work ” [21].

The effectiveness of lockdown measures and (more generally) the reasons for US state-level covid death variation are of vital importance for a number of reasons. Helping answer these questions can inform policy decisions for the current, covid-19 pandemic as well as for future pandemics involving pathogens with similar characteristics. If the evidence were to show that lockdowns were indeed ineffective, their undesirable side effects could be avoided in the future. On the other hand, evidence of efficacy of lockdowns obviously strengthens the case for their continued implementation in the future. Additionally, if the perception that lockdowns were ineffective is widespread but is in fact incorrect, confidence in government and public health institutions may be undermined unnecessarily. Public trust in public health measures is perhaps the area where a state-level analysis can make the greatest contribution. There is already a huge literature on the effectiveness of covid mitigation measures, but faith in the necessity of lockdown measures could still be bolstered by clear explanations, grounded in data and rigorous analysis, for the wide variability in covid death rates across U.S. states. Accurate assessments of cost-benefit tradeoffs of lockdown decisions can also help inform future elections, as voters attempt to evaluate the decisions of elected officials. Beyond simply looking at lockdowns, there is value in understanding which potential confounders are significant. Factors which may at first blush appear to be uncontrollable environmental characteristics (e.g. relative humidity) may in fact be subject to manipulation in some situations (e.g. indoors) which could help limit disease spread.

The public debates on the topic often lack rigor. For instance, Florida and California are commonly compared, but this level of analysis involves an  $N$  of 2 and typically little or no attempt to control for confounding variables aside from perhaps a casual appeal to similar temperatures. More rigorous statistical analysis is needed in order to inform such discussions. Some examples of such research already exist, such as [7]. This paper builds on such previous work. In addition to presenting analysis and results, another goal of the paper is to advertise a public data repository [18], which builds on a previously released repository presented at [10]. A Python notebook for reproducing the results in this paper is also available at [18].

The author acknowledges that modern academic standards for causal inference from data are high. No claim is being made that any of the results in this paper constitute rigorous proof of causal effects. Nonetheless, the explanatory variables have considerable causal plausibility and the statistical associations are highly significant after controlling for several confounders. It is possible that some or all of the associations presented here are not causal, but in that case, it is likely that knowledge of the dynamics of covid-19 spread will nonetheless be enhanced by explaining why these statistical associations exist.

**Outline** The remainder of this article is organized as follows. Section 2 describes prior, related research. Data collection and modeling choices are described in Section 3. Regression results and robustness checks are presented in Section 4. Possible reasons for the statistical associations are discussed in Section 5. Future research directions and broader lessons drawn from the work are in Section 6.

## 2 Previous work

State-level variations in covid deaths through Dec 1, 2020 are studied in [7]. Poverty and a few different measures of population density are found to be associated with more per-population covid deaths, while lockdown stringency is found to be negatively associated. Note that while the author of [7] also wrote an op-ed in the Orange County Register praising Florida’s approach as compared to California’s [8], [7] does not include humidity as an explanatory variable. This paper builds on the findings in [7] in various ways. Inequality is presented as a variable which captures far more variance in covid deaths

than poverty. Additionally, relative humidity and other explanatory variables not considered in [7] are found to be significant here. All-cause excess death models are presented as well as covid death models. Various checks for robustness and spatial correlation of residuals are presented. More recent data is used.

More informally, a recent Twitter thread [11] from a leading covid spread modeller looked at correlations between pairs of variables from the data source [10]. Substantial correlation between lockdown stringency and increases in unemployment rates was observed, while minimal correlation between covid death rates and any of the other variables examined was found. This analysis, however, did not include any multivariate regressions, nor did it look at metrics such as inequality, poverty, humidity, or changes in voting share for Trump vs. previous Republican presidential candidates. Furthermore, the District of Columbia was included in the correlation calculations even though DC has 9X the population density of any state.

The significance of humidity for airborne viral transmission has been noted in much research, for instance, [23] [15] and in a Twitter thread by a leading aerosol scientist [14]. The covid virus belongs to a class of viruses whose survival benefits from low humidity. Research indicates that high humidity leads to larger, heavy aerosols which fall out of the air sooner than in low-humidity conditions.

An association between inequality and covid mortality was noted early in the pandemic at the state level in [22] and at the county level in [9].

The state-level correlation between gini coefficient and covid deaths reported here was discovered contemporaneously by JM Pienaar and reported on Twitter at [16] a few days prior to the publication of the working paper version of this work.

### 3 Methods

39 variables (listed in table 1) measuring various socioeconomic, environmental and lockdown-policy traits of each of the 50 U.S. states (as well as DC) were gathered. 12 of the variables were taken from [10], a previously created repository. The remainder of the variables were collected from various publicly available sources, government and otherwise. The data is available for download in a spreadsheet at [18]. Although data for DC was collected, most of the analysis in the paper is performed with DC excluded. As [7] found,

population density is a crucial variable for modeling covid deaths since the start of the pandemic and DC has 9X the population density of the densest state (New Jersey). Such an extreme outlier renders population density essentially useless in a regression with  $N=51$ .

The meaning of most of the variables should be either self-evident from the variable name or easily discernible from following the web links in table 1. Clarifications, if necessary, can be obtained by contacting the author. However, a few of the variables which are central to the results of the paper are described here. The 'stringency\_index' variable is the average value of the Oxford stringency index [19] since March 1, 2020. The stringency index is a composite measure [3] which merges various policies such as bans on public gatherings, school closings, workplace closings, stay-at-home requirements and others into an overall one-number measure of lockdown stringency. The 'gini\_inequality' variable is the Gini coefficient [1] of each state, obtained from [4], the 2019 edition of the American Community Survey. The Gini coefficient summarizes inequality in a mathematically graceful way by using the entire income distribution rather than arbitrary cutoffs such as top 1%, top 10%, bottom 10%, etc. Trump16McCain08Shift is the difference between the share of the presidential vote received by Donald Trump in 2016 and the share received by John McCain in 2008. This measure, however imperfect, is intended to detect enthusiasm in the state's populace for Donald Trump specifically rather than general Republican lean of the state. The results are nearly identical if Mitt Romney's share in 2012 is used rather than McCain's share in 2008. However, Romney's share of the vote in Utah in 2012 is an outlier- much more of an outlier than McCain's home-state share in Arizona. The results are therefore slightly more robust if McCain in 2008 is used as the baseline.

| Variable                   | Source   |
|----------------------------|--|
| afr_amer_pct               | en.wikipedia.org/wiki/List_of_U.S._states_and_territories_by_African-American_population                             |
| cars_per_capita_pct        | en.wikipedia.org/wiki/List_of_U.S._states_by_vehicles_per_capita   |
| cig_use_pct                | https://www.cdc.gov/statesystem/cigaretteuseadult.html   |
| dem_margin_2020            | Gu github  |
| diabetes_pct               | ccd.cdc.gov/Toolkit/DiabetesBurden/Prevalence  |
| doctors_per_capita         | beckershospitalreview.com/workforce/50-states-ranked-by-most-active-physicians-per-100-000-population.html           |
| gini_inequality            | www.census.gov/content/dam/Census/library/publications/2020/acs/acsbr20-03.pdf                                       |
| high_blood_pressure_pct    | https://www.americashealthrankings.org/explore/annual/measure/Hypertension/state/ALL                                 |
| hispanic_latino_pct        | en.wikipedia.org/wiki/List_of_U.S._states_by_Hispanic_and_Latino_population  |
| income_per_capita          | Gu github  |
| mean_temperature           | Gu github  |
| median_age                 | Gu github  |
| mex_amer_pct               | en.wikipedia.org/wiki/List_of_U.S._states_by_Hispanic_and_Latino_population  |
| McCain08Share              | en.wikipedia.org/wiki/2008_United_States_presidential_election   |
| multi_gen_household_pct    | data.census.gov/cedsci/table?q=multigenerational&g=0100000US.04000.001&tid=ACSDT1Y2010.B11017                        |
| nursing_resid_per_pop      | www.kff.org/other/state-indicator/number-of-nursing-facility-residents   |
| obesity_rate               | Gu github  |
| over_65_pct                | www.prb.org/resources/which-us-states-are-the-oldest/  |
| pct_in_city_min_200k       | en.wikipedia.org/wiki/List_of_United_States_cities_by_population   |
| perc_25plus_with_bachelors | Gu github  |
| perc_blue_collar_jobs      | Gu github  |
| perc_pop_at_least_1_dose   | Gu github  |
| perc_pop_nonwhite          | Gu github  |
| perc_urban                 | Gu github  |
| population_per_sq_m        | Gu github  |
| poverty_rate               | en.wikipedia.org/wiki/List_of_U.S._states_and_territories_by_poverty_rate  |
| regular_church_pct         | www.pewforum.org/religious-landscape-study/compare/attendance-at-religious-services/by/state/                        |
| relative_humidity          | www.forbes.com/sites/brianbrettschneider/2018/08/23/oh-the-humidity-why-is-alaska-the-most-humid-state/?sh=6e6afa6a3 |
| res_per_household          | www.indexmundi.com/facts/united-states/quick-facts/all-states/average-household-size                                 |
| Romney12Share              | en.wikipedia.org/wiki/2012_United_States_presidential_election   |
| seldomornever_church_pct   | www.pewforum.org/religious-landscape-study/compare/attendance-at-religious-services/by/state/                        |
| share_in_apts_pct          | nmhc.org/research-insight/quick-facts-figures/quick-facts-resident-demographics/geography-of-apartment-residents/    |
| stringency_index           | raw.githubusercontent.com/OxCGRT/covid-policy-tracker/master/data/OxCGRT_latest.csv                                  |
| Trump16McCain08Shift       | -  |
| Trump16Share               | en.wikipedia.org/wiki/2016_United_States_presidential_election   |
| under_18_pct               | www.indexmundi.com/facts/united-states/quick-facts/all-states/percent-of-population-under-18                         |
| undocumented_pct           | www.pewresearch.org/hispanic/interactives/u-s-unauthorized-immigrants-by-state/                                      |
| uninsured_pct              | www.beckershospitalreview.com/rankings-and-ratings/states-ranked-by-uninsured-rates.html                             |
| urb_index_538              | fivethirtyeight.com/features/how-urban-or-rural-is-your-state-and-what-does-that-mean-for-the-2020-election/         |

Table 1: Variables and data sources

State-by-state covid death rates and all-cause excess death rates were obtained from publicly-available CDC websites. Covid death rates are as of 6/7/2021. Excess death rates are calculated through 4/1/2021 so as to avoid using premature excess death data from recent months.

The high correlation of Gini-coefficient inequality with covid deaths and excess deaths is one simple result presented in the paper (note: Gini coefficient, gini inequality and the simple term 'inequality' are used interchangeably in this paper). However, the bulk of the paper focuses on multivariate regressions with a handful of variables. Many methods exists for variable selection for multivariate regression. Fully automated one-by-one variable selection such as forward or backward stepwise selection based on p-values has been criticized as invalidating the statistical significance tests on the coefficients [17]. Such procedures can also be path-dependent, i.e., variables can switch from insignificant to significant depending on which other variables

have been included. Therefore, such fully-automated stepwise procedures based on p-values or adjusted R-squared were avoided. Alternatives include cross-validation [17] for confirming the predictive value of each variable and LASSO regression [20] [2], which provides a principled way to examine many variables at the same time, discarding some as irrelevant and retaining others. Prior knowledge and strength of belief in the predictive value of the variables is also often advocated as a guiding rule for deciding which variables to include.

In this paper, the primary model variables recurring in many of the models presented (Gini coefficient, stringency index, relative humidity, population density, Trump-McCain shift and nursing home residents as a share of population) were arrived at as follows. The model in [7], where population density, stringency and poverty were found to be significant, was taken as a starting point. Data for additional variables was collected and an informal search for additional variables with highly significant p-values and which increased R-squared ( as measured by both adjusted R-squared and cross-validation-estimated R-squared ) were sought. In particular, the gini coefficient relationship was discovered by adding `income_per_capita` to a model with poverty and finding a positive association with income after controlling for poverty and other factors. Subsequently, for confirmation purposes, a lasso regression with a cross-validation-optimized penalization term was conducted on the full set of 39 variables and found to give very similar results in terms of variable selection as the previous criteria had yielded.

Standard errors which are robust to heteroscedasticity have become standard usage within many social science disciplines and therefore the regressions results presented here use robust standard errors. However, the results do not change much when using traditional standard errors. The notebook code supplied allows the user to obtain standard errors either way.

Most of the regression results presented are ordinary least squares results with each of the 50 states equally weighted. Alternatively, a weighted least squares in which the states are weighted by population could be performed. However, it is far from clear that weighting by population is appropriate. The infections occurring in a state are not independent events, since individuals are infected by nearby individuals. Therefore, the state-level residual variance would not necessarily decrease like the reciprocal of the state's population. It also seems likely that idiosyncratic features of covid outbreaks in particular states (for instance, the local presence of a particular virus variant) contribute much more to the residual of the state than would the sampling-

style errors which decrease with population. Nonetheless, in section 4.1, a population-weighted least squares is presented for the purposes of confirming that the choice of whether to weight by population does not change the results much.

Although many public resources display results in terms of deaths per 100k population, the decision was made to present regression results here in terms of deaths per 330 million, i.e. deaths per total U.S. population. This mode of presentation makes the impact of the factors more relatable, since the big-picture headline U.S. death total of approximately 600K as of mid-June, 2021 is well known. Deaths per 330 million vs. per 100k is merely a matter of applying a different scale factor to the numbers, of course, and the choice does not change the substance of results in any way.

All independent variables are standardized to be have 0 mean and variance 1 (aka z-scored). The suffix '\_stdized' present in the covariate names indicates this standardization.

## 4 Results

The main results of the paper concern multivariate regressions, but the strikingly strong statistical relationships between Gini coefficient and both covid deaths and all-cause excess deaths are worthy of contemplation before digging into the regressions results. Figure 1 shows a scatter plot of deaths/100K vs. Gini coefficient for the 50 U.S. states since the start of the pandemic. Not only is 40% R-squared a substantial amount of explained variance- it is also visually arresting to note how completely empty the lower right corner of the plot is (where high-inequality, low-covid-death states hypothetically would be). The relationship with all-cause excess death percentage, shown in Figure 2, is even stronger (49% R-squared).



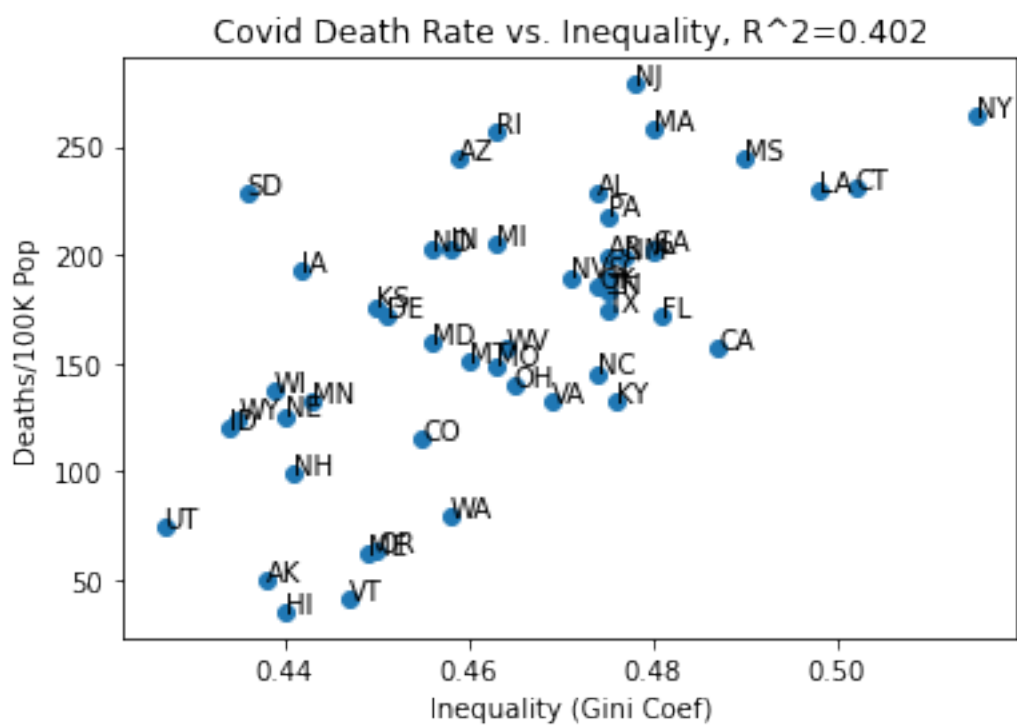


Figure 1: Covid death rate per 100K pop vs. Gini Coefficient

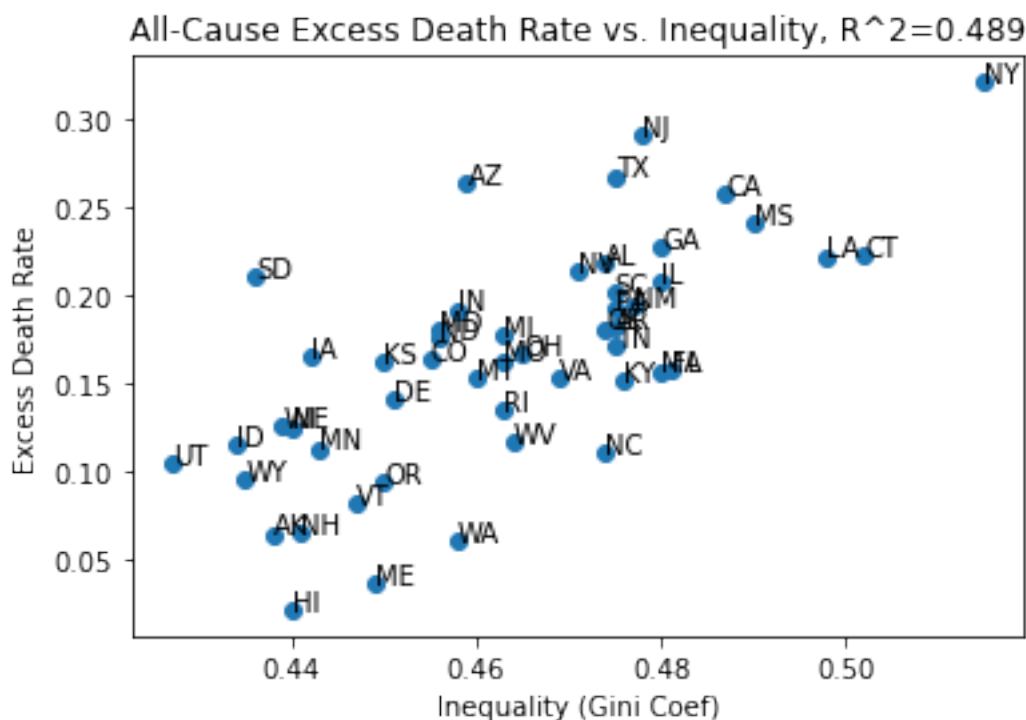


Figure 2: All-Cause Excess Death Rate vs. Gini Coefficient

Figure 3 shows regression results for covid deaths since March 1, 2020. All five factors are significant at the 99.9% level. An adjusted R-squared of 0.76 is achieved. 1 standard deviation in Gini coefficient is associated with 115k additional covid deaths nationally (per 330 million) after controlling for other factors. One standard deviation in population density is associated with 80k additional covid deaths. One standard deviation increase in lockdown stringency index is associated with 77k fewer covid deaths. One standard deviation increase in relative humidity is associated with 64k fewer deaths. One standard deviation increase in nursing home residents per population is associated with 67K additional deaths.

| OLS Regression Results                                  |  |                     |          |       |           |           |
|---|--|---------------------|----------|-------|-----------|-----------|
| Dep. Variable:  | deaths_since_march1_2020_per330million | R-squared:          | 0.786    |       |           |           |
| Model:  | OLS                                    | Adj. R-squared:     | 0.762    |       |           |           |
| Method:   | Least Squares                          | F-statistic:        | 47.82    |       |           |           |
| Date:   | Thu, 17 Jun 2021                       | Prob (F-statistic): | 1.08e-16 |       |           |           |
| Time:   | 16:40:26                               | Log-Likelihood:     | -642.68  |       |           |           |
| No. Observations:                                       | 50                                     | AIC:                | 1297.    |       |           |           |
| Df Residuals:   | 44                                     | BIC:                | 1309.    |       |           |           |
| Df Model:   | 5                                      |                     |          |       |           |           |
| Covariance Type:  | HC1                                    |                     |          |       |           |           |
|   | coef                                   | std err             | t        | P> t  | [0.025    | 0.975]    |
| Intercept   | 5.486e+05                              | 1.39e+04            | 39.355   | 0.000 | 5.21e+05  | 5.77e+05  |
| stringency_index_stdized                                | -7.663e+04                             | 1.77e+04            | -4.320   | 0.000 | -1.12e+05 | -4.09e+04 |
| relative_humidity_stdized                               | -6.395e+04                             | 1.83e+04            | -3.487   | 0.001 | -1.01e+05 | -2.7e+04  |
| nursing_resid_per_pop_stdized                           | 6.675e+04                              | 1.56e+04            | 4.282    | 0.000 | 3.53e+04  | 9.82e+04  |
| gini_inequality_stdized                                 | 1.149e+05                              | 1.51e+04            | 7.605    | 0.000 | 8.45e+04  | 1.45e+05  |
| population_per_sq_mi_stdized                            | 7.99e+04                               | 1.23e+04            | 6.470    | 0.000 | 5.5e+04   | 1.05e+05  |
| Omnibus:  | 1.791                                  | Durbin-Watson:      | 1.412    |       |           |           |
| Prob(Omnibus):  | 0.408                                  | Jarque-Bera (JB):   | 1.196    |       |           |           |
| Skew:   | 0.027                                  | Prob(JB):           | 0.550    |       |           |           |
| Kurtosis:   | 2.244                                  | Cond. No.           | 2.13     |       |           |           |
| Notes:  |  |                     |          |       |           |           |
| [1] Standard Errors are heteroscedasticity robust (HC1) |  |                     |          |       |           |           |

Figure 3: Regression results for Covid deaths since 3/1/2020

Figure 4 shows regression results for all-cause excess death percentage since March 1, 2020. One standard deviation increase in Gini coefficient is associated with a 5% increase in all-cause excess death percentage nationally. One standard deviation increase in lockdown stringency is associated with a 2.8% decrease nationally. One standard deviation increase in relative humidity is associated with a 2.1% decrease nationally. All 3 factors are significant at the 99.9% level. One standard deviation increase in population density is associated with a 1.4% increase in all-cause excess death percentage nationally. Population density is only significant at the 95% level. Population density was included because it decreases the spatial correlation of the residuals to a minimal level. It is unclear why nursing home population does not show up as significant here, but one possibility is that a portion of nursing home deaths were deaths of people who would otherwise have lived less than 15 additional months. Such deaths would not impact all-cause excess deaths measured in totality since March 1, 2020.

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=====
                        OLS Regression Results
=====
Dep. Variable:      pctExcessAfterMarch1_2020      R-squared:                0.737
Model:              OLS                          Adj. R-squared:           0.713
Method:             Least Squares                 F-statistic:              38.06
Date:               Thu, 17 Jun 2021               Prob (F-statistic):       6.66e-14
Time:               17:28:30                       Log-Likelihood:           99.843
No. Observations:   50                             AIC:                     -189.7
Df Residuals:       45                             BIC:                     -180.1
Df Model:            4
Covariance Type:    HC1

=====
                        coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept              0.1639         0.005     33.467      0.000         0.154         0.174
gini_inequality_stdized  0.0497         0.006      8.551      0.000         0.038         0.061
stringency_index_stdized -0.0288         0.005     -5.275      0.000        -0.039        -0.017
relative_humidity_stdized -0.0287         0.005     -4.563      0.000        -0.038        -0.012
population_per_sq_mi_stdized  0.0144         0.006      2.213      0.032         0.001         0.027
=====
Omnibus:              1.098      Durbin-Watson:           2.448
Prob(Omnibus):        0.577      Jarque-Bera (JB):        0.921
Skew:                 0.005      Prob(JB):                0.631
Kurtosis:             2.335      Cond. No.                1.76
=====

```

Figure 4: Regression results for all-cause excess deaths since 3/1/2020

Figure 5 shows regression results for covid deaths since October 1. The same stringency, humidity and Gini coefficient factors which appear in the covid death model since the start of the pandemic show up again here. Another significant factor is the difference between the share of vote received by Donald Trump in the 2016 presidential election vs. the share of vote for John McCain in 2008.

| OLS Regression Results       |                                      |                     |          |       |           |           |
|------------------------------|--------------------------------------|---------------------|----------|-------|-----------|-----------|
| Dep. Variable:               | deaths_since_oct1_2020_per330million | R-squared:          | 0.612    |       |           |           |
| Model:                       | OLS                                  | Adj. R-squared:     | 0.578    |       |           |           |
| Method:                      | Least Squares                        | F-statistic:        | 21.32    |       |           |           |
| Date:                        | Thu, 17 Jun 2021                     | Prob (F-statistic): | 6.46e-10 |       |           |           |
| Time:                        | 16:45:49                             | Log-Likelihood:     | -633.45  |       |           |           |
| No. Observations:            | 50                                   | AIC:                | 1277.    |       |           |           |
| Df Residuals:                | 45                                   | BIC:                | 1286.    |       |           |           |
| Df Model:                    | 4                                    |                     |          |       |           |           |
| Covariance Type:             | HC1                                  |                     |          |       |           |           |
|                              | coef                                 | std err             | t        | P> t  | [0.025    | 0.975]    |
| Intercept                    | 3.745e+05                            | 1.15e+04            | 32.673   | 0.000 | 3.51e+05  | 3.98e+05  |
| stringency_index_stdized     | -6.34e+04                            | 1.4e+04             | -4.528   | 0.000 | -9.16e+04 | -3.52e+04 |
| relative_humidity_stdized    | -4.45e+04                            | 1.79e+04            | -2.480   | 0.017 | -8.06e+04 | -8356.820 |
| Trump16McCain08Shift_stdized | 5.464e+04                            | 1.08e+04            | 5.054    | 0.000 | 3.29e+04  | 7.64e+04  |
| gini_inequality_stdized      | 5.667e+04                            | 9790.115            | 5.788    | 0.000 | 3.69e+04  | 7.64e+04  |
| Omnibus:                     | 1.496                                | Durbin-Watson:      | 1.648    |       |           |           |
| Prob(Omnibus):               | 0.473                                | Jarque-Bera (JB):   | 1.090    |       |           |           |
| Skew:                        | 0.362                                | Prob(JB):           | 0.580    |       |           |           |
| Kurtosis:                    | 3.014                                | Cond. No.           | 1.47     |       |           |           |

Figure 5: Regression results for Covid deaths since 10/1/2020

Figure 6 shows regression results for all-cause excess deaths since October 1. In addition to the 4 variables included for covid deaths since October 1, one standard deviation increase in the share of the population under 18 is associated with a 3% increase in excess death percentage. Unlike the other factors, no obvious causal mechanism for the under-18-share factor is apparent to the author of the paper, but it shows up as highly significant. The association may be due to some as-yet-undetermined confounder, but in any case, the association is worth further study.

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=====
                        OLS Regression Results
=====
Dep. Variable:      pctExcessAfterOct1_2020      R-squared:                0.725
Model:              OLS                        Adj. R-squared:           0.693
Method:              Least Squares              F-statistic:             30.15
Date:                Thu, 17 Jun 2021            Prob (F-statistic):       3.59e-13
Time:                17:36:25                    Log-Likelihood:           87.745
No. Observations:    50                        AIC:                     -163.5
Df Residuals:        44                        BIC:                     -152.0
Df Model:            5
Covariance Type:     HC1

=====
                        coef      std err          t      P>|t|      [0.025      0.975]
-----
Intercept              0.2049        0.006     32.481      0.000        0.192        0.218
stringency_index_stdized -0.0295        0.006    -4.559      0.000       -0.042       -0.016
gini_inequality_stdized  0.0451        0.007     6.341      0.000        0.031        0.059
relative_humidity_stdized -0.0342        0.007    -4.672      0.000       -0.049       -0.019
Trump16McCain08Shift_stdized 0.0295        0.009     3.426      0.001        0.012        0.047
under_18_pct_stdized    0.0309        0.007     4.131      0.000        0.016        0.046

=====
Omnibus:              10.804    Durbin-Watson:           2.128
Prob(Omnibus):        0.005    Jarque-Bera (JB):        12.159
Skew:                 0.808    Prob(JB):                 0.00229
Kurtosis:             4.795    Cond. No.                 2.56

=====
Notes:
[1] Standard Errors are heteroscedasticity robust (HC1)
=====

```

Figure 6: Regression results for all-cause excess deaths since 10/1/2020

## 4.1 Robustness Checks

Any regression conducted on data points with spatial relationships should check for spatial correlation of the residuals. Correlation of residuals invalidates the statistical significance tests of ordinary least squares, which relies on an assumption of i.i.d gaussian residuals. To check for spatial correlation, the correlation of each state's residual with the average residual of the states which border that state was computed and is shown in 4.1. Residual correlation with bordering states is minimal and is in fact slightly negative for 3 of the 4 models. The slight negative correlation is probably only small-sample noise but nonetheless strengthens evidence that positive spatial correlation is not a concern here. More formal spatial correlation tests may be conducted in future versions of this work. Figure 7 shows a color-coded representation of residuals for the covid-deaths-since-March1-2020 model, with blue indicating large negative residuals, red indicating large positive and purple indicating small magnitude residuals. No spatial pattern is evident. Analogous maps for the 3 other models are omitted for space consideration reasons but can be obtained by running the notebook code.

| Model                          | Corr of Resid with Avg Border State Resid |
|--------------------------------|---|
| covid deaths since March 1     | -0.04                                     |
| excess death pct since March 1 | 0.15                                      |
| covid deaths since Oct 1       | -0.16                                     |
| excess deaths since Oct 1      | -0.22                                     |

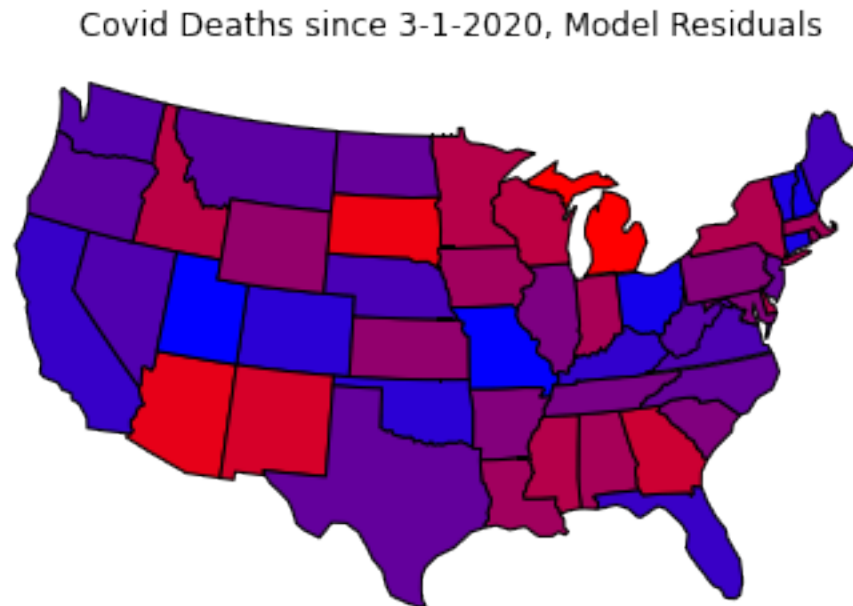


Figure 7: Model Resids, blue implies large neg, red implies large pos

When dealing with small sample sizes, additional confidence in regression results can be obtained by confirming that the results are not dependent on any individual data point. Accordingly, each of the 4 regressions were re-run with each of the 50 states removed and the highest p-value for any state removal was recorded for each coefficient. Tables 2, 3, 4, 5 display the results. In almost all cases, the coefficients remain highly significant with the removal of any state. The one major exception is that the significance of population density for all-cause excess since 3/1/2020 is not robust to the removal of New Jersey. While the intuitive case for the significance of population density for covid spread is clear, a case could be made that it cannot be distinguished

with high confidence from a story in which the densely populated northeastern states simply had the bad luck of getting hit hard with covid early on, when less was known about covid treatment, particularly given that density does not show up in the since-October models. While humidity is only significant at the 90% level for covid deaths since October 1 after removing Arizona, the overall case for the significance of humidity seems strong, given that it remains highly significant with the removal of any state for the other 3 models.

| Variable                      | Max PValue | State Removed |
|-------------------------------|------------|---------------|
| stringency_index_stdized      | 0.001      | HI            |
| relative_humidity_stdized     | 0.013      | NM            |
| nursing_resid_per_pop_stdized | 0.000      | AL            |
| gini_inequality_stdized       | 0.000      | AL            |
| population_per_sq_mi_stdized  | 0.000      | AL            |

Table 2: Max P\_values with any state removed, covid deaths since 3/1/2020

| Variable                      | Max PValue | State Removed |
|-------------------------------|------------|---------------|
| stringency_index_stdized      | 0.000      | AL            |
| relative_humidity_stdized     | 0.000      | NM            |
| nursing_resid_per_pop_stdized | 0.000      | AL            |
| gini_inequality_stdized       | 0.000      | AL            |
| population_per_sq_mi_stdized  | 0.199      | NJ            |

Table 3: Max P\_values with any state removed, excess death pct since 3/1/2020

| Variable                     | Max PValue | State Removed |
|------------------------------|------------|---------------|
| stringency_index_stdized     | 0.001      | HI            |
| relative_humidity_stdized    | 0.074      | AZ            |
| Trump16McCain08Shift_stdized | 0.000      | AL            |
| gini_inequality_stdized      | 0.000      | AL            |

Table 4: Max P\_values with any state removed, covid deaths since 10/1/2020



| Variable                     | Max PValue | State Removed |
|------------------------------|------------|---------------|
| stringency_index_stdized     | 0.001      | HI            |
| relative_humidity_stdized    | 0.001      | AZ            |
| Trump16McCain08Shift_stdized | 0.015      | UT            |
| gini_inequality_stdized      | 0.000      | AL            |
| under_18_pct_stdized         | 0.003      | NH            |

Table 5: Max Pvalues with any state removed, excess deaths since 10/1/2020

Alaska and Hawaii are both very high humidity, low covid death states and are arguably also outliers from the point of view of the difficulty of travel to and from those 2 states, so it is worth verifying whether the results are robust to removal of those 2 states. Figure 8 shows regression results for covid deaths since 3/1/2020 with both states removed. The coefficients are nearly identical to those with the states included.

| OLS Regression Results                                  |  |                     |          |       |           |           |
|---|--|---------------------|----------|-------|-----------|-----------|
| Dep. Variable:  | deaths_since_march1_2020_per330million | R-squared:          | 0.742    |       |           |           |
| Model:  | OLS                                    | Adj. R-squared:     | 0.711    |       |           |           |
| Method:   | Least Squares                          | F-statistic:        | 29.48    |       |           |           |
| Date:   | Fri, 18 Jun 2021                       | Prob (F-statistic): | 1.01e-12 |       |           |           |
| Time:   | 14:01:03                               | Log-Likelihood:     | -617.85  |       |           |           |
| No. Observations:                                       | 48                                     | AIC:                | 1248.    |       |           |           |
| Df Residuals:   | 42                                     | BIC:                | 1259.    |       |           |           |
| Df Model:   | 5                                      |                     |          |       |           |           |
| Covariance Type:  | HCl                                    |                     |          |       |           |           |
|   | coef                                   | std err             | t        | P> t  | [0.025    | 0.975]    |
| Intercept   | 5.657e+05                              | 1.45e+04            | 38.925   | 0.000 | 5.36e+05  | 5.95e+05  |
| stringency_index_stdized                                | -7.31e+04                              | 1.92e+04            | -3.802   | 0.000 | -1.12e+05 | -3.43e+04 |
| relative_humidity_stdized                               | -6.666e+04                             | 2.04e+04            | -3.269   | 0.002 | -1.08e+05 | -2.55e+04 |
| nursing_resid_per_pop_stdized                           | 6.732e+04                              | 1.73e+04            | 3.900    | 0.000 | 3.25e+04  | 1.02e+05  |
| gini_inequality_stdized                                 | 1.143e+05                              | 1.62e+04            | 7.066    | 0.000 | 8.17e+04  | 1.47e+05  |
| population_per_sq_mi_stdized                            | 8.126e+04                              | 1.26e+04            | 6.457    | 0.000 | 5.59e+04  | 1.07e+05  |
| Omnibus:  | 2.261                                  | Durbin-Watson:      | 1.335    |       |           |           |
| Prob(Omnibus):  | 0.323                                  | Jarque-Bera (JB):   | 1.346    |       |           |           |
| Skew:   | 0.045                                  | Prob(JB):           | 0.510    |       |           |           |
| Kurtosis:   | 2.185                                  | Cond. No.           | 2.19     |       |           |           |
| Notes:  |  |                     |          |       |           |           |
| [1] Standard Errors are heteroscedasticity robust (HCl) |  |                     |          |       |           |           |

Figure 8: regression results for covid deaths since 3/1/2020 with Hawaii and Alaska removed.

Figure 9 shows the results of a weighted least squares regression of covid deaths since 3/1/2020 with weights equal to state population. Here x\_1 is stringency\_index, x\_2 is relative\_humidity, x\_3 is nursing\_resid\_per\_pop, x\_4 is

gini\_inequality and x\_5 is population\_per\_sq\_mi, with all variables standardized (\_stdized) as in the other regressions. The results are quite similar to those presented in 3. The theoretically ideal weighting of the 50 data points is unclear and seems likely to lie somewhere between constant weighting and weighting by population. The results in 9 suggest the conclusions in the paper are unlikely to change materially, regardless of weighting choice.

```

=====
                        WLS Regression Results
=====
Dep. Variable:          y      R-squared:          0.697
Model:                  WLS    Adj. R-squared:      0.663
Method:                  Least Squares    F-statistic:    20.28
Date:                    Fri, 25 Jun 2021    Prob (F-statistic): 1.91e-10
Time:                    13:30:28    Log-Likelihood:    -653.35
No. Observations:        50    AIC:              1319.
Df Residuals:            44    BIC:              1330.
Df Model:                 5
Covariance Type:         nonrobust
=====
               coef      std err          t      P>|t|      [0.025      0.975]
-----
const          5.404e+05    1.63e+04    33.122    0.000    5.07e+05    5.73e+05
x1             -7.229e+04    1.98e+04    -3.648    0.001   -1.12e+05   -3.24e+04
x2             -5.746e+04    1.82e+04    -3.163    0.003   -9.41e+04   -2.08e+04
x3              7.034e+04    1.88e+04    3.740    0.001    3.24e+04    1.08e+05
x4              1.128e+05    1.76e+04    6.401    0.000    7.73e+04    1.48e+05
x5              7.491e+04    1.74e+04    4.313    0.000    3.99e+04    1.1e+05
=====
Omnibus:                  1.957    Durbin-Watson:      1.595
Prob(Omnibus):            0.376    Jarque-Bera (JB):    1.099
Skew:                     0.185    Prob(JB):            0.577
Kurtosis:                 3.625    Cond. No.             2.64
=====

Notes:
[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

```

Figure 9: population-weighted regression results for covid deaths since 3/1/2020

Age-adjustment is a commonly discussed topic when comparing covid deaths rates across states. Figure 10 shows the results of a regression where age-adjusted covid death rates since 3/1/2020 are used as a dependent variable. The coefficients are similar to those where unadjusted covid death rates were used, with the coefficient associated with the stringency index being even higher.

```

ageOnlyAdjDeathRate_per330million~stringency_index_stdized+relative_humidity_stdized+nursing_resid_per_pop_stdized+gini_ineq
uality_stdized+population_per_sq_mi_stdized

=====
OLS Regression Results
=====
Dep. Variable:    ageOnlyAdjDeathRate_per330million    R-squared:        0.796
Model:            OLS                                Adj. R-squared:    0.772
Method:            Least Squares                      F-statistic:       41.28
Date:              Thu, 27 Jan 2022                    Prob (F-statistic): 1.56e-15
Time:              12:58:19                            Log-Likelihood:    -640.18
No. Observations: 50                                  AIC:               1292.
Df Residuals:      44                                  BIC:               1304.
Df Model:           5
Covariance Type:   HCL1
=====
                    coef    std err          t      P>|t|      [0.025      0.975]
-----
Intercept              5.266e+05    1.33e+04    39.713    0.000      5e+05    5.53e+05
stringency_index_stdized -8.798e+04    1.65e+04   -5.329    0.000    -1.21e+05    -5.47e+04
relative_humidity_stdized -6.365e+04    1.28e+04   -4.985    0.000    -8.94e+04    -3.79e+04
nursing_resid_per_pop_stdized 5.233e+04    1.48e+04    3.535    0.001     2.25e+04    8.22e+04
gini_inequality_stdized  1.328e+05    1.55e+04    8.548    0.000    1.01e+05    1.64e+05
population_per_sq_mi_stdized 5.201e+04    1.45e+04    3.576    0.001     2.27e+04    8.13e+04
=====
Omnibus:            1.376    Durbin-Watson:      2.138
Prob(Omnibus):      0.503    Jarque-Bera (JB):    1.290
Skew:               -0.372    Prob(JB):            0.525
Kurtosis:           2.747    Cond. No.            2.12
=====

Notes:
[1] Standard Errors are heteroscedasticity robust (HCL1)

```

Figure 10: Regression results for age-adjusted covid deaths since 3/1/2020

To confirm that each variable adds out-of-sample predictive value to each model, 10-fold cross validated r-squareds were computed after removing each variable for each of the 4 models. Results are shown in tables 6, 7, 8 and 9. All variables add substantial predictive accuracy to all models, with the possible exception of density for excess deaths since 3/1/2020, which was included for the sake of reducing spatial correlation, as previously mentioned.

|                                   |       |
|-----------------------------------|-------|
| None                              | 0.708 |
| stringency_index_stdized          | 0.602 |
| relative_humidity_stdized removed | 0.620 |
| nursing_resid_per_pop_stdized     | 0.612 |
| gini_inequality_stdized           | 0.380 |
| population_per_sq_mi_stdized      | 0.592 |

Table 6: Cross-validated r\_squared with variable removed, covid deaths since 3/1/2020

|                              |       |
|------------------------------|-------|
| None                         | 0.653 |
| stringency_index_stdized     | 0.470 |
| relative_humidity_stdized    | 0.548 |
| gini_inequality_stdized      | 0.074 |
| population_per_sq_mi_stdized | 0.641 |

Table 7: Cross-validated r-squared with variable removed, excess death pct since 3/1/2020

|                                   |       |
|-----------------------------------|-------|
| None                              | 0.474 |
| stringency_index_stdized          | 0.237 |
| relative_humidity_stdized removed | 0.389 |
| gini_inequality_stdized           | 0.255 |
| Trump16McCain08Shift_stdized      | 0.292 |

Table 8: Cross-validated r-squared with variable removed, covid deaths since 10/1/2020

|                                   |       |
|-----------------------------------|-------|
| None                              | 0.602 |
| stringency_index_stdized          | 0.522 |
| relative_humidity_stdized removed | 0.426 |
| gini_inequality_stdized           | 0.216 |
| Trump16McCain08Shift_stdized      | 0.497 |
| under_18_pct_stdized              | 0.546 |

Table 9: Cross-validated r\_squared with variable removed, excess deaths since 10/1/2020

Tables 10, 11, 12, 13 show the coefficients obtained from a lasso regression with cross-validation-optimized penalization parameter. Any of the 39 variables not included in these tables was zeroed-out by the lasso regression. The significances of gini coefficient, lockdown stringency and humidity are confirmed by their consistent appearance among the largest-magnitude coefficients across all 4 models. The one exception is that lasso for covid deaths since October 1 assigns a much bigger coefficient to poverty than gini coefficient. It would be a defensible choice to include poverty rather than gini for this model, but these 2 variables are highly correlated and the overall themes

of the paper would not change much if poverty were swapped in for gini coefficient for that one model. Racial and ethnic population percentages do also show up with substantial coefficients in the lasso regressions. The absence of these variables in the regression models presented in the paper should not be taken as denying that racism (structural or otherwise) has played a significant role in covid death trends in the United States. Alternative models which include some of these factors would certainly be defensible. Structural racism and inequality are intertwined in the United States to a degree that they are difficult to disentangle statistically with a dataset where  $N=50$ . The regression models in this paper were chosen with parsimony as an important criterion.

| <b>Variable</b>               | <b>Lasso coef</b> |
|-------------------------------|-------------------|
| population_per_sq_mi_stdized  | 83956             |
| gini_inequality_stdized       | 61856             |
| stringency_index_stdized      | -59555            |
| nursing_resid_per_pop_stdized | 55570             |
| relative_humidity_stdized     | -47139            |
| poverty_rate_stdized          | 37433             |
| afr_amer_pct_stdized          | 26310             |
| hispanic_latino_pct_stdized   | 14395             |
| Trump16McCain08Shift_stdized  | 12064             |
| perc_urban_stdized            | 5475              |
| mean_temperature_stdized      | -5140             |
| median_age_stdized            | -3609             |
| share_in_apts_pct_stdized     | 502               |

Table 10: Lasso coefficients for covid deaths since 3/1/2020

| <b>Variable</b>                  | <b>Lasso coef</b> |
|----------------------------------|-------------------|
| gini_inequality_stdized          | 0.0356            |
| stringency_index_stdized         | -0.0158           |
| relative_humidity_stdized        | -0.0097           |
| hispanic_latino_pct_stdized      | 0.0089            |
| nursing_resid_per_pop_stdized    | 0.0066            |
| share_in_apts_pct_stdized        | 0.0062            |
| afr_amer_pct_stdized             | 0.0061            |
| under_18_pct_stdized             | 0.0054            |
| population_per_sq_mi_stdized     | 0.0053            |
| seldomornever_church_pct_stdized | -0.0034           |
| mex_amer_pct_stdized             | 0.0022            |

Table 11: Lasso coefficients for excess death pct since 3/1/2020

| <b>Variable</b>                  | <b>Lasso coef</b> |
|----------------------------------|-------------------|
| stringency_index_stdized         | -39371            |
| relative_humidity_stdized        | -32552            |
| nursing_resid_per_pop_stdized    | 30815             |
| Trump16McCain08Shift_stdized     | 29510             |
| hispanic_latino_pct_stdized      | 24536             |
| population_per_sq_mi_stdized     | 23324             |
| poverty_rate_stdized             | 22469             |
| res_per_household_stdized        | -17532            |
| diabetes_pct_stdized             | 14640             |
| seldomornever_church_pct_stdized | -13698            |
| median_age_stdized               | -12782            |
| gini_inequality_stdized          | 2278              |

Table 12: Lasso coefficients for covid deaths since 10/1/2020

| Variable                         | Lasso coef |
|----------------------------------|------------|
| mex_amer_pct_stdized             | 0.0338     |
| stringency_index_stdized         | -0.0273    |
| gini_inequality_stdized          | 0.0232     |
| relative_humidity_stdized        | -0.0147    |
| nursing_resid_per_pop_stdized    | 0.0147     |
| seldomornever_church_pct_stdized | -0.0110    |
| Trump16McCain08Shift_stdized     | 0.0110     |
| obesity_rate_stdized             | 0.0057     |
| under_18_pct_stdized             | 0.0028     |
| median_age_stdized               | -0.0025    |

Table 13: Lasso coefficients for excess death pct since 10/1/2020

## 5 Discussion

The author’s expertise is in machine learning and data science, rather than epidemiology, virology, aerosol science, sociology or political science. The statistical associations are offered up in this paper as a starting point for further study by experts in these various subfields. With that said, some cautious and tentative speculation for the reasons behind the associations are presented in this section.

The consistent association of lockdown stringency with fewer deaths is certainly at least suggestive of a causal effect. The association may not necessarily prove the causal effect specifically of mandates, rules and regulations, though. One possible confounder is that states where it is politically palatable to implement lockdown regulations are states where the population would have been more inclined to follow guidelines such as social distancing voluntarily, even in the absence of mandates and rules. The evidence for the efficacy of social distancing and related measures (whether voluntary or enforced), however, seems strong. Perhaps there is some confounder which would explain the association of stringency with fewer deaths while undermining the case for the efficacy of public health guidelines, but no such confounder is apparent to the author at this time.

The causal pathway from higher humidity to reduced viral transmission seems clear based on research such as [23].

The association between inequality and covid deaths is open to many

interpretations and it is possible that the relationship is not causal. With that said, many plausible explanations exist. Inequality may be associated with an increase in less-affluent workers providing various in-person services for more affluent residents (taxis and Ubers, restaurant waitstaff, house cleaners, hairdressers, etc). Inequality may also be associated with crowded living conditions in poor quality, poorly ventilated residencies. Inequality may also be associated with longer commute times either via public transportation or carpool, because high housing costs could lead to less affluent workers living farther away from their jobs. High-inequality state governments may also simply be less responsive to the needs and well-being of their less-affluent residents.

The reasons that Donald Trump gained vote share relative to previous Republican candidates are much studied and perhaps complex from a socioeconomic point of view. There are many possible causal pathways which could explain the statistical relationship between Trump vote gain and higher covid deaths. Socioeconomic conditions which led to Trump’s appeal may have also caused higher covid deaths. With that said, the rhetoric and messaging of President Trump and his administration regarding social distancing, masks, etc also seems like a plausible causal pathway here.

## 6 Conclusions

There are many possible future directions for this research. The Oxford stringency index can be decomposed into several more specific mitigation measures. Statistical associations (or lack thereof) for these more specific measures are worthy of further study. The time dynamics of lockdown measures could also be studied more. This paper simply looks at the statistical association of average lockdown stringency with deaths over the entire course of the pandemic (or after October 1). Correlations between stringency at one point in time and covid deaths after a time lag (e.g. one month) merit some analysis.

The tools of modern causal inference, beyond just multiple regressions, such as directed acyclic graphs, g-computation and related techniques may also be useful here.

This paper may also serve as a warning against drawing conclusions about the relevance (or lack thereof) of variables based on simple pairwise correlations with a variable of interest. Lockdown stringency and humidity both



show minimal correlation with covid deaths on their own, yet show up as extremely significant, in a very robust way, once a few other variables are included as controls.

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