EE-411: Fundamentals of inference and learning

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Lecture 8: A short course on statistical learning theory

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8.1 Intro

When performing a supervised learning task, one trains the model such that it minimizes the empirical risk. However, what we would like is for the population risk to minimized.

Definition 8.1 (Population Risk) Let $\mathbf{x} \in \mathcal{X}$ (population set), y be the true label and \hat{y} be the estimator,

$$\mathcal{R}_{pop}[\hat{y}] := \mathbb{E}_{\mathbf{x},y}[loss(\hat{y}(\mathbf{x}), y)]$$

Definition 8.2 (Empirical Risk) Let n be the number of observations,

$$\mathcal{R}_{emp}^{(n)}[\hat{y}] := \frac{1}{n} \sum_{i=1}^{n} loss(\hat{y}(\mathbf{x}_i), y_i)$$

By the LLN, we expect for any function \hat{y}

$$\mathcal{R}_{\mathrm{emp}}^{(n)}[\hat{y}] \stackrel{p}{\to} \mathcal{R}_{\mathrm{pop}}[\hat{y}]$$

Therefore, given a "large" dataset, minimizing the empirical risk should suffice (i.e. population risk is minimized as well).

The question now is, how large should n be?

This is the question that statictical learning theory aims to answer. More precisely, it aims to give some insight on the difference between $\mathcal{R}_{pop}[\hat{y}]$ and $\mathcal{R}_{emp}^{(n)}[\hat{y}]$.

For concreteness, we will be using using as loss function, the 0/1 loss:

$$\begin{aligned} \operatorname{loss}(\hat{y}(\mathbf{x}), y) &:= 1 - \mathbf{1}_{\hat{y}(\mathbf{x}) = y} \\ &= \begin{cases} 0, & \text{if } \hat{y}(\mathbf{x}) = y \\ 1, & \text{otherwise} \end{cases} \end{aligned}$$

8.2 For a given function $\hat{y}()$

Theorem 8.3 (Hoeffding bound Reminder) Let $Z_i = 0/1$ a binary variable with i = 1, ...n. Then,

$$\mathbb{P}\left(\frac{1}{n}\sum_{i}Z_{i} \leq \mathbb{E}[Z] - \epsilon\right) \leq \delta \tag{8.1}$$

with $\delta = e^{-2n\epsilon^2}$.

Developing this inequality, we get:

$$1 - \mathbb{P}\left(\frac{1}{n}\sum_{i} Z_{i} \leq \mathbb{E}[Z] - \epsilon\right) \geq 1 - \delta$$

$$\mathbb{P}\left(\frac{1}{n}\sum_{i} Z_{i} \geq \mathbb{E}[Z] - \epsilon\right) \geq 1 - \delta$$

$$\mathbb{P}\left(\mathbb{E}[Z] - \frac{1}{n}\sum_{i} Z_{i} \leq \epsilon\right) \geq 1 - \delta$$

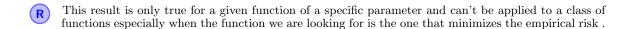
This means that for a random binary variable and for a probability at least δ , the difference between the true mean and the new estimate of the mean is smaller than ϵ . Another way to formulate this result would be to write

$$\delta = e^{-2n\epsilon^2} \to \log \frac{1}{\delta} = 2n\epsilon^2 \to \epsilon = \sqrt{\frac{1}{2n} \log \frac{1}{\delta}}$$

So if the probability \mathbb{P} is at least $1 - \delta$, then

$$\mathbb{E}[Z] - \frac{1}{n} \sum_{i} Z_{i} \le \sqrt{\frac{1}{2n} \log \frac{1}{\delta}} \approx \sqrt{\frac{1}{n}}$$

for a small δ according to PAC learning



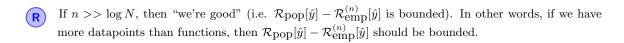
8.3 Finite number of functions

Definition 8.4 (Uniform Convergence) Let $\hat{y}_1, ..., \hat{y}_N$ be N estimators, n be the number of observations (i.e. datapoints) and $\epsilon > 0$.

Then, with probability of at least $1 - \delta$ (with $\delta := e^{-2n \cdot \epsilon^2}$), we have

1.
$$\forall \hat{y} \in \mathcal{F}, \quad \mathcal{R}_{pop}[\hat{y}] - \mathcal{R}_{emp}^{(n)}[\hat{y}] \le \sqrt{\frac{\log(N) + \log(1/\delta)}{2n}}$$

2. Equivalently,
$$\sup_{y \in F} \left\{ [\mathcal{R}_{pop}[\hat{y}] - \mathcal{R}_{emp}^{(n)}[\hat{y}] \right\} \le \sqrt{\frac{\log(N) + \log(1/\delta)}{2n}}$$



Proof:

$$\mathbb{P}\Big[\forall k \in \{1,..,N\}, \quad \mathbb{E}_{\mathbf{x},y}[\mathrm{loss}(\hat{y}_k(\mathbf{x}),y)] - \frac{1}{n} \sum_{i=1}^n \mathrm{loss}(\hat{y}_k(\mathbf{x}_i),y_i) \leq \epsilon\Big] = 1 - \mathbb{P}\Big[\exists k \in \{1,..,N\}, \quad \mathcal{R}_{\mathrm{pop}}[\hat{y}_k] - \mathcal{R}_{\mathrm{emp}}^{(n)}[\hat{y}_k] \geq \epsilon\Big]$$

What's more, we have that

$$\mathbb{P}\Big[\exists k \in \{1,..,N\}, \mathcal{R}_{\text{pop}}[\hat{y}_k] - \mathcal{R}_{\text{emp}}^{(n)}[\hat{y}_k]\Big] \leq \sum_{j=1}^{N} \mathbb{P}[\mathcal{R}_{\text{pop}}[\hat{y}_j] - \mathcal{R}_{\text{emp}}^{(n)}[\hat{y}_j] \geq \epsilon] \leq N \cdot e^{-2n \cdot \epsilon^2}$$

R The inequality on the left is obtained from Boole's inequality. As for the one on the right, it is given by Hoeffding's inequality

Which implies,

$$\mathbb{P}\Big[\forall k \in \{1, .., N\}, \quad \mathbb{E}_{\mathbf{x}, y}[\operatorname{loss}(\hat{y}_k(\mathbf{x}), y)] - \frac{1}{n} \sum_{i=1}^n \operatorname{loss}(\hat{y}_k(\mathbf{x}_i), y_i) \le \epsilon\Big] \ge 1 - N \cdot e^{-2n\epsilon^2}$$

8.4 What to do if N is infinite

Definition 8.5 (Growth Function) Let $\hat{y} \in \mathcal{F}$. We introduce $\Pi_{\mathcal{F}}[\vec{x_1},...,\vec{x_m}]$ as the number of possible classifications.

$$\Pi_{\mathcal{F}}\left[\vec{x_1},...,\vec{x_m}\right] \leq 2^m$$

We define the Growth Function:

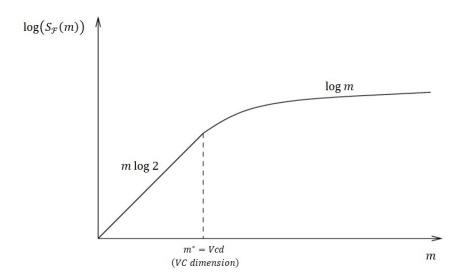
$$S_{\mathcal{F}}(m) = MAX \ all \ positions \ of \ m \ points \left[\Pi_{\mathcal{F}}[\vec{x_1},...,\vec{x_m}]\right] \le 2^m$$

R The Growth Function is a measure of complexity of the function class (a bigger Growth Function value translates into the function class being able to fit more complicated labels).

Using the works of Vapnik–Chervonenkis (VC), we can prove that if $\hat{y} \in \mathcal{F}$ with probability of at least $1 - \delta$, then:

$$\sup_{\hat{y}} \left\{ [\mathcal{R}_{\text{pop}}[\hat{y}] - \mathcal{R}_{\text{emp}}^{(n)}[\hat{y}] \right\} \le 2\sqrt{\frac{\log(S_{\mathcal{F}}(m)) + \frac{4}{\delta}}{n}}$$

Looking further into the (log of the) Growth Function, we can aslo notice that its typical behavior goes as follows:



This shows us that for any function class, there exists a point m^* beyond which we wouldn't be able to fit a possible separation.

Theorem 8.6 (Vapnik-Chervonen(VC), Sauer-Shelah, Sauer's lemma)

$$S_{\mathcal{F}}(m) < \left[\frac{em}{Vcd}\right]^{Vcd}$$

With $Vcd = \max_{m} such that S_{\mathcal{F}}(m) = 2^{m}$

(More intuitively, Vcd is the maximum number of points that the function class can fit perfectly, regardless of position or label.)



If \mathcal{F} is the set of linear function in dimension d, then $Vcd(\mathcal{F}) = d+1$

Definition 8.7 (VC Bound) Let $\hat{y} \in \mathcal{F}$, then with probability of at least $1 - \delta$:

$$\sup_{\hat{y}} \left\{ [\mathcal{R}_{pop}[\hat{y}] - \mathcal{R}_{emp}^{(n)}[\hat{y}] \right\} \leq 2\sqrt{\frac{2Vcd\log(\frac{2ne}{Vcd}) + \log(\frac{4}{\delta})}{n}} \leq \sqrt{\frac{Vcd\log(n)}{n}k}$$

Therefore, if n >> Vcd then $\mathcal{R}_{pop}[\hat{y}] - \mathcal{R}_{emp}^{(n)}[\hat{y}]$ is very small, allowing us to focus on optimising the empirical risk instead of the population risk.

8.5 More on Uniform Convergence

Let $f \in \mathcal{F}$ (a class of function) and $VCd(\mathcal{F}) = VC$ dimension of the function class. Then, from the uniform convergence theorem, we have

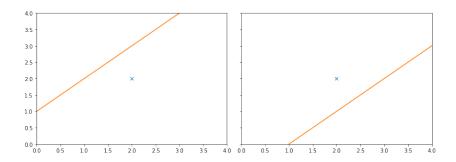
$$\sup_{f \in F} |\mathcal{R}_{\text{pop}}[\hat{y}] - \mathcal{R}_{\text{emp}}^{(n)}[\hat{y}]| \lesssim \sqrt{\frac{\text{VCd}(\mathcal{F})\log n}{n}}$$

The question now is, what is $VCd(\mathcal{F})$?

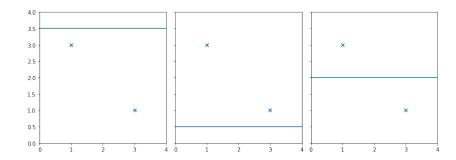
Definition 8.8 (Vapnik–Chervonenkis dimension) $VCd(\mathcal{F})$ is the largest number of points, such that it exists a set of points that can be fitted, no matter the assignment $f \in \mathcal{F}$.

8.5.1 Example 1: Linear separation in dimension d=2

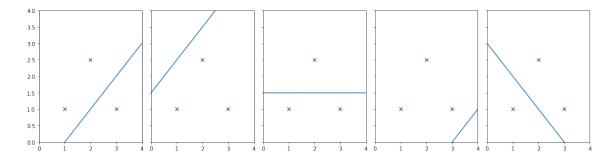
n = 1



n=2



n = 3



For n = 3, any set of unaligned points can be fitted, no matter the assignment $f \in \mathcal{F}$. On the contrary, for n = 4 there isn't a set of points that can be fitted no matter the assignment.

Therefore, VCd[linear - class, d = 2] = 3.

Claim 8.9 VCd[linear - class, d] = d + 1

Conclusion: For linear classifiers, we have $\sup_{f \in F} |\mathcal{R}_{pop}[\hat{y}] - \mathcal{R}_{emp}^{(n)}[\hat{y}]| \lesssim \sqrt{\frac{d+1}{n}}$