Heterogeneous Dynasties and Long-Run Mobility

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Overview

1 Introduction

2 Models

3 Calibrations and Results

1 Introduction

1.1 motivational empirical fact 1

 Positive correlation between grandparent-child wealth-rank (Boserup 2014)

FYI

- no effect of direct contacts between grandparents and grandchildren identified (Braun and Stuhler 2018)
- reasonable to assume the grandparents effects (of social status) operating through indirect mechanisms (through parents).

1.1 motivational empirical fact 2

 Positive correlation between dynastic wealth-ranks across almost 600 years (Barone and Mocetti 2016; I doubt this one)!

FYI

- Literature: the transmission of wealth, education or occupational status through an underlying and unobserved latent factor.
- reasonable to assume the grandparents effects (of social status) operating through indirect mechanisms (through parents).

1.2 motivational theoretical fact

- current optimal consumption-saving model with heterogeneous agents can identify the main drivers of wealth inequality (earnings and wealth return) and fit well with (P-C) inter-generational social mobility of wealth (Benhabib 2011 2019)
- but cannot well fit long-run (GP-GC) intergenerational mobility in two aspects:
 - a large enough coefficient for GP-C wealth-rank
 - a large enough correlation for dynastic wealth-ranks over very long time periods.

1.3 economics Question and possible directions

Questions:

 how we should extend our current optimal-saving model with HA to capture features of the long-run intergenerational mobility?

Directions:

- 1: assume high intergenerational autocorrelation in earnings and/or wealth return;
- 2: assume persistent heterogeneity in wealth return rate across generations
 - It allows households in some dynasties to have wealth grow faster on average than households in other dynasties.

2 Models

2.1 models 1: rank-based model

Setups:

- an economy populated by N households, indexed by $i=1,\cdots,N$
- rank households by their wealth
 - $\rho_t(i)$: wealth-rank of household i at time t
 - define ranked wealth processes $w_{(1)} \geqslant \cdots \geqslant w_{(N)}$ by $w_{(\rho_t(i))}(t) = w_i(t)$,
 - $\rho_t(i) < \rho_t(j)$ iff $w_i(t) > w_j(t)$ or $w_i(t) = w_j(t)$ and i < j,
 - $w(t) = w_1(t) + \cdots + w_N(t)$: aggregate wealth of economy

2.1 models 1: rank-based model

For households $i = 1, \dots, N$, wealth dynamics are given by

$$d\log w_i(t) = \alpha_{\rho_t(i)}dt + \sigma_{\rho_t(i)}dB_i(t)$$
 (2.1)

where

- B_i : Brownian motion
- $\alpha_1 + \cdots + \alpha_N$: normalise average growth rate of economy to 0.

2.1 models 1: rank-based model

Banner et al. (2005) prove that eq 2.1 admits a stationary distribution iff α_k satisfy eq 2.2

$$\alpha_1 + \dots + \alpha_k < 0, \text{ for } k < N \tag{2.2}$$

Proposition 1

consider a rank-based model (eq. 2.1) that satisfies (eq. 2.2) and also eq. 2.3, for all $k=2,\cdots,N-1$,

$$\sigma_{k+1}^2 - \sigma_k^2 = \sigma_k^2 - \sigma_{k-1}^2 \tag{2.3}$$

the ranked wealth processes satisfy eq. 2.4, for all $k = 1, \dots, N-1$.

$$\mathbb{E}[\log w_{(k)}(t) - \log w_{(k+1)}(t)] = \frac{\sigma_k^2 + \sigma_{k+1}^2}{-4(\alpha_1 + \dots + \alpha_k)}$$
(2.4)

2.2 models 2: simple HA consumption model (Benhabib 2019 & 2011)

- each household has CRRA preferences and a joy-of-giving bequest motive
- in equilibrium, intergenerational wealth dynamics for each household is

$$w_i(t+1) = \lambda(r_{i,t})w_i(t) + \beta(r_{i,t}, y_{i,t})$$
 (2.5)

- $y_{i,t}, r_{i,t}$: labor income and wealth return for household i in generation t,
- $w_i(t)$: wealth holdings of household i in generation t.
- functions λ and β : average wealth return and average labor income, adjusted for equilibrium household behavior.

2.3 approximate model 1 using model 2

- assume function $\pi_t(k)$ identify index i of the k-th ranked household at time t s.t. $\pi_t(k) = i$ iff $\rho_t(i) = k$.
- For each rank $k=1,\cdots,N$, rank-based approximation of eq. 2.5 is eq. 2.1 with new defined α_k and σ_k^2 eq. 2.6

$$\alpha_k = \mathbb{E}[\log(w_{\pi_t(k)}(t+1)/w(t+1)) - \log(w_{\pi_t(k)}(t)/w(t)))] \tag{2.6-a}$$

$$\sigma_k^2 = Var[\log(w_{\pi_t(k)}(t+1)/w(t+1)) - \log(w_{\pi_t(k)}(t)/w(t)))]$$
(2.6-b)

• α_k and σ_k : average and and variance of growth rate of wealth relative to the aggregate at each rank k of the distribution.

2.3 approximate model 1 using model 2

At stationary distribution of standard model, the rank-based relative growth rate parameters α_k satisfy, for each rank $k=1,\cdots,N$,

$$\alpha_k = \mathbb{E}[\log(w_{\pi_t(k)}(t+1)/w_{\pi_t(k)}(t))]$$

= E [
$$\lambda(r_{\pi_t(k),t}) + \beta(r_{\pi_t(k),t}, y_{\pi_t(k),t}) / w_{\pi_t(k)}(t)]$$
(2.7)

• expected value of aggregate wealth w satisfies $\mathbb{E}[\log w(t+1)] = \mathbb{E}[\log w(t)]$ by stationarity.

2.3 approximate model 1 using model 2

from eq. 2.7, we have

$$\alpha_k = \mathbb{E}[\log(\lambda(r_{\pi_t(k),t})(1 + \frac{\beta(r_{\pi_t(k),t}, y_{\pi_t(k),t})}{\lambda(r_{\pi_t(k),t})w_{\pi_t(k)}(t)}))]$$

$$= \mathbb{E}[\log(\lambda(r_{\pi_{t}(k),t}))] + \mathbb{E}[\log(1 + \frac{\beta(r_{\pi_{t}(k),t},y_{\pi_{t}(k),t})}{\lambda(r_{\pi_{t}(k),t})w_{\pi_{t}(k)}(t)})] (2.8)$$

• This decomposition enables us to map the stability condition for rank-based models eq. 2.2 into a condition in terms of β and λ in standard model eq. 2.5.

Proposition 2 (Proof-provided)

if standard model eq. 2.5 is stationary, then its rank-based approximation defined by eq. 2.1 and eq. 2.6 is also stationary.

2.3 model 3: persistently heterogeneous rank-based model

For each household, wealth dynamics are given by eq. 2.9

$$d\log w_i(t) = (\gamma_i + \hat{\alpha}_{\rho_t(i)})dt + \sigma_{\rho_t(i)}dB_i(t)$$
 (2.9)

- $\gamma_i \in \{\gamma_l, \gamma_h\}$ with $\gamma_h > \gamma_l$,
 - assume n of households are high types with $\gamma_i = \gamma_h$,
 - N-n households are low types with $\gamma_i=\gamma_l$,
- normalizing average growth rate of wealth to 0 requires $\sum_{k=1}^{N} \hat{\alpha}_k + \sum_{i=1}^{N} \gamma_i = \sum_{k=1}^{N} \hat{\alpha}_k + (N-n)\gamma_l + n\gamma_h = 0$

2.3 model 3: persistently heterogeneous rank-based model

Ichiba et al. (2011) show that this model admits a stationary distribution iff it satisfy a stability condition that generalizes the condition eq. 2.2 for standard rank-based model eq. 2.1 with no heterogeneity:

$$\sum_{k=1}^{m} \hat{\alpha}_k + \tilde{m}\gamma_h + (m - \tilde{m})\gamma_l < 0$$
 (2.10)

for all $m = 1, \dots, N-1$; $\tilde{m} = \min(m, n)$.

2.4 theoretical characterization of asymptotic wealth-rank model eq. 2.1 for mobility

define occupation times $\xi_{i,k}$ for all i,k, as the fraction of time household i occupies rank k, $\xi_{i,k} = \lim_{T \to \infty} \frac{1}{T} \int_0^T \mathbb{1} \{ \rho_t(i) = k \} dt$.

• by definition, the occupation times must add up to 1, s.t. $\sum_{i=1}^{N} \xi_{i,k} = \sum_{k=1}^{N} \xi_{i,k} = 1$.

Proposition 2.3

occupation times $\xi_{i,k}$ in standard rank-based model eq. 2.1 satisfy eq. 2.11

$$\xi_{i,k} = \frac{1}{N}, a.s., for all i, k \tag{2.11}$$

2.4 theoretical characterization of asymptotic wealth-rank model eq. 2.1 for mobility

Proposition 2.3 (continue

furthermore, for each household i, the asymptotic wealth-rank satisfies eq. 2.12

$$\lim_{\gamma \to \infty} \mathbb{E}[\rho_{t+\tau}(i)] = \frac{N+1}{2}$$
 (2.12)

2.5 theoretical characterization of asymptotic wealth-rank model eq. 2.9 for mobility

Define occupantion times for low-type and high-type households.

- define low-type household occupation times $\xi_{l,k}s.t.\xi_{l,k} = \xi_{i,k}$, for all ranks $k = 1, \dots, N$;
- define the high-type occupation times $\xi_{h,k}$ s.t. $\xi_{h,k} = \xi_{j,k}$, for all ranks $k = 1, \dots, N$.
- the low- and high-type occupation times $\xi_{l,k}$ and $\xi_{h,k}$ must satisfy eq. 2.13

$$(N-n)\xi_{l,k} + n\xi_{h,k} = 1 (2.13)$$

for all $k = 1, \dots, N$.

2.5 theoretical characterization of asymptotic wealth-rank model eq. 2.9 for mobility

Proposition 2.4 (Proof Provided)

consider a persistently heterogeneous rank-based model eq. 2.9 that satisfies eq. 2.3 and eq. 2.10,

then the low- and high-type occupation times $\xi_{l,k}$ and $\xi_{h,k}$ satisfy eq. 2.14 and eq. 2.15

$$0 < \xi_{l,1} < \xi_{l,2} < \dots < \xi_{l,N} < \frac{1}{N-n}, a.s.,$$
 (2.14)

and

$$\frac{1}{n} > \xi_{h,1} > \xi_{h,2} > \dots > \xi_{h,N} > 0, a.s.$$
 (2.15)

• It implies that $\xi_{l,1} < \xi_{h,1}$ and $\xi_{h,N} < \xi_{l,N}$ since $\xi_{i,1} + \cdots + \xi_{i,N} = 1$.

2.5 theoretical characterization of asymptotic wealth-rank model eq. 2.9 for mobility

Theorem 2.5

consider a persistently heterogeneous rank-based model eq. 2.9 that satisfies eq. 2.3 and eq. 2.10 then

$$\lim_{\gamma \to \infty} \mathbb{E}[\rho_{t+\tau}(i)] < \lim_{\tau \to \infty} \mathbb{E}[\rho_{t+\tau}(j)] \text{ iff } \rho_t(i) < \rho_t(j)$$
 (2.16)

for all households $i, j = 1, \dots, N$.

 expectations are taken w.r.t. the stationary distribution unconditional on the types of households i and j.

3 Calibrations and results

3.1 calibration 1: approximated rank-based model

aim

• construct eq. 2.1 using 2.6 to define rank-based parameters α_k and σ_k

Procedures:

- **step 1**: parametrize eq. 2.1 (follow B 2011 2019, find the interpretations of paras there)
- step 2: simulate this parameterization of standard model with 2000 generations and 10000 households.
- step 3:use simulated results and Fernholz (2017)'s econometric procedure to estimate the relative growth-rate parameters α_k

3.1 calibration 1: approximated rank-based model

For step 1,

- set household lifespan T equal to 45 years
- growth rate of labor earnings equal to 0.01
- preference parameters η , ψ , χ are set to 0.04, 2, and 0.25
- estate tax and capital income tax, b and ζ are 0.2 and 0.15
- model lifetime labor income $y_{i,t}$ by using a six-state Markov chain calibrated to US SCF with mean and variance of 6.4 and 16.1, and 1 unit is 10000 dollars
- represent the idiosyncratic lifetime return on wealth $r_{i,t}$ by a 4-state Markov chain with mean and variance matching Fagereng et al. (2020) for Norwegian

3.2 results for calibration 1

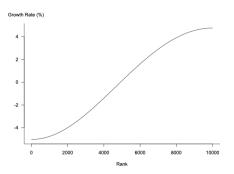


Figure 1: Annualized estimated parameters α_k for the approximated rank-based model.

3.2 results for calibration 1

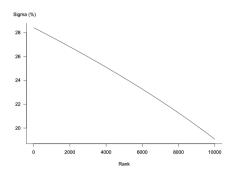


Figure 2: Annualized estimated parameters σ_k for the approximated rank-based model.

3.2 results for calibration 1

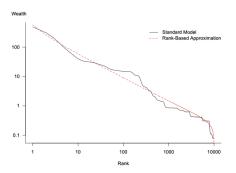


Figure 3: Log-log plot of wealth versus rank for the Standard model (average from 2,000 simulations) and its rank-based approximation.

aim:

 calibrate eq. 2.9 s.t. it maintains approximately the same realistic stationary wealth distribution as approximated rank-based model

assume

- 3000 of households are high-type, with $\gamma_h = 0.02$
- remaining 7000 low-type have $\gamma_1 \approx -0.0086$.
- they use the same estimated σ_k from calibration 1 here.

calibration for $\hat{\alpha}_k$ should be more delicate

• we cannot use α_k from calibration 1 here because the persistently heterogeneous parameters γ_i from eq. 2.9 lead to a more skewed stationary distribution than in the model eq. 2.1.

Methodology:

• adjust the estimated rank-based relative growth rate parameters α_k in fig1 s.t. $\hat{\alpha}_k \neq \alpha_k$, to maintain a similar stationary distribution for the two rank-based models.

consider the rank-based approximation eq. 2.1 of heterogeneous rank-based model eq. 2.9, where α_k are defined in eq. 2.6.

relative growth rate para α_k^\prime for the rank-based approximation are given by eq. 3.1

$$\alpha_k' = \hat{\alpha}_k + (N - n)\xi_{l,k}\gamma_l + n\xi_{h,k}\gamma_h \tag{3.1}$$

for all $k = 1, \dots, N$.

- we cannot directly solve for $\hat{\alpha}_k$ since we cannot directly solve for occupation time $\xi_{l,k}$ and $\xi_{h,k}$ in eq. 3.1.
- we can simulate the persistently heterogeneous rank-based model to generate estimates of these parameters.

Procedure of the simulation: generate estimates of the paras $\hat{\alpha}_k$ from the model eq. 2.9 s.t. α_k' is approximately equal α_k , for each rank k.

- step 1: use eq. 3.1 to guess values of paras $\hat{\alpha}_k$ s.t. $\alpha_k' \alpha_k \approx 0$, for all $k = 1, \cdots, N$,
- Step 2: simulate the persistently heterogeneous rank-based model with these parameters $\hat{\alpha}_k$ to
 - ullet generate estimates of the rank-based approximation paras $lpha_k'$
 - calculate the standard deviation of $\alpha_k' \alpha_k$ (error term),
- step 3: once the error term is calculated, we incrementally alter the values of $\hat{\alpha}_k$ by setting each equal to $x\hat{\alpha}_k$

Procedures (continue)

- step 4: re-estimate the parameters α_k' and again calculate the sum of squared values $\alpha_k' \alpha_k$.
- step 5: evaluate whether squared error with parameter values $x\hat{\alpha}_k$ is smaller.
 - if so, then keep the new paras and repeat the **steps 1-4** by altering the new paras in the same way;
 - if not, then consider a different value of x in **step 3** and repeat the **step 4-5**.
- step 6: this procedure repeats until the sum of squared values $\alpha_k' \alpha_k$ is larger for the paras $x\hat{\alpha}_k$, for both x = 1.001 and x = 0.999.

3.4 results for calibration 3

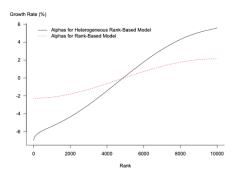


Figure 4: Annualized estimated parameters $\hat{\alpha}_k$ for the persistently heterogeneous rank-based model, and annualized estimated parameters α_k for the approximated rank-based model.

3.5 model 4 and calibration 3: auto-correlated returns model

- introduce autocorrelated returns into the standard model eq.
 2.5 to capture imperfect social mobility as in B (2011 2019)
- investigate how such returns impact long-run mobility

wealth returns follow a highly persistent AR-1 process with eq. 3.2

$$\log(1 + r_{i,t+1}) = \theta \log(1 + r_{i,t}) + \epsilon_{i,t}$$
 (3.2)

- $\epsilon_{i,t}$: normally distributed with mean and standard deviation equal to 0.0375 and 0.025;
- persistence parameter θ set to 0.9.
- labor earnings $\log y_{i,t}$ IID and drawn from a normal distribution with mean equal to 0.85 and sd equal to 1.416.

3.6 results table 1

	Data	Standard	Approximated	Persist. Heter.	Auto-Correlated
		Model	Rank-Based	Rank-Based	Returns
			Model	Model	Model
Wealth Distribution					
Top 1%	33.6%	33.1%	32.6%	35.1%	98.5%
Top 1-5%	26.7%	26.4%	17.1%	16.6%	0.6%
Top 5-10%	11.1%	6.3%	9.4%	9.1%	0.3%
Top 10-20%	12.0%	8.0%	11.0%	10.6%	0.3%
Top 20-40%	11.2%	10.5%	12.9%	12.4%	0.2%
Top 40-60%	4.5%	7.7%	8.2%	7.9%	0.1%
Bottom 60-100%	-0.1%	8.0%	8.7%	8.3%	0.1%
Wealth-Rank Correlations					
Parent-Child Rank Coeff.	0.191	0.177	0.187	0.218	0.252
Grandparent-Child Rank Coeff.	0.116	0.015	-0.004	0.089	0.147
Long-Run Persistence Coeff.	0.105	0.001	0.000	0.116	0.022

Table 1: Upper part: Average wealth shares from 1,000 simulations of the different models - data from the Survey of Consumer Finances. Lower part: Average coefficients from regressions of child rank on parent rank and grandparent rank from 1,000 simulations of the different model - data from Danish wealth holdings for three generations in Boserup et al. (2014). Average coefficient from regressions of household rank in generation t on household rank in generation t-23 (585 years) from 1,000 simulations of the different models - data from estimates of very long-run (585 years) dynastic wealth holdings in Florence, Italy, in Barone and Mocetti (2016).

3.6 results table 2

	Top 1%	Top 5%	Bottom 50%	Bottom 25%
High-Type Households	86.6%	72.7%	17.5%	12.0%
Low-Type Households	13.4%	27.3%	82.5%	88.0%

Table 2: Average composition of the top 1%, top 5%, bottom 50%, and bottom 25% of households from 1,000 simulations of the heterogeneous rank-based model.

3.6 results figure 5 & 6

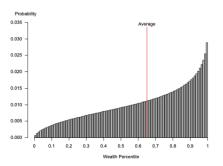


Figure 5: Average high-type household occupation times for different percentiles of the wealth distribution from 1,000 simulations of the heterogeneous rank-based model.

3.6 results figure 5 & 6

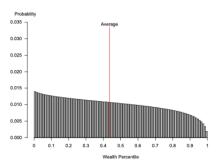


Figure 6: Average low-type household occupation times for different percentiles of the wealth distribution from 1,000 simulations of the heterogeneous rank-based model.

3.6 results figure 7

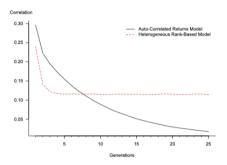


Figure 7: Rank correlations across multiple generations from 1,000 simulations of the heterogeneous rank-based and auto-correlated returns models.