

# REDSHIFT EVOLUTION OF HIGHER-ORDER CORRELATIONS IN COSMIC LARGE-SCALE STRUCTURE WITH MINKOWSKI FUNCTIONALS

JAMES M. SULLIVAN,<sup>1,2</sup> ALEXANDER WIEGAND,<sup>1</sup> AND DANIEL J. EISENSTEIN<sup>1</sup>

<sup>1</sup>*Harvard-Smithsonian Center for Astrophysics, 60 Garden St, Cambridge, MA, 02138, USA*

<sup>2</sup>*University of Texas at Austin, 110 Inner Campus Dr, Austin, TX 78705, USA*

## ABSTRACT

We probe the higher-order galaxy clustering in the final data release (DR12) of the Sloan Digital Sky Survey Baryon Oscillation Spectroscopic Survey (BOSS) using germ-grain Minkowski Functionals (MFs). Our data selection contains 979,430 BOSS galaxies from both the northern and southern galactic caps over the redshift range  $0.2 - 0.6$ . We extract the higher-order part of the MFs and find deviations from the case without higher order MFs with  $\chi^2$  values of  $\mathcal{O}(10^3)$  for 24 degrees of freedom across the entire data selection. We show the MFs to be sensitive to contributions up to the five-point correlation function across the entire data selection. We measure significant redshift evolution in the higher-order functionals for the first time, with a percentage growth between redshift bins of  $20\% - 25\%$  in both galactic caps. This is a factor of 2 greater than similar growth in the two-point correlation term, and may have implications for the growth factor and the redshift evolution of the three-point and higher functions.

*Keywords:* methods: data analysis, methods: statistical, cosmology: observations, large-scale structure of Universe

## 1. INTRODUCTION

Cosmic large-scale structure is largely understood by characterizing and modeling observed galaxy distributions, which strongly influence the accepted cosmological model. Recent galaxy redshift surveys contain spectroscopic redshifts for of order  $10^5 - 10^6$  galaxies, and their size makes them ideal testing grounds for exploring non-Gaussian features in this structure. Any quantitative analysis of such features must be firmly rooted in precise statistical measures. The most standardized of these measures is the spatial two-point correlation function  $\xi_2$  (the probability of finding two galaxies within a certain distance), which is commonly used to constrain cosmological model parameters. However, observed structure is more complex than the Gaussian structure described by two-point statistics, and is in fact more intricate than the structure described by three-point statistics. Evidently, to accurately measure large-scale structure we must look to higher-order information in relation to these correlations.

The simplest approach is to calculate the correlation functions directly. While this has been achieved for the two-point and three-point functions (Slepian & Eisenstein (2015) and Slepian et al. (2017)), fourth- and higher-order functions of a similar level of accuracy have yet to be determined. In fact, computing these functions directly is computationally not feasible (Wiegand & Eisenstein (2017)). There are various alternatives, and one of the most useful and rigorous is Minkowski Functional (MF) analysis. MFs quantitatively describe the geometry of extended bodies by mapping the shape of a body to real numbers. These functionals uniquely characterize the geometry and topology of a galaxy distribution, and contain information about all higher-order correlation functions.

MFs were first used to characterize large-scale structure by Mecke et al. (1994) in the form of the germ-grain model. This model pins down the morphology of the galaxy distribution by treating the survey galaxies as points (the germs) and decorating them with balls (the grains) whose scale-probing radius is the only model parameter. The union of these balls creates a set of extended bodies to which methods from integral geometry can be applied (For a review see Schmalzing (1999) or Schmalzing et al. (1996), and see Buchert (1995) for a short review.). We discuss this model further in the context of our analysis.

Another popular use of MFs lies in applying them to the isodensity contours of density fields, including galaxy and cluster surveys (e.g. Schmalzing et al. (1996), Kerscher et al. (1997), Kerscher et al. (1998), Kerscher et al. (2001a), and Kerscher et al. (2001b)), dark matter overdensity fields (Platzöder & Buchert (1996), Schmalzing & Buchert (1997), Sahni et al. (1998), Sathyaprakash et al. (1998), Schmalzing et al. (1999), Hikage et al. (2003), Nakagami et al. (2004),

Choi et al. (2013), and Blake et al. (2014)), and other astrophysical settings (Petri et al. (2013), Petri et al. (2013), Gleser et al. (2006), Einasto et al. (2014), and Yoshiura et al. (2017)). Still further work with MFs has recently used the isothermality contour maps of the CMB to constrain its Gaussianity (Ducoat et al. (2013), Planck Collaboration et al. (2014a), and Planck Collaboration et al. (2014b), Planck Collaboration et al. (2016)). Clearly MFs are a powerful and widely-used tool in subfields across cosmological disciplines.

Here we focus on the germ-grain model, because recent work using the 7th and 12th data releases of SDSS-III (DR7 and DR12) of the Baryon Oscillation Spectroscopic Survey (BOSS) (Dawson et al. (2013)) has shown its use in accessing correlation information that cannot be calculated directly (DR7 paper - (Wiegand et al. (2014)) & DR12 paper - (Wiegand & Eisenstein (2017))), SDSS-III - Eisenstein et al. (2011)). We use the largest complete spectroscopic redshift survey to date, the complete DR12 dataset, to access unprecedented accuracy in the higher-order correlation functions. We also explore the redshift evolution of the higher-order correlations, which are expected to grow nonlinearly in time with varying behavior for different orders.

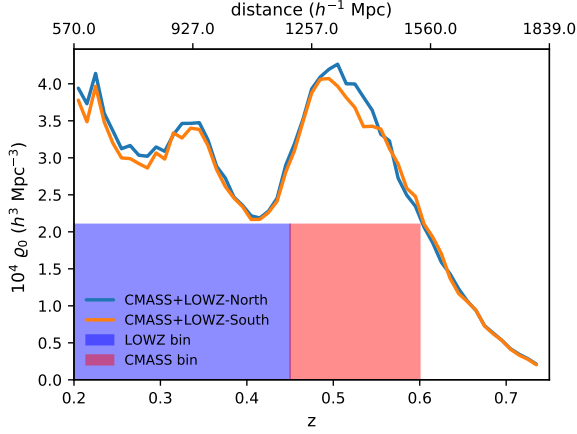
This paper is organized as follows. Section 2 describes our usage of the BOSS dataset. Section 3 and Section 4 describe our use of the germ-grain model and our transformations of the MFs. Section 5 describes higher-order correlations and relates our analysis of redshift evolution. Section 6 presents our conclusions.

## 2. BOSS DATA

Our sample was collected with the 2.5m Sloan Telescope (Gunn et al. (2006)). For specifics on photometry and instruments, refer to Gunn et al. (1998), Fukugita et al. (1996), Lupton et al. (2001), Smith et al. (2002), Pier et al. (2003), Padmanabhan et al. (2008), Doi et al. (2010), as well as the eighth data release (Eisenstein et al. (2011)). For details on spectroscopic redshift determination see Smee et al. (2013) and Bolton et al. (2012). Our samples are drawn from the third phase of the SDSS (York et al. (2000)) Luminous Red Galaxy (LRG) catalog (Eisenstein et al. (2001)) of the BOSS (Dawson et al. (2016)) in DR12 (Alam et al. (2015)). For specifics regarding the data, see Reid et al. (2016).

### 2.1. Redshift Samples - CMASS and LOWZ

By using the full DR12 dataset, we are able to access an unprecedented number of spectroscopic redshifts. We use the CMASS ("Constant (stellar) Mass") and LOWZ samples (referring to low redshifts, or Low- $z$ ) samples, which are defined in detail by Reid et al. (2016). The CMASS sample was designed to expand upon the color cuts towards the blue in the SDSS-I and SDSS-II LRG samples, and has a redshift

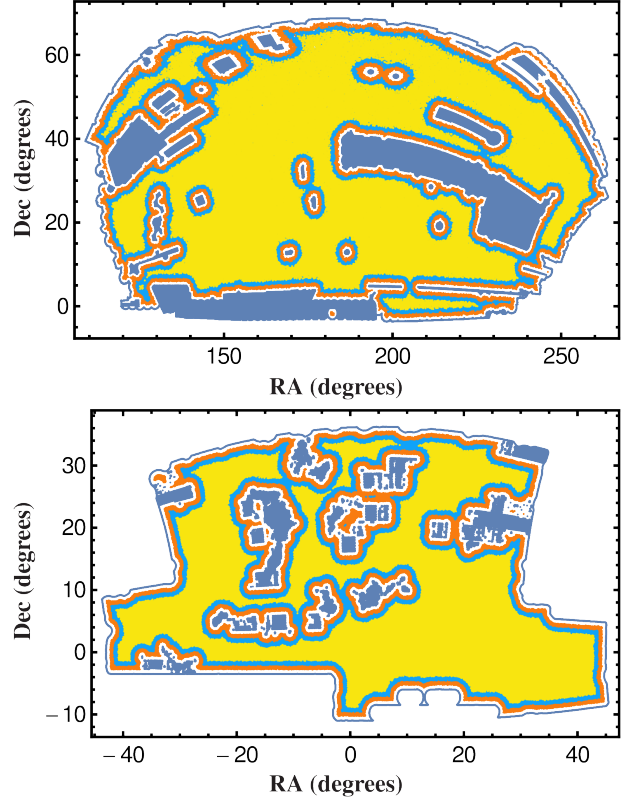


**Figure 1.** Minimum number density for the combined full CMASS, LOWZ, and LOWZE2 catalogs. By considering only values of  $\rho_0$  that are at or below the minimum value over the combined sample, we need not restrict our analysis to differing densities at differing redshifts. The bins we call CMASS and LOWZ in this work are shown here.

range of  $z \in [0.450, 0.750]$ . The LOWZ sample was designed to decrease the lower redshift bound of SDSS-I and SDSS-II down to  $z = 0.2$ , increasing the effective survey area by roughly a factor of 3. We consider two redshift bins (Fig. 1), which we will use CMASS ( $0.45 \leq z \leq 0.60$ ) and LOWZ ( $0.20 \leq z \leq 0.45$ ) to refer to from now on. The bins we use give a redshift range near triple that used in the DR12 paper. We use both the northern and the southern galactic caps (N- and SGC) of CMASS and LOWZ for an effective area of  $9,376 \text{ deg}^2$ . We also make use of the LOWZE2 sample (chunk 2 of the original LOWZ footprint), though it contributes a relatively small amount to the effective area and sample count ( $131 \text{ deg}^2$ , 2,985). There are a total of 979,430 galaxies in our two-bin sample, more than double the 410,615 used in the DR12 paper. The CMASS-South sample is approximately a factor of 2 smaller than its northern counterpart, and provides a much sparser coverage of the sky (Table 1). This has some effect on the behavior of the MFs, especially at large scales. LOWZ offers comparatively fewer observed galaxies, but has a better balance between the north and the south than CMASS does.

### 2.2. Survey and Corrections

We used a similar masking process to that of the DR12 paper. We considered only areas of the survey that were more than a distance of  $2R$  from any inner or outer boundary of the data, where  $R$  is the common ball radius around each galaxy. This restricted our use of the data, more so for large values of  $R$ , but for MFs boundary corrections using random distributions are not effective. In correcting for bad and missing data, we used veto masks to remove galaxies based on effects such as poor spectral seeing, extinction, areas without observations and poor photometric conditions. The masks for both

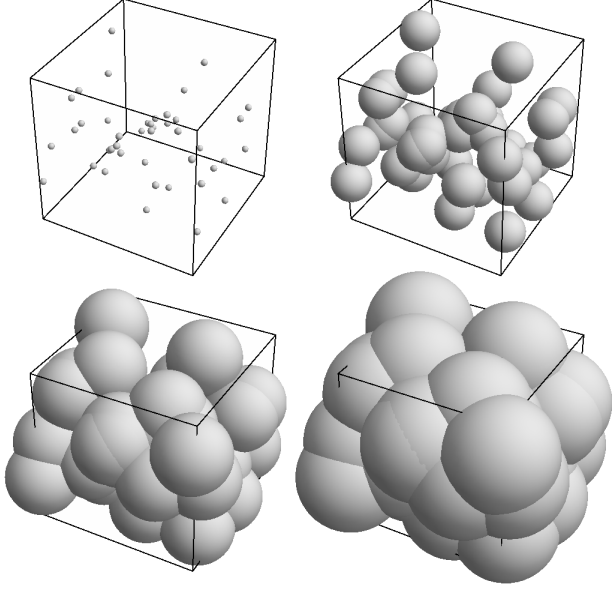


**Figure 2.** External and internal boundaries for the northern and southern patches analyzed. In blue the regions that we chose to define our boundary. In orange, light blue and yellow a projection of the regions 18, 36 and  $54 h^{-1} \text{ Mpc}$  away from the nearest boundary. The boundary for the northern high- $z$  bin is the same as in [Wiegand et al. \(2014\)](#).

caps are provided in Fig. 2. The LOWZ sample especially has a comparatively large number of holes and covers a smaller area of the sky. We see the effects of this on our analysis, especially at low densities.

### 2.3. MD Patchy Mocks

We use mock data realizations to compare the correlations in the BOSS data to those of the cosmological concordance model. The mocks used in our analysis are from the MultiDark(MD)-Patchy mock galaxy catalogs ([Kitauro et al. \(2016\)](#)). These mocks were specifically produced to reflect the number density of the data, and were produced by referencing a large N-body simulation (BigMultiDark). We used the same  $\Lambda \text{CDM}$  cosmology as the simulation, which is the Planck1 cosmology, with  $\Omega_M = 0.307115$ ,  $\Omega_\Lambda = 0.692885$ ,  $\Omega_b = 0.048$ ,  $\sigma_8 = 0.8288$ , and  $h = 0.6777$ . As indicated in the DR12, these mocks are very well suited for comparison when calculating the higher-order parts of the MFs, although there is light tension between the data and the MD-Patchy power spectrum that may be relieved by a 5% reduction in the mock power spectrum amplitude. We used 997 mock survey re-



**Figure 3.** Illustration of the germ-grain procedure to transform a set of galaxies into an extended body. The galaxy positions are surrounded by balls of a common radius  $R$ . Then the MFs of the body formed by the union of all the balls are studied as a function of  $R$ . From Wiegand et al. (2014).

alizations for CMASS-North (the same as for the DR12 paper), 399 mocks for CMASS-South and LOWZ-North, and 383 mocks for LOWZ-South in our calculations using the CHIPMINK code. These large numbers of mock realizations allow us to derive correlation-conscious uncertainties when comparing the data to the concordance model.

**Table 1.** Reduced set of basic parameters of the SDSS DR12 CMASS and LOWZ samples.

Sample	CMASS			LOWZ		
	NGC	SGC	total	NGC	SGC	total
$N_{\text{gal}}$	607,357	228,990	836,347	177,336	132,191	309,527
Effective area (deg <sup>2</sup> )	6,851	2,525	9,376	5,836	2,501	8,337
$N_{\text{gal}}(\text{our bins})$	410,617	294,091	704,708	151,003	123,719	274,722

### 3. THE GERM-GRAIN MODEL

In this Section, we provide an abbreviated description of Minkowski Functionals. For a more complete explanation of the germ-grain model, see the DR7 paper.

#### 3.1. Minkowski Functionals

We apply a result from integral geometry (Hadwiger 1957) to express the MFs as a linear combination of 4 base functionals in three-dimensional Euclidean space. We choose a

normalization of the base functionals (denoted  $V_0 - V_3$ ) that correspond to familiar geometric measures as follows:

$$V_0 = V ; V_1 = \frac{S}{6} ; V_2 = \frac{H}{3\pi} ; V_3 = \chi . \quad (1)$$

Here  $V$  is the volume of the extended body,  $S$  is its surface area,  $H$  is the integral mean curvature of the surface, and  $\chi$  is the topological Euler characteristic or integral Gaussian curvature. To refresh the reader,  $H$  is the average of the curvature of a surface over all angles and  $\chi$  is the sum of independent components and cavities minus the number of holes.

The germ-grain model connects MFs and observed galaxy distributions. The germ-grain procedure is to take a point distribution of "germs" and surround each point with a convex body, or "grain." By taking a galaxy distribution with positions  $\{\mathbf{x}_1, \dots, \mathbf{x}_N\}$  as the point distribution, and using balls (filled spheres) of radius  $R$  as the convex body, we create an extended body from the union of the spheres. The MFs can now be used to describe the relatively simple extended body that was transformed from the observed galaxy distribution (Fig. 3).

The MFs clearly depend on the common ball scaling radius  $R$ , but they also depend on the sample density  $\varrho_0$  of the observed distribution due to the germ-grain construction. The radius  $R$  is a tunable parameter that we alter in our analysis to probe different scales of structure. When exploring the density parameter space, we randomly sample from either the mock or data distribution at different values of  $\varrho_0$ . These values are determined by downsampling at fixed percentages from the height of our bins. At low densities, this tends to erase structure and the MF's approach Poissonity.

### 4. MEASURING MFS

In this Section, we describe the quantities we measure for each sample.

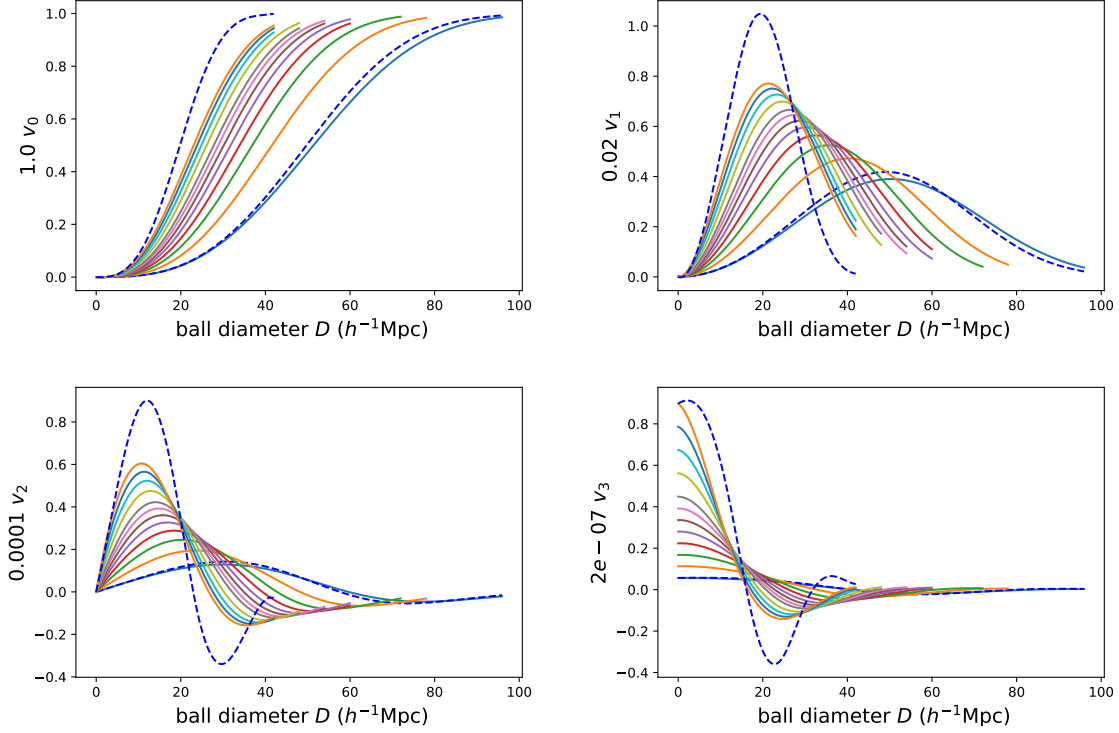
#### 4.1. Functional Densities

To measure the MFs for our samples, we must calculate the MFs of the individual balls. These *partial MFs* are then taken as an ensemble of statistical quantities. Rather than work with the extrinsic quantities  $V_\mu$ , we use the survey volume  $V_{\text{Survey}}$  to define the intrinsic functional densities  $v_\mu$ :

$$v_\mu \equiv V_\mu / V_{\text{Survey}} ; \mu \in \{0, \dots, 3\} \quad (2)$$

The analytic expressions for these  $v_\mu$  are given by:

$$\begin{aligned} \langle v_0 \rangle &= 1 - e^{-\varrho_0 \bar{V}_0} , \\ \langle v_1 \rangle &= \varrho_0 \bar{V}_1 e^{-\varrho_0 \bar{V}_0} , \\ \langle v_2 \rangle &= \left( \varrho_0 \bar{V}_2 - \frac{3\pi}{8} \varrho_0^2 \bar{V}_1^2 \right) e^{-\varrho_0 \bar{V}_0} , \\ \langle v_3 \rangle &= \left( \varrho_0 \bar{V}_3 - \frac{9}{2} \varrho_0^2 \bar{V}_1 \bar{V}_2 + \frac{9\pi}{16} \varrho_0^3 \bar{V}_1^3 \right) e^{-\varrho_0 \bar{V}_0} , \end{aligned} \quad (3)$$



**Figure 4.** MF densities  $v_\mu$  for the CMASS-North MD-Patchy mock files within a sample density range of 5% to 80% of the reference density  $\varrho_0 = 2.08 \times 10^{-4} h^3 \text{Mpc}^{-3}$ . The most extreme densities are plotted along with their theoretical Poisson point distributions (blue dotted lines). At the lowest density the MFs approach the Poisson curves, and at the highest density show the greatest difference from the Poisson curves.

where  $\langle v_\mu \rangle$  refer to the average functional densities over the distribution, and  $\bar{V}_\mu$  refer to the unnormalized modified MFs (Wiegand et al. (2014)). The  $v_\mu$  as measured for the mocks are shown in Fig. 4. Their behavior as a function of  $R$  (or diameter  $D$  in this case) and  $\varrho_0$  is intuitive: as  $R$  increases the volume of the survey box occupied by the balls is saturated as  $v_0$  approaches unity, and an decrease in density delays this saturation. The other functionals are slightly less transparent, but show that there is some maximum in the uncovered surface area when  $R$  is increased, and a similar maximum and minimum in curvature and Euler characteristic. For the surface area, a decrease in density will delay the maximum as a function of radius, and for the remaining two MF densities, the density affects the amplitude of the functional.

#### 4.2. Transforming the Densities

Observed galaxy distributions have significant structure, and thus must employ the full form of  $\bar{V}_\mu$  in Eq. (3). But for the case of a Poisson distribution, the unnormalized modified MFs  $\bar{V}_\mu$  are given for a ball  $B$  by the familiar quantities:

$$\begin{aligned} V_0(B) &= \frac{4\pi}{3} R^3 ; & V_1(B) &= \frac{2}{3} \pi R^2 ; \\ V_2(B) &= \frac{4}{3} R ; & V_3(B) &= 1 . \end{aligned} \quad (4)$$

These are intuitive and convenient to work with, and we use the Poisson case as a reference in analyzing distributions with more structure. Following from this choice, we adopt the dimensionless transformed MFs  $\eta_\mu$ :

$$\eta_\mu \equiv \bar{V}_\mu / V_\mu(B) . \quad (5)$$

The mapping from  $v_\mu$  to  $\eta_\mu$  retains the information about the MFs as a function of  $R$  and  $\varrho_0$ , but the latter provides a better perspective in terms of the higher-order correlations. To see this, we can rewrite the dimensionless MFs as the following power series:

$$\eta_\mu = \sum_{n=0}^{\infty} \frac{c_{\mu,n+1}}{(n+1)!} (-\varrho_0 V_0(B))^n \quad (6)$$

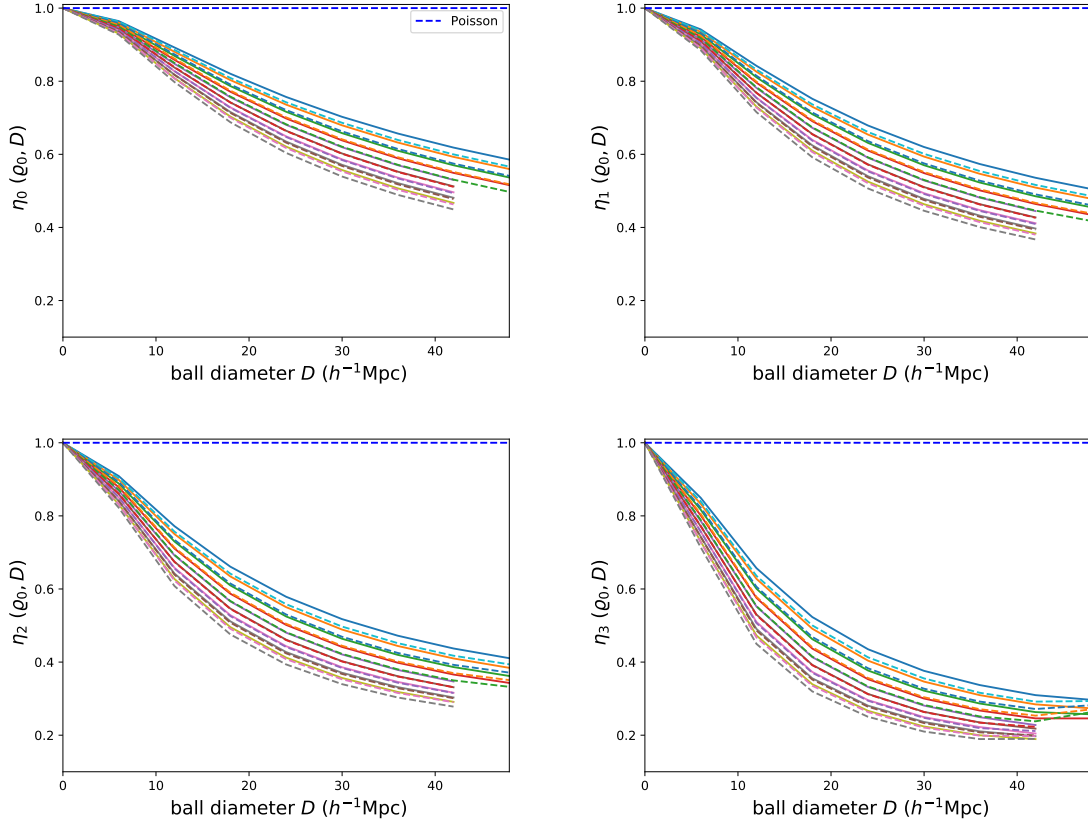
where  $c_{\mu,1} = 1$  and

$$c_{\mu,n+1}(R) = V_0^{-n}(B) \int \xi_{n+1}(0, \mathbf{x}_1, \dots, \mathbf{x}_n) \times \frac{V_\mu(B \cap B_{\mathbf{x}_1} \cap \dots \cap B_{\mathbf{x}_n})}{V_\mu(B)} d^3 x_1 \dots d^3 x_n \quad (7)$$

where  $V_\mu(B \cap B_{\mathbf{x}_1} \cap \dots \cap B_{\mathbf{x}_n})$  are the weighting functions).

Here, each term of the series is the integral of the product of the  $n$ -point correlation function  $\xi_{n+1}$  and a weighting function that depends on the balls (and therefore  $R$ ). There is also





**Figure 5.** Dimensionless transformed MFs in the Southern Galactic Cap for evaluated for 399 SDSS DR12 MD-Patchy mocks plotted as a function of ball diameter. CMASS-South values are solid, LOWZ-South values are dashed. These quantities can be more directly related to the correlation functions than the  $v_\mu$  can. The difference in the functionals between the redshift bins is visible across all the MFs.

a factor that includes the sample density  $\varrho$  dependence that may be considered independent of the correlation-dependent integral. Now we have a quantity conveniently defined on the unit interval  $[0, 1]$  (for standard cosmological structure) that can be more easily compared in terms of our parameters. Fig. 5 shows the nine highest density samples (40% - 80% of  $\varrho_0 = 2.08 \times 10^{-4} h^3 \text{ Mpc}^{-3}$ ) of  $\eta_\mu$  as a function of  $D$  and now gives a measure of the non-Poisson behavior of the MFs. The MFs deviate further from Poisson behavior as  $D$  and  $\varrho_0$  increase, approaching the structure of the full galaxy distribution, and illustrating the increased non-Poissonity at large scales. Visually the MFs are on the order of 1% different between the redshift bins, so already the redshift evolution in the MFs is evident. However, we have yet to isolate higher-order contributions of the MFs, which we discuss in the next section.

## 5. HIGHER-ORDER CONTRIBUTIONS TO THE MFS

This Section describes our analysis of the isolated higher-order contributions to the MFs. In the interest of conserving page space, we will often choose  $\eta_1$  as a proxy for the behavior of all the MFs, which is similar to that of  $\eta_1$ .

### 5.1. The Two-point Function

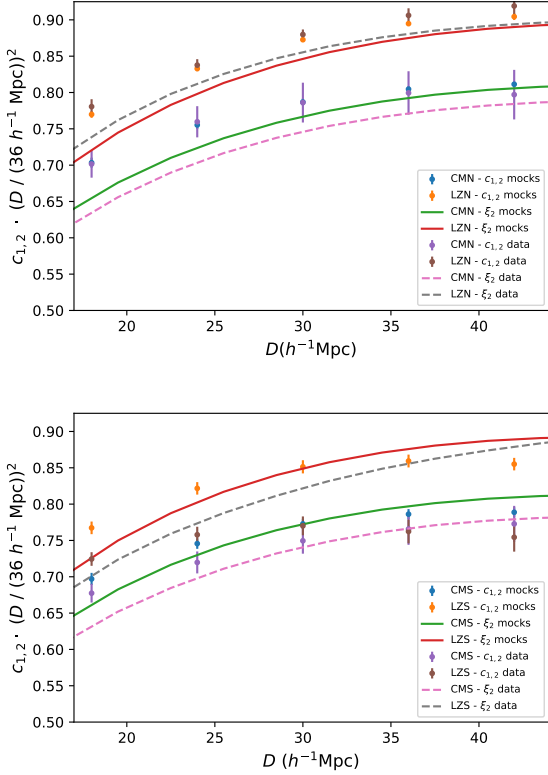
To measure higher-order correlations, we subtract the two-point correlation function contribution to the MFs. To do this, we must first calculate the two-point term. Ignoring the higher-order terms ( $n \geq 3$ ) of Eqn. (6) we are left with:

$$\eta_\mu = 1 - \frac{\varrho_0}{2} \int d^3x_1 \xi_2(0, \mathbf{x}_1) \frac{V_\mu(B \cap B_{\mathbf{x}_1})}{V_\mu(B)}. \quad (8)$$

We assume isotropic correlations, and upon transforming to spherical coordinates (8) becomes:

$$\eta_\mu = 1 - 2\pi\varrho_0 \int_0^{2R} \frac{V_\mu(R, r)}{V_\mu(B(R))} \xi_2(r) r^2 dr. \quad (9)$$

where the integral is zero if the centers of the balls are more than  $2R$  apart. This is the same as Eqn. (6) for  $n = 1$  barring the factor of  $V_0(B)$ .



**Figure 6.** Comparison between the MF coefficient  $c_{1,2}$  and the independently calculated integrated two-point correlation function, represented here by  $\xi_2$ . The true integrated quantity is given by Eqn. (9) where the second term is multiplied by a factor of  $\frac{\theta_0}{2}$  to directly compare with  $c_{1,2}$ . Both values are scaled by the square of the scale diameter to better show the deviations in the mocks and data.

We can write the weight functions  $V_\mu(B \cap B_1)$  as:

$$V_0(r) = \frac{1}{12} \pi (2R - r)^2 (r + 4R) ; \quad (10)$$

$$V_1(r) = \frac{1}{3} \pi R (2R - r) ; \quad (11)$$

$$V_2(r) = \frac{2}{3} (2R - r) + \frac{2}{3} R \sqrt{1 - \left(\frac{r}{2R}\right)^2} \arcsin\left(\frac{r}{2R}\right) ; \quad (12)$$

$$V_3(r) = 1 , \quad (13)$$

where  $R$  is again the radius of the balls and  $r$  is the integration variable. The weights then conveniently define integration windows for  $\xi_2$  for the different functionals, and as  $\mu$  increases, so does the  $R$  at which the structure is probed.

Using independently determined values of  $\xi_2$ , we compare the two-point coefficient of Eqn. (8) to the mathematically equivalent integrated two-point function Eqn. (9). The integrated two-point correlation functions match up fairly well to the  $c_{\mu,2}$  coefficients at all but the highest scales we probed (Fig. 6). It is reassuring to find this match in both redshift bins and in both galactic caps, because the MD-Patchy mocks were designed to match the two-point correlation function.

The agreement is not quite as good as in the DR12 paper, however this is mostly due to our use of a smaller covariance matrix. We employ a  $24 \times 24$  covariance matrix in our determination of the  $c_{\mu,2}$  (see Section 5.3 for details), while the matrix used in the DR12 paper is  $171 \times 171$ . Also the lower number of mocks used has a slight effect on the agreement between the mocks and data in LOWZ curves and the CMASS-South curves. Redshift evolution between the CMASS and LOWZ is visible in the two-point function, with percentage growth between bins of approximately 10%.

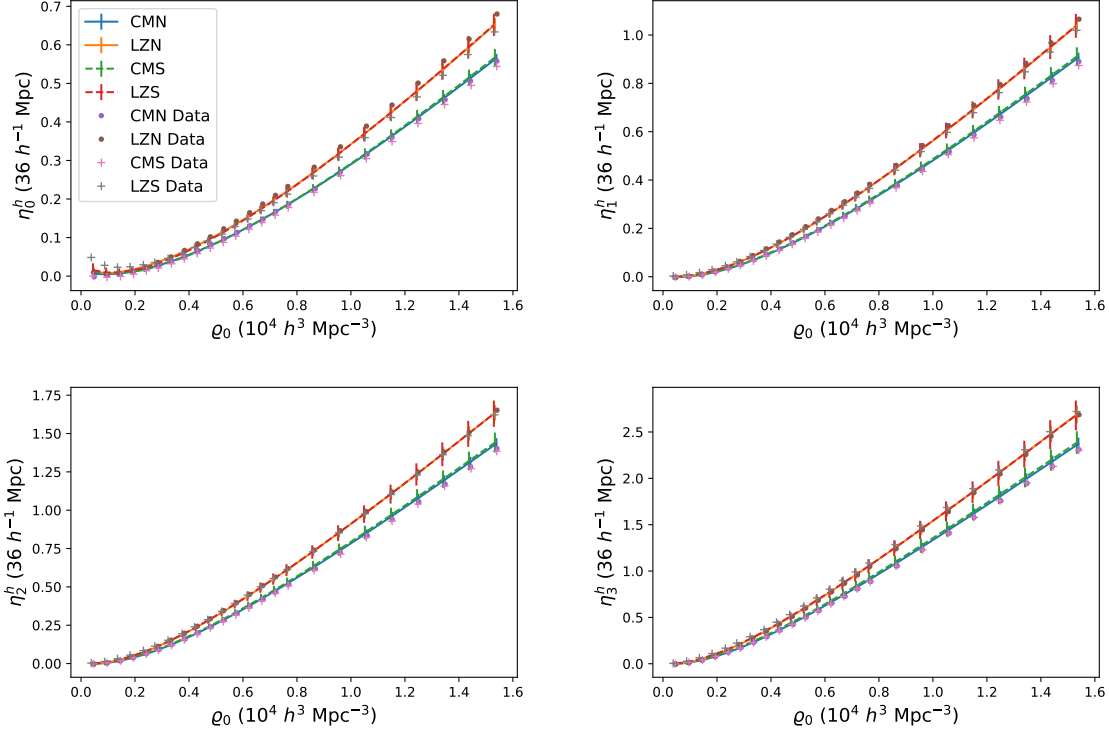
## 5.2. Higher-order Functionals

We isolate only the higher-order contributions to the MFs to determine their effect on the higher order correlations. Subtracting the two-point contribution Eqn. (9) from the power series expression for  $\eta_\mu$ , we take only the terms with  $n \geq 2$  in Eqn. (6) as the higher-order part of the functionals, which we define as  $\eta_\mu^h$ .

The higher-order functionals  $\eta_\mu^h$  are a significant portion of the calculated functionals  $\eta_\mu$ . Fig. 7 shows  $\eta_\mu^h(D = 36 h^{-1} \text{Mpc})$  for both redshift samples. We measure the significance of the detection by computing  $\chi^2$  values with 24 degrees of freedom for the mocks and the zero line and find values of order  $10^3$ , indicating a highly significant contribution of higher-order correlations to the structure of the distribution. In the limit of low sampling density, the higher-order part seems negligible, reflecting the near two-point structure characteristic of a completely Gaussian point distribution. Though it seems intuitive to expect this in the low-density limit, there is no theorem that would require a structured point process to pass through a stage of Gaussianity when the density is reduced. Regardless, the two-point behavior increasingly fails to describe the higher-order MFs as the sample density increases, meaning the integrated three and higher point functions must be non-zero, quantitatively capturing non-Gaussianity of the density field.

The data fit very well to the mocks, and nearly all the points lie within the  $1\sigma$  error bars. We draw from this agreement the conclusion that the higher-order correlations are well described by the Planck1 concordance model, insofar as the MD simulation reflects it. Quantifying this, we find  $\chi^2$  values for 24 degrees of freedom of order 10 when comparing  $\eta_\mu^h$  for the mocks and data. The complete set of  $\chi^2$  values is given in Table 2. There is a known issue with  $\eta_0^h$  in the limit of low  $\varrho_0$  where the values very slightly differ from zero, which should not occur by design. This is most visible for LOWZ-South, and may be attributed to the mask construction.

The branching between the curves for the different redshift bins exhibits the redshift evolution in the structure, which grows at higher densities. However, a more targeted view of the redshift evolution shows that the percentage difference in  $\eta_\mu^h$  is relatively constant as a function of density. Fig. 8 shows



**Figure 7.** The higher-order contribution to the MFs  $\eta_\mu^h$  as a function of density  $\rho_0$ . The solid lines are the MD-Patchy mock values with  $1\sigma$  error bars.

**Table 2.** The  $\chi^2$  values of significance for the MFs with 24 degrees of freedom. The first column gives the values for the deviation from the zero line and gives a measure of the significance of our detection of  $\eta_\mu^h$ . The second column gives the values for the deviation between the mocks and the data for  $\eta_\mu^h$ . Similarly, the third column gives the deviation between the mocks and the data for  $\eta_\mu$ .

	$\chi^2(\text{detection})$				$\chi^2(\eta_\mu^h)$				$\chi^2(\eta_\mu)$			
$\mu$	CMN	LZN	CMS	LZS	CMN	LZN	CMS	LZS	CMN	LZN	CMS	LZS
0	6293	3619	2036	1030	55	21	14	8	189	31	28	31
1	5925	3520	1792	1125	54	18	12	8	105	23	12	18
2	4417	2728	1398	911	39	14	10	6	58	15	6	10
3	2759	1804	903	641	26	9	7	5	37	14	4	6

this percentage growth from the higher redshift CMASS bin to the lower redshift LOWZ bin. The percentage growth in fact decreases from near 25% as the density increases. Compared to the two-point percentage growth of  $\approx 10\%$ , this is a factor of two greater for the higher-order contribution to the MFs.

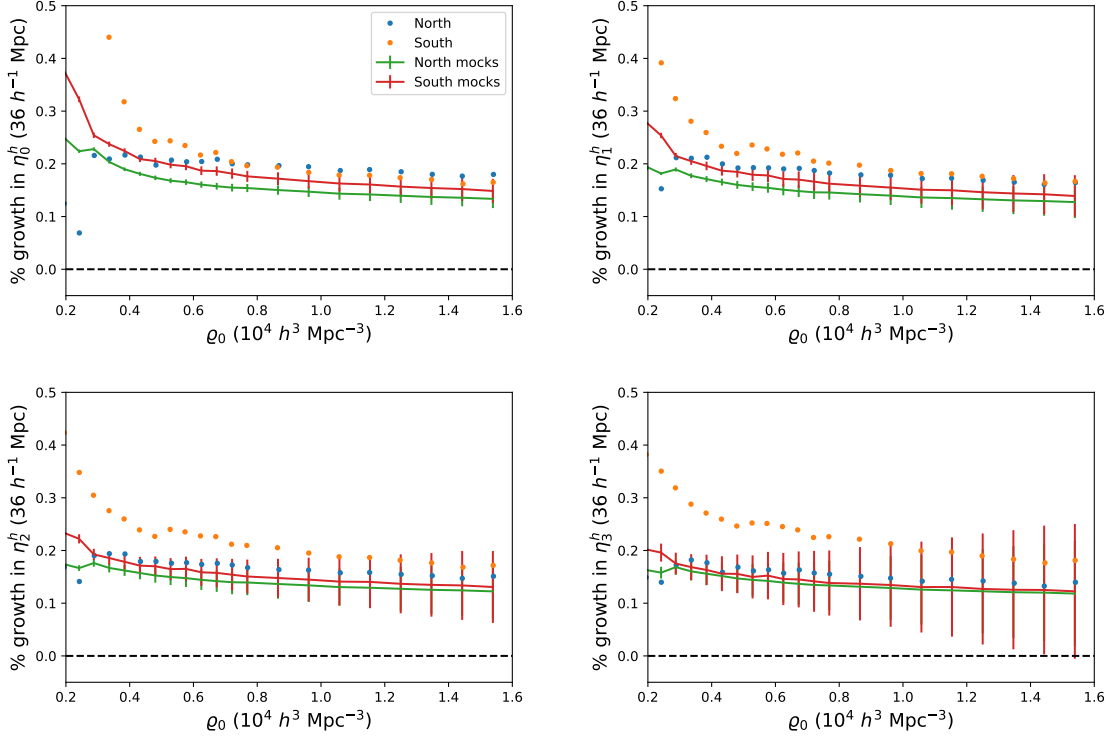
The percentage growth values for the NGC seem to agree with the mocks very well, and are close, if not within, the  $1\sigma$  error bars. For the SGC the agreement is weaker, but seems to but strengthens as the density increases. Even if we were to take the trend of the SGC values literally they are generally greater than those of the NGC or the mocks, and would indicate a higher percentage growth if anything.

However, we expect this deviating behavior to be a result of our treatment of the SGC.

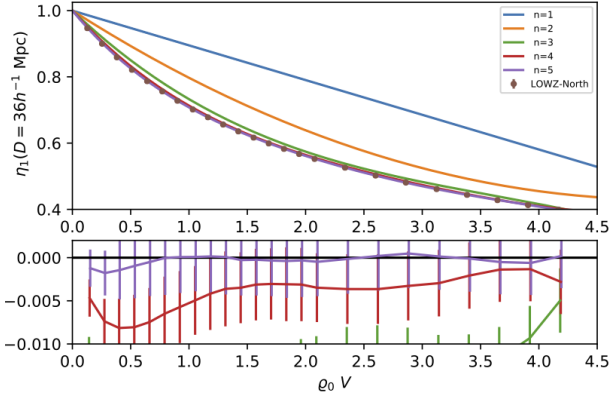
### 5.3. Correlation Coefficients

To better describe the contribution of the higher-order correlations to the MFs, we now explore the relative contributions of different orders. We employ the power series decomposition in Eqn. (6) and fit a polynomial of increasing degree  $n$  to the first  $n$  terms of the series expansion. We use the MD-Patchy mocks to produce a  $24 \times 24$  covariance matrix for  $\eta_\mu$  to account for correlation between the data points. We account for the bias from using finite terms by inflating the error bars using the correction specified by Percival et al. (2014). We used a  $\chi^2$  fitting method with 24 degrees





**Figure 8.** The percentage growth of the higher-order term of the functionals between redshift bins. Visually there is a 20%–25% growth from the higher redshift bin to the lower redshift bin. The MD-Patch mocks are the solid lines with  $1\sigma$  error bars, and the data are given by the points. The zero line is shown in black for comparison to the no-growth case.



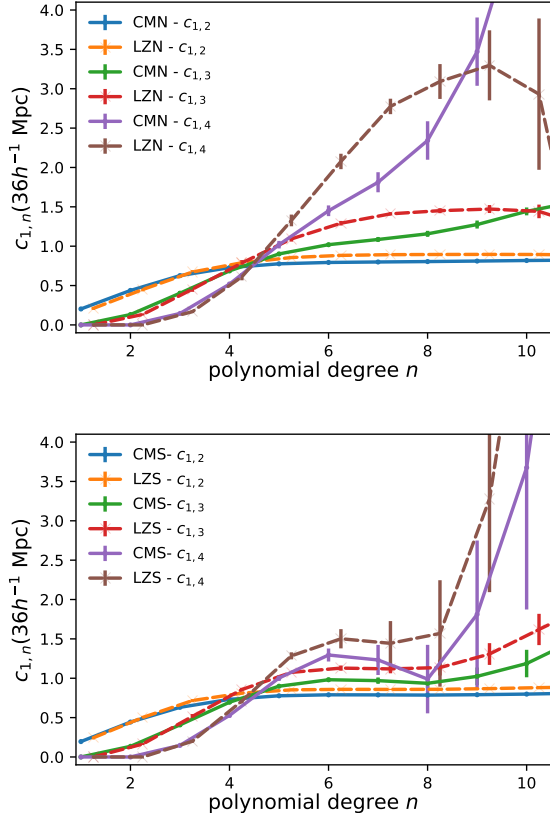
**Figure 9.** Polynomial fit for the southern sample up to degree 6 for a ball scale diameter  $D = 36 h^{-1}$  Mpc. The residual plots indicate no convergence before degree 4.

of freedom using the covariance matrix to obtain the  $c_{\mu,n}$  coefficients. Recalling Eqn (8), these coefficients are directly related to the integrated correlation functions of order  $n$ .

Fig. 9 shows the convergence of the fit in LOWZ-North, which is similar to the convergence for CMASS-North and occurs after degree  $n = 4$ . That the linear term, or two-point correlation term, provides a very bad fit is not a surprise, as the three-point correlation function has been well

established. However, the convergence after the fourth degree term, or fourth order correlation function term, shows the MFs are sensitive to fourth order statistics. Furthermore, we compute  $\chi^2$  values of order 10 for the fourth-order coefficient, and values of  $\approx 50$  and  $\approx 300$  for the third- and second-order correlations. This is true for SGC and LOWZ-North. In the DR12 paper, the CMASS-North fit showed sensitivity up to the sixth degree coefficient, and the LOWZ sensitivity is a degree lower than that. This is very likely due to the reduced size of the covariance matrix and the decrease by more than a factor of 2 in the survey area (between CMASS and LOWZ). Although the fit convergence occurs at a lower degree polynomial, it is still at a high enough degree to be interesting, due to the confirmation of the four- and five-point contributions. We expect the convergence to occur at closer to the same rate if we were to use a larger covariance matrix, but we can confidently state that fifth-order sensitivity is constant across both redshift bins.

Iterating further, we find the coefficients of the fit of the polynomial, which can be written in terms of the integrated  $n$ -point correlation functions  $\xi_n$  by Eqn. (8). Fig. 10 shows  $c_{1,n}$  for all four samples, and the stability of the two point coefficient is evident after  $n = 4$ . The fit performs very poorly after  $n = 8$ , and we expect this since the fit seems to converge at a degree of 4. There is also no restrictive prior put on



**Figure 10.** The first three coefficients  $c_{1,n+1}$  of Eqn. (8) for a fit of degree  $n$  and ball diameter of  $36 h^{-1} \text{ Mpc}$ .

the coefficients of the fourth order term as in the DR12 paper. Instead, we attempted L1 and L2 likelihood penalization regularization methods using machine learning techniques, but were unable to produce convincing values for the coefficients even at the price of an extra regularization parameter. It is evident that in each galactic cap, the LOWZ sample values are generally greater than those of the CMASS samples, potentially reflecting the more developed structure at lower redshifts.

## 6. SUMMARY & CONCLUSIONS

In this paper, we perform an unprecedented analysis of higher-order correlation information on the largest spectroscopic redshift survey to date. We use Minkowski Functional analysis along with independently calculated two-point correlation functions to measure non-Gaussianity in large-scale

structure. Our detection of non-Gaussianity is highly significant, with  $\chi^2$  values of  $\mathcal{O}(10^3)$  for 24 degrees of freedom across the NGC and SGC of the SDSS-III DR12.

We also measure the percentage growth in redshift of the higher-order part of MFs for the first time. This growth is measured to be 20%–25%, depending on density. We show the redshift evolution of non-Gaussianity is greater in the sum of the higher-order contributions to the MFs than in the two-point term by a factor of 2.

Our results motivate multiple avenues for future analysis of higher-order correlations, including measurement of the growth of the three-point and four-point functions. The growth in the higher-order correlations may also be applied in relation to the linear growth rate.

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*Software:* CHIPMINK

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