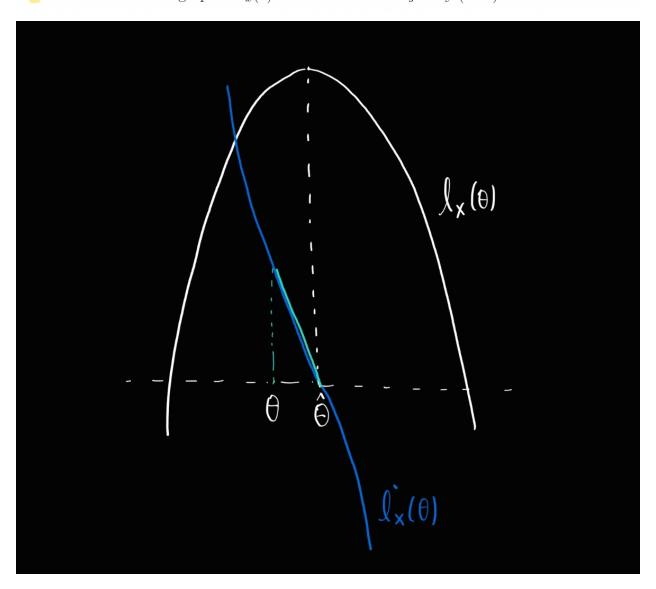
## Advanced Statistical Methods HW3

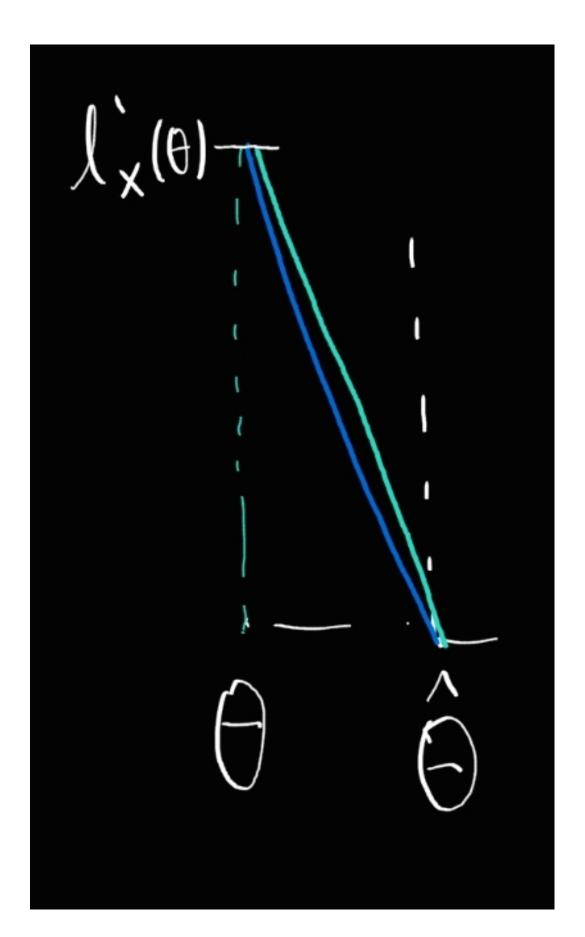
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10/3/2021

### Exercise 4.2

2. Draw a schematic graph of  $\dot{l}_x(\theta)$  versus  $\theta$ . Use it to justify (4.25).





Since loglikelihood  $l_x(\theta)$  is typically concave, we can draw a schematic graph of  $l_x(\theta)$  as a concave function of  $\theta$  which has a maximum value at  $\theta = \hat{\theta}^{MLE}$ . Near  $\hat{\theta}^{MLE}$ ,  $\dot{l}_x(\theta)$  has positive value at  $\theta < \hat{\theta}^{MLE}$ , negaive value at  $\theta > \hat{\theta}^{MLE}$ , and zero value at  $\theta = \hat{\theta}^{MLE}$  as drawn in the first figure. Now, from the second figure, we can justify (4.25)

$$\hat{\theta}^{MLE} \doteq \theta + \frac{\dot{l}_x(\theta)}{-\ddot{l}_x(\theta)}$$

The slope of  $\dot{l}_x(\theta)$  is  $\ddot{l}_x(\theta)$ , which can be approximated by

$$\ddot{l}_x(\theta) \doteq \frac{\dot{l}_x(\theta)}{\theta - \hat{\theta}^{MLE}} \quad :: \quad \dot{l}_x(\hat{\theta}^{MLE}) = 0$$

By multiplying  $\theta - \hat{\theta}^{MLE}$  and dividing by  $\ddot{l}_x(\theta)$  on both sides, we get the desired result.

#### Exercise 4.3

3. You observe  $x_1 \sim \text{Bin}(20, \theta)$  and, independently,  $x_2 \sim \text{Poi}(10 \cdot \theta)$ . Numerically compute the Cramér–Rao lower bound (4.33). Hint: Fisher information adds for independent observations.

let  $x = (x_1, x_2)$ .

$$I(\theta) = E\left[\left(\frac{\partial}{\partial \theta}\log f(x;\theta)\right)^2\right] = -E\left[\frac{\partial^2}{\partial \theta^2}\log f(x;\theta)\right] = -E\left[\frac{\partial^2}{\partial \theta^2}\log f(x_1;\theta)\right] - E\left[\frac{\partial^2}{\partial \theta^2}\log f(x_2;\theta)\right] \quad \therefore \ x_1, x_2 \quad indep.$$

We can analytically derive a closed form of  $I(\theta)$  and Cramer-Rao lower bound which is  $I(\theta)^{-1}$ 

$$f(x_1;\theta) = \binom{20}{x_1} \theta^{x_1} (1-\theta)^{20-x_1}$$

$$\log f(x_1;\theta) = \log \binom{20}{x_1} + x_1 \log \theta + (20-x_1) \log(1-\theta)$$

$$\frac{\partial}{\partial \theta} \log f(x_1;\theta) = \frac{x_1}{\theta} - \frac{20-x_1}{1-\theta}$$

$$\frac{\partial^2}{\partial \theta^2} \log f(x_1;\theta) = -\frac{x_1}{\theta^2} - \frac{20-x_1}{(1-\theta)^2}$$

$$-E\left[\frac{\partial^2}{\partial \theta^2} \log f(x_1;\theta)\right] = \frac{20\theta}{\theta^2} + \frac{20(1-\theta)}{(1-\theta)^2} = \frac{20}{\theta(1-\theta)}$$

$$f(x_2;\theta) = \frac{e^{-10\theta}(10\theta)^{x_2}}{x_2!}$$

$$\log f(x_2;\theta) = -10\theta + x_2 \log(10\theta) - \log x_2!$$

$$\frac{\partial}{\partial \theta} \log f(x_2;\theta) = -10 + \frac{x_2}{\theta}$$

$$\frac{\partial^2}{\partial \theta^2} \log f(x_2;\theta) = -\frac{x_2}{\theta^2}$$

$$-E\left[\frac{\partial^2}{\partial \theta^2} \log f(x_1;\theta)\right] = \frac{10\theta}{\theta^2} = \frac{10}{\theta}$$

Hence we have

$$I(\theta) = \frac{20}{\theta(1-\theta)} + \frac{10}{\theta} = \frac{10(3-\theta)}{\theta(1-\theta)}$$
$$CRLB(\theta) = \frac{1}{I(\theta)} = \frac{\theta(1-\theta)}{10(3-\theta)}$$

Now we have to numerically compute CRLB given observed data  $x_1$  and  $x_2$ . We shall find  $\hat{\theta}^{MLE}$  and then CRLB can be computed as  $CRLB(\hat{\theta}^{MLE})$  or 1/I(x) where

$$I(x) = -\frac{\partial^2}{\partial \theta^2} \log f(x;\theta) \Big|_{\theta = \hat{\theta}^{MLE}} = - \Big( -\frac{x_1}{\theta^2} - \frac{20 - x_1}{(1 - \theta)^2} - \frac{x_2}{\theta^2} \Big) \; \Big|_{\theta = \hat{\theta}^{MLE}} = \frac{x_1 + x_2}{\theta^2} + \frac{20 - x_1}{(1 - \theta)^2} \; \Big|_{\theta = \hat{\theta}^{MLE}}$$

is an observed Fisher information.

We shall now consider an algorithm to find  $\hat{\theta}^{MLE}$  numerically.

$$\begin{split} \dot{l}_x(\theta) &= \frac{x_1}{\theta} - \frac{20 - x_1}{1 - \theta} - 10 + \frac{x_2}{\theta} = \frac{x_1 + x_2}{\theta} - \frac{20 - x_1}{1 - \theta} - 10 \\ \ddot{l}_x(\theta) &= -\frac{x_1 + x_2}{\theta^2} - \frac{20 - x_1}{(1 - \theta)^2} < 0 \quad \forall \ 0 < \theta < 1 \\ \dot{l}_x(\theta) &= 0 \Leftrightarrow (1 - \theta)(x_1 + x_2) - \theta(20 - x_1) - 10\theta(1 - \theta) = 0 \Leftrightarrow 10\theta^2 - (30 + x_2)\theta + (x_1 + x_2) = 0 \end{split}$$

Define a function g by

$$g(\theta) = 10\theta^2 - (30 + x_2)\theta + (x_1 + x_2) = 10\left(\theta - \frac{30 + x_2}{20}\right)^2 + constant$$

Note that  $\dot{l}_x(\theta) = 0 \Leftrightarrow g(\theta) = 0$  on  $\theta \in (0,1)$ . Hence, if we can find a solution  $\theta \in (0,1)$  of  $g(\theta) = 0$ , then it is a solution of likelihood equation so that it is  $\hat{\theta}^{MLE}$ . Note that since  $\frac{30+x_2}{20} > 1.5$ , by the shape of g,  $g(\theta) = 0$  must have a root bigger than 1.5, which implies that we should guide the algorithm to do not find a solution outside of desired area (0,1).

Here we can show an algorithm to find  $\hat{\theta}^{MLE}$ . To generate a sample  $x_1, x_2$ , we set a true value of  $\theta$  as 0.7. To find a solution  $\theta$  in (0,1), we shall use Newton's method with an initial value

$$\hat{\theta}^{MME} = \frac{x_1 + 2x_2}{40} \quad \because \quad \theta = E\left[\frac{X_1}{20}\right], \ \theta = E\left[\frac{X_2}{10}\right] \Rightarrow 2\theta = E\left[\frac{X_1}{20} + \frac{X_2}{10}\right]$$

```
theta=0.7

set.seed(100)
x1=rbinom(1, size=20, prob=theta)
x2=rpois(1, lambda=10*theta)
print(c(x1, x2))
```

#### ## [1] 15 5

```
# error tolerance value for terminating the alogirthm
tol=1e-8

# function g defined above
g=function(theta, x1, x2){
    10*theta^2-(30+x2)*theta+(x1+x2)
}

# derivative of g with respect to theta
gprime=function(theta, x1, x2){
    20*theta-(30+x2)
}

# find mle given data and error tolerance
# 'mix' stands for
binompois_mle<-function(x1,x2, tol){
    # initial value is theta.mme
    theta.current=(x1+2*x2)/40</pre>
```

```
while(TRUE){
    # update theta value by newton's method
    theta.new = theta.current - g(theta.current, x1,x2) / gprime(theta.current, x1,x2)

# break the iteration if |g(theta)| < error tolerance
    if(abs(g(theta.new,x1,x2))<tol ) break

# If theta deviates from (0,1) then end the iteration and print the message
    if(theta.new > 1 | theta.new < 0) {
        print("out of parameter sapce (0,1)")
        break
    }
    theta.current = theta.new
}
return(theta.new)
}
theta.hat=binompois_mle(x1, x2, tol)
theta.hat</pre>
```

#### ## [1] 0.7192236

Given  $\theta = 0.7$ , the sample  $x_1 = 15$ ,  $x_2 = 5$  is generated. From this sample, we calculate  $\hat{\theta}^{MLE} = 0.71922$  which is quite close to the true value of  $\theta$ .

```
CRLB.true=theta*(1-theta)/(10*(3-theta))
CRLB.plugin=theta.hat*(1-theta.hat)/(10*(3-theta.hat))
obsinfo=(x1+x2)/theta.hat^2+(20-x1)/(1-theta.hat)^2
CRLB.obsinfo=1/obsinfo
print(cbind(CRLB.true, CRLB.plugin, CRLB.obsinfo))
```

```
## CRLB.true CRLB.plugin CRLB.obsinfo
## [1,] 0.009130435 0.008854047 0.009795578
```

CRLB.plugin equals to  $CRLB(\hat{\theta}^{MLE})$  and CRLB.obsinfo equals to 1/I(x) We can check that RCLB.plugin and RCLB.obsinfo, which are the resulted computation for RCLB by our numerical method, are pretty close to the true value of Cramer Rao lower bound.

#### Problem 2

### 2. Recreate Figure 4.2 with n=20 and n=100.

We shall generate random n samples from Cauchy(0,1) distribution for 10000 times. Each time we will find MLE of  $\theta$  which is a location parameter of cauchy distribution. (Note that the true value is set as 0). True

value of CRLB is  $1/I(\theta)$  where

$$f(x_1; \theta) = \frac{1}{\pi} \frac{1}{1 + (x_1 - \theta)^2}$$

$$\log f(x_1; \theta) = -\log \pi - \log\{1 + (x_1 - \theta)^2\}$$

$$\frac{\partial}{\partial \theta} \log f(x_1; \theta) = \frac{2(x_1 - \theta)}{1 + (x_1 - \theta)^2}$$

$$\frac{\partial^2}{\partial \theta^2} \log f(x_1; \theta) = \frac{2\{(x_1 - \theta)^2 - 1\}}{\{1 + (x_1 - \theta)^2\}^2}$$

$$I_1(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(x_1; \theta)\right] = \frac{1}{2}$$

$$I(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(x_1; \theta)\right] = \sum_{i=1}^n -E\left[\frac{\partial^2}{\partial \theta^2} \log f(x_i; \theta)\right] = n \cdot \frac{1}{2}$$

Thus, we have the true value of CRLB (asymptotic variance of MLE) given by

$$CRLB(\theta) = \frac{1}{n \cdot \frac{1}{2}} = \frac{2}{n}$$

Note that  $CRLB(\hat{\theta}^{MLE}) = CRLB(\theta) \Big|_{\theta = \hat{\theta}^{MLE}}$  is independent of observed data x in this case because  $CRLB(\theta)$  is constant. Fisher suggested the observed information bound 1/I(x) as an estimate for asymptotic variance of MLE, which is given as

$$I(x) = \sum_{i=1}^{n} -\frac{\partial^{2}}{\partial \theta^{2}} \log f(x_{i}; \theta) \bigg|_{\theta = \hat{\theta}^{MLE}} = \sum_{i=1}^{n} \frac{2\{1 - (x_{i} - \theta)^{2}\}}{\{1 + (x_{i} - \theta)^{2}\}^{2}} \bigg|_{\theta = \hat{\theta}^{MLE}}$$

Unlike plug in estimate of CRLB, the observed information bound 1/I(x) depends on the observed data x. Now, we shall generate a list of MLE's and the observed information bounds corresponding to 10000 times of randomly generated data samples of size n.

```
# l'(theta) where l(theta) is loglikelihood function
cauchy_lprime<-function(theta, x){</pre>
  n=length(x)
  sum=0
  for(i in 1:n){
    numerator=2*(x[i]-theta)
    denominator=1+(x[i]-theta)^2
    sum = sum + numerator/denominator
  }
  return(sum)
}
# l''(theta) where l(theta) is loglikelihood function
cauchy_ldprime<-function(theta,x){</pre>
  n=length(x)
  sum=0
  for(i in 1:n){
    numerator = 2*((x[i]-theta)^2-1)
    denominator = (1+(x[i]-theta)^2)^2
    sum = sum + numerator/denominator
  }
  return(sum)
}
```

```
# Function to find mle using Newton's method
cauchy_mle<-function(x, tol){</pre>
  # Use initial value as median since location parameter of cauchy distribution is population median
  theta.current=median(x)
  iter=0
  while(TRUE){
    theta.new = theta.current - cauchy_lprime(theta.current, x) / cauchy_ldprime(theta.current, x)
    if(abs(cauchy lprime(theta.new,x))<tol ) break</pre>
    theta.current = theta.new
    iter=iter+1
    # Print a message if the algorithm converges too slow
    if(iter>1e+5) {
      print('too slow for iteration to converge')
      break
    }
 }
  return(theta.new)
# Function to calculate observed information bound
cauchy_obs_info_bound<-function(x){</pre>
 n=length(x)
 obs_info = - cauchy_ldprime(cauchy_mle(x, tol), x)
 return(1/obs_info)
To find MLE, we need to find a solution of likelihood equation. First, we have tried to use Newton's method.
# We will generate 10000 datasets.
N=10000
\# List of MLEs and observed Information bounds will be stored in X
X=matrix(0, ncol=2, nrow=N)
colnames(X)<-c('MLE', 'Obs.Info.Bound')</pre>
# sample size
n=20
#error tolerance
tol=1e-6
set.seed(100)
for(i in 1:N){
  # Generate sample of size n from cauchy(0,1)
  x=rcauchy(n, location=0, scale=1)
  mle = cauchy_mle(x, tol)
  # Too large value of theta may be a wrong answer
  if(abs(mle)>1e+03) {
    print("Abnormal value of mle is yielded")
    # Draw another sample and find mle again
    while(abs(mle)>1e+03){
      x=rcauchy(n, location=0, scale=1)
      mle=cauchy_mle(x, tol)
```

```
}
  }
  # Calculate observed information bound
  obs_info_bound = cauchy_obs_info_bound(x)
  X[i,]=c(mle, obs_info_bound)
## [1] "Abnormal value of mle is yielded"
## [1] "too slow for iteration to converge"
## [1] "too slow for iteration to converge"
## [1] "Abnormal value of mle is yielded"
## [1] "too slow for iteration to converge"
## [1] "Abnormal value of mle is yielded"
## [1] "Abnormal value of mle is yielded"
X=as.data.frame(X)
# Sorting the list by the order statistics of observed information bound
X=X[order(X$Obs.Info.Bound),]
head(X, 10)
                 MLE Obs.Info.Bound
##
## 6155 -31.97291766
                        -3.33141318
## 5206
         15.33740838
                        -1.81613613
## 8575
         -6.82551482
                        -0.96347451
## 2856
        -0.23648680
                         0.03724461
## 6964
        -0.08596687
                         0.04185940
## 6361
          0.38495939
                         0.04203593
## 5730
        -0.22412033
                         0.04216019
## 3055
          0.20981631
                         0.04228430
## 5542
        -0.03992384
                         0.04234671
## 8977
          0.24383388
                         0.04259172
```

We can see that performance of Newton's method for finding MLE is not quite good in this case.

- Since we know the true value of  $\theta$ , we can say that too large  $\hat{\theta}$  value is abnormal to estimate  $\theta$ . Probably what the algorithm found is local maximizer of  $\dot{l}_x(\theta)$  other than  $\hat{\theta}^{MLE}$ .
- For some sample data, the algorithm does not converge even after 100000 iterations. It is too slow. As an alternative, we shall use Dekker's method which is a hybrid method combining a bisection method with a secant method.

```
# To use a function `swap`
library(seqinr)
```

```
cauchy_mle<-function(x, tol){</pre>
  # Start an algorithm with initial interval [a,b]
  # Here [a,b] is chosen as [-k, k] where k=|med(X_i)|+3
  # For bisection method to begin, l'(a)l'(b) < 0 should be satisfied.
  a = -(abs(median(x))+3)
  theta.past = a
  theta.current=abs(median(x))+3
  while(cauchy_lprime(theta.past, x)*cauchy_lprime(theta.current, x ) > 0){
    # If bisection l'(a)l'(b) < 0 is not satisfied then
    # give small fluctuation to a and b to attain the condition
   warning("the initial values does not saitsfy starting condition
   of bisection method. Shifting it a little...")
   theta.past = theta.past + rnorm(1)
   theta.current = theta.current + rnorm(1)
   a=theta.past
    # If this procedure takes too much iteration, return NA
   iter=iter+1
    if(iter>1e+3) {
     print('initial value shifing is too complicated')
     return(NA)
     break
   }
  }
  iter=0
  while(TRUE){
    # Calculate the ratio which is used instead of l''(theta)
    # This is a difference between Newton's method and secant method
   ratio = (cauchy_lprime(theta.current, x)-cauchy_lprime(theta.past, x)) / (theta.current-theta.past)
    # For current bisection interval [a,b], calculate middle point m
   middle=(a + theta.current)/2
    # Propose a updated theta value , which is mainly done by a secant method
    # If two previous theta value are the same, then secant method cannot be used
    # so that middle point is proposed as new theta.
   proposal = ifelse(cauchy_lprime(theta.current, x)-cauchy_lprime(theta.past, x) !=0 ,
                      theta.current-cauchy_lprime(theta.current, x) / ratio, middle)
    # Determine an updated theta. Proposed theta becomes updated theta if it lies between m and b
    # Other wise, middle point becomes an updated theta
   if(middle<theta.current){</pre>
      theta.new=ifelse(proposal<=theta.current & proposal>=middle, proposal, middle)
    else{
     theta.new=ifelse(proposal<=middle & proposal>=theta.current, proposal, middle)
    # End the iteration if " |l'(theta)| < error tolerance " is attained.
   if(abs(cauchy_lprime(theta.new,x))<tol ) break</pre>
    # Setting a bisection interval [a,b] for the next step
    if(cauchy_lprime(theta.new, x)*cauchy_lprime(a, x) > 0) a=theta.current
```

```
if(abs(cauchy_lprime(theta.new,x)) > abs(cauchy_lprime(a, x)) ) swap(a, theta.new)
   theta.past=theta.current
   theta.current=theta.new
    # Print a message if the algorithm converges too slow
   iter=iter+1
   if(iter>1e+5) {
      print('too slow for iteration to converge')
      break
   }
 }
 return(theta.new)
# We will generate 10000 datasets.
N=10000
\# List of MLEs and observed Information bounds will be stored in X
X=matrix(0, ncol=2, nrow=N)
colnames(X)<-c('MLE', 'Obs.Info.Bound')</pre>
# sample size
n=20
#error tolerance
tol=1e-6
set.seed(100)
for(i in 1:N){
  # Generate sample of size n from cauchy(0,1)
  x=rcauchy(n, location=0, scale=1)
  mle = cauchy_mle(x, tol)
  # Too large value of theta may be a wrong answer
  if(abs(mle)>1e+03) {
   print("Abnormal value of mle is yielded")
    \# Draw another sample and find mle again
   while(abs(mle)>1e+03){
     x=rcauchy(n, location=0, scale=1)
      mle=cauchy_mle(x, tol)
   }
  # Calculate observed information bound
  obs_info_bound = cauchy_obs_info_bound(x)
  X[i,]=c(mle, obs_info_bound)
}
## Warning in cauchy_mle(x, tol): the initial values does not saitsfy starting condition
       of bisection method. Shifting it a little...
## Warning in cauchy_mle(x, tol): the initial values does not saitsfy starting condition
       of bisection method. Shifting it a little...
```

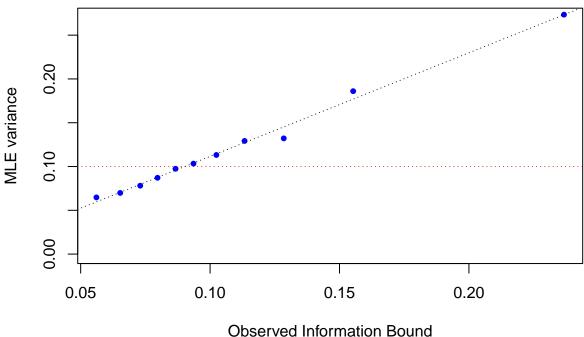
```
## Warning in cauchy_mle(x, tol): the initial values does not saitsfy starting condition
      of bisection method. Shifting it a little...
## Warning in cauchy_mle(x, tol): the initial values does not saitsfy starting condition
       of bisection method. Shifting it a little...
X=as.data.frame(X)
# Sorting the list by the order statistics of observed information bound
X=X[order(X$Obs.Info.Bound),]
head(X, 10)
##
                 MLE Obs.Info.Bound
## 8568 0.079729895
                         0.03870160
## 5734 -0.314433264
                         0.04078561
## 5546 0.080828287
                         0.04257231
## 3384 0.006496629
                         0.04289137
## 8520 -0.126163832
                         0.04293726
## 2857 -0.195008345
                         0.04328865
## 2436 0.116693888
                         0.04333881
## 2124 0.383324723
                         0.04378321
## 6973 -0.096978053
                         0.04393224
## 7833 -0.017384145
                         0.04415765
```

Thanks to Dekker's method, the problems we had when we used Newton's method are solved. We can guess that this is because we start the algorithm in restricted interval which  $\hat{\theta}^{MLE}$  is most likely to lie in. Now it is time to recreate Figure 4.2

```
# Create a function which splits the list of MLE's and observed information bounds
# by the size of observation information bounds
# and then calculate variance of MLE & median of observed information bounds in each group
splitter<-function(X){</pre>
  percent=c(0,5,15,25,35,45,55,65,75,85,95,100)
  N=nrow(X)
 last.indices=N*percent/100
  var_theta=rep(0, 11)
  med_infobound=rep(0,11)
  for(i in 2:12){
    indices=(last.indices[i-1]+1):last.indices[i]
   thetas=X[indices, 1]
   infobounds=X[indices,2]
   var_theta[i-1]=var(thetas)
   med_infobound[i-1]=median(infobounds)
  as.data.frame(cbind(var_theta, med_infobound))
}
S=splitter(X)
# Recreate Figure 4.2 with n = 20
plot(S$med_infobound, S$var_theta, col='blue', pch=20,
     xlab='Observed Information Bound', ylab='MLE variance', main='Reproducing Figure 4.2 with n=20',
     ylim=c(0,round(max(S$var_theta),2))
# CRLB(theta) which does not depend on the data
```

```
unconditional.variance= 1/(1/2 * n)
abline(h=unconditional.variance, lty='dotted', col='red')
# linear regression fit of Var(theta) ~ Med(Obs.Info.Bound)
fit<-lm(var_theta~med_infobound, data=S)</pre>
abline(fit$coefficients, lty='dotted')
```

## Reproducing Figure 4.2 with n=20



Repeat this procedure for n = 100 case.

```
# We will generate 10000 datasets.
N=10000
\# List of MLEs and observed Information bounds will be stored in X
X=matrix(0, ncol=2, nrow=N)
colnames(X)<-c('MLE', 'Obs.Info.Bound')</pre>
# sample size
n=100
#error tolerance
tol=1e-6
set.seed(100)
for(i in 1:N){
  # Generate sample of size n from cauchy(0,1)
  x=rcauchy(n, location=0, scale=1)
  mle = cauchy_mle(x, tol)
  # Too large value of theta may be a wrong answer
  if(abs(mle)>1e+03) {
```

```
print("Abnormal value of mle is yielded")
    # Draw another sample and find mle again
   while(abs(mle)>1e+03){
     x=rcauchy(n, location=0, scale=1)
     mle=cauchy_mle(x, tol)
 }
 # Calculate observed information bound
 obs_info_bound = cauchy_obs_info_bound(x)
 X[i,]=c(mle, obs_info_bound)
}
X=as.data.frame(X)
# Sorting the list by the order statistics of observed information bound
X=X[order(X$Obs.Info.Bound),]
head(X, 10)
##
                MLE Obs.Info.Bound
## 8269 0.103122404
                        0.01170760
## 3807 0.047390131
                        0.01227333
## 2657 0.147277084 0.01228737
                     0.01244905
## 302 0.146465417
                     0.01246538
## 3671 0.012381201
## 5655 0.004597226 0.01251603
## 7423  0.045802189  0.01256538
                     0.01270758
## 8301 0.045404875
## 6780 -0.155045171 0.01287148
## 9111 0.090526844
                    0.01318322
S=splitter(X)
# Recreate Figure 4.2 with n = 20
plot(S$med_infobound, S$var_theta, col='blue', pch=20,
    xlab='Observed Information Bound', ylab='MLE variance', main='Reproducing Figure 4.2 with n=100',
    ylim=c(0,round(max(S$var_theta),2))
# CRLB(theta) which does not depend on the data
unconditional.variance= 1/(1/2 * n)
abline(h=unconditional.variance, lty='dotted', col='red')
# linear regression fit of Var(theta) ~ Med(Obs.Info.Bound)
fit<-lm(var theta~med infobound, data=S)</pre>
abline(fit$coefficients, lty='dotted')
```

# Reproducing Figure 4.2 with n=100

