Spatial Geostatistical Model

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1 Full Dimensional Model Statement

$$\begin{split} \boldsymbol{y}_t &= \boldsymbol{H}_t \boldsymbol{X}_t \boldsymbol{\beta}_t + \boldsymbol{\eta}_t + \boldsymbol{\epsilon}_t \\ \boldsymbol{\beta}_t &\sim N(\boldsymbol{\mu}_{\beta}, \boldsymbol{\Sigma}_{\beta}) \\ \boldsymbol{\eta}_t &\sim N(0, \sigma_{\eta}^2 \boldsymbol{R}(\phi)) \\ \boldsymbol{\epsilon}_t &\sim N(0, \sigma_{\epsilon}^2) \\ \boldsymbol{\mu}_{\beta} &\sim N(\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0) \\ \boldsymbol{\Sigma}_{\beta} &\sim \sigma_{\beta}^2 \boldsymbol{I} \\ \sigma_{\beta}^2 &\sim IG(\alpha_{\beta}, \beta_{\beta}) \\ \sigma_{\eta}^2 &\sim IG(\alpha_{\eta}, \beta_{\eta}) \\ \sigma_{\epsilon}^2 &\sim IG(\alpha_{\epsilon}, \beta_{\epsilon}) \\ \phi &\sim IG(\alpha_{\phi}, \beta_{\phi}) \end{split}$$

2 Posterior

$$\prod_{t=1}^{T} [\boldsymbol{\beta}_t, \boldsymbol{\mu}_t, \sigma_{\beta}^2, \sigma_{\eta}^2, \sigma_{\epsilon}^2, \phi | \boldsymbol{y}_t] \propto \prod_{t=1}^{T} [\boldsymbol{y}_t | \boldsymbol{\beta}_t, \sigma_{\eta}^2, \phi, \sigma_{\epsilon}^2] [\boldsymbol{\beta}_t | \boldsymbol{\mu}_{\beta}, \sigma_{\beta}^2] [\boldsymbol{\mu}_{\beta}] [\sigma_{\beta}^2] [\sigma_{\epsilon}^2] [\phi]$$

3 Full Conditionals

3.1 Full Conditional for β_t

For t = 1, ..., T,

$$\begin{split} [\boldsymbol{\beta}_t|\cdot] &\propto [\boldsymbol{y}_t|\boldsymbol{\beta}_t, \sigma_{\eta}^2, \sigma_{\epsilon}^2, \phi][\boldsymbol{\beta}_t|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^2] \\ &\propto e^{-\frac{1}{2}(\boldsymbol{y}_t - \boldsymbol{K}_t \tilde{\boldsymbol{X}} \boldsymbol{\beta}_t)^T (\sigma_{\eta}^2 \boldsymbol{R}_t(\phi) + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} (\boldsymbol{y}_t - \boldsymbol{K}_t \tilde{\boldsymbol{X}} \boldsymbol{\beta}_t)_e - \frac{1}{2}(\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta})^T (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta})} \\ &\propto e^{-\frac{1}{2}\{\boldsymbol{\beta}_t^T (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\sigma_{\eta}^2 \boldsymbol{R}_t(\phi) + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{K}_t \tilde{\boldsymbol{X}} + (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1}) \boldsymbol{\beta}_t - 2\boldsymbol{\beta}_t^T (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\sigma_{\eta}^2 \boldsymbol{R}_t(\phi) + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{\mu}_{\beta})\}} \end{split}$$

which is Normal with mean $A^{-1}b$ and variance A^{-1} .

$$\boldsymbol{A}^{-1} = (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\sigma_{\eta}^2 \boldsymbol{R}_t(\phi) + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{K}_t \tilde{\boldsymbol{X}} + (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1})^{-1}$$

$$\approx (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{K}_t \tilde{\boldsymbol{X}} + (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1})^{-1}$$

$$\begin{split} \boldsymbol{b} &= (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\sigma_{\eta}^2 \boldsymbol{R}_t(\phi) + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{y}_t + (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1} \boldsymbol{\mu}_{\beta}) \\ &\approx (\tilde{\boldsymbol{X}}^T \boldsymbol{K}_t^T (\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^2 \boldsymbol{I})^{-1} \boldsymbol{y}_t + (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1} \boldsymbol{\mu}_{\beta}) \end{split}$$

3.2 Full Conditional for μ_{β}

$$\begin{split} [\boldsymbol{\mu}_{\beta}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{\beta}_{t}|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^{2}] [\boldsymbol{\mu}_{\beta}] \\ &\propto e^{-\frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta})^{T} (\sigma_{\beta}^{2} \boldsymbol{\Lambda})^{-1} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta})} e^{-\frac{1}{2} (\boldsymbol{\mu}_{\beta} - \boldsymbol{\mu}_{0})^{T} \sum_{0}^{-1} (\boldsymbol{\mu}_{\beta} - \boldsymbol{\mu}_{0})} \\ &\propto e^{-\frac{1}{2} (\boldsymbol{\mu}_{\beta}^{T} (T*(\sigma_{\beta}^{2} \boldsymbol{\Lambda})^{-1} + \boldsymbol{\Sigma}_{0}^{-1}) \boldsymbol{\mu}_{\beta} - 2\boldsymbol{\mu}_{\beta}^{T} (\sum_{t=1}^{T} (\sigma_{\beta}^{2} \boldsymbol{\Lambda})^{-1} \boldsymbol{\beta}_{t} + \boldsymbol{\Sigma}_{0}^{-1} \boldsymbol{\mu}_{0})) \end{split}$$

which is multivariate normal with mean $(T*(\sigma_{\beta}^2\boldsymbol{\Lambda})^{-1}+\boldsymbol{\Sigma}_0^{-1})^{-1}(\sum_{t=1}^T(\sigma_{\beta}^2\boldsymbol{\Lambda})^{-1}\boldsymbol{\beta}_t+\boldsymbol{\Sigma}_0^{-1}\boldsymbol{\mu}_0)$ and variance $(T*(\sigma_{\beta}^2\boldsymbol{\Lambda})^{-1}+\boldsymbol{\Sigma}_0^{-1})^{-1}$

3.3 Full Conditional for σ_{β}^2

$$\begin{split} [\sigma_{\beta}^2|\cdot] &\propto \prod_{t=1}^T [\boldsymbol{\beta}_t|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^2] [\sigma_{\beta}^2] \\ &\propto (\prod_{t=1}^T |\sigma_{\beta}^2 \boldsymbol{\Lambda}|^{-\frac{1}{2}}) e^{-\frac{1}{2} \sum_{t=1}^T (\boldsymbol{\beta}_t - \boldsymbol{\mu}_t)^T (\sigma_{\beta}^2 \boldsymbol{\Lambda})^{-1} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_t) \left(\sigma_{\beta}^2\right)^{-(\alpha_{\beta} + 1)} e^{-\frac{\beta_{\beta}}{\sigma_{\beta}^2}} \\ &\propto (\sigma_{\beta}^2)^{-(\alpha_{\beta} + \frac{T*|\boldsymbol{\Lambda}|}{2} + 1)} e^{-\frac{1}{\sigma_{\beta}^2} (\frac{1}{2} \sum_{t=1}^T (\boldsymbol{\beta}_t - \boldsymbol{\mu}_t)^T (\boldsymbol{\Lambda})^{-1} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_t) + \beta_{\beta})} \end{split}$$

which is $\mathrm{IG}(\alpha_{\beta} + \frac{T*|\mathbf{\Lambda}|}{2}, \frac{1}{2}\sum_{t=1}^{T}(\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{t})^{T}(\mathbf{\Lambda})^{-1}(\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{t}) + \beta_{\beta})$ where $|\mathbf{\Lambda}|$ is the determinant $\det(\mathbf{\Lambda}) = \prod_{i=1}^{J} \lambda_{j}^{2}$ where the λ_{j} are the singular values of \boldsymbol{X} .

3.4 Full Conditional for σ_{η}^2

$$\begin{split} [\sigma_{\eta}^{2}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{y}_{t}|\boldsymbol{\beta}_{t}, \sigma_{\eta}^{2}, \boldsymbol{\phi}, \sigma_{\epsilon}^{2}][\sigma_{\eta}^{2}] \\ &\propto (\prod_{t=1}^{T} |\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{1}{2}})e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\eta}^{2})^{-\alpha_{\eta}+1}e^{-\frac{\beta_{\eta}}{\sigma_{\eta}^{2}}} \\ &\propto |\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{T}{2}}e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\eta}^{2})^{-\alpha_{\eta}+1}e^{-\frac{\beta_{\eta}}{\sigma_{\eta}^{2}}} \\ &\approx |\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{T}{2}}e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\eta}^{2})^{-\alpha_{\eta}+1}e^{-\frac{\beta_{\eta}}{\sigma_{\eta}^{2}}} \end{split}$$

which can be sampled using a Metropolis-Hastings step

3.5 Full Conditional for σ_{ϵ}^2

$$\begin{split} [\sigma_{\epsilon}^{2}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{y}_{t}|\boldsymbol{\beta}_{t}, \sigma_{\eta}^{2}, \boldsymbol{\phi}, \sigma_{\epsilon}^{2}][\sigma_{\epsilon}^{2}] \\ &\propto (\prod_{t=1}^{T} |\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{1}{2}})e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\epsilon}^{2})^{-\alpha_{\epsilon}+1}e^{-\frac{\beta_{\epsilon}}{\sigma_{\epsilon}^{2}}} \\ &\propto |\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{T}{2}}e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\sigma_{\eta}^{2}\boldsymbol{R}_{t}(\boldsymbol{\phi}) + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\epsilon}^{2})^{-\alpha_{\epsilon}+1}e^{-\frac{\beta_{\epsilon}}{\sigma_{\epsilon}^{2}}} \\ &\approx |\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda}|^{-\frac{T}{2}}e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})^{T}(\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2}\boldsymbol{\Lambda})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{K}_{t}\tilde{\boldsymbol{X}}\boldsymbol{\beta}_{t})(\sigma_{\epsilon}^{2})^{-\alpha_{\epsilon}+1}e^{-\frac{\beta_{\epsilon}}{\sigma_{\epsilon}^{2}}} \end{split}$$

which can be sampled using a Metropolis-Hastings step

3.6 Full Conditional for ϕ

$$\begin{split} [\phi|\cdot] &\propto \prod_{t=1}^{T} [y_{t}|\beta_{t},\sigma_{\eta}^{2},\phi,\sigma_{\epsilon}^{2}][\phi] \\ &\propto \prod_{t=1}^{T} |\sigma_{\eta}^{2} \boldsymbol{R}(\phi) + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda}|^{-\frac{1}{2}} e^{-\frac{1}{2}} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t})^{T} (\sigma_{\eta}^{2} \boldsymbol{R}(\phi) + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda})^{-1} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t}) \phi^{-\alpha_{\phi} + 1} e^{-\frac{\beta_{\phi}}{\phi}} \\ &\propto |\sigma_{\eta}^{2} \boldsymbol{R}(\phi) + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda}||^{-\frac{T}{2}} e^{-\frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t})^{T} (\sigma_{\eta}^{2} \boldsymbol{R}(\phi) + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda})^{-1} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t}) \phi^{-\alpha_{\phi} + 1} e^{-\frac{\beta_{\phi}}{\phi}} \\ &\approx |\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda}|^{-\frac{T}{2}} e^{-\frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t})^{T} (\tilde{\boldsymbol{C}} + \sigma_{\epsilon}^{2} \boldsymbol{\Lambda})^{-1} (\boldsymbol{y}_{t} - \boldsymbol{K}_{t} \tilde{\boldsymbol{X}} \boldsymbol{\beta}_{t}) \phi^{-\alpha_{\phi} + 1} e^{-\frac{\beta_{\phi}}{\phi}} \end{split}$$

which can be sampled using a Metropolis-Hastings step