1 The Model

1.1 Data Model

$$[y_t | \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma_{\epsilon}^2] \sim \mathrm{N}\left(\boldsymbol{H}_t \boldsymbol{X}_t \boldsymbol{\beta}_{\boldsymbol{\gamma}}, \sigma_{\epsilon}^2\right)$$

where
$$\boldsymbol{eta_{\gamma}} = \left[egin{array}{c} eta_0 \\ eta_1 \gamma_1 \\ \vdots \\ eta_p \gamma_p \end{array}
ight]$$
 and $\boldsymbol{y} = \left[egin{array}{c} y_1 \\ y_2 \\ \vdots \\ y_T \end{array}
ight]$.

1.2 Process Model - climate variable of interest is AR(1) in time

$$[y_t - \mu | y_{t-1}, \alpha, \sigma_{\eta}^2] \sim N(\alpha (y_{t-1} - \mu), \sigma_{\eta}^2)$$

where μ is an intercept term, assumed to be known for the PDSI drought index ($\mu = 0$)

1.3 Parameter Model

$$[\beta_0] \propto 1$$

$$[\beta_j | \sigma_{\epsilon}^2, \gamma_j] \stackrel{iid}{\sim} N\left(0, \frac{\sigma_{\epsilon}^2 \gamma_j}{\lambda_j}\right) \qquad \text{for } j = 1, \dots, p$$

$$[\sigma_{\epsilon}^2] \propto \frac{1}{\sigma_{\epsilon}^2}$$

$$[\sigma_{\eta}^2] \propto \frac{1}{\sigma_{\eta}^2}$$

$$[\gamma_j] \propto \text{Bern}(\pi_j) \qquad \text{for } j = 1, \dots, p$$

$$[\alpha] \propto \text{Unif}(0, 1)$$

where π_j and λ_j are fixed hyperpriors for $j = 1, \dots, p$.

1.4 Posterior

$$\left[\boldsymbol{\beta}_{\boldsymbol{\gamma}}, \boldsymbol{\gamma}, \sigma_{\epsilon}^{2}, \sigma_{\eta}^{2}, \alpha \middle| \boldsymbol{y} \right] = \left(\prod_{t=1}^{T} \left[y_{t}, \middle| \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma_{\epsilon}^{2} \right] \right) \left(\prod_{t=2}^{T} \left[y_{t} - \mu, \middle| y_{t-1}, \alpha, \sigma_{\eta}^{2} \right] \right) \left[\boldsymbol{\beta} \middle| \sigma_{\epsilon}^{2}, \boldsymbol{\gamma} \right] \left[\sigma_{\epsilon}^{2} \right] \left[\sigma_{\eta}^{2} \right] \left(\prod_{j=1}^{p} \left[\gamma_{j} \right] \right) \left[\alpha \right]$$

2 Full Conditionals

2.1 Full Conditional for unobserved y_t

For t = 1,

$$[y_{1}|\cdot] \propto [y_{1}|\boldsymbol{\beta},\boldsymbol{\gamma},\sigma_{\epsilon}^{2}] [y_{2} - \mu|y_{1},\alpha,\sigma_{\eta}^{2}]$$

$$\propto \exp\left\{-\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{1} - \boldsymbol{H}_{1}\boldsymbol{\gamma}\boldsymbol{X}_{1}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}\right)^{2}\right\} \exp\left\{-\frac{1}{2\sigma_{\eta}^{2}} \left((y_{2} - \mu) - \alpha(y_{1} - \mu))^{2}\right\}$$

$$\propto \exp\left\{-\frac{1}{2} \left(y_{1}^{2} \left(\frac{1}{\sigma_{\epsilon}^{2}} + \frac{\alpha^{2}}{\sigma_{\eta}^{2}}\right) - 2y_{1} \left(\frac{\boldsymbol{H}_{1}\boldsymbol{\gamma}\boldsymbol{X}_{1}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}}{\sigma_{\epsilon}^{2}} + \frac{\alpha(y_{2} - (1 - \alpha)\mu)}{\sigma_{\eta}^{2}}\right)\right)(\right\}$$

which is
$$N(A^{-1}b, A^{-1})$$
 where $A^{-1} = \left(\frac{1}{\sigma_{\epsilon}^2} + \frac{\alpha^2}{\sigma_{\eta}^2}\right)^{-1}$ and $b = \frac{H_1 \gamma X_1 \gamma \beta \gamma}{\sigma_{\epsilon}^2} + \frac{\alpha(y_2 - (1 - \alpha)\mu)}{\sigma_{\eta}^2}$

For t = 2, ..., T - 1,

$$[y_{t}|\cdot] \propto [y_{t}|\boldsymbol{\beta},\boldsymbol{\gamma},\sigma_{\epsilon}^{2}] [y_{t}-\mu|y_{t-1},\alpha,\sigma_{\eta}^{2}] [y_{t+1}-\mu|y_{t},\alpha,\sigma_{\eta}^{2}]$$

$$\propto \exp\left\{-\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{t}-\boldsymbol{H}_{t}\boldsymbol{\gamma}\boldsymbol{X}_{t}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}\right)^{2}\right\} \exp\left\{-\frac{1}{2\sigma_{\eta}^{2}} \left((y_{t}-\mu)-\alpha\left(y_{t-1}-\mu\right)\right)^{2}\right\} \exp\left\{-\frac{1}{2\sigma_{\eta}^{2}} \left((y_{t+1}-\mu)-\alpha\left(y_{t}-\mu\right)\right)^{2}\right\} \exp\left\{-\frac{1}{2\sigma_{\eta}^{2}} \left((y_{t}-\mu)-\alpha\left(y_{t}-\mu\right)\right)^{2}\right\} \exp\left\{-\frac{1}{2\sigma_{\eta}^{2}} \left((y_{t}-\mu)-\alpha\left(y_{t}-\mu\right)\right)^{2}\right$$

which is N $\left(A^{-1}b, A^{-1}\right)$ where $A^{-1} = \left(\frac{1}{\sigma_{\epsilon}^2} + \frac{\alpha^2 + 1}{\sigma_{\eta}^2}\right)^{-1}$ and $b = \frac{\mathbf{H}_t \gamma \mathbf{X}_t \gamma \beta \gamma}{\sigma_{\epsilon}^2} + \frac{\alpha y_{t-1} + (1-\alpha)\mu}{\sigma_{\eta}^2} + \frac{\alpha (y_{t+1} - (1-\alpha)\mu)}{\sigma_{\eta}^2}$. For t = T,

$$\begin{aligned} &[y_T|\cdot] \propto \left[y_T \middle| \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma_{\epsilon}^2 \right] \left[y_T - \mu \middle| y_{T-1}, \alpha, \sigma_{\eta}^2 \right] \\ &\propto \exp\left\{ -\frac{1}{2\sigma_{\epsilon}^2} \left(y_T - \boldsymbol{H}_T \boldsymbol{\gamma} \boldsymbol{X}_T \boldsymbol{\gamma} \boldsymbol{\beta}_{\boldsymbol{\gamma}} \right)^2 \right\} \exp\left\{ -\frac{1}{2\sigma_{\eta}^2} \left((y_T - \mu) - \alpha \left(y_{T-1} - \mu \right) \right)^2 \right\} \\ &\propto \exp\left\{ -\frac{1}{2} \left(y_T^2 \left(\frac{1}{\sigma_{\epsilon}^2} + \frac{1}{\sigma_{\eta}^2} \right) - 2y_T \left(\frac{\boldsymbol{H}_T \boldsymbol{\gamma} \boldsymbol{X}_T \boldsymbol{\gamma} \boldsymbol{\beta}_{\boldsymbol{\gamma}}}{\sigma_{\epsilon}^2} + \frac{\alpha y_{T-1} + (1 - \alpha) \mu}{\sigma_{\eta}^2} \right) \right) \right\} \end{aligned}$$

which is N
$$\left(A^{-1}b,A^{-1}\right)$$
 where $A^{-1}=\left(\frac{1}{\sigma_{\epsilon}^2}+\frac{1}{\sigma_{\eta}^2}\right)^{-1}$ and $b=\frac{\boldsymbol{H}_T\boldsymbol{\gamma}\boldsymbol{X}_T\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}}{\sigma_{\epsilon}^2}+\frac{\alpha y_{T-1}+(1-\alpha)\mu}{\sigma_{\eta}^2}$.

2.2 Full Conditional for β

$$\begin{split} [\boldsymbol{\beta}|\cdot] &\propto \prod_{t=1}^{T} \left[y_{t} \middle| \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma_{\epsilon}^{2} \right] \left[\boldsymbol{\beta} \middle| \boldsymbol{\gamma}, \sigma_{\eta}^{2} \right] \\ &\propto \prod_{t=1}^{T} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{t} - \boldsymbol{H}_{t} \boldsymbol{\gamma} \boldsymbol{X}_{t} \boldsymbol{\gamma} \boldsymbol{\beta} \boldsymbol{\gamma} \right)^{2} \right\} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T} \boldsymbol{\Lambda} \boldsymbol{\gamma} \boldsymbol{\beta} \boldsymbol{\gamma} \right\} \\ &\propto \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \left[\boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T} \left(\sum_{t=1}^{T} \boldsymbol{X}_{t} \boldsymbol{\gamma}^{T} \boldsymbol{H}_{t} \boldsymbol{\gamma}^{T} \boldsymbol{H}_{t} \boldsymbol{\gamma} \boldsymbol{X}_{t} \boldsymbol{\gamma} + \boldsymbol{\Lambda} \boldsymbol{\gamma} \right) \boldsymbol{\beta}_{\boldsymbol{\gamma}} - 2 \boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T} \left(\sum_{t=1}^{T} \boldsymbol{X}_{t} \boldsymbol{\gamma}^{T} \boldsymbol{H}_{t} \boldsymbol{\gamma}^{T} \boldsymbol{y}_{t} \right) \right] \right\} \end{split}$$

which is MVN
$$(\boldsymbol{A}^{-1}\boldsymbol{b}, \boldsymbol{A}^{-1})$$
 where $\boldsymbol{A}^{-1} = \left(\frac{\boldsymbol{X}_{t}^{T}\boldsymbol{\gamma}\boldsymbol{H}_{t}^{T}\boldsymbol{\gamma}\boldsymbol{H}_{t}\boldsymbol{\gamma}\boldsymbol{X}_{t}\boldsymbol{\gamma} + \boldsymbol{\Lambda}\boldsymbol{\gamma}}{\sigma_{\epsilon}^{2}}\right)^{-1}$ and $\boldsymbol{b} = \frac{\boldsymbol{X}_{t}^{T}\boldsymbol{\gamma}\boldsymbol{H}_{t}^{T}\boldsymbol{\gamma}\boldsymbol{y}_{t}}{\sigma_{\epsilon}^{2}}$

2.3 Full Conditional for σ_{ϵ}^2

$$\begin{split} \left[\sigma_{\epsilon}^{2}\big|\cdot\right] &\propto \prod_{t=1}^{T}\left[y_{t}\big|\boldsymbol{\beta},\boldsymbol{\gamma},\sigma_{\epsilon}^{2}\right]\left[\boldsymbol{\beta}\big|\boldsymbol{\gamma},\sigma_{\epsilon}^{2}\right]\left[\sigma_{\epsilon}^{2}\right] \\ &\propto \prod_{t=1}^{T}\left(\sigma_{\epsilon}^{2}\right)^{-\frac{1}{2}}\exp\left\{-\frac{1}{\sigma_{\epsilon}^{2}}\frac{\left(y_{t}-\boldsymbol{H}_{t}\boldsymbol{\gamma}\boldsymbol{X}_{t}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}\right)^{2}}{2}\right\}\big|\sigma_{\epsilon}^{2}\boldsymbol{\Lambda}\boldsymbol{\gamma}\big|^{-\frac{1}{2}}\exp\left\{-\frac{1}{\sigma_{\epsilon}^{2}}\frac{\boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T}\boldsymbol{\Lambda}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}}{2}\right\}\left(\sigma_{\epsilon}^{2}\right)^{-1} \\ &\propto \left(\sigma_{\epsilon}^{2}\right)^{-\frac{T+p\boldsymbol{\gamma}}{2}-1}\exp\left\{-\frac{1}{\sigma_{\epsilon}^{2}}\frac{\sum_{t=1}^{T}\left(y_{t}-\boldsymbol{H}_{t}\boldsymbol{\gamma}\boldsymbol{X}_{t}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}\right)^{2}+\boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T}\boldsymbol{\Lambda}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}}{2}\right\} \end{split}$$
 which is IG $\left(\frac{T+p\boldsymbol{\gamma}}{2},\frac{\sum_{t=1}^{T}\left(y_{t}-\boldsymbol{H}_{t}\boldsymbol{\gamma}\boldsymbol{X}_{t}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}\right)^{2}+\boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T}\boldsymbol{\Lambda}\boldsymbol{\gamma}\boldsymbol{\beta}\boldsymbol{\gamma}}{2}\right)$

2.4 Full Conditional for σ_{η}^2

$$\begin{split} \left[\sigma_{\eta}^{2}\middle|\cdot\right] &\propto \prod_{t=2}^{T} \left[y_{t}\middle|y_{t-1}, \alpha, \sigma_{\eta}^{2}\right] \left[\sigma_{\eta}^{2}\right] \\ &\propto \prod_{t=2}^{T} \left(\sigma_{\eta}^{2}\right)^{-\frac{1}{2}} \exp\left\{-\frac{1}{\sigma_{\eta}^{2}} \frac{\left(\left(y_{t}-\mu\right)-\alpha\left(y_{t-1}-\mu\right)\right)^{2}}{2}\right\} \left(\sigma_{\eta}^{2}\right)^{-1} \\ &\propto \left(\sigma_{\eta}^{2}\right)^{-\frac{T-1}{2}-1} \exp\left\{-\frac{1}{\sigma_{\eta}^{2}} \frac{\sum_{t=2}^{T} \left(\left(y_{t}-\mu\right)-\alpha\left(y_{t-1}-\mu\right)\right)^{2}}{2}\right\} \end{split}$$

which is IG $\left(\frac{T-1}{2}, \frac{\sum_{t=2}^{T}((y_t-\mu)-\alpha(y_{t-1}-\mu))^2}{2}\right)$

2.5 Full Conditional for α

$$\begin{split} & [\alpha|\cdot] \propto \prod_{t=2}^{T} \left[y_{t} \middle| y_{t-1}, \alpha, \sigma_{\eta}^{2} \right] [\alpha] \\ & \propto \prod_{t=2}^{T} \exp \left\{ -\frac{1}{2\sigma_{\eta}^{2}} \left((y_{t} - \mu) - \alpha \left(y_{t-1} - \mu \right) \right)^{2} \right\} \boldsymbol{I} \left\{ \alpha \in (0, 1) \right\} \\ & \propto \exp \left\{ -\frac{1}{2} \left(\alpha^{2} \frac{\sum_{t=2}^{T} \left(y_{t-1} - \mu \right)^{2}}{\sigma_{\eta}^{2}} - 2\alpha \frac{\sum_{t=2}^{T} \left(y_{t-1} - \mu \right) \left(y_{t} - \mu \right)}{\sigma_{\eta}^{2}} \right) \right\} \boldsymbol{I} \left\{ \alpha \in (0, 1) \right\} \end{split}$$

which is Truncated Normal $(A^{-1}b, A^{-1})$ where $A^{-1} = \left(\frac{\sum_{t=2}^{T}(y_{t-1}-\mu)^2}{\sigma_{\eta}^2}\right)^{-1}$ and $b = \frac{\sum_{t=2}^{T}(y_{t-1}-\mu)(y_t-\mu)}{\sigma_{\eta}^2}$ restricted to the interval (0,1).

2.6 Full Conditional for γ_j

$$\begin{aligned} \left[\boldsymbol{\gamma} | \cdot \right] &\propto \prod_{t=1}^{T} \left[y_{t} \middle| \boldsymbol{\beta}, \boldsymbol{\gamma}, \sigma_{\epsilon}^{2} \right] \left[\boldsymbol{\gamma} \right] \\ &\propto \prod_{t=1}^{T} \left(\sigma^{2} \right)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{t} - \boldsymbol{H}_{t} \boldsymbol{\gamma} \boldsymbol{X}_{t} \boldsymbol{\gamma} \boldsymbol{\beta}_{\boldsymbol{\gamma}} \right)^{2} \right\} \left| \sigma_{\epsilon}^{2} \boldsymbol{\Lambda}_{\boldsymbol{\gamma}} \right|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T} \boldsymbol{\Lambda}_{\boldsymbol{\gamma}} \boldsymbol{\beta}_{\boldsymbol{\gamma}} \right\} \left(\prod_{j=1}^{p} \pi_{j}^{\gamma_{j}} \left(1 - \pi_{j} \right)^{(1 - \gamma_{j})} \right) \\ &\propto \left(\sigma^{2} \right)^{-\frac{T + p}{2}} \left| \boldsymbol{\Lambda}_{\boldsymbol{\gamma}} \right|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{t} - \boldsymbol{H}_{t} \boldsymbol{\gamma} \boldsymbol{X}_{t} \boldsymbol{\gamma} \boldsymbol{\beta}_{\boldsymbol{\gamma}} \right)^{2} \right\} \exp \left\{ -\frac{1}{2\sigma_{\epsilon}^{2}} \boldsymbol{\beta}_{\boldsymbol{\gamma}}^{T} \boldsymbol{\Lambda}_{\boldsymbol{\gamma}} \boldsymbol{\beta}_{\boldsymbol{\gamma}} \right\} \left(\prod_{j=1}^{p} \pi_{j}^{\gamma_{j}} \left(1 - \pi_{j} \right)^{(1 - \gamma_{j})} \right) \end{aligned}$$

which can be sampled using a Metropolis Hastings proposal. This requires a smart choice of proposal distribution.

If H_tX_t has orthogonal columns (e.g $X_t^TH_t^TH_tX_t$ is diagonal) then **ADD IN A SUM OVER ALL T**

$$\begin{split} & [\gamma_{j}|\cdot] \propto \prod_{t=1}^{T} \left[y_{t}, \left|\beta_{j}, \gamma_{j}, \sigma_{\epsilon}^{2}\right] \left[\beta_{j} \middle| \gamma_{j}, \sigma^{2}\right] \left[\gamma_{j}\right] \\ & \propto \exp\left\{-\frac{1}{2\sigma_{\epsilon}^{2}} \left(y_{t} - \boldsymbol{H}_{tj} \boldsymbol{X}_{tj} \gamma_{j} \beta_{j}\right)^{2}\right\} \left(\frac{\sigma_{\epsilon}^{2}}{\lambda_{j}}\right)^{-\frac{\gamma_{j}}{2}} \exp\left\{-\frac{\gamma_{j} \lambda_{j} \beta_{j}^{2}}{2\sigma^{2}}\right\} \pi^{\gamma_{j}} \left(1 - \pi\right)^{1 - \gamma_{j}} \\ & \propto \left(\frac{\sigma_{\epsilon}^{2}}{\lambda_{j}}\right)^{-\frac{\gamma_{j}}{2}} \exp\left\{-\frac{1}{2\sigma_{\epsilon}^{2}} \gamma_{j} \left(\beta_{j}^{2} \left(\boldsymbol{X}_{tj}^{T} \boldsymbol{H}_{tj}^{T} \boldsymbol{H}_{tj} \boldsymbol{X}_{tj} \lambda_{j}\right) - 2y_{t} \boldsymbol{H}_{tj} \boldsymbol{X}_{tj} \beta_{j}\right)\right\} \pi^{\gamma_{j}} \left(1 - \pi\right)^{1 - \gamma_{j}} \\ & \propto \left(\left(\frac{\sigma_{\epsilon}^{2}}{\lambda_{j}}\right)^{-\frac{1}{2}} \exp\left\{-\frac{1}{2\sigma_{\epsilon}^{2}} \left(\beta_{j}^{2} \left(\boldsymbol{X}_{tj}^{T} \boldsymbol{H}_{tj}^{T} \boldsymbol{H}_{tj} \boldsymbol{X}_{tj} \lambda_{j}\right) - 2y_{t} \boldsymbol{H}_{tj} \boldsymbol{X}_{tj} \beta_{j}\right)\right\} \frac{\pi_{j}}{1 - \pi_{j}}\right)^{\gamma_{j}} \\ & \propto \Psi^{\gamma_{j}} \end{split}$$

so
$$\gamma_j$$
 is Bern $\left(\frac{\Psi_j}{1+\Psi_j}\right)$ where $\Psi_j = \left(\frac{\lambda_j}{\sigma^2}\right)^{\frac{1}{2}} \exp\left\{-\frac{1}{2\sigma^2}\left[\beta_j^2\left(\delta_j + \lambda_j\right) - 2\delta_j\beta_j^2\right]\right\} \frac{\pi}{1-\pi}$