# Spatial Predictive Process Model

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### 1 Full Dimensional Model Statement

#### 1.1 Data Model

$$oldsymbol{y}_t = oldsymbol{H}_t oldsymbol{X} oldsymbol{eta}_t + oldsymbol{\eta}_t + oldsymbol{\epsilon}_t$$

### 1.2 Process Model

$$\begin{split} \boldsymbol{\beta}_t &\sim N(\boldsymbol{\mu}_{\beta}, \boldsymbol{\Sigma}_{\beta}) & \boldsymbol{\Sigma}_{\beta} = \sigma_{\beta}^2 \boldsymbol{I}_t \\ \boldsymbol{\eta}_t &\sim N(0, \boldsymbol{\Sigma}_{\eta}) & \boldsymbol{\Sigma}_{\eta} = \sigma_{\eta}^2 \boldsymbol{R}(\phi) & \boldsymbol{R}(\phi) = \exp\left(-\boldsymbol{D}_t/\phi\right) \\ \boldsymbol{\epsilon}_t &\sim N(0, \boldsymbol{\Sigma}_{\epsilon}) & \boldsymbol{\Sigma}_{\epsilon} = \sigma_{\epsilon}^2 \boldsymbol{I}_t \end{split}$$

#### 1.3 Parameter Model

$$\begin{split} & \boldsymbol{\mu}_{\beta} \sim N(\boldsymbol{\mu}_{0}, \boldsymbol{\Sigma}_{0}) \\ & \sigma_{\beta}^{2} \sim IG(\alpha_{\beta}, \beta_{\beta}) \\ & \sigma_{\eta}^{2} \sim IG(\alpha_{\eta}, \beta_{\eta}) \\ & \sigma_{\epsilon}^{2} \sim IG(\alpha_{\epsilon}, \beta_{\epsilon}) \\ & \phi \sim IG(\alpha_{\phi}, \beta_{\phi}) \end{split}$$

where  $I_{\beta}$  is the identity matrix of size  $\tau \times \tau$  where  $\tau$  is the number of parameters in  $\beta_t$ ,  $I_t$  is the identity matrix of size  $n_t \times n_t$  and  $n_t$  is the number of samples of  $y_t$  at time t and  $D_t$  is the distance matrix between locations observed at time t. Define  $\Sigma = \Sigma_{\eta} + \Sigma_{\epsilon}$ 

### 2 Predictive Process

For large dimensional spatial processes it can be computationally expensive to invert  $\Sigma_t$  over the set of all locations  $s_t$  at time t from the set of all locations  $s_t$ . This motivates the use of a predictive process  $\tilde{\eta}$  to approximate  $\eta_t$  over a set of knots  $s_t$  where  $\tilde{\eta} = c_t(s_t, s^* | \sigma_{\eta}^2, \phi)^T C^{*-1}(s^*, s^* | \sigma_{\eta}^2, \phi) \eta^*$ . The covariance between the desired locations  $s_t \in S$  at time t and the set of knots  $s_t \in S^*$  is  $c_t(s_t, s^* | \sigma_{\eta}^2, \phi)$ . The covariance matrix over the set of knots is  $C^{*-1}(s^*, s^* | \sigma_{\eta}^2, \phi)$ . The lower dimensional  $\eta^* \sim \text{MVN}(0, C^{*-1}(s^*, s^* | \sigma_{\eta}^2, \phi))$ . Equivalently,  $\tilde{\eta} \sim MVN(0, c_t^T C^{*-1} c_t)$ . This model needs to be modified so as to not underestimate the variance. This is done by replacing  $\tilde{\eta}_t$  with  $\tilde{\eta}_t + \tilde{\epsilon}_t$  where  $\tilde{\epsilon}_t = MVN(0, \sigma_{\eta}^2 I_{n_t} - \text{diag}(c_t^T C^{*-1} c_t))$ 

### 2.1 Data Model

$$oldsymbol{y}_t = oldsymbol{H}_t oldsymbol{X} oldsymbol{eta}_t + ilde{oldsymbol{\eta}}_t + ilde{oldsymbol{\epsilon}}_t$$

### 2.2 Process Model

$$\begin{split} \boldsymbol{\beta}_t &\sim N(\boldsymbol{\mu}_{\beta}, \boldsymbol{\Sigma}_{\beta}) & \boldsymbol{\Sigma}_{\beta} = \sigma_{\beta}^2 \boldsymbol{I}_t \\ \tilde{\boldsymbol{\eta}}_t &\sim N(0, \boldsymbol{\Sigma}_{\tilde{\eta}}) & \boldsymbol{\Sigma}_{\tilde{\eta}} = \boldsymbol{c}^T \boldsymbol{C}^{*-1} \boldsymbol{c} \\ \tilde{\boldsymbol{\epsilon}}_t &\sim N(0, \boldsymbol{\Sigma}_{\tilde{\epsilon}}) & \boldsymbol{\Sigma}_{\tilde{\epsilon}} = (\sigma_{\epsilon}^2 + \sigma_{\eta}^2) \boldsymbol{I}_t - \boldsymbol{c}^T \boldsymbol{C}^{*-1} \boldsymbol{c} \end{split}$$

### 2.3 Parameter Model

$$\mu_{\beta} \sim N(\mu_{0}, \Sigma_{0})$$

$$\sigma_{\beta}^{2} \sim IG(\alpha_{\beta}, \beta_{\beta})$$

$$\sigma_{\eta}^{2} \sim IG(\alpha_{\eta}, \beta_{\eta})$$

$$\sigma_{\epsilon}^{2} \sim IG(\alpha_{\epsilon}, \beta_{\epsilon})$$

$$\phi \sim IG(\alpha_{\phi}, \beta_{\phi})$$

### 3 Posterior

$$\prod_{t=1}^{T} [\boldsymbol{\beta}_{t}, \boldsymbol{\mu}_{\beta}, \sigma_{\beta}^{2}, \sigma_{\eta}^{2}, \sigma_{\epsilon}^{2}, \phi | \boldsymbol{y}_{t}] \propto \prod_{t=1}^{T} [\boldsymbol{y}_{t} | \boldsymbol{\beta}_{t}, \sigma_{\eta}^{2}, \phi, \sigma_{\epsilon}^{2}] [\boldsymbol{\beta}_{t} | \boldsymbol{\mu}_{\beta}, \sigma_{\beta}^{2}] [\boldsymbol{\mu}_{\beta}] [\sigma_{\beta}^{2}] [\sigma_{\eta}^{2}] [\sigma_{\epsilon}^{2}] [\phi]$$

## 4 Full Conditionals

### 4.1 Full Conditional for $\beta_t$

For t = 1, ..., T,

$$\begin{split} [\boldsymbol{\beta}_t|\cdot] &\propto [\boldsymbol{y}_t|\boldsymbol{\beta}_t, \sigma_{\eta}^2, \sigma_{\epsilon}^2, \boldsymbol{\phi}] [\boldsymbol{\beta}_t|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^2] \\ &\propto e^{-\frac{1}{2}} (\boldsymbol{y}_t - \boldsymbol{H}_t \boldsymbol{X} \boldsymbol{\beta}_t)^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{y}_t - \boldsymbol{H}_t \boldsymbol{X} \boldsymbol{\beta}_t)_e - \frac{1}{2} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta})^T \boldsymbol{\Sigma}_{\beta}^{-1} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta}) \\ &\propto e^{-\frac{1}{2}} \{\boldsymbol{\beta}_t^T (\boldsymbol{X}^T \boldsymbol{H}_t^T \boldsymbol{\Sigma}^{-1} \boldsymbol{H}_t \boldsymbol{X} + \boldsymbol{\Sigma}_{\beta}^{-1}) \boldsymbol{\beta}_t - 2 \boldsymbol{\beta}_t^T (\boldsymbol{X}^T \boldsymbol{H}_t^T \boldsymbol{\Sigma}^{-1} \boldsymbol{y}_t + \boldsymbol{\Sigma}_{\beta}^{-1} \boldsymbol{\mu}_{\beta}) \} \end{split}$$

which is Normal with mean  $\boldsymbol{A}^{-1}\boldsymbol{b}$  and variance  $\boldsymbol{A}^{-1}$  where

$$\begin{split} \boldsymbol{A}^{-1} &= (\boldsymbol{X}^T \boldsymbol{H}_t^T \boldsymbol{\Sigma}^{-1} \boldsymbol{H}_t \boldsymbol{X} + \boldsymbol{\Sigma}_{\beta}^{-1})^{-1} \\ \boldsymbol{b} &= (\boldsymbol{X}^T \boldsymbol{H}_t^T \boldsymbol{\Sigma}^{-1} \boldsymbol{y}_t + \boldsymbol{\Sigma}_{\beta}^{-1} \boldsymbol{\mu}_{\beta}) \end{split}$$

### 4.2 Full Conditional for $\mu_{\beta}$

$$\begin{split} [\boldsymbol{\mu}_{\beta}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{\beta}_{t}|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^{2}][\boldsymbol{\mu}_{\beta}] \\ &\propto e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta})^{T} \boldsymbol{\Sigma}_{\beta}^{-1} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta})} e^{-\frac{1}{2} (\boldsymbol{\mu}_{\beta} - \boldsymbol{\mu}_{0})^{T} \boldsymbol{\Sigma}_{0}^{-1} (\boldsymbol{\mu}_{\beta} - \boldsymbol{\mu}_{0})} \\ &\propto e^{-\frac{1}{2} (\boldsymbol{\mu}_{\beta}^{T} (T \boldsymbol{\Sigma}_{\beta}^{-1} + \boldsymbol{\Sigma}_{0}^{-1}) \boldsymbol{\mu}_{\beta} - 2 \boldsymbol{\mu}_{\beta}^{T} (\sum_{t=1}^{T} \boldsymbol{\Sigma}_{\beta}^{-1} \boldsymbol{\beta}_{t} + \boldsymbol{\Sigma}_{0}^{-1} \boldsymbol{\mu}_{0})) \end{split}$$

which is multivariate normal with mean  $(T\Sigma_{\beta}^{-1} + \Sigma_{0}^{-1})^{-1}(\sum_{t=1}^{T}\Sigma_{\beta}^{-1}\beta_{t} + \Sigma_{0}^{-1}\mu_{0})$  and variance  $(T\Sigma_{\beta}^{-1} + \Sigma_{0}^{-1})^{-1}$ 

## 4.3 Full Conditional for $\sigma_{\beta}^2$

$$\begin{split} [\sigma_{\beta}^2|\cdot] &\propto \prod_{t=1}^T [\boldsymbol{\beta}_t|\boldsymbol{\mu}_{\beta}, \sigma_{\beta}^2] [\sigma_{\beta}^2] \\ &\propto (\prod_{t=1}^T |\boldsymbol{\Sigma}_{\beta}|^{-\frac{1}{2}}) e^{-\frac{1}{2} \sum_{t=1}^T (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta})^T \boldsymbol{\Sigma}_{\beta}^{-1} (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta}) (\sigma_{\beta}^2)^{-(\alpha_{\beta}+1)} e^{-\frac{\beta_{\beta}}{\sigma_{\beta}^2}} \\ &\propto (\sigma_{\beta}^2)^{-(\alpha_{\beta} + \frac{\sum_{t=1}^T n_t}{2} + 1)} e^{-\frac{1}{\sigma_{\beta}^2} (\frac{1}{2} \sum_{t=1}^T (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta})^T (\boldsymbol{\beta}_t - \boldsymbol{\mu}_{\beta}) + \beta_{\beta})} \end{split}$$

which is  $IG(\alpha_{\beta} + \frac{\sum_{t=1}^{T} n_{t}}{2}, \frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta})^{T} (\boldsymbol{\beta}_{t} - \boldsymbol{\mu}_{\beta}) + \beta_{\beta})$  since the determinant  $|\boldsymbol{\Sigma}_{\beta}| = (\sigma_{\beta}^{2})^{n_{t}}$  and  $\boldsymbol{\Sigma}_{\beta}^{-1} = \frac{1}{\sigma_{\beta}^{2}} \boldsymbol{I}_{t}$ 

## 4.4 Full Conditional for $\sigma_{\eta}^2$

$$\begin{split} [\sigma_{\eta}^{2}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{y}_{t}|\boldsymbol{\beta}_{t}, \sigma_{\eta}^{2}, \phi, \sigma_{\epsilon}^{2}][\sigma_{\eta}^{2}] \\ &\propto (\prod_{t=1}^{T} |\boldsymbol{\Sigma}|^{-\frac{1}{2}}) e^{-\frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t} \boldsymbol{X} \boldsymbol{\beta}_{t})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t} \boldsymbol{X} \boldsymbol{\beta}_{t}) (\sigma_{\eta}^{2})^{-\alpha_{\eta} + 1} e^{-\frac{\beta_{\eta}}{\sigma_{\eta}^{2}}} \\ &\propto |\boldsymbol{\Sigma}|^{-\frac{T}{2}} e^{-\frac{1}{2} \sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t} \boldsymbol{X} \boldsymbol{\beta}_{t})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t} \boldsymbol{X} \boldsymbol{\beta}_{t}) (\sigma_{\eta}^{2})^{-\alpha_{\eta} + 1} e^{-\frac{\beta_{\eta}}{\sigma_{\eta}^{2}}} \end{split}$$

which can be sampled using a Metropolis-Hastings step

# 4.5 Full Conditional for $\sigma_{\epsilon}^2$

$$\begin{split} [\sigma_{\epsilon}^{2}|\cdot] &\propto \prod_{t=1}^{T} [\boldsymbol{y}_{t}|\boldsymbol{\beta}_{t}, \sigma_{\eta}^{2}, \phi, \sigma_{\epsilon}^{2}][\sigma_{\epsilon}^{2}] \\ &\propto (\prod_{t=1}^{T} |\boldsymbol{\Sigma}|^{-\frac{1}{2}}) e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t}\boldsymbol{X}\boldsymbol{\beta}_{t})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t}\boldsymbol{X}\boldsymbol{\beta}_{t}) (\sigma_{\epsilon}^{2})^{-\alpha_{\epsilon} + 1} e^{-\frac{\beta_{\epsilon}}{\sigma_{\epsilon}^{2}}} \\ &\propto |\boldsymbol{\Sigma}|^{-\frac{T}{2}} e^{-\frac{1}{2}\sum_{t=1}^{T} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t}\boldsymbol{X}\boldsymbol{\beta}_{t})^{T} \boldsymbol{\Sigma}^{-1} (\boldsymbol{y}_{t} - \boldsymbol{H}_{t}\boldsymbol{X}\boldsymbol{\beta}_{t}) (\sigma_{\epsilon}^{2})^{-\alpha_{\epsilon} + 1} e^{-\frac{\beta_{\epsilon}}{\sigma_{\epsilon}^{2}}} \end{split}$$

which can be sampled using a Metropolis-Hastings step

### 4.6 Full Conditional for $\phi$

$$\begin{split} [\phi|\cdot] &\propto \prod_{t=1}^{T} [y_t|\beta_t, \sigma_{\eta}^2, \phi, \sigma_{\epsilon}^2][\phi] \\ &\propto \prod_{t=1}^{T} |\mathbf{\Sigma}|^{-\frac{1}{2}} e^{-\frac{1}{2}} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \mathbf{\Sigma}^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \phi^{-\alpha_{\phi} + 1} e^{-\frac{\beta_{\phi}}{\phi}} \\ &\propto |\mathbf{\Sigma}|^{-\frac{T}{2}} e^{-\frac{1}{2}} \sum_{t=1}^{T} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \mathbf{\Sigma}^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \phi^{-\alpha_{\phi} + 1} e^{-\frac{\beta_{\phi}}{\phi}} \end{split}$$

which can be sampled using a Metropolis-Hastings step

### 5 Posterior Predictive Distribution

The posterior predictive distribution for  $\boldsymbol{y}_t$  is sampled a each MCMC iteration k by

$$\boldsymbol{y}_t^{(k)} \sim N(\boldsymbol{H_t} \boldsymbol{X} \boldsymbol{\beta}_t^{(k)}, \boldsymbol{\Sigma}^{(k)})$$