

Spatial Geostatistical Model

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1 Full Dimensional Model Statement

1.1 Data Model

$$\mathbf{y}_t = \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t + \boldsymbol{\eta}_t + \boldsymbol{\epsilon}_t$$

1.2 Process Model

$$\begin{aligned} \boldsymbol{\beta}_t &\sim N(\mathbf{0}, \boldsymbol{\Sigma}_\beta) & \boldsymbol{\Sigma}_\beta &= \sigma_\epsilon^2 \mathbf{D}_\gamma & \mathbf{D}_\gamma &= \text{diag}(\gamma_1^2, \dots, \gamma_\tau^2) \\ \boldsymbol{\eta}_t &\sim N(0, \boldsymbol{\Sigma}_{\eta_t}) & \boldsymbol{\Sigma}_{\eta_t} &= \sigma_\eta^2 \mathbf{R}_t(\phi) & \mathbf{R}_t(\phi) &= \exp(-\mathbf{D}_t/\phi) \\ \boldsymbol{\epsilon}_t &\sim N(0, \boldsymbol{\Sigma}_{\epsilon_t}) & \boldsymbol{\Sigma}_{\epsilon_t} &= \sigma_\epsilon^2 \mathbf{I}_{n_t} \\ \gamma_i^2 &\sim \text{Exp}(\lambda^2/2) \text{ Equivalently } & \frac{1}{\gamma_i^2} &\sim IG(1, \lambda^2/2) & \text{for } i = 1 \dots, \tau \end{aligned}$$

1.3 Parameter Model

$$\begin{aligned} \sigma_\eta^2 &\sim IG(\alpha_\eta, \beta_\eta) \\ \sigma_\epsilon^2 &\sim IG(\alpha_\epsilon, \beta_\epsilon) \\ \phi &\sim IG(\alpha_\phi, \beta_\phi) \\ \lambda^2 &\sim \Gamma(\alpha_\lambda, \beta_\lambda) \end{aligned}$$

where \mathbf{I}_{n_t} is the identity matrix of size $n_t \times n_t$ where τ is the number of parameters in $\boldsymbol{\beta}_t$ and n_t is the number of samples of \mathbf{y}_t at time t and \mathbf{D}_t is the distance matrix between locations observed at time t . Define $\boldsymbol{\Sigma}_t = \boldsymbol{\Sigma}_{\eta_t} + \boldsymbol{\Sigma}_{\epsilon_t}$

2 Posterior

$$\prod_{t=1}^T [\boldsymbol{\beta}_t, \gamma^2, \sigma_\eta^2, \sigma_\epsilon^2, \phi, \lambda^2 | \mathbf{y}_t] \propto \prod_{t=1}^T [\mathbf{y}_t | \boldsymbol{\beta}_t, \sigma_\eta^2, \phi, \sigma_\epsilon^2] [\boldsymbol{\beta}_t | \sigma_\epsilon^2, \gamma^2] [\gamma | \lambda^2] [\sigma_\eta^2] [\sigma_\epsilon^2] [\phi] [\lambda^2]$$

3 Full Conditionals

3.1 Full Conditional for β_t

For $t = 1, \dots, T$,

$$\begin{aligned} [\beta_t | \cdot] &\propto [\mathbf{y}_t | \beta_t, \sigma_\eta^2, \sigma_\epsilon^2, \phi][\beta_t | \boldsymbol{\mu}_\beta, \sigma_\beta^2] \\ &\propto \exp\left(-\frac{1}{2}(\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)^T \boldsymbol{\Sigma}_t^{-1}(\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)\right) \exp\left(-\frac{1}{2}\beta_t^T \boldsymbol{\Sigma}_\beta^{-1} \beta_t\right) \\ &\propto \exp\left(-\frac{1}{2}\{\beta_t^T (\mathbf{X}^T \mathbf{H}_t^T \boldsymbol{\Sigma}_t^{-1} \mathbf{H}_t \mathbf{X} + \boldsymbol{\Sigma}_\beta^{-1}) \beta_t - 2\beta_t^T (\mathbf{X}^T \mathbf{H}_t^T \boldsymbol{\Sigma}_t^{-1} \mathbf{y}_t)\}\right) \end{aligned}$$

which is Normal with mean $\mathbf{A}^{-1} \mathbf{b}$ and variance \mathbf{A}^{-1} where

$$\begin{aligned} \mathbf{A}^{-1} &= (\mathbf{X}^T \mathbf{H}_t^T \boldsymbol{\Sigma}_t^{-1} \mathbf{H}_t \mathbf{X} + \boldsymbol{\Sigma}_\beta^{-1})^{-1} \\ \mathbf{b} &= (\mathbf{X}^T \mathbf{H}_t^T \boldsymbol{\Sigma}_t^{-1} \mathbf{y}_t) \end{aligned}$$

3.2 Full Conditional for γ^2

$$\begin{aligned} [\gamma^2 | \cdot] &\propto \prod_{t=1}^T [\beta_t | \sigma_\epsilon^2, \gamma^2] \prod_{i=1}^\tau [\gamma_i^2 | \lambda^2] \\ &\propto \prod_{t=1}^T |\boldsymbol{\Sigma}_\beta|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}\beta_t^T \boldsymbol{\Sigma}_\beta^{-1} \beta_t\right) \prod_{i=1}^\tau \exp\left(-\frac{\lambda^2}{2}\gamma_i^2\right) \\ &\propto |\mathbf{D}_{\gamma^2}|^{-\frac{\tau}{2}} \exp\left(-\frac{1}{2} \sum_{t=1}^T \beta_t^T \boldsymbol{\Sigma}_\beta^{-1} \beta_t\right) \prod_{i=1}^\tau \exp\left(-\frac{\lambda^2}{2}\gamma_i^2\right) \end{aligned}$$

which can be sampled using a Metropolis-Hastings step.

3.3 Full Conditional for σ_η^2

$$\begin{aligned} [\sigma_\eta^2 | \cdot] &\propto \prod_{t=1}^T [\mathbf{y}_t | \beta_t, \sigma_\eta^2, \phi, \sigma_\epsilon^2][\sigma_\eta^2] \\ &\propto \left(\prod_{t=1}^T |\boldsymbol{\Sigma}_t|^{-\frac{1}{2}}\right) \exp\left(-\frac{1}{2} \sum_{t=1}^T (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)\right) (\sigma_\eta^2)^{-\alpha_\eta - 1} \exp\left(-\frac{\beta_\eta}{\sigma_\eta^2}\right) \\ &\propto |\boldsymbol{\Sigma}_t|^{-\frac{\tau}{2}} \exp\left(-\frac{1}{2} \sum_{t=1}^T (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \beta_t)\right) (\sigma_\eta^2)^{-\alpha_\eta - 1} \exp\left(-\frac{\beta_\eta}{\sigma_\eta^2}\right) \end{aligned}$$

which can be sampled using a Metropolis-Hastings step

3.4 Full Conditional for σ_ϵ^2

$$\begin{aligned}
[\sigma_\epsilon^2|\cdot] &\propto \prod_{t=1}^T [\mathbf{y}_t|\boldsymbol{\beta}_t, \sigma_\eta^2, \phi, \sigma_\epsilon^2][\boldsymbol{\beta}_t|\sigma_\epsilon^2, \boldsymbol{\gamma}][\sigma_\epsilon^2] \\
&\propto \left(\prod_{t=1}^T |\boldsymbol{\Sigma}_t|^{-\frac{1}{2}} \exp \left(-\frac{1}{2} \sum_{t=1}^T (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \right) |\boldsymbol{\Sigma}_\beta|^{-\frac{1}{2}} \exp \left(-\frac{1}{2} \sum_{t=1}^T \boldsymbol{\beta}_t^T \boldsymbol{\Sigma}_\beta^{-1} \boldsymbol{\beta}_t \right) \right. \\
&\quad \times (\sigma_\epsilon^2)^{-\alpha_\epsilon - 1} \exp \left(-\frac{\beta_\epsilon}{\sigma_\epsilon^2} \right) \\
&\propto |\boldsymbol{\Sigma}_t|^{-\frac{T}{2}} \exp \left(-\frac{1}{2} \sum_{t=1}^T (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \right) |\boldsymbol{\Sigma}_\beta|^{-\frac{1}{2}} \exp \left(-\frac{1}{2} \sum_{t=1}^T \boldsymbol{\beta}_t^T \boldsymbol{\Sigma}_\beta^{-1} \boldsymbol{\beta}_t \right) \\
&\quad \times (\sigma_\epsilon^2)^{-\alpha_\epsilon - 1} \exp \left(-\frac{\beta_\epsilon}{\sigma_\epsilon^2} \right)
\end{aligned}$$

which can be sampled using a Metropolis-Hastings step

3.5 Full Conditional for ϕ

$$\begin{aligned}
[\phi|\cdot] &\propto \prod_{t=1}^T [\mathbf{y}_t|\boldsymbol{\beta}_t, \sigma_\eta^2, \phi, \sigma_\epsilon^2][\phi] \\
&\propto \prod_{t=1}^T |\boldsymbol{\Sigma}_t|^{-\frac{1}{2}} \exp \left(-\frac{1}{2} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \right) \phi^{-\alpha_\phi - 1} \exp \left(-\frac{\beta_\phi}{\phi} \right) \\
&\propto |\boldsymbol{\Sigma}_t|^{-\frac{T}{2}} \exp \left(-\frac{1}{2} \sum_{t=1}^T (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t)^T \boldsymbol{\Sigma}_t^{-1} (\mathbf{y}_t - \mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t) \right) \phi^{-\alpha_\phi - 1} \exp \left(-\frac{\beta_\phi}{\phi} \right)
\end{aligned}$$

which can be sampled using a Metropolis-Hastings step

Full Conditional for λ^2

$$\begin{aligned}
[\lambda^2|\cdot] &\propto \prod_{i=1}^\tau [\gamma_i^2|\lambda^2][\lambda^2] \\
[\lambda^2|\cdot] &\propto \prod_{i=1}^\tau \left(\frac{\lambda^2}{2} \exp \left(-\frac{\lambda^2}{2} \gamma_i^2 \right) \right) \left(\frac{\lambda^2}{2} \right)^{\alpha_\lambda - 1} \exp \left(-\beta_\lambda (\lambda^2/2) \right) \\
&\propto (\lambda^2)^{\alpha_\lambda + \tau - 1} \exp \left(-\lambda^2 (\beta_\lambda/2 + \sum_{j=1}^\tau \gamma_j^2/2) \right)
\end{aligned}$$

which is $\text{Gamma}(\alpha_\lambda + \tau, \beta_\lambda/2 + \sum_{j=1}^\tau \gamma_j^2/2)$

4 Posterior Predictive Distribution

The posterior predictive distribution for \mathbf{y}_t is sampled at each MCMC iteration k by

$$\mathbf{y}_t^{(k)} \sim N(\mathbf{H}_t \mathbf{X} \boldsymbol{\beta}_t^{(k)}, \boldsymbol{\Sigma}^{(k)})$$