信号检测与估值

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2019年9月

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ch3. 信号检测与估计理论的基础知识

ch3-3. 贝叶斯准则—例题(续)及性能分析

🕕 贝叶斯准则例题 4

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贝叶斯准则例题 4

考虑以下信号检测问题:

$$H_0: x_k = 1 + n_k,$$
 $k = 1, 2, ..., N$
 $H_1: x_k = -1 + n_k,$ $k = 1, 2, ..., N$

其中 n_k 是均值为零, 方差为 $\sigma_n^2 = 1/2$ 的高斯随机变量, 且不同采样时刻的加性噪声之间是相互统计独立的。

若两种假设先验等概的,且代价因子为 $c_{00} = 1, c_{10} = 4, c_{11} = 2, c_{01} = 3$ 。请给出上述问题的贝叶斯检测准则和平均代价 C。

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贝叶斯准则例题 4: 判决概率分布分析

$$H_0: x_k = 1 + n_k, \quad H_1: x_k = -1 + n_k$$

 $k = 1, 2, ..., N, \quad \mathbf{x} = (x_1, x_2, ..., x_N)^T$
 $P(H_i|H_j) = \int_{\mathcal{D}} p(\mathbf{x}|H_j) d\mathbf{x}$

$$\begin{array}{c|c} p(I|H_1) & & p(I|H_0) \\ \hline p(I|H_1) & & p(I|H_0) \\ \hline P(H_1|H_1) & & P(H_0|H_0) \\ \hline & & & & \\ \hline P(H_1|H_0) & P(H_0|H_1) \\ \hline \end{array}$$

$$P(H_0|H_0) = \int_{L_0} p(\mathbf{x}|H_0) d\mathbf{x}, \qquad P(H_1|H_0) = \int_{L_1} p(\mathbf{x}|H_0) d\mathbf{x}$$

$$P(H_0|H_1) = \int_{L_0} p(\mathbf{x}|H_1) d\mathbf{x}, \qquad P(H_1|H_1) = \int_{L_1} p(\mathbf{x}|H_1) d\mathbf{x}$$

$$\mathbf{L} = L_0 \cup L_1, \quad L_0 \cap L_1 = \emptyset, \quad \int_{\mathbf{L}} p(\mathbf{x}|H_j) d\mathbf{x} = 1$$

$$P(H_0|H_0) + P(H_1|H_0) = \int_{L_0} p(\mathbf{x}|H_0) d\mathbf{x} + \int_{L_1} p(\mathbf{x}|H_0) d\mathbf{x} = \int_{\mathbf{L}} p(\mathbf{x}|H_0) d\mathbf{x} = 1$$

$$P(H_0|H_1) + P(H_1|H_1) = \int_{\mathbf{L}} p(\mathbf{x}|H_1) d\mathbf{x} + \int_{\mathbf{L}} p(\mathbf{x}|H_1) d\mathbf{x} = \int_{\mathbf{L}} p(\mathbf{x}|H_1) d\mathbf{x} = 1$$

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贝叶斯准则例题 4: 解

解: N 次独立采样, 样本为 $x_k(k=1,2,\cdots,N)$:

$$H_0: x_k = 1 + n_k$$
 $k = 1, 2, \dots, N$
 $H_1: x_k = -1 + n_k$ $k = 1, 2, \dots, N$

步骤 1: 计算两个似然函数, 构建似然比

由于n是高斯分布随机变量,因此在 H_0 假设下,第k次采样值 x_k 服从高斯分布,

且均值为 1, 方差为 σ_n^2 ; 在 H_1 假设下, 第 k 次采样值 x_k 服从均值为 -1, 方差为 σ_n^2 的高斯分布。

$$p(x_k|H_0) = \left(\frac{1}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{(x_k - 1)^2}{2\sigma_n^2}\right) \implies p(\mathbf{x}|H_0) = \prod_{k=1}^N \left(\frac{1}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{(x_k - 1)^2}{2\sigma_n^2}\right)$$

$$p(x_k|H_1) = \left(\frac{1}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{(x_k + 1)^2}{2\sigma_n^2}\right) \implies p(\mathbf{x}|H_1) = \prod_{k=1}^N \left(\frac{1}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{(x_k + 1)^2}{2\sigma_n^2}\right)$$

$$\lambda(\mathbf{x}) = \frac{p(\mathbf{x}|H_1)}{p(\mathbf{x}|H_0)} = \exp\left(\frac{\sum_{k=1}^N \left((x_k - 1)^2 - (x_k + 1)^2\right)}{2\sigma_n^2}\right)$$

贝叶斯准则例题 4: 解续(1)

步骤 2: 根据两个假设的先验概率和代价因子, 计算判决门限

$$\eta \stackrel{\text{def}}{=} \frac{P(H_0)(c_{10} - c_{00})}{P(H_1)(c_{01} - c_{11})} = \frac{4 - 1}{3 - 2} = 3$$

步骤 3: 形成贝叶斯检测基本表达式

$$\lambda(\mathbf{x}) = \frac{p(\mathbf{x}|H_1)}{p(\mathbf{x}|H_0)} \mathop{\gtrsim}_{H_0}^{H_1} \eta$$

$$\exp\left(\frac{\sum\limits_{k=1}^{N}\left((x_{k}-1)^{2}-(x_{k}+1)^{2}\right)}{2\sigma_{n}^{2}}\right)\underset{H_{0}}{\overset{H_{1}}{\geqslant}}\eta$$

步骤 4: 化简, 形成贝叶斯检测判决表达式

$$-4\sum_{k=1}^{N} x_{k} \underset{H_{0}}{\gtrless} 2\sigma_{n}^{2} \ln \eta \implies \sum_{k=1}^{N} x_{k} \underset{H_{1}}{\gtrless} -\frac{\sigma_{n}^{2} \ln \eta}{2}$$

$$l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^{N} x_k \underset{H_1}{\overset{H_0}{\geq}} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{def}{=} \gamma$$

贝叶斯准则例题 4: 解续(2)

经过上述化简,信号检测的判决式由似然比检验的形式,简化为检验统计量 l(x) 与检测门限 γ 相比较的形式,形成贝叶斯检测判决表达式:

$$l(\mathbf{x}) \stackrel{\text{def}}{=} \frac{1}{N} \sum_{k=1}^{N} x_k \underset{H_1}{\gtrless} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{\text{def}}{=} \gamma$$

检验统计量 $l(x) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^{N} x_k$ 是观测信号 $x_k (k = 1, 2, ..., N)$ 的求和取平均值的结果,即它是 $x_k (k = 1, 2, ..., N)$ 的函数,是一个随机变量。

因为高斯随机变量的线性组合还是高斯随机变量,所以两种假设下的观测量 $(l|H_0),(l|H_1)$ 也是服从高斯分布的随机变量。

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_0)$

$$\begin{array}{l} H_0: x_k = 1 + n_k \\ H_1: x_k = -1 + n_k \end{array} \ l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k \underset{H_1}{\stackrel{H_0}{\geqslant}} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{def}{=} \gamma \quad$$
 统计量: $l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k$

假设 H_0 条件下, 统计量 l(x) 为高斯分布, 均值和方差分别为

$$E[l|H_0] = E\left[\frac{1}{N}\sum_{k=1}^{N}(x_k|H_0)\right] = E\left[\frac{1}{N}\sum_{k=1}^{N}(1+n_k)\right] = \frac{1}{N}\sum_{k=1}^{N}E[1+n_k] = 1$$

$$Var[l|H_0] = E\left[(l|H_0 - E(l|H_0))^2\right] = E\left[\left(\frac{1}{N}\sum_{k=1}^{N}(1+n_k) - 1\right)^2\right]$$

$$= \frac{1}{N^2}\sum_{k=1}^{N}E[n_k^2] = \frac{1}{N^2}\sum_{k=1}^{N}\sigma_n^2 = \frac{\sigma_n^2}{N}$$

因此,
$$(l|H_0) \sim \mathcal{N}(1, \frac{\sigma_n^2}{N})$$

$$p(l|H_0) = \left(\frac{1}{2\pi Var[l|H_0]}\right)^{1/2} \exp\left(-\frac{(l-E[l|H_0])^2}{2Var[l|H_0]}\right) = \left(\frac{N}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{N(l-1)^2}{2\sigma_n^2}\right)$$

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_0)$

$$H_{0}: x_{k} = 1 + n_{k}$$

$$H_{1}: x_{k} = -1 + n_{k}$$

$$l(\mathbf{x}) \stackrel{\text{def}}{=} \frac{1}{N} \sum_{k=1}^{N} x_{k} \stackrel{H_{0}}{\geq} -\frac{\sigma_{n}^{2} \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{\text{def}}{=} \gamma$$

$$f(l|H_{0}) = \left(\frac{N}{2\pi\sigma_{n}^{2}}\right)^{1/2} \exp\left(-\frac{N(l-1)^{2}}{2\sigma_{n}^{2}}\right)$$

$$P(H_{1}|H_{0}) = 1 - \int_{\gamma}^{\infty} p(l|H_{0})dl \implies Q(u) = \int_{x}^{\infty} \left(\frac{1}{2\pi}\right)^{1/2} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= 1 - \int_{\gamma}^{\infty} \left(\frac{N}{2\pi\sigma_{n}^{2}}\right)^{1/2} \exp\left(-\frac{N(l-1)^{2}}{2\sigma_{n}^{2}}\right) dl \qquad \text{by } l = \frac{\sigma_{n}u}{\sqrt{N}} + 1$$

$$= 1 - \int_{\frac{\sqrt{N}(\gamma-1)}{\sigma_{n}}}^{\infty} \left(\frac{1}{2\pi}\right)^{1/2} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= 1 - Q\left(\frac{\sqrt{N}(\gamma-1)}{\sigma_{n}}\right) = 1 - Q\left(\frac{\sqrt{N}\left(-\frac{\sigma_{n}^{2} \ln \eta}{2N} - 1\right)}{\sigma_{n}}\right)$$

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_0)$

$$\begin{split} H_0: x_k &= 1 + n_k \\ H_1: x_k &= -1 + n_k \end{split} \quad l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k \stackrel{H_0}{\geq} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{def}{=} \gamma \end{split} \quad \mathbf{统} \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k \\ P(H_1|H_0) &= 1 - \mathcal{Q}\left(\frac{\sqrt{N}\left(-\frac{\sigma_n^2 \ln \eta}{2N} - 1\right)}{\sigma_n}\right) \\ &= 1 - \mathcal{Q}\left(-\frac{\sigma_n \ln \eta}{2\sqrt{N}} - \frac{\sqrt{N}}{\sigma_n}\right) \\ &= 1 - \mathcal{Q}\left(-\frac{\ln \eta}{d} - \frac{d}{2}\right) \\ P(H_0|H_0) &= 1 - P(H_1|H_0) \end{split}$$

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_1)$

$$H_0: x_k = 1 + n_k$$
 $l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k \underset{H_1}{\gtrless} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{def}{=} \gamma$ 统计量: $l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k$

假设 H1 条件下, 统计量 l(x) 为高斯分布, 均值和方差分别为

$$E[l|H_1] = E\left[\frac{1}{N}\sum_{k=1}^{N}(x_k|H_1)\right] = E\left[\frac{1}{N}\sum_{k=1}^{N}(-1+n_k)\right] = -1 + \frac{1}{N}\sum_{k=1}^{N}E[n_k] = -1$$

$$Var[l|H_1] = E\left[(l|H_1 - E(l|H_1))^2\right] = E\left[\left(\frac{1}{N}\sum_{k=1}^{N}(-1+n_k) + 1\right)^2\right]$$

$$= \frac{1}{N^2}\sum_{k=1}^{N}E[n_k^2] = \frac{1}{N^2}\sum_{k=1}^{N}\sigma_n^2 = \frac{\sigma_n^2}{N}$$

因此, $(l|H_1) \sim \mathcal{N}(-1, \frac{\sigma_n^2}{N})$

$$p(l|H_1) = \left(\frac{1}{2\pi Var[l|H_1]}\right)^{1/2} \exp\left(-\frac{(l-E[l|H_1])^2}{2Var[l|H_1]}\right) = \left(\frac{N}{2\pi\sigma_n^2}\right)^{1/2} \exp\left(-\frac{N(l+1)^2}{2\sigma_n^2}\right)$$

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_1)$

$$H_{0}: x_{k} = 1 + n_{k}$$

$$H_{1}: x_{k} = -1 + n_{k}$$

$$l(\mathbf{x}) \stackrel{\text{def}}{=} \frac{1}{N} \sum_{k=1}^{N} x_{k} \stackrel{H_{0}}{\geq} -\frac{\sigma_{n}^{2} \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{\text{def}}{=} \gamma$$

$$f(l|H_{1}) = \left(\frac{N}{2\pi\sigma_{n}^{2}}\right)^{1/2} \exp\left(-\frac{N(l+1)^{2}}{2\sigma_{n}^{2}}\right)$$

$$P(H_{0}|H_{1}) = \int_{\gamma}^{\infty} p(l|H_{1})dl \implies Q(u) = \int_{x}^{\infty} \left(\frac{1}{2\pi}\right)^{1/2} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= \int_{\gamma}^{\infty} \left(\frac{N}{2\pi\sigma_{n}^{2}}\right)^{1/2} \exp\left(-\frac{N(l+1)^{2}}{2\sigma_{n}^{2}}\right) dl \qquad \text{by } l = \frac{\sigma_{n}u}{\sqrt{N}} - 1$$

$$= \int_{\frac{\sqrt{N}(\gamma+1)}{\sigma_{n}}}^{\infty} \left(\frac{1}{2\pi}\right)^{1/2} \exp\left(-\frac{u^{2}}{2}\right) du$$

$$= Q\left(\frac{\sqrt{N}(\gamma+1)}{\sigma_{n}}\right) = Q\left(\frac{\sqrt{N}\left(-\frac{\sigma_{n}^{2} \ln \eta}{2N} + 1\right)}{\sigma_{n}}\right)$$

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贝叶斯准则例题 4: 性能分析—观测量 $(l|H_1)$

$$\begin{split} H_0: x_k &= 1 + n_k \\ H_1: x_k &= -1 + n_k \end{split} \ l(\mathbf{x}) \stackrel{\text{def}}{=} \frac{1}{N} \sum_{k=1}^N x_k \stackrel{H_0}{\geqslant} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{\text{def}}{=} \gamma \end{split} \ \text{统计量: } l(\mathbf{x}) \stackrel{\text{def}}{=} \frac{1}{N} \sum_{k=1}^N x_k \\ P(H_0|H_1) &= Q\left(\frac{\sqrt{N} \left(-\frac{\sigma_n^2 \ln \eta}{2N} + 1\right)}{\sigma_n}\right) \\ &= Q\left(-\frac{\sigma_n \ln \eta}{2\sqrt{N}} + \frac{\sqrt{N}}{\sigma_n}\right) \qquad \text{by } d^2 = \frac{4N}{\sigma_n^2} \\ &= Q\left(-\frac{\ln \eta}{d} + \frac{d}{2}\right) \\ P(H_1|H_1) &= 1 - P(H_0|H_1) \end{split}$$

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贝叶斯准则例题 4: 平均代价

$$H_0: x_k = 1 + n_k H_1: x_k = -1 + n_k$$
 $l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k \underset{H_1}{\overset{H_0}{\geqslant}} - \frac{\sigma_n^2 \ln \eta}{2N} = -\frac{\ln 3}{4N} \stackrel{def}{=} \gamma$ 统计量: $l(\mathbf{x}) \stackrel{def}{=} \frac{1}{N} \sum_{k=1}^N x_k$

判决概率: (式中, 信噪比 $d^2 = \frac{4N}{\sigma_a^2}$)

$$\begin{split} P(H_1|H_0) &= 1 - Q\left(-\frac{\ln \eta}{d} - \frac{d}{2}\right), \qquad P(H_0|H_0) = 1 - P(H_1|H_0) \\ P(H_0|H_1) &= Q\left(\frac{-\ln \eta}{d} + \frac{d}{2}\right), \qquad P(H_1|H_1) = 1 - P(H_0|H_1) \end{split}$$

两种假设先验等概 \Longrightarrow $P(H_0) = P(H_1) = \frac{1}{2}$

代价因子为 $c_{00} = 1$, $c_{10} = 4$, $c_{11} = 2$, $c_{01} = 3$

平均代价:

$$C = P(H_0)c_{00}P(H_0|H_0) + c_{10}P(H_1|H_0) + P(H_1)c_{01}P(H_0|H_1) + c_{11}P(H_1|H_1)$$

四个判决概率与信噪比 d 有关,只需要设计信噪比,可得到所需性能。

欢迎批评指正!