

Thesis Defence

Simon Fraser University



Parameter Estimation and Uncertainty Quantification Applied to
Advection-Diffusion Problems Arising in Atmospheric Source
Inversion

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Content

1 Overview

2 Setting up the Problem

- Finding a Surrogate for F
- Optimizing the Surrogate Interpolation Capabilities
- Reducing the Complexity of the Model

3 Introducing the Bayesian Framework

4 Decoding the Posterior

Problem of Interest

Given a function

$$F : A \times \Theta \rightarrow \mathbb{R},$$

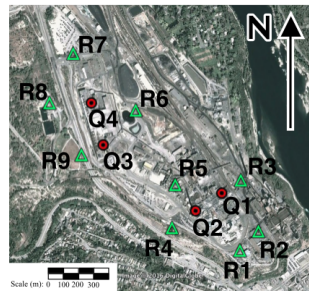
where $A \subset \mathbb{R}^n$ and Θ is a set of parameters. Given experimental (noisy) measures of F at known points $\mathbf{x}_1, \dots, \mathbf{x}_n \in A$. How to infer the values of the parameters in Θ and their uncertainties when F is computationally expensive?

Case Study

Consider the model of pollutant transport for the concentration c of a pollutant

$$\partial_t c + L(\theta)c = f$$

Goal: Estimate Q_i and θ using measurements of deposition in R_j .



Mathematical Model

$$\partial_t c(\mathbf{x}, t) + \nabla \cdot (\mathbf{u}(\mathbf{x}, t) c + \mathbf{S}(\mathbf{x}, t) \nabla c) = q(\mathbf{x}, t) \quad \text{on } \mathbb{R}^2 \times \mathbb{R}_{\geq 0} \times (0, T).$$

- $\mathbf{u}(\mathbf{x}, t) = (u_x(z, t), u_y(z, t), u_{set})$ (Wind velocity field)
- $\|(u_x, u_y)\|_2 \propto \left(\frac{z}{z_r}\right)^\gamma$,
- $\mathbf{S} = \text{diag}(s_x, s_y, s_z)$ (Eddy diffusion matrix),
- $s_z = f(L, z_{cut})$,
- $s_x = s_y = g(z_i, L)$, with z_i Mixing layer height.

Boundary Conditions

- Far-field boundary condition

$$c(\mathbf{x}, t) \rightarrow 0 \text{ as } \|\mathbf{x}\| \rightarrow \infty$$

- Robin boundary conditions at $z = 0$

$$\left(u_{\text{set}} c + s_z \frac{\partial c}{\partial z} \right) |_{z=0} = u_{\text{dep}} c |_{z=0}$$

- Concentration and deposition are related via

$$w(x, y, T) = \int_0^T c(x, y, 0, t) u_{set} dt. \quad (1)$$

$$F(x_i, y_i, T) = \int_{R_i} w(x, y, T) dx dy \approx w(x_i, y_i, T) \Delta A,$$

- ΔA is the cross-sectional area of the dust-fall jar
- T is taken to be one month.

- Numerical solution of the concentration c was obtained via a finite volume solver¹ using a 30x30 resolution grid on the domain.
- Solving for one set of parameters can take up to half an hour.

Conclusion: Finding the deposition F is computationally very expensive.

¹Hosseini, Bamdad and Stockie, J.M. Estimating airborne particulate emissions using a finite-volume forward solver coupled with a Bayesian inversion approach.

Roadmap

- 1 Find a surrogate for F ,
- 2 Locate the optimal points to evaluate F so that the surrogate is accurate,
- 3 See if it is possible to do dimensionality reduction,
- 4 Use the Bayesian framework to obtain the posterior distribution of the parameters in the light of the experimental data,
- 5 Perform inference on the posterior using numerical methods.

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Gaussian Process

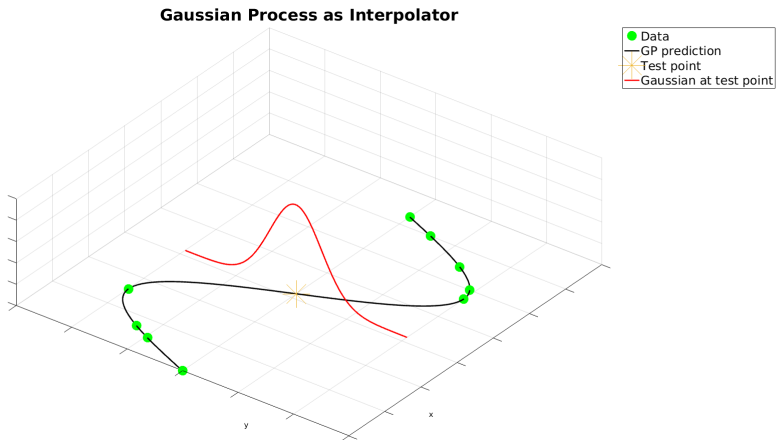
A Gaussian process (GP) is a collection of random variables $\{g(x)\}_{x \in A}$, for some set A , possibly uncountable, such that any finite subset of random variables $\{g(x_k)\}_{k=1}^N \subset \{g(x)\}_{x \in A}$ for $\{x_k\}_{k=1}^N \subset A$ are jointly Gaussian.

A GP is completely defined by its mean $m(x)$ and covariance operator $k(x, x')$:

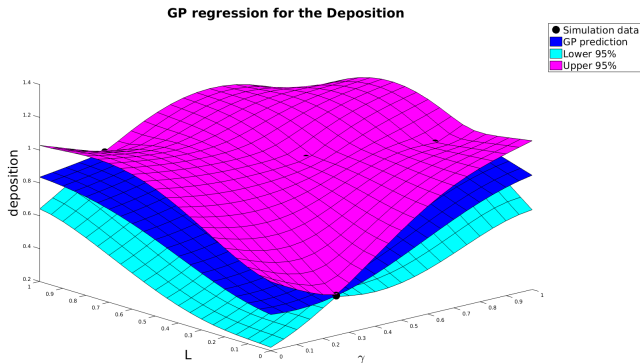
$$m(x) = \mathbb{E}(g(x)),$$

$$k(x, x') = \mathbb{E}((x - m(x))(x' - m(x'))).$$

Gaussian Process as Interpolator

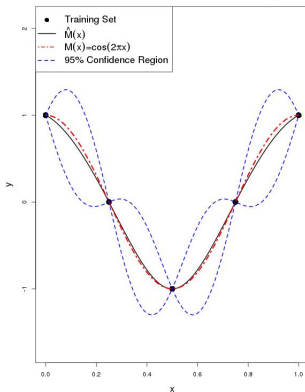


Interpolation with Real Data

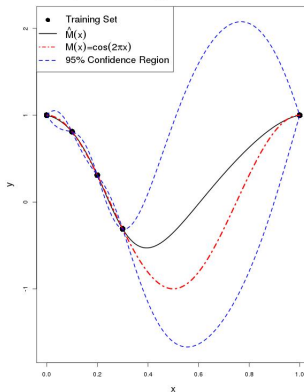


Why Experimental Design

Interpolation Using a Maximin Design



Interpolation Using an Arbitrary Partition

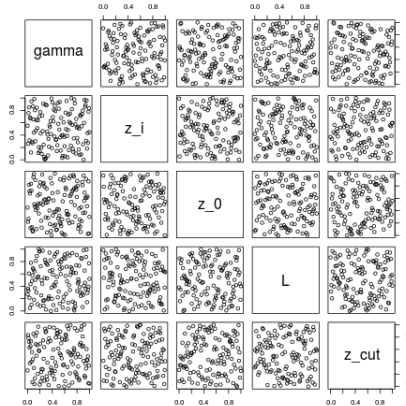


Maximin Design

Given $T \subset \mathbb{R}^n$ and a subset S of T , with finite (fixed) cardinality, say $|S| = n$. A maximin distance design S^o is a collection of points of T such that

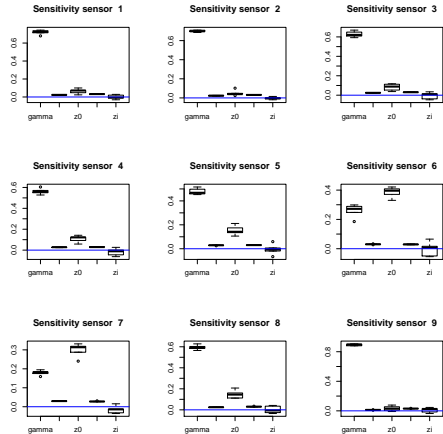
$$\max_{(S \subset T, |S|=n)} \min_{(s, s' \in S)} \|s - s'\| = \min_{s, s' \in S^o} \|s - s'\| = \max!,$$

Example of a Design



Sensitivity Analysis

- γ : Fitting parameter for the z dependence of the velocity.
- z_0 : Roughness length.
- z_i : Mixing layer height.
- L : Monin-Obukhov length.
- z_{cut} : cutoff height.



What About the Sources?

The deposition behaves linearly with respect to the values of the 4 sources, hence

$$\begin{bmatrix} R_1 \\ R_2 \\ \vdots \\ R_9 \end{bmatrix} = \mathcal{A}(\gamma, z_0, L) \begin{bmatrix} q_1 \\ \vdots \\ q_4 \end{bmatrix}$$

To evaluate the matrix $\mathcal{A}(\gamma, z_0, L)$ we approximate its entries by Gaussian process and create a surrogate $A(\gamma, z_0, L)$.

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Bayes' Rule

Given a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and two events $A, B \in \mathcal{F}$, with $\mathbb{P}(B) \neq 0$, we define the conditional probability of A given B by

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}.$$

With the definitions above, we are now in a position to state Bayes' formula as

$$\mathbb{P}_{post}(A|B) \propto \mathbb{P}_{like}(B|A)\mathbb{P}_{prior}(A). \quad (2)$$

Looking at the Stochastic Model

To account for the uncertainties in the model and in the experimental measurements we propose the model

$$\begin{bmatrix} R_1 \\ R_2 \\ \vdots \\ R_9 \end{bmatrix} = A(\gamma, z_0, L) \begin{bmatrix} q_1 \\ \vdots \\ q_4 \end{bmatrix} + \epsilon, \quad \text{where } \epsilon \sim \mathcal{N}(0, \lambda_\epsilon I_{9 \times 9})$$

Probabilistic Model

Our goal is to estimate the values of $\omega := (\gamma, z_0, L)$ and $q := (q_1, q_2, q_3, q_4)$ given the measurements \vec{R} . Mathematically we want to estimate:

$$\mathbb{P}_{post}(\omega, q | \vec{R}) \propto \underbrace{\mathbb{P}_{like}(\vec{R} | \omega, q) \mathbb{P}_{prior}(\omega) \mathbb{P}_{prior}(q)}_{\text{Assuming } p \text{ and } q \text{ independent}}$$

We assume $\omega \sim \text{Uniform}$ over the domain of definition of the parameters.

What About q ?

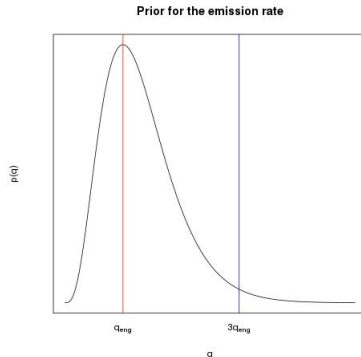
- $q_k > 0$ for $k = 1, 2, 3, 4$.
- If we trust the engineers the most likely value for q is the engineers estimate and the true value cannot be very far away from those estimates.

Source	Estimated Emission Rate [ton/yr]
q_1	35
q_2	80
q_3	5
q_4	5

Choosing a Prior for q

A consistent assumption is $q_k \sim Ga(\alpha_k, \beta_k)$, for $k = 1, 2, 3, 4$, the following conditions that define α_k and β_k for all k uniquely.

- $\beta_k(\alpha_k - 1) = q_{eng,k}$
- $qgamma(0.99, \alpha_k, \beta_k) = 3q_{eng,k}$



Likelihood

Recall that $\epsilon \sim \mathcal{N}(0, \lambda_\epsilon I_{9 \times 9})$ then

- $\mathbb{P}(\epsilon) \propto \exp\left(-\frac{\|\epsilon\|_2^2}{2\lambda_\epsilon^2}\right)$
- $\mathbb{P}_{like}(\vec{R}|\omega, q) \propto \exp\left(-\frac{1}{2\lambda_\epsilon^2}\|\vec{R} - A(\omega)q\|_2^2\right)$

Calculating the Posterior

- $\mathbb{P}_{like}(\vec{R}|\omega, q) \propto \exp\left(-\frac{1}{2\lambda_\epsilon^2}\|\vec{R} - A(\omega)q\|_2^2\right)$
- $\mathbb{P}_{prior}(q_k) \propto q_k^{\alpha_k-1} \exp(\beta_k q_k)$
- $\mathbb{P}_{prior}(\omega) \propto \mathbf{1}_{[0.1, 0.6] \times [0, 2] \times [-600, 0]}$

Then

$$\mathbb{P}_{post}(\omega, q|\vec{R}) \propto \mathbb{P}_{like}(\vec{R}|\omega, q)\mathbb{P}_{prior}(\omega)\mathbb{P}_{prior}(q)$$

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Sampling From a Probability Distribution

Algorithm 1 Metropolis-Hastings Algorithm

```
1: pick a point  $q_1$  in the support of the distribution
2: for  $j=2:N$  do
3:   Draw  $u \sim U([0, \alpha])$ 
4:    $q_j \leftarrow q_{j-1} + u$ 
5:    $\beta \leftarrow \min(1, \frac{\mathbb{P}_{post}(q_j|D)}{\mathbb{P}_{post}(q_{j-1}|\mathbf{y})})$ 
6:   Draw  $w \sim U([0, 1])$ 
7:   if  $w < \beta$  then
8:      $q_{j-1} = q_j$       (Accept the move)
9:   else
10:     $q_{j-1} = q_{j-1}$     (Reject the move)
11:  end if
12: end for
```

Setting λ_ϵ

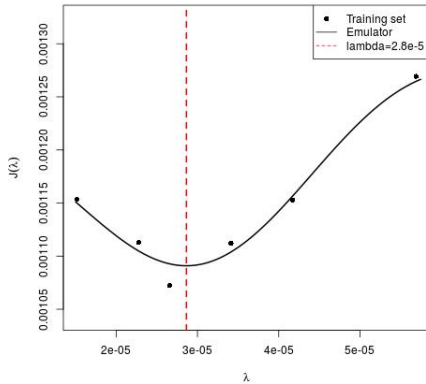
We define

$$J(\lambda_\epsilon) = \frac{1}{2} \int \left(\|A(p)q - \vec{R}\|_2 + \|q - q_{est}\|_2 \right) d\mathbb{P}_{post}^{\lambda_\epsilon},$$

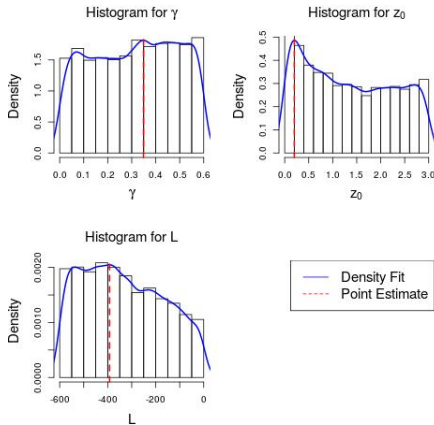
and choose a minimizer of J

$$\hat{\lambda}_\epsilon = \operatorname{argmin} J(\lambda_\epsilon).$$

Calculating J

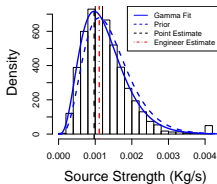


Histograms for the Parameters

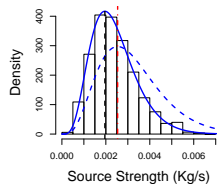


Histograms for the Sources

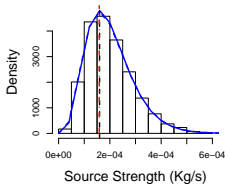
Histogram for Source 1



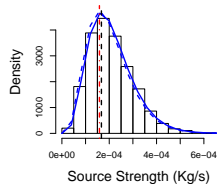
Histogram for Source 2



Histogram for Source 3



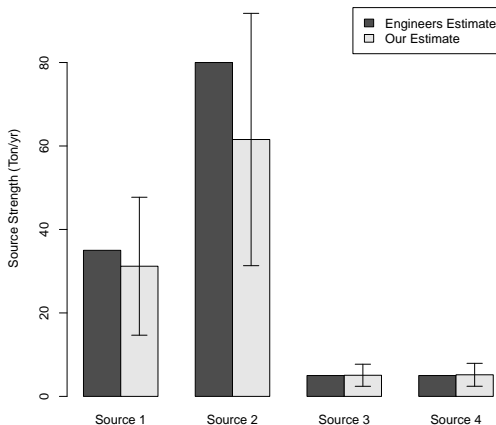
Histogram for Source 4



Results for the Parameters

Parameter	Point Estimate	68% Confidence Interval
γ	0.3478	[0.1498, 0.5458]
z_0	0.0811	[0, 1.5781]
L	-379.45	[-195.86, -563.04]

Results for the Sources



Comparison with Related Work

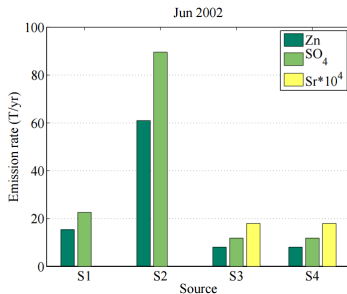


Figure: Hosseini, Bamdad and Stockie, J.M. Estimating airborne particulate emissions using a finite-volume forward solver coupled with a Bayesian inversion approach.

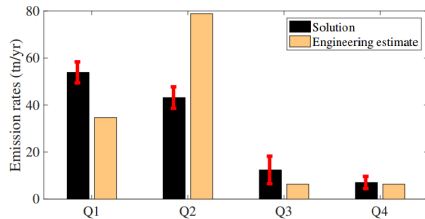


Figure: Lushi, E. and Stockie, J.M. An inverse Gaussian plume approach for estimating atmospheric pollutant emissions from multiple point sources

Conclusions

- We have developed a method to cheaply estimate parameters in computationally expensive models.
- Instead of trial and error, we propose a methodology that allows to estimate parameters in complex models using experimental data.
- Besides a point estimate we are able to obtain a confidence interval for it.
- Our results qualitatively agree with previous results.