

MATH 418/544 Assignment 3

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Problem 1(a).

Proof. $\{(-\infty, x] | x \in \mathbb{R}^2\}$ is a π -system since for arbitrary set $A = (-\infty, a]$ and $B = (-\infty, b]$ where $a = (a_1, a_2), b = (b_1, b_2)$, we have $A \cap B = (-\infty, c] \in \{(-\infty, x] | x \in \mathbb{R}^2\}$ for some $c = (\min(a_1, b_1), \min(a_2, b_2))$. We will next show that $\int 1_B(y)f(y)dy$ is a probability measure. Because $1_B(y)$ and $f(y)$ are both non-negative functions by definition, we have $\int 1_B(y)f(y)dy \geq 0$ for all sets in $(\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$, and $\int 1_\emptyset(y)f(y)dy = 0$. To show countable additivity, without loss of generality, let $\{B_i\}_{i=1}^\infty$ be a collection of disjoint set in $(\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$, denote $B = \bigcup_{i=1}^\infty B_i$, the disjoint union of B_i . Then by the properties of indicator functions we have

$$\sum_{i=1}^\infty \int 1_{B_i}(y)f(y)dy = \int 1_{\bigcup_{i=1}^\infty B_i}(y)f(y)dy = \int 1_B(y)f(y)dy$$

Therefore we have shown the countable additivity. Also, we have $\int 1_{\mathbb{R}^2}(y)f(y)dy = \int f(y)dy = 1$. Hence we've proven that $\int 1_B(y)f(y)dy$ is a probability measure defined on $(\mathbb{R}^2, \mathcal{B}(\mathbb{R}^2))$. Thus we know $\int 1_B(y)f(y)dy$ and P are probabilities that agree on π -system $\{(-\infty, x] | x \in \mathbb{R}^2\}$ where $\{(-\infty, x] | x \in \mathbb{R}^2\} \subset \mathcal{B}(\mathbb{R}^2)$, and by *proposition 1.16* in class we know $\mathcal{B}(\mathbb{R}^2) = \sigma(\mathcal{S}_2^{-\infty})$, $\mathcal{S}_2^{-\infty} = \{(-\infty, b] | b \in \mathbb{R}^2\}$. According to the *corollary 1.6* of the π - λ theorem, we know that $\int 1_B(y)f(y)dy$ and P agrees on all Borel sets B in $\mathcal{B}(\mathbb{R}^2)$, in other words, $P(X \in B) = \int 1_B(y)f(y)dy$. \square

Problem 1(b).

Proof. The distribution function of Y is

$$\begin{aligned} F_Y(z) &= P(Y \leq z) = P(X_1 + X_2 \leq z) \\ &= \int_{\mathbb{R}^2} 1_{\{x_1+x_2 \leq z\}} f(x_1, x_2) dx_1 dx_2 \\ &= \int_{\mathbb{R}} dx_2 \int_{-\infty}^{z-x_1} \frac{1}{4} 1_{[0,2]}(x_1) 1_{[0,2]}(x_2) dx_1 \\ &= \frac{1}{4} \int_0^2 |(-\infty, z-x_2) \cap [0, 2]| dx_2 \end{aligned} \tag{1}$$

where $|(-\infty, z-x_2) \cap [0, 2]|$ is the length of the interval. Then we have the distribution function for Y :

$$F_Y(z) = \begin{cases} 0 & \text{if } z < 0 \\ \frac{1}{8}z^2 & \text{if } 0 \leq z \leq 2 \\ -\frac{1}{8}z^2 + z - 1 & \text{if } 2 < z \leq 4 \\ 1 & \text{if } z > 4 \end{cases}$$

By differentiating the distribution function, we can get the probability density function of Y :

$$f_Y(z) = \begin{cases} \frac{1}{4}z & \text{if } 0 \leq z \leq 2 \\ 1 - \frac{1}{4}z & \text{if } 2 < z \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

\square

Problem 2.

Proof. Since $Q \sim U(-\frac{\pi}{2}, \frac{\pi}{2})$, then we have

$$f_Q(q) = \begin{cases} \frac{1}{\pi} & \text{if } q \in (-\frac{\pi}{2}, \frac{\pi}{2}) \\ 0 & \text{otherwise} \end{cases}$$

The distribution function for X : $F_X(x) = P(X \leq x) = P(\tan \theta \leq x) = P(\theta \leq \arctan x)$. So we have $F_X(x) = \frac{\arctan x}{\pi} + \frac{1}{2}$ for $x \in \mathbb{R}$. $F'_X(x) = \frac{1}{\pi x^2 + \pi}$ is continuous and by fundamental theorem of calculus, $F_X(x) = \int_{-\infty}^x \frac{1}{\pi y^2 + \pi} dy$. Thus X has the probability density function $f_X(x) = F'_X(x) = \frac{1}{\pi x^2 + \pi}$ for $x \in \mathbb{R}$ \square

Problem 3.

Proof. Need to show $X_{N(\omega)}(\omega)$ is a random variable, then by definition, it is equivalent to show for all set $E \in \mathcal{B}(\mathbb{R})$, $\{\omega : X_{N(\omega)}(\omega) \in E\} \subset \mathcal{F}$. We can rewrite $\{\omega : X_{N(\omega)}(\omega) \in E\}$ as

$$\bigcup_{n=1}^{\infty} \left\{ \omega : (N(\omega) = n) \cap (X_n(\omega) \in E) \right\}$$

Since N is an \mathbb{N} -valued random variable and X_n is random variable for all $n \in \mathbb{N}$, we know that $\{\omega : N(\omega) = n\}$ and $\{\omega : X_n(\omega) \in E\}$ are all in \mathcal{F} . It is also known that \mathcal{F} is a σ -field, which is closed under intersection and countable unions. Therefore we have shown that $\bigcup_{n=1}^{\infty} \left\{ \omega : (N(\omega) = n) \cap (X_n(\omega) \in E) \right\} \subset \mathcal{F}$, hence we have proven that $X_{N(\omega)}(\omega)$ is a random variable. \square

Problem 4(a).

Proof. Since $X_n(\omega) = \omega_n \in \{0, 1\}$, $X(\omega) = \sum_{n=1}^{\infty} X_n(\omega) 2^{-n} \in \left[\sum_{n=1}^{\infty} 0 \times 2^{-n}, \sum_{n=1}^{\infty} 1 \times 2^{-n} \right] = [0, 1]$. Therefore the series converges for every $\omega \in \Omega$. For each single $X_n(\omega)$, it is Borel measurable. Then by *theorem 1.3.4* from textbook, the finite sum $\sum_{i=1}^N X_i(\omega)$ is also Borel measurable, denote the sum by g_N . Since the $\sum_{n=1}^{\infty} X_n(\omega) 2^{-n}$ converges, we have $\lim_{N \rightarrow \infty} g_N = \limsup_{N \rightarrow \infty} g_N$, where $\limsup_{N \rightarrow \infty} g_N$ is also Borel measurable by *theorem 1.3.5* from textbook. Therefore the infinite sum $X(\omega) = \sum_{n=1}^{\infty} X_n(\omega) 2^{-n} = \lim_{N \rightarrow \infty} g_N$ is Borel measurable, in other words, we have shown that $X(\omega)$ is a random variable taking values in $[0, 1]$ \square

Problem 4(b).

Proof. A number $x \in \mathbb{R}$ is dyadic rational number if x has a finite binary expansion, in other words, there exists an integer $m \geq 1$ and $\omega_1, \omega_2, \dots, \omega_m$ such that $x = \sum_{n=1}^m \omega_n 2^{-n}$. Then we can easily show the set of dyadic rational numbers is dense in $\mathbb{R} \cap [0, 1]$ (For every $x \in \mathbb{R} \cap [0, 1]$ and every $\epsilon > 0$, we can find a dyadic rational number d such that $|d - x| \leq \epsilon$). Also notice that every number $x \in \mathbb{R} \cap [0, 1]$ either (i) has a unique, non-terminating binary expansion, or (ii) it has two binary expansions, one ending in an infinite sequence of 0s and the other ending in an infinite sequence of 1s, in either case of (ii) we have $P(X = x) = \lim_{n \rightarrow \infty} (\frac{1}{2})^n = 0$. Take an example of dyadic rational numbers $\frac{0}{2}, \frac{1}{2}, \frac{2}{2}$, the case where $m = 1, i = 0, 1$:

$$X \in [\frac{0}{2}, \frac{1}{2}] = \{X_1 = 0\} \cup \{X_1 = 1, X_2 = X_3 = \dots = 0\}, X \in [\frac{1}{2}, \frac{2}{2}] = \{X_1 = 1\} \cup \{X_1 = 0, X_2 = X_3 = \dots = 1\}$$

$$\begin{aligned} P(X \in [\frac{0}{2}, \frac{1}{2}]) &= P(\{X_1 = 0\}) + P(\{X_1 = 1, X_2 = X_3 = \dots = 0\}) \\ &= \frac{1}{2} + \lim_{n \rightarrow \infty} (\frac{1}{2})^n = \frac{1}{2} \end{aligned} \tag{2}$$

Similarly $P(X \in [\frac{1}{2}, \frac{2}{2}]) = \frac{1}{2}$. Generalizing this result by induction, assume this holds for $m = k$, such that

$$\begin{aligned} P(X \in [i2^{-k}, (i+1)2^{-k}]) &= P(\{X_1 = w_1, X_2 = w_2, \dots, X_k = w_k\}) \\ &\quad + P(\{X_1 = w'_1, X_2 = w'_2, \dots, X_k = w'_k, X_{k+1} = \dots = w\}) \\ &= 2^{-k}, \forall i = 0, \dots, 2^m - 1 \end{aligned} \quad (3)$$

Where the second part include the cases where the sequence ends with either a infinite sequence of 0s or an infinite sequence of 1s and has probability 0. Then for $m = k + 1$, we have

$$\begin{aligned} P(X \in [i2^{-k+1}, (i+1)2^{-k+1}]) &= P(\{X_1 = 0, X_2 = w_1, \dots, X_k = w_{k-1}, X_{k+1} = w_k\}) \\ &\quad + P(\{X_1 = 0, X_2 = w'_1, \dots, X_k = w'_{k-1}, X_{k+1} = w'_k, X_{k+2} = \dots = w\}) \\ &= \frac{1}{2} \times 2^{-k} + 0 = 2^{-k+1}, \forall i = 0, \dots, 2^m - 1 \end{aligned} \quad (4)$$

Where now the whole binary sequence is shifted to the right by 1 position (think of this as dividing the binary number by 2), where numbers fall between $[i2^{-k}, (i+1)2^{-k}]$ will fall between $[i2^{-k+1}, (i+1)2^{-k+1}]$ now and the rest is same as above. Then for dyadic rational numbers $i2^{-m}$ and $(i+1)2^{-m}$:

$$P(X \in [i2^{-m}, (i+1)2^{-m}]) = 2^{-m}$$

for every $m \geq 0$ and $i = 0, \dots, 2^m - 1$. Hence for any dyadic rational numbers a, b in $[0, 1]$ where $a \leq b$, we have $P(X \in [a, b]) = b - a$ since a, b can be written as disjoint union of intervals of dyadic rational numbers. Recall that dyadic rational numbers are also dense in $\mathbb{R} \cap [0, 1]$. Thus, given any two numbers a, b in $[0, 1]$ where $a \leq b$ (not necessarily dyadic), we can find a decreasing sequence $\{a_j\}$ and an increasing sequence $\{b_j\}$ of dyadic rational numbers that converges to a and b respectively. Then we have

$$P(X \in (a, b)) = P(X \in \bigcup_{j=1}^{\infty} [a_j, b_j]) = \lim_{j \rightarrow \infty} P(X \in [a_j, b_j]) = \lim_{j \rightarrow \infty} (b_j - a_j) = b - a$$

by continuity from below (Note that endpoints does not matters, since $P(X = 1) = P(X = 0) = 0$). Therefore by definition we have proven that X has a uniform distribution on $[0, 1]$. \square