

PS1, 1, 2, 3, 4, 5

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$$F_X(x) = \mathbb{P}(X \leq x), \quad A_x = \{\omega \in \Omega : X(\omega) \leq x\}$$

$$1. \mathbb{P}(X > x) = 1 - F_X(x) \quad | \quad A_x^c = \{\omega \dots : X(\omega) > x\}$$

$$2. \mathbb{P}(X = x) = F_X(x) - F_X(x^-) \quad | \quad \lim_{z \uparrow x} F_X(z)$$

$$3. \mathbb{P}(X < x) = F_X(x^-) \quad \mathbb{P}(X = x) = \mathbb{P}(\{\omega \in \Omega : X(\omega) = x\})$$
$$\mathbb{P}(X < x) = \mathbb{P}(\{\omega \in \Omega : X(\omega) < x\})$$

$$4. \mathbb{P}(a < X \leq b) = F_X(b) - F_X(a) \quad | \quad \{\omega \dots : a < X(\omega) \leq b\} =$$
$$\{\omega \dots : X(\omega) > a\} \cap \{\omega \dots : X(\omega) \leq b\}$$

## Discrete RV

$X$  is discrete if it takes countably many values  $\underbrace{\{x_1, x_2, \dots\}}_S$

• Probability Mass Function (PMF)

$$\begin{aligned} f_X(x) &= \mathbb{P}(X=x) \\ &= F_X(x) - F_X(x^-) \end{aligned}$$

$$\left| \begin{array}{l} f_X(x) \in [0, 1] \quad \forall x \\ \sum_x f_X(x) = 1 \end{array} \right.$$

$x \in S, 0$  otherwise

We can recover the CDF:  $F_X(x) = \sum_{z: z \leq x} f_X(z)$

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## Continuous RVs

$X$  is cts if its CDF can be written as

$$F_X(x) = \int_{-\infty}^x f_X(z) dz$$

where  $f_X(x)$  is an integrable function.  $f_X(x)$  is the probability density function (PDF)

$$\begin{aligned} f_X(x) &= \frac{d}{dx} F_X(x) = \lim_{h \rightarrow 0} \frac{F_X(x+h) - F_X(x)}{h} \\ &= \lim_{h \rightarrow 0} \frac{\mathbb{P}(x < X \leq x+h)}{h}, \quad \mathbb{P}(X=x) = 0 \end{aligned}$$

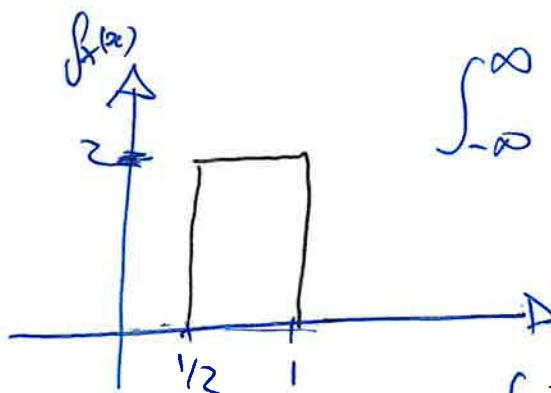
PDF

$$f_X(x) \geq 0$$

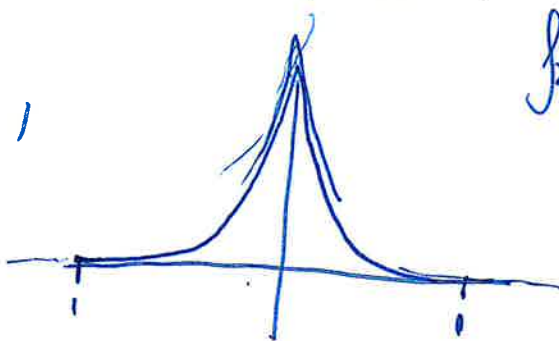
$$f_X(x) \leq 1$$

$$\int_{-\infty}^{\infty} f_X(x) dx = 1$$

PDF is not bounded by 1



$$f_X(x) = \begin{cases} 2 & x \in [1/2, 1] \\ 0 & \text{otherwise} \end{cases}$$



$$\int_{-\infty}^{\infty} = \int_{1/2}^1$$

## Expectation

Example:  $X$  takes  $n$  different values all with probability  $\frac{1}{n}$   
 $\{x_1, \dots, x_n\}$

What do you expect of  $X$ ?

$$\frac{1}{n} \sum_{i=1}^n x_i$$

$X$  is discrete

$X$  is continuous

$$\mathbb{E}[X] = \sum_{x \in S} x f_X(x), \quad \mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

$$\mathbb{E}[X] \text{ I assume } \mathbb{E}[|X|] < \infty$$

# Law of the Unconscious Statistician (LOTUS)

## Unconscious

Suppose we have  $g: \mathbb{R} \rightarrow \mathbb{R}$  that is "nice"

Given  $X$   
Then  $g(X)$  is going to be a RV

$$\mathbb{E}[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx \quad \text{if continuous}$$

$$\mathbb{E}[g(X)] = \sum_{x \in S} g(x) f_X(x)$$

$$\int a f(x) dx = a \int f(x) dx$$

Assuming  $\mathbb{E}[|g(X)|] < \infty$

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$$\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2 - 2X\mathbb{E}[X] + \mathbb{E}[X]^2]$$

$$= \mathbb{E}[X^2] - \mathbb{E}[2X\mathbb{E}[X]] + \mathbb{E}[\mathbb{E}[X]^2]$$

$$= \mathbb{E}[X^2] - 2\mathbb{E}[X]^2 + \mathbb{E}[X]^2$$

$$= \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

$$\int_{-\infty}^{\infty} x^2 dF_X(x) = \begin{cases} \int_{-\infty}^{\infty} x^2 f_X(x) dx \\ \sum x^2 f_X(x) \end{cases}$$



$$k\text{-th moment: } \int_{-\infty}^{\infty} x^k dF_X(x)$$

$$k\text{-th central moment: } \int_{-\infty}^{\infty} (x - E[X])^k dF_X(x)$$

~~Jensen's~~

Jensen's inequality

Suppose  $g: \mathbb{R} \rightarrow \mathbb{R}$  is a convex function,  $X$  is a RV

$$\text{Then } E[g(X)] \geq g(E[X])$$



Proof: Since  $g$  is convex you can find  $m$  :  $g(x) \geq g(x_0) + m(x - x_0)$   
 $x_0 = E[X]$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx \geq \int_{-\infty}^{\infty} [g(E[X]) + m(x - E[X])] f_X(x) dx$$

$$= g(E[X]) \underbrace{\int_{-\infty}^{\infty} f_X(x) dx}_1 + m \underbrace{\int_{-\infty}^{\infty} (x - E[X]) f_X(x) dx}_{E[X] - E[X] \int_{-\infty}^{\infty} f_X(x) dx}$$

$$= g(E[X]) \quad \left| \quad g(x) = x^2, \text{ then } E[X^2] \geq E[X]^2 \Rightarrow \text{Var}(X) \geq 0 \right.$$

## Moment Generating Function (MGF)

$X$  the MGF is defined as

$$M_X(t) = \mathbb{E}[e^{tx}]$$

if it exists for ~~some~~  $t \in (-h, h)$ ,  $h > 0$

$$\left. \frac{d^k}{dt^k} M_X(t) \right|_{t=0} = \mathbb{E}[X^k]$$

• MGF uniquely determines  $X$ , so if  $X_1, X_2$  are RV with

$$M_{X_1}(t) = M_{X_2}(t) \quad \forall t$$

$$\cancel{X_1 \neq X_2} \Rightarrow X_1 = X_2$$

$$\frac{d^k}{dt^k} M_X(t) \Big|_{t=0} = \mathbb{E}[X^k]$$

$$\mathbb{E}[e^{tx}] \quad , \quad \boxed{e^x = \sum_{j=0}^{\infty} \frac{x^j}{j!}}$$

$$\begin{aligned} \mathbb{E}[e^{tx}] &= \mathbb{E}\left[\sum_{j=0}^{\infty} \frac{(tx)^j}{j!}\right] & \mathbb{E}[X^0] &= \int_{-\infty}^{\infty} 1 dF_X(x) = 1 \\ &= \sum_{j=0}^{\infty} \frac{t^j}{j!} \mathbb{E}[X^j] \end{aligned}$$

$$\frac{d}{dt} M_X(t) \Big|_{t=0} = \mathbb{E}[X] + \underbrace{\sum_{j=2}^{\infty} \frac{t^{j-1}}{(j-1)!} \mathbb{E}[X^j]}_0 \Big|_{t=0}$$

$$\begin{aligned} \frac{d^k}{dt^k} M_X(t) \Big|_{t=0} &= \underbrace{\frac{k(k-1)\dots 1}{k!}}_1 \mathbb{E}[X^k] + \underbrace{\sum_{j=k+1}^{\infty} \frac{t^{j-k}}{(j-k)!} \mathbb{E}[X^j]}_0 \Big|_{t=0} \\ &= \mathbb{E}[X^k] \end{aligned}$$

