## STAT 656: Bayesian Data Analysis Fall 2024 Homework 2

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```
library("bayesplot")
library("ggplot2")
library("rstan")
options(repr.plot.width = 6, repr.plot.height = 4)
bayesplot_theme_set(theme_default(base_size = 24, base_family = "sans"))
```

## Synthetic Data

The file hw2\_synthetic.csv is a dataset of count-valued measurements  $\mathbf{y} = \{y_1, \dots, y_n\}$ , with  $y_i \in 0, 1, \dots$ Each output  $y_i$  has an associated  $x_i = (x_{i,1}, x_{i,2}) \in \mathbb{R}^2$ , and write  $\mathbf{x} = \{x_1, \dots, x_n\}$ 's as  $\mathbf{x}$ . We model  $y_i$  as

$$y_i|\beta \sim \text{Poisson}(e^{f(x_i,\beta)}).$$

Here, the exponential is to ensure the Poisson rate is always positive, and the function  $f(x_i, \beta) = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \beta_3 x_{i,1}^2 + \beta_4 x_{i,2}^2 + \beta_5 x_{i,1} x_{i,2}$ .

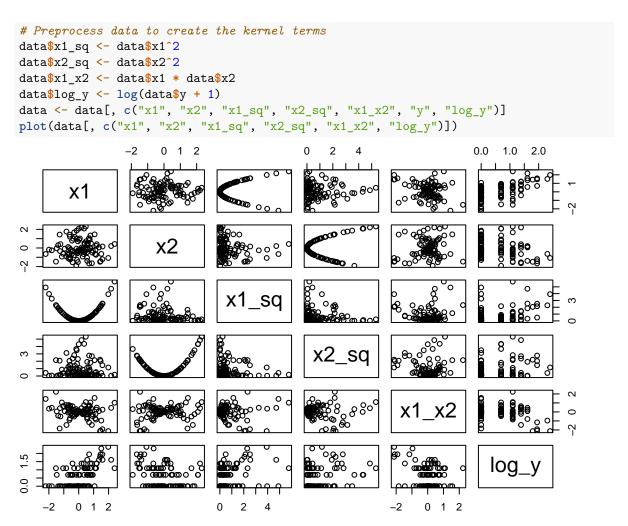
1. (25 points) Solution:

```
# Read the data
if (file.exists("data/hw2_synthetic.csv")) {
   data <- read.csv("data/hw2_synthetic.csv", header = TRUE)
} else {
   stop("FileNotFound: data file not found at 'data/hw2_synthetic.csv'.")
}
summary(data)</pre>
```

```
##
                             x2
          x1
                                                  : 0.00
   Min.
           :-2.2147
                              :-1.91436
                       Min.
                                           1st Qu.: 0.00
   1st Qu.:-0.4942
                       1st Qu.:-0.65105
   Median : 0.1139
                       Median :-0.17722
                                           Median: 1.00
    Mean
           : 0.1089
                              :-0.03781
                                                   : 1.29
                       Mean
                                           Mean
    3rd Qu.: 0.6915
                       3rd Qu.: 0.50090
                                           3rd Qu.: 2.00
           : 2.4016
                              : 2.30798
                                                   :10.00
   Max.
                       Max.
                                           Max.
```

Since the model is a Poisson regression, the logarithm of the response variable is assumed to be a linear combination of the kernel features. The visualization below shows the relationship between the logarithm of the response variable and the kernel features. It is shown that  $\log y$  is roughly positively correlated with  $x_1$  and  $x_1x_2$ , and roughly negatively correlated with  $x_2$ ,  $x_1^2$ , and  $x_2^2$ .

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Therefore, the weights  $\beta_i$  can be both positive and negative. Given no prior knowledge on the features, I choose a non-informative prior for the weights, *i.e.*, the isotropic standard normal distribution:  $\mathcal{N}(0, \mathbf{I})$ . The Stan model is implemented as follows:

```
linreg_poisson_code <- "</pre>
    // Input arguments to the model
    data {
        int<lower=0> n;
                                // Number of observations
        int<lower=0> k;
                                // Number of features
        real<lower=0> pr_std;
                               // Prior coefficients standard deviation
        matrix[n, k] x;
                                // Observation matrix
        int<lower=0> y[n];
                                            // Integer response vector
    }
    // Latent parameters of interests
    parameters {
        vector[k] beta;
                                // Coefficients
    // Transformed parameters for MCMC sampling
   transformed parameters {
        vector[n] lambda = exp(x * beta);  // Poisson rate
```

```
// PoissoXn Linear Regression Model
    model {
        beta ~ normal(0, pr_std);  // Prior on the coefficients
                                             // Poisson emission
        y ~ poisson(lambda);
    }
    // Retrieve MCMC samples
    generated quantities {
        real y_hat[n];
        y_hat = poisson_rng(lambda);
    }
linreg_poisson_model <- stan_model(</pre>
 model_name = "poisson_regression",
 model_code = linreg_poisson_code
x <- data[, c("x1", "x2", "x1_sq", "x2_sq", "x1_x2")]
x[, "intercept"] <- 1
y <- data$y
reg_data \leftarrow list(n = nrow(x), k = ncol(x), pr_std = 1.0, x = x, y = y)
samples <- sampling(</pre>
 linreg_poisson_model,
 data = reg_data,
 iter = 10000,
 warmup = 2000,
 chains = 4,
 seed = 42,
  show_messages = FALSE,
```

The visualization below shows the posterior distributions of the coefficients  $\beta = \{\beta_0, \dots, \beta_5\}$  with 95% intervals. Based on the results, the coefficients of  $x_1^2$ , and  $x_1x_2$  are close to zero (*i.e.*, with zero within its 95% posterior interval), while the coefficients of  $x_1$ ,  $x_2$ , and  $x_2^2$  are significantly different from zero. Therefore, the terms  $x_1$ ,  $x_2$ , and  $x_2^2$  are important in the expression for f(x).

```
samples <- as.data.frame(samples)
colnames(samples)[seq_len(ncol(x))] <- colnames(x)
fig <- mcmc_areas(
    samples[, seq_len(ncol(x))],
    pars = colnames(x),
    prob = 0.95,
)
fig + scale_y_discrete(
    labels = parse(text = paste0("beta[", c(1:(ncol(x) - 1), 0), "]"))
)</pre>
```

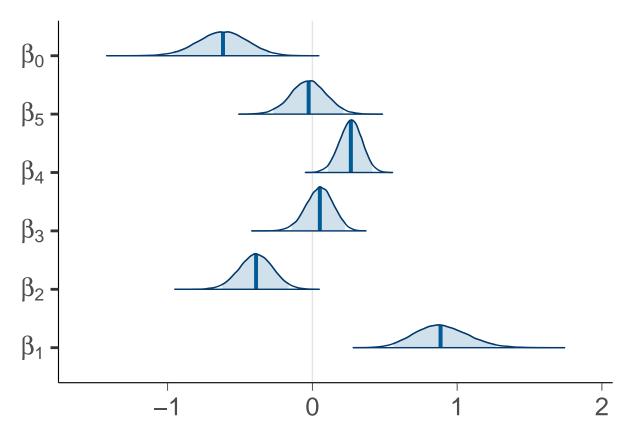


Figure 1: Posterior Distributions of the Coefficients.