Data analysis in Astronomy Homework 2 due 11/5 11:59 pm

params = {'legend.fontsize': 20,

20.0

17.5

15.0

12.5

def my_plot_style():

params = {'legend.fontsize': 20,

'axes.labelsize': 20, 'axes.titlesize':20,

'figure.facecolor':'w', #'lines.linewidth' : 1.5, 'xtick.major.width':1.5, 'ytick.major.width':1.5, 'xtick.minor.width':1.5, 'ytick.minor.width':1.5,

'text.usetex' : False, 'font.family': 'sans-serif'}

plt.rcParams.update(params)

mean_precision = np.zeros(npoints) gmean_precision = np.zeros(npoints)

for ipoint in range(0, npoints): ndata = (ipoint+1)*1000xaxis[ipoint] = ndata

for isim in range(0, nsim):

randomdata = np.random.standard_cauchy(ndata)

randomgauss = np.random.normal(0, 1, ndata)

simmean[isim] = np.mean(randomdata)

simgauss[isim] = np.mean(randomgauss)

simmean = np.zeros(nsim) simgauss = np.zeros(nsim)

xaxis = np.zeros(npoints) std = np.zeros(npoints) gstd = np.zeros(npoints)

nsim = 1000

npoints = 40

 10^{-5}

calculation

ISM = data[1].dataEBV = ISM['SFD'] HI = ISM['HI']/1e21

To do:

0

data = pf.open('sky_maps_new_64_v6.fits')

 $conversion_factor = 2*1e20/1e21$ H2 = ISM['C010']*conversion_factor

Pearson(EBV, H2) uncertainty: 0.0041 Spearman(EBV, H2) uncertainty: 0.0046

1. (Monte Carlo Exercise) The variance of the sample variance

A sample of N data points drawn from a normal distribution has a variance. This variance also has a variance.

Based on theoretical calculation, the variance of variance with N data points is —— where is the standard deviation of the normal

Name:

distribution.

To do:

1. write a code and use Monte Carlo method to validate the theoretical expection. (15 points)

2. Produce a plot with x-axis (N data point) and y-axis (variance of variance) with two curves showing a. your Monte Carlo simulation and b. theoretical calculation. (10 points) (You can just use a normal distribution with mean=0 and sigma=10) In []: import numpy as np import matplotlib.pyplot as plt def my_plot_style():

'axes.labelsize': 20, 'axes.titlesize':20, 'xtick.labelsize':16, 'ytick.labelsize':16, 'xtick.major.size':5, 'xtick.minor.size':2.5, 'ytick.major.size':5, 'ytick.minor.size':2.5, 'figure.facecolor':'w', #'lines.linewidth' : 1.5, 'xtick.major.width':1.5, 'ytick.major.width':1.5, 'xtick.minor.width':1.5, 'ytick.minor.width':1.5, 'xtick.major.pad': 12, 'ytick.major.pad': 8, 'axes.linewidth':1.5, 'xtick.direction':'in', 'ytick.direction':'in', 'ytick.labelleft':**True**, 'text.usetex' : False, 'font.family': 'sans-serif'} plt.rcParams.update(params) mean = 0sigma = 10nsim = 1000simvar = np.zeros(nsim) npoints = 100xaxis = np.zeros(npoints) var2 = np.zeros(npoints) theo_var2 = np.zeros(npoints) for ipoint in range(0, npoints): ndata = (ipoint+1)*1000xaxis[ipoint] = ndatafor isim in range(0, nsim): randomdata = np.random.normal(mean, sigma, ndata) simvar[isim] = np.var(randomdata) var2[ipoint] = np.var(simvar) theo_var2[ipoint] = 2*pow(sigma, 4)/ndata my_plot_style() plt.figure(figsize = (8, 8)) plt.scatter(xaxis, var2, s=15, c="#67ccc1", label="My simulation") plt.scatter(xaxis, theo_var2, s=15, c="#ff3360", label="Theoretical calculation") plt.legend() plt.xlabel("N data point") plt.ylabel("Var of Var") plt.show() In [3]: from IPython import display display.Image("/home/judy/Astrophys/Week5/HW2/Monte_Carlo_1.png") Out[3]:

My simulation

Theoretical calculation

10.0 7.5 5.0 2.5 0.0 60000 20000 40000 0 80000 100000 N data point 2. When the central limit theorem breaks down The central limit theorem states that given a sample with N data points drawn from some distributions with mean and standard deviation , the precision of the mean of the sample will scale with sigma/sqrt(N). In other words, if you have a larger sample, you can obtain a more precise estimation of the mean value. However, this is not always true. To do: a. Find a distribution that will break the central limit theorem and write a code using Monte Carlo method to demonstrate that. (10 points) b. Make a plot showing your result with that distribution and the expected trend based on the central limit theorem. (10 points) Cauchy distribution will break the central limit theorem since it's variance cannot be defined. In []: import numpy as np import matplotlib.pyplot as plt import math

'xtick.labelsize':16, 'ytick.labelsize':16, 'xtick.major.size':5, 'xtick.minor.size':2.5, 'ytick.major.size':5, 'ytick.minor.size':2.5,

'xtick.major.pad': 12, 'ytick.major.pad': 8, 'axes.linewidth':1.5, 'xtick.direction':'in', 'ytick.direction':'in', 'ytick.labelleft':**True**,

std[ipoint] = np.std(simmean) mean_precision[ipoint] = std[ipoint]/math.sqrt(ndata) gstd[ipoint] = np.std(simgauss) gmean_precision[ipoint] = gstd[ipoint]/math.sqrt(ndata) my_plot_style() plt.figure(figsize = (8, 8)) plt.scatter(xaxis, mean_precision, s=15, c="#67ccc1", label="Cauchy distribution") plt.scatter(xaxis, gmean_precision, s=15, c="#ff3360", label="Gaussain distribution") plt.legend() plt.yscale("log") ax = plt.gca()ax.set_ylim([0.00001, 10]) plt.xlabel("N data point") plt.ylabel(r'\$\sigma\$/\$\sqrt{N}\$') plt.show() In [4]: from IPython import display display.Image("/home/judy/Astrophys/Week5/HW2/Central_limit_2.png") Out[4]: 10^{1} 10⁰ 10^{-1} Cauchy distribution Gaussain distribution 10^{-3} 10^{-4}

b. Write a code to estimate the uncertainty of the Pearson and Spearman correlation coefficients for (EBV, HI) and (EBV, H2) (15 points) In [10]: import numpy as np import matplotlib.pyplot as plt import astropy.io.fits as pf import scipy.stats as ss

5000 10000 15000 20000 25000 30000 35000 40000

N data point

3. Finishing your Pearson and Spearman correlation coefficients

In class, you have learned how to calculate Pearson and Spearman correlation coefficientts.

https://www.dropbox.com/s/h7545q0vzcqhi38/sky_maps_new_64_v6.fits?dl=0

a. Finishing your Pearson and Spearmn correlation coefficients (10 points)

def pearson_correlation(par1, par2): std1 = np.std(par1)std2 = np.std(par2)mean1 = np.mean(par1)mean2 = np.mean(par2)ndata = len(par1)cov = (1/ndata)*(np.sum((par1-mean1)*(par2-mean2)))rho = cov/(std1*std2)return rho def spearman_correlation(par1, par2): rank_par1 = len(par1) - rankdata(par1).astype(int) +1 rank_par2 = len(par2) - rankdata(par2).astype(int) +1 std1 = np.std(rank_par1) std2 = np.std(rank_par2) $mean1 = np.mean(rank_par1)$ $mean2 = np.mean(rank_par2)$ $ndata = len(rank_par1)$ cov = (1/ndata)*(np.sum((rank_par1-mean1)*(rank_par2-mean2))) rho = cov/(std1*std2)return rho bootstrap_time = 500 Pbootstrap_EBV_HI = np.zeros(bootstrap_time) Sbootstrap_EBV_HI = np.zeros(bootstrap_time) Pbootstrap_EBV_H2 = np.zeros(bootstrap_time) Sbootstrap_EBV_H2 = np.zeros(bootstrap_time) for iboot in range(0, bootstrap_time): random_EBV = np.random.randint(0, len(EBV), len(EBV)) random_HI = np.random.randint(0, len(HI), len(HI)) $random_H2 = np.random.randint(0, len(H2), len(H2))$ new_random_EBV = EBV[random_EBV] new_random_HI = HI[random_HI] $new_random_H2 = H2[random_H2]$ Pbootstrap_EBV_HI[iboot] = ss.pearsonr(new_random_EBV, new_random_HI)[0] Sbootstrap_EBV_HI[iboot] = ss.spearmanr(new_random_EBV, new_random_HI)[0] Pbootstrap_EBV_H2[iboot] = ss.pearsonr(new_random_EBV, new_random_H2)[0] Sbootstrap_EBV_H2[iboot] = ss.spearmanr(new_random_EBV, new_random_H2)[0] Pbperr_EBV_HI = np.std(Pbootstrap_EBV_HI) Sbperr_EBV_HI = np.std(Sbootstrap_EBV_HI) Pbperr_EBV_H2 = np.std(Pbootstrap_EBV_H2) $Sbperr_EBV_H2 = np.std(Sbootstrap_EBV_H2)$ print(f'Pearson(EBV, HI) uncertainty : {Pbperr_EBV_HI:.4f}') print(f'Spearman(EBV, HI) uncertainty : {Sbperr_EBV_HI:.4f}') print(f'Pearson(EBV, H2) uncertainty : {Pbperr_EBV_H2:.4f}') print(f'Spearman(EBV, H2) uncertainty : {Sbperr_EBV_H2:.4f}') Pearson(EBV, HI) uncertainty: 0.0043 Spearman(EBV, HI) uncertainty: 0.0044