

Homework assignment 1

1. (20 points) Let X be a \mathbb{R} -valued continuous random variable with Probability Density Function (PDF)

$$f(x) = \begin{cases} \exp(\theta - x), & x > \theta, \\ 0, & x \leq \theta. \end{cases} \quad (1)$$

- (a) (3 points) Compute $\mathbb{E}[X]$ and $\text{Var}(X)$.
 (b) (3 points) Let X_1, \dots, X_n be i.i.d. copies of X . Show that

$$\hat{\theta}_n := \frac{1}{n} \sum_{i=1}^n (X_i - 1), \quad (2)$$

is an unbiased estimator of θ .

- (c) (7 points) Derive the $100(1 - \alpha)\%$ confidence interval of θ for $\alpha \in (0, 1)$.
 (d) (7 points) Assume we have a set of observations $\mathcal{D} = \{10.0, 12.0, 15.0\}$. Using the fact that

$$\int_{-\infty}^{1.96} \mathcal{N}(x|0, 1) dx \approx 0.975, \quad (3)$$

compute the 95% confidence interval of θ . Do you find anything counter-intuitive? If so, what is weird about the confidence interval?

2. (10 points) Let (E, \mathcal{E}, μ) be a measure space. Using the definition of measures, show that

1. For any $A \subset B$, $\mu(A) \leq \mu(B)$.
2. For any A_1, \dots, A_n , $\mu(\bigcup_{i=1}^n A_i) \leq \sum_{i=1}^n \mu(A_i)$.

3. (5 points) Let $(X_i)_{i \geq 1}$ be an i.i.d. sequence of random variables with $\mathbb{E}[X_1] < \infty$ and $\text{Var}(X_1) < \infty$. Show that

$$\frac{X_1 + \dots + X_n}{n} \xrightarrow{\text{P}} \mathbb{E}[X_1]. \quad (4)$$

Hint: use Chebyshev's inequality, for a random variable X with $\mathbb{E}[X] = \mu < \infty$ and $\text{Var}(X) = \sigma^2 < \infty$,

$$\mathbb{P}(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}. \quad (5)$$

4. (15 points) Let $(X_n)_{n \geq 1}$ and $(Y_n)_{n \geq 1}$ be a sequence of random variables defined on a common probability space. Show that if $X_n \xrightarrow{\text{P}} X$ and $Y_n \xrightarrow{\text{P}} Y$ for some random variables X and Y , then $X_n + Y_n \xrightarrow{\text{P}} X + Y$.
 5. (15 points) Let $X_i \sim \text{Gamma}(\alpha_i, 1)$ for $i = 1, \dots, n$ with PDF

$$f_X(x) = \frac{x^{\alpha_i-1} e^{-x}}{\Gamma(\alpha_i)}, \quad (6)$$

where

$$\Gamma(z) := \int_0^\infty t^{z-1} e^{-t} dt, \quad (7)$$

denotes the gamma function. Show that

$$Y := \left(\frac{X_1}{\sum_{i=1}^n X_i}, \dots, \frac{X_n}{\sum_{i=1}^n X_i} \right), \quad (8)$$

is a Dirichlet random variable with PDF

$$f(y) = \frac{\Gamma(\sum_{i=1}^n \alpha_i)}{\prod_{i=1}^n \Gamma(\alpha_i)} \prod_{i=1}^n y_i^{\alpha_i-1}. \quad (9)$$

Hint: use the change of variable

$$(X_1, \dots, X_n) \rightarrow (Y_1, \dots, Y_{n-1}, Z), \quad (10)$$

where

$$Y_i = \frac{X_i}{Z} \text{ for } i = 1, \dots, n-1, \quad Z = \sum_{i=1}^n X_i. \quad (11)$$

6. (20 points) Consider the following Gaussian-Gaussian-inverse-Wishart model.

$$\theta := (\mu, \Sigma), \quad (12)$$

$$p(x|\theta) = \mathcal{N}(x|\mu, \Sigma) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left(-\frac{1}{2} (x - \mu)^\top \Sigma^{-1} (x - \mu) \right) \quad (13)$$

$$\begin{aligned} p(\theta) &= \mathcal{N}(\mu|m, \Sigma/r) \mathcal{W}^{-1}(\Sigma|\Psi, \nu) \\ &= \frac{r^{d/2}}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left(-\frac{r}{2} (\mu - m)^\top \Sigma^{-1} (\mu - m) \right) \\ &\quad \times \frac{|\Psi|^{\nu/2}}{2^{\nu d/2} \Gamma_d(\nu/2)} |\Sigma|^{-\frac{\nu+d+1}{2}} \exp \left(-\frac{1}{2} \text{tr}(\Psi \Sigma^{-1}) \right), \end{aligned} \quad (14)$$

where

$$\Gamma_d(z) := \pi^{\frac{d(d-1)}{4}} \prod_{j=1}^d \Gamma(z + (1-j)/2). \quad (15)$$

Let $X := \{x_1, \dots, x_n\}$ be a set of observations.

- (a) (7 points) Compute the posterior $p(\theta|X)$. (Hint: use the fact $x^\top A y = \text{tr}(A y x^\top)$).
- (b) (6 points) Compute the marginal likelihood $p(X)$.
- (c) (7 points) Show that $p(x|\theta)$ and $p(\theta)$ belong to conjugate exponential families. That is, show that they can be represented as

$$p(x|\theta) = \exp(T(x)^\top \eta - \mathbf{1}^\top A(\eta) - B(x)) \quad (16)$$

$$p(\theta) = \exp(\eta^\top \chi - \zeta^\top A(\eta) - \mathbf{1}^\top C(\chi, \zeta)). \quad (17)$$

(Hint: use the fact $\text{tr}(A^\top B) = \text{vec}(A)^\top \text{vec}(B)$).

7. (15 points) Define U_1, U_2 as follows:

$$U_1, U_2 \stackrel{\text{i.i.d.}}{\sim} \text{Unif}(-1, 1), \quad \text{accept } (U_1, U_2) \text{ if } U_1^2 + U_2^2 \leq 1. \quad (18)$$

Show that

$$X_1 = U_1 \sqrt{\frac{-2 \log(U_1^2 + U_2^2)}{U_1^2 + U_2^2}} \sim \mathcal{N}(0, 1), \quad X_2 = U_2 \sqrt{\frac{-2 \log(U_1^2 + U_2^2)}{U_1^2 + U_2^2}} \sim \mathcal{N}(0, 1). \quad (19)$$

Hint: use the intermediate transform

$$U_1 = R \cos \Theta, \quad U_2 = R \sin \Theta, \quad 0 \leq R \leq 1. \quad (20)$$