**Detecting exceptional temporal changes in genetic diversity using limited information.**

Julian Wittische, Pierre Legendre, Patrick M. A. James

Département de Sciences Biologiques, CP 6128 Succursale Centre-Ville, Université de Montréal, Pavillon Marie-Victorin, Montréal, QC, Canada, H3C 3J7

Correspondence: Julian Wittische; E-mail: [jwittische@gmail.com](mailto:jwittische@gmail.com)

Running title: Unique spatio-temporal genetic change

**INTRODUCTION**

Spatial and temporal variation in genetic information can tell us a great deal about demography and movement of populations. Population genetics approaches are, and will continue to be, widely used for conservation biology purposes as the approach that is used to translate these genetic data into meaningful inference (Allendorf, Hohenlohe, & Luikart, 2010; Harrisson, Pavlova, Telonis-Scott, & Sunnucks, 2014; Segelbacher et al., 2010). Among population genetic approaches, landscape genetics approaches bridge an important gap in the field of molecular ecology by providing information about the interaction between micro-evolutionary processes and landscape features (Balkenhol, Cushman, Storfer, & Waits, 2015; Manel & Holderegger, 2013; Manel, Schwartz, Luikart, & Taberlet, 2003; Wagner & Fortin, 2013). The vast majority of studies using these approaches focus on explaining the spatial variation rather than the temporal variation in genetic diversity. However, temporal gains and losses of genetic diversity are at the crux of many conservation issues because they influence the evolution and persistence of a species through processes such as local adaptation, maladaptation, or divergent natural selection.

One of the main ongoing challenges for biologists, is therefore to detect both when and where in the landscape, a significant gain or loss of genetic diversity occurs. Once detected, those changes in genetic diversity may be associated with natural or anthropogenic landscape changes, from local and abrupt like a wildfire, to global and long-term like climate warming. However, it is rarely possible to observe the effects of these events instantaneously and researchers are often left with spatial legacies in xYZ… which may not be readily observable from demographic data alone. When a demographic event does not constitute a selective pressure, alleles are randomly transferred from a generation to the next and genetic drift happens leading to a loss of diversity. Common examples of situations where genetic drift occurs include geographic isolation, population bottleneck and massive migrations from previously isolated populations, which would substantially reduce or alter local genetic variation. The result of such events in a local population tend to alter the genetic distance of this population with surrounding populations (Segelbacher et al., 2010). Detecting changes in the genetic make-up of a population through time, including the nature of those changes, may describe what is happening at the demographic level, and therefore serve as an alarm for managers.

Few methods currently exist for the temporal comparison of multivariate community and genetic data. For example… (*now go through existing tools – Then provide critique*). Very sophisticated methods exist to infer demographic history from genetic data, even from static genetic data. However, those methods are designed for very large genetic datasets with tens of thousands of loci and known ascertainment. This paper does not aim to infer demographic histories accurately, rather it aims to help researchers with two limited sets of genetic data to identify whether substantial change has occurred in one of the population they studied. It seems to me that you are trying to identify demographic histories using the TBI approach… right?  
Other methods do exist with which one can infer demographic history using (static) genetic data  
<https://journals.plos.org/plosgenetics/article?id=10.1371/journal.pgen.1003905> The relevance and performance of traditional approaches (e.g. PCA-based) to test change using temporal genetic datasets, where the objective is to find which population has indeed changed more significantly than others in the landscape, has not been evaluated. The rapid pace of global loss of genetic diversity (Leigh, Hendry, Vázquez‐Domínguez, & Friesen, 2019), is making it increasingly important to move beyond, single sampling/time, snapshot research (Draheim, Moore, Fortin, & Scribner, 2018).

A permutation-based statistical inference method for the analysis of spatial-temporal changes in community composition have recently been proposed (Legendre & Gauthier, 2014; Shimadzu, Dornelas, & Magurran, 2015). Temporal Beta-diversity Indices (TBI; Legendre 2018) were designed to assess whether there are sites where the difference in community composition between survey times seems exceptionally large. This approach has not yet been tested nor applied to the question of temporal variation in genetic data. The method involves estimating temporal change in each sampling site between two dates using a dissimilarity index/distance, testing the significance of each change through permutations, and partitioning the change into losses and gains. Comparing genetic data at two different dates, whether or not they were separated by an *a priori* known event, may help us understand more about the ecological processes shaping the system. A strong genetic change would also indicate the parts of the landscape where an event had the strongest effects, or highlight which sites should be investigated if managers are not aware of an *a priori* known event.

In this study, we describe a method to identify locations that have undergone significant genetic change through time. Identify such locations, and quantifying other locations relative temporal genetic change, is important because… ??? . To demonstrate the effectiveness and applicability of the approach we use both empirical and simulation data. Empricial data represent… Simualtion data were generated using… We simulated scenarios where part of the landscape is affected by non-selective demographic changes mimicking the effects of common demographic event events. We then used TBI to measure changes in genetic make-up of our populations, and evaluated the power and error rates associated with this approach. Finally, we illustrated the possibilities of this approach through applications on two real genetic datasets. We predict that dispersal will affect our ability to detect the genetic legacies of an event, we predict that the higher the number of populations affected by extraordinary events, the lower the performance of the TBI testing procedure, and finally we predict that the longer the time between samplings, the harder it will be to identify where and when a demographic event occurred. is to assess which is the most appropriate for genetic data and this type of question. Is TBI applicable?

**METHODS**

*Simulation framework*

To simulate changes in genetic information through time, we used the spatially-explicit gene flow simulation software CDMetaPOP (Landguth, Bearlin, Day, & Dunham, 2017). CDMetaPOP simulates dispersal and mating of individuals across a landscape, and allows to define the initial genetic structure, spatial distribution of individuals, dispersal characteristics, and life history traits of the population.

Loci were modelled after single nuclear polymorphism (SNP) and therefore are bi-allelic. The mutation rate was set as to reflect empirically-derived mutation rates found in many taxa (REF I gave to Ryan). Simulated individuals each carried a genome of 100 neutral loci without linkage disequilibrium. Each simulated population in the landscape had a maximum carrying capacity of 50 individuals, and each simulated landscape comprised 25 (a grid of 5 by 5) interconnected such populations with structural connectivity only reflecting geographical distance (Fig. 2). That corresponds to a maximum of 1250 individuals in the landscape. Each simulation was run for 100 generations before a demographic event was forced on up to three populations in the landscape. 10 more generations were simulated after the event.

We simulated 180 replicates for each scenario, with the new allocation of allelic frequencies for each replicate. Those parameters were chosen as a compromise between realism and computational time limitations, and we believe they were appropriate to produce the complex evolutionary dynamics necessary to reasonably realistic and useful genetic data.

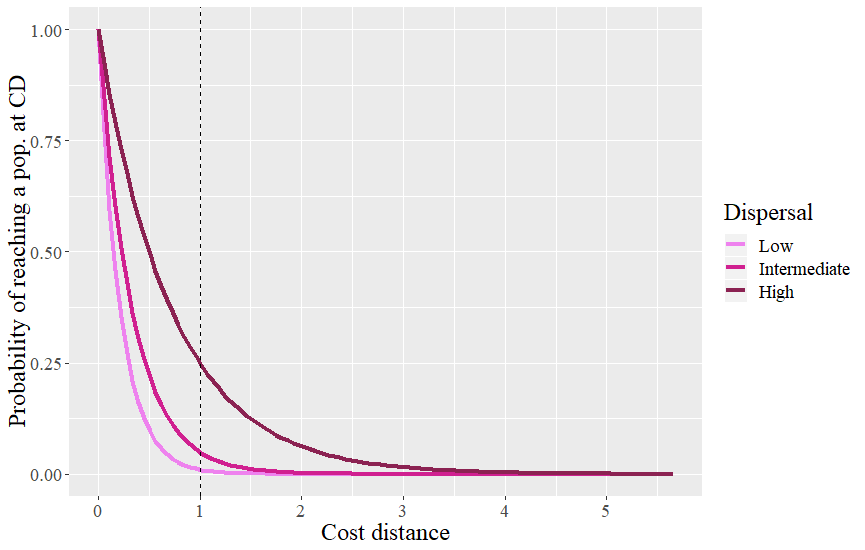
We examined the influence of dispersal, demographic event type, and demographic event spatial extent on the persistence of genetic spatial legacies using this simulation model. With 3 dispersal regimes, 2 different demographic event types and 3 different numbers of populations affected, we have 18 different scenarios giving us a total of 3240 (18 x 180) simulations. In the next sections, we detail how we modelled the aforementioned three factors. Additional material concerning the use of the simulator can be found in Supplementary Materials.

*Dispersal regimes*

The dispersal of individuals was controlled through a dispersal kernel based on a negative exponential distribution from which the distance realised by an individual () is sampled. This distribution transforms the cost distance of travel (CD) between cells according to a single parameter (*B*): . Cost distances used here are simply the geographical distances between the centroids of the populations. The values created through the use of the negative exponential distribution can then be rescaled using the maximum and the minimum distance possible in the landscape, which gives us the probability that an individual reaches a population beyond a certain CD:

When is randomly sampled as being higher than 1, the target population to which an individual travels, was selected randomly from the set of populations available at the distance selected in the previous step. Otherwise, the individual stays within its original population. We chose this way of modelling dispersal so that most individuals stay within their original population, that is more individuals randomly travel a distance below 1 than higher, while keeping opportunities for occasional long distance dispersal. This holds advantages compared to simpler approaches such as nearest neighbours or linear probability (REF).

In order to investigate the effect of different levels of dispersal, we changed the dispersal kernel by choosing values of *B* which would give us low, intermediate and high dispersal (Fig.1). We considered the % of individuals reaching an adjacent population as an indicator of the intensity of dispersal. We therefore respectively chose 1% (*B* = 2), 5% (*B* = 1.301), and 25% (*B* = 0.6015).



**Fig.1:** Probability of dispersal of an individual at all possible distances in the landscape, for three different dispersal scenarios. If an individual disperse below a distance of 1 (dashed line), then it does not leave its original population.

*Demographic events design*

The first demographic event we considered involves modelling an exogeneous immigration from a previously isolated population otherwise sharing the same characteristics as other populations. This population was simulated during the same number of generations and the cost distance from the isolated population to the target population(s) and was set to 0 between the 100th and 101st generations, mimicking a mass immigration event between the two. The cost distances are then set back to normal.

The second scenario involves a demographic bottleneck through massive mortality. To do that, the carrying capacity of the disturbed population was set to 20% of its original value between the 100th and 101st generations.

*Number and position of target populations*

Beyond the dispersal intensity and the demographic event type, we wanted to evaluate how the number of target populations affected the performance of our testing procedure. To achieve this, we disturbed from 1 to 3 populations among the 25. When only 1 population was disturbed we partitioned the 180 replicates of that scenario equally among 6 populations in the landscape. Because our landscape is homogenous and symmetric, only 6 positions need to be assessed (Fig. 2). When several (*k*) populations were disturbed, we randomly sampled 1 position among the 6 previously described and randomly picked 1 or 2 additional populations directly adjacent (when possible) to it. We did this 6 times (30 replicates for each set of targeted populations). We choose to pick target populations this way to respect the spatial autocorrelation often exhibited in demographic events.

*Controls*

To further the quality and transparency of our simulation experiments, we used simulations designed to serve as controls for the rest of the scenarios. Those control populations are never affected by any event and therefore only display other sources of genetic variation such as gene flow, drift, and mutation. Dispersal ability was therefore the only parameter to change for the controls, giving us 3 control scenarios. We evaluated the FPR of those three control scenarios (no need for FNR because there are no true positives/false negatives so it was always equal to 0). When describing the performance of other scenarios with similar dispersal parameters, we always put control values as a reference.

*Genetic dissimilarity*

The Chord distance has been commonly used in both community ecology (Orlóci 1967; Legendre & Borcard 2018) and population genetics (Cavalli-Sforza & Edwards 1967; Balkenhol et al. 2016). We chose chord distance because it has already been tested for use with TBI with non-genetic data (Legendre 2019) and because it may be more appropriate than other indices of genetic dissimilarity when most of the variation among populations is due to recent changes (Takezaki & Nei 1996; Kalinowski 2002) as it does not assume populations are in drift-mutation equilibrium. Here we use the Chord distance to calculate genetic dissimilarity of a single site sampled at two different points in (simulated) time.

*Timing*

To assess the influence of the timing of sampling on our ability to detect significant temporal change, we used TBI on simulation data collected each year, up to 5 years after the event, and compared them with data from the event year. We did the same with the earliest sampling period, that is we used simulation data dating 5 years before the event, as well as each year until the event. We used two out of the eighteen scenarios, in order to save computation time, and chose the most extreme according to the results, to represent the “easiest” and “hardest” contexts to detect change. We used the ubiquitous 0.05 p-value threshold as it was a good compromise between decent FPR and FNR in our initial results, and because it is the most likely threshold users would pick.

*Permutation approaches*

One of the most crucial steps in describing change is to evaluate the significance of the change. Indeed, without a mean to distinguish typical variation from truly atypical change, decision makers and researchers would be left to arbitrarily set thresholds for what constitute change. Permutation-based approaches may help us to achieve this by creating a distribution of values which can then be compared to the measured value of change. Most calculations used in this paper are based on the *TBI()* function available in the *R* package *adespatial* (Dray et al., 2019). Three permutation approaches were considered to test the significance of TBI, but only one was kept in the final version of the *TBI()* function (Legendre, 2019). Because they were tested on very a different type of data, we used an older version of the *TBI()* function to tests which one should be kept for genetic data. The first permutation approach consisted in permuting a locus in the same way in both (original sampling and resampling) gene frequency data frames. The second permutation approach consisted in permuting a locus independently in both data frames. The third permutation approach consisted in permuting sampling sites in both data frames. We summarized statistical performance per permutation approach, and used the best approach to answer all other questions.

*Statistical performance*

We used the False Positive Rate (FPR) and False Negative Rate (FNR) frameworks to assess statistical performance of the TBI testing procedure and to evaluate which of the permutation procedures, and permutation p-value thresholds, is most appropriate. A false positive is a population that we *a priori* know did not undergo any specific demographic event, but has been classified as having experienced one of the two simulated demographic events by the testing procedure. A false negative is a population that we had set as target for demographic event but that was not classified as having been disturbed by the testing procedure. FPR represents the number of false positives over the total number of negatives, and FNR represents the number of false negative over the total number of positives. A high FPR means that we often select the wrong population(s). A high FNR means that we often miss the right population(s). The higher the FNR, the lower the power of our testing procedure. Because choosing a proper threshold for the TBI permutation tests is important in order to find a compromise between power and selectivity, we evaluated statistical performance across a range of thresholds: 0.0001, 0.00025, 0.0005, 0.00075, 0.001, 0.0025, 0.005, 0.0075, 0.01, 0.025, 0.05, 0.075, 0.1.

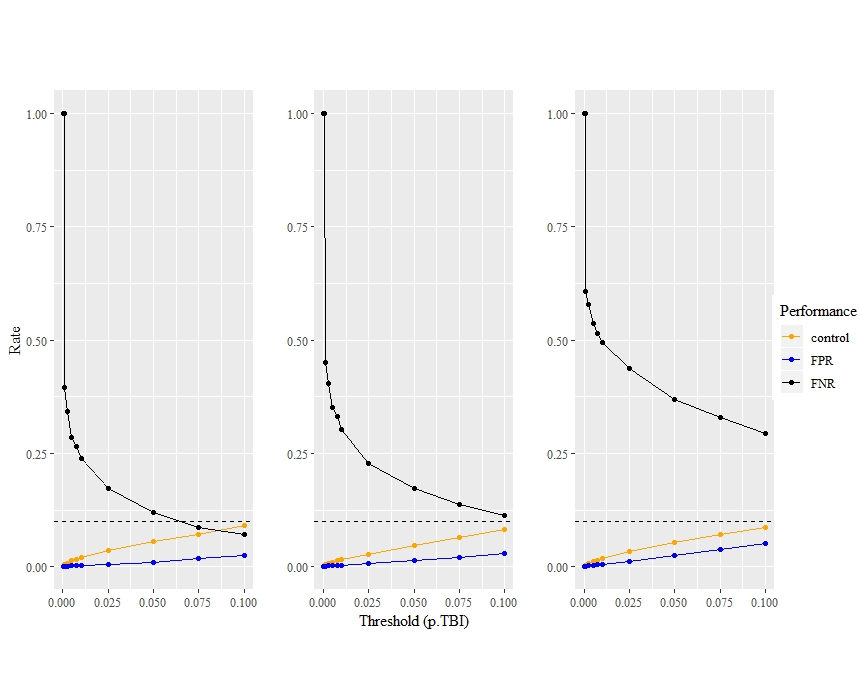
**RESULTS**

*Permutation approach*

The first permutation approach is the only one that is functional with genetic data. Indeed, the second and third approaches most often failed to find any significant change. This means that they never found any false positive (FPR = 0), which is great, but also that they very rarely found any true positive (FNR > 0.9), regardless of the scenario or the p-value threshold we used. Because only the first approach was suitable to study simulation outputs, we used it for the rest of the analyses.

*Dispersal ability*

As hypothesized, the dispersal ability of an organism, relative to its landscape, greatly affects our ability to detect exceptional temporal changes from limited genetic datasets. Indeed, when we group scenarios with the same dispersal parameters (low, intermediate, high) together, FNR and FPR substantially increase with dispersal intensity (Fig. 3). This is true regardless of the used threshold and the bigger the threshold, the larger the difference between average values of FPR of the three scenarios (sup. Fig. S1). For example, at the ubiquitous 0.05 threshold, which here seems to be a decent compromise between low FNR and FPR, average FNR values are 0.1210, 0.1727 and 0.3702, for the low, intermediate and high dispersal scenarios respectively. At this threshold and for the same scenario groups, FPR also increases, from 0.0107 to 0.0138 and 0.0244.



C

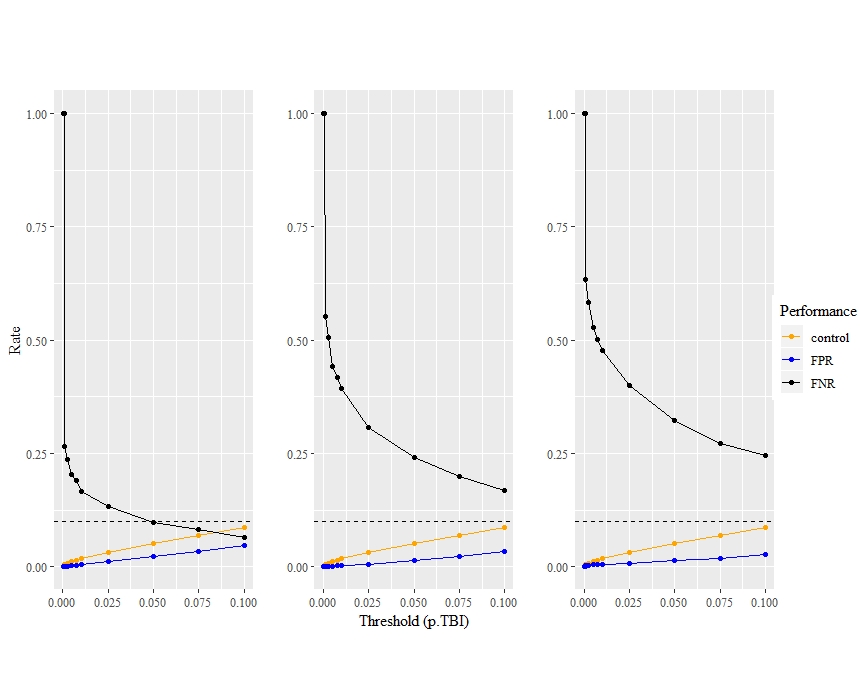
B

A

**Fig. 3.** Influence of dispersal ability on our ability to detect exceptional change. FPR and FNR values at 13 different p.TBI thresholds for low (A), intermediate (B), and high (C) dispersal scenarios. Control FPR values, from scenarios with identical dispersal parameters, are also featured. The dashed horizontal line indicates 0.1 which is the maximum threshold value used, for comparison with FPR values.

*Number of population affected*

The number of populations affected by an event also affects our ability to detect exceptional temporal change. When looking at groups of scenarios with the same number of affected populations (1, 2, and 3 populations), we can see that FNR increases with additional affected populations, regardless of which threshold is considered (Fig. 4). FPR values from scenarios with 2 affected populations are consistently higher than values from one affected population scenarios. FPR values from scenarios with 3 affected populations are on average higher than values from other scenarios up to a threshold of 0.01, and are on average lower for thresholds above 0.05 (sup. Fig. S2), therefore indicating an interaction between the number of affected populations and the threshold used in the permutation procedure. However, for thresholds that would be considered suitable regarding power (*e.g.* power > 50%), a higher number of populations leads to a lower FPR.



A

C

B

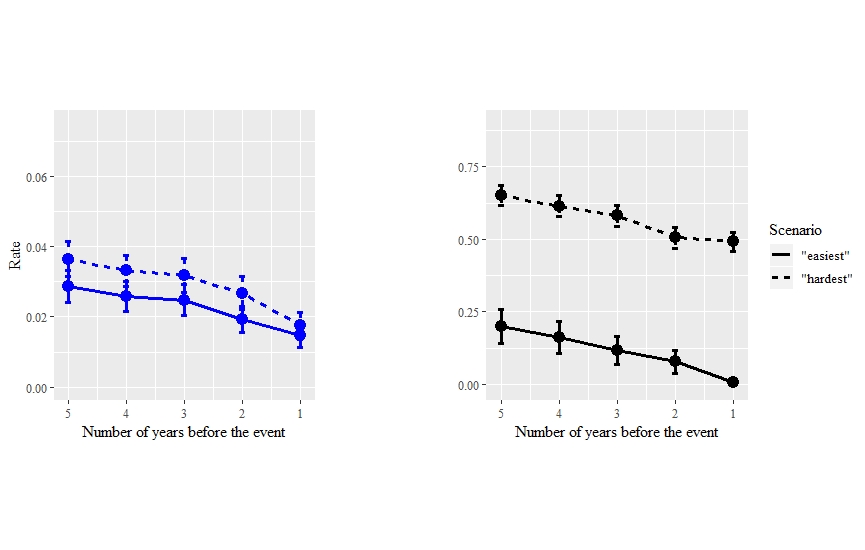
**Fig. 4.** Influence of the number of affected populations on our ability to detect exceptional change. FPR and FNR values at 13 different p.TBI thresholds for 1 (A), 2 (B), and 3 (C) affected populations scenarios. Control FPR values, from scenarios with identical dispersal parameters, are also featured. The dashed horizontal line indicates 0.1 which is the maximum threshold value used, for comparison with FPR values.

*Lag time between pre-event-sampling and event*

We can see in Fig. 5 A & B, that the longest the pre-event sampling is from the event, the less power and the more false positives we get. Sampling done 5 years before the event led to about twice as much false positives as sampling done the year before the event. The effect of time on FPR or FNR is similar regardless of the scenarios (Fig. 5 A & B), however FNR variation increased with time for the “easiest” scenario (immigration event + 1 population + low dispersal) whereas FPR variation did not for the “hardest” scenario (immigration event + 3 populations + “easiest” and the “hardest” scenario sharply change between 1 and 2 years, it stays about the same for longer periods between samplings.

*Lag time between event and post-event sampling*

As hypothesized from the nature of genetic processes in connected populations, the signal of the demographic event inflicted upon populations disappears gradually over time. When considering the scenario most likely to preserve the signal according to earlier results on FNR and FPR, the TBI approach was still able to avoid false adequately two years after the event (Fig. 5 C) but average FPR sharply increased at the three years’ mark, then increased linearly again in the following years. For the harder scenario FPR increased much faster with the years, following a slightly saturated curve, and reaching 5% of false positives after only two years (Fig. 5 C). Average FNR, and the width of its confidence intervals, increased linearly for the easier scenario, but (Fig. 5 D). Beyond the fact that its starting FNR at 0.05 was much higher for the harder scenario (Fig. 3; Fig. 5 D), it also increased much faster with time, reaching a plateau at unacceptable power values after a 4 years. With the harder scenario, almost 25% of power is lost as the result of only two generations.



A

B

C

D

**Fig. 5.** Influence of number years between the event and a pre-event sampling (A, B) or a post-event sampling (C, D) on averages and confidence intervals of FPR (A, C) and FNR (B, D), for two extreme scenarios with the 0.05 p-value threshold.

*Threshold and general performance*

Stricter values (lower values) for the TBI p-value threshold expectedly bring a better FPR but also bring a pathological FNR (low power). Indeed, across all scenarios, the FNR decreases exponentially when threshold values increase, while the FPR increases linearly (e.g. Fig. 3; Fig. 4). Notably, FPR values never surpassed 0.1, which was the maximum threshold chosen in our testing, which makes them acceptable (REF). The decrease in average FNR across all scenarios associated with an increase in the threshold value, is accompanied by a decrease of the associated standard variation, as soon as variation exists (FNR not equal to 1): from 0.3749 (0.001) to 0.2471 (0.1), considering all scenarios. In contrast, the increase in average FPR concurrent with an increasing of its variation: from 0 (0.0001) to 0.0377 (0.1).

*Control simulations*

Experimental FPR values consistently stayed below control FPR values, also the difference generally diminished with the intensity of dispersal (Fig. 3). This means that in the presence of an actual event, we were less likely to wrongfully choose a population as having been affected. Control FPR values did not vary between scenario groups (ANOVA; p-value = 0.353), which means that dispersal does not affect the selection of a random population as a positive. Finally, control FPR values never passed 0.1, which was the maximum threshold chosen in our testing.

**DISCUSSION**

A few words about why we are doing it again.

*TBI applicability to genetic data*

Our analyses have shown that TBI is applicable to genetic data under certain conditions. First, only one permutation approach (permutations done locus by locus, and in the same way for both samples) is suitable when using gene frequency data. The other permutation approaches were incontrovertibly poor in their ability to pick up on the genetic signal left by the demographic events occurring in some populations, as they almost never select any.

*A tool for future users of TBI on genetic data*

Simulations provide a very useful tool for the planning researchers who would want to   
  
*Limits*

How would population size or amount of genomic information affect results (recall the first review of Paul’s paper asked for other factos like this)

In other words, demographic processes are quickly diluting the signal by transferring the initial effect on genetic diversity to other populations.

Although signal of a past demographic event can be kept in richer genomic data (e.g. probability of mutational configurations in sequence blocks), gene frequency data in a high-dispersion species and connected landscape will not keep the signal beyond a year. To be investigated for non-SNP gene frequency data.

* Provided the landscape was resampled the year of the event, the closer the date of the first sampling, the better performance-wise. However, reasonable performance can be expected even if the first sampling was a few years before the event.
* Liberal thresholds should be used in order to obtain reasonable statistical power, while keeping the FPR low.
* In other terms, the more an organism disperses, the lower our power in identifying significant temporal change in genetic diversity.

**Paragraph discussing the influence of dispersal on the detectability of demographic event from gene frequency data.**

**Paragraph discussing the influence of the fraction of the landscape affected by a demographic event on the detectability of demographic event from gene frequency data.**

**Paragraph discussing how we were able to use TBI to test and illustrate temporal change** (results for empirical data not ready yet)**.**

**Paragraph discussing the limits of TBI use on genetic data, including the fact that it may need to be parameterized (e.g. choosing a threshold) based on landscape or taxa characteristics.**

**Paragraph discussing further investigation of the relative importance of genetic drift, gene flow and other forces, in shaping temporal variation.**

**Paragraph discussing the importance of LTER, exhaustive sampling, and the need to move beyond single-time snapshot studies of landscape genetics.**

*Tool for future field research*

While we used simulations to …, we do not advise future users to choose a default p-value threshold and

**ACKNOWLEDGEMENTS**

This research was supported by a grant to PMAJ and the TRIA Network from the Natural Sciences and Engineering Research Council of Canada (grant no. NET GP 434810-12), with contributions from Alberta Agriculture and Forestry, fRI Research, Manitoba Conservation and Water Stewardship, Canadian Forest Service (Natural Resources Canada), Northwest Territories Environment and Natural Resources, Ontario Ministry of Natural Resources and Forestry, Saskatchewan Ministry of Environment, West Fraser, and Weyerhaeuser. JW was also supported by a scholarship from the Forest Complexity Modelling (FCM) NSERC CREATE. We thank Hinatea Ariey for help with the creation of the first figure. Finally, we thank Jeremy Larroque for his comments on an earlier version of the manuscript.

**REFERENCES**

Allendorf, F. W., Hohenlohe, P. A., & Luikart, G. (2010). Genomics and the future of conservation genetics. *Nature Reviews. Genetics*, *11*(10), 697–709. https://doi.org/10.1038/nrg2844

Balkenhol, N., Cushman, S., Storfer, A., & Waits, L. (2015). *Landscape Genetics: Concepts, Methods, Applications*. Wiley-Blackwell.

Draheim, H. M., Moore, J. A., Fortin, M. J., & Scribner, K. T. (2018). Beyond the snapshot: Landscape genetic analysis of time series data reveal responses of American black bears to landscape change. *Evolutionary Applications*, *11*(8), 1219–1230. https://doi.org/10.1111/eva.12617

Harrisson, K. A., Pavlova, A., Telonis-Scott, M., & Sunnucks, P. (2014). Using genomics to characterize evolutionary potential for conservation of wild populations. *Evolutionary Applications*, *7*(9), 1008–1025. https://doi.org/10.1111/eva.12149

Landguth, E. L., Bearlin, A., Day, C. C., & Dunham, J. (2017). CDMetaPOP: an individual-based, eco-evolutionary model for spatially explicit simulation of landscape demogenetics. *Methods in Ecology and Evolution*, *8*(1), 4–11. https://doi.org/10.1111/2041-210X.12608

Legendre, P. (2019). A temporal beta-diversity index to identify sites that have changed in exceptional ways in space–time surveys. *Ecology and Evolution*, *9*(6), 3500–3514. https://doi.org/10.1002/ece3.4984

Legendre, P., & Gauthier, O. (2014). Statistical methods for temporal and space-time analysis of community composition data. *Proceedings of the Royal Society B: Biological Sciences*, *281*(1778). https://doi.org/10.1098/rspb.2013.2728

Leigh, D. M., Hendry, A. P., Vázquez‐Domínguez, E., & Friesen, V. L. (2019). Estimated six percent loss of genetic variation in wild populations since the industrial revolution. *Evolutionary Applications*, (April), 1–8. https://doi.org/10.1111/eva.12810

Manel, S., & Holderegger, R. (2013). Ten years of landscape genetics. *Trends in Ecology & Evolution*, *28*(10), 614–621. https://doi.org/10.1016/j.tree.2013.05.012

Manel, S., Schwartz, M. K., Luikart, G., & Taberlet, P. (2003). Landscape genetics: combining landscape ecology and population genetics. *Trends in Ecology & Evolution*, *18*(4), 189–197. https://doi.org/10.1016/S0169-5347(03)00008-9

Segelbacher, G., Cushman, S. A., Epperson, B. K., Fortin, M. J., Francois, O., Hardy, O. J., … Manel, S. (2010). Applications of landscape genetics in conservation biology: Concepts and challenges. *Conservation Genetics*, *11*(2), 375–385. https://doi.org/10.1007/s10592-009-0044-5

Shimadzu, H., Dornelas, M., & Magurran, A. E. (2015). Measuring temporal turnover in ecological communities. *Methods in Ecology and Evolution*, *6*(12), 1384–1394. https://doi.org/10.1111/2041-210X.12438

Wagner, H. H., & Fortin, M.-J. (2013). A conceptual framework for the spatial analysis of landscape genetic data. *Conservation Genetics*, *14*(2), 253–261. https://doi.org/10.1007/s10592-012-0391-5