Structural Estimation

Thomas H. Jørgensen

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### Outline

Introduction

Structural Estimation

# Stochastic Dynamic Programming

• Last time: Dynamic Programming Backwards induction Grids Interpolation

Structural Estimation

# Stochastic Dynamic Programming

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- Today:
  - Uncertainty:

Future income is uncertain

- + Another state variable: Permanent income
- + "Normalization" of one state variable.

# Stochastic Dynamic Programming

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Simulated Method of Moments (SMM/SMD)

Relate to "reduced-form"

Can we combine approaches?

See Eisenhauer, Heckman and Mosso (2015) Todd and Wolpin (2023)

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- Example: Buffer Stock model of Deaton (1991); Carroll (1992)
  - Estimated in Gourinchas and Parker (2002)

# Gourinchas and Parker (2002)

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Structural Estimation

- Approach:
  - 1. **Estimate model** with 2+ motives:

Buffer-stock motive: Income risk while working. Life cycle motive: Consumption in retirement.

# Gourinchas and Parker (2002)

- Research question: "Which savings motives dominate across life?"
- Approach:
  - 1. **Estimate model** with 2+ motives: Buffer-stock motive: Income risk while working. Life cycle motive: Consumption in retirement.
  - 2. Quantify importance of these motives over life Counterfactual simulations

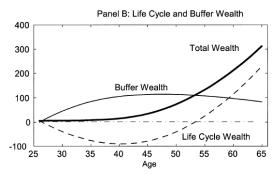


FIGURE 7.—The role of risk in saving and wealth accumulation.

#### Outline

Stochastic DP

# Buffer-stock model (Deaton-Carroll) Bellman equation

• Simplest version of the buffer-stock model is

$$V_{t}(M_{t}, P_{t}) = \max_{C_{t}} \frac{C_{t}^{1-\rho}}{1-\rho} + \beta \mathbb{E}_{t} \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right]$$
s.t.
$$A_{t} = M_{t} - C_{t} \quad \text{(assets)}$$

$$M_{t+1} = RA_{t} + Y_{t+1} \quad \text{(resources/cash-on-hand)}$$

$$Y_{t+1} = P_{t+1}\xi_{t+1} \quad \text{(income)}$$

$$P_{t+1} = GP_{t}\psi_{t+1} \quad \text{(perm. income)}$$

$$A_{t} \geq 0, \forall t \quad \text{(no borrowing)}$$

where  $\mathbb{E}_{t}[V_{t+1}(M_{t+1}, P_{t+1})] = \mathbb{E}[V_{t+1}(M_{t+1}, P_{t+1}) | M_{t}, P_{t}, C_{t}]$ are expectations over perm. and trans. income shocks,

$$\log \xi_{t+1} \sim \mathcal{N}(\mu_{\xi}, \sigma_{\xi}^2), \ \log \psi_{t+1} \sim \mathcal{N}(\mu_{\psi}, \sigma_{\psi}^2)$$

# Buffer-stock model (Deaton-Carroll) Bellman equation

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$$\begin{array}{lll} V_t(\textit{M}_t,\textit{P}_t) & = & \max_{\textit{C}_t} \frac{\textit{C}_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(\textit{M}_{t+1},\textit{P}_{t+1}) \right] \\ & \text{s.t.} \\ & A_t & = & \textit{M}_t - \textit{C}_t \qquad \text{(assets)} \\ & \textit{M}_{t+1} & = & \textit{R}A_t + Y_{t+1} \quad \text{(resources/cash-on-hand)} \\ & Y_{t+1} & = & \textit{P}_{t+1}\xi_{t+1} \quad \text{(income)} \\ & \textit{P}_{t+1} & = & \textit{GP}_t\psi_{t+1} \quad \text{(perm. income)} \\ & \textit{A}_t & \geq & \textit{0}, \forall t \qquad \text{(no borrowing)} \\ \end{array}$$

where  $\mathbb{E}_{t}[V_{t+1}(M_{t+1}, P_{t+1})] = \mathbb{E}[V_{t+1}(M_{t+1}, P_{t+1}) | M_{t}, P_{t}, C_{t}]$ are expectations over perm. and trans. income shocks,

$$\log \xi_{t+1} \sim \mathcal{N}(\mu_{\xi}, \sigma_{\xi}^2), \ \log \psi_{t+1} \sim \mathcal{N}(\mu_{\psi}, \sigma_{\psi}^2)$$

 Gourinchas and Parker (2002): "natural" borrowing constraint. mass-point at zero in trans. income shock distribution,  $\xi_{t+1}$ 

• Last period: Everything is consumed,

$$C_T^*(M_T, P_T) = M_T$$
$$V_T(M_T, P_T) = \frac{M_T^{1-\rho}}{1-\rho}$$

# Buffer-stock model (Deaton-Carroll) Bellman equation

• Last period: Everything is consumed,

$$C_T^{\star}(M_T, P_T) = M_T$$

$$V_T(M_T, P_T) = \frac{M_T^{1-\rho}}{1-\rho}$$

• Gourinchas and Parker (2002): Retirement-periods Assumes a linear post-retirement value (w.  $P_{T+1} = P_T$ )

$$V_{T+1}(M_{T+1}, P_{T+1}) = \kappa \cdot (M_{T+1} + h \cdot P_{T+1})$$

Motivated by a deterministic perfect credit market solution (estimate  $\kappa$  and h, through  $\gamma_0$  and  $\gamma_1$  – see e.g. Jørgensen and Tô, 2020)

• They also allow for time-varying taste-shifters,  $v_t(Z_t)$ .

• **Defining**  $c_t \equiv C_t/P_t$ ,  $m_t \equiv M_t/P_t$  etc. implies

$$A_t = M_t - C_t$$

$$A_t/P_t = M_t/P_t - C_t/P_t$$

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#### Normalization I

• **Defining**  $c_t \equiv C_t/P_t$ ,  $m_t \equiv M_t/P_t$  etc. implies

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and state transition

$$M_{t+1} = RA_t + Y_{t+1} \ M_{t+1}/P_{t+1} = RA_t/P_{t+1} + Y_{t+1}/P_{t+1} \ m_{t+1} = Ra_tP_t/P_{t+1} + \xi_{t+1} \ m_{t+1} = rac{R}{G\psi_{t+1}}a_t + \xi_{t+1}$$

The **adjustment factor**  $\frac{1}{G\psi_{t+1}}$  is due to changes in permanent income

#### Normalization II

• **Defining**  $v_t(m_t) = V_t(M_t, P_t) / P_t^{1-\rho}$  implies

$$\begin{split} V_t(M_t, P_t) &= \max_{C_t} \frac{C_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right] \\ &= \max_{c_t} \frac{(c_t P_t)^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right] \Leftrightarrow \\ V_t(M_t, P_t) / P_t^{1-\rho} &= \max_{c_t} \frac{(c_t P_t)^{1-\rho} / P_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) / P_t^{1-\rho} \right] \Leftrightarrow \\ v_t(m_t) &= \max_{c_t} \frac{c_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ \underbrace{V_{t+1}(M_{t+1}, P_{t+1}) / P_{t+1}^{1-\rho}}_{t-1} \cdot \underbrace{P_{t+1}^{1-\rho} / P_t^{1-\rho}}_{=(G\psi_{t+1})^{1-\rho}} \right] \\ &= \max_{c_t} \frac{c_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right] \end{split}$$

## Bellman equation in ratio form

$$v_{t}(m_{t}) = \max_{c_{t}} \frac{c_{t}^{1-\rho}}{1-\rho} + \beta \mathbb{E}_{t} \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right]$$
s.t.
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$$m_{t+1} = \frac{1}{G\psi_{t+1}} Ra_{t} + \xi_{t+1}$$

$$a_{t} \geq 0$$

Structural Estimation

• **Benefit:** Dimensionality of state space reduced,  $2 \rightarrow 1$ . Can this always be done?

#### Bellman equation in ratio form

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- **Benefit:** Dimensionality of state space reduced,  $2 \rightarrow 1$ . Can this always be done?
- No... Uses that utility is homothetic (budget constraint also important)

$$V_T(M_T, P_T) = \frac{M_T^{1-\rho}}{1-\rho} = \frac{(m_T P_T)^{1-\rho}}{1-\rho} = \frac{m_T^{1-\rho}}{1-\rho} P_T^{1-\rho}$$

such that  $v_T(m_T) = V_T(M_T, P_T)/P_T^{1-\rho}$  holds!

## Solving the model: Numerical Integration

Solved by backwards induction

Terminal period:

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Structural Estimation

For t < T:

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• How to evaluate expectations?

$$\mathbb{E}_{t}\left[\bullet\right] = \int_{\psi_{t+1}} \int_{\xi_{t+1}} \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right] f(d\psi_{t+1}, d\xi_{t+1})$$

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• Numerical Integration: Discretize into sum (Gauss-Hermite)

$$\mathbb{E}_{t}\left[\bullet\right] \approx \sum_{j=1}^{J} \sum_{k=1}^{K} [(G\psi^{(j)})^{1-\rho} v_{t+1}(m^{(j,k)})] \omega_{j} \omega_{k}$$

and interpolate  $v_{t+1}(\bullet)$  for values  $m^{(j,k)} = \frac{1}{Gt^{(j)}}Ra_t + \xi^{(k)}$  of  $\overrightarrow{m}$  grid.

Outline

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  - "All econometric work relies heavily on a priori assumptions. The main difference between structural and experimental (or "atheoretic") approaches is not in the number of assumptions but the extent to which they are made explicit." (Keane, 2010)

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Structural Estimation

#### Benefit of models:

- 1. Ensure *consistent* world view
- 2. Assumptions are clear: Better models are well defined.
- 3. Hopefully "deep" policy-invariant parameters (Lucas critique).

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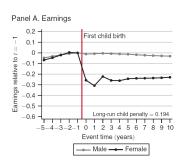
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- **Frontier:** Use exogenous variation to estimate structural model.

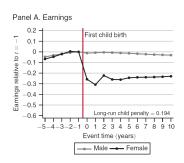
• Example: Event-studies (child-birth, Kleven, Landais and Søgaard, 2019)

- Reduced-form to be causal: "statistical" assumptions
  - No self-selection (timing)
  - No anticipation effects.
  - Parallel trends.



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- A model can allow for these assumptions to be violated But only through the chosen functional forms and mechanisms
  - "Economic" assumptions
  - Easier to debate and approve upon (?)

#### Structural Estimation

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Structural Estimation

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  - 1. Maximum Simulated Likelihood (MSL, SML)
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- Example model: Life-cycle buffer-stock model
  - States:  $M_{it}$ ,  $P_{it}$
  - Choice: Cit
- **Parameters** to estimate:  $\theta = \{\beta, \rho\}$ 
  - Calibration: G,  $\sigma_{\psi}$ ,  $\sigma_{\xi}$ , R, and  $\lambda$  ("known")

# Simulated Method of Moments (SMM/SMD)

•  $\Lambda^d = \frac{1}{N} \sum_{i=1}^N \Lambda^d_i$  are some moments in the data Could be avg., var, cov, regression-coefs, etc.

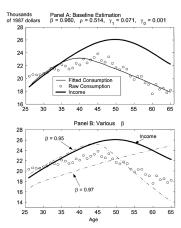
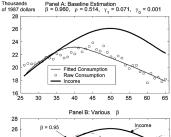


FIGURE 5.—The fitted consumption profile.

- $\Lambda^d = \frac{1}{N} \sum_{i=1}^N \Lambda^d_i$  are some moments in the data Could be avg., var, cov, regression-coefs, etc.
- $\Lambda^m(\theta)$  are the same moments calculated on simulated data



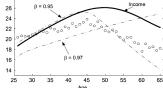
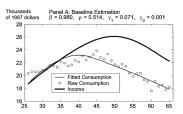


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- The difference is then

$$g(\theta) = \Lambda^d - \Lambda^m(\theta)$$



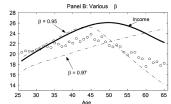


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SMM then is

$$\hat{\theta} = \arg\min_{\theta} g(\theta)' Wg(\theta)$$

 $\beta = 0.960$ ,  $\rho = 0.514$ ,  $\gamma_{s} = 0.071$ ,  $\gamma_{s} = 0.001$ 26 24 22 20 18 16 <sup>L</sup> Panel B: Various B

Panel A: Baseline Estimation

Structural Estimation 

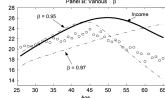


FIGURE 5.—The fitted consumption profile.

where W is **weighting matrix**.

## Weighting Matrix, W

 Common weighting matrices, W, are (should be positive-definite)

#### 1. Theoretically optimal

Inverse of covariance matrix of empirical moments Can cause problems in finite samples

#### 2. Identity, /

Equal weighting.

Does not take level-differences out of moments

3. Diagonal matrix with inverse of empirical moment variances Removes "level" differences. Scales with uncertainty about empirical moments Popular

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#### 4. Freely chosen

Focus on fitting some specific dimensions of the data

- 1. **Solve** the buffer-stock model and **simulate** a full panel
- 2. Construct a data set from the simulated data
- 3. Try to **estimate**  $\theta = \{\beta, \rho\}$ using as moments the average wealth for each age between 40 and 55  $\Lambda^d = (A_{40}, A_{41}, \dots, A_{55})$

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I will now describe how to calculate the objective function

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W\left(\Lambda^d - \Lambda^m(\theta)\right)$$

for a given value of  $\theta$ .

This function should then be minimized to get

$$\hat{\theta} = \arg\min_{\theta} \, Q(\theta)$$

1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)

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- 2. For s = 1, ..., S:
  - 2.1 Simulate N agents for T periods to get

$$C_{it}^{(s)}(\theta) = P_{it}^{(s)} \cdot \check{c}_{t}^{\star}(M_{it}^{(s)}(\theta)/P_{it}^{(s)};\theta)$$

$$M_{it}^{(s)}(\theta) = RA_{it-1}^{(s)}(\theta) + Y_{it}^{(s)}$$

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for some initial  $A_{i0}$  and  $P_{i0}$  and draws of ?

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$$\begin{split} C_{it}^{(s)}(\theta) &= P_{it}^{(s)} \cdot \check{c}_{t}^{\star} (M_{it}^{(s)}(\theta) / P_{it}^{(s)}; \theta) \\ M_{it}^{(s)}(\theta) &= RA_{it-1}^{(s)}(\theta) + Y_{it}^{(s)} \\ A_{it-1}^{(s)}(\theta) &= M_{it-1}^{(s)}(\theta) - C_{it-1}^{(s)}(\theta) \\ Y_{it}^{(s)} &= P_{it}^{(s)} \xi_{it}^{(s)} \\ P_{it}^{(s)} &= GP_{it-1}^{(s)} \psi_{it}^{(s)} \end{split}$$

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for some initial  $A_{i0}$  and  $P_{i0}$  and draws of  $\xi_{it}^{(s)}$  and  $\psi_{it}^{(s)}$ .

2.2 Calculate moments using simulated data,  $\Lambda_s(\theta) = \{\frac{1}{N} \sum_{i=1}^{N} A_{i+}^{(s)}(\theta)\}_{t-10}^{55}$ 

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- 2.2 Calculate moments using simulated data,  $\Lambda_s(\theta) = \{\frac{1}{N} \sum_{i=1}^{N} A_{it}^{(s)}(\theta)\}_{t=40}^{55}$
- 3. Calculate the objective function with  $\Lambda^m(\theta) = \frac{1}{5} \sum_{s=1}^{5} \Lambda_s(\theta)$

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W \left(\Lambda^d - \Lambda^m(\theta)\right)$$

1. Solve model to get  $c_t^*(m; \theta)$  for all t on a grid of m (2-dim array)

Structural Estimation

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- 1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)
- 2. Simulate  $\tilde{S} = SN$  agents for T periods to get

$$C_{t}^{(s)}(\theta) = P_{t}^{(s)} \cdot \check{c}_{t}^{\star}(M_{i}^{(s)}(\theta)/P_{t}^{(s)};\theta)$$

$$M_{t}^{(s)}(\theta) = RA_{t-1}^{(s)}(\theta) + Y_{t}^{(s)}$$

$$A_{t-1}^{(s)}(\theta) = M_{t-1}^{(s)}(\theta) - C_{t-1}^{(s)}(\theta)$$

$$Y_{t}^{(s)} = P_{t}^{(s)}\xi_{t}^{(s)}$$

$$P_{t}^{(s)} = GP_{t-1}^{(s)}\psi_{t}^{(s)}$$

Structural Estimation 0000000000000000000

for some initial  $A_0$  and  $P_0$  and draws of  $\mathcal{E}_{+}^{(s)}$  and  $\psi_{+}^{(s)}$ .

- 1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)
- 2. Simulate  $\tilde{S} = SN$  agents for T periods to get

$$C_{t}^{(s)}(\theta) = P_{t}^{(s)} \cdot \check{c}_{t}^{\star}(M_{i}^{(s)}(\theta)/P_{t}^{(s)}; \theta)$$

$$M_{t}^{(s)}(\theta) = RA_{t-1}^{(s)}(\theta) + Y_{t}^{(s)}$$

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Structural Estimation 0000000000000000000

for some initial  $A_0$  and  $P_0$  and draws of  $\xi_{\star}^{(s)}$  and  $\psi_{\star}^{(s)}$ 

3. Calculate simulated moments,  $\Lambda^m(\theta) = \{\frac{1}{5} \sum_{s=1}^{\tilde{S}} A_t^{(s)}(\theta)\}_{t=40}^{55}$  now

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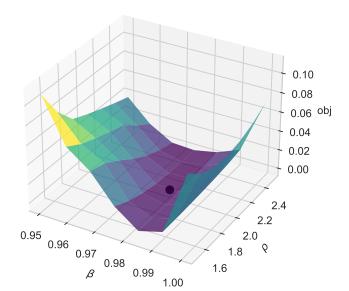
Structural Estimation

for some initial  $A_0$  and  $P_0$  and draws of  $\zeta_{+}^{(s)}$  and  $\psi_{+}^{(s)}$ 

- 3. Calculate simulated moments,  $\Lambda^m(\theta) = \{\frac{1}{\xi} \sum_{s=1}^{\tilde{\xi}} A_t^{(s)}(\theta)\}_{t=40}^{55}$  now
- 4. Calculate the objective function

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W\left(\Lambda^d - \Lambda^m(\theta)\right)$$

### Buffer-stock: MSM



### Indirect inference / minimum distance

- Many different names for very similar approaches
  - McFadden (1989): Method of Simulated Moments (MSM)
  - Duffie and Singleton (1993): Simulated Minimum Distance (SMD)

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• Gouriéroux, Monfort and Renault (1993) + Smith (1993): Indirect Inference (II)

### Indirect inference / minimum distance

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- Gouriéroux, Monfort and Renault (1993) + Smith (1993): Indirect Inference (II)
- SMD/II rely on an auxillary statistical model
  - ullet Let  $\Lambda^d$  be the parameters of the auxillary model when estimated on the actual data
  - Let  $\Lambda_s(\theta)$  be the parameters of the auxiliary model when estimated on simulated data
- Note: The auxiliary statistical model is misspecified and its parameters are thus typically not interpretable

### Simulation Pitfalls

- FIX the seed (or draws!)
- Flat objective function!
  - Discrete choices: Taking a mean of an indicator function

- Gradient based numerical optimization will likely FAIL!
  - Use, e.g., scipy.optimize.minimize(fun , method='Nelder-Mead') (Nelder-Mead)
  - Or some smoothing device (e.g. Logit)
- As  $N, S \rightarrow \infty$  this problem vanishes
- The problem is also less severe around  $\theta_0$
- Continuous outcomes do not have this problem

### Asymptotics

 MSM is consistent and asymptotically normal under standard assumptions

$$\sqrt{\textit{N}}(\hat{ heta}- heta_0) 
ightarrow \mathcal{N}( exttt{0,} (1+\textit{S}^{-1})\textit{V})$$

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where  $\theta_0$  is vector of true parameters

Standard formulas for V:

$$V = (G'WG)^{-1}G'W\Omega W'G(G'WG)^{-1}$$

where  $G = -\frac{\partial \Lambda^m(\theta)}{\partial \theta}$  is the Jacobian of the objective function.  $\Omega = Var(\Lambda_i^d)$  is the variance of the (individual) moments in the data. **Remember:** Standard errors are large if large changes in  $\theta$  imply small changes in the objective function

### Identification

• Is there enough variation in the data to identify  $\theta$ ? Very hard to prove anything because the model is typically strongly non-linear

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  - "How much does the variance of  $\theta$  increase if we drop a (set of) moments?"

non-linear

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  - "How much does the variance of  $\theta$  increase if we drop a (set of) moments?"
- Graphical inspection is useful: Plot the objective function in the neighborhood of the found optimum
- Problems:
  - The objective function might have multiple minima (no global solver exists)
  - 2. The objective function could be very flat in some directions (increasing S might help)

## Robustness/Sensitivity

- Curse of dimensionality and lack of identification
  - ⇒ we cannot estimate all the parameters of the model

- ⇒ first step estimation/calibration is often necessary
  - 1. Calculations on own data (e.g. exogenous processes)
  - 2. References to previous estimates
  - 3. Standard choices

## Robustness/Sensitivity

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  - $\Rightarrow$  we cannot estimate all the parameters of the model
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    - 1. Calculations on own data (e.g. exogenous processes)
    - 2. References to previous estimates
    - 3. Standard choices
- Robustness: Can we vary the calibration choices without changing the result substantially?

- "Sensitivity to Calibration": (Jørgensen, 2023)
- "How much does estimates of  $\theta$  change when 1. step calibrations change?"

### Calibration vs. Estimation

• **Estimation** or **calibration**: What is the difference? (my take)

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- Estimation: "systematic" Use a solver to minimize a criteria function wrt.  $\theta$ Report standard errors on  $\hat{\theta}$ Time-consuming!

### Calibration vs. Estimation

- **Estimation** or **calibration**: What is the difference? (my take)
- Estimation: "systematic" Use a solver to minimize a criteria function wrt.  $\theta$ Report standard errors on  $\hat{\theta}$ Time-consuming!
- Calibration: "hand-held" Use a (small) grid of values for  $\theta$ to minimize some moment(s). Sometimes eyeballing, sequentially for each parameter. Do not report standard errors Less time-consuming!

### Next Time

#### Next time:

Static and dynamic labor supply Recap for some + new stuff for most.

#### Literature:

Keane (2011, sections 1–5): "Labor Supply and Taxes: A Survey"

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- Read before lecture
- Reading guide:
  - Section 1: short Introduction
  - Section 2: Optimal Taxation, Motivation. Skim fast.
  - Section 3: Basic model. Key, focus here.
  - Section 4: Econometric issues. Skim.
  - Section 5: Roadmap of empirical literature. Short, read.

(Remaining: empirical literature.)

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