## Parenthood Timing and Gender Inequality

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#### Motivation

Gender inequality in Western labor markets emerges when individuals become parents (Kleven et al., 2024)

Quantifying the causal effect of parenthood is central to understanding gender inequity and informing policy

Existing evidence is highly conflicting

The conflict is tied to two factors that complicate the analysis:

- 1. Parenthood (timing) may be selective: human capital, wealth, health, career prospects
- 2. Effects may depend on timing: age of children, career stage at childbirth

Existing methods only address 1. or 2.

### This Paper

How much can we say about the causal effect of parenthood?

- ▶ How would labor market outcomes of parents change if they did not have children
- 1. New methodological approach robust to selection and dynamic effects
  - Leverage quasi-experimental variation in assisted conception procedures
- 2. Empirical evidence using administrative Dutch data
  - ► Focus on couples undergoing artificial insemination
- Unified formal framework to disentangle the extent of selective timing and timingdependent effects
  - Quantify bias in leading methods that disregard either factor

#### Preview of Main Results

- ▶ Parenthood persistently reduces women's work hours and income
  - ► Yearly reductions between 9 and 24 percent
- Parenthood causes a large share, but far from all, of post-child gender inequality
  - ▶ Between 36 and 54 percent in work hours and up to 46 percent in income
- ▶ Both selection and dynamic effects are substantial
  - Bias as extreme as attributing all or none of gender inequality to parenthood

#### Literature and Contribution

- 1. Existing methods for quantifying parenthood's impacts rely on restrictive assumptions about dynamic effects or selection
  - ► Address selection: Hotz et al. (2005); Agüero & Marks (2008); Cristia (2008); Miller (2011); Lundborg et al. (2017); Bensnes et al. (2023); Gallen et al. (2023); Lundborg et al. (2024)
  - Address dynamic effects: Fitzenberger et al. (2013); Angelov et al. (2016); Adda et al. (2017); Bütikofer et al. (2018); Kleven et al. (2019); Melentyeva & Riedel (2023); Kleven et al. (2024)

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# I provide a unified comparison framework and demonstrate that accounting for selection and dynamic effects is sufficient to reconcile the conflicting results

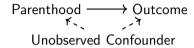
- 3. Addressing selection and dynamic effects is a common challenge
  - ▶ Education programs with multiple admission cycles, assignment to judges, promotion tournaments

### Method applicable to many other settings with sequential quasi-experiments

### Identification Challenge

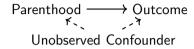
 $\begin{array}{c} \text{Parenthood} & \longrightarrow \text{Outcome} \\ \text{``} & \text{``} & \text{`'} \end{array}$  Unobserved Confounder

### Identification Challenge

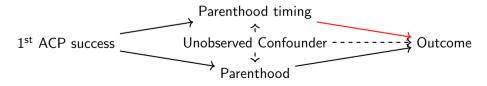


- Addressing selection requires a natural experiment
- Consider couples undergoing assisted conception procedures (ACP) for first child
- Sperm is inserted into uterus—assume conceptionis as-good-as-random
  - Uncorrelated with past labor market outcomes conditional on age
- ▶ 75% become mothers after their first ACP fails

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Theoretical bias

### Outline of the Approach

Objective is a clean comparison uncontaminated by timing differences:

Women conceiving at first ACP vs similar childless women with failed first ACP

Challenge: women remaining childless after first ACP fails may be a selective group:

- 1. Decision to undergo more ACPs may be selective—complicates using later ACPs
- 2. Conception via non-ACP means might be selective

Step 1: selection via subsequent ACPs can be addressed using the realized ACP count

▶ Women who tried more were less likely to never succeed ex-ante: weight them more

Step 2: non-ACP conceptions can be addressed using a worst-case bounding approach

Assume most extreme selection into non-ACP births consistent with the data

Crucial assumption: ACP outcomes are as good as random (cond. on age)

#### Model: Outcomes

Moment  $t \in \{1, ..., T\}$  since woman's first ACP

Outcome when motherhood begins at first ACP:

$$Y_t(1)$$

Childless outcome:

$$Y_t(0)$$

Outcome when motherhood begins in period k:

$$Y_t(k)$$

These scenarios involve women trying to conceive through ACPs

▶ I will first focus on quantifying impacts in these scenarios

To simplify exposition:

- ightharpoonup We are at t=T
- $ightharpoonup Y_T(k) = Y_T(later) \text{ for all } k > 1$



#### Model: Latent Variables and Treatment Effect

Women differ in two unobserved characteristics:

- "Willingness" to undergo ACPs,  $W \in \{1, \dots, \overline{w}\}$ 
  - ▶ Would undergo W ACPs for the first child if all ACPs failed
- ▶ "Reliance" on ACPs,  $R \in \{0, 1\}$ 
  - No child if all ACPs fail, R = 1
  - ightharpoonup "Reliers"  $\supseteq$  "compliers" (no child if first ACP fails)

Average treatment effect for reliers:

$$au_{ATR} = \mathbb{E}[Y(1) - Y(0)|R = 1]$$

- ▶ Reliers are a more general group than compliers
- As compliers, reliers may change over time, which I will address in extensions

### Model: Observables

- $\triangleright$  ACP j success indicator,  $Z_j$ 
  - $ightharpoonup Z_i = 0$  if failed or did not happen
  - Only ACPs before the first child
- Number of realized ACPs:

$$A = \min \left( \{j : Z_j = 1\} \cup \{W\} \right)$$

Parenthood indicator:

$$D = Z_A + (1 - Z_A)(1 - R)$$

- One-sided non-compliance
- Realized outcome:

$$Y = Y(0)(1 - D) + Y(1)DZ_1 + Y(later)D(1 - Z_1)$$

### Sequential Unconfoundedness

### Assumption (Sequential Unconfoundedness)

$$(Y(1), Y(0), R, W) \perp \!\!\! \perp \!\!\! Z_j | A \geq j.$$

In words: once sperm/embryo at ACP j are implanted, whether this results in a conception is as-good-as-random

- The decision to undergo the procedure can be endogenous
- ightharpoonup Y(1), Y(0), R and W can be related
- Main method relaxes to covariate-conditional version
  - E.g. age at procedure, technology

W=1 (willing to try once)

$$W=1$$
 (willing to try once)

$$Z_1 = 1$$

$$Z_1 = 0$$

$$W=1$$
 (willing to try once)

$$Z_1 = 1$$

$$\mathbb{E}[Y(1)]$$

$$Z_1 = 0$$

$$\mathbb{E}[Y(0)]$$

$$W=1 \ ext{(willing to try once)}$$

$$Z_1=1$$

$$Z_1 = 0$$

$$W=1$$
 (willing to try once)

$$Z_1 = 1$$

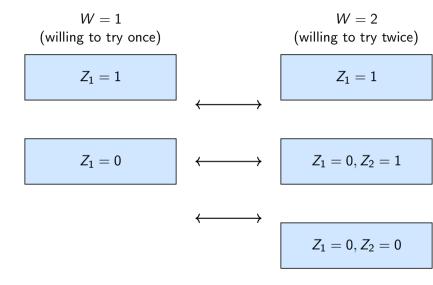
$$Z_1=0$$

$$W = 2$$
 (willing to try twice)

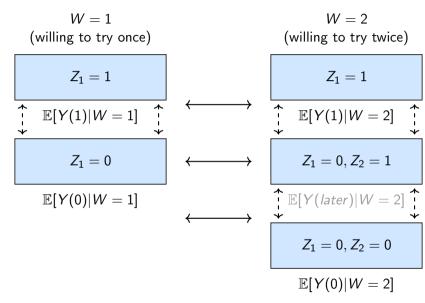
$$Z_1 = 1$$

$$Z_1=0,Z_2=1$$

$$Z_1=0,Z_2=0$$



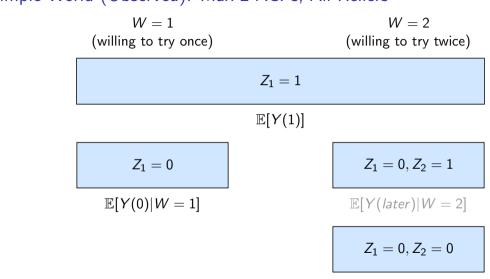




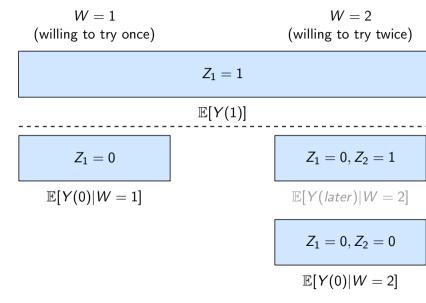
$$W=1$$
 (willing to try once)  $W=2$  (willing to try twice)  $Z_1=1$ 

$$Z_1=0$$

$$Z_1=0,Z_2=1$$



 $\mathbb{E}[Y(0)|W=2]$ 



$$W=1$$
 (willing to try once)  $W=2$  (willing to try twice)  $Z_1=1$   $\mathbb{E}[Y(1)]$   $Z_1=0$   $Z_1=0$   $Z_1=0$   $\mathbb{E}[Y(0)|W=1]$   $\mathbb{E}[Y(ater)|W=2]$   $Z_1=0,Z_2=0$   $\mathbb{E}[Y(0)|W=2]$ 

$$R=1$$
 (no child if fail)  $R=0$  (child if fail)

$$R=1$$
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$$Z_1=1$$

$$R=1$$
 (no child if fail)

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 (child if fail)

$$Z_1=1$$

$$Z_1=0,D=0$$

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$$R=1$$
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Distribution of 
$$Y(1)$$

$$Z_1=0,D=0$$

$$\mathbb{E}[Y(0)|R=1]$$

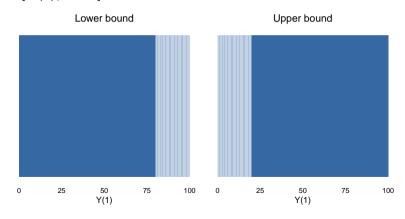
$$Z_1=0,D=1$$

$$\mathbb{E}[Y(later)|R=0]$$

$$Pr(R=1) =$$

### Intuition: Motherhood Outcome Y(1)

- 1. Treated group is a representative sample but their types are unobserved
- 2. Identify Pr(R = 1) = 0.8 on control group
- 3. Assume most extreme distributions of types in treated group
- 4. Bound  $\mathbb{E}[Y(1)|R=1]$



#### Technical Details

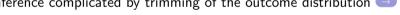
#### Formal identification

- Covariate-conditional sequential unconfoundedness
- Combine the two steps in a semi-parametric moment equation

Using covariates to narrow the bounds

► The bounds are sharp

Inference complicated by trimming of the outcome distribution



- ▶ Build on a double/debiased machine learning approach by Semenova (2023)
- Construct Neyman-orthogonal moment functions that are robust to first-stage nonparametric estimation errors in quantile and other nuisance functions

### Background and Data

Assisted conception procedures

- ▶ Intrauterine insemination: direct sperm injection, minimally invasive, free
- ▶ In-vitro fertilization: invasive medical procedure, first 3 free

Dutch family policies similar to OECD average

- ▶ 16 weeks maternity + pregnancy leave, 1 week paternity leave
- ▶ Average net childcare cost 10% median household income
- Part-time work more common, average work hours similar to OECD average

Novel dataset combining Dutch hospital medical records on ACPs with tax records



- ▶ Data on treatment dates and types; success imputed as birth within 10 months without additional ACPs
- Work hours and income include leave; results for hours corrected for uncertainty
- ▶ 15,523 cohabiting opposite-sex couples undergoing intrauterine insemination for their first child between 2013 and 2016
- ▶ Balance: ACP success at each attempt uncorr, with past outcomes cond, on age

#### Results: Bounds

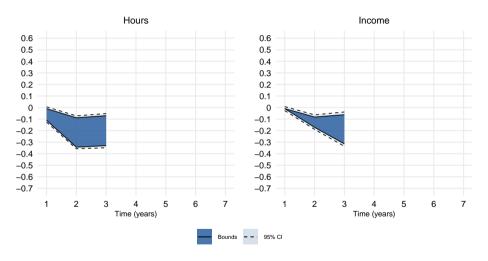


Figure 1: Bounds for Effects on Women Each Year Since Parenthood (% Relative to Childless)

#### Results: Bounds

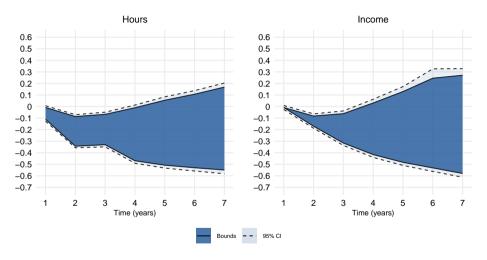


Figure 1: Bounds for Effects on Women Each Year Since Parenthood (% Relative to Childless)

## Narrowing the Bounds Further

Use additional information which on mother are reliers:

- Some women have non-ACP children after ACP succeeds
- ▶ May be reasonable to assume they are not reliant on ACPs
- Consistent with being determined to have at least one child
- ▶ Reduces uncertainty around which women are reliers



## Bounds with Monotonicity



Figure 2: Bounds for Women Under Monotonicity

#### Bounds for Men



Figure 3: Bounds fo Men

## **Gender Inequality**

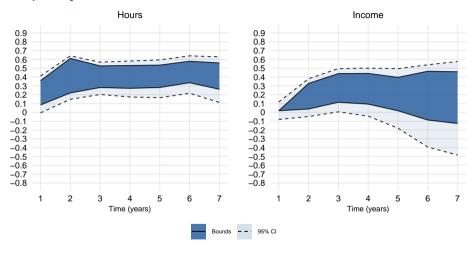


Figure 4: Share of Gender Inequality Caused by Parenthood (Effect on Gender Gap relative to Gap After Parenthood)

#### Extensions

Bias in leading methods Naive comp. Less naive comp. Formal procedure Bias in IV Bias in ES Mental health side effects Discussion Bounds for non-depresses Relation to methodological literature Theoretical comparison Results Confidence inveral comparison Confidence intervals Inequality correcting for age De-aging partners Stable complier group Childless final period Estimator without DMI Estimates Monotonicity Discussion Direction Partnered only Partnership and depression Test Testing Bensnes et al. (2023); Gallen et al. (2023) Estimates Heterogeneity Willingness to try Population imputation\* ES pop. Mother. imp. Childless imp. Effect imp.

#### Conclusion

Method for evaluating the career cost of parenthood:

- First evidence robust to selective fertility and dynamic effects
- ▶ Applicable to various settings with sequential treatment assignment and selection:
  - Assignment to job training and educational programs, legal settings with assignment to varying leniency "judges", promotion tournaments, clinical trials in extension phase
  - ▶ Also works with two-sided non-compliance and without sequential randomization

#### Application to Dutch data:

- Motherhood reduces work hours and income by 9% to 24%
- Parenthood causes up to 50% of post-child gender inequality

I show that accounting for bias can reconcile conflicting findings in literature:

- Naive: IV might understate the role of parenthood in gender ineq., ES overstates it
- ▶ This also holds after after accounting for differences in population and effect definition

#### External relevance:

- Alternative methods give very similar results on ACP sample and representative sample
- ▶ My method can be used to validate alternatives (I propose imputation; could go structural)

#### Policy:

Large share of gender inequality may not be due to parenthood per se

# Instrumental Variable vs Event Study: Percent Reduction in Earnings



Source: (Lundborg et al., 2024)

▶ "Naive" comparison with differing sub-populations and treatment definitions





#### Theoretical Bias



Figure 5: Motherhood among unsuccessfully treated  $\tau_{RF} = 0.25 \tau_{Parenthood} - 0.75 \tau_{Delay}$   $\tau_{IV} = \tau_{Parenthood} - 3 \tau_{Delay}$ 

Back (for illustration restricting heterogeneity between individuals)

## Bounding $\tau_{ATR}$

Construct the moment:

$$m^{L}(G, \eta^{0}) = Y1_{\{Y < q(r(X_{1}), X_{1})\}} \frac{Z_{1}}{e_{1}(X_{1})} - Y(1 - D) \prod_{i=1}^{A} \frac{(1 - Z_{i})}{(1 - e_{i}(X_{i}))}$$

- ► G is the observed data vector
- $\triangleright \eta^0$  contains the following:

  - $q(r(X_1), X_1)$  is the  $r(X_1)$ -th quantile of Y given  $X_1$  and  $Z_1 = 1$  $r(X_1)$  identifies the covariate-conditional relier share

# Assumption (Conditional Sequential Unconfoundedness)

 $(Y(k), R, W) \perp \!\!\! \perp Z_j \mid X_j \text{ for all } j, k, \text{ and } X_j, A \geq j.$ 

#### Theorem (Lower Bound)

Under conditional sequential unconfoundedness and regularity, the sharp lower bound on  $\tau_{ATR}$  is  $\mathbb{E}[m^L(G,\eta^0)]/\mathbb{E}[r(X_1)]$ .

## Intuition: Motherhood Outcome Y(1)—Covariates

Pre-ACP covariates can help narrow the bounds:

- ► Can identify relier share at each covariate value
- Baseline bounds assume extreme scenarios where reliers have highest or lowest treated outcomes
- These distributions of treated outcomes might be inconsistent with conditional relier shares



# Estimating the Bounds

Distribution of  $m^L(G, \eta^0)$  is complicated by  $q(r(X_1), X_1)$ 

- ► Semenova (2023) addresses a closely related inference challenge
- Double/debiased machine learning approach
- 1. Adjust  $m^L(G,\eta^0)$  to make it insensitive to small error in  $q(r(X_1),X_1)$
- Asymptotic inference as if  $q(r(X_1), X_1)$  was known

New moment:

$$\psi^L(G,\xi^0) = m^L(G,\eta^0) + corr(G,\xi^0)$$

Identifies same parameter:

2. Sample splitting

$$\mathbb{E}[\psi^L(G,\xi^0)] = \mathbb{E}[m^L(G,\eta^0)]$$

Insensitive to estimation error in  $q(r(X_1), X_1)$ :

$$\partial_{\sigma(\cdot)} \mathbb{E}[\psi^{L+}(G,\xi_r)|X_1]|_{\xi_r=\xi^0} = 0$$
 a.s.

## Assisted Conception Procedures

- ▶ IUI (main procedure): sperm injected into uterus
  - ▶ Minimally invasive, primary ACP in most countries
  - "Free" in NL
- ▶ IVF (secondary procedure): embryo inserted into uterus
  - Invasive treatment, performed under sedation/anesthesia
  - Eggs retrieved through the vaginal wall using a specialized needle
  - ▶ In NL, first 3 free; each subsequent costs between 1000 and 4000 EUR



#### Institutions

- Dutch family friendly polices similar to OECD average
  - ▶ 16 weeks of fully paid pregnancy+maternity leave
  - ▶ 1 week of paternity leave
  - Average time in child care similar to OECD average
  - ▶ Net child care cost 10% median household income
- Dutch employment intensity similar to OECD average
  - Employment among parents and non-parents relatively high
  - Part time work much more common
  - ▶ Approximately 15% two-parent families have both partners working part-time

Back

#### Data

#### Administrative data from Statistics Netherlands

- Comprehensive hospital records cover fertility treatments from 2012 to 2017: procedure date and type
  - Success imputed as having child born within 10 months
- ► Tax records cover work hours and income from 2011 to 2023
  - Include maternity leave and pay
  - Main bounds account for uncertainty around actual work hours
- Birth dates, legal family connections, cohabitation
- Dispensed medication registry

Main sample: cohabiting opposite-sex couples undergoing IUI for their first child between 2013 and 2016: 15,523

Back

ToC

## Overview of Descriptives

- ► First and subsequent ACP success uncorrelated with past labor market outcomes condiditional on age Table first Table later
  - ▶ Support for independence of  $Z_j$  and (Y(1), Y(0))
- Success probability stable across ACPs conditional on age Figure
  - ▶ Support for independence of  $Z_j$  and W
- Representative sample worked less and had lower income before parenthood, but differences relatively small Table
  - ACP sample older before parenthood

Back (summary)

#### Balance in 1st ACP

Table 1: First ACP Outcomes and Descriptives

	Success (1)	Fail (2)	Difference (1)-(2)	Dif. cond. age & educ (1)-(2) cond.
Work (W)	0.882	0.863	0.019	0.008
	[0.323]	[0.344]	(0.009)	(0.009)
Work (P)	0.884	0.865	0.019	0.013
, ,	[0.320]	[0.342]	(0.009)	(0.009)
Hours (W)	1240.315	1207.860	32.455	18.702
	[604.666]	[635.194]	(16.183)	(16.560)
Hours (P)	1474.530	1438.590	35.940	18.579
	[658.231]	[695.692]	(17.713)	(17.870)
Income 1000s € (W)	28.065	27.418	0.647	0.745
	[19.559]	[20.219]	(0.516)	(0.546)
Income 1000s € (P)	37.205	36.952	0.252	0.364
	[26.482]	[29.452]	(0.746)	(0.730)
Bachelor deg. (W)	0.480	0.451	0.029	
	[0.500]	[0.498]	(0.013)	
Bachelor deg. (P)	0.394	0.381	0.013	
	[0.489]	[0.486]	(0.012)	
Age (W)	31.638	32.388	-0.750	
	[4.015]	[4.383]	(0.111)	
Age (P)	34.675	35.461	-0.786	
	[5.513]	[5.996]	(0.152)	
Observations	1,714	13,809		
Joint p-val.			0.000	0.928

Note: Labor market outcomes measured year before first ACP. (W) - woman, (P) - partner. Last column uses inverse prbability weights for the first ACP that follow the main specificaition. Standard deviations in brackets. Standard errors in parentheses.

## Balance in Subsequent ACPs

Table 2: Balance in Later ACPs

	$Z_2$	$Z_3$	$Z_4$	$Z_5$	$Z_6$	<b>Z</b> 7	$Z_8$	$Z_9$	$Z_{10}$
Work (W)	0.009	-0.004	0.022	0.014	0.039	-0.003	-0.011	0.022	0.030
	(0.010)	(0.011)	(0.011)	(0.012)	(0.012)	(0.017)	(0.018)	(0.019)	(0.024)
Work (P)	0.006	0.016	0.012	0.020	-0.004	-0.004	-0.019	0.017	0.030
	(0.010)	(0.010)	(0.012)	(0.012)	(0.015)	(0.015)	(0.019)	(0.020)	(0.027)
Hours (W)	32.885	-4.482	52.999	41.332	81.957	11.894	-18.836	72.659	24.819
	(18.721)	(20.032)	(21.045)	(22.686)	(25.131)	(31.187)	(32.937)	(38.210)	(48.490
Hours (P)	21.655	24.730	23.756	38.965	9.666	-6.580	-28.458	30.525	43.722
	(21.018)	(21.089)	(23.574)	(25.255)	(30.585)	(31.513)	(37.976)	(44.856)	(52.821
Income 1000s € (W)	1.481	-0.015	1.685	1.802	2.086	0.150	-0.043	0.866	-0.444
	(0.615)	(0.624)	(0.767)	(0.830)	(0.913)	(1.000)	(1.092)	(1.234)	(1.629)
Income 1000s € (P)	-0.749	1.002	2.040	0.800	0.774	0.025	0.259	-0.324	0.149
	(0.835)	(0.912)	(1.066)	(1.115)	(1.424)	(1.424)	(1.563)	(1.737)	(2.203)
Observations	12,974	10,774	8,726	6,977	5,411	3,944	2,723	1,850	1,174
Joint <i>p</i> -val.	0.175	0.976	0.234	0.303	0.140	1.000	0.956	0.704	0.917

Note: Each column describes the difference in average characteristics between women for whom the respective ACP succeeds and those for whom it fails, among those who undergo the procedure, using inverse probability weights for each ACP following the main specification. Labor market outcomes and age measured year before first treatment. (W) - woman, (P) - partner. Standard errors in parentheses.

ACP histories Back (summary) Back (detailed descriptives)

#### **Estimated Success Probabilities**



Figure 6: Estimated Success Probabilities

## Comparison to Representative Sample

Table 3: Full Sample, Reliers, and Representative Sample

	Success (1)	Fail (2)	Reliers (3)	Rep. (4)	Success vs rep. (1)-(4)	Rel. vs rep (3)-(4)
Work (W)	0.882	0.863	0.820	0.801	0.080	0.019
	[0.323]	[0.344]	[0.333]	[0.399]	(0.010)	(0.005)
Work (P)	0.884	0.865	0.849	0.783	0.101	0.066
	[0.320]	[0.342]	[0.344]	[0.412]	(0.010)	(0.005)
Hours (W)	1240.315	1207.860	1117.711	1076.204	164.111	41.508
	[604.666]	[635.194]	[582.334]	[696.245]	(16.856)	(8.412)
Hours (P)	1474.530	1438.590	1390.699	1250.948	223.582	139.752
	[658.231]	[695.692]	[662.920]	[793.536]	(19.211)	(9.576)
Income 1000s € (W)	28.065	27.418	24.976	21.362	6.703	3.615
	[19.559]	[20.219]	[15.359]	[18.330]	(0.444)	(0.222)
Income 1000s € (P)	37.205	36.952	35.299	28.107	9.098	7.193
	[26.482]	[29.452]	[24.304]	[29.076]	(0.704)	(0.351)
Bachelor deg. (W)	0.480	0.451	0.398	0.411	0.069	-0.012
	[0.500]	[0.498]	[0.411]	[0.492]	(0.012)	(0.006)
Bachelor deg. (P)	0.394	0.381	0.329	0.345	0.049	-0.015
	[0.489]	[0.486]	[0.397]	[0.475]	(0.012)	(0.006)
Age (W)	31.638	32.388	33.480	28.713	2.926	4.767
	[4.015]	[4.383]	[3.897]	[4.658]	(0.113)	(0.056)
Age (P)	34.675	35.461	36.580	28.713	5.962	7.868
	[5.513]	[5.996]	[3.928]	[4.665]	(0.113)	(0.057)
Observations	1,714	13,809	4,882	376,152		

Note: Labor market outcomes measured year before first ACP for main smple and year and 9 months before brith of first child for the representative sample is expected to match the main sample by year of conception. Average relier outcomes are based on sample of women who remain childless 7 years after their first ACP with weights described under implementation. (W) - woman, (P) - partner. Standard deviations in brackets. Standard errors in parentheses.

#### **ACP Histories**



Figure 7: ACP Histories and Reliance

#### What are the counterfactuals?

#### Broadly:

- Do not want/plan children
- ▶ Want/plan children

#### Motherhood outcome:

- Get immediately
- Get naturally after few attempts
- Get with medical assistance

#### Childless outcome:

- Do not try
- ► Try and fail naturally
- Try and fail with medical assistance (+ naturally?)

Extrapolation requires carefully addressing mental health consequences of failure (and medical procedures)

Back (model) Back (extensions) Depr. effect Bounds non-depr. Arguments

## Relation to Methodological Literature



## Comparison with Lee (2009)



Figure 8: Comparison with Lee (2009) Bounds for Effects on Women



# Less Naive Comparison to Existing Methods

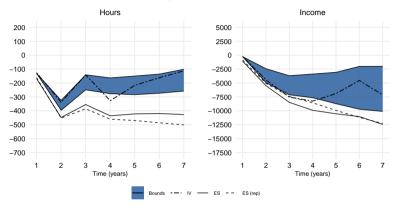


Figure 9: Estimates Using on Different Methods (Absolute Effects)

- ▶ Using women whose ACP succeeds for ES makes treatment definition consistent
- ▶ The three methods still target different sub-populations



# Quantifying Bias in Existing Methods

Instrumental Variable (Lundborg et al., 2017):

- ▶ I have bounded  $\tau_{ATR}$  allowing for dynamic effects
- lacktriangle I can point-identify  $au_{ATR}$  assuming static effects à la instrumental variable
- lacktriangle Subtracting it from bounds on  $au_{ATR}$  gives bounds on bias

Event Study (Kleven et al., 2019):

- ▶ ES imputes childless career trajectories from pre-parenthood outcomes of later mothers
- I proxy when women chose to have children with ACP timing
- Relier childless trajectories from my method vs imputed à la event study from older reliers
- Comparing the two identifies the bias
- **Same population and treatment definition as**  $\tau_{ATR}$  allows for comparison



## Effect of Delaying Motherhood

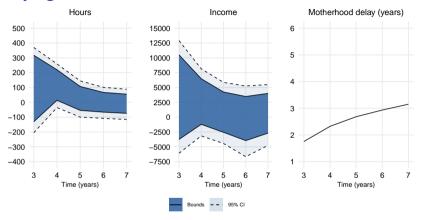


Figure 10: Effect of Delaying Relative to Motherhood at First Attempt (Absolute)

$$au_{IV} = au_{Parenthood} - 3 au_{Delay}$$



#### Placebo Event

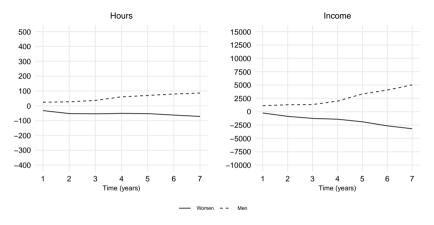


Figure 11: Placebo Event Study (Absolute)

▶ Negative selection of early mothers and positive selection of fathers



## Gender Inquality: Parenthood and Selection

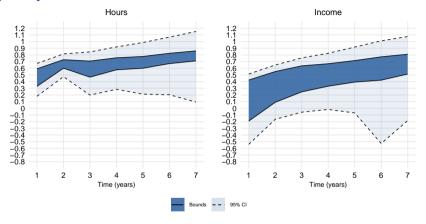


Figure 12: Share of Gender Inequality Explained by Section and Parenthood

▶ Consistent with ES estimates attributing almost all gender inequality to parenthood



#### Mental Health and ACPs

Mental health consequences associated with failure to conceive are a part of the story:

 Unmet fertility goals may negatively impact mental health, and in turn, labor market outcomes

There are, however, additional concerns:

- ▶ Mental health issues caused specifically by failed conception or ACPs (external)
  - Focusing on artificial insemination helps mitigate this
- Large impacts unique to ACP families (external)
- Worsened mental health by threatening monotonicity (internal)

In practice, these impacts are likely small (Lundborg et al., 2024)

Antidepressant uptake Back (extensions) Conclusion

### Monotone Bounds for Non-depressed Childless Women



Figure 13: Monotone Bounds for Women Who Would Not Uptake Antidepressants if They Were to Remain Childless



#### Confidence Intervals



Figure 14: 95% CI for Different Methods (Absolute)

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### Monotonicity

#### Is monotonicity realistic?

- Yes, if families are determined to have at least one child
  - Decreasing marginal returns to children.
  - Stronger sufficient assumption: success cannot increase natural births
- No, if first treatment success increases the likelihood of attempting to conceive naturally
  - Couples may realize they are fertile and try more
  - First child may "save the relationship" resulting in more attempts to conceive



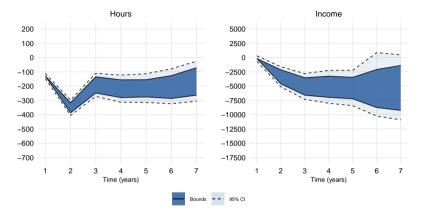
#### Monotone Bounds: Women who Remain Childless



Figure 15: Monotone Bounds Using Completed Fertility (Absolute)

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#### Simple estimator



- $ightharpoonup \mathbb{E}[g(X_1)|R=1]$  identified on chillness reliers using baseline method
- ▶ Only need to bound  $\mathbb{E}[\varepsilon|R=1]$

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### Relaxing Monotonicity Direction





ToC

### Relaxing Monotonicity to Partnered Women

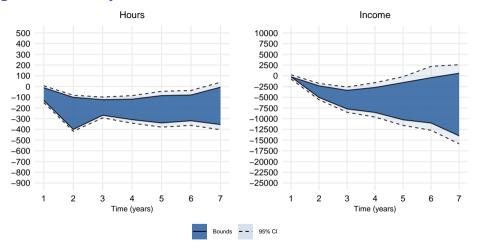


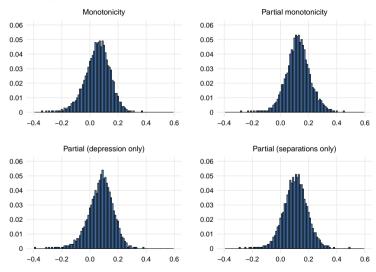
Figure 16: Monotone Bounds Using Women Who Stay Partnered

## Relaxing Monotonicity for Depression and Partnership



Figure 17: Monotone Bounds For Women Who Stay Partnered and Do Not Uptake Antidepress.

# **Testing Monotonicity**



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### Heterogeneity by Willingness to Undergo Procedures



Figure 18: Cumulative Outcomes 6 Years After, G Above or Below 6



## Monotone Bounds: Excluding Depression



Figure 19: Effect on Antidepressant Take-Up

# Monotone Bounds: Correcting for Partner's age



Figure 20: Monotone Bounds Using Male Income at Same Age as Female

# Testing the Plug-in Approach

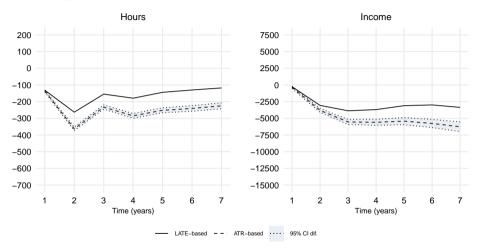


Figure 21: Plug-In Estimators Exploiting Different Numbers of Treatments



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# Event Study: Population vs IUI Sample



Figure 22: ES for Population and Women with First IUI Success

## Imputing Population Motherhood Outcomes Using IUI Sample

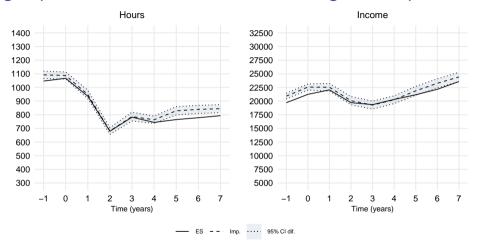


Figure 23: Population Outcomes vs IUI-Imputation (Age & Education)

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### Imputing Population Childless Outcomes Using IUI Sample



Figure 24: Population Outcomes vs IUI-Imputation (Age & Education)

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# Event Study vs IUI-imputation for Population



Figure 25: Event Study vs IUI-Imputation for Population (Age & Education)

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