

Binomial	$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$	E(X) = np, Var(X) = np(1-p)
Poisson	$P(X = x) = \frac{\lambda^x e^{-\lambda}}{x!}$	E(X) = λ, Var(X) = λ
Normal	$f(x) = \left(\frac{1}{\sigma\sqrt{2\pi}}\right) e^{-\left(\frac{1}{2}\right)\left(\frac{x-\mu}{\sigma}\right)^2}$	E(X) = μ, Var(X) = σ <sup>2</sup>
Exponential	$f(x) = \lambda e^{-\lambda x}$ for $x \geq 0$	E(X) = $\frac{1}{\lambda}$ , Var(X) = $\frac{1}{\lambda^2}$
Uniform	$f(x) = \frac{1}{b-a}$ for $a \leq x \leq b$	E(X) = $\frac{a+b}{2}$ , Var(X) = $\frac{(b-a)^2}{12}$
Bernoulli	$P(X = 1) = p$ $P(X = 0) = 1 - p$	E(X) = p, Var(X) = p(1 - p)
Geometric	$P(X = x) = (1 - p)^{x-1} p$	E(X) = $\frac{1}{p}$ , Var(X) = $\frac{1-p}{p^2}$
Negative Binomial	$P(X = x) = \binom{x-1}{r-1} p^r (1 - p)^{x-r}$	E(X) = $\frac{r}{p}$ , Var(X) = $\frac{r(1-p)}{p^2}$
Hypergeometric	$P(X = x) = \frac{\binom{K}{x} \binom{N-K}{n-x}}{\binom{N}{n}}$	E(X) = $n \left(\frac{K}{N}\right)$ , Var(X) = $n \left(\frac{K}{N}\right) \left(1 - \frac{K}{N}\right) \left(\frac{N-n}{N-1}\right)$

- MME: equate sample mean to population mean
- 1st moment:  $\bar{X} = E(X)$
- 2nd moment:  $\left(\frac{1}{n}\right) \sum_{i=1}^n (X_i^2) = E(X^2)$
- MLE: maximise the likelihood function(for n independent observations:  $L(\theta) = \prod_{i=1}^n f(X_i; \theta)$ ). Also show the maxima using the 2nd derivative test.

	$H_0$ true	$H_0$ false
Fail to reject $H_0$ (i.e. -ve test)	Correct $1 - \alpha$	Type II error
Reject $H_0$ (i.e. +ve test)	Type I error $\alpha$	Correct Power of Test

1. HT for population mean

**Use when:** Testing if a population mean differs from a hypothesized value (use z if σ known, t if unknown); e.g. has the avg rate gone up (given μ) - this will b a one tailed test. e.g. one-taild test

- $H_0 : \mu = \mu_0$  |  $H_1 : \mu > \mu_0$
- calculate sample mean  $\bar{x}$  and sample standard deviation  $s_x$
- calculate test statistic:  $t_{\text{calc}} = \frac{\bar{x} - \mu_0}{\frac{s_x}{\sqrt{n}}}$  with df = n-1 (or use z if σ known)
- determine critical value from t-table (or z-table) at significance level α
- compare  $t_{\text{calc}}$  with  $t_{\text{crit}}$
- If  $t_{\text{calc}} > t_{\text{crit}}$  (for right-tailed test), reject  $H_0$ ; If  $t_{\text{calc}} < -t_{\text{crit}}$  (for left-tailed test), reject  $H_0$

Comparison of one-tailed tests:

- **Right-tailed** ( $H_1 : \mu > \mu_0$ ): Used when testing for an **increase** (e.g., “has performance improved?”). Reject if  $t_{\text{calc}} > t_{\text{crit}}$  (positive critical value)
- **Left-tailed** ( $H_1 : \mu < \mu_0$ ): Used when testing for a **decrease** (e.g., “has cost reduced?”). Reject if  $t_{\text{calc}} < -t_{\text{crit}}$  (negative critical value)

2. HT for population proportion

**Use when:** Testing if a population proportion differs from a hypothesized value; e.g., testing if the proportion of car owners in a region differs from a claimed percentage.

- State hypotheses:  $H_0 : P = p_0$  (population proportion equals hypothesized value);  $H_1$ : (choose:  $P \neq p_0$  or  $P > p_0$  or  $P < p_0$ )
- Calculate sample proportion:  $\hat{p} = \frac{x}{n}$  where  $x$  is number of successes and  $n$  is sample size
- Check conditions:  $np_0 \geq 10$  and  $n(1 - p_0) \geq 10$  for normal approximation validity
- Specify significance level α (typically 0.05) and calculate test statistic:  $z_{\text{calc}} = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$
- Determine critical value(s) from z-table: Two-tailed:  $\pm z_{\text{crit}}$  (e.g.,  $\pm 1.96$  for α = 0.05); Right-tailed:  $z_{\text{crit}}$  (e.g., 1.645 for α = 0.05); Left-tailed:  $-z_{\text{crit}}$  (e.g., -1.645 for α = 0.05)
- Make decision: Two-tailed: If  $|z_{\text{calc}}| > z_{\text{crit}}$ , reject  $H_0$ ; Right-tailed: If  $z_{\text{calc}} > z_{\text{crit}}$ , reject  $H_0$ ; Left-tailed: If  $z_{\text{calc}} < -z_{\text{crit}}$ , reject  $H_0$

3. HT for population variance

**Use when:** Testing if a population variance differs from a hypothesized value, or when you need to verify assumptions about data variability (e.g., quality control - has the variance in product weights changed?).

- $H_0 : \sigma^2 = \sigma_0^2$  (population variance equals hypothesized value);  $H_1$ : (choose:  $\sigma^2 \neq \sigma_0^2$  or  $\sigma^2 > \sigma_0^2$  or  $\sigma^2 < \sigma_0^2$ )
- Calculate sample variance:  $s^2 = \left(\frac{1}{n-1}\right) \sum_{i=1}^n (X_i - \bar{X})^2$
- Specify significance level α (typically 0.05) and calculate test statistic:  $\chi_{\text{calc}}^2 = \frac{(n-1)s^2}{\sigma_0^2}$  with df = n-1
- Determine critical value(s) from  $\chi^2$  table: Two-tailed:  $\chi_{\text{lower}}^2$  and  $\chi_{\text{upper}}^2$  (e.g., at  $\frac{\alpha}{2}$  and  $1 - \frac{\alpha}{2}$ ); Right-tailed:  $\chi_{\text{crit}}^2$  at  $1 - \alpha$ ; Left-tailed:  $\chi_{\text{crit}}^2$  at α
- Make decision: Two-tailed: If  $\chi_{\text{calc}}^2 < \chi_{\text{lower}}^2$  or  $\chi_{\text{calc}}^2 > \chi_{\text{upper}}^2$ , reject  $H_0$ ; Right-tailed: If  $\chi_{\text{calc}}^2 > \chi_{\text{crit}}^2$ , reject  $H_0$ ; Left-tailed: If  $\chi_{\text{calc}}^2 < \chi_{\text{crit}}^2$ , reject  $H_0$

4. 2 sample test + paired t test

**Use when:** Testing if the mean difference between paired observations is zero; e.g., before-after measurements on the same subjects (drug effectiveness, training impact), matched pairs (father-son heights, twin studies).

- Calculate differences:  $d_i = x_i - y_i$  for each paired observation
- Calculate sample mean of differences:  $\bar{d} = \frac{\sum_{i=1}^n d_i}{n}$
- Calculate sample standard deviation of differences:  $s_d = \sqrt{\frac{\sum_{i=1}^n (d_i - \bar{d})^2}{n-1}}$
- State hypotheses:  $H_0 : \mu_d = 0$  (no difference/drug has no effect);  $H_1$ : (choose:  $\mu_d \neq 0$  or  $\mu_d > 0$  (i.e. drug improves score) or  $\mu_d < 0$ )
- Specify significance level α (typically 0.05) and calculate test statistic:  $t_{\text{calc}} = \frac{\bar{d} - \mu_{d0}}{\frac{s_d}{\sqrt{n}}}$  with df = n-1
- Determine critical value from t-table at chosen α and df

- Make decision: Two-tailed: If  $|t_{\text{calc}}| > t_{\text{crit}}$ , reject  $H_0$ ; Right-tailed: If  $t_{\text{calc}} > t_{\text{crit}}$ , reject  $H_0$ ; Left-tailed: If  $t_{\text{calc}} < -t_{\text{crit}}$ , reject  $H_0$

5. 2 sample test + comparing mean

**Use when:** Testing if the means of two independent populations differ; e.g., comparing average scores between two groups, testing if mean salaries differ between two departments, comparing average heights between two different populations; e.g. TGA and TGB

(pooled variance - when σ<sup>2</sup> known to be equal)

- State hypotheses:  $H_0 : \mu_1 = \mu_2$  or  $H_0 : \mu_1 - \mu_2 = 0$ ;  $H_1$ : (choose:  $\mu_1 \neq \mu_2$  or  $\mu_1 > \mu_2$  or  $\mu_1 < \mu_2$ )
- Calculate pooled variance:  $s_p^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1 + n_2 - 2}$
- Calculate test statistic:  $t_{\text{calc}} = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{s_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$  with df =  $n_1 + n_2 - 2$

- Determine critical value from t-table at significance level α and df
- Make decision: Two-tailed: If  $|t_{\text{calc}}| > t_{\text{crit}}$ , reject  $H_0$ ; Right-tailed: If  $t_{\text{calc}} > t_{\text{crit}}$ , reject  $H_0$ ; Left-tailed: If  $t_{\text{calc}} < -t_{\text{crit}}$ , reject  $H_0$

(pooled variance) CI =  $\left( (\bar{x}_1 - \bar{x}_2) \pm t_{\text{crit}} * \sqrt{s_p^2 * \left(\frac{1}{n_1} + \frac{1}{n_2}\right)} \right)$

(unpooled variance - Welch’s t-test - when σ<sup>2</sup> not known to be equal)

- State hypotheses:  $H_0 : \mu_1 = \mu_2$ ;  $H_1$ : (choose:  $\mu_1 \neq \mu_2$  or  $\mu_1 > \mu_2$  or  $\mu_1 < \mu_2$ )
- Calculate test statistic:  $t_{\text{calc}} = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\left(\frac{s_1^2}{n_1}\right) + \left(\frac{s_2^2}{n_2}\right)}}$
- Calculate degrees of freedom:  $df = \frac{\left(\frac{s_1^2}{n_1}\right)^2 + \left(\frac{s_2^2}{n_2}\right)^2}{\frac{\left(\frac{s_1^2}{n_1}\right)^2}{n_1-1} + \frac{\left(\frac{s_2^2}{n_2}\right)^2}{n_2-1}}$

- Determine critical value from t-table at significance level α and calculated df
- Make decision: Two-tailed: If  $|t_{\text{calc}}| > t_{\text{crit}}$ , reject  $H_0$ ; Right-tailed: If  $t_{\text{calc}} > t_{\text{crit}}$ , reject  $H_0$ ; Left-tailed: If  $t_{\text{calc}} < -t_{\text{crit}}$ , reject  $H_0$

(unpooled variance) CI =  $\left( (\bar{x}_1 - \bar{x}_2) \pm t_{\text{crit}} * \sqrt{\left(\frac{s_1^2}{n_1}\right) + \left(\frac{s_2^2}{n_2}\right)} \right)$  and df

calculation for unpooled case:  $df = \frac{\left(\left(\frac{s_1^2}{n_1}\right)^2 + \left(\frac{s_2^2}{n_2}\right)^2\right)}{\left(\frac{\left(\frac{s_1^2}{n_1}\right)^2}{n_1-1}\right) + \left(\frac{\left(\frac{s_2^2}{n_2}\right)^2}{n_2-1}\right)}$

6. 2 sample test + comparing proportion

**Use when:** Testing if proportions differ between two independent groups; e.g., comparing preference rates, success rates, or occurrence rates between two populations; e.g. PP non-smoking rooms preference | e.g. the cardiac arrests during day and night number of deaths question | e.g. elem and high school teacher preference.

Decision Rule:

- Use **pooled proportion** for hypothesis testing (assumes  $p_1 = p_2$  under  $H_0$ )
- Use **unpooled proportion** for confidence intervals (no assumption about equality)

(unpooled proportion - for confidence intervals) CI =  $(\hat{p}_1 - \hat{p}_2) \pm z_{\text{crit}} * \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$  (pooled proportion - for hypothesis testing)

- State hypotheses:  $H_0 : p_1 = p_2$  or  $H_0 : p_1 - p_2 = 0$ ;  $H_1$ : (choose:  $p_1 \neq p_2$  or  $p_1 > p_2$  or  $p_1 < p_2$ )
- Calculate sample proportions:  $\hat{p}_1 = \frac{x_1}{n_1}$  and  $\hat{p}_2 = \frac{x_2}{n_2}$  and pooled proportion:  $\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$

- Calculate test statistic:  $z_{\text{calc}} = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$
- Determine critical value from z-table based on significance level  $\alpha$  and test type (one-tailed or two-tailed)
- Make decision: If  $|z_{\text{calc}}| > z_{\text{crit}}$  (two-tailed) or  $z_{\text{calc}} > z_{\text{crit}}$  (right-tailed) or  $z_{\text{calc}} < -z_{\text{crit}}$  (left-tailed), reject  $H_0$ .

### 7. 2 sample test + comparing variance

**Use when:** Testing if the variances of two independent populations differ; e.g., comparing variability in quality between two manufacturing processes, testing if variance in test scores differs between two teaching methods.

- State hypotheses:  $H_0 : \sigma_1^2 = \sigma_2^2$  or  $H_0 : \frac{\sigma_1^2}{\sigma_2^2} = 1$ ;  $H_1 : (\text{choose: } \sigma_1^2 \neq \sigma_2^2 \text{ or } \sigma_1^2 > \sigma_2^2 \text{ or } \sigma_1^2 < \sigma_2^2)$
- Calculate sample variances:  $s_1^2$  and  $s_2^2$  from each sample
- Specify significance level  $\alpha$  (typically 0.05)
- Calculate test statistic:  $F_{\text{calc}} = \frac{s_1^2}{s_2^2}$  (always put larger variance in numerator for one-tailed test, or follow hypothesis direction for two-tailed)
- Determine degrees of freedom:  $df_1 = n_1 - 1$  (numerator) and  $df_2 = n_2 - 1$  (denominator)
- Determine critical value(s) from F-distribution table: Two-tailed:  $F_{\text{lower}}$  at  $\frac{\alpha}{2}$  and  $F_{\text{upper}}$  at  $1 - \frac{\alpha}{2}$ ; Right-tailed:  $F_{\text{crit}}$  at  $1 - \alpha$ ; Left-tailed:  $F_{\text{crit}}$  at  $\alpha$
- Make decision: Two-tailed: If  $F_{\text{calc}} < F_{\text{lower}}$  or  $F_{\text{calc}} > F_{\text{upper}}$ , reject  $H_0$ ; Right-tailed: If  $F_{\text{calc}} > F_{\text{crit}}$ , reject  $H_0$ ; Left-tailed: If  $F_{\text{calc}} < \frac{1}{F_{\text{crit}}}$ , reject  $H_0$

**Note:** F-test is sensitive to non-normality. Consider Levene’s test or Bartlett’s test as alternatives when normality is questionable.

### 8. ANOVA

**Use when:** Testing if means of three or more independent groups differ; e.g., comparing average test scores across multiple teaching methods, testing if mean sales differ across different regions. **Note:** ANOVA assumes normality, independence, and equal variances across groups. If assumptions are violated, consider Kruskal-Wallis test.

- State hypotheses:**  $H_0 : \mu_1 = \mu_2 = \mu_3 = \dots = \mu_k$  (all group means are equal);  $H_1 : \text{at least one group mean differs from the others}$
- Calculate group means  $\bar{X}_i$  for each of the  $k$  groups and overall mean  $\bar{X}$
- Calculate Sum of Squares Between groups:  $SSB = \sum_{i=1}^k n_i (\bar{X}_i - \bar{X})^2$
- Calculate Sum of Squares Within groups:  $SSW = \sum_{i=1}^k \sum_{j=1}^{n_i} (X_{ij} - \bar{X}_i)^2$
- Calculate Mean Squares:  $MSB = \frac{SSB}{k-1}$  and  $MSW = \frac{SSW}{N-k}$  where  $N = \sum n_i$
- Calculate test statistic:  $F_{\text{calc}} = \frac{MSB}{MSW}$  and determine degrees of freedom:  $df_1 = k - 1$  (between groups) and  $df_2 = N - k$  (within groups)
- Specify significance level  $\alpha$  (typically 0.05) and find critical value  $F_{\text{crit}}$  from F-table
- Make decision: If  $F_{\text{calc}} > F_{\text{crit}}$ , reject  $H_0$  (at least one group mean differs)

**ANOVA Table:**

Source	df	Sum of Squares	Mean Square	F-statistic
Between Groups	$k - 1$	SSB	$MSB = \frac{SSB}{k-1}$	$F = \frac{MSB}{MSW}$
Within Groups	$N - k$	SSW	$MSW = \frac{SSW}{N-k}$	
Total	$N - 1$	SST		

where  $SST = SSB + SSW$

### 9. Bartlett’s Test for Homogeneity of Variances

**Use when:** Testing if variances are equal across multiple groups ( $k \geq 2$ ); e.g., comparing variability in yields across different fertilizer treatments (additives A, B, C). **Note:** Bartlett’s test is sensitive to non-normality. Consider Levene’s test as a more robust alternative.

- State hypotheses:**  $H_0 : \sigma_1^2 = \sigma_2^2 = \dots = \sigma_k^2$  (all population variances are equal);  $H_1 : \text{at least one variance differs}$ .
- Calculate pooled variance:**  $S_p^2 = \frac{\sum_{i=1}^k (n_i - 1)s_i^2}{N - k}$  where  $N = \sum_{i=1}^k n_i$  (total sample size)
- Calculate test statistic:**  $\chi_{\text{calc}}^2 = \frac{(N - k) \ln(S_p^2) - \sum_{i=1}^k (n_i - 1) \ln(s_i^2)}{C}$  where  $C = 1 + \left(\frac{1}{3(k-1)}\right) \left(\sum_{i=1}^k \left(\frac{1}{n_i - 1}\right) - \left(\frac{1}{N - k}\right)\right)$  and **Determine degrees of freedom:**  $df = k - 1$
- Specify significance level  $\alpha$**  (typically 0.05) and find critical value  $\chi_{\text{crit}}^2$  from  $\chi^2$  table
- Make decision:** If  $\chi_{\text{calc}}^2 > \chi_{\text{crit}}^2$ , reject  $H_0$  (variances are not equal)

### 10. 2 way ANOVA

**Use when:** Testing the effects of two categorical independent variables (factors) on a continuous dependent variable, including their interaction; e.g., studying effect of fertilizer type AND irrigation method on crop yield, or teaching method AND class size on test scores. **Assumptions:** Independent observations + Normally distributed residuals + Homogeneity of variance across groups + Balanced design (equal sample sizes per cell) preferred

- Hypotheses:** Factor A:  $H_0 : \alpha_1 = \alpha_2 = \dots = \alpha_r = 0$  (no effect of factor A);  $H_1 : \text{at least one } \alpha_i \neq 0$  | Factor B:  $H_0 : \beta_1 = \beta_2 = \dots = \beta_c = 0$  (no effect of factor B);  $H_1 : \text{at least one } \beta_j \neq 0$  | Interaction:  $H_0 : \text{no interaction between A and B}$ ;  $H_1 : \text{interaction exists}$
- Calculate cell means**  $\bar{X}_{ij}$ , row means  $\bar{X}_{i.}$ , column means  $\bar{X}_{.j}$ , and grand mean  $\bar{X}$
- Calculate Sum of Squares:**
  - Total:  $SST = \sum_{i=1}^r \sum_{j=1}^c \sum_{k=1}^n (X_{ijk} - \bar{X})^2$
  - Factor A:  $SS_A = cn \sum_{i=1}^r (\bar{X}_{i.} - \bar{X})^2$
  - Factor B:  $SS_B = rn \sum_{j=1}^c (\bar{X}_{.j} - \bar{X})^2$
  - Interaction:  $SS_{AB} = n \sum_{i=1}^r \sum_{j=1}^c (\bar{X}_{ij} - \bar{X}_{i.} - \bar{X}_{.j} + \bar{X})^2$
  - Error:  $SSE = \sum_{i=1}^r \sum_{j=1}^c \sum_{k=1}^n (X_{ijk} - \bar{X}_{ij})^2$
- Calculate Mean Squares:**
  - $MS_A = \frac{SS_A}{r-1}$
  - $MS_B = \frac{SS_B}{c-1}$
  - $MS_{AB} = \frac{SS_{AB}}{(r-1)(c-1)}$
  - $MSE = \frac{SSE}{rc(n-1)}$
- Calculate F-statistics:**
  - Factor A:  $F_A = \frac{MS_A}{MSE}$  with  $df = (r - 1, rc(n - 1))$
  - Factor B:  $F_B = \frac{MS_B}{MSE}$  with  $df = (c - 1, rc(n - 1))$
  - Interaction:  $F_{AB} = \frac{MS_{AB}}{MSE}$  with  $df = ((r - 1)(c - 1), rc(n - 1))$
- Specify significance level  $\alpha$**  (typically 0.05) and **Find critical values** from F-table for each test
- Make decisions:** For each test, if  $F_{\text{calc}} > F_{\text{crit}}$ , reject  $H_0$

**Two-Way ANOVA Table:**

Source	df	Sum of Squares	Mean Square	F-statistic
Factor A	$r - 1$	$SS_A$	$MS_A$	$F_A = \frac{MS_A}{MSE}$

Factor B	$c - 1$	$SS_B$	$MS_B$	$F_B = \frac{MS_B}{MSE}$
Interaction A×B	$(r - 1)(c - 1)$	$SS_{AB}$	$MS_{AB}$	$F_{AB} = \frac{MS_{AB}}{MSE}$
Error	$rc(n - 1)$	SSE	MSE	
Total	$rcn - 1$	SST		

where  $r$  = number of levels of factor A,  $c$  = number of levels of factor B,  $n$  = number of observations per cell.

**Note:** If interaction is significant, interpret main effects with caution as the effect of one factor depends on the level of the other factor.

### 11. $\chi^2$ test + goodness of fit

**Use when:** Testing if observed categorical data follows a specified theoretical distribution; e.g., testing if die rolls are fair, if color preferences match expected proportions, or if genotypes follow Mendelian ratios.

- e.g. rock paper scissor
- State hypotheses:  $H_0 : \text{the observed frequencies fit the expected frequencies (data follows the specified distribution)}$ ;  $H_1 : \text{they do not fit}$
- Calculate expected counts:  $E_i = \text{Total} \times \text{Probability}_i$  for each category
- Check condition: Ensure all expected frequencies  $E_i \geq 5$  for validity
- Specify significance level  $\alpha$  (typically 0.05)
- Calculate test statistic:  $\chi_{\text{calc}}^2 = \sum_{i=1}^c \left(\frac{(O_i - E_i)^2}{E_i}\right)$  where  $O_i$  are observed frequencies
- Determine degrees of freedom:  $df = c - 1$  where  $c$  is the number of categories
- Find critical value  $\chi_{\text{crit}}^2$  from  $\chi^2$  table at chosen  $\alpha$  and  $df$
- Make decision: If  $\chi_{\text{calc}}^2 > \chi_{\text{crit}}^2$ , reject  $H_0$  (observed data does not fit expected distribution). If  $\chi_{\text{calc}}^2 \leq \chi_{\text{crit}}^2$ , fail to reject  $H_0$  (insufficient evidence against the expected distribution)

### 12. $\chi^2$ test + test of independence

- e.g. illegal piracy law and attitude | e.g. accident and seatbelt usage
- $H_0 : \text{the two variables are independent (no association)}$ ;  $H_1 : \text{the two variables are not independent (there is an association)}$
- Create contingency table** and calculate totals for rows, columns, and grand total  $N$
- Calculate expected frequencies:**  $E_{ij} = \frac{\text{row}_i \text{ total} \times \text{column}_j \text{ total}}{N}$
- Calculate test statistic:**  $\chi_{\text{calc}}^2 = \sum_{i,j} \frac{(O_{ij} - E_{ij})^2}{E_{ij}}$  where  $O_{ij}$  are observed frequencies and  $E_{ij}$  are expected frequencies
- Determine degrees of freedom:**  $df = (r - 1)(c - 1)$  where  $r$  = number of rows and  $c$  = number of columns
- Specify significance level  $\alpha$**  (typically 0.05) and **Find critical value**  $\chi_{\text{crit}}^2$  from  $\chi^2$  table at chosen  $\alpha$  and  $df$
- Make decision:** If  $\chi_{\text{calc}}^2 > \chi_{\text{crit}}^2$ , reject  $H_0$  (variables are dependent). If  $\chi_{\text{calc}}^2 \leq \chi_{\text{crit}}^2$ , fail to reject  $H_0$  (insufficient evidence of dependence)

**Note:** Ensure all expected frequencies  $E_{ij} \geq 5$  for validity of chi-square approximation.

### 13. Wilcoxon Signed-Rank Test

**Use when:** Testing if the median of paired differences equals zero (or a hypothesized value) when normality assumption is violated; e.g., non-parametric alternative to paired t-test for before-after measurements, matched pairs when data is skewed or ordinal.

- **State hypotheses:**  $H_0 : \tilde{\mu}_d = 0$  (median difference is zero);  $H_1 :$  (choose:  $\tilde{\mu}_d \neq 0$  or  $\tilde{\mu}_d > 0$  or  $\tilde{\mu}_d < 0$ )
- **Calculate differences:**  $d_i = x_i - y_i$  for each paired observation
- **Rank absolute differences:** Ignore zero differences, rank  $|d_i|$  from smallest to largest (assign average ranks for ties)
- **Assign signs:** Attach the original sign of each difference to its rank
- **Calculate test statistic:**  $W = \sum$  of positive ranks (or sum of negative ranks - use smaller for two-tailed test)
- **Determine critical value:** For small samples ( $n \leq 25$ ), use Wilcoxon signed-rank table; for large samples ( $n > 25$ ), use normal approximation:  $Z = \frac{W - \frac{n(n+1)}{4}}{\sqrt{\frac{n(n+1)(2n+1)}{24}}}$
- **Specify significance level  $\alpha$**  (typically 0.05)
- **Make decision:**
  - Small sample: If  $W \leq W_{\text{crit}}$  from table, reject  $H_0$
  - Large sample: If  $|Z| > z_{\text{crit}}$  (two-tailed) or  $Z > z_{\text{crit}}$  (right-tailed) or  $Z < -z_{\text{crit}}$  (left-tailed), reject  $H_0$

**Note:** More robust than paired t-test for non-normal data or outliers. Assumes symmetric distribution of differences around the median.

### 14. Wilcoxon Mann Whitney U Test

**Use when:** Testing if two independent samples come from populations with the same distribution (non-parametric alternative to two-sample t-test); e.g., comparing treatment effects when normality assumption is violated, or with ordinal data.

- **State hypotheses:**  $H_0 :$  the two populations have identical distributions (or equal medians);  $H_1 :$  the distributions differ (or medians differ)
- **Rank all observations:** Combine both samples, rank from smallest to largest (assign average ranks for ties)
- **Calculate rank sums:**  $R_1 =$  sum of ranks for sample 1;  $R_2 =$  sum of ranks for sample 2
- **Calculate U statistics:**  $U_1 = n_1n_2 + \frac{n_1(n_1+1)}{2} - R_1$  and  $U_2 = n_1n_2 + \frac{n_2(n_2+1)}{2} - R_2$
- **Test statistic:** Use  $U = \min(U_1, U_2)$  (or equivalently, can use  $U_1$  or  $U_2$  with appropriate critical values)
- **Determine critical value:**
  - Small samples ( $n_1, n_2 \leq 20$ ): Use Mann-Whitney U table
  - Large samples: Use normal approximation:  $Z = \frac{U - \frac{n_1n_2}{2}}{\sqrt{\frac{n_1n_2(n_1+n_2+1)}{12}}}$
- **Specify significance level  $\alpha$**  (typically 0.05)
- **Make decision:**
  - Small sample: If  $U \leq U_{\text{crit}}$  from table, reject  $H_0$
  - Large sample: If  $|Z| > z_{\text{crit}}$  (two-tailed) or  $Z > z_{\text{crit}}$  (right-tailed) or  $Z < -z_{\text{crit}}$  (left-tailed), reject  $H_0$

**Note:** More robust than two-sample t-test for non-normal data, outliers, or ordinal data. Does not assume equal variances.

### 15. Kruskal-Wallis Test

**Use when:** Comparing medians across 3+ independent groups with non-normal data or ordinal scales (non-parametric alternative to one-way ANOVA); e.g., comparing satisfaction ratings across multiple brands, test scores across teaching methods when normality is violated; e.g. MM kitkat calorie question

- $H_0 :$  the medians are equal across all groups (or distributions are identical);  $H_1 :$  not all medians are equal (at least one group differs)

- Combine all observations from all groups and rank them from smallest to largest (assign average ranks for ties)
- Sum the ranks for each group:  $R_1, R_2, ..., R_k$
- Calculate test statistic:  $H = \left(\frac{12}{N(N+1)}\right) \sum_{j=1}^k \left(\frac{R_j^2}{n_j}\right) - 3(N+1)$  where  $N$  is total sample size and  $n_j$  is size of group  $j$
- Determine degrees of freedom:  $df = k - 1$  where  $k$  is number of groups
- Find critical value from  $\chi^2$  table at chosen significance level  $\alpha$
- Decision: If  $H_{\text{calc}} > \chi_{\text{crit}}^2$ , reject  $H_0$

**Note:** Use when ANOVA assumptions are violated (non-normal distributions, unequal variances) or with ordinal data. For large samples,  $H$  follows  $\chi^2$  distribution approximately.

## On Regression

Total Variance (SST) = Variance Explained (SSR) + Variance Unexplained (SSE)

$$\text{SST} = \sum_{i=1}^n \left(Y_i - \bar{Y}\right)^2 \quad \text{SSR} = \sum_{i=1}^n \left(\hat{Y}_i - \bar{Y}\right)^2 \quad \text{SSE} = \sum_{i=1}^n \left(Y_i - \hat{Y}_i\right)^2 \\ \text{SST} = \text{SSR} + \text{SSE}$$

$$\text{Coefficient of Determination } R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \left(\frac{\text{SSE}}{\text{SST}}\right)$$

#### F-distributed test statistic

$$H_0 : \beta_1 = \beta_2 = \beta_3 = c... = \beta_k = 0$$

$$H_1 : \text{at least one } \beta_j \neq 0$$

$k$  is the number of independent variables

$$\text{MSR} = \frac{\text{SSR}}{k}$$

$$\text{MSE} = \frac{\text{SSE}}{n-k-1} \text{ where } n-k-1 \text{ is the degrees of freedom}$$

$n$  is the number of observations and  $k$  is the number of independent variables

$$F = \frac{\frac{\text{SSR}}{k}}{\frac{\text{SSE}}{n-k-1}} = \frac{\text{MSR}}{\text{MSE}}$$

- The F-statistic in regression analysis is one-sided with the rejection region in the right tail of the F-distribution.
- If  $F_{\text{calc}} > F_{\text{crit}}$ , reject  $H_0$ .

### Hypothesis test for individual regression coefficients

#### Test of slope coefficient

$$t = \frac{\hat{b}_1 - B_1}{s_{\hat{b}_1}}$$

This is t distributed with  $n - k - 1$  degrees of freedom.

where  $s_{\hat{b}_1}$  is the standard error of the slope coefficient.

$$s_{\hat{b}_1} = \sqrt{\frac{\text{MSE}}{\sum_{i=1}^n \left(X_i - \bar{X}\right)^2}}$$

e.g. of linear regression b/w X and Y - test whether the slope coefferient is different from 0 to confirm if there is a signifiant relationship between X and Y.

- state hypothesis

$$H_0 : \beta_1 = 0 \quad H_1 : \beta_1 \neq 0$$

- Identify appropriate test statistic with  $t = \frac{\hat{b}_1 - B_1}{s_{\hat{b}_1}}$  with n - k - 1 = n - 2 degrees of freedom (with k = 1).
- specify significance level  $\alpha$  (say 0.05)

- state the decision rule by determining the critical value from t-distribution table. In this case, critical t-values =  $\pm 2.776$ . Reject  $H_0$  if  $t_{\text{calc}} < -2.776$  or  $t_{\text{calc}} > 2.776$ .
- Calculate test statistic. Suppose slope coefferient  $\hat{b}_1 = 1.25$ , MSE = 11.9688 and variation of X = 122.640. Then,  $s_e = \sqrt{\text{MSE}} = \sqrt{11.9688} = 3.46$ .  $s_{\hat{b}_1} = \frac{3.46}{\sqrt{122.640}} = 0.31$ .  $t = \frac{1.25-0}{0.31} = 4$
- Make a decision. Since  $t_{\text{calc}} = 4$  is greater than 2.776, we reject  $H_0$  of zero slope and conclude that there is a significant linear relationship between X and Y.

#### Test of correlation

- State the hypothesis.  $H_0 : \rho = 0 \quad H_1 : \rho \neq 0$

- Identify appropriate test statistic with

$$t = r\sqrt{\frac{n-2}{1-r^2}}$$

with n - k - 1 degrees of freedom = n - 2 (with k = 1).

- specify significance level  $\alpha$  (say 0.05)

- state the decision rule by determining the critical value from t-distribution table. In this case, critical t-values =  $\pm 2.776$ . Reject  $H_0$  if  $t_{\text{calc}} < -2.776$  or  $t_{\text{calc}} > 2.776$ .
- calculate the test statistic. correlation (r) is 0.8945 and n = 6.  $t = 0.8945 \frac{\sqrt{4}}{\sqrt{1-0.8945^2}} = 4$ .
- Make a decision. Since  $t_{\text{calc}} = 4$  is greater than 2.776, we reject  $H_0$  of zero correlation and conclude that there is significant evidence of correlation between X and Y.

e.g. test whether there is a positive slope or positive correlation. same steps as above. Except this is a one tailed test. So only check if  $t_{\text{calc}} > t_{\text{crit}}$ .  $t_{\text{crit}} = +2.312$  for 5% level of significance with 4 degrees of freedom.

#### Tests of the intercept

$$t_{\text{intercept}} = \frac{\hat{b}_0 - B_0}{s_{\hat{b}_0}}$$

where  $s_{\hat{b}_0} = S_e \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n \left(X_i - \bar{X}\right)^2}}$

## On $R^2$ and Adjusted $R^2$ scores

### $R^2$ (Coefficient of Determination)

**Formula:**

$$R^2 = \frac{\text{SSR}}{\text{SST}} = 1 - \left(\frac{\text{SSE}}{\text{SST}}\right)$$

where:

- $\text{SST} = \sum_{i=1}^n \left(Y_i - \bar{Y}\right)^2$  (Total Sum of Squares)
- $\text{SSR} = \sum_{i=1}^n \left(\hat{Y}_i - \bar{Y}\right)^2$  (Regression Sum of Squares - Explained Variation)
- $\text{SSE} = \sum_{i=1}^n \left(Y_i - \hat{Y}_i\right)^2$  (Error Sum of Squares - Unexplained Variation)

**Meaning:**

- Measures the proportion of total variation in the dependent variable (Y) that is explained by the independent variable(s)
- Range:  $0 \leq R^2 \leq 1$

- $R^2 = 0.89$  means 89% of variation in  $Y$  is explained by the model
- Higher  $R^2$  indicates better model fit
- $R^2$  always increases (or stays the same) when adding more predictors, even if they are not meaningful

Adjusted  $R^2$

**Formula:**  $R^2_{\text{adj}} = 1 - \left(\frac{n-1}{n-k-1}\right)(1 - R^2)$

or equivalently:  $R^2_{\text{adj}} = 1 - \left(\frac{\frac{\text{SSE}}{\text{SSE}}}{\frac{\text{SSE}}{\text{SSE}}}\right)$

- where:
- $n$  = number of observations
  - $k$  = number of independent variables (predictors)
  - $n - k - 1$  = degrees of freedom for residuals

Meaning:

- Adjusts  $R^2$  for the number of predictors in the model
- Penalizes for adding unnecessary variables
- Can decrease when adding predictors that don't improve the model enough
- Better for comparing models with different numbers of predictors
- Always:  $R^2_{\text{adj}} \leq R^2$
- Use Adjusted  $R^2$  to avoid overfitting and to compare models fairly

NOTE: relation b/w correlation and  $R^2$  score;  $r = \pm\sqrt{R^2}$  score. The sign is same as that of the slope coefficient.

G-test (Likelihood Ratio Test) for Logistic Regression

The G-test evaluates whether a logistic regression model with predictors significantly improves fit compared to a null model (intercept only).

Hypothesis

- $H_0$ : The predictor variable(s) provide no predictive value (null model is adequate)
- $H_1$ : At least one predictor significantly improves the model

Test Statistic

$G = 2(\ln L_{\text{full}} - \ln L_{\text{null}}) = \text{Null Deviance} - \text{Residual Deviance}$

- where:
- $\ln L_{\text{full}}$  = log-likelihood of model with predictors
  - $\ln L_{\text{null}}$  = log-likelihood of null model (intercept only)
  - Deviance =  $-2 \times \ln L$

Distribution:  $G \sim \chi^2_k$  where  $k$  = number of predictors added

Decision Rule

- Calculate  $G_{\text{calc}}$  from model output
- Compare with  $\chi^2_{\text{crit}}$  at significance level  $\alpha$  with  $k$  degrees of freedom
- Reject  $H_0$  if  $G_{\text{calc}} > \chi^2_{\text{crit}}$  or  $p\text{-value} < \alpha$

Example: Predicting Transmission Type from Weight (mtcars)

**Model:** Predict transmission type (am: 0=automatic, 1=manual) from car weight (wt)

Output Interpretation:

Null deviance: 43.230 on 31 degrees of freedom  
Residual deviance: 19.176 on 30 degrees of freedom

Calculation:

1. Calculate G-statistic:

$G = 43.230 - 19.176 = 24.054$

2. Degrees of freedom:  $k = 1$  (one predictor: weight)

3. Critical value:  $\chi^2_{0.05(1)} = 3.841$

4. Decision:  $G_{\text{calc}} = 24.054 > 3.841 = \chi^2_{\text{crit}}$

Therefore, **reject  $H_0$** . Weight is a significant predictor of transmission type.

5. P-value:  $p < 0.001$  (highly significant)

Interpretation:

- The model with weight fits significantly better than just predicting the overall proportion of manual vs automatic cars
- Adding weight reduces unexplained deviance by 24.054 units
- This improvement is statistically significant at any reasonable  $\alpha$  level

When to Use G-test:

- Overall model significance testing (is the model better than guessing?)
- Comparing nested models (e.g., model with 3 predictors vs model with 1 predictor)
- More reliable than Wald test for small samples or extreme coefficients
- Particularly important in logistic regression where residuals are not normally distributed

Model Selection Criteria: AIC and BIC

AIC (Akaike Information Criterion)

**Formula:**  $\text{AIC} = 2k - 2 \ln(L)$

or for linear regression:  $\text{AIC} = n \ln\left(\frac{\text{SSE}}{n}\right) + 2k$

- where:
- $k$  = number of parameters (including intercept):  $k = p + 1$  where  $p$  is number of predictors
  - $L$  = maximum likelihood of the model
  - $n$  = number of observations
  - SSE = sum of squared errors

Meaning & Interpretation:

- Measures model quality by balancing goodness of fit and model complexity
- Lower AIC indicates better model
- Penalizes for adding too many parameters (prevents overfitting)

**Decision Rule:** Compare AIC values across multiple candidate models and Choose the model with the **minimum AIC**

BIC (Bayesian Information Criterion)

**Formula:**  $\text{BIC} = k \ln(n) - 2 \ln(L)$

or for linear regression:  $\text{BIC} = n \ln\left(\frac{\text{SSE}}{n}\right) + k \ln(n)$

Meaning:

- Similar to AIC but with stronger penalty for additional parameters
- Lower BIC indicates better model
- BIC penalizes complexity more heavily than AIC (especially for large  $n$ )
- Tends to select simpler models than AIC
- Choose model with **minimum BIC**

Comparison:  $R^2$  vs Adjusted  $R^2$  vs AIC vs BIC

Criterion	Goal	Penalty for Complexity	Best Value	Use Case
$R^2$	Maximize	None	Higher = Better	Overall fit assessment

Adjusted $R^2$	Maximize	Moderate	Higher = Better	Compare models with different # of predictors
AIC	Minimize	Moderate (2k)	Lower = Better	Model selection, n balance fit & complexity
BIC	Minimize	Strong (k ln(n))	Lower = Better	Model selection, n prefer simpler models

There is a difference between regression model and the fitted regression equation.

e.g. model

$\text{Price} = \beta_0 + \beta_1 \text{Area} + \beta_2 \text{Bedrooms} + \varepsilon$

fitted regression equation

$\widehat{\text{Price}} = \widehat{\beta}_0 + \widehat{\beta}_1 \text{Area} + \widehat{\beta}_2 \text{Bedrooms}$

On deriving standard regression anova table from the sequential (type I) anova table

input

NOTE: this table is the output given by anova(model) in R

Source	Df	Sum Sq	Mean Sq	F value	Pr(>F)
SqFt	1	82.778	82.778	27.8040	1.842e-05
Bedrooms	1	13.833	13.833	4.6464	0.04094
Bathrooms	1	14.021	14.021	4.7095	0.03970
factor(Mall)	1	37.137	37.137	12.4740	0.00163
Residuals	25	74.430	2.977		

output

Source	Df	Sum Sq	Mean Sq	F value
Regression	4	SSR = 82.778 + 13.833 + 14.021 + 37.137 = 147.769	MSR = SSR / DF = 147.769 / 4 = 36.942	$F_{\text{obs}} = \frac{\text{MSR}}{\text{MSE}} = \frac{36.942}{2.977} = 12.41$
Residuals	25	SSE = 74.430	2.977	
Total	4 + 25 = 29	SST = SSR + SSE = 147.769 +		

		74.430 = 222.199		
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some more notes on this regd hypothesis testing

5% significance level with 4 and 25 degrees of freedom

From F-table, critical value = 2.76

Since  $F_{\text{obs}} = 12.41 > 2.76 = F_{\text{crit}}$ , we reject  $H_0$  and conclude that at least one of the regression coefficients is significantly different from zero.

on hypothesis testing for individual regression coefficients (say for SqFt)

$$H_0 : \beta_1 = 0 \quad H_1 : \beta_1 \neq 0$$

n = 30, k = 4 (no. of independent variables), df = n - k - 1 = 30 - 4 - 1 = 25 this is a 2 tailed test. so critical t values =  $\pm 2.060$  (from t table for 25 df at 5% significance level). The given t-value in the table is 1.230.  $1.230 < 2.060$ , so we fail to reject  $H_0$  and conclude that the regression coefficient for SqFt is not significantly different from zero. Alternatively, we can check the p-value =  $0.0.23028 > 0.05$ , so we fail to reject  $H_0$ .

### Another interesting question on ANOVA and regression

## Logistic Regression

### Generic Fitted Model for Logistic Regression

The logistic regression model predicts the probability that a binary outcome  $Y$  equals 1 given predictor variables  $X_1, X_2, \dots, X_p$ .

**Fitted Model:**

$$\hat{P}(Y = 1 \mid X) = \frac{1}{1 + e^{-(\hat{\beta}_0 + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \dots + \hat{\beta}_p X_p)}}$$

Or equivalently:

$$\hat{P}(Y = 1 \mid X) = \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \dots + \hat{\beta}_p X_p}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \dots + \hat{\beta}_p X_p}}$$

**Fitted Logit (Log-Odds) Equation:**

$$\text{logit}(\hat{P}) = \ln\left(\frac{\hat{P}}{1 - \hat{P}}\right) = \hat{\beta}_0 + \hat{\beta}_1 X_1 + \hat{\beta}_2 X_2 + \dots + \hat{\beta}_p X_p$$

Where:

- $\hat{P}(Y = 1 \mid X)$  is the estimated probability that  $Y = 1$
- $\hat{\beta}_0$  is the estimated intercept
- $\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_p$  are the estimated coefficients
- $X_1, X_2, \dots, X_p$  are the predictor variables
- The logit function transforms probabilities  $(0, 1)$  to the real line  $(-\infty, \infty)$

### VIF (Variance Inflation Factor) in Logistic Regression

VIF measures multicollinearity among independent variables - how much one predictor can be explained by other predictors in the model.

**Formula:**

For each independent variable  $X_j$ :  $\text{VIF}_j = \frac{1}{1 - R_j^2}$

where  $R_j^2$  is the coefficient of determination when  $X_j$  is regressed on all other independent variables.

**Interpretation:**

- VIF = 1:** No correlation with other predictors (ideal)
- VIF = 1-5:** Moderate correlation (generally acceptable)
- VIF = 5-10:** High correlation (concerning - potential multicollinearity problem)
- VIF > 10:** Severe multicollinearity (problematic - action needed)

NOTE: VIF applies to both linear and logistic regression because it measures relationships among the independent variables, not between predictors and the outcome.

## Important question

Output from `lm` command in R:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	(A)	20.673	6.578	3.26e-08 ***
SocialMedia	(B)	4.798	3.172	0.00264 **

Residual standard error: 80.97 on 48 degrees of freedom

Multiple R-squared: (C), Adjusted R-squared: 0.1561

F-statistic: (D) on 1 and 48 DF, p-value: 0.002637

Output from `anova` command in R:

	Df	Sum Sq
SocialMedia	1	65987
Residuals	48	314733

We need to find values of (A), (B), (C), and (D).

$A = b_0 =$ $s.e.(b_0) \times t_{b_0} =$ 135.987	$B = b_1 =$ $s.e.(b_1) \times t_{b_1} =$ 15.21926	$C = R^2 =$ $\frac{SST - SSE}{SST} =$ $\frac{65987}{380720} =$ 0.1733216	$D = F_{\text{obs}} =$ $\frac{MSR}{MSE} =$ $\frac{65987}{6556.938} =$ 10.06369
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For C and D: Regression ANOVA table:

Source of variation	Degrees of freedom	Sum of squares	Mean square	F statistic
Regression	1	$SSR =$ 65987	$MSR =$ $\frac{SSR}{1} =$ 65987	$D = F_{\text{obs}} =$ $\frac{MSR}{MSE} =$ $\frac{65987}{6556.938} =$ 10.06369
Residual or Error	$48 = n - 2$	$SSE =$ 314733	$MSE =$ $\frac{SSE}{48} =$ $\frac{314733}{48} =$ 6556.938	
Total	$49 = n - 1$	$SST =$ 380720		

## Summary Table

	Small Sample Size	Large Sample Size
Normal + Known Variance	z	z
Normal + Unknown Variance	t	t
Non-Normal + Known Variance	N/A	z
Non-Normal + Unknown Variance	N/A	t

checking for unbiased estimator... An estimator  $\hat{\theta}$  is unbiased for parameter  $\theta$  if:  $E[\hat{\theta}] = \theta$ . If  $E[\hat{Q}] \neq Q$ , the estimator is biased, with bias given by:  $\text{Bias}(\hat{Q}) = E[\hat{Q}] - Q$ .

MSE measures the average squared difference between the estimator and the true parameter value:  $\text{MSE}(\hat{Q}) = E[(\hat{Q} - Q)^2]$ . It can be decomposed into variance and bias components:  $\text{MSE}(\hat{Q}) = \text{Var}(\hat{Q}) + (\text{Bias}(\hat{Q}))^2$ . MSE measures the efficiency of an estimator; lower MSE indicates a more efficient estimator.

### MME Estimates for Normal Distribution

Given:  $X_1, X_2, \dots, X_n$  are IID from  $N(\mu, \sigma^2)$

Population moments:

- First moment:  $E[X] = \mu$
- Second moment:  $E[X^2] = \mu^2 + \sigma^2$

Sample moments:

- First sample moment:  $\bar{X} = \left(\frac{1}{n}\right) \sum_{i=1}^n X_i$
- Second sample moment:  $\left(\frac{1}{n}\right) \sum_{i=1}^n X_i^2$

MME estimates (equate sample and population moments):

$$\hat{\mu}_{\text{MME}} = \bar{X} = \left(\frac{1}{n}\right) \sum_{i=1}^n X_i$$

$$\hat{\sigma}_{\text{MME}}^2 = \left(\frac{1}{n}\right) \sum_{i=1}^n (X_i - \bar{X})^2$$

Properties:

- $\hat{\mu}_{\text{MME}}$  is unbiased:  $E[\hat{\mu}] = \mu$  ✓
- $\hat{\sigma}_{\text{MME}}^2$  is BIASED:  $E[\hat{\sigma}_{\text{MME}}^2] = \left(\frac{n-1}{n}\right)\sigma^2 \neq \sigma^2$  with bias =  $-\frac{\sigma^2}{n}$
- Unbiased estimator for  $\sigma^2$  uses  $(n - 1)$ :  $S^2 = \left(\frac{1}{n-1}\right) \sum_{i=1}^n (X_i - \bar{X})^2$
- For Normal distribution: MME = MLE for  $\mu$  and  $\sigma^2$

#	What we want to test?	Probability distribution of the statistic	Degrees of freedom	Test Statistic
1.	Population Mean	Normal / t distribution	n-1 for t	$z = \frac{\bar{x} - \mu_0}{\frac{\sigma}{\sqrt{n}}}$ or $t = \frac{\bar{x} - \mu_0}{\frac{s}{\sqrt{n}}}$
2.	Population Proportion	...	...	$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$ where $\hat{p} = \frac{x}{n}$
3.	Population Variance	Chi-square distribution	n-1	$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2}$
4.	mean of differences (paired t test)	t distribution	n-1	$t = \frac{\bar{d} - \mu_{d0}}{\frac{s_d}{\sqrt{n}}}$
5.	test of difference in means of 2 populations (assume equal variance)	Normal / t distribution	$n_1 + n_2 - 2$ for t	$t, z = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{s_p^2 \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}}$ (for pooled variance) $t, z = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\left( \frac{s_1^2}{n_1} \right) + \left( \frac{s_2^2}{n_2} \right)}}$ (for unpooled variance)
6.	test of difference in proportions of 2 populations	...	...	$z_{\text{calc}} = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1-\hat{p}) \left( \frac{1}{n_1} + \frac{1}{n_2} \right)}}$
7.	test of difference in variances of 2 populations	F distribution	$n_1 - 1$ and $n_2 - 1$	$F = \frac{s_1^2}{s_2^2}$
8.	difference in means of more than 2 populations (ANOVA)	F distribution	$k - 1$ and $N - k$	$F = \frac{MS_{\text{between}}}{MS_{\text{within}}}$
9.	barlett's test for equal variances	Chi-square distribution	$k - 1$	$\chi^2 = \frac{\left( \ln(S_p^2) - \sum \left( (n_i - 1) \ln \left( \frac{s_i^2}{n - k} \right) \right) \right) * (n - k)}{1 + \left( \frac{1}{3(k-1)} \right) * \left( \sum \left( \frac{1}{n_i - 1} \right) - \left( \frac{1}{n - k} \right) \right)}$
10.	two way ANOVA	F distribution	$(r - 1), (c - 1), (rc(n - 1))$	$F = \frac{MS_{\text{factor}}}{MS_{\text{error}}}$
11.	goodness of fit	Chi-square distribution	$c - 1$	$\chi^2 = \sum \left( \frac{(O_i - E_i)^2}{E_i} \right)$
12.	test of independence	Chi-square distribution	$(r - 1)(c - 1)$	$\chi^2 = \sum \left( \frac{(O_i - E_i)^2}{E_i} \right)$
13.	wilcoxon signed rank test	approximate Normal distribution for large n	...	$Z = \frac{W - \left( \frac{n(n+1)}{4} \right)}{\sqrt{n(n+1) \frac{2n+1}{24}}}$
14.	wilcoxon mann whitney test	approximate Normal distribution for large n	...	$Z = \frac{U - \left( n_1 \frac{n_2 + 1}{2} \right)}{\sqrt{n_1 n_2 \frac{n_1 + n_2 + 1}{12}}}$
15.	kruskal wallies test	approximate $\chi^2$ distribution for large n $H \sim \chi_{k-1}^2$	$k - 1$	$H = \left( \frac{12}{N(N+1)} \right) \sum \left( \left( \frac{R_j^2}{n_j} \right) \right) - 3(N + 1)$

NOTE: For  $\chi^2$  tests, ensure that the expected frequencies are sufficiently large (usually at least 5) to validate the use of the chi-square approximation.

## Statistical Tests – Comprehensive Comparison

Test Name	Parametric or Not	Decision Criteria	When it is used	When it is not used	Places of confusion	Related to which test	Examples from the reading material
One Sample t-test	Parametric	$t = \frac{\bar{X} - \mu_0}{SE}, df = n - 1$	To test a hypothesis about one mean in a single population where data is normally distributed (interval/ratio scale).	When data is not normally distributed or is non-metric (nominal/ordinal).	Deciding whether to use a t-test or Z-test depending on the situation (e.g., sample size).	One Sample Z-test	Testing whether the familiarity with the internet is high (>4) or not, where familiarity is measured on an interval scale of 1–7.
Two Independent Sample t-test	Parametric	t-statistic, F-test (Levene's) for equality of variances	To test whether the mean of a variable for two independent groups is different or the same.	When samples are related/paired, or when data is not normally distributed.	Whether to assume equal variances or not; requires testing for sample variance equality first.	Mann-Whitney U test (non-parametric equivalent)	Testing whether the mean familiarity with the internet is different or the same for males versus females.
Paired Sample t-test	Parametric	t-statistic based on mean and variance of paired differences ( $D$ ), $df = n - 1$	When data for two samples relate to the same group of respondents (e.g., measuring attitude toward two different things by the same sample).	When samples are drawn randomly from different populations (independent).	Confusing independent vs. paired samples.	Wilcoxon matched-pairs signed-ranks test, Sign test	Testing whether the mean of attitude towards the internet and attitude towards technology is the same or not within a single sample.
Kolmogorov-Smirnov (K-S) One-Sample Test	Non-parametric	K-S Z, maximum absolute difference between observed and theoretical distributions	To test whether observations for a single variable could reasonably have come from a particular distribution (e.g., Normal).	When you are comparing two independent samples (uses the 2-sample K-S test instead).	Applying Lilliefors significance correction.	Chi-square goodness-of-fit	Testing the hypothesis that Internet Usage (Hrs/Week) is normally distributed.
Chi-Square Goodness-of-Fit	Non-parametric	$\chi^2 = \sum \frac{(f_o - f_e)^2}{f_e}$	On a single categorical variable to determine if observed frequencies are similar to expected frequencies.	When expected cell frequencies are too small (e.g., minimum expected frequency < 5).	Ensuring the sum of percentages for the expected distribution equals exactly 100% (requires rescaling if not).	K-S test, Binomial test	Testing if the ratings of familiarity with the internet are equally distributed, or testing if the ratings follow a specific expected distribution like 10%, 20%, 20%, 10%, 25%, 15%.
Binomial Test	Non-parametric	Exact Sig. (comparing observed proportion to test proportion, e.g., $p = 0.5$ )	As a goodness-of-fit test specifically for dichotomous (two-category) variables.	When the variable has more than two categories.	Only applicable to variables that can be split into exactly two groups (e.g., Yes/No).	Chi-square goodness-of-fit	Testing if the proportion of people utilizing "Internet Shopping" (Yes/No) or "Internet Banking" (Yes/No) is significantly different from a test proportion of 0.50.
Runs Test	Non-parametric	Number of runs, Z-statistic	To determine whether the order or sequence in which observations are obtained is random.	When the actual shape of the distribution is being tested rather than the sequence of elements.	Determining what constitutes a 'run' (a maximal non-empty segment of adjacent equal elements).	—	Evaluating a sequence of positive and negative elements (e.g., "++++ +++++++ + -"), or testing whether the observations for familiarity with the internet were generated randomly.
Chi-Square Test for Independence	Non-parametric	$\chi^2 = \sum \frac{(f_o - f_e)^2}{f_e}, df = (r - 1) \times (c - 1)$	To determine whether frequencies observed at the combination of levels of two categorical variables are associated or independent.	When expected counts in cells are less than 5.	Maximum value of contingency coefficient depends on table size, so it should only be used to compare tables of the same size.	Phi coefficient, Contingency coefficient, Cramer's V	Testing if the proportion of respondents using the internet for shopping (light vs. heavy usage) is indifferent to gender (males vs. females).
Mann-Whitney U Test	Non-parametric	U statistic, Wilcoxon W, Z-statistic	To compare difference in location of two populations using two independent samples, measured on an ordinal scale.	When data is paired/related, or when testing more than two groups.	It uses rank combinations instead of absolute values (cases are combined and ranked).	Two Independent Sample t-test, Kruskal-Wallis	Testing the hypothesis that two populations (male and female) are identical with respect to their mean rank of familiarity with the internet.
Wilcoxon Matched-Pairs Signed-Ranks Test	Non-parametric	Z-statistic based on positive/negative ranks	Analyzes differences between paired observations, taking into account the magnitude of the differences.	When samples are independent.	Handling ties (differences of zero) in the ranking process.	Paired Sample t-test, Sign test	Testing whether the mean rank of attitude towards the internet is the same as the mean rank of attitude towards technology.
Paired Sample Sign Test	Non-parametric	Exact Sig. based on binomial distribution	Analyzes differences between paired observations, taking into account only the sign (pluses vs. minuses) of the differences.	When the magnitude of the differences is available and meaningful (Wilcoxon is better).	It is less powerful than Wilcoxon because it ignores the magnitude of the difference.	Wilcoxon matched-pairs signed-ranks test	Analyzing the frequencies of positive differences, negative differences, and ties between the paired variables "Attitude toward Technology" and "Attitude toward Internet".
One-way ANOVA	Parametric	$F = \frac{MS_{\text{between}}}{MS_{\text{within}}}, df = (c - 1) \text{ and } (N - c)$	Test of means for two or more populations with one categorical independent variable and a metric dependent variable.	When error terms are correlated, variances are not constant, or distribution is not normal.	Rejecting the null hypothesis only tells you that not all means are equal; you must use multiple comparison tests (like LSD) to find which specific means differ.	Kruskal-Wallis test (non-parametric equivalent)	A department store determining the effect of different levels of in-store promotion (High, Medium, Low) on store sales.

Test Name	Parametric or Not	Decision Criteria	When it is used	When it is not used	Places of confusion	Related to which test	Examples from the reading material
n-way (Factorial) ANOVA	Parametric	F-statistics for main effects and interaction effects	To measure the effects of two or more independent categorical variables at various levels on a metric dependent variable.	—	Interpreting interaction effects (where the effect of one variable depends on the level of another).	One-way ANOVA	Measuring the combined effect of two independent variables: the amount of humor (No, Medium, High) and the amount of store information (Low, Medium, High). Another example tests the effect of in-store promotion and couponing on sales.
Analysis of Covariance (ANCOVA)	Parametric	F-statistics	To account for the influence of uncontrolled independent variables (covariates) while examining differences in mean values related to controlled independent variables.	—	Differentiating the impact of the categorical factors from the continuous covariates.	ANOVA	Determining how different price levels affect cereal consumption while controlling for the covariate of household size. Another example determines the effect of in-store promotion and couponing on sales while controlling for the clientele effect.
Kruskal-Wallis Test	Non-parametric	H statistic (Chi-Square distribution)	Extension of Mann-Whitney test. Examines difference in medians across more than two groups (one-way ANOVA on ranks).	When parametric assumptions are fully met (ANOVA is more powerful).	Cases from all k groups are ordered into a single ranking to calculate rank sums.	One-way ANOVA, Mann-Whitney U	Examining differences in sales by ordering and comparing the mean ranks across three different levels of promotion.
k-Sample Median Test	Non-parametric	Pearson's Chi-Square based on median split	Tests null hypothesis that medians of populations from 2 or more samples are identical by grouping data into "> Median" and "≤ Median".	When there are few tied rankings (Kruskal-Wallis is more powerful).	It is less powerful than Kruskal-Wallis because it ignores exact rank and only uses location relative to the overall median.	Kruskal-Wallis Test	Grouping observed sales data into "> Median" or "≤ Median" frequencies and testing for differences across three levels of promotion.