

# ST2131/MA2216 AY1718 Sem 2 Answers

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1. (a)

$$P(X = i) = pq^{i-1} = (1 - q)q^{i-1} = q^{i-1} - q^i$$

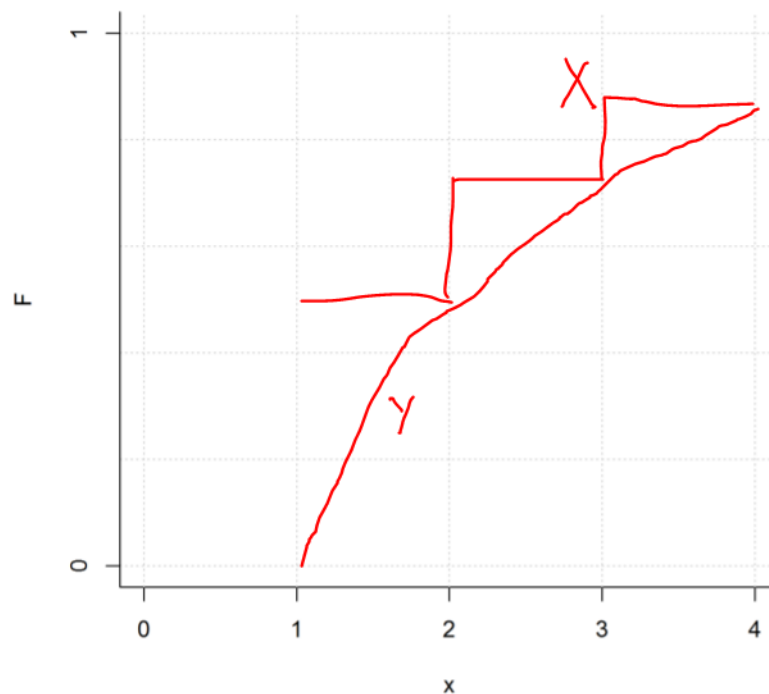
(b)

$$\begin{aligned} P(X \leq i) &= (q^0 - q^1) + (q^1 - q^2) + \cdots + (q^{i-1} - q^i) \\ &= 1 - q^i \end{aligned}$$

(c)

$$\begin{aligned} P(W = w) &= P(w \leq F_Y^{-1}(U) < w + 1) \\ &= P(F_Y(w) \leq U < F_Y(w + 1)) \\ &= P(1 - q^{w-1} \leq U < 1 - q^w) \\ &= q^{w-1} - q^w \\ &= \text{Geo}(p) \end{aligned}$$

(d)



2. Let  $C$  be the total sum of car weights

$$C \sim Z(115 \times 3, 115 \times 0.3^2)$$

$$W - C \sim Z(400 - 115 \times 3, 40^2 + 115 \times 0.3^2)$$

$$\begin{aligned} P(W < C) &= P(W - C < 0) \\ &= P\left(Z < \frac{0 - 400 - 115 \times 3}{\sqrt{40^2 + 115 \times 0.3^2}}\right) \\ &= P(Z < -1.3706) \\ &= 0.0852 \end{aligned}$$

3. (a)

$$\begin{aligned} E(Y) &= \int_0^1 E(N(x, x^2)) \, dx \\ &= \int_0^1 x \, dx \\ &= \frac{1}{2} \end{aligned}$$

$$\begin{aligned} E(Y^2) &= \int_0^1 E(N(x, x^2)^2) \, dx \\ &= \int_0^1 2x^2 \, dx \\ &= \frac{2}{3} \end{aligned}$$

$$\begin{aligned} Var(Y) &= \frac{2}{3} - \left(\frac{1}{2}\right)^2 \\ &= \frac{5}{12} \end{aligned}$$

$$E(X) = \frac{1}{2}$$

$$\begin{aligned} E(XY) &= \int_0^1 E(XY|Y=x) \, dx \\ &= \int_0^1 x^2 \, dx \\ &= \frac{1}{3} \end{aligned}$$

$$\begin{aligned} cov(X, Y) &= \frac{1}{3} - \left(\frac{1}{2}\right)\left(\frac{1}{2}\right) \\ &= \frac{1}{12} \end{aligned}$$

(b)

$$U = g(X, Y) = Y/X, V = h(X, Y) = X$$

$$x = v, y = uv$$

$$J(x, y) = \begin{vmatrix} \frac{\partial g}{\partial x} & \frac{\partial g}{\partial y} \\ \frac{\partial h}{\partial x} & \frac{\partial h}{\partial y} \end{vmatrix} = \begin{vmatrix} -\frac{y}{x^2} & \frac{1}{x} \\ 1 & 0 \end{vmatrix} = -\frac{1}{x}$$

$$\begin{aligned}
f_{U,V}(u, v) &= f_{X,Y}(x, y) | -x| \\
&= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{1}{2x^2}(y-x)^2\right\} \\
&= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{1}{2v^2}(uv-v)^2\right\} \\
&= \frac{1}{\sqrt{2\pi}} \exp\left\{-\frac{1}{2v^2}(u^2v^2 - 2uv^2 + v^2)\right\} \\
&= \frac{1}{\sqrt{2\pi}} \exp\{u - u^2/2 - 1/2\}
\end{aligned}$$

Which is not even dependent on  $v$ . Hence, they are independent.

4. Note that  $|Y| \sim \text{Unif}(0, 1)$

$$\begin{aligned}
P(Z < k) &= \int_0^1 P(X/y < k) \, dy \\
&= \int_0^1 P(X < ky) \, dy
\end{aligned}$$

If  $0 < k < 2$ , then  $ky < 2$ , hence,  $P(X < ky) = ky/2$

$$\begin{aligned}
P(Z < k) &= \int_0^1 P(X < ky) \, dy = \int_0^1 ky/2 \, dy = k/4 \\
&\therefore f_Z(k) = 1/4
\end{aligned}$$

Else, if  $k \geq 2$ , then  $2/k \leq 1$

$$\begin{aligned}
P(Z < k) &= \int_0^1 P(X < ky) \, dy \\
&= \int_0^{2/k} P(X < ky) \, dy + \int_{2/k}^1 P(X < ky) \, dy \\
&= \int_0^{2/k} ky/2 \, dy + \int_{2/k}^1 1 \, dy \\
&= (k/4)(2/k)^2 + (1 - 2/k) \\
&= 1 - 1/k \\
&\therefore f_Z(k) = 1/k^2
\end{aligned}$$

5. Define  $R_0 := 0$

Suppose  $E(R_i) = i$  for all  $i = 0, 1, \dots, n-1$ . We want to show that  $E(R_n) = n$  by induction. The base case of  $i = 0$  is trivial.

$$\begin{aligned}
E(R_n) &= E\left(1 + \sum_{i=0}^n P(X_n = i)R_{n-i}\right) \\
&= 1 + \sum_{i=0}^n P(X_n = i)E(R_{n-i}) \\
&= 1 + \sum_{i=0}^n P(X_n = i)(n - i) + P(X_n = 0)(E(R_n) - n) \\
&= 1 + n \sum_{i=0}^n P(X_n = i) - \sum_{i=0}^n iP(X_n = i) + P(X_n = 0)(E(R_n) - n) \\
&= 1 + n(1) - E(X_n) + P(X_n = 0)(E(R_n) - n) \\
&= 1 + n(1) - 1 + P(X_n = 0)(E(R_n) - n) \\
&= n + P(X_n = 0)(E(R_n) - n)
\end{aligned}$$

$$\therefore E(R_n) = n + P(X_n = 0)(E(R_n) - n)$$

$$\therefore E(R_n) = \frac{n - nP(X_n = 0)}{1 - P(X_n = 0)} = n$$