#### Markov chain Monte Carlo methods

#### Goal

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z},$$

where  $Z = \int \gamma(x) dx$ .

#### Goal

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z},$$

where  $Z = \int \gamma(x) dx$ .

We want to approximate complex integral: let  $x_1,...,x_N\sim p(x)$ , then for any test function h(x):

$$\mathbb{E}_{X \sim p}[h(X)] = \int h(x)p(x)dx \approx \frac{1}{N} \sum_{n=1}^{N} h(x_n).$$

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z}$$

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z}$$

• p(x) = p(x|y): posterior

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z}$$

- p(x) = p(x|y): posterior
- $\gamma(x) = p(x, y)$ : joint likelihood

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z}$$

- p(x) = p(x|y): posterior
- $\gamma(x) = p(x, y)$ : joint likelihood
- $Z=p(y)=\int p(x,y)dx$ : marginal likelihood, which is usually intractable.

Sample from a distribution:

$$p(x) = \frac{\gamma(x)}{Z}$$

- p(x) = p(x|y): posterior
- $\gamma(x) = p(x, y)$ : joint likelihood
- $Z = p(y) = \int p(x,y) dx$ : marginal likelihood, which is usually intractable.

Here, x can be parameters or any latent variables of interest.

### Importance sampling

Instead of sampling from p (hard to do because Z is unknown), sample  $x \sim q$  and adjust for the difference between  $\gamma$  and q:

$$\int h(x)p(x)dx \approx \sum \bar{w}(x)h(x),$$

where 
$$w(x) = \gamma(x)/q(x)$$
 and  $\bar{w}(x) = w(x)/\sum_n w(x)$ .

If  $\boldsymbol{x}=(x_1,...,x_D)$  is high-dimensional, we can sample each component sequentially:

If  $\boldsymbol{x}=(x_1,...,x_D)$  is high-dimensional, we can sample each component sequentially:

$$\bullet \ x_d^n \sim q_d(x_d|x_{1:d-1}).$$

If  $\boldsymbol{x}=(x_1,...,x_D)$  is high-dimensional, we can sample each component sequentially:

- $x_d^n \sim q_d(x_d|x_{1:d-1})$ .
- Interleave resampling step to maintain particle diversity and prune unpromising particles.

SMC methods work well when there is a temporal structure in x, where it is natural to sample one dimension at a time.

SMC methods work well when there is a temporal structure in x, where it is natural to sample one dimension at a time.

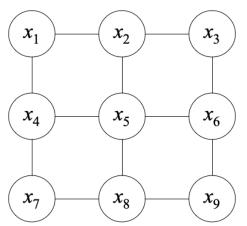
The main theoretical driver for IS/SMC methods is the law of large numbers.

. .

So why do we need another algorithm/method?

- Large variance associated with (poor) choice of proposal distribution.
- Curse of dimensionality may still manifest and approximation can be poor.

Consider a  $K \times K$  2-dimensional lattice G = (V, E).



 $\bullet$  Ising model: Each node represents a RV  $X_v$  which takes a value in  $\{-1,+1\},$  denoting the "spin" of an atom/molecule.

- Ising model: Each node represents a RV  $X_v$  which takes a value in  $\{-1,+1\}$ , denoting the "spin" of an atom/molecule.
- Spatial analysis: Each node represents a spatial coordinate (spatial statistics) and  $X_v \in \{ \text{no gold}, \text{gold!} \}$  or  $X_v \in \mathbb{R}^+$  some measurement of the amount of gold at location v.

- Ising model: Each node represents a RV  $X_v$  which takes a value in  $\{-1,+1\}$ , denoting the "spin" of an atom/molecule.
- Spatial analysis: Each node represents a spatial coordinate (spatial statistics) and  $X_v \in \{ \text{no gold}, \text{gold!} \}$  or  $X_v \in \mathbb{R}^+$  some measurement of the amount of gold at location v.
- Image processing: Each node represents a pixel of an image.  $X_v \in \{ \text{black}, \text{white} \}$  or gray scale  $X_v \in [0,1]$  or RGB color.

### Example: Ising model

Let  $X=(X_v).$  The "energy" function for the Ising model is defined as:

$$H(x) = \sum_v \phi(x_v) + \sum_{(u,v) \in E} \psi(x_u,x_v),$$

- $(u, v) \in E$  denote neighbors (adjacent nodes),
- $\phi$ : unary potential,
- $\psi$ : pairwise potential (measuring interaction strength).

Example:  $\phi(x_v) = \beta x_v$  and  $\psi(x_u, x_v) = \kappa x_u x_v$  for  $\beta, \kappa \in \mathbb{R}$ .

### Example: Ising model

Tthe probability distribution on X is defined as,

$$p(x) = \frac{1}{Z} \exp(-H(x))$$

where

$$Z = \sum_{x_n: v \in V} \exp(-H(x)).$$

 ${\cal Z}$  in statistical physics is referred to as "partition function". Essentially a normalization constant.

### SMC for Ising model?

- Not obvious what order to sample the variables.
- $\bullet$  It could lead to very few unique values for  $x_v$  sampled earlier in the SMC iteration.
- Leads to poor approximation involving those sampled earlier on.

# SMC for Ising model?

- Not obvious what order to sample the variables.
- $\bullet$  It could lead to very few unique values for  $x_v$  sampled earlier in the SMC iteration.
- Leads to poor approximation involving those sampled earlier on.
- $\bullet$  For the Ising model, maybe it makese more sense to continually sample new values for  $x_v$  given  $x_{-v}$  until we are satisfied.

Initialize  $x^0$ .

Initialize  $x^0$ .

For t = 1, ..., T:

Initialize  $x^0$ .

For t = 1, ..., T:

 $\bullet \text{ Propose a value } x' \sim q(\cdot|x_{t-1}).$ 

Initialize  $x^0$ .

For t = 1, ..., T:

- Propose a value  $x' \sim q(\cdot|x_{t-1})$ .
- Compute acceptance probability:

$$A(x'|x) = \min\left(1, \frac{\gamma(x')}{\gamma(x_{t-1})} \frac{q(x_{t-1}|x')}{q(x'|x_{t-1})}\right).$$

Initialize  $x^0$ .

For t = 1, ..., T:

- Propose a value  $x' \sim q(\cdot|x_{t-1})$ .
- Compute acceptance probability:

$$A(x'|x) = \min\left(1, \frac{\gamma(x')}{\gamma(x_{t-1})} \frac{q(x_{t-1}|x')}{q(x'|x_{t-1})}\right).$$

• Sample  $u \sim U(0,1)$ 

Initialize  $x^0$ .

For t = 1, ..., T:

- Propose a value  $x' \sim q(\cdot|x_{t-1})$ .
- Compute acceptance probability:

$$A(x'|x) = \min\left(1, \frac{\gamma(x')}{\gamma(x_{t-1})} \frac{q(x_{t-1}|x')}{q(x'|x_{t-1})}\right).$$

- Sample  $u \sim U(0,1)$
- Set,

$$x_t = \left\{ \begin{array}{ll} x' & \text{if } u < A \\ x_{t-1} & \text{otherwise} \end{array} \right.$$

# Why does MH work?

• If we take samples  $x_1,...,x_N$  using MH algorithm, why is this equivalent to taking samples from the target distribution p(x)?

#### Markov chain

Markov chain  $\{X_t\}$  is a stochastic process modeling a sequence of events where the probability of each event depends only on the previous event.

 $\bullet \ \ \text{Markov property:} \ p(x_t|x_{1:t-1}) = p(x_t|x_{t-1}).$ 

#### Markov chain

Given measurable space  $(\mathcal{X},\mathcal{F})$ ,

$$K: \mathcal{X} \times \mathcal{F} \to [0,1]$$

is referred to as the Markov kernel (a probability measure).

- ullet Each random variable  $X_t \in \mathcal{X}$
- K(x,A) specifies the probability of moving to a set  $F \in \mathcal{F}$  given that the chain is in state  $x \in \mathcal{X}$ .

# Markov chain: continuous state space

For continuous state space,  $\mathcal{X}=\mathbb{R}$ , the transition probability can be described using a density function  $K(x_{t-1},x_t)=k(x_t|x_{t-1})$ .

### Markov chain: discrete state space

For discrete state space, the Markov chain is described using a transition matrix P, where  $P_{ij}$  represents the probability of transitioning from state  $P(x_t=j|x_{t-1}=i)$ .

# Markov chain: stationary distribution

The Markov chain  $\{X_t\}$  converges to unique **stationary** distribution as  $t\to\infty$  if some conditions are satisfied.

A probability distribution  $\pi$  defined on  $\mathcal X$  is invariant (stationary) under a Markov kernel K if for all  $F\in\mathcal F$ 

$$\pi(A) = \int \pi(x)K(x,F)dx.$$

For discrete case:  $\pi = \pi P$ .

# Markov chain: detailed balance (reversibility)

A Markov chain with kernel  $K: \mathcal{X} \times \mathcal{F}$  satisfies the detailed balance condition with respect to a probability distribution  $\pi$  if,

$$\pi(x)k(x'|x) = \pi(x')k(x|x').$$

**Reversibility**: probability of being in state x and moving to x' from x is the same as being in state x' and moving to x.

• Note: detailed balance is a stronger condition than stationary condition: if detailed balance is satisfied,  $\pi$  is a stationary distribution of the Markov chain with kernel K.

# Markov chain: Ergodicity

- Aperiodic: Markov chain does not return to the same state at some fixed interval.
- Positive recurrent: the expected number of steps for returning to the same state is finite.

### Metropolis-Hastings is a Markov chain

Claim: MH algorithm constructs a Markov chain on  $\mathcal X$  whose stationary distribution is p(x).

### Metropolis-Hastings is a Markov chain

Claim: MH algorithm constructs a Markov chain on  $\mathcal X$  whose stationary distribution is p(x).

#### Markov transition kernel:

ullet given current state x, we move to a new state x' with probability

$$q(x^{\prime}|x)A(x^{\prime}|x)$$

## Metropolis-Hastings is a Markov chain

Claim: MH algorithm constructs a Markov chain on  $\mathcal X$  whose stationary distribution is p(x).

#### Markov transition kernel:

ullet given current state x, we move to a new state x' with probability

$$q(x'|x)A(x'|x)$$

stay at current state x with probability

$$q(x|x) + q(x'|x)(1 - A(x)).$$

To prove: p(x)k(x'|x) = p(x')k(x|x').

To prove: 
$$p(x)k(x'|x) = p(x')k(x|x')$$
.

To move from state x to x', we must first propose x' and accept x'.

To prove: p(x)k(x'|x) = p(x')k(x|x').

To move from state x to x', we must first propose x' and accept x'.

Case 1: 
$$A(x'|x) = \frac{p(x')q(x|x')}{p(x)q(x'|x)} < 1$$
.

$$p(x)q(x'|x)A(x'|x) = p(x)q(x'|x)\frac{p(x')q(x|x')}{p(x)q(x'|x)}$$
(1)

$$= p(x')q(x|x'). (2)$$

Case 2:  $A(x'|x) \ge 1$ .

$$p(x')q(x|x')A(x|x') = p(x')q(x|x')\frac{p(x)q(x'|x)}{p(x')q(x|x')}$$
(3)

$$= p(x)q(x'|x). (4)$$

#### Is MH an ergodic Markov chain?

Yes, as long as we choose our proposal carefully.

ullet Randomness in k is needed to prevent apriodicity. This is built-in the MH kernel where we reject proposed values with some chance.

#### Is MH an ergodic Markov chain?

Yes, as long as we choose our proposal carefully.

- ullet Randomness in k is needed to prevent apriodicity. This is built-in the MH kernel where we reject proposed values with some chance.
- Recurrent: ensure that we choose a proposal that allows to visit every state  $x \in \mathcal{X}$ . For example, Gaussian random walk  $q(x'|x) = N(x'|x, \sigma^2 I)$  satisfies this.

#### Is MH an ergodic Markov chain?

Yes, as long as we choose our proposal carefully.

- ullet Randomness in k is needed to prevent apriodicity. This is built-in the MH kernel where we reject proposed values with some chance.
- Recurrent: ensure that we choose a proposal that allows to visit every state  $x \in \mathcal{X}$ . For example, Gaussian random walk  $q(x'|x) = N(x'|x, \sigma^2 I)$  satisfies this.
- For discrete case, ensure q(x'|x) > 0 for all  $x', x \in \mathcal{X}$ .

 $\bullet$  Choosing a suitable Metropolis-Hastings proposal distribution q(x'|x) is crucial.

- Choosing a suitable Metropolis-Hastings proposal distribution  $q(x^\prime|x)$  is crucial.
- A **local proposal** (e.g., small Guassian perturbation) allows gradual exploration and prevents the chain from getting stuck.

- $\bullet$  Choosing a suitable Metropolis-Hastings proposal distribution q(x'|x) is crucial.
- A **local proposal** (e.g., small Guassian perturbation) allows gradual exploration and prevents the chain from getting stuck.
- Independent Metropolis refers to the case where a global proposal is used q(x'|x) = q(x'), which can lead to high rejection rates (why?).

- $\bullet$  Choosing a suitable Metropolis-Hastings proposal distribution q(x'|x) is crucial.
- A **local proposal** (e.g., small Guassian perturbation) allows gradual exploration and prevents the chain from getting stuck.
- Independent Metropolis refers to the case where a global proposal is used q(x'|x) = q(x'), which can lead to high rejection rates (why?).
- Large global proposals tend to be rejected, causing the chain to get stuck at a point for long periods.

Gibbs sampling is an MCMC algorithm, which is well suited for high-dimensional distributions where sampling directly from the joint distribution is difficult.

• Initialize  $x^0$ .

- Initialize  $x^0$ .
- ② For t = 1, ..., T:
  - Iterate over each variable  $x_i$ :
    - $\blacktriangleright$  Sample  $x_i^{(t)} \sim p(x_i|x_{-i}^{(t)})$  , where  $x_{-i}$  refers to all other variables except  $x_i.$

- Initialize  $x^0$ .
- ② For t = 1, ..., T:
  - Iterate over each variable  $x_i$ :
    - $\blacktriangleright$  Sample  $x_i^{(t)} \sim p(x_i|x_{-i}^{(t)}),$  where  $x_{-i}$  refers to all other variables except  $x_i.$

This means that each variable is sampled from its conditional distribution given the current values of all other variables.

- Initialize  $x^0$ .
- ② For t = 1, ..., T:
  - Iterate over each variable  $x_i$ :
    - $\blacktriangleright$  Sample  $x_i^{(t)} \sim p(x_i|x_{-i}^{(t)})$  , where  $x_{-i}$  refers to all other variables except  $x_i.$

This means that each variable is sampled from its conditional distribution given the current values of all other variables.

Gibbs sampling is particularly effective when the conditional distributions  $p(x_i|x_{-i})$  are easy to sample from.

Claim: Gibbs sampling can be viewed as a special case of the Metropolis-Hastings algorithm where the proposal distribution is always accepted.

Claim: Gibbs sampling can be viewed as a special case of the Metropolis-Hastings algorithm where the proposal distribution is always accepted.

Proof: Suppose we are proposing a new value for  $x_i.$  Let  $x'=(x_1,...,x_{i-1},x_i',x_{i+1},...,x_N).$ 

Claim: Gibbs sampling can be viewed as a special case of the Metropolis-Hastings algorithm where the proposal distribution is always accepted.

Proof: Suppose we are proposing a new value for  $x_i.$  Let  $x'=(x_1,...,x_{i-1},x_i',x_{i+1},...,x_N).$ 

$$\begin{split} A(x'|x) &= \frac{p(x')q(x|x')}{p(x)q(x'|x)} \\ &= \frac{p(x_i'|x_{-i})p(x_{-i})q(x|x')}{p(x_i|x_{-i})p(x_{-i})q(x'|x)}. \end{split}$$

Claim: Gibbs sampling can be viewed as a special case of the Metropolis-Hastings algorithm where the proposal distribution is always accepted.

Proof: Suppose we are proposing a new value for  $x_i.$  Let  $x'=(x_1,...,x_{i-1},x_i',x_{i+1},...,x_N).$ 

$$\begin{split} A(x'|x) &= \frac{p(x')q(x|x')}{p(x)q(x'|x)} \\ &= \frac{p(x_i'|x_{-i})p(x_{-i})q(x|x')}{p(x_i|x_{-i})p(x_{-i})q(x'|x)}. \end{split}$$

Since  $q(x'|x)=p(x_i'|x_{-i})$  and  $q(x|x')=p(x_i|x_{-i})$ , the acceptance probability simplifies to 1.

## Gibbs for Ising model

For t = 1, ..., T:

 $\bullet \ \mbox{Sample} \ x_v \sim p(x_v|x_{-v}) \ \mbox{for each} \ v \in V.$ 

Sample each variable in turn, conditioned on the values of all of the other variables.

## Gibbs for Ising model

For t = 1, ..., T:

 $\bullet \ \mbox{Sample} \ x_v \sim p(x_v|x_{-v}) \ \mbox{for each} \ v \in V.$ 

Sample each variable in turn, conditioned on the values of all of the other variables.

What does  $p(x_v|x_{-v})$  look like?

# Gibbs sampling for Ising model

$$p(x_v|x_{-v}) = \frac{p(x_v, x_{-v})}{\sum_{x_v'} p(x_v', x_{-v})}$$
 (5)

$$\propto \exp(-\phi(x_v) - \sum_{(u,v) \in E} \psi(x_u,x_v)). \tag{6}$$

### Example: image denoising

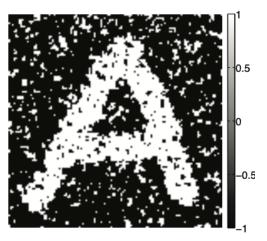


Figure 1: Fig 12.3 (a), PML 2

If all of the neighbors of  $x_v$  is white/black,  $x_v$  is likely to be white/black.

# Example: image denoising

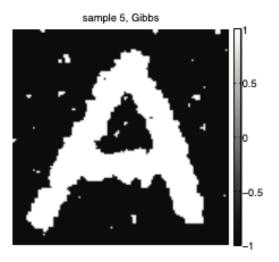


Figure 2: Fig 12.3 (b), PML 2

#### Example: image denoising

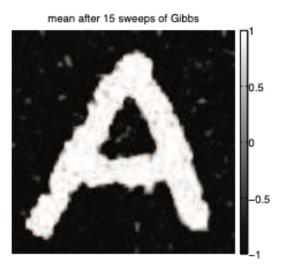


Figure 3: Fig 12.3 (c), PML 2

Undirected graphical models (UGM):

ullet Each node  $v \in V$  represents a random variable.

#### Undirected graphical models (UGM):

- Each node  $v \in V$  represents a random variable.
- ullet Edge between  $u,v\in V$  is denoted (u,v). Presence of an edge indicates that there is a symmetric relationship between u and v but we cannot easily pinpoint directionality.

• UGMs are commonly referred to as Markov Random Field (MRF).

- UGMs are commonly referred to as Markov Random Field (MRF).
- Commonly used for modeling dependence structure where directionality is unclear.

- UGMs are commonly referred to as Markov Random Field (MRF).
- Commonly used for modeling dependence structure where directionality is unclear.
- Example: The value taken at each pixel (random variable  $X_v$ ) is related to the value taken by its neighbors but it is not causal.

• Pairwise Markov property: For any two non-adjacent nodes u,v,  $X_u \perp X_v | X_{rest}$ .

- Pairwise Markov property: For any two non-adjacent nodes u,v,  $X_u \perp X_v | X_{rest}$ .
- Local Markov property:  $X_u \perp X_{rest} | X_{nbr(u)}$ , where  $nbr(u) = \{v: (u,v) \in E\}.$

- Pairwise Markov property: For any two non-adjacent nodes u,v,  $X_u \perp X_v | X_{rest}$ .
- $\bullet$  Local Markov property:  $X_u \bot X_{rest} | X_{nbr(u)}$  , where  $nbr(u) = \{v: (u,v) \in E\}.$
- Global Markov property:

Any two sets  $A,B\subset V$ , are conditionally independent given a separating set S, i.e.,  $X_A\bot X_B|X_S$ , if S separates A and B in G.

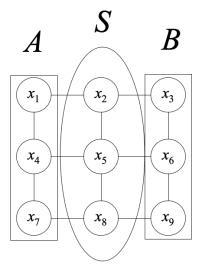


Figure 4: Global Markov property

**Markov blanket** of v is defined as a minimal set of nodes that separates v from the rest of the nodes. It is given by, MB(v) = nbr(v).

**Markov blanket** of v is defined as a minimal set of nodes that separates v from the rest of the nodes. It is given by, MB(v) = nbr(v).

MB plays a central role in determining efficient inference algorithm. Example, Gibbs sampling update of a variable  $X_v$  is conditioned on its MB and nothing else.

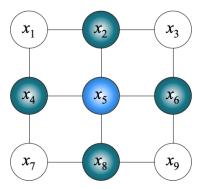


Figure 5: Markov blanket

# Undirected graphical models: Hammersley-Clifford Theorem

A strictly positive probability distribution  $p(x_V)$  satisfies the global Markov property with respect to G if and only if it can be factorized as,

$$p(x_V) = \frac{1}{Z} \prod_{C \in \mathcal{C}} \psi_C(x_C),$$

- ullet C denotes the set of (maximal) cliques,
- $\psi_C$  denotes potential function for clicque C,
- ullet Z is normalization constant also referred to as partition function.

Clique  $C\subseteq V$  of G=(V,E) is a fully connected subgraph of G such that every pair of nodes  $u,v\in C$  are adjacent i.e.,  $\{u,v\}\in E$ .

A clique  ${\cal C}$  is maximal if adding a node to  ${\cal C}$  does not preserve full connectivity.

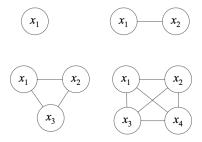


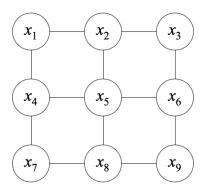
Figure 6: Example: Markov blanket

An edge  $\{u,v\}$  is a clique. A fully connected set of nodes is a clique.

Computing partition function is a source of great computational challenge:

$$Z = \int_{\mathcal{X}_V} \prod_{C \in \mathcal{C}} \psi_C(x_C).$$

In most cases, the inference involving UGM requires approximate methods.



What are the maximal cliques in this graph?

# Back to Gibbs sampling

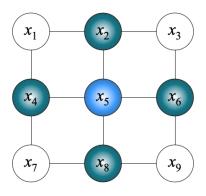
Given an UGM, determine the Markov blanket for each node v.

# Back to Gibbs sampling

Given an UGM, determine the Markov blanket for each node v.

Determine the conditional  $p(x_v|x_{mb(v)})$ .

# Back to Gibbs sampling



• Partition the nodes into disjoint sets  $A,B\subset V$  such that

$$x_u \perp x_v | B, \quad u, v \in A$$

and

$$x_u \perp x_v | A, \quad u, v \in B.$$

ullet Partition the nodes into disjoint sets  $A,B\subset V$  such that

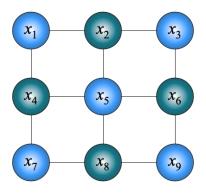
$$x_u \perp x_v | B, \quad u, v \in A$$

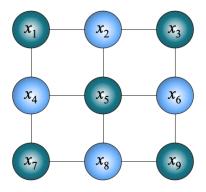
and

$$x_u \perp x_v | A, \quad u, v \in B.$$

At each iteration t = 1, ..., T:

- $\bullet \ \ \mathsf{Sample} \ p(x_A|x_{-A})\text{,}$
- Sample  $p(x_B|x_{-B})$ .





• Selecting the right inference algorithm depends on the problem's structure and computational constraints.

- Selecting the right inference algorithm depends on the problem's structure and computational constraints.
- MCMC methods can be utilized to sample from intractable distributions.

- Selecting the right inference algorithm depends on the problem's structure and computational constraints.
- MCMC methods can be utilized to sample from intractable distributions.
- Metropolis-Hastings provides general sampling but requires careful proposal design for efficiency.

- Selecting the right inference algorithm depends on the problem's structure and computational constraints.
- MCMC methods can be utilized to sample from intractable distributions.
- Metropolis-Hastings provides general sampling but requires careful proposal design for efficiency.
- Gibbs sampling is efficient when conditional distributions are easy to sample from, leveraging local dependencies.

- Selecting the right inference algorithm depends on the problem's structure and computational constraints.
- MCMC methods can be utilized to sample from intractable distributions.
- Metropolis-Hastings provides general sampling but requires careful proposal design for efficiency.
- Gibbs sampling is efficient when conditional distributions are easy to sample from, leveraging local dependencies.
- MRFs serve as a foundation for probabilistic inference, particularly in structured probabilistic models.

#### Applications of UGMs

- Neuroscience and associative memory: Hopfield networks (1982, 1984).
  - ▶ Energy-based models for pattern recognition and memory retrieval.

## Applications of UGMs

- Neuroscience and associative memory: Hopfield networks (1982, 1984).
  - ▶ Energy-based models for pattern recognition and memory retrieval.
- Deep learning: Restricted Boltzmann Machines (1986,2006).
  - Probabilistic generative models used in unsupervised pre-training of deep networks. Inspired contrastive divergence and other energy-based models in deep learning: the first "deep" neural network.

## Applications of UGMs

- Neuroscience and associative memory: Hopfield networks (1982, 1984).
  - ▶ Energy-based models for pattern recognition and memory retrieval.
- Deep learning: Restricted Boltzmann Machines (1986,2006).
  - Probabilistic generative models used in unsupervised pre-training of deep networks. Inspired contrastive divergence and other energy-based models in deep learning: the first "deep" neural network.
- Natural language processing and large language models (2017)
  - ▶ GPT-based large language models capture token dependencies.
  - ▶ The architecture is not a UGM (transformers use self-attention) but GPT learns long-range dependencies between tokens (subword) in non-sequential manner (not directional).