# Probability Theory – Exercise 1

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# Problem 1

Let  $\{\mathcal{M}_{\alpha}\}$  be an arbitrary collection of  $\sigma$ -fields of E. Show that  $\bigcap_{\alpha} \mathcal{M}_{\alpha}$  is a  $\sigma$ -field.

Proof. Let  $\mathcal{M} = \bigcap_{\alpha} \mathcal{M}_{\alpha}$ . To show  $\mathcal{M}$  is  $\sigma$ -field, we need to check that  $\mathcal{M}$  contains  $\emptyset$  and closed under countable unions and complements. Since  $\mathcal{M}_{\alpha}$  is  $\sigma$ -field,  $\emptyset \in \mathcal{M}_{\alpha}$  for all  $\alpha$ . So,  $\emptyset \in \mathcal{M} = \bigcap_{\alpha} \mathcal{M}_{\alpha}$ . Now let  $A_i \in \mathcal{M}$  for  $i = 1, 2, 3, \cdots$ . If we show that  $\bigcup_i A_i \in \mathcal{M}$  and  $A_i^c \in \mathcal{M}$  for any i, we are done. Since  $A_i \in \mathcal{M}$ ,  $A_i \in \mathcal{M}_{\alpha}$  for all  $\alpha$ . Then  $\bigcup_i A_i$  and  $A_i^c$  are contained in  $\mathcal{M}_{\alpha}$  for all  $\alpha$ , i. Hence  $\{\bigcup_i A_i, A_i^c\} \subset \mathcal{M}$  for any i.

#### Problem 2

Show that if  $E_1$  and  $E_2$  are measurable sets in  $\mathbb{R}$ , then  $m(E_1 \cup E_2) + m(E_1 \cap E_2) = m(E_1) + m(E_2)$ .

Proof. If either  $E_1$  or  $E_2$  have infinite measure, then given equality holds since  $m(E_1 \cup E_2) + m(E_1 \cap E_2) = \infty = m(E_1) + m(E_2)$ . Without loss of generality, assume  $E_1$  and  $E_2$  have finite measure. Note that  $\mathcal{M}$  is closed under countable unions, countable intersections, complements and we can write  $E_1 \cup E_2$  as a union of pairwise disjoint sets  $E_1 \cup E_2 = (E_1 \setminus E_2) \cup (E_1 \cap E_2) \cup (E_2 \setminus E_1)$ . Thus by countable additivity,

$$m(E_1) + m(E_2) = m((E_1 \setminus E_2) \cup (E_1 \cap E_2)) + m((E_2 \setminus E_1) \cup (E_1 \cap E_2))$$

$$= m(E_1 \setminus E_2) + m(E_1 \cap E_2) + m(E_2 \setminus E_1) + m(E_1 \cap E_2)$$

$$= m((E_1 \setminus E_2) \cup (E_1 \cap E_2) \cup (E_2 \setminus E_1)) + m(E_1 \cap E_2)$$

$$= m(E_1 \cup E_2) + m(E_1 \cap E_2)$$

# Problem 3

Let  $(\Omega, \mathcal{F}, P)$  be a probability space. Prove that if  $H_i$  are pairwise disjoint events such that  $\bigcup_{i=1} H_i = \Omega$ ,  $P(H_i) \neq 0$ , then

$$P(A) = \sum_{i=1}^{\infty} P(A|H_i)P(H_i).$$

*Proof.* Since  $A \subset \Omega$ ,  $P(A) = P(A \cap \Omega) = P(A \cap (\bigcup_{i=1}^{\infty} H_i))$ . Since  $H_i$  are pairwise disjoint,  $A \cap H_i$  are also pairwise disjoint. By countable additivity of probability measure and definition of conditional probability,

$$P(A) = P(A \cap (\bigcup_{i=1}^{\infty} H_i)) = P(\bigcup_{i=1}^{\infty} (A \cap H_i)) = \sum_{i=1}^{\infty} P(A \cap H_i) = \sum_{i=1}^{\infty} P(A|H_i)P(H_i).$$

#### Problem 4

Let  $\{f_n\}$  be a sequence of measurable functions defined on an interval [a,b]. Suppose that there exists an itegrable function g on [a,b] such that  $f_n \leq g$  for all n. Show that

$$\int_{a}^{b} \limsup_{n \to \infty} f_n \, \mathrm{d}m \ge \limsup_{n \to \infty} \int_{a}^{b} f_n \, \mathrm{d}m.$$

*Proof.* Let E be an interval [a, b]. Since  $g - f_n$  is non-negative measurable functions, by Fatou's lemma, we get

$$\liminf_{n \to \infty} \int_{E} (g - f_n) dm \ge \int_{E} \left( \liminf_{n \to \infty} (g - f_n) \right) dm.$$

By the fact that  $\liminf(-f_n) = -\limsup f_n$ , we can write LHS as

$$\liminf_{n \to \infty} \int_E (g - f_n) \, \mathrm{d}m = \liminf_{n \to \infty} \left( \int_E g \, \mathrm{d}m - \int_E f_n \, \mathrm{d}m \right) = \int_E g \, \mathrm{d}m - \limsup_{n \to \infty} \int_E f_n \, \mathrm{d}m$$

and RHS as

$$\int_{E} \left( \liminf_{n \to \infty} (g - f_n) \right) dm = \int_{E} \left( g - \limsup_{n \to \infty} f_n \right) dm = \int_{E} g dm - \int_{E} \limsup_{n \to \infty} f_n dm.$$

Hence, we get

$$\int_{E} \limsup_{n \to \infty} f_n \, \mathrm{d}m \ge \limsup_{n \to \infty} \int_{E} f_n \, \mathrm{d}m.$$

## Problem 5

Let  $\{f_n\}$  be a sequence of integrable functions on a set E such that  $f_n \to f$  a.e. with f integrable. Show that

$$\int_{E} |f_{n} - f| \, \mathrm{d}m \to 0 \text{ if and only if } \int_{E} |f_{n}| \, \mathrm{d}m \to \int_{E} |f| \, \mathrm{d}m.$$

Proof. By triangle inequality,  $|f_n| = |f + f_n - f| \le |f| + |f_n - f|$ . Let  $g_n = |f_n| - |f_n - f| - |f| \le 0$ . Note that  $g_n$  converges to 0 a.e. since  $f_n \to f$  a.e. Then  $|g_n| = |f| + |f_n - f| - |f_n| \le |f| + |f_n| + |f| - |f_n| = 2|f|$ . Since f is integrable,  $|g_n|$  is also integrable. Then, by the LDCT,

$$\lim_{n \to \infty} \int_E |g_n| \, \mathrm{d}m = \int_E \lim_{n \to \infty} |g_n| \, \mathrm{d}m = \int_E 0 \, \mathrm{d}m = 0.$$

Thus we get,

$$\begin{split} \lim_{n \to \infty} \int_E |g_n| \, \mathrm{d}m &= \lim_{n \to \infty} \int_E (|f| + |f_n - f| - |f_n|) \, \mathrm{d}m \\ &= \lim_{n \to \infty} \Big( \int_E |f| \, \mathrm{d}m + \int_E |f_n - f| \, \mathrm{d}m - \int_E |f_n| \, \mathrm{d}m \Big) \\ &= \int_E |f| \, \mathrm{d}m + \lim_{n \to \infty} \int_E |f_n - f| \, \mathrm{d}m - \lim_{n \to \infty} \int_E |f_n| \, \mathrm{d}m = 0. \end{split}$$

This exactly proves our claim.

## Problem 6

Let  $(\Omega, \mathcal{F}, P)$  be a probability space and  $X : \Omega \to \mathbb{R}$  a random variable. Define a distribution function of X by

$$F_X(y) = P(\{\omega : X(\omega) \le y\}) = P_X((-\infty, y]).$$

Prove that  $F_X$  is continuous if and only if  $P_X(\{y\}) = 0$  for all  $y \in \mathbb{R}$ .

*Proof.* Since  $F_X$  is right continuous, to show  $F_X$  is continuous, ETS  $F_X$  is left continuous. Let  $\{a_n\}$  be an arbitrary sequences such that  $a_n \geq 0$  and  $\lim_{n \to \infty} a_n = 0$ . If  $F_X$  is continuous,  $F_X$  is left continuous. Then  $F_X(y) - \lim_{n \to \infty} F_X(y - a_n) = 0$ , which is  $P_X(\{y\})$ . Therefore,  $P_X(\{y\}) = 0$  for all  $y \in \mathbb{R}$ .

Conversely, if 
$$P_X(\{y\}) = 0$$
,  $P_X(\{y\}) = \lim_{n \to \infty} P_X((y - a_n, y]) = F_X(y) - \lim_{n \to \infty} F_X(y - a_n) = 0$ . Thus,  $F_X$  is left continuous.

### Problem 7

Find the distribution function  $F_X$  and the expectation for a random variable X on a probability space  $([0,1], \mathcal{M}_{[0,1]}, m_{[0,1]})$  where  $X(\omega) = \min\{\omega, 1 - \omega\}$ .

*Proof.* Let  $P = m_{[0,1]}$  and  $P_X(B) = P(X^{-1}(B))$  for Borel set B.

$$\begin{split} F_X(y) &= P(\{\omega \in [0,1] : \min\{\omega, 1-\omega\} \leq y\}) = P_X([-\infty,y]) = 2 \int_{-\infty}^y \mathbf{1}_{[0,\frac{1}{2}]}(\omega) \, \mathrm{d}\omega \\ &= \begin{cases} 0 & \text{if } y < 0 \\ 2y & \text{if } 0 \leq y \leq \frac{1}{2} \\ 1 & \text{if } \frac{1}{2} < y \end{cases} \\ f_X(y) &= \frac{\mathrm{d}F_X(y)}{\mathrm{d}y} = 2\mathbf{1}_{[0,\frac{1}{2}]}(y) \\ \mathbb{E}(X) &= \int_0^1 X \, \mathrm{d}P = \int_{\mathbb{R}} x \, \mathrm{d}P_X(x) = \int x f_X(x) \, \mathrm{d}x = \int x 2\mathbf{1}_{[0,\frac{1}{2}]}(x) \, \mathrm{d}x = \int_0^{\frac{1}{2}} 2x \, \mathrm{d}x = x^2 \big|_0^{\frac{1}{2}} = \frac{1}{4} \end{cases} \end{split}$$

Since the density  $f_X(x)=2\mathbf{1}_{[0,\frac{1}{2}]}(x),\,X$  follows uniform distribution within an interval  $[0,\frac{1}{2}].$