

Mathematical Statistics II

Ch.5 Distributions of Functions of Random Variables

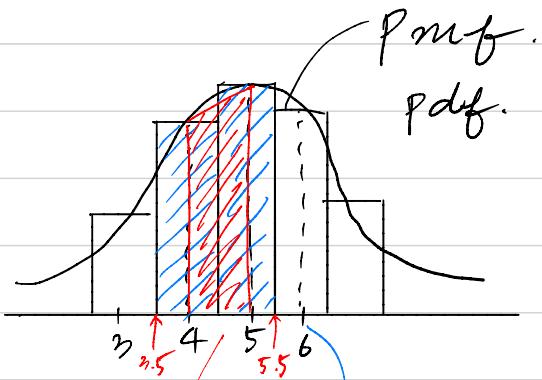
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Ch5.7 Approximations for Discrete Distributions



Qf). In discrete case,

$$P(4 \leq X \leq 5) = P(X=4) + P(X=5)$$

In continuous case,

$$P(\underset{>0}{4} \leq X \leq \underset{=0}{5}) + P(\underset{-0}{X=4}) + P(\underset{-0}{X=5}).$$

discrete case $\xrightarrow{\text{adjust}} \text{continuous case}$
adjust . prob

In discrete case,
 $P(4 < X < 5) = P(3.5 \leq X \leq 5.5)$

Apply CLT

* When we apply the CLT in discrete case,
we need to adjust

$$P(X=4) = P(3.5 \leq X \leq 4.5) \quad \text{apply CLT}$$

$$P(\underset{= P(X=5)}{4 < X \leq 5}) = P(4.5 \leq X \leq 5.5).$$

* $P(X=4)$. $P(Z \geq Z_i = 4)$

$$\text{CLT 用途 } P\left(\frac{\sum X_i - \mu}{\sigma/\sqrt{n}} = Z = 4\right) = 0$$

zero 71-84%.

Approximation of Binomial distribution

Thm.

Let $X_1, \dots, X_n \sim \text{Ber}(p)$. Then $Y = \sum_{i=1}^n X_i \sim \text{Bin}(n, p)$.

$$W = \frac{Y - np}{\sqrt{np(1-p)}} = \frac{\bar{X} - p}{\sqrt{p(1-p)/n}} \rightarrow N(0, 1) \text{ as } n \rightarrow \infty,$$

That means that $Y \sim N(np, np(1-p))$ as $n \rightarrow \infty$. ($np \geq 5$ and $n(1-p) \geq 5$)

6/1, 1 X .

Ber(p).
 $X_1 \dots X_n$
 \star 独立 \rightarrow Bin
 $\frac{Y}{n} \sim \text{Bin}(n, p)$

$$\frac{\sum_{i=1}^n X_i - np}{\sqrt{np(1-p)}} \xrightarrow{n \rightarrow \infty} N(0, 1).$$

$$P(\sum X_i > 0).$$

$$Y = \sum_{i=1}^n X_i \sim \text{Bin}(n, p)$$

$$E[Y] = np, \quad \text{Var}[Y] = np(1-p)$$

$$\frac{Y - np}{\sqrt{np(1-p)}} \sim N(0, 1) \text{ as } n \rightarrow \infty \text{ by CLT.}$$

$$\frac{\sum X_i - np}{\sqrt{np(1-p)}} = \frac{\bar{X} - p}{\sqrt{p(1-p)/n}}$$

$$Y = \sum_{i=1}^n X_i \sim N(np, np(1-p))$$

Note - correction for continuity

In general, $Y \sim \text{Bin}(n, p)$ where n is quite large.

Oliver
9/7/14 X

$$P(Y \leq k) \approx \Phi\left(\frac{k - np + 0.5}{\sqrt{np(1-p)}}\right)$$

$$P(Y < k) \approx \Phi\left(\frac{k - np - 0.5}{\sqrt{np(1-p)}}\right)$$

Example 5.7-4

$Y \sim \text{Bin}(36, 0.5)$. Compute the approximate probability of
 $P(Y = 20)$ and $P(12 < Y \leq 18)$.

have to adjust

$$X_i \sim \text{Ber}(0.5)$$

$$Y = \sum_{i=1}^n X_i \sim (np, np(1-p)).$$
$$(36 \times 0.5, 36 \times 0.5^2)$$

$$P(Y = 20) = P(19.5 \leq Y \leq 20.5)$$

$$= P\left(\frac{19.5 - 18}{\sqrt{36 \times 0.5^2}} \leq \frac{Y - 18}{\sqrt{36 \times 0.5^2}} \leq \frac{20.5 - 18}{\sqrt{36 \times 0.5^2}}\right)$$
$$\frac{Y - 18}{\sqrt{36 \times 0.5^2}} \sim N(0, 1)$$

Apply CLT,

$$\approx P\left(\frac{1.5}{3} \leq Z \leq \frac{2.5}{3}\right)$$

Approximation of Poisson distribution

Example 5.7-5

Let $X_1, \dots, X_{20} \sim \text{Poi}(1)$ and $Y = \sum_{i=1}^n X_i$. Compute the approximate probability of $P(16 < Y \leq 21)$.

$$X_1, \dots, X_{20} \sim \text{Poi}(1) \quad Y = \sum_{i=1}^n X_i$$

$$P(16 < Y \leq 21).$$

pmf of $\text{Poi}(\lambda)$ = $\frac{e^{-\lambda} \cdot \lambda^x}{x!}$

mgt of $\text{Poi}(\lambda)$ = $\exp\{\lambda(e^t - 1)\}$

$$M_Y(t) = E[e^{tY}] = E[e^{t \sum X_i}]$$

$$= E[e^{tx_1 + tx_2 + \dots + tx_n}]$$

$$= \prod_{i=1}^n E[e^{tx_i}] = \prod_{i=1}^n M_{X_i}(t) = [M_{X_i}(t)]^n$$

$$= [\exp\{\lambda(e^t - 1)\}]^n$$

= $\exp\{n\lambda(e^t - 1)\}$

: mgt of $\text{Poi}(n\lambda)$

$$\therefore Y \sim \text{Poi}(20).$$

$$P(16 < Y \leq 21) = P(17 \leq Y \leq 21) = P(16.5 \leq Y \leq 21.5)$$

$$E[Y] = 20, \quad \text{Var}[Y] = 20.$$

$$P\left(\frac{16.5 - 20}{\sqrt{20}} \leq Z \leq \frac{21.5 - 20}{\sqrt{20}}\right)$$

$$= P\left(\frac{-3.5}{\sqrt{20}} \leq Z \leq \frac{1.5}{\sqrt{20}}\right).$$

$$Z \sim N(0,1).$$

Ch5.8 Chebyshev's Inequality

Chebyshev's Inequality

Thm. 5.8-1 (Chebyshev's Inequality)

If the random variable X has a mean μ and variance σ^2 , then for every $k \geq 1$,

$$P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$



Corollary 5.8-1

If $\epsilon = k\sigma$, then

$$P(|X - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2}$$

$$\frac{1}{k^2} = \frac{\sigma^2}{\epsilon^2} \quad k\sigma = \epsilon \quad k^2 = \frac{\epsilon^2}{\sigma^2}$$

In computing a probability

- i) if dist^n is known, we can obtain the exact one.
- ii) if "unknown but n is large,
 $P(X)$ or $P(I_{X_i})$ can be approximately calculated (by CLT).
- iii) if σ unknown & n is small or single r.v
is contained in prob.

$$P(X \geq 3.5)$$

$$X \sim N(3, 2^2)$$

$$X \sim (\mu, \sigma^2)$$

$$\underline{P(|X-\mu| \geq k\sigma) \leq \frac{1}{k^2}}$$

pf). Assume that X is discrete r.v., following $f_X(x)$ pmf.

$$\sigma^2 = \text{Var}(X) = E[(X-\mu)^2] = \sum_{x \in X} (x-\mu)^2 f_X(x)$$

$$= \underbrace{\sum_{x \in A} (x-\mu)^2 f_X(x)}_{+ \geq 0 +} + \underbrace{\sum_{x \in A^c} (x-\mu)^2 f_X(x)}_{+ \geq 0 +}$$

$$\geq \sum_{x \in A} (x-\mu)^2 f_X(x) \quad \dots \textcircled{1}$$

$$\text{where } A = \{x \mid |x-\mu| \geq k\sigma\}.$$

Within A ,

$$(x-\mu)^2 \geq (k\sigma)^2, \text{ for } \forall x \in A$$

$$\Rightarrow (x-\mu)^2 f_X(x) \geq k^2 \sigma^2 f_X(x) \text{ for } \forall x \in A$$

$$\Rightarrow \sum_{x \in A} (x-\mu)^2 f_X(x) \geq \sum_{x \in A} k^2 \sigma^2 f_X(x) = k^2 \sigma^2 P(X \in A) \dots \textcircled{2}$$

$$\sum_{x \in A} f(x)$$

By $\textcircled{1}$ & $\textcircled{2}$

$$\sigma^2 \geq \sum_{x \in A} (x-\mu)^2 f_X(x) \geq k^2 \sigma^2 P(X \in A)$$

$$\therefore P(X \in A) = P(|X-\mu| \geq k\sigma) \leq \frac{1}{k^2} \quad \text{回}$$

$$\epsilon = k\sigma \Rightarrow k = \frac{\epsilon}{\sigma}$$

$$P(X \in A) \leq \frac{1}{k^2}$$

$$P(|X-\mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2}$$

Example 5.8-1

Let X have a mean of 25 and a variance of 16. Find the lower bound for $P(17 < X < 33)$.

$$(25, 4^2). \quad \sigma = 4.$$

$$P(|X-25| \geq 4 \cdot k) \leq \frac{1}{k^2} \quad \dots \textcircled{Q}$$

$$P(17 < X < 33)$$

$$= P(-8 < X - 25 < 8) = P(|X - 25| < 8)$$

$$\text{by } \textcircled{Q} \quad P(|X - 25| \geq 8) \leq \frac{1}{4}$$
$$P(|X - 25| < 8) \geq \frac{3}{4}$$

$$= P(|X - 25| < 2 \cdot 6) \geq 1 - \frac{1}{4} = \frac{3}{4} = 0.75$$

$$P(|X - 25| \geq 12)$$
$$= P(|X - \mu| \geq 3 \cdot 6) \leq \frac{1}{3^2} = \frac{1}{9}$$

$$P(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

$$\text{if } \epsilon = k\sigma, P(|X - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2}$$

$$X \sim (25, 4^2)$$

$$\begin{aligned} P(17 < X < 33) &= P(17 - 25 < X - 25 < 33 - 25) \\ &= P(|X - 25| < 8) = 1 - P(|X - 25| \geq 8) \leq \frac{1}{4} \\ \therefore P(|X - 25| < 8) &\geq 1 - \frac{1}{4} = \frac{3}{4} \end{aligned}$$

[Definition] Convergence in probability

For any $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} P(|X_n - X| < \epsilon) = 1$$

Then, X_n converges in probability to X .

$$X_n \xrightarrow{P} X \iff \lim_{n \rightarrow \infty} P(|X_n - X| < \varepsilon) = 1$$

\bar{X}_n : sample mean from $X_i \sim (\mu, \sigma^2)$

$$\bar{X}_n \sim ?(\mu, \frac{\sigma^2}{n})$$

$$P(|\bar{X}_n - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{n\varepsilon^2}$$

$$0 \leq \lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| \geq \varepsilon) \leq \lim_{n \rightarrow \infty} \frac{\sigma^2}{n \cdot \varepsilon^2} = 0$$

$$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| \geq \varepsilon) = 0$$

$$\iff \lim_{n \rightarrow \infty} P(|\bar{X} - \mu| < \varepsilon) = 1.$$

