

Homework 3

Junwoo Yang

October 18, 2020

Problem 1. Let $\text{Var}(X)$ be the variance of a random variable X .

- (a) Find $\text{Var}(aX)$ in terms of $\text{Var}(X)$.
- (b) Find $\text{Var}(X)$ of $X : [0, 1] \rightarrow \mathbb{R}$ given by $X(\omega) = \min\{\omega, 1 - \omega\}$.
- (c) If a_1, \dots, a_n, b are arbitrary real numbers and X_1, \dots, X_n are random variables, show

$$\text{Var}(a_1X_1 + \dots + a_nX_n + b) = \sum_{i=1}^n a_i^2 \text{Var}(X_i) + 2 \sum_{1 \leq i < j \leq n} a_i a_j \text{Cov}(X_i, X_j).$$

Proof. (a)

$$\text{Var}(aX) = \mathbb{E}(a^2X^2) - \mathbb{E}(aX)^2 = a^2\mathbb{E}(X^2) - a^2\mathbb{E}(X)^2 = a^2(\mathbb{E}(X^2) - \mathbb{E}(X)^2) = a^2\text{Var}(X). \quad \square$$

(b)

$$F_X(y) = P(\{\omega \in [0, 1] : \min\{\omega, 1 - \omega\} \leq y\}) = 2 \int_{-\infty}^y \mathbf{1}_{[0, \frac{1}{2}]}(\omega) d\omega$$

$$= \begin{cases} 0 & \text{if } y < 0 \\ 2y & \text{if } 0 \leq y \leq \frac{1}{2} \\ 1 & \text{if } \frac{1}{2} < y \end{cases}$$

$$f_X(y) = \frac{dF_X(y)}{dy} = 2\mathbf{1}_{[0, \frac{1}{2}]}(y)$$

$$\begin{aligned} \mathbb{E}(X) &= \int_0^1 X dP = \int_{\mathbb{R}} x dP_X(x) = \int x f_X(x) dx = \int x 2\mathbf{1}_{[0, \frac{1}{2}]}(x) dx \\ &= \int_0^{\frac{1}{2}} 2x dx = x^2 \Big|_0^{\frac{1}{2}} = \frac{1}{4} \end{aligned}$$

$$\mathbb{E}(X^2) = \int x^2 f_X(x) dx = \int x^2 2\mathbf{1}_{[0, \frac{1}{2}]}(x) dx = \int_0^{\frac{1}{2}} 2x^2 dx = \frac{2}{3} x^3 \Big|_0^{\frac{1}{2}} = \frac{1}{12}$$

$$\text{Var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 = \frac{1}{12} - \frac{1}{16} = \frac{1}{48}$$

- (c) Let $Z = a_1X_1 + \dots + a_nX_n + b = \sum_{i=1}^n a_iX_i + b$. Then, we get that

$$\mathbb{E}(Z) = \sum_{i=1}^n a_i \mathbb{E}(X_i) + b$$

$$Z^2 = \sum_{i=1}^n a_i^2 X_i^2 + 2 \sum_{1 \leq i < j \leq n} a_i a_j X_i X_j + 2b \sum_{i=1}^n a_i X_i + b^2$$

$$\begin{aligned}
\mathbb{E}(Z^2) &= \sum_{i=1}^n a_i^2 \mathbb{E}(X_i^2) + 2 \sum_{1 \leq i < j \leq n} a_i a_j \mathbb{E}(X_i X_j) + 2b \sum_{i=1}^n a_i \mathbb{E}(X_i) + b^2 \\
\mathbb{E}(Z)^2 &= \sum_{i=1}^n a_i^2 \mathbb{E}(X_i)^2 + 2 \sum_{1 \leq i < j \leq n} a_i a_j \mathbb{E}(X_i) \mathbb{E}(X_j) + 2b \sum_{i=1}^n a_i \mathbb{E}(X_i) + b^2 \\
\text{Var}(Z) &= \mathbb{E}(Z^2) - \mathbb{E}(Z)^2 \\
&= \sum_{i=1}^n a_i^2 (\mathbb{E}(X_i^2) - \mathbb{E}(X_i)^2) + 2 \sum_{1 \leq i < j \leq n} a_i a_j (\mathbb{E}(X_i X_j) - \mathbb{E}(X_i) \mathbb{E}(X_j)) \\
&= \sum_{i=1}^n a_i^2 \text{Var}(X_i) + 2 \sum_{1 \leq i < j \leq n} a_i a_j \text{Cov}(X_i, X_j).
\end{aligned}$$

Problem 2. Find the correlation $\rho_{X,Y}$ if $X = 2Y + 1$.

Proof.

$$\begin{aligned}
\mathbb{E}(X) &= \mathbb{E}(2Y + 1) = 2\mathbb{E}(Y) + 1 \\
X - \mathbb{E}(X) &= 2(Y - \mathbb{E}(Y)) \\
\text{Var}(X) &= \mathbb{E}((X - \mathbb{E}(X))^2) = 4\mathbb{E}((Y - \mathbb{E}(Y))^2) = 4\text{Var}(Y) \\
\text{Cov}(X, Y) &= \mathbb{E}((X - \mathbb{E}(X))(Y - \mathbb{E}(Y))) = 2\mathbb{E}((Y - \mathbb{E}(Y))^2) = 2\text{Var}(Y) \\
\rho_{X,Y} &= \frac{\text{Cov}(X, Y)}{\text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}}} = \frac{2\text{Var}(Y)}{2\text{Var}(Y)} = 1 \quad \square
\end{aligned}$$

Problem 3. Find F_X the distribution function of a random variable $X : [0, 1] \rightarrow \mathbb{R}$ given by $X(\omega) = \min\{\omega, 1 - \omega\}$.

Proof.

$$F_X(y) = P(\{\omega \in [0, 1] : \min\{\omega, 1 - \omega\} \leq y\}) = 2 \int_{-\infty}^y \mathbf{1}_{[0, \frac{1}{2}]}(\omega) d\omega = \begin{cases} 0 & \text{if } y < 0 \\ 2y & \text{if } 0 \leq y \leq \frac{1}{2} \\ 1 & \text{if } \frac{1}{2} < y \end{cases} \quad \square$$

Problem 4. Suppose that X, Y are independent random variables and that f, g are Borel measurable functions on \mathbb{R} . Show that the random variables $f(X), g(Y)$ are independent.

Proof. Let (Ω, \mathcal{F}, P) be probability space. We want to show that

$$P(\{\omega \in \Omega : f(X(\omega)) \in B, g(Y(\omega)) \in C\}) = P(\{\omega : f(X(\omega)) \in B\})P(\{\omega : g(Y(\omega)) \in C\})$$

for all Borel sets B, C . Note that $f^{-1}(B)$ and $g^{-1}(C)$ are Borel sets because f, g are Borel measurable functions. Then, by definition of independence of random variables, we get that

$$\begin{aligned}
P(f(X) \in B \cap g(Y) \in C) &= P(X \in f^{-1}(B) \cap Y \in g^{-1}(C)) \\
&= P(X \in f^{-1}(B))P(Y \in g^{-1}(C)) \\
&= P(f(X) \in B)P(g(Y) \in C).
\end{aligned}$$

Therefore, random variables $f(X), g(Y)$ are independent. \square

Problem 5. Show that $|\rho_{X,Y}| = 1$ if and only if $X_c = X - \mathbb{E}(X)$ and $Y_c = Y - \mathbb{E}(Y)$ are linearly dependent, that is, $P(\{\omega \in \Omega : aX_c(\omega) + bY_c(\omega) = 0\}) = 1$ for some real numbers a and b , not both 0.

Proof. Without loss of generality, assume that $\|X_c\|_2$ and $\|Y_c\|_2$ are non-zero.

(\Rightarrow) The correlation $\rho_{X,Y}$ is cosine of angle between X_c and Y_c . $\rho_{X,Y} = \cos \theta = 1$ means that θ is an even multiple of π , i.e. X_c and Y_c have same direction. In this case, Y_c is just a positive scalar multiple of X_c , i.e. $Y_c = tX_c$ for $t \in \mathbb{R}, t > 0$. If $\rho_{X,Y} = \cos \theta = -1$, θ is an odd multiple of π . This means X_c and Y_c have opposite direction and Y_c can be written as tX_c for $t \in \mathbb{R}, t < 0$. Thus X_c and Y_c are linearly dependent.

(\Leftarrow) TFAE

$$\begin{aligned} & P(\{\omega \in \Omega : aX_c(\omega) + bY_c(\omega) = 0\}) = 1 \text{ for some } a, b \in \mathbb{R} \setminus \{0\} \\ \iff & \forall \omega \in \Omega, Y_c(\omega) = tX_c(\omega) \text{ for some } t \in \mathbb{R} \setminus \{0\} \\ \iff & Y - \mathbb{E}(Y) = t(X - \mathbb{E}(X)) \\ \iff & Y = tX - t\mathbb{E}(X) + \mathbb{E}(Y) = tX + c \text{ where } c = -t\mathbb{E}(X) + \mathbb{E}(Y) ; \text{ constant.} \end{aligned}$$

We want to show that $|\text{Cov}(X, Y)| = \text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}}$.

$$\because |\rho_{X,Y}| = \left| \frac{\text{Cov}(X, Y)}{\text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}}} \right| = \frac{|\text{Cov}(X, Y)|}{\text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}}}$$

$$\begin{aligned} \text{Cov}(X, Y) &= \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) \\ &= \mathbb{E}(tX^2 + cX) - \mathbb{E}(X)\mathbb{E}(tX + c) \\ &= t\mathbb{E}(X^2) + c\mathbb{E}(X) - t\mathbb{E}(X)^2 - c\mathbb{E}(X) \\ &= t(\mathbb{E}(X^2) - \mathbb{E}(X)^2) \\ &= t\text{Var}(X). \end{aligned}$$

$$|\text{Cov}(X, Y)| = \begin{cases} t\text{Var}(X) & \text{if } t > 0 \\ -t\text{Var}(X) & \text{if } t < 0 \end{cases}$$

$$\begin{aligned} \text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}} &= \text{Var}(X)^{\frac{1}{2}} \text{Var}(tX + c)^{\frac{1}{2}} \\ &= \text{Var}(X)^{\frac{1}{2}} \text{Var}(tX)^{\frac{1}{2}} \\ &= \sqrt{t^2} \text{Var}(X) \\ &= \begin{cases} t\text{Var}(X) & \text{if } t > 0 \\ -t\text{Var}(X) & \text{if } t < 0 \end{cases} \end{aligned}$$

Therefore, $|\text{Cov}(X, Y)| = \text{Var}(X)^{\frac{1}{2}} \text{Var}(Y)^{\frac{1}{2}}$.

□