

# Mathematical Statistics 1

## Ch.3 Continuous Distributions

Jungsoon Choi

jungsoonchoi@hanyang.ac.kr

# Table of Contents

- Uniform Distributions
  - Exponential Distributions
  - Gamma and Chi-Square Distributions
    - ( general )
- Gramma's special Distribution.*

# (Continuous) Uniform Distribution

## 2.1 Uniform Distribution, $\text{Unif}(a, b)$

interval.

- Let r.v.  $X$  denote the outcome when a point is selected at random from an interval  $[a, b]$ .
- $pdf$

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{o.w.} \end{cases}$$

$$f(x) = \frac{1}{b-a}, \quad a \leq x \leq b$$

- $cdf$

$$F(x) = \begin{cases} 0 & x < a, \\ \frac{x-a}{b-a} & a \leq x < b \\ 1 & b \leq x. \end{cases}$$

$X \sim$  discrete Uniform distribution.

$$P(X=x) = \frac{1}{6} \quad x=1, \dots, 6.$$

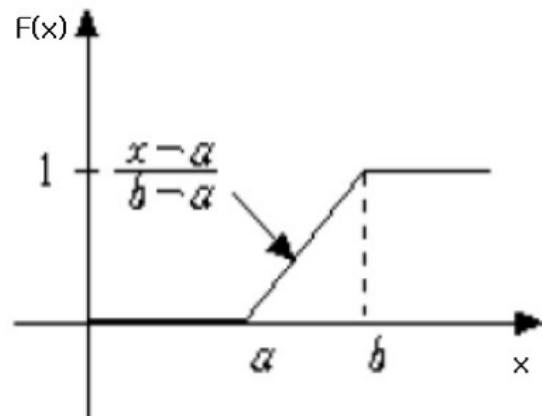
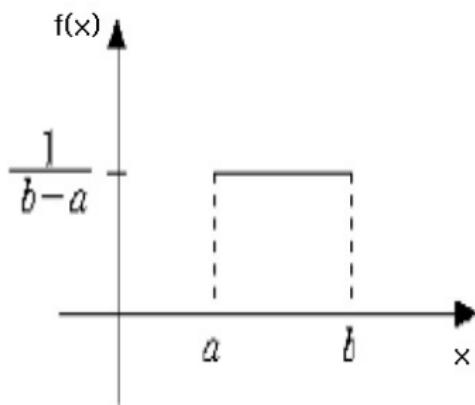
$$F_X(x) = P(X \leq x) = \int_{-\infty}^x f_X(t) dt$$

$$x < a \rightarrow 0$$

$$a \leq x \leq b \rightarrow \int_{-\infty}^x f_X(t) dt = \int_{-\infty}^a f_X(t) dt + \int_a^x f_X(t) dt \\ = \int_a^x \frac{1}{b-a} dt = \frac{x-a}{b-a}$$

$$x > b \rightarrow 1$$

## Uniform pdf and distribution function



- Mean and Variance

$$E(X) = \int_a^b \frac{x}{b-a} dx = \frac{x^2}{2(b-a)} \Big|_a^b = \frac{a+b}{2}$$

$$E(X^2) = \int_a^b \frac{x^2}{b-a} dx = \frac{x^3}{3(b-a)} \Big|_a^b = \frac{a^2 + ab + b^2}{3}$$

$$\text{Var}(X) = E(X^2) - \{E(X)\}^2 = \frac{(b-a)^2}{12}$$

- mgf

$$M(t) = E(e^{tX}) = \int_a^b \frac{e^{tx}}{b-a} dx = \frac{e^{tx}}{t(b-a)} \Big|_a^b = \frac{e^{tb} - e^{ta}}{t(b-a)}$$

Therefore,

$$M(t) = \begin{cases} \frac{e^{tb} - e^{ta}}{t(b-a)} & t \neq 0, \\ 1 & t = 0. \end{cases}$$

$X \sim \text{Unif}(a, b)$ .

$$f_X(x) = \frac{1}{b-a}$$

$$E(X) = \int_a^b \frac{x}{b-a} dx = \frac{1}{b-a} \left[ \frac{1}{2}x^2 \right]_a^b = \frac{b^2 - a^2}{2(b-a)} = \frac{b+a}{2}$$

$$E(X^2) = \int_a^b \frac{x^2}{b-a} dx = \frac{1}{b-a} \left[ \frac{1}{3}x^3 \right]_a^b = \frac{b^3 - a^3}{3(b-a)} = \frac{(b^2 + ab + a^2)}{3}$$

$$\text{Var}(X) = \frac{a^2 + ab + b^2}{3} - \frac{a^2 + 2ab + b^2}{4} = \frac{1}{12} (4a^2 + 4ab + 4b^2 - 3a^2 - 6ab - 3b^2) \\ = \frac{1}{12} (a^2 + b^2 - 2ab)$$

$$M_X(t) = E(e^{tx}) = \int_a^b e^{tx} \frac{1}{b-a} dx = \frac{1}{b-a} (a - b)^2$$

$$= \frac{1}{b-a} \int_a^b e^{tx} dx = \frac{1}{b-a} \cdot \frac{1}{t} [e^{tx}]_a^b$$

$$= \frac{e^{tb} - e^{ta}}{t(b-a)}$$

$$M(t) = \begin{cases} \frac{e^{tb} - e^{ta}}{t(b-a)} & t \neq 0 \\ 1 & t=0 \end{cases}$$

↙ definition  
→ M(t)

$$M_X^{(1)}(0) = \left. \frac{(b e^{tb} - a e^{ta}) \cdot t (b-a) - (e^{tb} - e^{ta})(b-a)}{t^2 (b-a)} \right|_{t=0}$$

=

$$f_X(x) = \frac{1}{b-a} \quad (a \leq x \leq b).$$

$$F_X(x) = \frac{x-a}{b-a}$$

$$\begin{aligned} E(X) &= \int_a^b \frac{x}{b-a} dx = \frac{1}{b-a} \left[ \frac{1}{2}x^2 \right]_a^b = \frac{b^2 - a^2}{2(b-a)} \\ &= \frac{a+b}{2} \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_a^b \frac{x^2}{b-a} dx = \frac{1}{3(b-a)} [x^3]_a^b = \frac{b^3 - a^3}{3(b-a)} \\ &= \frac{a^2 + ab + b^2}{3} \end{aligned}$$

$$\begin{aligned} \text{Var}(X) &= \frac{a^2 + ab + b^2}{3} - \frac{(a^2 + 2ab + b^2)}{4} \\ &= \frac{4a^2 + 4ab + 4b^2 - 3(a^2 + 2ab + b^2)}{12} \\ &= \frac{a^2 - 2ab + b^2}{12} = \frac{(a-b)^2}{12} = \frac{(b-a)^2}{12} \end{aligned}$$

$$M(t) = E(e^{tx}) = \frac{1}{b-a} \int_a^b e^{tx} dx = \frac{1}{t(b-a)} [e^{tx}]_a^b$$

$$= \frac{e^{tb} - e^{ta}}{t(b-a)}$$

$$\begin{cases} \frac{e^{tb} - e^{ta}}{t(b-a)} & t \neq 0 \\ 1 & t = 0 \end{cases}$$

$$M'(t) = \frac{d}{dt} \left( \frac{e^{tb} - e^{ta}}{t(b-a)} \right)$$

$$= \frac{(b e^{tb} - a e^{ta})(t(b-a)) - (e^{tb} - e^{ta})(b-a)}{t^2(b-a)^2}$$

$$X \sim \text{Unif}(a, b)$$

$$M_X(t) = \begin{cases} \frac{e^{tb} - e^{ta}}{t(b-a)}, & t \neq 0 \\ 1, & t = 0 \end{cases}$$

$$g(t) \stackrel{!}{=} e^{tb} = 1 + bt + \frac{b^2 t^2}{2!} + \frac{b^3 t^3}{3!} + \dots$$

$$\begin{aligned} \text{For } t \neq 0, \quad M_X(t) &= \frac{1}{t(b-a)} \left\{ (1 + bt + \frac{b^2 t^2}{2!} + \frac{b^3 t^3}{3!} + \dots) - \right. \\ &\quad \left. (1 + at + \frac{a^2 t^2}{2!} + \frac{a^3 t^3}{3!} + \dots) \right\} \\ &= \frac{1}{t(b-a)} \left\{ (b-a)t + \frac{(b^2 - a^2)t^2}{2!} + \frac{(b^3 - a^3)t^3}{3!} + \dots \right\} \\ &= \frac{1}{b-a} \left\{ (b-a) + \frac{(b^2 - a^2)t}{2!} + \frac{(b^3 - a^3)t^2}{3!} + \dots \right\} \\ \Rightarrow M_X^{(1)}(t) &= \frac{1}{b-a} \left\{ \frac{(b^2 - a^2)}{2!} + \frac{(b^3 - a^3)2t}{3!} + \dots \right\} \end{aligned}$$

$$\begin{aligned} M_X^{(1)}(0) &= \lim_{t \rightarrow 0} \frac{M_X(t) - M_X(0)}{t - 0} \\ &= \lim_{t \rightarrow 0} \frac{1}{b-a} \left\{ \frac{(b^2 - a^2)t}{2!} + \frac{(b^3 - a^3)t^2}{3!} + \dots \right\} - 1 \\ &= \lim_{t \rightarrow 0} \frac{\frac{(b^2 - a^2)}{2!} + \frac{(b^3 - a^3)t}{3!} + \dots}{b-a} \\ &= \frac{a+b}{2} = E(X) \end{aligned}$$

$$\begin{aligned}
 E(X^2) &= Mx^{(2)}(0) = \lim_{t \rightarrow 0} \frac{Mx''(t) - Mx''(0)}{t - 0} \quad \text{E(X) circled} \\
 &= \lim_{t \rightarrow 0} \frac{\frac{1}{b-a} \left\{ \frac{(b^2-a^2)}{2} + \frac{(b^3-a^3)(2t)}{3!} + \dots \right\} - \frac{(a+b)}{2}}{t} \\
 &= \lim_{t \rightarrow 0} \frac{\frac{(b^3-a^3)(2t)}{3!} + \frac{(b^4-a^4)(3t^2)}{4!} + \dots}{t(b-a)} \\
 &= \lim_{t \rightarrow 0} \frac{\frac{(b^3-a^3)}{3} + \frac{(b^4-a^4)(3t)}{4!} + \dots}{b-a} \\
 &= \frac{a^2+ab+b^2}{3}
 \end{aligned}$$

$$\text{Var}(X) = E(X^2) - \{E(X)\}^2 = \frac{(a-b)^2}{12}$$

# Exponential Distribution

## 2.2 Exponential Distribution, $\text{Exp}(\beta)$

- Let r.v  $X$  be **the waiting time until the first event occurs.**

- pdf*

$$\frac{1}{\beta} e^{-\frac{x}{\beta}}$$

$$f(x) = \frac{1}{\beta} e^{-x/\beta}, \quad 0 \leq x < \infty, \quad 0 < \beta.$$

- cdf*

$$\frac{1}{\beta} \int_0^x e^{-\frac{t}{\beta}} dt = \frac{1}{\beta} \left[ -\beta e^{-\frac{t}{\beta}} \right]_0^x$$

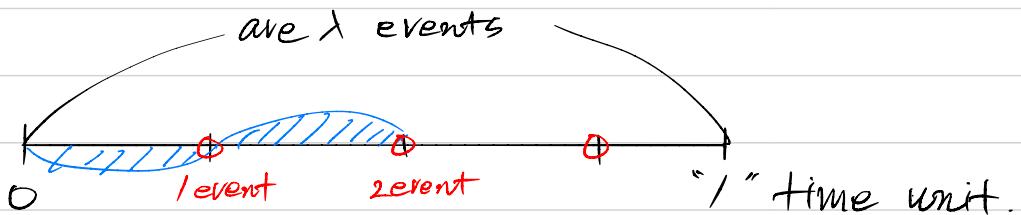
$$F(x) = \int_0^x \frac{1}{\beta} e^{-t/\beta} dt = \begin{cases} 0, & x < 0, \\ 1 - e^{-x/\beta}, & 0 \leq x < \infty. \end{cases}$$

- `dexp(x, rate=1/beta)`:  $f_X(x)$ ,  $X \sim \text{Exp}(\beta)$  in R

- `pexp(x, rate=1/beta)`:  $P(X \leq x)$ ,  $X \sim \text{Exp}(\beta)$  in R

In Poisson,  $T$ : # of events in "1" time unit with a mean of  $\lambda$   
 $\Rightarrow T \sim \text{Poi}(\lambda)$ ,  $f_T(t) = \frac{e^{-\lambda} \lambda^t}{t!}$ ,  $t=0,1,2,3,\dots$

Next, the waiting time between successive events is of interest.  
 the waiting time until the first event happens.



$X$ : the waiting time until the first event happens

$$\begin{aligned} F_X(x) &= P(X \leq x) = 1 - P(X > x) \\ &= 1 - P(\text{waiting time until the } 1^{\text{st}} \text{ event happens} \\ &\quad \text{is greater than } x) \\ &= 1 - P(\text{No event in the time interval } [0,x]). \end{aligned}$$

are  $\lambda$  events / 1 time unit \$X \sim \text{Exp}(\beta)\$

$\Leftrightarrow$  are  $\lambda x$  events /  $x$  time unit.

Assume  $T_2$ : # of events in " $x$ " time unit

$$T_2 \sim \text{Poi}(\lambda x) \rightarrow f_{T_2}(t) = \frac{e^{-\lambda x} (\lambda x)^t}{t!}, t=0,1,2,\dots$$

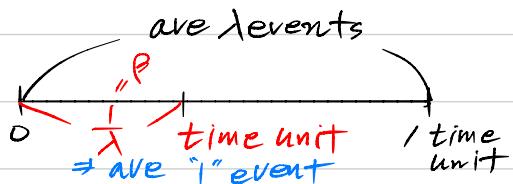
$$P(\text{No event in } [0,x]) = P(T_2=0) = e^{-\lambda x}$$

$$\therefore F_X(x) = 1 - e^{-\lambda x}$$

$$f_X(x) = F'_X(x) = \lambda e^{-\lambda x}, x \geq 0$$

$$\text{let } \frac{1}{\lambda} = \beta \quad f_X(x) = \frac{1}{\beta} e^{-\frac{x}{\beta}}, x \geq 0$$

$\beta$ : average waiting time until the  $1^{\text{st}}$  event happens.



- The first two moments

$$\begin{aligned} E(X) &= \int_0^\infty \frac{x}{\beta} \exp\left(-\frac{x}{\beta}\right) dx = \int_0^\infty \exp\left(-\frac{x}{\beta}\right) dx \\ &= -\beta \exp\left(-\frac{x}{\beta}\right) \Big|_0^\infty = \beta. \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_0^\infty \frac{x^2}{\beta} \exp\left(-\frac{x}{\beta}\right) dx = 2\beta \int_0^\infty \exp\left(-\frac{x}{\beta}\right) dx \\ &= -2\beta^2 \exp\left(-\frac{x}{\beta}\right) \Big|_0^\infty = 2\beta^2. \end{aligned}$$

- Mean:  $E(X) = \beta$

- Variance:

$$\text{Var}(X) = E(X^2) - \{E(X)\}^2 = \beta^2$$

- $\beta$ : the average waiting time until the first event occurs

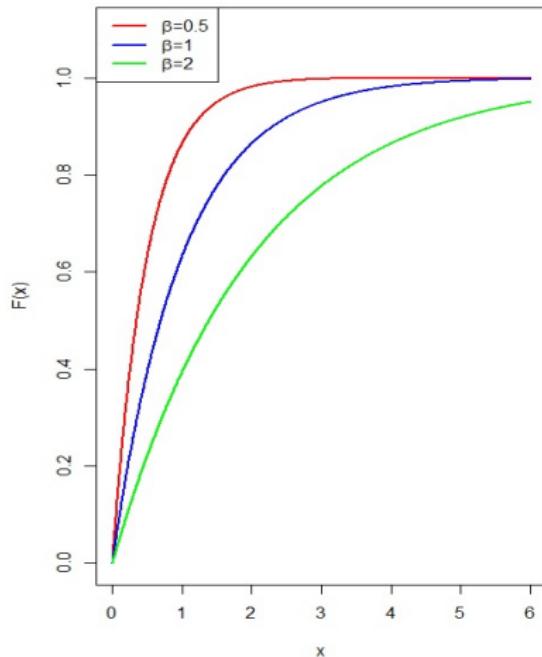
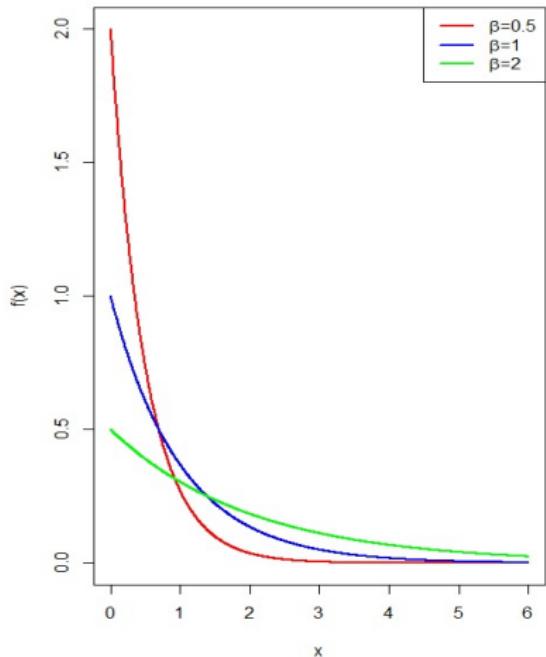
- mgf

$$\begin{aligned}
 M(t) &= \int_0^\infty \frac{e^{tx}}{\beta} \exp\left(-\frac{x}{\beta}\right) dx = \int_0^\infty \frac{1}{\beta} \exp\left[-\left(\frac{1-\beta t}{\beta}\right)x\right] dx \\
 &= \frac{1}{1-\beta t} \int_0^\infty \frac{1-\beta t}{\beta} \exp\left[-\left(\frac{1-\beta t}{\beta}\right)x\right] dx \\
 &= \frac{1}{1-\beta t}, \quad t < \frac{1}{\beta}.
 \end{aligned}$$

$$\begin{aligned}
 M^{(1)}(0) &= \frac{d}{dt} M(t) \Big|_{t=0} = \frac{d}{dt} (1-\beta t)^{-1} \Big|_{t=0} = \beta(1-\beta t)^{-2} \Big|_{t=0} = \beta \\
 M^{(2)}(0) &= \frac{d^2}{dt^2} M(t) \Big|_{t=0} = \frac{d^2}{dt^2} (1-\beta t)^{-1} \Big|_{t=0} = \frac{d}{dt} \beta(1-\beta t)^{-2} \Big|_{t=0} \\
 &= 2\beta^2(1-\beta t)^{-3} \Big|_{t=0} = 2\beta^2
 \end{aligned}$$

$$\text{Var}(X) = M^{(2)}(0) - \{M^{(1)}(0)\}^2 = \beta^2.$$

## Exponential pdf and distribution function



### Example 3.2-1

Let  $X$  have an exponential distribution with a mean of 20. Find the probability that  $X$  is less than 18. Also, find the value of  $m$  satisfying  $F(m) = 0.5$ .

$$\frac{1}{\lambda} e^{-\frac{x}{\lambda}}$$
$$\lambda e^{-\lambda x}$$
$$\cancel{\lambda e^{-\lambda x}}$$

## Relationship between Poisson and Exponential distributions

- $T \sim \text{Poi}(\lambda)$ ,  $T$ : the number of events occurring in a given time interval

$$f(t) = \frac{e^{-\lambda} \lambda^t}{t!}, \quad t = 0, 1, \dots; \quad 0 \leq \lambda < \infty$$

- The waiting times between successive events are also a random variable.
- Let  $X$  denote the waiting time until the first change occurs when we know that the mean number of events in the unit interval,  $\lambda$ .
- Then the random variable  $X$  has an Exponential distribution with parameter  $\beta = 1/\lambda$ , which is the mean waiting time for the first event.

### Example 3.2-2

Customers arrive in a certain shop according to an approximate Poisson process at a mean rate of 20 per hour. What is the probability that the shopkeeper will have to wait more than 5 minutes for the arrival of the first customer? Also, find the median time until the first arrival.

$$X \sim \text{Exp}(20)$$

$$f_X(x) = \frac{1}{20} e^{-\frac{x}{20}}, x \geq 0$$

$$1) P(X < 18) = \int_0^{18} \frac{1}{20} e^{-\frac{x}{20}} dx =$$

$$2) F_X(m) = 1 - e^{-\frac{m}{20}} = 0.5$$

$$\Rightarrow m = -20 \ln(0.5).$$

$$X \sim \text{Exp}(3)$$

$$\frac{1}{3} = m.$$

$$\int_0^{\infty} \frac{1}{3} e^{-\frac{x}{3}} dx$$

$$\text{Exp}(3)$$

$X$ : # of customers per "1" hour

$$X \sim \text{Poi}(20)$$

$$\text{Poi}(3)$$

$y$ : waiting time (min) until the first customer arrives.

$$\begin{aligned} &\text{ave 20 customers / 60 min} \\ \Leftrightarrow &\text{ave 1 customer / } \underline{3 \text{ min}} \end{aligned}$$

$$y \sim \text{Exp}(3)$$

$$P(y > 5), F_Y(m) = 0.5$$

$$\frac{1}{\sqrt{n}} \cdot \frac{1}{\sqrt{m}}$$

$$f(x) = \frac{1}{3} e^{-\frac{1}{3}x}, 0 \leq x < \infty.$$

$$\text{Exp}(3)$$

$$P(X > 5) = \int_5^\infty \frac{1}{3} e^{-\frac{1}{3}x} dx$$

$$-\left[e^{-\frac{1}{3}x}\right]_5^\infty = -e^{-\frac{5}{3}}$$

$$m =$$

$$\int_0^m \frac{1}{3} e^{-\frac{1}{3}x} dx = 0.5$$

$$-\left(e^{-\frac{m}{3}} - 1\right)$$

$$-\ln 2 = -\frac{m}{3}$$

$$1 - e^{-\frac{m}{3}} = \frac{1}{2}$$

$$\frac{1}{2} = e^{-\frac{m}{3}}$$

$$m = \frac{3}{2} \ln 2$$

# Gamma and Chi-Square Distributions

## 2.3.1 Gamma Distribution, $\text{Gamma}(\alpha, \beta)$

- Gamma function

$$\Gamma(\alpha) = \int_0^{\infty} t^{\alpha-1} e^{-t} dt > 0, \quad \alpha > 0$$

- If  $\alpha > 1$ ,  $\Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1)$

$$\begin{aligned}\Gamma(\alpha) &= [-t^{\alpha-1} e^{-t}]_0^{\infty} + \int_0^{\infty} (\alpha - 1)t^{\alpha-2} e^{-t} dt \\ &= (\alpha - 1) \int_0^{\infty} t^{\alpha-2} e^{-t} dt = (\alpha - 1)\Gamma(\alpha - 1).\end{aligned}$$

- For integer  $n$ ,  $\Gamma(n) = (n - 1)!$ ,  $\Gamma(1) = 1$ .

- Let r.v.  $X$  be **the waiting time until the  $\alpha$ th event occurs.**
- pdf**

$$f(x) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}, \quad 0 < x < \infty, 0 < \alpha, \beta,$$

where  $\Gamma(\alpha) = (\alpha - 1)!$  if  $\alpha$  is an integer.

- $Exp(\beta) \equiv Gamma(1, \beta)$
- `dgamma(x, shape=alpha, scale=beta):  $f_X(x)$  in R`
- `pgamma(x, shape=alpha, scale=beta):  $P(X \leq x)$  in R`

- The first two moments

$$\begin{aligned}
 E(X) &= \int_0^\infty x f(x) dx = \int_0^\infty \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{(\alpha+1)-1} \exp\left(-\frac{x}{\beta}\right) dx \\
 &= \alpha\beta \int_0^\infty \frac{1}{\Gamma(\alpha+1)\beta^{\alpha+1}} x^{(\alpha+1)-1} \exp\left(-\frac{x}{\beta}\right) dx = \alpha\beta. \\
 E(X^2) &= \int_0^\infty x^2 f(x) dx = \int_0^\infty \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{(\alpha+2)-1} \exp\left(-\frac{x}{\beta}\right) dx \\
 &= \alpha(\alpha+1)\beta^2 \int_0^\infty \frac{1}{\Gamma(\alpha+2)\beta^{\alpha+2}} x^{(\alpha+2)-1} \exp\left(-\frac{x}{\beta}\right) dx \\
 &= \alpha(\alpha+1)\beta^2.
 \end{aligned}$$

- Variance

$$\text{Var}(X) = E(X^2) - \{E(X)\}^2 = \alpha(\alpha+1)\beta^2 - (\alpha\beta)^2 = \alpha\beta^2.$$

$$\sum f(x) = 1. \quad \int f(x) dx = 1$$

$\rightarrow X$ : waiting time until the  $\alpha^{\text{th}}$  event happens

$$E(X) = \int_0^\infty \frac{1}{T(x) \cdot \beta^x x^\alpha} e^{-\frac{x}{\beta}} dx$$

$$\begin{aligned} &= \frac{T(\alpha+1) \beta^{\alpha+1}}{T(\alpha) \cdot \beta^\alpha} \int_0^\infty \frac{1}{T(\alpha+1) \beta^{\alpha+1}} x^{(\alpha+1)-1} e^{-\frac{x}{\beta}} dx \\ &= \alpha \beta \quad E(X) = 2\beta \end{aligned}$$

$\downarrow$  pdf of Gamma ( $\alpha+1, \beta$ )

$$X \sim \text{Exp}(\beta)$$

$X$ : waiting time until 1<sup>st</sup> event happens.

$$\Rightarrow E(X) = \beta$$

$\uparrow$  average waiting time until the 1<sup>st</sup> event.

$$\int \underline{\quad} dx$$

$\downarrow$  pdf of Gamma

$$1 - \beta t > 0$$

$$1 > \beta t \rightarrow t < \frac{1}{\beta}$$

$$(\alpha, \frac{\beta}{1 - \beta t})$$

$$X \sim \text{Gamma}(\alpha, \beta)$$

$$X > 0 \sim \chi^2(1)$$

$\left[ \begin{array}{l} \alpha : \text{shape parameter} \\ \beta : \text{scale} \end{array} \right]$

$$\text{Gamma}\left(\frac{r}{2}, 2\right)$$

• mgf

$$\begin{aligned}
 M(t) &= \int_0^\infty \frac{e^{tx}}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} \exp\left(-\frac{x}{\beta}\right) dx \\
 &= \int_0^\infty \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} \exp\left[-\left(\frac{1-\beta t}{\beta}\right)x\right] dx \\
 &= \left(\frac{1}{1-\beta t}\right)^\alpha \int_0^\infty \frac{1}{\Gamma(\alpha)} \left(\frac{1-\beta t}{\beta}\right)^\alpha x^{\alpha-1} \exp\left[-\left(\frac{1-\beta t}{\beta}\right)x\right] dx \\
 &= \left(\frac{1}{1-\beta t}\right)^\alpha, \quad t < \frac{1}{\beta}.
 \end{aligned}$$

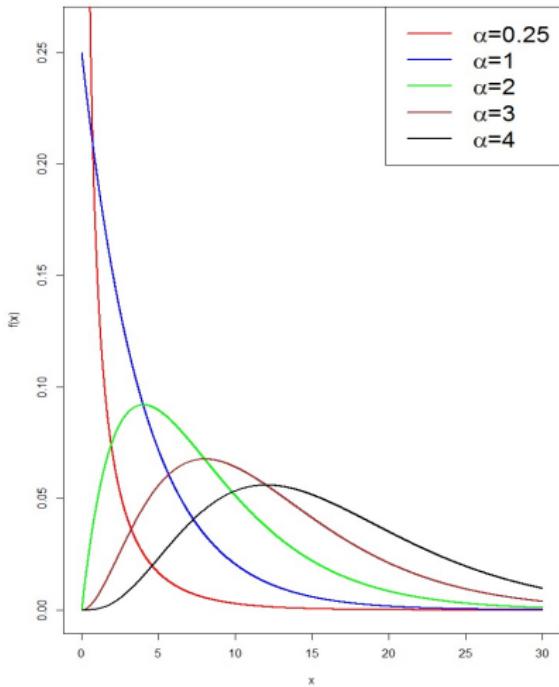
$$\begin{aligned}
 M^{(1)}(0) &= \frac{d}{dt} M(t) \Big|_{t=0} = \frac{d}{dt} (1-\beta t)^{-\alpha} \Big|_{t=0} = \alpha \beta (1-\beta t)^{-\alpha-1} \Big|_{t=0} = \alpha \beta \\
 M^{(2)}(0) &= \frac{d^2}{dt^2} M(t) \Big|_{t=0} = \frac{d^2}{dt^2} (1-\beta t)^{-\alpha} \Big|_{t=0} = \frac{d}{dt} \alpha \beta (1-\beta t)^{-\alpha-1} \Big|_{t=0} \\
 &= \alpha(\alpha+1) \beta^2 (1-\beta t)^{-\alpha-2} \Big|_{t=0} = \alpha(\alpha+1) \beta^2
 \end{aligned}$$

$$\text{Var}(X) = M^{(2)}(0) - \left\{M^{(1)}(0)\right\}^2 = \alpha(\alpha+1)\beta^2 - (\alpha\beta)^2 = \alpha\beta^2.$$

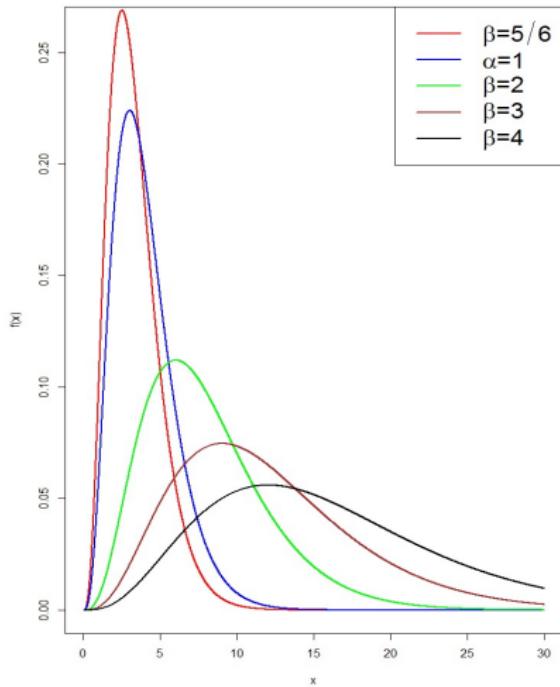


## Gamma pdf

Gamma pdf with  $\beta=4$  and various values of  $\alpha$ .



Gamma pdf with  $\alpha=4$  and various values of  $\beta$



### Example 3.2-4

Suppose the number of customers per hour arriving at a shop follows a Poisson process with mean 30. What is the probability that the shopkeeper will wait more than 5 minutes before both of the first two customers arrive?

$X$ : # of customers per hour ( $= 60 \text{ min}$ )  $\sim \text{Poi}(30)$

$Y$ : waiting time until the first two customers arrive  
are 30 customers / 60 min  $\alpha = 2$

$\Rightarrow$  ave 1 customer / 2 min

$$\Rightarrow \beta = 2$$

$\therefore Y \sim \text{Gamma}(2, 2)$

$$P(Y > 5) = \int_5^\infty \frac{1}{\Gamma(2)} 2^2 y e^{-\frac{y}{2}} dy = \dots = \frac{1}{2} e^{-\frac{5}{2}}$$

## 2.3.2 Chi-Square Distribution, $\chi^2(r)$

- Special case of Gamma distribution:  $\chi^2(r) \equiv \text{Gamma}(r/2, 2)$
- $r$ : the number of degrees of freedom
- pdf

positive  
2nd kind

$$f(x) = \frac{1}{\Gamma(r/2)2^{r/2}} x^{r/2-1} e^{-x/2}, \quad 0 < x < \infty, 0 < r.$$

- The mean and variance

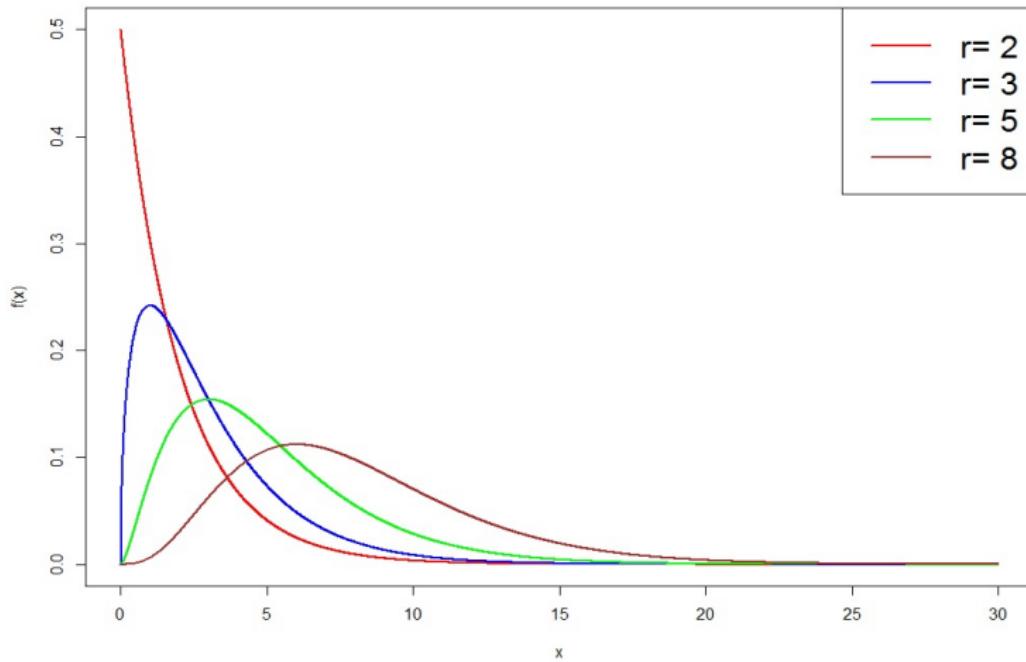
$$E(X) = \alpha\beta = r, \quad \text{Var}(X) = \alpha\beta^2 = 2r.$$

- mgf

$$M(t) = (1 - 2t)^{-r/2}, \quad t < 1/2$$

## Chi-square pdf

Chi-square pdfs with  $r=2,3,5,8$



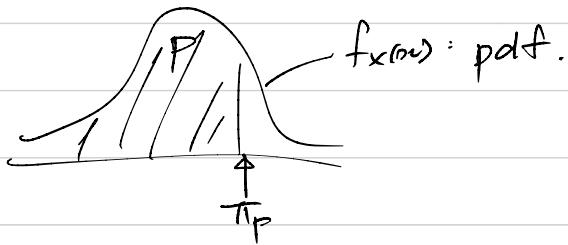
- $\chi^2_\alpha(r)$ : the  $100(1 - \alpha)$ th percentile of the chi-square distribution with  $r$  degrees of freedom such that

$$P[X \leq \chi^2_\alpha(r)] = 1 - \alpha$$

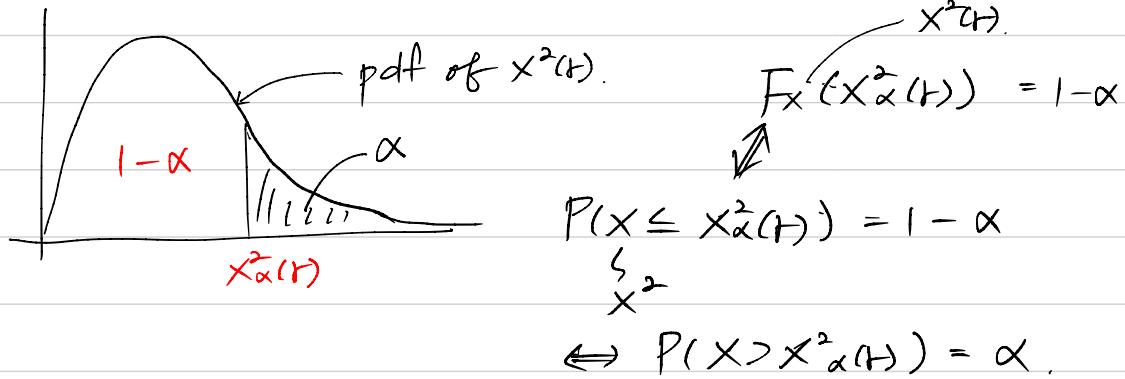
$\pi_p$  : 100  $p^{\text{th}}$  percentile of  $X \sim f_X(x)$

$$F_X(\pi_p) = p$$

$$P(X \leq \pi_p) = p$$



$\chi_{\alpha}(t)$  : 100  $(1-\alpha)^{\text{th}}$  percentile of  $X \sim \chi^2(t)$



$\chi^2(t)$ . one parameter. ( $t > 0$ ).

### Example 3.2-8

Let  $X$  have a chi-square distribution with five degrees of freedom.  
Find the values of  $\chi^2_{0.10}(5)$  and  $\chi^2_{0.90}(5)$ .

### Example 3.2-9

If customers arrive at a shop on the average of 30 per hour in accordance with a Poisson process, what is the probability that the shopkeeper will have to wait longer than 9.390 minutes for the first nine customers to arrive?

$X$ : # of customers per 1 hour ( $= 60 \text{ min}$ ).

$$X \sim \text{Poi}(30)$$

ave 30 customers / 60 min

$$\Rightarrow " 1 " / 2 "$$

$Y$ : waiting time (min) for the first nine  $\cancel{x}$ .  
customers to arrive

$$Y \sim \text{Gamma}(9, 2) \equiv \chi^2(18)$$

$$P(Y > 9.39) = 0.95$$