Lecture 4. Order Flow and the Probability of Informed Trading

organity informed a uninformed traders *121 TOURSON STONE

1. The Distribution of Buys and Sells

Assume that conditional on V, the probability of a current order direction q_t (i.e., a buy or a sell) does not depend on the direction of prior orders: i.e.,

$$\Pr(q_t|V,q_{t-1},q_{t-2},\ldots) = \Pr(q_t|V).$$

Let p denote the conditional probability of incoming a buy order given V. Then one sees that the conditional distributions of b buys given n trades and V are binomial:

$$\Pr(b|n,V) = \binom{n}{b} p^b (1-p)^{n-b}. \tag{1.1}$$

• Extending the previous Glosten and Milgrom sequential trading model, we compute p as:

$$\Pr(Buy|\underline{V}) = \frac{\Pr(Buy,\underline{V})}{\Pr(\underline{V})} \circ = \frac{\Pr(Buy,\underline{V},I) + \Pr(Buy,\underline{V},U)}{\Pr(\underline{V})}$$

$$= \frac{0 + \delta(1-\mu)/2}{\delta}$$

$$= \frac{1-\mu}{2}$$

and

$$\Pr(Buy|\overline{V}) = \underbrace{1+\mu}_{2},$$

both of which are "constant" over successive trades.

Remark 1.1. Advancing (1.1), one sees that the number of buys conditional only on n trades is a mixture of binomials:

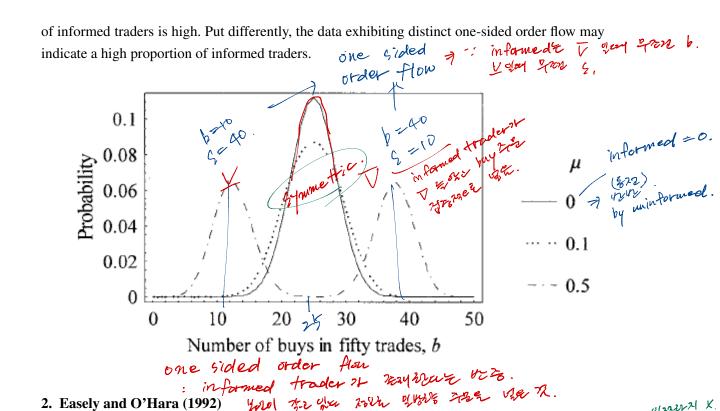
dvancing (1.1), one sees that the number of buys conditional only on
$$n$$
 trades is a smials:

$$\Pr(b|n) = \Pr(V = \underline{V}) \times \Pr(b|n,\underline{V}) + \Pr(V = \overline{V}) \times \Pr(b|n,\overline{V}).$$

$$= \delta \binom{n}{b} \left(\frac{1-\mu}{2}\right)^b \left(1-\frac{1-\mu}{2}\right)^{n-b}$$

$$+(1-\delta) \binom{n}{b} \left(\frac{1+\mu}{2}\right)^b \left(1-\frac{1+\mu}{2}\right)^{n-b}.$$
(1.2)

For n = 50 and $\delta = 0.5$, the following plot shows (1.2) for different μ . With no informed trading (i.e., $\mu = 0$), the distribution is close to normal. With $\mu = 0.5$, the distribution is bimodal; that is to say, "one-sided" order flow (i.e, a preponderance of buys or sells) emerges if a proportion



Easely and O'Hara (1992) extend the basic sequential trade model by introducing information revent uncertainty. Assume that the number of arriving traders in any finite interval follows a Poisson arrival process; specifically, the Poisson arrival intensities for informed and uninformed traders are μ and ε , respectively.

Informed buyers arrive with intensity μ only if \overline{V} occurs, and informed sellers arrive with intensity μ only if \underline{V} occurs. In all states, uninformed buyers and sellers arrive with intensity μ .

Proposition 2.1. Let $N_1(t)$ and $N_2(t)$ denote the number of type I and type II events, respectively, occurring in [0,t] Assume that $N_1(t)$ and $N_2(t)$ follow Poisson arrival processes with λ_1 and λ_2 , respectively. Suppose that $N_1(t)$ and $N_2(t)$ are independent. Define $N(t) = N_1(t) + N_2(t)$. Then, N(t) follows a Poisson arrival process with $\lambda_1 + \lambda_2$.

• If \underline{V} occurs, the total arrival intensity of buyers is ε and that of sellers is $\mu + \varepsilon$. If \overline{V} occurs, the total arrival intensity of buyers is $\mu + \varepsilon$ and that of sellers is ε .

Information intensity $\begin{array}{c}
\alpha \\
1-\delta
\end{array}$ $\begin{array}{c}
V \\
(\varepsilon, \mu+\varepsilon) \\
\text{un informed} \\
\text{un informed} \\
\text{un not ruled} \\
\text{un tormed} \\
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\end{array}$ $\begin{array}{c}
1-\delta \\
1-\delta
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V \\
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1-\delta
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Given that an information event occurs with the probability of α , the unconditional number of buys and sells on any day are jointly distributed as a <u>Poisson mixture</u>:

$$\Pr(b,s) = \Pr(b,s|\text{No information event}) + \Pr(b,s|\text{Information event},\underline{V}) + \Pr(b,s|\text{Information event},\overline{V}) = (1-\alpha)\Pr(b;\varepsilon)\Pr(s;\varepsilon) + \alpha\delta\Pr(b;\varepsilon)\Pr(s;\mu+\varepsilon) + \alpha(1-\delta)\Pr(b;\mu+\varepsilon)\Pr(s;\varepsilon),$$
 (2.1)

where $Pr(n; \lambda)$ denotes the probability of *n* arrivals when λ is the intensity parameter.

• The joint probability Pr(b,s) can be used to construct a <u>sample likelihood function</u>, and one can estimate all parameters of $\alpha, \delta, \mu, \varepsilon$ by <u>maximum likelihood estimation</u>. Specifically, one writes (2.1) as

$$L(b,s|\theta) = (1-\alpha)e^{-\varepsilon t} \frac{(\varepsilon t)^b}{b!} e^{-\varepsilon t} \frac{(\varepsilon t)^s}{s!} + \alpha \delta e^{-\varepsilon t} \frac{(\varepsilon t)^b}{b!} e^{-(\varepsilon + \mu)t} \frac{((\varepsilon + \mu)t)^s}{s!} + \alpha (1-\delta)e^{-(\varepsilon + \mu)t} \frac{((\varepsilon + \mu)t)^b}{b!} e^{-\varepsilon t} \frac{(\varepsilon t)^s}{s!}.$$

Over h independent days, the likelihood of observing the data $M = (b_i, s_i)_{i=1}^h$ is the product of the daily likelihoods

$$L(M|\theta) = \prod_{i=1}^{h} L(b_i, s_i|\theta). \iff \text{in dependent}.$$

Remark 2.2. The *probability of informed trading* (PIN) is defined as the unconditional probability that a randomly chosen trader is informed on a randomly chosen day: i.e.,

PIN =
$$\frac{E[\text{Informed order arrival intensity}]}{E[\text{Total order arrival intensity}]}$$

$$= \frac{\alpha(\delta\mu + (1-\delta)\mu)}{(1-\alpha)(\varepsilon+\varepsilon) + \alpha[\delta(\varepsilon+\mu+\varepsilon) + (1-\delta)(\mu+\varepsilon+\varepsilon)]}$$

$$= \frac{\alpha\mu}{\alpha\mu + 2\varepsilon}$$

The processes PIN increases Recall that high μ corresponds to a strong tendency of one-sided

As μ increases, PIN increases. Recall that high μ corresponds to a strong tendency of one-sided order flow. Therefore, the observation suggests that order flow tends to be one-sided on a trading day with high PIN.

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