

# Basic Fundamentals of Structural Equation Modeling (part 2)

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# Outline (continued)

## III. Additional Topics and lavaan Commands

- I. More lavaan syntax

- J. Composite variables

- K. Robust estimators

- L. Bootstrapping

- M. Multigroup comparisons

# III. Additional Topics and lavaan Commands

## I. More lavaan syntax

1. lavaan has a number of operators and syntax options.  
(see “Basic\_lavaan\_Syntax\_Guide\_Aug1\_2013.pdf”)

formula type	operator	operator stands for
regression	$\sim$	“regressed on”
correlation	$\sim\sim$	“correlated with”
intercept	$\sim 1$	“estimates intercept”
latent variable definition	$=\sim$	“is measured by”
create a composite	$<\sim$	“is caused by”

## 2. You can work with individual parameters by naming them.

Lavaan names parameters as “**y1 ~ x1**”.

We assign names by pre-multiplying a predictor with the name being assigned.

```
model.2a <- 'y1 ~ b1*x1'
```

Note, parameter labels must start with a letter!

### 3. Assigning values to parameters by naming them.

#### c. Fixing Parameter Values to Specific Quantities

There are times when we want to be able to specify that particular parameters have fixed quantitative values. Lavaan allows us to do this using various options. Here is one approach:

```
model.2b <- 'y1 ~ 0*x1 + x2
            y2 ~ x2
            y1 ~~ y2'
```

In this model statement, x1 is pre-multiplied by zero to set its value to zero. We can also accomplish this using a more elaborate and more flexible approach:

```
model.2c <- 'y1 ~ b1*x1 + x2
            y2 ~ x2
            y1 ~~ y2
            b1 == 0'
```

Now we have labeled the parameter “b1” and then assigned it a value of 0 in a separate statement. This second specification will actually result in an explicit test of the constraint.

#### 4. Correlations/Covariances between exogenous variables are not usually estimated, but we can.

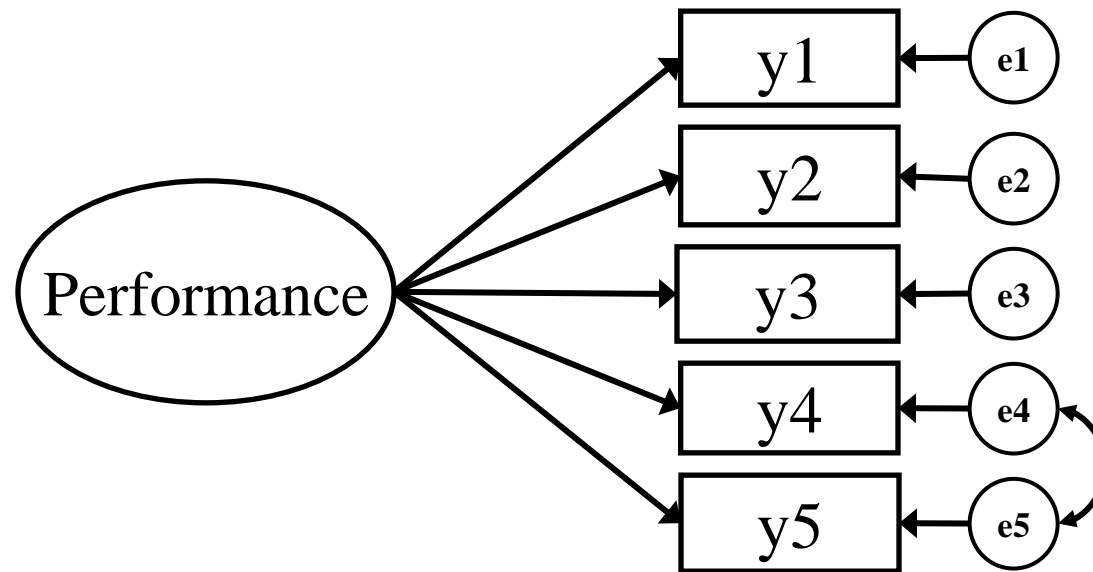
```
#estimating the model  
model.2d.ests <- sem(model.2, data = data.mod1, fixed.x=FALSE)
```

Now we obtain an estimate of the covariance in our lavaan output, as shown in bold below.

---

	Estimate	Std.err	Z-value	P(> z )
Regressions:				
y1 ~				
x1	-0.003	0.004	-0.763	0.446
x2	-0.087	0.019	-4.643	0.000
y2 ~				
x2	-3.363	0.896	-3.752	0.000
Covariances:				
y1 ~~				
y2	0.945	0.432	2.189	0.029
<b>x1 ~~</b>				
<b>x2</b>	<b>-2.651</b>	<b>1.352</b>	<b>-1.961</b>	<b>0.050</b>

5. Lavaan creates latent variables by declaring them in the absence of any known values.



```
lvmod.2 <- ' # Latent variable definition
             Perform =~ stems + infls + clonedia
             + leafht + leafwdth

             # Error Covariances
             leafht ~~ leafwdth'
```



# J. Composite Variables

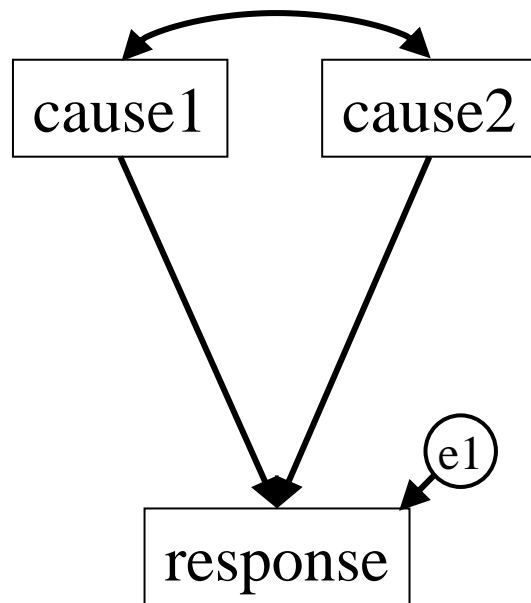
Environ Ecol Stat  
DOI 10.1007/s10651-007-0047-7

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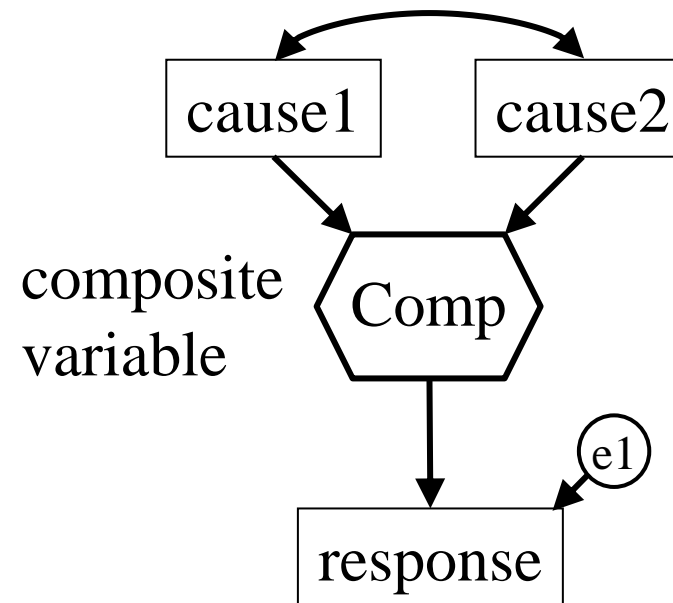
**Representing general theoretical concepts in structural equation models: the role of composite variables**

**James B. Grace • Kenneth A. Bollen**

1. In the simple case, we can use composites to represent collective effects of a set of variables.



First step in compositing process.

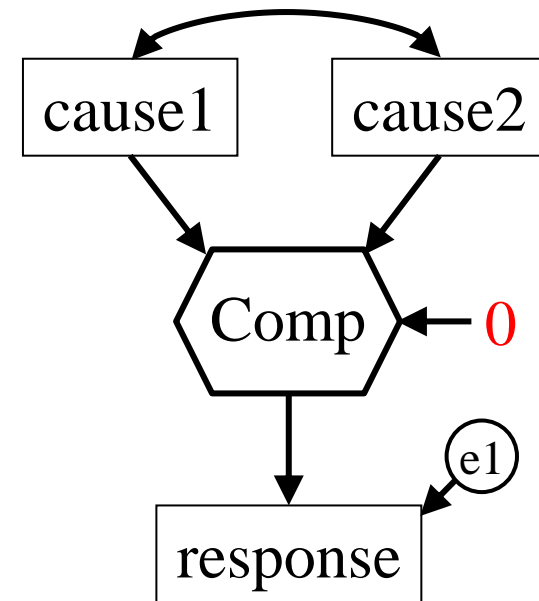


Second step in compositing process.

## 2. Composites are abstractions, and we must declare them.

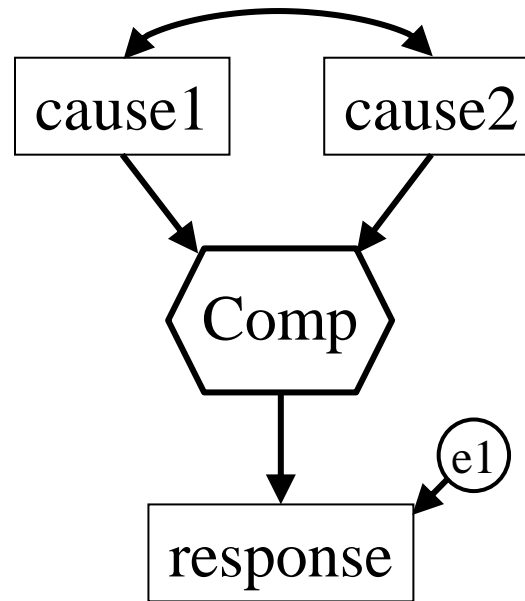
To specify composites, we need to conceptualize a composite as a latent variable with no error (error set = 0).

Composites were not part of the original LISREL conceptualization.



The composite captures the combined effects of its causes on its responses (akin to a multiple regression predictor).

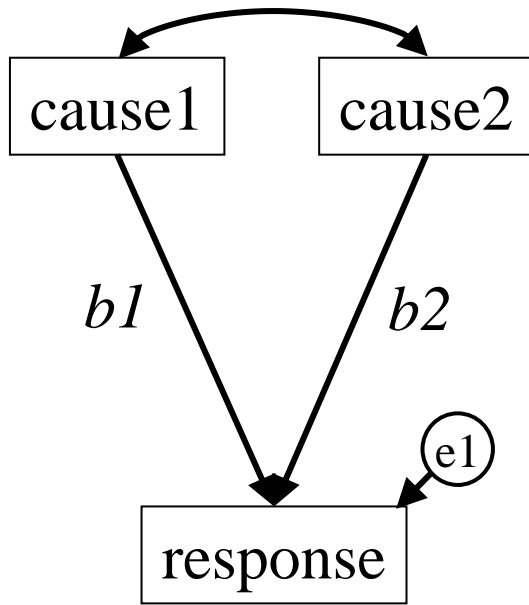
3. We can create composites with lavaan syntax.



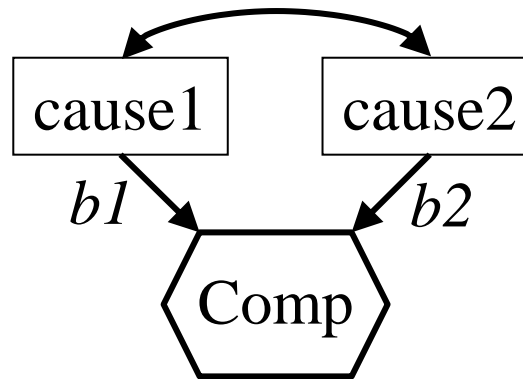
```
# we declare composite as effects of two variables  
compmod.z <- 'Comp <~ 1*cause1 + cause2  
  
response ~ Comp'
```

We use '1\*cause2'to set scale of composite.

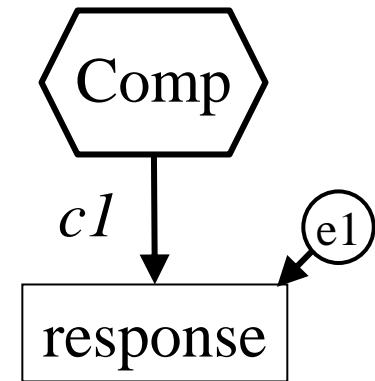
4. We can also compute composite scores by hand.



First, run model without composite and obtain coefficients.



Use coefficients to calculate composite scores.



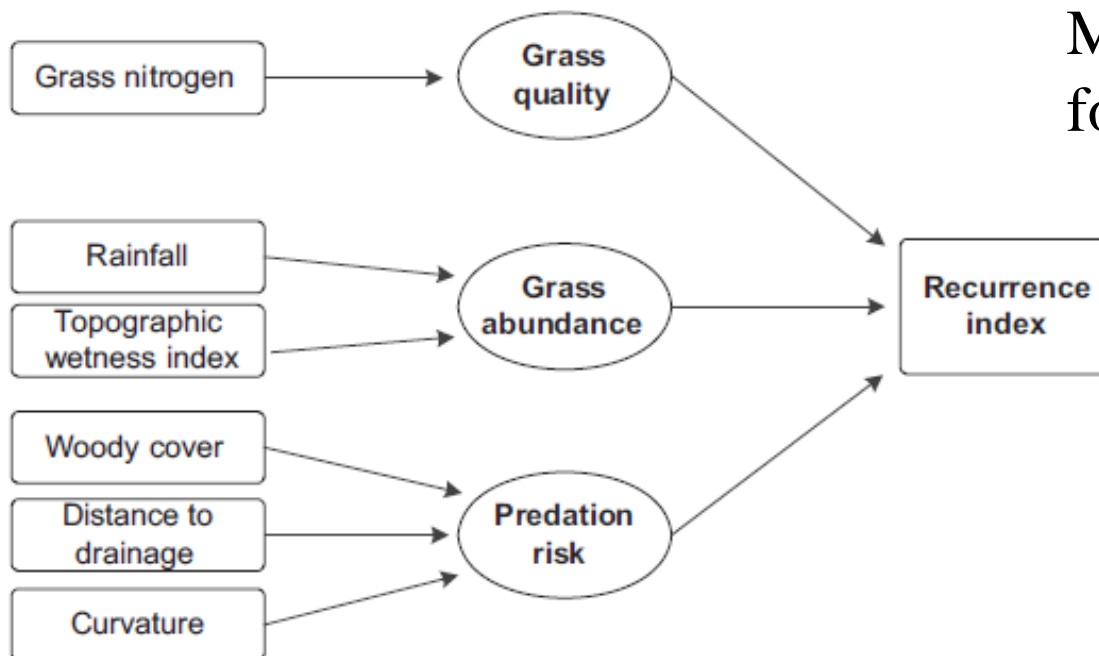
Use composite scores as a variable.

The coefficient  $c1$  represents collective effects and is only meaningful in standardized units (normally, its raw value is 1.0).

## 5. We can use composites to represent general concepts.

### Body size and the division of niche space: food and predation differentially shape the distribution of Serengeti grazers

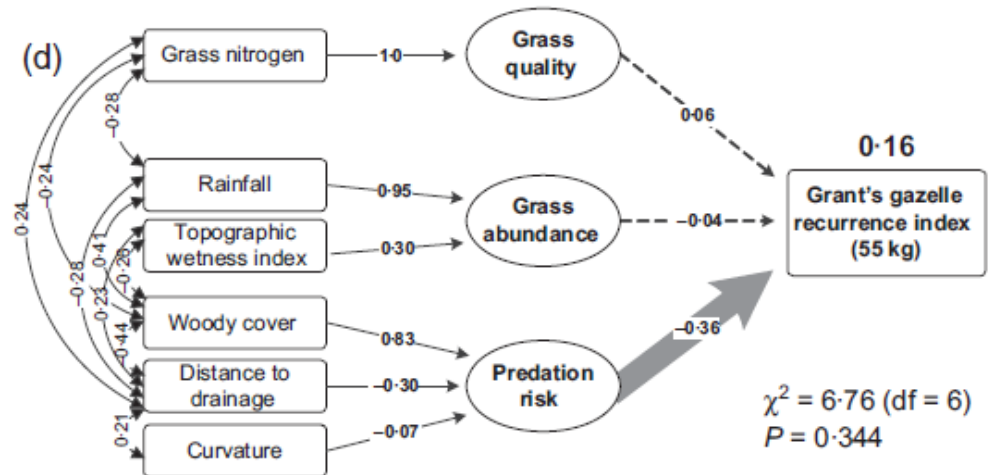
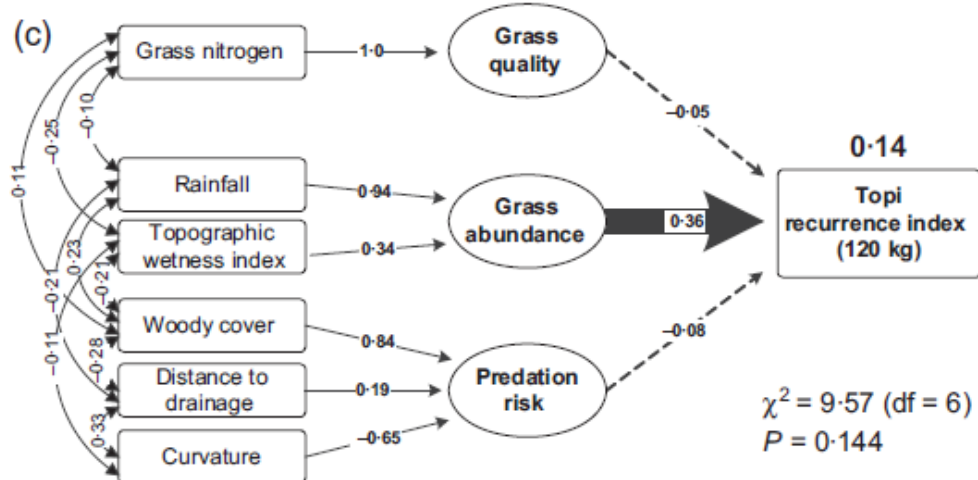
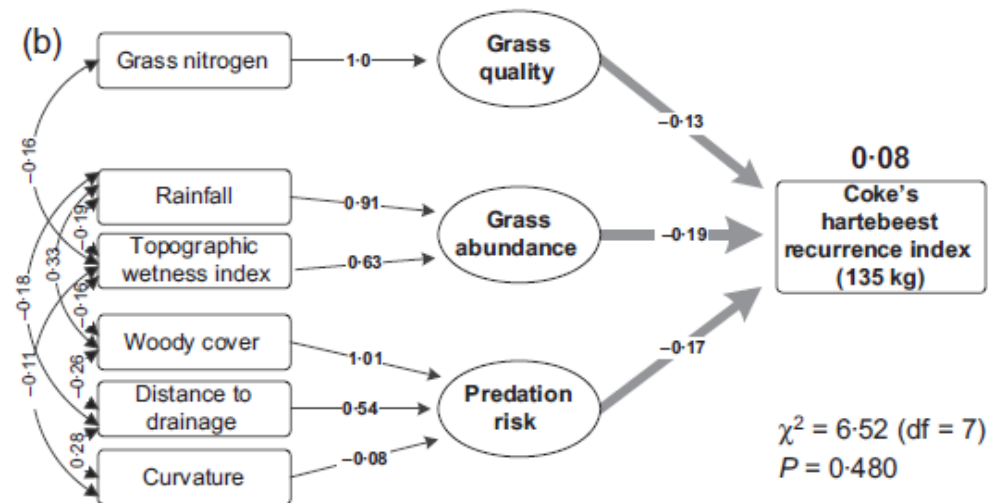
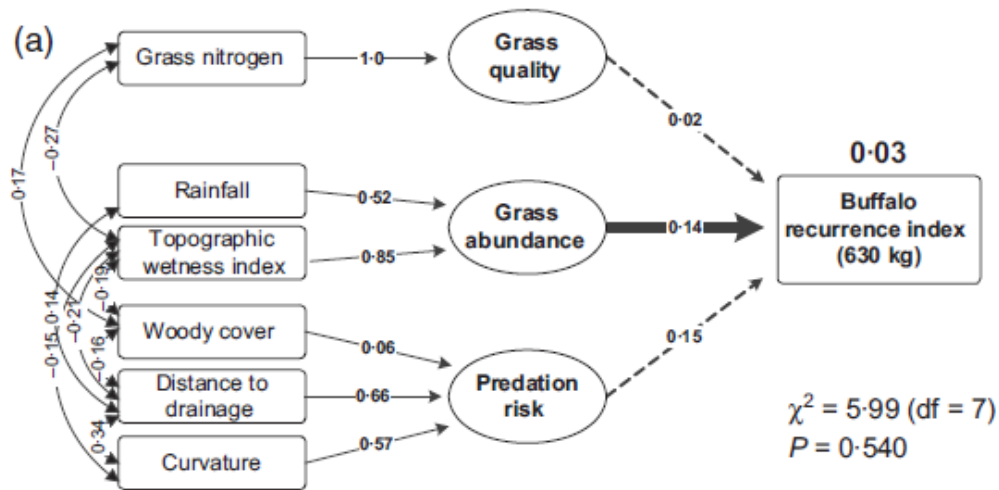
J. Grant C. Hopcraft<sup>1,2\*</sup>, T. Michael Anderson<sup>3</sup>, Saleta Pérez-Vila<sup>4</sup>, Emilian Mayemba<sup>5</sup> and Han Olff<sup>1</sup>



Model of predation risk versus food rewards for grazers.

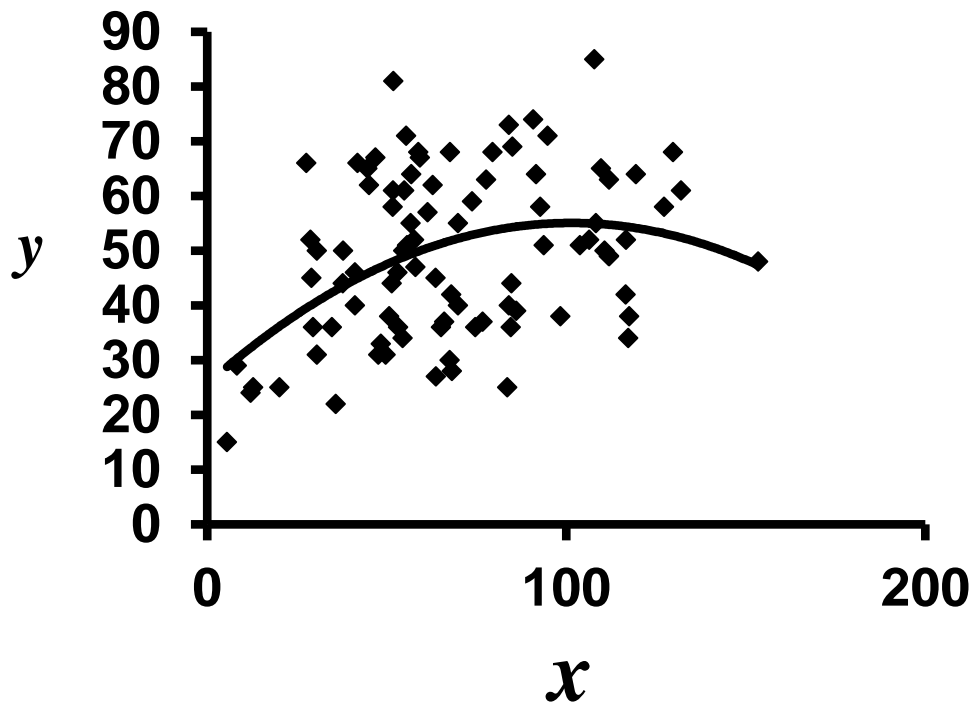
Fig. 2. The *a priori* structural equation model used to assess the

6. And we can use them to compare apples and oranges (with care in interpretation).



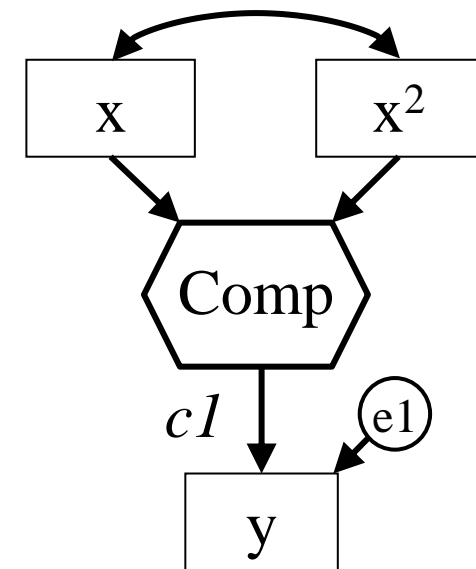
7. We can also use composites as devices.

Consider this nonlinear relationship.



We can model this using polynomial regression:

$$y = \gamma x + \gamma x^2 + \varepsilon$$



Centering a variable before squaring it reduces the autocorrelation between polynomial terms.

```
# center x before squaring
x <- x - mean(x)
x2 <- x^2
```



# Discussion

# K. Robust Estimators

# 1. Lavaan permits use of “robust” estimation.

Lavaan has two main options for robust estimation:

MLM – produces chi-squares and standard errors robust to non-normality. AKA the Satorra-Bentler correction.

MLR – similar to MLM, but uses the Yuan-Bentler method so that missing data can be accommodated.

see discussion in:

Yuan & Bentler. 2000. In Sobel & Becker (eds.) Sociological Methodology (pp 165-200)

## 2. Robust estimation invoked with ‘estimator =’ command.

```
# create model
mod <- 'y2 ~ y1
        y1 ~ x1'

# estimate model
mod.fit <- sem(mod, data=dat, fixed.x=F,
               estimator="mlm")

# get results
summary(mod.fit)
```

“fixed.x=F” is required when using “mlm” option.

### 3. Results

Estimator	ML	Robust
Chi-square	4.213	4.082
Degrees of freedom	1	1
P-value	0.040	0.043
Scaling correction factor for the Yuan-Bentler correction		1.032

Standard Errors	Robust.mlm			
	Estimate	Std.err	Z-value	P(> z )
Regressions:				
y2 ~				
y1	-0.154	0.024	-6.386	0.000
y1 ~				
x1	1.185	0.290	4.091	0.000

# L. Bootstrapping

# 1. Lavaan has resampling methods for non-normal data.

```
# fit model and request bootstrapped results  
  
mod.fit <- sem(mod, dat=dat,  
               test="boot", se="boot", bootstrap=200)
```

## 2. Results

Estimator	ML
Chi-square	4.213
Degrees of freedom	1
P-value	0.040
P-value (Bollen-Stine)	0.053

Estimate	Std.err	Z-value	P(> z )
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Regressions with Robust.mlm standard errors:

y2.ln ~

y1.ln	-0.154	0.024	-6.386	0.000
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y1.ln ~

x1.ln	1.185	0.290	4.091	0.000
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Regressions with bootstrapped standard errors:

y2.ln ~

y1.ln	-0.154	0.027	-5.750	0.000
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y1.ln ~

x1.ln	1.185	0.423	2.800	0.005
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# Discussion

# M. Multigroup Modeling

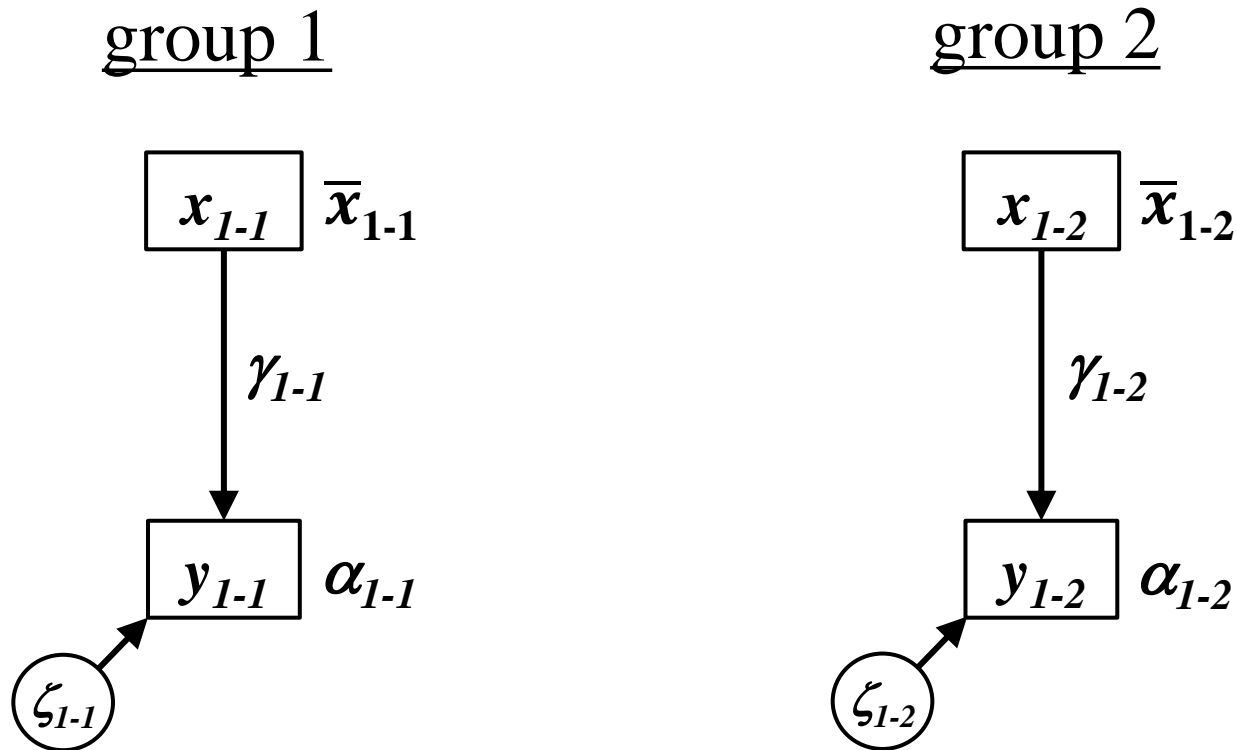
1. Multigroup modeling involves situations where there are discrete groups in the data that you want to compare.

- males versus females in a population
- treated versus control plots
- areas with different disturbance histories

In multigroup modeling, we develop a common model for different groups and then ask what parameters are the same or different between groups.

We can use the classical likelihood-based measures (e.g., model chi-square) to test for constraints across groups.

2. There are several parameters that can be compared.



A number of hypotheses we can test:

- (1) equal slopes:  $\gamma_{1-1} = \gamma_{1-2}$
- (2) equal intercepts:  $\alpha_{1-1} = \alpha_{1-2}$
- (3) equal means:  $\bar{x}_{1-1} = \bar{x}_{1-2}$
- (4) equal errors:  $\zeta_{1-1} = \zeta_{1-2}$

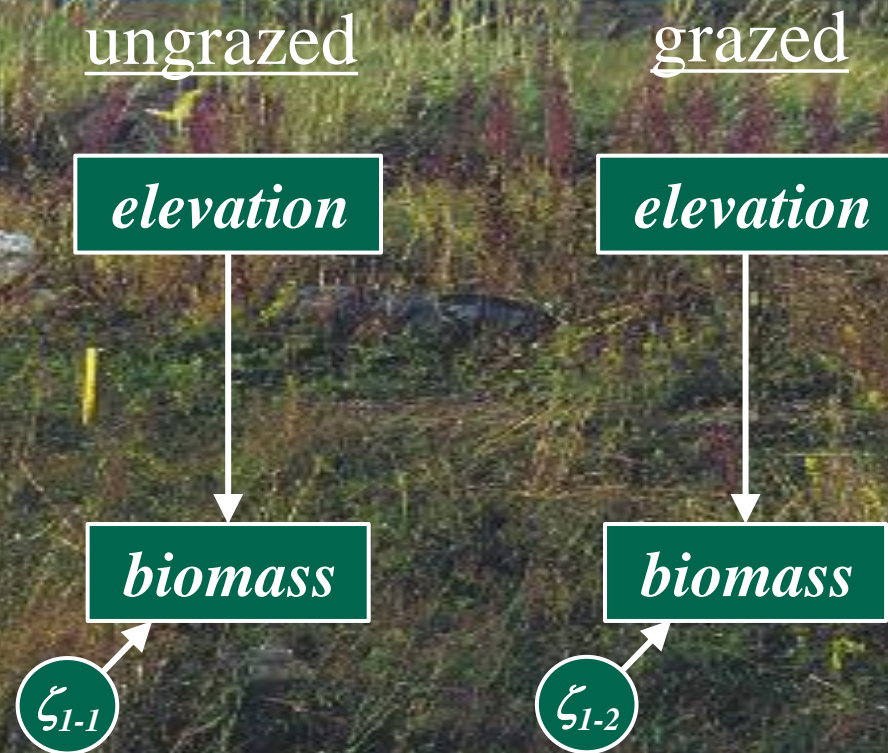
increasingly strong  
constraints

↓

# Effects of Grazing on Finnish Coastal Meadows\*

How does biomass respond to the interaction between grazing & elevation?

Data from 1-m<sup>2</sup> plots arrayed along an elevation gradient in each of several paired grazed and ungrazed meadows in SW Finland.



Grace, J.B. and Jutila, H. (1999) The relationship between species density and community biomass in grazed and ungrazed coastal meadows. *Oikos*, 85:398-408.

1. Lavaan uses a 'group=' command to invoke a multigroup analysis.

```
#lavaan code for basic model
mod1 <- 'biomass ~ elev'

#fit the model, specifying groups
mod1.fit <- sem(mod1, data=sem.dat, group="grazed")

#request output
summary(mod1.fit)
```

## 2. Default allows all parameters to differ between groups.

```
> summary(mod1.fit)
lavaan (0.5-12) converged normally after 24 iterations

Number of observations per group
1                                165
0                                189

Estimator                        ML
Minimum Function Test Statistic  0.000
Degrees of freedom                0
P-value (Chi-square)             0.000

Chi-square for each group:

1                                0.000
0                                0.000
```

There are no equality constraints and therefore no chi-square tests.

### 3. Default allows all parameters to differ between groups.

Group 1 [1]: (notice Group=1 = grazed)

	Estimate	Std.err	Z-value	P(> z )
Regressions:				
biomass ~				
elev	-0.474	0.205	-2.311	0.021
Intercepts:				
biomass	5.263	0.117	45.072	0.000
Variances:				
biomass	0.534	0.059		

Group 2 [0]:

	Estimate	Std.err	Z-value	P(> z )
Regressions:				
biomass ~				
elev	-0.798	0.145	-5.523	0.000
Intercepts:				
biomass	5.926	0.065	91.069	0.000
Variances:				
biomass	0.296	0.030	0.000	



4. We can test equality constraints by labeling parameters.

```
#lavaan code naming the path coefficient "b1"  
mod2 <- 'biomass ~ c("b1","b1")*elev'  
  
#fit the model, specifying groups  
mod2.fit <- sem(mod2, data=sem.dat, group="grazed")
```

When you label a parameter across groups, you have to pass to lavaan a vector of labels, one for each group. Here, `c("b1","b1")` is a vector of labels.

5. With constraints imposed, we can test if models sig. different.

```
> summary(mod2.fit)
lavaan (0.5-12) converged normally after 19 iterations

Estimator ML
Minimum Function Test Statistic 1.668
Degrees of freedom 1
P-value (Chi-square) 0.197

Chi-square for each group:

1 1.116
0 0.552
```

The overall model chi-square of 1.668 with 1 df would traditionally be interpreted as a non-significant difference between models.

6. With constraints imposed, we can test if parameters different.

Group 1 [1]:

		Estimate	Std.err	Z-value	P(> z )
Regressions:					
biomass ~					
elev	(b1)	-0.691	0.118	-5.836	0.000
Intercepts:					
biomass		5.371	0.082	65.493	0.000
Variances:					
biomass		0.538	0.059		

**We get one best estimate for both groups.**

Group 2 [0]:

		Estimate	Std.err	Z-value	P(> z )
Regressions:					
biomass ~					
elev	(b1)	-0.691	0.118	-5.836	0.000
Intercepts:					
biomass		5.888	0.058	101.552	0.000
Variances:					
biomass		0.297	0.031		

7. We can also use an 'equal' or a 'group.equal' command.

```
mod2a <- 'biom.log ~ equal("b1")*elev.m'  
mod2a.fit <- sem(mod2a, data=sem.dat, group="grazed")
```

Produces exactly the same results as the previous command.

```
mod3 <- 'biom.log ~ elev.m'  
mod3.fit <- sem(mod3, data=sem.dat, group="grazed",  
group.equal="regressions")
```

8. There are a number of helpful options for the 'group.equal' command.

```
group.equal=c(  
  "intercepts",  
  "means",  
  "regressions",  
  "residuals",  
  "residual.covariances")
```

This gives you a taste of the possibilities for testing equality constraints across groups.

For more, consult the tutorial

<http://lavaan.ugent.be/tutorial/tutorial.pdf>

# Discussion