

06051540-MATH70076-assessment-1

MSc in Statistics 2025/26, Imperial College London

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Question 1

For $\xi \neq 0$:

We know from the cumulative distribution function that:

$$F(x; \sigma, \xi, u) = 1 - \left(1 + \frac{\xi(x-u)}{\sigma}\right)_+^{-1/\xi} \quad (1)$$

Rewriting to make x the subject:

$$\left(1 + \frac{\xi(x-u)}{\sigma}\right)_+^{-1/\xi} = (1 - F) \quad (2)$$

$$1 = (1 - F) \left(1 + \frac{\xi(x-u)}{\sigma}\right)_+^{1/\xi} \quad (3)$$

$$(1 - F)^{-1} = \left(1 + \frac{\xi(x-u)}{\sigma}\right)_+^{1/\xi} \quad (4)$$

$$(5)$$

For $\xi > 0$, we see that $\left(1 + \frac{\xi(x-u)}{\sigma}\right)_+^{1/\xi} > 0$, so we get:

$$(1 - F)^{-\xi} = 1 + \frac{\xi(x-u)}{\sigma} \quad (6)$$

$$(1 - F)^{-\xi} - 1 = \frac{\xi(x-u)}{\sigma} \quad (7)$$

$$\left((1 - F)^{-\xi} - 1\right) \times \frac{\sigma}{\xi} = x - u \quad (8)$$

$$u + \left((1 - F)^{-\xi} - 1\right) \times \frac{\sigma}{\xi} = x \quad (9)$$

$$(10)$$

So $F_X^{-1} = u + ((1-x)^{-\xi} - 1) \times \frac{\sigma}{\xi}$ But given our inputs of $F_X^{-1}(x)$ vary between 0 and 1, we can write that:

$$F_X^{-1}(x) = u + ((x)^{-\xi} - 1) \times \frac{\sigma}{\xi} \quad (11)$$

For $\xi < 0$, we know that we have quickly decaying tails with finite upper endpoint. With $x > u$, this finite endpoint is met when

$$1 + \frac{\xi(x-u)}{\sigma} = 0 \quad (12)$$

$$\frac{\xi(x-u)}{\sigma} = -1 \quad (13)$$

$$x - u = -\frac{\sigma}{\xi} \quad (14)$$

$$x = u - \frac{\sigma}{\xi} \quad (15)$$

$$(16)$$

So for $x > u - \frac{\sigma}{\xi}$, produced by the inverse function above we discard the values of x .

For $\xi = 0$:

$$F = 1 - \exp\left(-\frac{x-u}{\sigma}\right) \quad (17)$$

$$\exp\left(-\frac{x-u}{\sigma}\right) = 1 - F \quad (18)$$

$$-\frac{x-u}{\sigma} = \ln(1-F) \quad (19)$$

$$-x + u = \sigma \ln(1-F) \quad (20)$$

$$u - \sigma \ln(1-F) = x \quad (21)$$

$$(22)$$

So $F_X^{-1} = u - \sigma \ln(1-x)$

But given our inputs of $F_X^{-1}(x)$ vary between 0 and 1, we can write that:

$$F_X^{-1}(x) = u - \sigma \ln(x) \quad (23)$$

Question 2a

Defining the quantile function, and using the basis of the cdf from question 1, we get that:

```

qgpd <- function (p, sigma=1, xi=0, u=0){
  if (p<0||p>1){
    return(warning("NaNs produced - p must be between 0 and 1"))
  }
  else if (sigma<=0){
    return(warning("NaNs produced - sigma must be greater than 0"))
  }
  else if (xi != 0){
    return(u + ((1-p)^(-xi)-1) * sigma / xi)
  } else {
    return(u - sigma * log(1-p))
  }
}

```

By default, the expected inputs for the function are $\sigma=1$, $\xi=0$, $u=0$. The code also prevents inputs where p values are less than 0, where p values exceed 1, or where σ is less than or equal to 0.

The expected output is a real number greater than u that is unbounded if ξ is greater than or equal to 0. The expected output is less than $u - \sigma/\xi$ if ξ is less than 0.

Regarding the behaviours of the quantile function: The larger the value of ξ , the slower the tail decays. In this for $\xi \geq 0$ the functions output approaches infinity as $p \rightarrow 1$. For $\xi < 0$, the functions output has a maximum at $u - \sigma/\xi$ (when $p \rightarrow 1$). For larger values of σ , the slower the tail decays.

Question 2b

```
qgpd(0.5,2,-0.4,1.5)
```

```
[1] 2.710709
```

```
qgpd(0.75,2,-0.4,1.5)
```

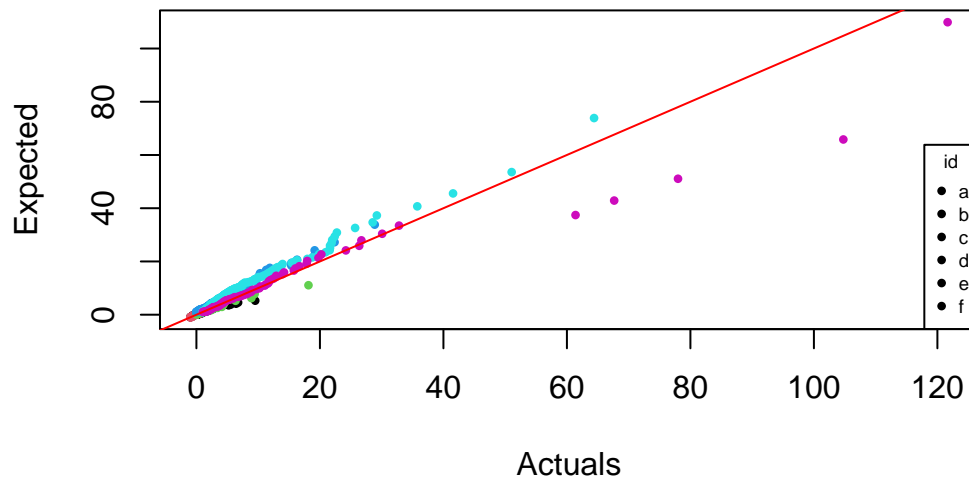
```
[1] 3.628254
```

```
qgpd(0.99,2,-0.4,1.5)
```

```
[1] 5.707553
```

Question 3

Graphing our actual vs expected split by id, we see that:



From the figure above, we see some interesting results – although most actuals seem in line with the expecteds (being close to the figures red line), this is not always the case. Ultimately, however, the different ids need to be separated so as to provide a more granular analysis of the distributions:

