

# Detection of causality in nonlinear time series using extremal value theory

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Date of the Master thesis defense: 21.6.2021

# Outline

1. Preliminaries
2. Main idea
3. Other approaches
4. Simulations and application

# Granger causality

Let  $X = (X_t, t \in \mathbb{Z})$  and  $Y = (Y_t, t \in \mathbb{Z})$  be time series.  
 $X$  causes  $Y$  if there exists measurable set  $A$  such that

$$P(Y_{t+1} \in A \mid \mathbb{U}_t) \neq P(Y_{t+1} \in A \mid \mathbb{U}_t \setminus \sigma(X_s, s \leq t)),$$

where  $\mathbb{U}_t$  is all the information in the universe until time  $t$ .

# VAR model

Bivariate time series  $(X, Y)$  follows VAR model if

$$\begin{aligned}X_t &= \alpha_1 X_{t-1} + \cdots + \alpha_q X_{t-q} + \gamma_1 Y_{t-1} + \cdots + \gamma_q Y_{t-q} + \varepsilon_t^X, \\Y_t &= \beta_1 Y_{t-1} + \cdots + \beta_q Y_{t-q} + \delta_1 X_{t-1} + \cdots + \delta_q X_{t-q} + \varepsilon_t^Y.\end{aligned}$$

It can be rewritten as

$$\begin{aligned}X_t &= \sum_{i=0}^{\infty} a_i \varepsilon_{t-i}^X + \sum_{i=0}^{\infty} c_i \varepsilon_{t-i}^Y, \\Y_t &= \sum_{i=0}^{\infty} b_i \varepsilon_{t-i}^Y + \sum_{i=0}^{\infty} d_i \varepsilon_{t-i}^X.\end{aligned}$$

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# Nonlinear model (NAR)

Let  $Z = (Z_t, t \in \mathbb{Z})$  be  $d$ -dimensional time series.

**General form**  $Z_t = f(Z_{t-1}, \dots, Z_{t-q}, \varepsilon_t)$ .

**Additive form**  $Z_t = f_1(Z_{t-1}) + \dots + f_q(Z_{t-q}) + \varepsilon_t$ .

For bivariate  $Z = (X, Y)$  and  $q = 1$  is

$$\begin{aligned} X_t &= f_1(X_{t-1}) + f_2(Y_{t-1}) + \varepsilon_t^X, \\ Y_t &= g_1(Y_{t-1}) + g_2(X_{t-1}) + \varepsilon_t^Y. \end{aligned}$$

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# Regular variation

We will always assume  $(\varepsilon_i, i \in \mathbb{Z}) \stackrel{\text{iid}}{\sim} RV(\theta)$ , which means

$$P(\varepsilon_t > x) \sim x^{-\theta} L(x)$$

for some slowly varying  $L$ .

Examples include Cauchy or Pareto distributions.



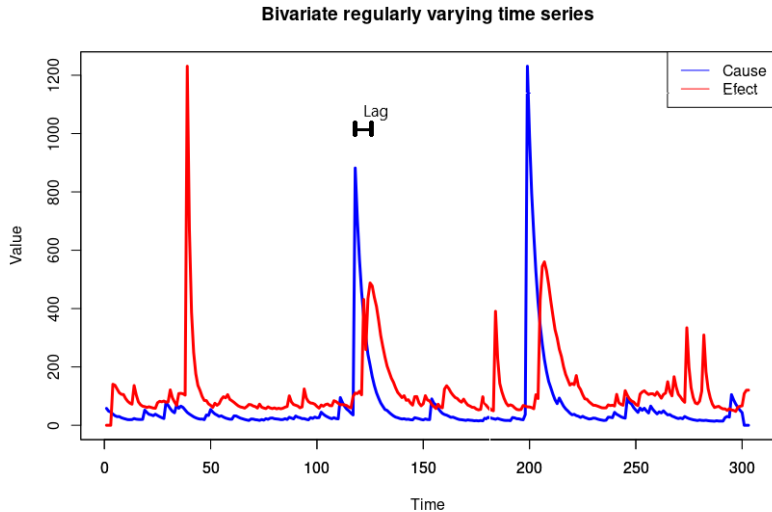
# Main idea

Consider NAR model

$$\begin{aligned}X_t &= \frac{1}{2}X_{t-1} + \varepsilon_t^X, \\Y_t &= \frac{1}{2}Y_{t-1} + \sqrt{X_{t-5}} + \varepsilon_t^Y,\end{aligned}$$

with  $\varepsilon_t^X, \varepsilon_t^Y \stackrel{\text{iid}}{\sim} \text{Pareto}$ .

# Sample realization



# Causal tail coefficient for stationary time series

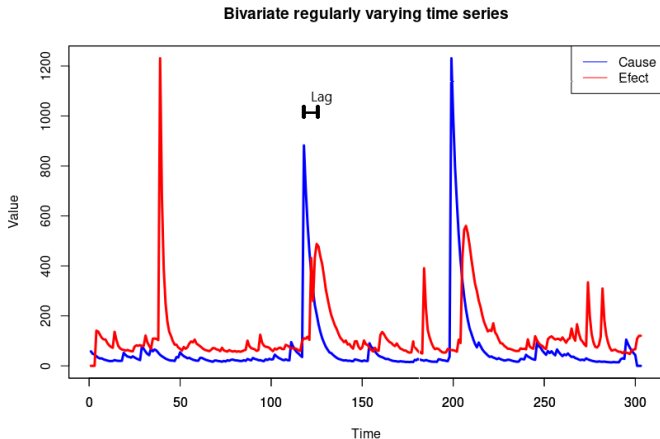
Therefore, we propose the following coefficient,

$$\Gamma_{X,Y}^{time}(q) := \lim_{u \rightarrow 1^-} \mathbb{E}[\max\{F_Y(Y_0), \dots, F_Y(Y_q)\} \mid F_X(X_0) > u],$$

where, in our case, is  $q = 5$ . Here, we denote by  $F_X, F_Y$  the stationary distributions of  $X_t$  and  $Y_t$ , respectively.

# Causal tail coefficient for stationary time series

$$\Gamma_{X,Y}^{time}(q) = \lim_{u \rightarrow 1^-} \mathbb{E}[\max\{F_Y(Y_0), \dots, F_Y(Y_q)\} \mid F_X(X_0) > u]$$



# Main theorem

$$\Gamma_{X,Y}^{time}(q) = \lim_{u \rightarrow 1^-} \mathbb{E}[\max\{F_Y(Y_0), \dots, F_Y(Y_q)\} \mid F_X(X_0) > u]$$

## Theorem

$(X, Y)^\top$  follow either heavy-tailed VAR or heavy-tailed NAR model of order  $q$ . Let “extremal causal condition” hold. Then,

- $X$  causes  $Y \implies \Gamma_{X,Y}^{time}(q) = 1$
- $X$  does not cause  $Y \implies \Gamma_{X,Y}^{time}(q) < 1$

# Idea why theorem holds

We have a heavy-tailed VAR model where  $X$  does not cause  $Y$ :

$$X_t = \sum_{i=0}^{\infty} a_i \varepsilon_{t-i}^X + \sum_{i=0}^{\infty} c_i \varepsilon_{t-i}^Y$$

$$Y_t = \sum_{i=0}^{\infty} b_i \varepsilon_{t-i}^Y.$$

Then,

$$\Gamma_{X,Y}^{time}(q) = \lim_{u \rightarrow 1^-} \mathbb{E}[\max\{F_Y(Y_0), \dots, F_Y(Y_q)\} \mid F_X(X_0) > u] < 1$$

if and only if

$$\lim_{u \rightarrow \infty} \mathbb{E}[f(\varepsilon_q^Y, \varepsilon_{q-1}^Y, \varepsilon_{q-2}^Y, \dots) \mid \sum_{i=0}^{\infty} a_i \varepsilon_{-i}^X + \sum_{i=0}^{\infty} c_i \varepsilon_{-i}^Y > u] < 1.$$

We apply a principle of the single big jump.

# Application of Theorem

## Theorem

$(X, Y)^\top$  follow either heavy-tailed VAR or heavy-tailed NAR model of order  $q$ . Let “extremal causal condition” hold. Then,

- $X$  causes  $Y \implies \Gamma_{X,Y}^{time}(q) = 1$
- $X$  does not cause  $Y \implies \Gamma_{X,Y}^{time}(q) < 1$

We will apply this in the following.

# Lag

## Definition (Minimal lag)

Let  $(X, Y)^\top$  follow stable  $VAR(q)$  model, specified by

$$\begin{aligned}X_t &= \alpha_1 X_{t-1} + \cdots + \alpha_q X_{t-q} + \gamma_1 Y_{t-1} + \cdots + \gamma_q Y_{t-q} + \varepsilon_t^X, \\Y_t &= \beta_1 Y_{t-1} + \cdots + \beta_q Y_{t-q} + \delta_1 X_{t-1} + \cdots + \delta_q X_{t-q} + \varepsilon_t^Y.\end{aligned}$$

We call  $p \in \mathbb{N}$  the *minimal lag*, if  $\gamma_1 = \cdots = \gamma_{p-1} = \delta_1 = \cdots = \delta_{p-1} = 0$  and either  $\delta_p \neq 0$  or  $\gamma_p \neq 0$ . If such  $p$  does not exist, we define the minimal lag as  $+\infty$ .

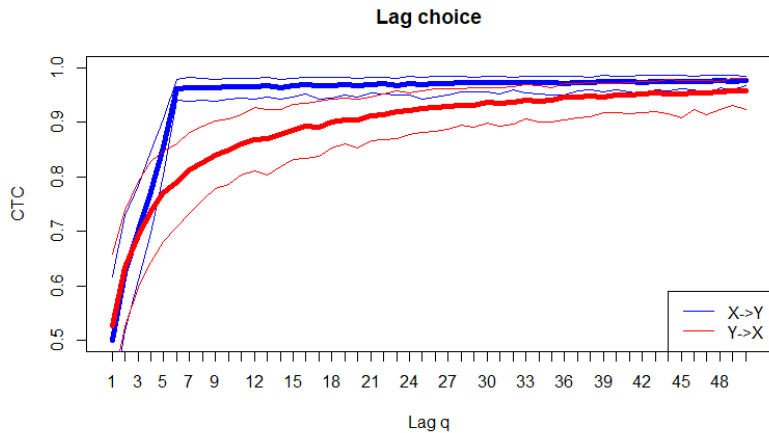
## Lemma (Specification of our Theorem)

Let the assumptions from the previous theorem hold. Let  $p$  be the minimal lag. Then,  $\Gamma_{X,Y}^{time}(r) < 1$  for all  $r < p$ , and  $\Gamma_{X,Y}^{time}(r) = 1$  for all  $r \geq p$ .

How to estimate the minimal lag? Just take minimal  $p$  for which is  $\Gamma_{X,Y}^{time}(p) = 1$ .



# Minimal lag



$$\Gamma_{X,Y}^{time}(q) = \lim_{u \rightarrow 1^-} \mathbb{E}[\max\{F_Y(Y_0), \dots, F_Y(Y_q)\} \mid F_X(X_0) > u]$$

## Definition

$$\hat{\Gamma}_{X,Y}^{time}(q) := \frac{1}{k} \sum_{i: X_i \geq \tau_k^X} \max\{\hat{F}_Y(Y_i), \dots, \hat{F}_Y(Y_{i+q})\},$$

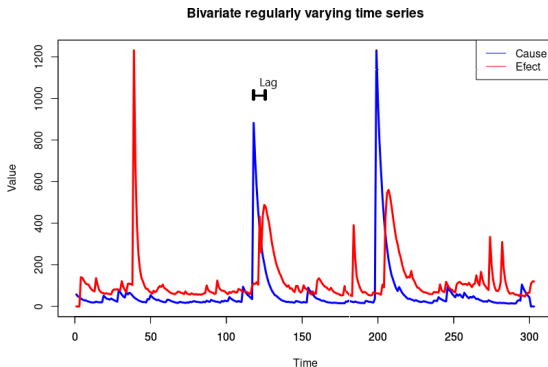
where  $\tau_k^X = X_{(n-k+1)}$  is the  $k$ -th largest value of  $X_i$ , and  $\hat{F}_Y(Y_i) = \frac{1}{n} \sum_{j=1}^n 1\{Y_j \leq Y_i\}$ .

Here,

$$k_n \rightarrow \infty, \frac{k_n}{n} \rightarrow 0, \text{ as } n \rightarrow \infty.$$

# Bootstrap

Distribution of  $\hat{\Gamma}_{X,Y}^{time}(q)$  is hard to compute. We used block bootstrap method for computing confidence intervals <sup>1</sup>.

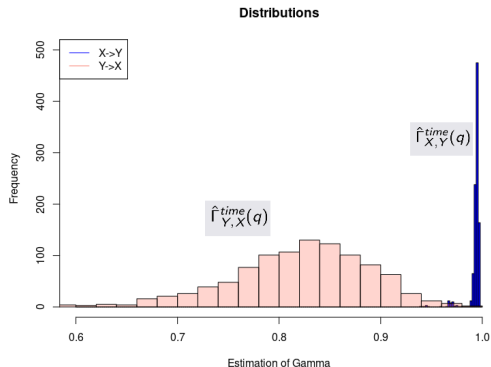


<sup>1</sup>More precisely so-called Reverse Bootstrap Percentile Interval.

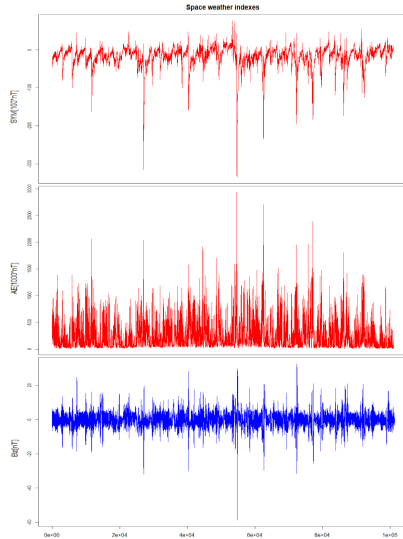
# Simulations

$$X_t = 0.5X_{t-1} + \varepsilon_t^X,$$

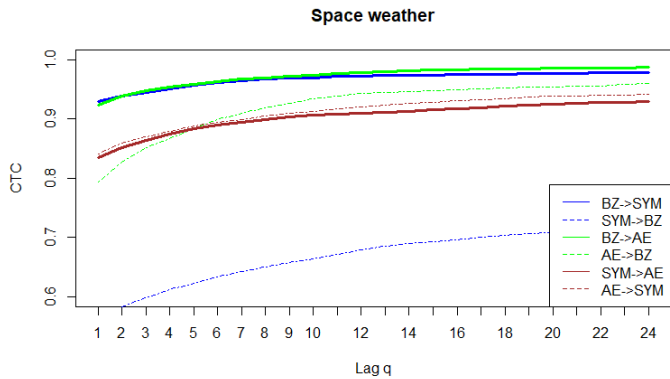
$$Y_t = 0.5Y_{t-1} + 0.5X_{t-2} + \varepsilon_t^Y.$$



# Application



# Application



Thank you!