

Heterogeneity in Firms: A Proxy Variable Approach for Quantile Production Functions

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Abstract

We propose a new approach to estimate production functions in which output elasticities are heterogeneous across the firm size distribution. This paper extends the proxy variable approach for estimating production functions to the conditional quantiles of firm production. Production function parameters are estimated in a simple two-stage approach which relies a location-shift assumption on unobserved productivity. We show that this method allows us to capture heterogeneity in output elasticities along the firm size distribution that may not be found in conditional mean estimates of Levinsohn and Petrin (2003) and Ackerberg *et al.* (2015). We provide small-sample evidence in a Monte Carlo study to show that this approach is robust compared to other production function estimators. The method is applied to plant-level manufacturing data from Chile and Colombia.

Keywords: Production functions, Heterogeneous elasticity, Quantile regression

JEL Classification: C14, C36, D24

1 Introduction

Production function estimation is an ongoing and historical empirical research topic that links firm's input to output decisions. Identification of the output elasticities and consequently the distribution of firm-level productivity is constrained by endogeneity issues. This is because productivity is unobserved by the econometrician, but observed by the firm when making input decisions.

A popular approach to address this issue is to introduce a proxy variable such as investment, made popular by Olley and Pakes (1996) (OP) or an intermediate input using Levinsohn and Petrin (2003) (LP) or Ackerberg *et al.* (2015) (ACF). These proxies are functions of a state variable such as capital and unobserved productivity. Under certain assumptions, this function is strictly increasing in its scalar unobserved productivity component. Inverting this function controls for unobserved

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productivity and the production function parameters can be estimated with a simple two-stage estimator.

While these methods have been useful in identifying the production function parameters and recovering consistent estimates of total factor productivity (TFP), resulting estimates may be biased if there is additional heterogeneity in production technology across firms. Thus, allowing for heterogeneous coefficients is one possible way to capture these differences. The literature on heterogeneous production functions is small relative to the empirical research using the homogeneous coefficient model, even though many empirical studies have found heterogeneity in firm behavior and decisions.¹ This is because estimating the homogeneous coefficient model by itself is very difficult due to the issue of unobserved productivity.

In our approach, we allow for non-neutrality of the unobserved idiosyncratic production shock, but require the component of productivity that is anticipated by firms to be Hicks-neutral. We use the proxy variable approach in this framework in order to control for the part of production unobservables that are correlated with input decisions. The literature on control function approaches for quantile regression models is still a developing area therefore it is not straightforward to estimate production functions by allowing for endogenous inputs and their heterogeneous coefficients.² We are not aware of any published paper which takes into account the endogeneity issue of production functions in the conventional quantile regression framework. We fill the gap in this paper by proposing an easy-to-implement estimator.

We show through simulation, that our proposed two-step estimator performs well when productivity is estimated using the static control function approaches of either Levinsohn and Petrin (2003) or Ackerberg *et al.* (2015). The estimator is successful in capturing both heterogeneous output elasticities and controlling for unobserved productivity. In our empirical application, we consider several popular firm and plant-level manufacturing datasets and compare our estimator to the ACF estimator. We show that heterogeneity in these estimates implies differences in other features of firm production, such as returns to scale, capital intensity and productivity.

The rest of the paper is organized as follows. Section 2 reviews prior approaches for production function estimation and the interpretations of quantile production functions. Section 3 introduces the econometric model and the proposed estimator. Section 4 presents finite-sample behaviors of the estimator via Monte Carlo experiments and Section 5 applies this estimator to US, Chilean, and Colombian manufacturing datasets. Section 6 concludes with directions for future research.

¹Some notable examples are Kasahara, Schrimpf and Suzuki (2017), Balat, Brambilla and Sasaki (2018) and Li and Sasaki (2017) to name of few. Also Gandhi *et al.* (2020) who estimate a nonparametric production function and obtain heterogeneous estimates.

²See for example Chesher (2003), Ma and Koenker (2006), Lee (2007) and ?

2 Literature Review

2.1 Production Function Estimation

We briefly review the ACF (2015) procedure for estimating a *value-added* production function (in logs):

$$y_{it} = \beta_k k_{it} + \beta_l l_{it} + \omega_{it} + \eta_{it}, \quad (1)$$

where y_{it} denotes value-added output for firm i at time t , l_{it} denotes labor input, k_{it} denotes capital input, ω_{it} is unobserved productivity and η_{it} denotes an iid shock to production.³

To control for the correlation between ω_{it} and inputs k_{it} and l_{it} , ACF introduces an intermediate input demand defined as

$$m_{it} = m_t(k_{it}, l_{it}, \omega_{it}), \quad (2)$$

where the function m is strictly increasing in ω_{it} for all k_{it} and l_{it} . Productivity can then be expressed as

$$\omega_{it} = m_t^{-1}(k_{it}, l_{it}, m_{it}). \quad (3)$$

Substituting this equation into the production function they obtain

$$y_{it} = \beta_0 + \beta_k k_{it} + \beta_l l_{it} + m_t^{-1}(k_{it}, l_{it}, m_{it}) + \eta_{it} = \Phi(k_{it}, l_{it}, m_{it}) + \eta_{it}. \quad (4)$$

The function, $\Phi_t(k_{it}, l_{it}, m_{it})$, is identified by the following first stage moment restriction

$$\mathbb{E}[\eta_{it} | \mathcal{I}_{it}] = \mathbb{E}[y_{it} - \Phi_t(k_{it}, l_{it}, m_{it}) | \mathcal{I}_{it}] = 0, \quad (5)$$

where \mathcal{I}_{it} denotes the firm's information at time t . The first stage estimate of Φ_t can be obtained by a local linear regression or a polynomial regression in (k_{it}, l_{it}, m_{it}) . This step controls for the unobservable η_{it} .

A second stage moment restriction identifies the coefficient on capital and labor. Assume that productivity follows an auto-regressive process

$$\omega_{it} = \mathbb{E}[\omega_{it} | \omega_{it-1}] + \xi_{it} = g(\omega_{it-1}) + \xi_{it} \quad (6)$$

where ξ_{it} denotes an innovation to productivity and satisfies $\mathbb{E}[\xi_{it} | \mathcal{I}_{it-1}] = 0$.

Then, the production function parameters can be estimated from the moment restrictions corresponding to orthogonality assumptions on the productivity innovation, ξ_{it} by expressing produc-

³We omit the constant β_0 because it is not separately identified from the mean of productivity

tivity as a function of the parameters β_k and β_l

$$\mathbb{E} \left[\hat{\xi}_{it}(\beta_k, \beta_l) \otimes \binom{k_{it}}{l_{it-1}} \right] = 0 \quad (7)$$

where $\hat{\xi}_{it}(\beta_k, \beta_l)$ is constructed from the residuals of a regression of $(\widehat{\beta_0 + \omega_{it}}) = \widehat{\Phi}_t(k_{it}, l_{it}, m_{it}) - \beta_k k_{it} - \beta_l l_{it}$ on its lagged value for a hypothetical value of (β_k, β_l) .

2.2 Production Functions and Quantile Regression

A quantile regression framework for production functions may facilitate many different interpretations. Since this framework has been applied in the production frontier literature, we briefly review how this may be a natural interpretation and the limitations of such a model. A (stochastic) frontier (SFA) model of production proposed by [Aigner et al. \(1977\)](#) introduces statistical error into a frontier model. Frontier models assume firms deviate from an optimal frontier of production. The SFA model is typically written as

$$y_i = f(x_i, \beta) + \varepsilon_i, \quad (8)$$

where $\varepsilon_i = \eta_i - u_i$, x_i are inputs to production and β are the parameters. The error term η_i denotes the statistical noise in the model such as measurement error and u_{it} represents one-sided deviations from the production frontier. Estimates of β are typically obtained using maximum likelihood which requires strong distributional assumptions on the error terms. [Bernini et al. \(2004\)](#) suggests quantile regression could be used to estimate the highest percentiles of the conditional output distribution as an estimate of the production frontier. [Aragon et al. \(2005\)](#) uses a characterization of a deterministic frontier that interprets production functions as being from a continuous interval $\tau \in [0, 1]$ where $\tau = 1$ corresponds to the efficient frontier. The difficulty with these approaches is choosing which quantile to estimate as the frontier and performing inference on these extremal quantiles.⁴ Predicting differences between the frontier and a production process for a given firm is also complicated since the predicted error is a composite of technical inefficiency and noise. Other econometric issues, such as endogeneity of input choices with respect to inefficiency are difficult to incorporate in this framework. The purpose of this paper is not to debate the advantages and disadvantages of production frontier models so we leave this discussion for future studies and acknowledge the theoretical challenges of quantile frontiers and instead focus on the standard production function model.

There are two main challenges with implementing a quantile regression framework to the standard production function model. First, addressing the endogeneity of ω_{it} using traditional panel

⁴See [Chernozhukov \(2005\)](#) on how to extend inference in quantile regression to extreme tails in the conditional distribution

data methods or control function approaches for quantile models is still a developing field that faces many challenges such as incidental parameters or non-smooth criterion functions. Second, it is not straightforward to link firms output decisions to variations in the ex-post shock η_{it} without imposing some strict structural model of firm production.

Addressing the first point, we will argue that recent quantile panel data and quantile IV models should not be used for the same reasons panel data and IV models are not used for the conditional mean model. Quantile panel data models allow for flexible interactions between unobserved heterogeneity and the quantiles of the conditional response function. Some well known approaches assume a time-invariant fixed effect such as [Koenker \(2004\)](#), [Lamarche \(2010\)](#), or [Canay \(2011\)](#) which assumes the fixed effect shifts only the location of the conditional quantile function. This approach has two disadvantages. First, assuming the unobservable is time-invariant is restrictive. [Griliches and Hausman \(1986\)](#) show that a fixed effect for productivity, ω_i , leads to low estimates of the capital elasticity. Second, with the exception of the [Canay \(2011\)](#) estimator, the fixed effects of these models are incidental parameters so as the sample size grows, so does the number of parameters that need to be estimated which makes it computationally costly. An alternative to fixed effect estimation is to model the unobserved heterogeneity as a projection onto the observables plus a disturbance in the spirit of [Chamberlain \(1984\)](#). This correlated random effect (CRE) approach was used by [Abrevaya and Dahl \(2008\)](#) and variations of this type of estimator have been developed by other researchers. One issue with the CRE approach is that identification of the conditional quantile function is difficult because it now depends on the joint distribution of unobservables in the response function and the random effect.

Another alternative is to make use of valid instruments if they are available. The conventional argument for using input prices p_{it}^k and p_{it}^l as instruments is that they must be uncorrelated with the error term $\omega_{it} + \eta_{it}$ and correlated with input choices for capital and labor. Then one could use two-stage least squares to obtain consistent estimates of β_k and β_l . This idea can be extended to quantile-IV models such as [Chernozhukov and Hansen \(2005\)](#). In their identification arguments, one would need to strengthen assumptions to conditional independence as well as monotonicity of a quantile structural function (QSF) in $U_{it} = \omega_{it} + \eta_{it}$. Then if one writes the QSF for the production function as $y_{it} = Q(k_{it}, l_{it}, U_{it})$, their the model is identified using

$$P[y_{it} \leq Q_\tau(y_{it}|k_{it}, l_{it})|k_{it}, l_{it}, p_{it}^k, p_{it}^l] = \tau, \quad \tau \in (0, 1]. \quad (9)$$

We do not use this procedure since input prices may not have enough variation across firms and exogeneity can be violated if they capture input quality differences as shown by [Griliches and Hausman \(1986\)](#). Therefore, it is not straightforward to develop an identification strategy that also leads to a computationally simple estimator. This is a common challenge in the quantile IV literature. However, we will show that a simple adaption of the [Canay \(2011\)](#) estimator works well in practice.

3 A Random Coefficient Production Function

We specify a *value-added* production function as a random coefficient model:

$$y_{it} = \beta_k(\eta_{it})k_{it} + \beta_l(\eta_{it})l_{it} + \omega_{it} \quad (10)$$

The variables in equation (10) have the same interpretation as the ones we introduced in the ACF model. In this specification we allow the output elasticities to be functionally dependent on the production shock η_{it} while productivity ω_{it} maintains its additive separability. The model presented here assumes that ω_{it} is only a location-shift of the conditional output distribution. We assume the constant in the production function does not vary over τ since the constant is not separately identified from the mean of productivity. The location-shift specification is crucial in our approach since $(\beta_0 + \omega_{it})$ also appears in the conditional mean counterpart of equation (10) so that if this object can be consistently estimated from popular proxy variable approaches such as OP, LP or ACF, then with additional restrictions from the quantile IV literature, the production function parameters $\beta_k(\tau)$ and $\beta_l(\tau)$ can also be consistently estimated.

A value-added specification in equation (10) is non-trivial. Value-added production functions are common in the empirical literature, however the objects recovered from a value-added model such as the output elasticities and TFP can only be mapped to its gross-output counterpart under special structural production functions such as Leontief value-added. For example, a Leontief value-added production function in our framework could be written as:

$$y_{it} = \min\{\beta_k(\eta_{it})k_{it} + \beta_l(\eta_{it})l_{it} + \omega_{it}, \beta_m + m_{it}\},$$

where, thanks to the equivariance properties of quantiles we have:

$$Q_\tau(y_{it}|\mathcal{I}_{it}) = \min\{\beta_k(\tau)k_{it} + \beta_l(\tau)l_{it} + \omega_{it}, \beta_m + m_{it}\},$$

so that the conditional quantiles of a structural-value added production function could be written as:

$$Q_\tau(y_{it}|\mathcal{I}_{it}) = \beta_k(\tau)k_{it} + \beta_l(\tau)l_{it} + \omega_{it},$$

which corresponds directly to the conditional quantiles of equation 10. The value-added approach avoids the non-identification results of [Gandhi et al. \(2020\)](#). One consequence of this approach is that estimates of productivity may be more disperse than gross-output estimates which controls for material inputs. In the Appendix, we provide results using estimates from a gross-output production function, one which includes material inputs, to examine whether our estimator proposed in the next section can still reveal heterogeneity in the quantiles.

A natural question that follows from our model specification in 10 is the interpretation of the

unobservable η_{it} . Typically, in the production function literature, this shock has been interchangeably called an “ex-post shock” or measurement error in output. However, in the quantile regression literature, we must consider these interpretations separately because measurement error in the dependent variable will lead to bias in our estimator. For example, [Hausman et al. \(2021\)](#) shows that this bias (in a bivariate setting) understates the true dispersion in quantile regression estimates which they refer to as compression bias. In the Appendix, we discuss how our estimator may be adapted to measurement error in output using their approach. This type of measurement error may be non-negligible in many applications where deflated sales is used as the output measure where the deflator is only industry-specific (common to firms).

Having eliminated measurement error as an interpretation for η_{it} , another possibility is to prescribe the relationship between input usage and output variation that is contributed by uncertainty in realizations of η_{it} . For example consider the location-scale model, which is a special case of (10):

$$y_{it} = \beta_k k_{it} + \beta_l l_{it} + \omega_{it} + (\mu_k k_{it} + \mu_l l_{it})\eta_{it} \quad (11)$$

Which implies that the τ -th conditional quantile of y_{it} is given by

$$Q_\tau(y_{it} | \mathcal{I}_{it}) = \beta_k k_{it} + \beta_l l_{it} + \omega_{it} + (\mu_k k_{it} + \mu_l l_{it})F^{-1}(\tau) \quad (12)$$

where $F^{-1}(\tau)$ is the quantile function of production shocks η_{it} .

The idea that input choices can impact firm’s production beyond the conditional mean has important consequences for firm’s attitude towards production risk. A volume of literature that originated in the late 1970’s challenged the standard stochastic specifications of production functions ([Just and Pope, 1978, 1979](#)) by considering a specification that allowed firm’s inputs to both increase or decrease the marginal variability of final output. These models are commonly applied to studying the agricultural industry where the variance on the yield of harvested crops could be increased by adverse weather or decreased by pesticide usage. Since manufacturing businesses tend to operate in a more controlled environment, risk is less prevalent in these industries so the conditional variance of η_{it} may be smaller.

The connection between production risk and heterogeneity in quantiles is interesting so we briefly mention that under quantile preferences, a firm who maximizes the τ level of utility of profits could explain heterogeneity in the output distribution. Unlike risk-neutral firms, firms could have a utility function that is represented by preferences of the firm manager(s) who decides the optimal expenditure on inputs. Different managers may have different preferences for risk. A short list of papers have considered quantile utility maximization such as [Manski \(1988\)](#), [Rostek \(2009\)](#), [Chambers \(2007\)](#), and [Bhattacharya \(2009\)](#). Dynamic input choices such as investment are much more difficult to solve using the quantile utility framework and the reader can refer to [de Castro and Galvao \(2017\)](#) for a treatment of dynamic quantile utility models. As far as we

know, the quantile utility framework has not been applied to firm decision problems and a more thorough treatment of such is outside the scope of this paper.

A general quantile model such as the one specified in (10) can be seen as an extension of the higher-order moment estimation of risk initiated by Antle (1983). However, it can also be seen purely as an econometric specification issue as we are unaware of any tests that could distinguish between higher order moment production risk and misspecification. We choose the latter interpretation for our model. We briefly mention in the Appendix how a quantile maximization framework may be used in both the proxy variable approach and the approach of Gandhi *et al.* (2020).

3.1 Identification

We show that the production function coefficients $\beta(\tau) = (\beta_k(\tau), \beta_l(\tau))$ are non-parametrically identified with $T = 2$ under conditional independence assumptions and other mild regularity conditions used in conditional deconvolution arguments. Similar ideas have been used in panel data models such as Canay (2011), Neumann (2007), and Evdokimov (2010). To simplify notation, we let $x_{i1} = (k_{i1}, l_{i1})$, $x_{i2} = (k_{i2}, l_{i2})$ and $x_i = (x_{i1}, x_{i2})$. Let $Z_{it} = \beta_k(\eta_{it})k_{it} + \beta_l(\eta_{it})l_{it}$ for $t \in \{1, 2\}$. We then write two consecutive period of output as $y_{i1} = Z_{i1} + \omega_{i1}$ and $y_{i2} = Z_{i2} + \omega_{i2}$. We assume a linear AR(1) process for productivity $\omega_{i2} = \rho\omega_{i1} + \xi_{i2}$. Plugging into the second line of the previous equation allows us to write $y_{i2} = Z_{i2} + \rho\omega_{i1} + \xi_{i2}$. Dividing by ρ we write $\tilde{y}_{i2} = y_{i2}/\rho = \tilde{Z}_{i2} + \tilde{\xi}_{i2} + \omega_{i1}$. Therefore, our two repeated measures of productivity are given by:

$$\begin{aligned} y_{i1} &= Z_{i1} + \omega_{i1} \\ \tilde{y}_{i2} &= \tilde{Z}_{i2} + \tilde{\xi}_{i2} + \omega_{i1} \end{aligned} \tag{13}$$

We present our identification argument in two steps. In the first steps, we utilize conditional deconvolution arguments to identify relevant conditional characteristic functions which are known up to parameters ρ as well as the mean production coefficients $\beta^\mu = (\beta_k^\mu, \beta_l^\mu) = \mathbb{E}[\beta(\eta_{it})]$. In the second step, we argue that these parameters are identified by standard assumptions in the proxy variable approach of Olley and Pakes (1996), Levinsohn and Petrin (2003), or Ackerberg *et al.* (2015) which are both based on conditional mean restrictions. We begin with some assumptions.

Assumption 3.1 Suppose the following set of assumptions

- (a) *Random Sample:* The random variables $(y_i, Z_i, \omega_i)_{i=1}^N$ are identically and independently distributed and $T = 2$
- (b) *Conditional Independence:* (i) $f_{Z_{it}|x_i, Z_{is}, \omega_{it}}(Z_{it}|x_i, Z_{is}, \omega_{it}) = f_{Z_{it}|x_i}(Z_{it}|x_i) \forall t \neq s$ and $t \in \{1, 2\}$ and (ii) $f_{\xi_{i2}|x_i, Z_{it}, \omega_{i1}}(\xi_{i2}|x_i, Z_{it}, \omega_{i1}) = f_{\xi_{i2}|x_i}(\xi_{i2}|x_i)$ for $t \in \{1, 2\}$
- (c) *Characteristic Functions:* The conditional characteristic functions $\phi_{Z_{i1}|x_i}(s|x_i)$, $\phi_{\omega_{i1}|x_i}(s|x_i)$ and $\phi_{\xi_{i2}|x_i}(s|x_i)$ do not vanish $\forall s \in \mathbb{R}$

(d) *Quantiles:* $\eta_{it} \perp\!\!\!\perp (x_i, \omega_{it})$ where $\eta_{it} \sim U[0, 1]$ and Z_{it} is strictly increasing in η_{it}

In addition to Assumption 3.1 we restate and augment the assumptions in Ackerberg *et al.* (2015)

Assumption 3.2 *Suppose the following set of assumptions*

- (a) *Information Set:* The firm's information set at time t is given by \mathcal{I}_{it} and includes current and past productivity shocks, but does not include future productivity shocks
- (b) *Productivity:* The evolution of productivity follows a linear AR(1) process: $\omega_{it} = \rho\omega_{it-1} + \xi_{it}$ where $\mathbb{E}[\xi_{it}|\mathcal{I}_{it-1}]$ and ω_{it} is strictly stationary, $f_{\omega_{i1}}(\omega_{i1}) = f_{\omega_{i2}}(\omega_{i2})$
- (c) *Timing of Input Choices:* Firms accumulate capital according to

$$k_{it} = \kappa(k_{it-1}, k_{it-1})$$

(d) *Scalar Unobservable:* Firm's intermediate input demand is given by

$$m_{it} = m_t(k_{it}.l_{it}, \omega_{it})$$

(e) *Strict Monotonicity:* $m_t(k_{it}.l_{it}, \omega_{it})$ is strictly increasing in ω_{it}

Assumption 3.1(b)(i) implies that the ex-post production shocks are mutually independent over time and independent from productivity ω_{it} . This is a potentially strong condition, but commonly used in nonseparable panel data models. Related conditions are used in Hu *et al.* (2019) who assume that η_{it} is independent of ω_{it} conditional on x_i . Part (ii) of this assumption implies that the innovation shocks to productivity ξ_{it} are independent of the ex-post shocks η_{it} and that these shocks are independent of lagged productivity ω_{it-1} . Assumption 3.1(c) are mild technical conditions on characteristic functions. Most standard families of distributions satisfy this property. Lastly, 3.1 (d) extends the independence assumptions of η_{it} and ω_{it} to x_i . A similar independence assumption is used in Gandhi *et al.* (2020). This assumption is used to identify the relevant conditional quantiles.

Assumption 3.2 extends some of the assumptions in Ackerberg *et al.* (2015) to identify β^μ and ρ which are necessary in our deconvolution argument. 3.2(b) requires productivity to follow a linear AR(1) process. Although restrictive, this allows us to easily adapt a repeated measure type argument used for deconvolution. Strict stationarity is also restrictive and is used to achieve identification of the marginal distribution of productivity for each time period. With these assumptions stated we are now ready to state the main identification result.

Theorem 3.1 *Under Assumptions 3.1 and 3.2, the location β^μ , coefficient ρ and function $\beta(\tau)$ are identified*

Proof: Using Assumption 3.1(b)(i) we can write the characteristic functions of (13) as

$$\begin{aligned}\phi_{y_{i1}|x_i}(s|x_i) &= \phi_{Z_{i1}|x_i}(s|x_i)\phi_{\omega_{i1}|x_i}(s|x_i) \\ \phi_{\tilde{y}_{i2}|x_i}(s|x_i) &= \phi_{\tilde{Z}_{i2}|x_i}(s|x_i)\phi_{\xi_{i2}|x_i}(s|x_i)\phi_{\omega_{i1}|x_i}(s|x_i)\end{aligned}\tag{14}$$

The joint characteristic function of $(y_{i1}, \tilde{y}_{i2})|x_i$ can then be written as:

$$\begin{aligned}\Psi_{(y_{i1}, \tilde{y}_{i2})|x_i}(s_1, s_2|x_i) &= \mathbb{E}[\exp(i[s_1 y_{i1} + s_2 \tilde{y}_{i2}])|x_i] \\ &= \mathbb{E}[\exp(i[s_1 Z_{i1} + s_2 \tilde{Z}_{i2} + s_2 \xi_{i2} + (s_1 + s_2) \omega_{i1}])|x_i] \\ &= \phi_{Z_{i1}|x_i}(s_1|x_i)\phi_{\tilde{Z}_{i2}|x_i}(s_1|x_i)\phi_{\xi_{i2}|x_i}(s_1|x_i)\phi_{\omega_{i1}|x_i}((s_1 + s_2)|x_i)\end{aligned}\tag{15}$$

where the last equality of (15) uses Assumption 3.1(b)(i-ii). Taking the derivative of (15) with respect to its first component yields

$$\begin{aligned}\partial_{s_1} \Psi_{(y_{i1}, \tilde{y}_{i2})|x_i}(s_1, s_2|x_i) &= \partial_{s_1} \phi_{Z_{i1}|x_i}(s_1|x_i)\phi_{\tilde{Z}_{i2}|x_i}(s_1|x_i)\phi_{\xi_{i2}|x_i}(s_1|x_i)\phi_{\omega_{i1}|x_i}((s_1 + s_2)|x_i) \\ &\quad + \phi_{Z_{i1}|x_i}(s_1|x_i)\phi_{\tilde{Z}_{i2}|x_i}(s_1|x_i)\phi_{\xi_{i2}|x_i}(s_1|x_i)\partial_{s_1} \phi_{\omega_{i1}|x_i}(s_1 + s_2|x_i)\end{aligned}$$

Thus,

$$\frac{\partial_{s_1} \Psi_{(y_{i1}, \tilde{y}_{i2})|x_i}(s, -s|x_i)}{\Psi_{(y_{i1}, \tilde{y}_{i2})|x_i}(s, -s|x_i)} = \frac{\phi'_{Z_{i1}|x_i}(s|x_i)}{\phi_{Z_{i1}|x_i}(s|x_i)} + \frac{\phi'_{\omega_{i1}|x_i}(0|x_i)}{\phi_{\omega_{i1}|x_i}(0|x_i)}$$

where $\phi_{\omega_{i1}|x_i}(0|x_i) = 1$ and $\phi'_{\omega_{i1}|x_i}(0|x_i) = i\mathbb{E}[\omega_{i1}|x_i]$ by properties of characteristic functions. We first discuss identification of $\mathbb{E}[\omega_{i1}|x_i]$. Note that,

$$\begin{aligned}\mathbb{E}[\omega_{i1}|x_i] &= \mathbb{E}[y_{i1} - Z_{i1}|x_i] = \mathbb{E}[y_{i1} - x_{i1}\beta(\eta_{i1})|x_i] = \mathbb{E}[y_{i1}|x_i] - x_{i1}\mathbb{E}[\beta(\eta_{i1})|x_i] \\ &= \mathbb{E}[y_{i1}|x_i] - x_{i1}\mathbb{E}[\beta(\eta_{i1})] = \mathbb{E}[y_{i1}|x_i] - x_{i1}\beta^\mu\end{aligned}$$

where the second to last equality follows from the independence assumption in (3.1)(d). In order to identify $\mathbb{E}[\omega_{i1}|x_i]$ we must identify β^μ . The location parameters are identified by Assumption 3.2. Then the conditional characteristic function $\phi_{Z_{i1}|x_i}(s|x_i)$ is identified through

$$\phi_{Z_{i1}|x_i}(s|x_i) = \exp\left(\int_0^s \frac{i\mathbb{E}[y_{i1} \exp(i\gamma(y_{i1} - \tilde{y}_{i2}))|x_i]}{\phi_{y_{i1}-\tilde{y}_{i2}|x_i}(\gamma|x_i)} d\gamma - is\mathbb{E}[\omega_{i1}|x_i]\right)$$

Then the characteristic function $\phi_{\omega_{i1}|x_i}(s|x_i)$ is identified using

$$\phi_{\omega_{i1}|x_i}(s|x_i) = \frac{\phi_{y_{i1}|x_i}(s|x_i)}{\phi_{Z_{i1}|x_i}(s|x_i)}$$

From the productivity process, we have $\tilde{\omega}_{i1} = \omega_{i2}/\rho = \omega_{i1} + \tilde{\xi}_{i2}$. Using Assumption 3.1(b)(ii) we

can write the characteristic function of $\phi_{\omega_{i2}|x_i}(s|x_i)$ as

$$\phi_{\tilde{\omega}_{i2}|x_i}(s|x_i) = \phi_{\omega_{i1}|x_i}(s|x_i)\phi_{\tilde{\xi}_{i2}|x_i}(s|x_i)$$

Then using Assumption 3.2(b) the characteristic functions $\phi_{\omega_{i1}|x_i}(s|x_i) = \phi_{\omega_{i2}|x_i}(s|x_i)$ so we can write

$$\phi_{\xi_{i2}|x_i}(s|x_i) = \frac{\phi_{\omega_{i1}|x_i}(s|x_i)}{\phi_{\omega_{i1}|x_i}(\rho s|x_i)}$$

Finally, $\phi_{Z_{i2}|x_i}(s|x_i)$ is identified using

$$\phi_{Z_{i2}|x_i}(s|x_i) = \frac{\phi_{y_{i2}|x_i}(s|x_i)}{\phi_{\xi_{i2}|x_i}(s|x_i)\phi_{\omega_{i1}|x_i}(\rho s|x_i)}$$

The characteristic functions are identified up to an unknown parameter ρ . This parameter is identified from Assumption 3.2(b) via the moment condition $\mathbb{E}[\xi_{it}|\mathcal{I}_{it-1}] = \mathbb{E}[\omega_{it} - \rho\omega_{it-1}|\mathcal{I}_{it-1}]$ where $\omega_{it} = \hat{\Phi}_t(k_{it}, l_{it}, m_{it}) - \beta_k^\mu k_{it} - \beta_l^\mu l_{it}$ using ACF. Identification of the conditional characteristic function implies identification of the conditional distributions. For example, the distribution of $Z_{i1}|x_i$ and $Z_{i2}|x_i$ can be recovered using the result of ?

$$F_{Z_{i1}|x_i}(Z_{i1}|x_i) = \frac{1}{2} - \lim_{v \rightarrow \infty} \int_{-v}^v \frac{e^{-isZ_{i1}}}{2\pi is} \phi_{Z_{i1}|x_i}(s|x_i) ds$$

$$F_{Z_{i2}|x_i}(Z_{i2}|x_i) = \frac{1}{2} - \lim_{v \rightarrow \infty} \int_{-v}^v \frac{e^{-isZ_{i2}}}{2\pi is} \phi_{Z_{i2}|x_i}(s|x_i) ds$$

The conditional quantiles are the inverse of the conditional distribution functions given by

$$Q_\tau(Z_{i1}|x_i) = Q_\tau(x_{i1}\boldsymbol{\beta}(\eta_{i1})|x_i) = x_{i1}\boldsymbol{\beta}(\tau)$$

$$Q_\tau(Z_{i2}|x_i) = Q_\tau(x_{i2}\boldsymbol{\beta}(\eta_{i2})|x_i) = x_{i2}\boldsymbol{\beta}(\tau)$$

where the third equality in each equation follows from Assumption (3.1)(d)

3.2 A Two-Step Estimator for Quantile Production Functions

Instead of using sample analogues based on the identification approach in the previous section, we propose a more simple estimation routine. Let $\varepsilon_{it} = k_{it}[\beta_k(\eta_{it}) - \beta_k^\mu] + l_{it}[\beta_l(\eta_{it}) - \beta_l^\mu]$ where $(\beta_k^\mu, \beta_l^\mu) = \boldsymbol{\beta}^\mu = \mathbb{E}[\boldsymbol{\beta}(\eta_{it})]$ is the mean of the random coefficients. The conditional mean version of the random coefficient production function can then be written as

$$y_{it} = \beta_k^\mu k_{it} + \beta_l^\mu l_{it} + \omega_{it} + \varepsilon_{it}, \quad (16)$$

where $\mathbb{E}[\varepsilon_{it}|\mathcal{I}_{it}] = 0$. This identity is useful since we can now use a consistent estimator of ω_{it} which can be plugged into the random coefficient production function in equation (10). In the ACF

approach, productivity is estimated as $\omega_{it} = \hat{\Phi}_t(k_{it}, l_{it}, m_{it}) - \hat{\beta}_k^\mu k_{it} - \hat{\beta}_l^\mu l_{it}$. Although we use ACF to identify the location in our model, we can easily extend the results of OP or LP when investment is used as a proxy or when the timing of labor matters. Recently, there has been criticism of these proxy variable approaches as they do not identify the elasticity of the flexible input in a gross-output production function. Therefore, it would be important to include the identification results of [Gandhi et al. \(2020\)](#) in our model. We briefly outline how to do this in the Appendix and focus on estimation of productivity using ACF for simplicity.

Once estimates of ω_{it} are obtained, estimates of $\beta_k(\tau)$ and $\beta_l(\tau)$ can be obtained from a quantile regression of $\hat{y}_{it} = y_{it} - \hat{\omega}_{it}$ with a constant omitted in estimation. To summarize these steps:

1. Let $\hat{\beta}_k^\mu$ and $\hat{\beta}_l^\mu$ be consistent estimators of β_k^μ and β_l^μ from a value-added production function. For example, using the OP, LP, or ACF estimation strategy. Construct the estimator, $\hat{\omega}_{it}$ using these estimates
2. Let $\beta(\tau) = (\beta_k(\tau), \beta_l(\tau))$ and $\hat{y}_{it} = y_{it} - \hat{\omega}_{it}$. Define the two-step estimator of $\beta_k(\tau)$ and $\beta_l(\tau)$ as:

$$\hat{\beta}(\tau) = \underset{\beta \in \mathcal{B}}{\operatorname{argmin}} \mathbb{E}[\rho_\tau(\hat{y}_{it} - \beta_k k_{it} - \beta_l l_{it})],$$

where \mathcal{B} is a compact and convex parameter space, $\rho_\tau(u) = u[\tau - \mathbb{I}\{u < 0\}]$ and $\mathbb{I}\{\cdot\}$ denotes the indicator function.

3.3 Large Sample Properties

Our two-step estimator in Section 3.2 relies on a consistent estimator of the first stage estimates of productivity. Standard errors from the asymptotic covariance matrix also include the variance imparted from these preliminary estimates. Therefore, our model falls under a class of generated dependent variables in quantile regression.⁵ The main challenge of our approach is that the first stage is semi-parametric due to the non-parametric function, $\Phi_t(k_{it}, l_{it}, m_{it})$ and the finite parameters β_k^μ and β_l^μ in the ACF procedure. Consistency and asymptotic normality of estimators of two-step conditional moment models is addressed by [Newey and Powell \(2003\)](#) and asymptotic efficiency by [Ai and Chen \(2003\)](#). One could use any linear sieve with known basis functions to estimate these parameters using a sieve-minimum distance (SMD) criterion function. Then consistency and asymptotic normality follow directly from assumptions in those papers.

A consistent estimate for $\Phi_t(k_{it}, l_{it}, m_{it})$ can be obtained from a linear sieve estimator. We let $\Phi_t(k_{it}, l_{it}, m_{it}; \theta) = p^{k_n}(z_{it})\theta_{(t)}$ where $z_{it} = (k_{it}, l_{it}, m_{it})$ and $p^{k_n}(z_{it})$ is chosen to be a tensor-product linear sieve basis, for example, B-spline, Fourier, or power series. The term k_n is the smoothing parameter which is required to grow with the sample size. In practice, we use a power series basis

⁵[Canay \(2011\)](#) is an example of this type of generated dependent variable problem and proves consistency and asymptotic normality. Issues with his derivation is addressed by [Besstremyannaya and Golovan \(2019\)](#). See also [Bhattacharya \(2019\)](#) for a general treatment of these types of models.

function of degree two which is consistent with the empirical implementations of Levinsohn and Petrin (2003) and Ackerberg *et al.* (2015). For deriving the consistency and asymptotic normality of our estimator, we require the following assumptions.

Assumption 3.3

(a) *Sampling:* $(y_{it}^0, k_{it}, l_{it}, m_{it}, \omega_{it})$ are iid defined on the probability space (Ω, \mathcal{F}, P) and take values in a compact set

(b) *Compactness and Convexity:* For all $\tau \in \mathcal{T}$, $\beta(\tau) \in \text{int}\mathcal{B}$, where \mathcal{B} is compact and convex

(c) *Full Rank and Continuity:* y^0 has bounded conditional density a.s. $\sup_{y \in \mathbb{R}} f_{y|k,l,\omega}(y) < K$ and

$$\Pi(\beta, \tau, \omega) = \mathbb{E}[g_\tau(y^0, x, \omega, \beta(\tau))] = \mathbb{E}[(\tau - \mathbb{1}\{y < \beta_k k + \beta_l l + \omega\})(k, l)]$$

The Jacobian matrix $J_\beta(\beta, \tau, \omega) = \frac{\partial}{\partial \beta} \Pi(\beta, \tau, \omega)$ is continuous and has full rank uniformly over $\mathcal{B} \times \mathcal{T} \times \mathcal{W}$ and $J_\omega(\beta, \tau, \omega) = \frac{\partial}{\partial \omega} \Pi(\beta, \tau, \omega)$ is uniformly continuous over $\mathcal{B} \times \mathcal{T} \times \mathcal{W}$

(d) *First Stage Estimates:* The first stage estimates $\hat{\alpha}^\mu = (\hat{\Phi}, \hat{\beta}^\mu)$ admits the following expansion:

$$\sqrt{nT}(\hat{\alpha}^\mu - \alpha^\mu) = \sqrt{nT}\mathbb{E}_{nT}[\psi_{it}] + o_p(1), \quad \mathbb{E}[\psi_{it}] = 0, \quad \mathbb{E}[\psi_{it}\psi_{it}'] < \infty$$

Assumption 3.3 is similar to Assumption 2 in Chernozhukov and Hansen (2006). The main differences lie in the conditions we place on the first stage estimates in our model. Assumption 3.3 is used to derive the influence of the variance of the first stage estimates in the covariance matrix of our quantile regression estimates. In the Appendix, we show how estimates of ψ_{it} can be obtained using Ackerberg *et al.* (2014). We now state the main asymptotic results:

Theorem 3.2 Suppose Assumption 3.3 holds, then for any T , as $n \rightarrow \infty$

$$\sup_{\tau \in \mathcal{T}} \|\hat{\beta}(\tau) - \beta(\tau)\| \rightarrow_p 0,$$

and

$$\sqrt{nT}(\hat{\beta}(\cdot) - \beta(\cdot)) = [-J_\beta(\cdot)]^{-1} \mathbb{E}_{nT}[\varphi_\tau(u_{it}(\tau))x_{it} + J_\omega(\cdot)\delta_{it}] + o_p(1) \rightarrow_d \mathbb{G}(\cdot) \quad \text{in } \ell^\infty(\mathcal{T})$$

where $u_{it}(\tau) \equiv y_{it}^0 - x_{it}\beta(\tau)$, $\delta_{it} \equiv \mu_{xz}\psi_{it}$, $\mu_{xz} \equiv (E[x_{it}], E[p^{k_n}(z_{it})])$, $\varphi_\tau(u) \equiv \tau - \mathbb{1}\{u < 0\}$, $J_\beta(\tau) \equiv J_\beta(\beta, \tau, 0)$, $J_\omega(\tau) \equiv J_\omega(\beta, \tau, 0)$, \mathbb{G} is a mean-zero Gaussian process with variance function $E[\mathbb{G}(\tau)\mathbb{G}(\tau')] = J_\beta(\tau)^{-1}\Sigma(\tau, \tau')[J_\beta(\tau)^{-1}]'$ where $\Sigma(\tau, \tau')$ is given by

$$\Sigma(\tau, \tau') = S(\tau, \tau') + J_\omega(\tau)\Gamma_{\delta g}(\tau') + \Gamma_{g\delta}(\tau)J_\omega(\tau') + J_\omega(\tau)\Gamma_{\delta\delta}J_\omega(\tau')'$$

where $S(\tau, \tau') = (\min\{\tau, \tau'\} - \tau\tau')E[xx']$, $\Gamma_{g\delta}(\tau) = E[g_\tau(y^0, x, \beta(\tau))\delta]$, and $\Gamma_{\delta\delta} = E[\delta^2]$

4 Monte Carlo Experiments

In this section we investigate the performance of our estimator in a set of Monte Carlo experiments. We use a location-scale model for production given by Equation (11) which we provide again in Equation (17) below.

$$y_{it} = \beta_k k_{it} + \beta_l l_{it} + \omega_{it} + (\gamma_k k_{it} + \gamma_l l_{it})\eta_{it}, \quad (17)$$

where $\beta_k = 0.4$ and $\beta_l = 0.6$. We replicate Ackerberg *et al.* (2015) simulations by sampling 1000 datasets consisting of 1000 firms. We simulate optimal input choices for 100 time periods, using the last 10 periods for estimation. We consider three different data generating processes (DGPs) for the scale parameters and the distribution of η_{it} . For DGP1, we let $\gamma_k = 0.6$, $\gamma_l = -0.6$ and $\eta_{it} \sim N(0, \sigma_\eta^2)$ where $\sigma_\eta^2 = 0.01$. This follows the original specification for the error term in Ackerberg *et al.* (2015). In DGP2, $\gamma_k = 0.4$, $\gamma_l = -0.4$ and $\eta_{it} \sim (\sqrt{3\sigma_\eta^2/5})t_5$ so that the error variance is the same as that of DGP1. In DGP3, we add considerable skewness to the ex-post shocks by letting $\gamma_k = 0.5$, $\gamma_l = -0.5$ and $\eta_{it} \sim Lognormal(0.15, \sigma_\eta^2)$.

We show that our estimator is flexible using previously proposed proxy variable estimators. It should only matter that $\hat{\beta}_k^\mu$ and $\hat{\beta}_l^\mu$ can be identified and consistently estimated. It is well known that under certain timing assumptions, that the labor elasticity may not be identified using the first stage of Levinsohn and Petrin (2003). Clearly, this would affect our estimates of productivity which we use in the second stage of our estimator. Therefore, we follow Ackerberg *et al.* (2015) in considering two DGPs where labor can and cannot be identified⁶. In the DGPs prefixed by “LP” in Table 1, we add optimization error to labor which breaks the functional dependence problem argued by ACF. Therefore, LP and ACF estimates of labor and hence productivity, should be consistent. We use the LP estimator for this DGP. In the DGPs prefixed by “ACF” we add serially correlated wages and allow for labor to be chosen in a sub-period $t-b$, where $b = 0.5$. In this setup, we know that the ACF estimates will be consistent since their estimator conditions on labor in the first stage whereas LP does not. Therefore, we use the ACF estimator for this DGP. An AR(1) process is specified for productivity $\omega_{it} = \rho\omega_{it-1} + \xi_{it}$ where $\rho = 0.7$. The variance of ξ_{it} and initial value ω_{i0} is set so that the standard deviation of ω_{it} is constant over time and equal to 0.3.

Table 1 provides estimates of the bias and mean-squared error (MSE) for each DGP of our proposed estimator (DS) and naive quantile regression (QR) that does not control for simultaneity bias. Our estimator performs very well. The bias of our estimator remains low for each τ under each DGP for the error distribution and each DGP under LP and ACF estimates for productivity. The MSE is also small and reaches its minimum at the median ($\tau = 0.5$). Not surprisingly, the QR estimates do not perform well in any situation. Both the bias and MSE are large compared to our estimates and the sign of the bias is positive for capital estimates and negative for labor estimates.

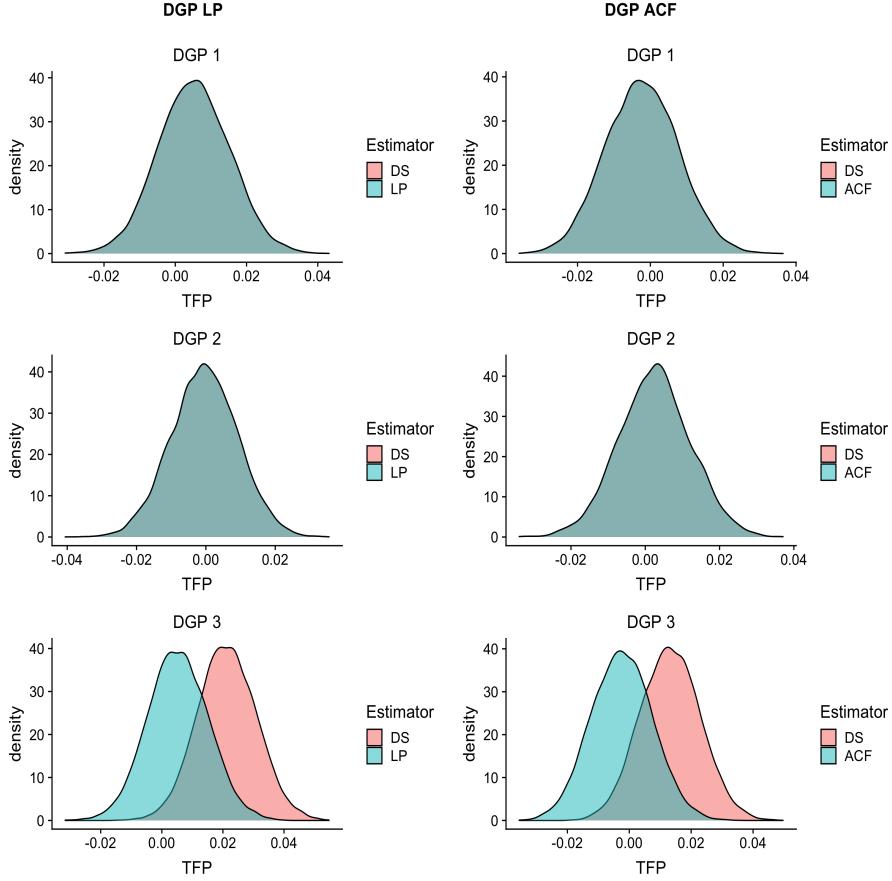
⁶Refer to Ackerberg *et al.* (2015) for a detailed discussion on how to generate the simulated data.

Table 1: Monte Carlo Results for Capital and Labor Estimates

DGP	τ	DS				QR			
		Capital		Labor		Capital		Labor	
		Bias	MSE	Bias	MSE	Bias	MSE	Bias	MSE
LP 1	0.10	-0.0149	0.0004	0.0199	0.0004	0.2691	0.0830	-0.2771	0.0865
	0.25	-0.0045	0.0002	0.0065	0.0000	0.2905	0.0971	-0.2975	0.1008
	0.50	0.0010	0.0002	0.0000	0.0000	0.3140	0.1143	-0.3210	0.1183
	0.75	0.0065	0.0002	-0.0065	0.0000	0.3375	0.1329	-0.3445	0.1380
	0.90	0.0169	0.0005	-0.0199	0.0004	0.3589	0.1513	-0.3659	0.1560
LP 2	0.10	-0.0097	0.0003	0.0127	0.0002	0.2833	0.0610	-0.2903	0.0635
	0.25	-0.0035	0.0002	0.0035	0.0000	0.2985	0.0870	-0.3065	0.0905
	0.50	0.0000	0.0002	0.0000	0.0000	0.3140	0.1149	-0.3210	0.1190
	0.75	0.0035	0.0002	-0.0035	0.0000	0.3285	0.1460	-0.3365	0.1506
	0.90	0.0097	0.0003	-0.0127	0.0002	0.3447	0.1849	-0.3517	0.1910
LP 3	0.10	-0.0053	0.0002	0.0083	0.0001	0.2857	0.0975	-0.2927	0.1013
	0.25	-0.0015	0.0002	0.0035	0.0000	0.2935	0.0998	-0.3005	0.1036
	0.50	0.0004	0.0002	-0.0004	0.0000	0.3064	0.1056	-0.3144	0.1096
	0.75	0.0049	0.0002	-0.0059	0.0000	0.3269	0.1196	-0.3349	0.1245
	0.90	0.0167	0.0005	-0.0197	0.0004	0.3537	0.1459	-0.3617	0.1513
ACF 1	0.10	-0.0099	0.0003	0.0129	0.0003	0.2881	0.0831	-0.2941	0.0865
	0.25	-0.0035	0.0003	0.0045	0.0001	0.3115	0.0971	-0.3175	0.1008
	0.50	-0.0010	0.0002	0.0000	0.0001	0.3380	0.1143	-0.3440	0.1184
	0.75	0.0025	0.0003	-0.0035	0.0001	0.3645	0.1329	-0.3715	0.1380
	0.90	0.0089	0.0003	-0.0119	0.0002	0.3889	0.1513	-0.3949	0.1560
ACF 2	0.10	-0.0057	0.0002	0.0087	0.0002	0.3053	0.0611	-0.3103	0.0635
	0.25	-0.0015	0.0002	0.0025	0.0001	0.3225	0.0870	-0.3285	0.0906
	0.50	0.0000	0.0002	0.0000	0.0001	0.3390	0.1150	-0.3450	0.1190
	0.75	0.0025	0.0002	-0.0025	0.0001	0.3555	0.1461	-0.3615	0.1506
	0.90	0.0067	0.0003	-0.0077	0.0001	0.3727	0.1850	-0.3787	0.1910
ACF 3	0.10	-0.0043	0.0003	0.0043	0.0001	0.3067	0.0975	-0.3127	0.1013
	0.25	-0.0015	0.0003	0.0015	0.0001	0.3145	0.0998	-0.3205	0.1036
	0.50	-0.0006	0.0003	0.0006	0.0001	0.3294	0.1057	-0.3354	0.1096
	0.75	0.0019	0.0003	-0.0029	0.0001	0.3519	0.1196	-0.3579	0.1245
	0.90	0.0097	0.0004	-0.0137	0.0003	0.3817	0.1459	-0.3877	0.1513

1000 replications; “DS” is our estimator; “QR” is quantile regression with no correction for simultaneity bias; “LP” uses estimates of productivity from Levinsohn and Petrin (2003) under the DGPs that are favorable to their procedure; “ACF” uses estimates of productivity from Ackerberg *et al.* (2015) under the DGPs that are favorable to their procedure

Figure 1: Monte Carlo Results for Total Factor Productivity Estimates



We are also interested in the differences in Total Factor Productivity (TFP) at the median and the mean. For Figure 1 we plot densities of TFP using our method and the method of LP and ACF. It is clear for the first two rows DGP1 and DGP2 produce equivalent estimates of TFP since the error distribution was specified as symmetric (Normal for DGP1 and Student-t for DGP2). In DGP3, where the error distribution was specified as log-normal, there are differences between the two estimates. This implies that when the distribution of production shocks are skewed, as is often the case in empirical applications, the median estimate of TFP will be robust to outliers.

5 Application

We apply our estimator to popular plant level manufacturing datasets from Chile and Colombia to examine heterogeneity in the output distribution. For both countries, we examine estimates across different manufacturing industries as well as how these estimates have changed over time. We use the DS estimator presented in this paper and compare it to the ACF estimates. We also compare our estimates to the quantile regression estimates without controlling for productivity. In the ACF

estimation procedure, we estimate the function Φ with a 3rd degree polynomial with interactions in capital, labor and materials. To estimate the coefficient on capital and labor, we use the ACF criterion function mentioned earlier. Since we only need a consistent estimator of the production function parameters, we do not consider any over-identification conditions in this step. We use bootstrap to estimate standard errors of $\beta_k(\tau)$ and $\beta_l(\tau)$ with the number of iterations set to 500.

5.1 US Compustat

The source for the US manufacturing data is from Compustat which covers publicly traded firms and contains data from their financial statements. We collect a sample between 1961 and 2010 on sales, capital expenditures, number of workers, and other expenses to construct measures of output, capital, labor, and material inputs using 3-digit deflators from [Bartelsman and Gray \(1996\)](#). Data preparation follows [Keller and Yeaple \(2009\)](#) and [Dermirer \(2020\)](#). Some issues regarding the Compustat dataset is that since the data is reported in the firm's financial statements, deflated output and input measures may not completely capture firm's actual usage. Also, since this sample only contains publicly traded firms, it is only a fraction of all manufacturing firms in the US. Summary statistics for these deflated values are provided in Table 2. We present a series of output elasticity estimates in Table 3 which are illustrated graphically in Figures 20, 21, 22, and 23.

Estimates of the capital elasticity are slightly increasing in the firm size distribution in every industry as well as the combined sample. The estimates for labor elasticity for each industry and the combined sample are decreasing. For NAICS 33, the labor elasticity increases at very small τ , but then decreases. In each industry, there is evidence that our model captures some heterogeneity compared to the LP model. In NAICS 31, the capital elasticity is smaller and the labor elasticity is larger than the LP estimates for very small τ . In NAICS 32, there is more heterogeneity in the tails of the firm size distribution for capital estimates, but only the labor estimates are lower than the LP estimate for very large τ . The same relationship in these estimates is true in NAICS 33 and in the combined sample.

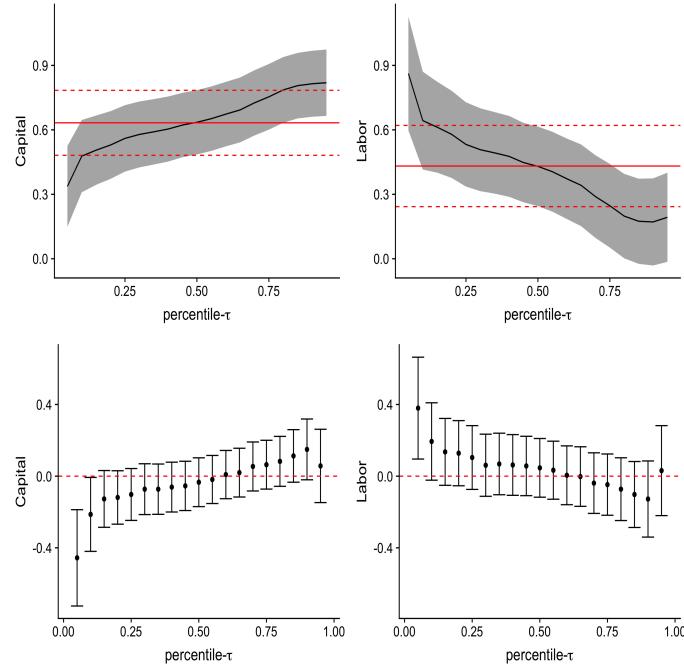


Figure 2: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

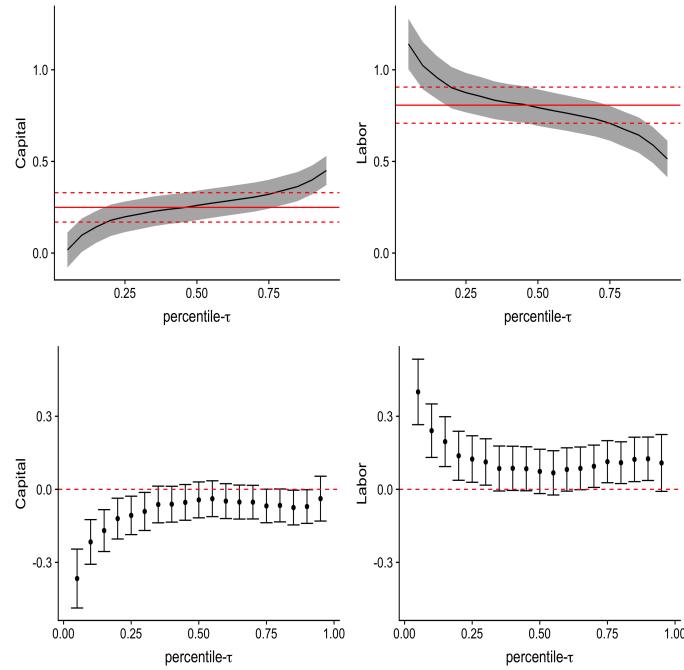


Figure 3: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

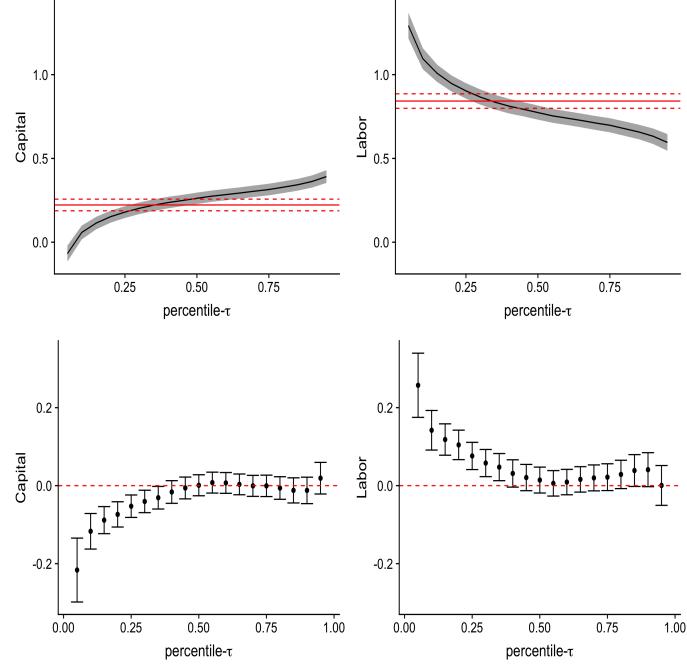


Figure 4: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

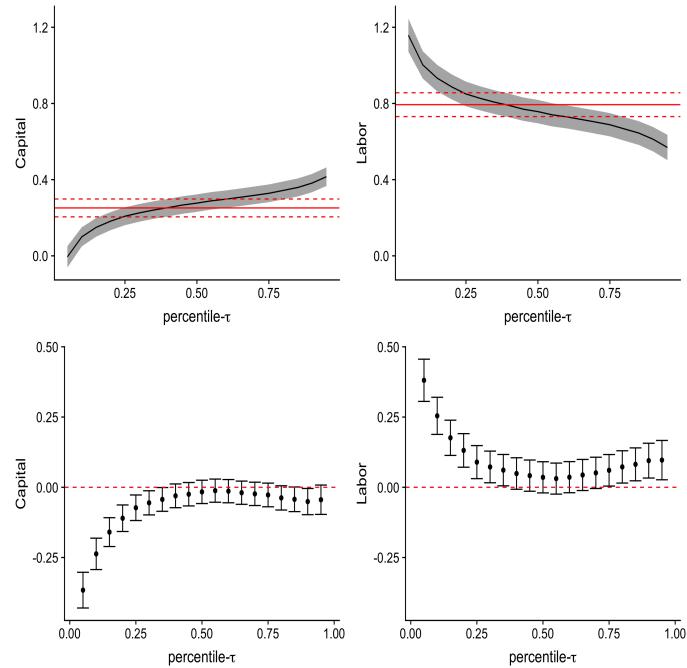


Figure 5: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

In each industry we compare the differences between QLP and QR estimates to test whether our model corrects for endogeneity from unobserved productivity. Bootstrap is used to construct confidence intervals of the difference between the two estimates. We find that there are significant differences between these estimates in all of the samples. This suggests that even after controlling for productivity differences across firms, there is still unaccounted heterogeneity between firms.

We use the estimates from the output elasticities to construct measures of returns to scale and capital intensity in Table 3. The results for returns to scale are puzzling as they are all significantly different from constant returns to scale. Each sample and each quantile of firm size exhibits large decreasing returns to scale and these returns decrease when τ increases. Previous papers that estimate returns to scale using the Compustat dataset such as [Keller and Yeaple \(2009\)](#) and [Dermirer \(2020\)](#) show constant returns to scale using a gross-output production function. Therefore it is possible that the empirical value-added (deflated sales minus intermediate input expenditure) is a poor proxy for value-added in our model and that value-added biases the returns to scale estimates. Differences in returns to scale in value-added and gross-output production functions are explored by [Basu and Fernald \(1997\)](#). We also report estimates of capital intensity measured by the ratio of capital to labor elasticity for each quantile. Each industry has estimates of capital intensity that increases with τ . This result is consistent with previous findings such as [Holmes and Mitchell \(2008\)](#), [Kumar *et al.* \(1999\)](#) and [Dermirer \(2020\)](#).

Table 2: Coefficient Estimates and Standard Errors for US Manufacturing Firms

NAICS	τ	Capital		Labor		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
31	0.10	0.478	0.1022	0.643	0.1385	1.121	0.0504	0.743	0.2789
	0.25	0.560	0.0939	0.532	0.1192	1.092	0.0399	1.053	0.3465
	0.50	0.635	0.0911	0.430	0.1125	1.065	0.0363	1.477	0.5403
	0.90	0.815	0.0934	0.171	0.1232	0.986	0.0463	4.755	224.0343
32	0.10	0.097	0.0555	1.023	0.0785	1.120	0.0297	0.095	0.0635
	0.25	0.198	0.0507	0.875	0.0655	1.073	0.0212	0.226	0.0783
	0.50	0.260	0.0489	0.794	0.0606	1.054	0.0184	0.327	0.0916
	0.90	0.399	0.0480	0.589	0.0595	0.988	0.0185	0.677	0.1641
33	0.10	0.058	0.0250	1.095	0.0383	1.154	0.0169	0.053	0.0248
	0.25	0.180	0.0213	0.903	0.0280	1.083	0.0106	0.199	0.0299
	0.50	0.262	0.0207	0.774	0.0257	1.036	0.0089	0.339	0.0380
	0.90	0.362	0.0221	0.633	0.0281	0.995	0.0095	0.573	0.0606
All	0.10	0.100	0.0304	1.002	0.0436	1.102	0.0161	0.100	0.0350
	0.25	0.208	0.0286	0.850	0.0394	1.058	0.0136	0.245	0.0451
	0.50	0.277	0.0279	0.757	0.0372	1.034	0.0120	0.366	0.0554
	0.90	0.382	0.0291	0.612	0.0389	0.994	0.0122	0.624	0.0885

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using ACF

Table 3: ACF Coefficient Estimates and Standard Errors for US Manufacturing Firms

NAICS	Capital		Labor		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
31	0.633	0.0920	0.432	0.1149	1.064	0.0381	1.466	0.5068
32	0.249	0.0488	0.807	0.0599	1.056	0.0172	0.309	0.0877
33	0.222	0.0211	0.842	0.0263	1.065	0.0089	0.264	0.0334
All	0.252	0.0285	0.794	0.0380	1.045	0.0118	0.317	0.0515

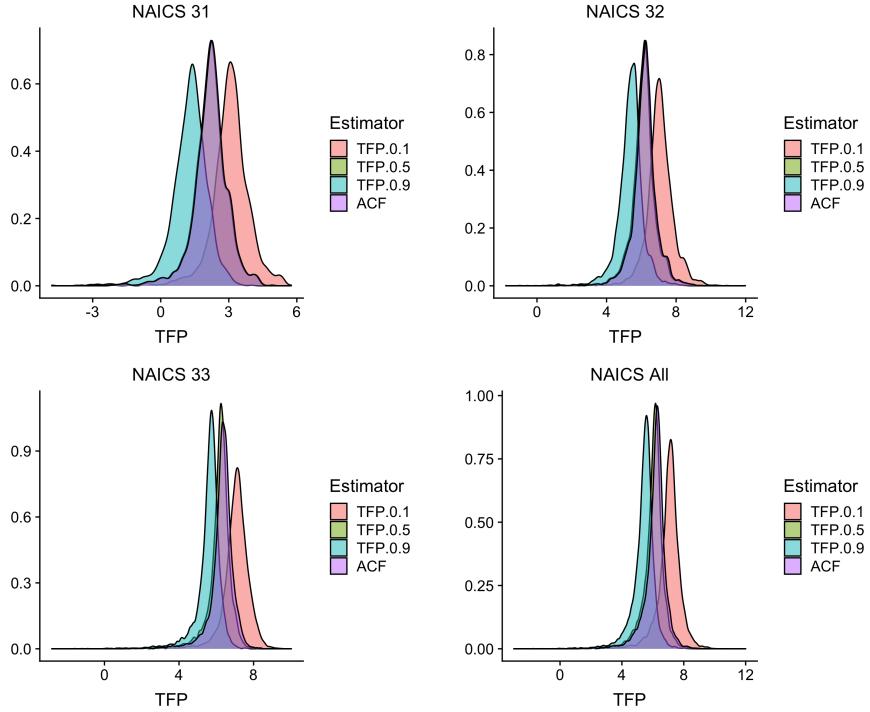
We also use our quantile production function estimates to construct measures of firm level productivity which we define as

$$\hat{w}_{it,\tau} = \exp(y_{it} - \hat{\beta}_k(\tau)k_{it} - \hat{\beta}_l(\tau)l_{it}) \quad (18)$$

We use these measures to compare productivity growth over time to LP estimates over the distribution of firm size. Figure 25 reports average productivity for all US firms in the sample with

the base year of the sample period set to 100. We can see that productivity growth was rapid in the beginning of the sample period but then declined after 1970, but then increases after 1980. Growth trends for each percentile of firm size were similar although larger firms in this sample were more productive than smaller ones. The LP estimates are close to the productivity estimates for medium-sized firms at $\tau = 0.5$.

Figure 6: Mean and Median Estimates of Total Factor Productivity



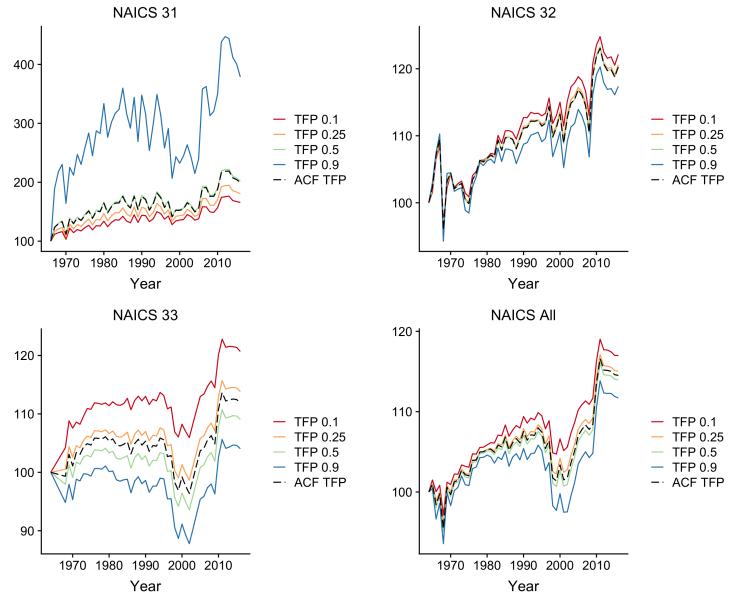


Figure 7: Estimated average productivity over time for the US. Base productivity in 1961 is set to 100.

Table 4: Productivity Differentials for US Manufacturing Firms

NAICS	τ	R&D		Advertising	
		Coef.	s.e	Coef.	s.e
31	0.10	0.028	0.0191	0.044	0.0218
	0.25	0.033	0.0160	0.047	0.0169
	0.50	0.038	0.0155	0.050	0.0158
	0.90	0.058	0.0189	0.067	0.0206
32	0.10	-0.003	0.0142	0.026	0.0139
	0.25	0.018	0.0099	0.043	0.0097
	0.50	0.025	0.0085	0.047	0.0086
	0.90	0.054	0.0101	0.070	0.0107
33	0.10	-0.029	0.0069	-0.006	0.0065
	0.25	0.004	0.0040	0.016	0.0043
	0.50	0.025	0.0038	0.030	0.0039
	0.90	0.039	0.0052	0.037	0.0043
All	0.10	0.004	0.0061	0.035	0.0069
	0.25	0.019	0.0048	0.044	0.0054
	0.50	0.025	0.0046	0.047	0.0045
	0.90	0.037	0.0052	0.054	0.0049

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using the DS-ACF estimator

Table 5: Productivity Differentials for US Manufacturing Firms

NAICS	R&D		Advertising	
	Coef.	s.e	Coef.	s.e
31	0.038	0.0143	0.051	0.0156
32	0.024	0.0082	0.046	0.0084
33	0.010	0.0034	0.019	0.0037
All	0.021	0.0042	0.045	0.0046

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using the ACF estimator

5.2 Chilean Manufacturing

This data comes from the census of Chilean manufacturing plants conducted by the Instituto Nacional de Estadística (INE). The sample is collected between 1979 and 1996 for firms with more

than 10 employees. We divide our estimates into the three largest manufacturing industries: Food (ISIC 311), Fabricated Metals (ISIC 381), and Textiles (ISIC 321). We also aggregate the three industries with the other smaller industries to obtain estimates from the entire sample. Summary statistics for the data we use are provided in Table 15.

Figures 8, 9 and 10 illustrate our estimates from our model compared to ACF estimates as well as their differences to QR estimates. In each industry and the combined sample, estimates of capital elasticity are increasing in firm size. In ISIC 311 and ISIC 381, our model estimates lower capital elasticities than the ACF estimate for small τ . In ISIC 321 and the combined sample, there are differences between these estimates for both small and large τ . The labor estimates in ISIC 311 are increasing but not different from the ACF estimator. In ISIC 381, labor estimates are mostly constant and not different from the ACF estimator. In ISIC 321, labor estimates are decreasing but not different from the ACF estimator. Finally, in the combined sample, our estimator is mostly constant and not different from the ACF estimator. Comparing our estimator and QR estimates, we find that there are differences in the estimates for capital and some evidence of differences in labor for small τ in ISIC 311, 381, and the combined sample. From Table 6, returns to scale are increasing with firm size. Capital intensity is increasing as well except for large firms at $\tau = 0.9$ in ISIC 311.

Figure 13 reports average productivity for all Chilean plants in the sample with base period set to 100. Productivity decreases in the beginning of the 1980s but then increases for the rest of the sample period. These estimates suggest that small firms had higher productivity than large firms. The ACF estimates are similar to productivity of medium-sized firms at $\tau = 0.5$. Figure ?? shows the time trends in output elasticities. The estimates for capital and labor elasticities follow similar trends for each quantile. Capital estimates decrease in the beginning of the sample period then increase, reaching their peak before 1990. These estimates increase in τ . Labor estimates are not much different across quantiles and decrease over time.

Figure 8: Estimated Coefficients of Capital and Labor for Chile: ISIC 311

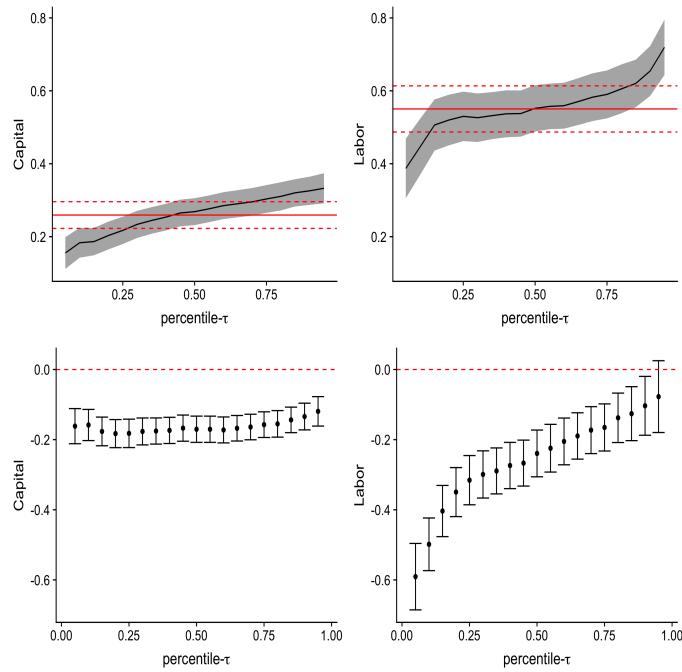


Figure 9: Estimated Coefficients of Capital and Labor for Chile: ISIC 381

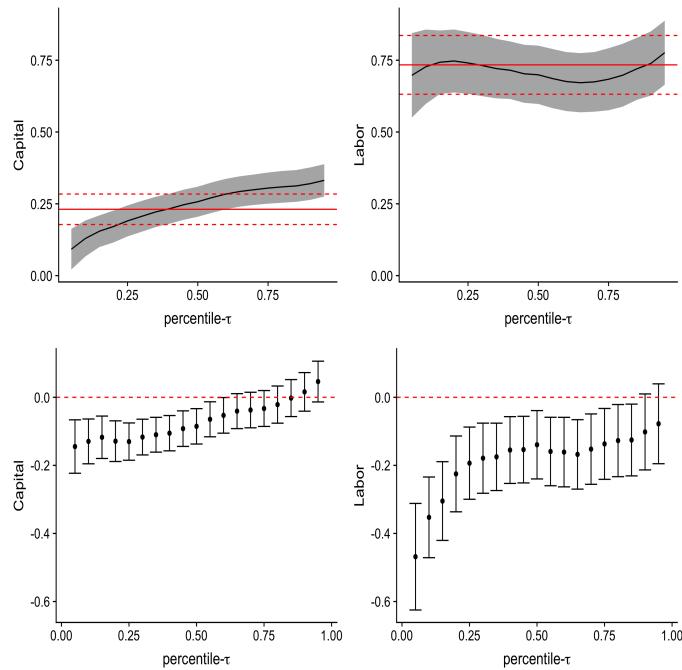


Figure 10: Estimated Coefficients of Capital and Labor for Chile: ISIC 321

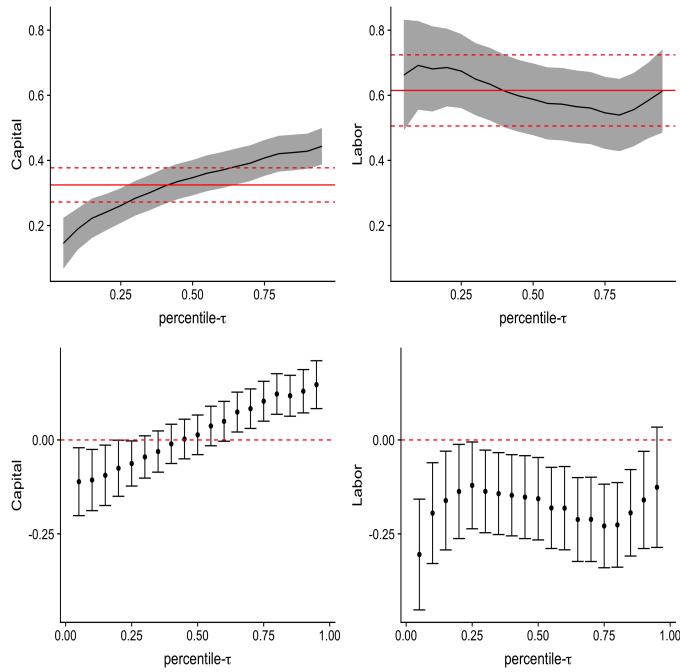


Figure 11: Estimated Coefficients of Capital and Labor for all Chilean Manufacturing Plants

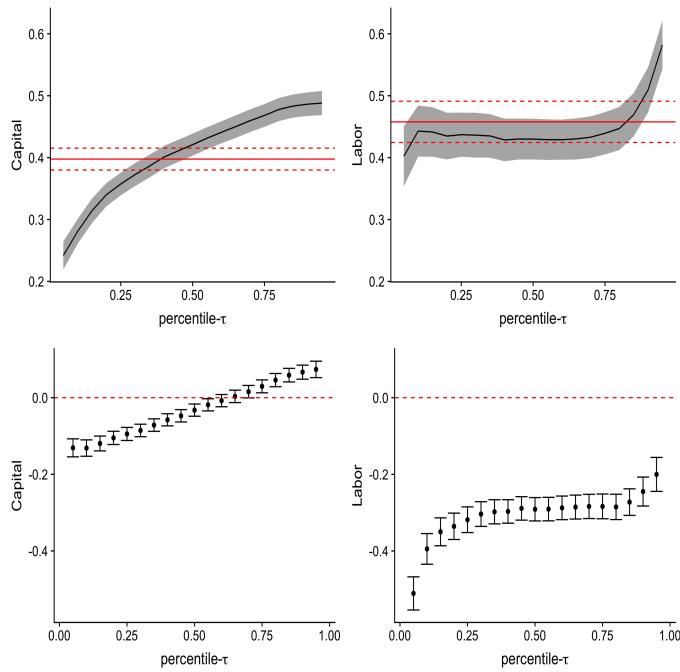


Table 6: Coefficient Estimates and Standard Errors for Chilean Manufacturing Plants

ISIC	τ	Capital		Labor		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.183	0.0249	0.447	0.0468	0.630	0.0321	0.410	0.0760
	0.25	0.217	0.0228	0.530	0.0412	0.747	0.0298	0.410	0.0613
	0.50	0.269	0.0224	0.552	0.0387	0.821	0.0279	0.487	0.0601
	0.90	0.326	0.0235	0.654	0.0416	0.980	0.0292	0.498	0.0546
381	0.10	0.129	0.0380	0.728	0.0786	0.857	0.0535	0.178	0.0626
	0.25	0.190	0.0328	0.740	0.0655	0.931	0.0469	0.257	0.0584
	0.50	0.257	0.0318	0.699	0.0617	0.957	0.0447	0.368	0.0674
	0.90	0.320	0.0344	0.738	0.0679	1.059	0.0473	0.434	0.0744
321	0.10	0.189	0.0387	0.692	0.0828	0.881	0.0578	0.273	0.0789
	0.25	0.261	0.0325	0.674	0.0690	0.935	0.0522	0.387	0.0750
	0.50	0.347	0.0330	0.588	0.0669	0.935	0.0502	0.591	0.0989
	0.90	0.428	0.0328	0.584	0.0705	1.012	0.0537	0.734	0.1220
All	0.10	0.281	0.0123	0.443	0.0251	0.724	0.0182	0.634	0.0441
	0.25	0.357	0.0111	0.437	0.0217	0.795	0.0166	0.818	0.0444
	0.50	0.421	0.0108	0.430	0.0200	0.851	0.0155	0.979	0.0468
	0.90	0.486	0.0116	0.510	0.0221	0.996	0.0163	0.954	0.0453

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using ACF

Table 7: ACF Coefficient Estimates and Standard Errors for Chilean Manufacturing Plants

ISIC	Capital		Labor		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.259	0.0222	0.550	0.0385	0.810	0.0280	0.471	0.0588
381	0.231	0.0323	0.734	0.0622	0.965	0.0447	0.315	0.0615
321	0.325	0.0319	0.615	0.0664	0.940	0.0508	0.528	0.0886
All	0.398	0.0107	0.458	0.0202	0.855	0.0157	0.869	0.0418

*Standard errors are obtained using bootstrap with 500 replications

Figure 12: Mean and Median Estimates of Total Factor Productivity

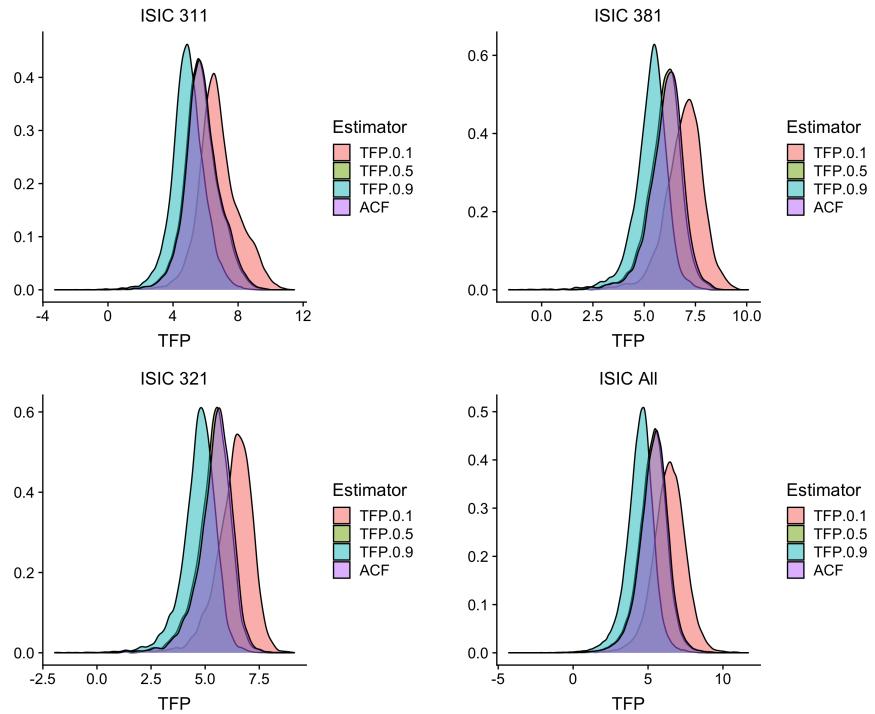


Figure 13: Chile Productivity Over Time

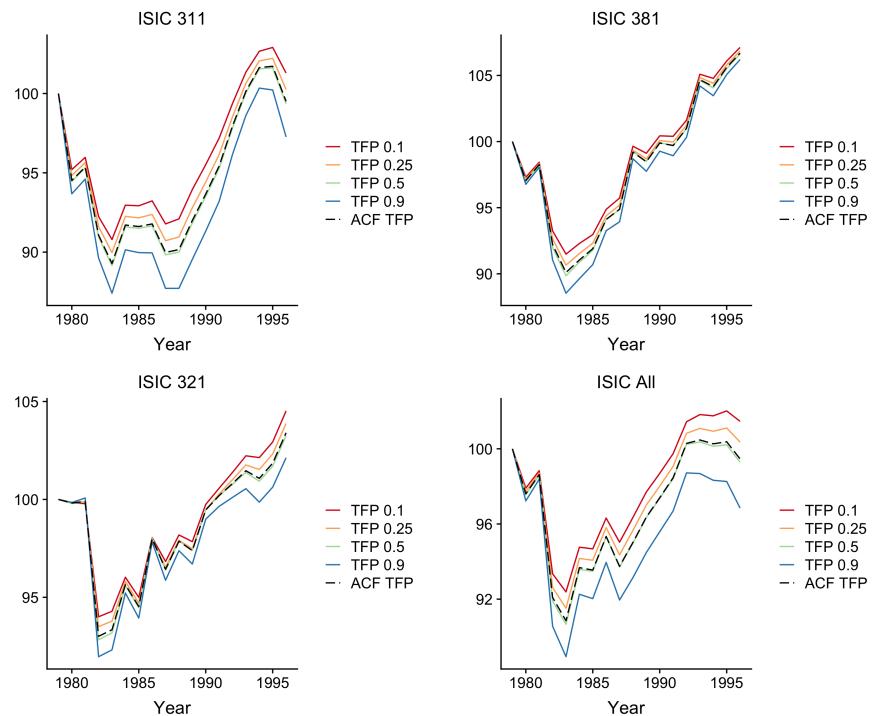


Table 8: Productivity Differentials for Chilean Manufacturing Plants using DS-ACF

ISIC	τ	Exporter		Importer		Advertiser	
		Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.069	0.0377	0.184	0.0397	0.076	0.0279
	0.25	0.057	0.0363	0.144	0.0384	0.054	0.0270
	0.50	0.044	0.0358	0.120	0.0376	0.041	0.0268
	0.90	0.026	0.0369	0.065	0.0374	0.012	0.0270
381	0.10	0.099	0.0299	0.192	0.0348	0.130	0.0377
	0.25	0.068	0.0283	0.149	0.0354	0.105	0.0348
	0.50	0.048	0.0275	0.122	0.0364	0.090	0.0332
	0.90	0.010	0.0277	0.069	0.0388	0.058	0.0338
321	0.10	0.021	0.0278	0.044	0.0374	0.074	0.0327
	0.25	0.005	0.0276	0.018	0.0366	0.056	0.0317
	0.50	0.007	0.0283	0.017	0.0371	0.055	0.0323
	0.90	-0.017	0.0297	-0.020	0.0399	0.030	0.0350
All	0.10	0.101	0.0118	0.192	0.0155	0.143	0.0121
	0.25	0.073	0.0116	0.156	0.0151	0.124	0.0116
	0.50	0.049	0.0116	0.127	0.0150	0.109	0.0114
	0.90	0.005	0.0117	0.067	0.0151	0.073	0.0116

Standard errors are obtained using bootstrap with 500 replications. Log TFP is regressed on Exporter, Importer, and Advertiser which are continuous variables

Table 9: Productivity Differentials for Chilean Manufacturing Plants using ACF

ISIC	Exporter		Importer		Advertiser	
	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.046	0.0359	0.123	0.0377	0.043	0.0269
381	0.051	0.0278	0.126	0.0362	0.092	0.0338
321	0.005	0.0277	0.015	0.0366	0.054	0.0322
All	0.052	0.0116	0.129	0.0149	0.109	0.0115

Standard errors are obtained using bootstrap with 500 replications. Log TFP is regressed on Exporter, Importer, and Advertiser which are continuous variables

5.3 Colombian Manufacturing

This data comes from the Colombian manufacturing census conducted by the Departamento Administrativo Nacional de Estadística. The sample is collected between 1977 and 1991. We divide our estimates into the three largest manufacturing industries: Food (ISIC 311), Apparel (ISIC 322), and Fabricated Metals (ISIC 381). As we did with the Chilean sample, we also aggregate the three industries with other smaller industries to obtain estimates from the entire sample of manufacturing plants. Summary statistics for this data is provided in Table 16.

Figures 14, 15, and 16 illustrate estimates from our model compared to the ACF estimates as well as their differences from QR estimates. The capital estimates in each industry are increasing and are significantly different than the ACF estimates. The labor estimates in each industry are decreasing, however, only in the combined sample are these estimates different from ACF for both low and high τ . Comparing our results to those obtained using QR, there is statistical difference in ISIC 311, difference in the labor estimate in ISIC 321 and ISIC 381 and differences in both capital and labor in the combined sample.

Figure 14: Estimated Coefficients of Capital and Labor for Colombia: ISIC 311

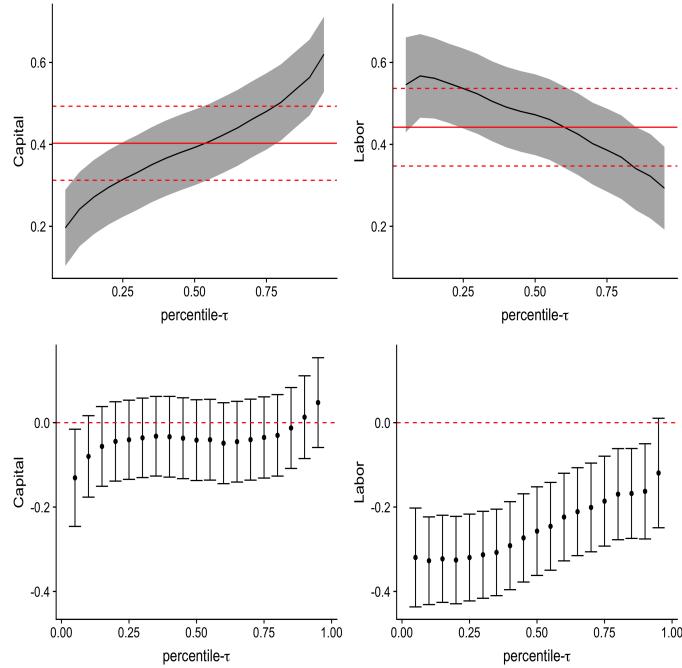


Figure 15: Estimated Coefficients of Capital and Labor for Colombia: ISIC 321

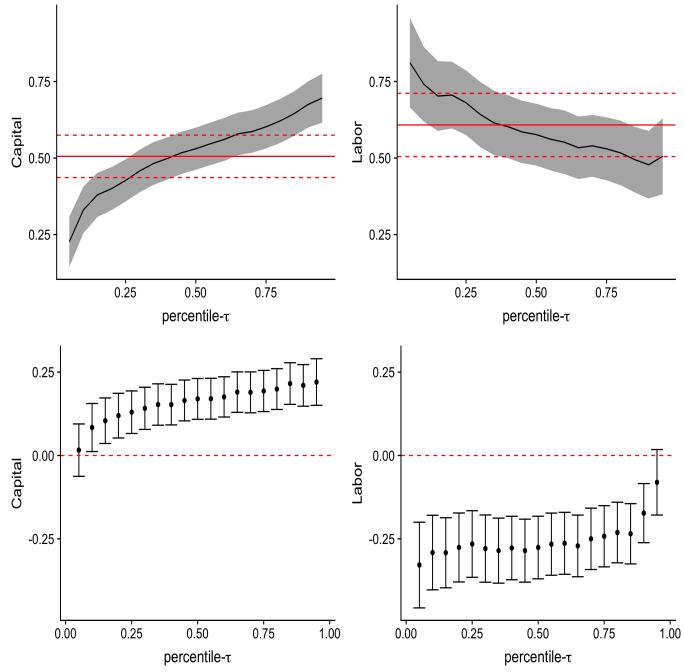


Figure 16: Estimated Coefficients of Capital and Labor for Colombia: ISIC 381

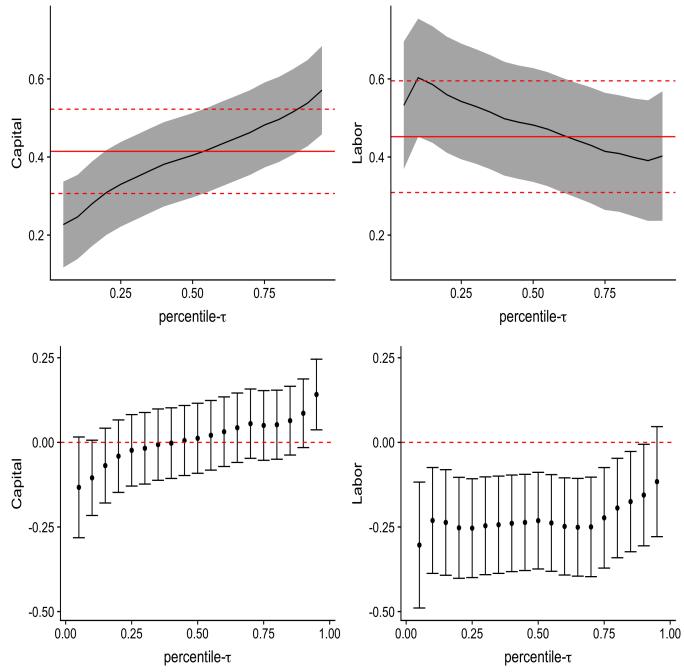
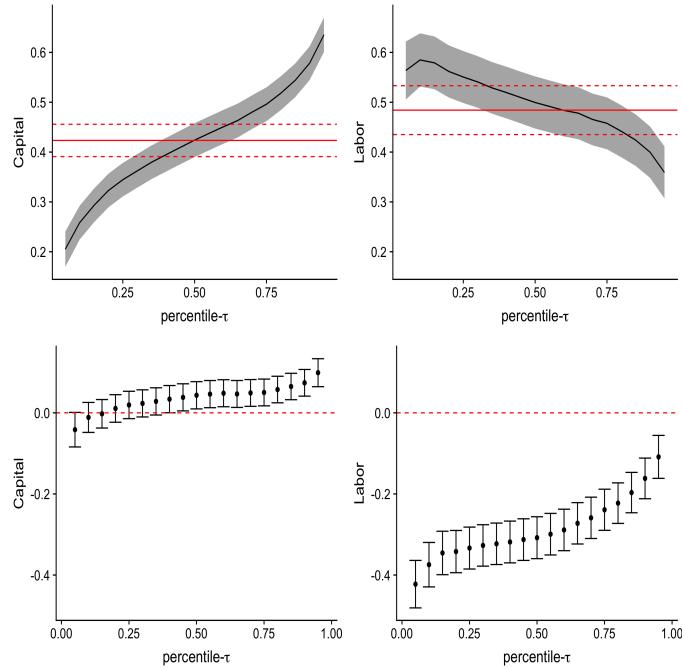


Figure 17: Estimated Coefficients of Capital and Labor for all Colombian Manufacturing Plants



Using these estimates we construct measures of returns to scale and capital intensity for each industry in Table 10. Most firms experience constant or slightly decreasing returns to scale. Returns to scale are smallest in ISIC 311. Returns to scale and capital intensity are increasing in firm size.

Table 10: Coefficient Estimates and Standard Errors for Colombian Manufacturing Plants

ISIC	τ	Capital		Labor		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.242	0.0551	0.567	0.0621	0.809	0.0391	0.427	0.1058
	0.25	0.314	0.0555	0.536	0.0594	0.851	0.0352	0.586	0.1221
	0.50	0.393	0.0562	0.472	0.0600	0.865	0.0344	0.831	0.1575
	0.90	0.563	0.0562	0.323	0.0626	0.886	0.0368	1.746	0.3270
322	0.10	0.331	0.0461	0.741	0.0741	1.071	0.0544	0.446	0.0902
	0.25	0.427	0.0419	0.681	0.0641	1.107	0.0508	0.627	0.1011
	0.50	0.530	0.0419	0.577	0.0614	1.107	0.0489	0.920	0.1409
	0.90	0.675	0.0458	0.478	0.0670	1.153	0.0489	1.413	0.2480
381	0.10	0.247	0.0655	0.603	0.0919	0.850	0.0687	0.409	0.1910
	0.25	0.330	0.0659	0.542	0.0901	0.872	0.0668	0.609	0.2568
	0.50	0.404	0.0656	0.482	0.0886	0.886	0.0660	0.839	0.3616
	0.90	0.538	0.0670	0.391	0.0939	0.929	0.0679	1.377	0.7882
All	0.10	0.258	0.0206	0.585	0.0325	0.843	0.0191	0.442	0.0494
	0.25	0.345	0.0205	0.551	0.0317	0.895	0.0185	0.626	0.0602
	0.50	0.424	0.0207	0.499	0.0318	0.923	0.0183	0.849	0.0780
	0.90	0.578	0.0205	0.400	0.0313	0.978	0.0182	1.447	0.1311

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using ACF

Table 11: ACF Coefficient Estimates and Standard Errors for Colombian Manufacturing Plants

ISIC	Capital		Labor		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.403	0.0550	0.442	0.0576	0.845	0.0346	0.911	0.1660
322	0.506	0.0421	0.608	0.0629	1.114	0.0497	0.831	0.1279
381	0.414	0.0656	0.452	0.0869	0.866	0.0647	0.917	0.4060
All	0.423	0.0197	0.484	0.0298	0.908	0.0177	0.874	0.0765

Figure ?? reports average productivity for all Colombian plants in the sample with base period set to 100. Productivity decreases in the beginning of the sample period, but then increases for the rest of the sample period after 1980 with some periods of sharp changes. Each percentile of firm size has similar productivity levels throughout the sample period. Figure ?? shows the time trends in output elasticities. Capital estimates are generally increases except for a slight dip in 1985. These estimates are increasing in τ . The estimates of labor elasticity are similar in the beginning of the sample period. They increase past 1980 where dispersion across quantiles is largest, then decrease

and dispersion lessens.

Figure 18: Mean and Median Estimates of Total Factor Productivity

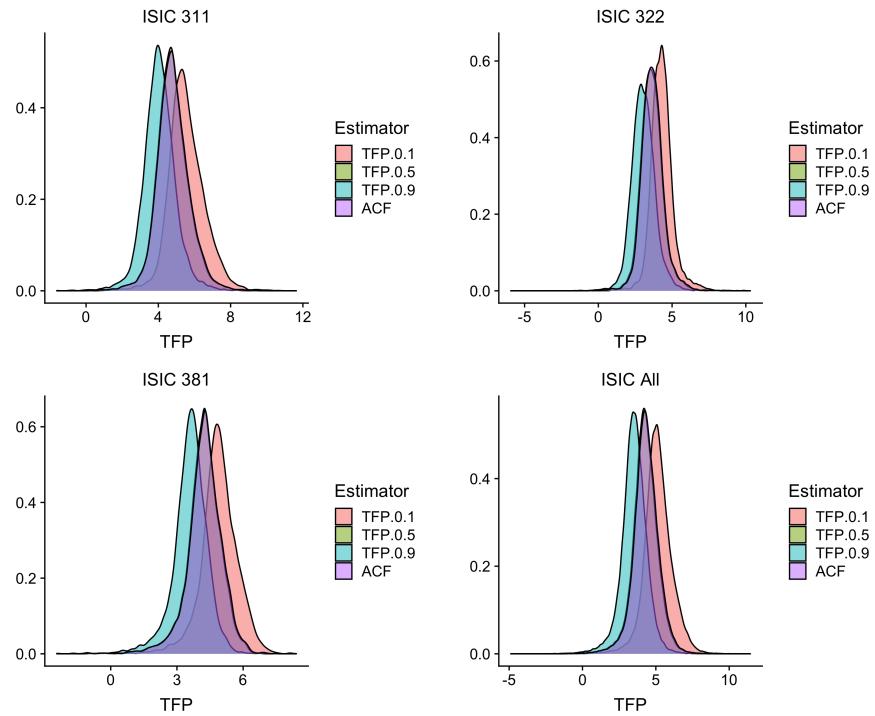


Figure 19: Chile Productivity Over Time

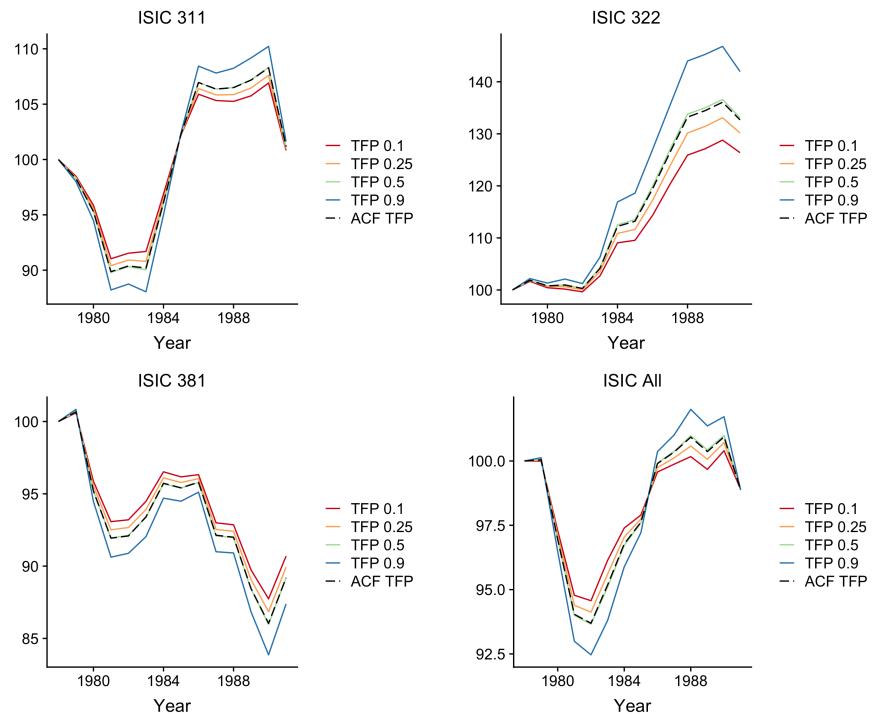


Table 12: Productivity Differentials for Colombian Manufacturing Plants using DS-ACF

ISIC	τ	Exporter		Importer		Advertiser	
		Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.090	0.0323	0.208	0.0420	0.158	0.0350
	0.25	0.074	0.0308	0.194	0.0387	0.151	0.0324
	0.50	0.061	0.0301	0.186	0.0353	0.149	0.0301
	0.90	0.034	0.0300	0.171	0.0305	0.145	0.0271
322	0.10	0.002	0.0243	0.040	0.0294	0.057	0.0254
	0.25	-0.010	0.0230	0.015	0.0291	0.036	0.0253
	0.50	-0.017	0.0225	-0.000	0.0298	0.020	0.0258
	0.90	-0.034	0.0223	-0.035	0.0319	-0.011	0.0274
381	0.10	0.107	0.0384	0.158	0.0424	0.169	0.0372
	0.25	0.091	0.0372	0.141	0.0407	0.151	0.0362
	0.50	0.079	0.0364	0.128	0.0394	0.136	0.0354
	0.90	0.051	0.0347	0.099	0.0369	0.106	0.0341
All	0.10	0.087	0.0118	0.176	0.0117	0.157	0.0096
	0.25	0.068	0.0111	0.148	0.0113	0.136	0.0093
	0.50	0.053	0.0106	0.126	0.0111	0.122	0.0091
	0.90	0.025	0.0099	0.085	0.0110	0.094	0.0091

Standard errors are obtained using bootstrap with 500 replications. Log TFP is regressed on Exporter, Importer, and Advertiser which are continuous variables

Table 13: Productivity Differentials for Colombian Manufacturing Plants using ACF

ISIC	Exporter		Importer		Advertiser	
	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.063	0.0303	0.190	0.0349	0.152	0.0301
322	-0.016	0.0225	0.002	0.0296	0.023	0.0256
381	0.082	0.0363	0.133	0.0391	0.140	0.0352
All	0.056	0.0106	0.130	0.0112	0.125	0.0091

Standard errors are obtained using bootstrap with 500 replications. Log TFP is regressed on Exporter, Importer, and Advertiser which are continuous variables

6 Conclusions

We propose a method that extends the proxy variable approach to estimating the conditional quantiles of firm size. The method is computationally attractive as it resembles the two-stage estimator proposed by [Canay \(2011\)](#). As a result, practitioners are able to easily apply the proposed estimator to production function models where the data may reveal significant heterogeneous output elasticities along the conditional firm size distribution. We showed that this estimator works well in finite samples and showed that it captures heterogeneity in firm size under different data generating processes. An application to widely used datasets from the US, Chile, and Colombia showed that in some industries, our estimator captures unobserved heterogeneity that the LP estimator does not.

Improvements and extensions of this estimator are currently being explored. For example, using a value-added production function may show more heterogeneity in estimates of elasticities and productivity than a gross-output production function. However, using a gross-output production function with an intermediate input proxy variable suffers from non-identification. We are currently working on an identification strategy that allows for different types of production function estimators such as the one proposed by [Gandhi *et al.* \(2020\)](#). This paper also makes an interesting but brief connection between the literature on production risk and quantile utility maximization. Currently, quantile utility maximization problems and estimation of these models are being studied by [de Castro and Galvao \(2017\)](#) and [de Castro *et al.* \(2018\)](#) in the context of dynamic consumption problems. It would be interesting to explore a model for a firm who maximizes quantile utility of profits which could provide an alternative interpretation for unobserved heterogeneity that are obtained from quantile regression estimates.

This paper contributes to the growing literature on production functions with unobserved heterogeneity. We show that differences in firm size correspond to the rank of the ex-post shock. The location-shift model for productivity we propose here restricts us from examining other dimensions of firm heterogeneity. Therefore, allowing richer distributional effects of productivity would be an interesting extension. This approach also restricts us from examining non-Hicks neutral productivity shocks such as factor-augmenting productivity. We are currently working on an extension of this paper to a non-separable model to address these last two points, but the estimator we propose here is computationally attractive and easy to implement in empirical research.

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Appendix

6.1 Summary Statistics

6.1.1 US

Table 14: Summary Statistics (in logs) for US Manufacturing Firms

Industry (NAICS code)		1st Qu.	Median	3rd Qu.	Mean	sd
31 (Total=1788)	Output	17.02	18.86	20.3	18.58	2.11
	Capital	15.65	17.61	19.58	17.53	2.49
	Labor	15.12	16.7	18.28	16.74	2.2
	Materials	15.42	17.09	18.89	17.15	2.35
32 (Total=5070)	Output	15.65	17.76	19.38	17.45	2.29
	Capital	14.24	16.34	18.54	16.35	2.71
	Labor	13.79	15.39	17.06	15.46	2.4
	Materials	14.04	15.8	17.81	15.91	2.56
33 (Total=11117)	Output	7.47	8.85	10.13	8.61	2
	Capital	5.59	7.32	9.5	7.43	2.37
	Labor	5.37	6.9	8.27	6.93	2.05
	Materials	5.55	7.17	8.85	7.24	2.2
All (Total=17975)	Output	16.72	18.54	19.96	18.26	2.1
	Capital	15.26	17.27	19.17	17.09	2.52
	Labor	14.64	16.23	17.83	16.25	2.25
	Materials	14.96	16.66	18.45	16.69	2.4

6.1.2 Chile

Table 15: Summary Statistics (in logs) for Chilean Manufacturing Plants

Industry (ISIC code)		1st Qu.	Median	3rd Qu.	Mean	sd
311 (Total=18562)	Output	10.17	10.76	12.13	11.3	1.59
	Capital	10.43	11.33	12.31	11.43	1.38
	Labor	10.45	11.34	12.44	11.48	1.41
	Materials	10.31	11.18	12.44	11.48	1.61
381 (Total=5289)	Output	6.63	7.54	8.96	7.94	1.98
	Capital	7.36	8.4	9.58	8.51	1.74
	Labor	7.17	8.26	9.46	8.35	1.7
	Materials	7.1	8.22	9.6	8.44	1.92
321 (Total=5413)	Output	3.07	3.48	4.1	3.72	0.96
	Capital	2.85	3.4	4.12	3.65	1.09
	Labor	2.97	3.46	4.27	3.67	0.94
	Materials	2.99	3.51	4.29	3.74	1.04
All (Total=65886)	Output	9.81	10.4	11.7	10.9	1.57
	Capital	9.81	10.7	11.71	10.79	1.42
	Labor	9.68	10.56	11.68	10.72	1.47
	Materials	9.76	10.6	11.82	10.88	1.63

6.1.3 Colombia

Table 16: Summary Statistics (in logs) for Colombian Manufacturing Plants

Industry (ISIC code)		1st Qu.	Median	3rd Qu.	Mean	sd
311 (Total=12631)	Output	9.06	10.25	11.65	10.46	1.8
	Capital	8.73	9.42	10.25	9.54	1.17
	Labor	8.55	9.33	10.36	9.58	1.43
	Materials	8.75	9.67	10.93	9.97	1.67
322 (Total=11600)	Output	6.04	7.09	8.38	7.27	1.78
	Capital	5.48	6.15	6.95	6.25	1.21
	Labor	5.94	6.8	7.86	6.99	1.54
	Materials	5.93	6.93	8.21	7.16	1.77
381 (Total=7070)	Output	2.56	3.14	4.01	3.38	1.11
	Capital	2.77	3.3	3.97	3.4	1
	Labor	2.71	3.22	3.95	3.39	1
	Materials	2.71	3.3	4.14	3.51	1.11
All (Total=83684)	Output	8.43	9.77	11.31	9.91	2.01
	Capital	7.68	8.57	9.5	8.53	1.51
	Labor	7.81	8.69	9.77	8.84	1.58
	Materials	7.94	8.98	10.32	9.19	1.88

6.2 Estimates using Gross-Output Production Function

6.2.1 US

Table 17: Coefficient Estimates and Standard Errors for US Manufacturing Firms

NAICS	τ	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
31	0.10	0.138	0.0482	0.101	0.0188	0.718	0.0741	0.958	0.0387	1.362	0.5810
	0.25	0.159	0.0479	0.088	0.0151	0.708	0.0738	0.956	0.0389	1.817	0.6859
	0.50	0.180	0.0494	0.074	0.0141	0.700	0.0741	0.954	0.0396	2.418	0.9637
	0.90	0.247	0.0603	0.043	0.0312	0.667	0.0780	0.957	0.0409	5.679	172.2314
32	0.10	0.134	0.0220	0.221	0.0185	0.623	0.0389	0.978	0.0221	0.605	0.1186
	0.25	0.148	0.0219	0.206	0.0165	0.623	0.0386	0.976	0.0219	0.717	0.1304
	0.50	0.160	0.0232	0.190	0.0146	0.624	0.0394	0.975	0.0223	0.842	0.1497
	0.90	0.252	0.0317	0.097	0.0202	0.596	0.0431	0.945	0.0234	2.607	0.9094
33	0.10	0.082	0.0179	0.300	0.0119	0.585	0.0343	0.967	0.0167	0.273	0.0610
	0.25	0.118	0.0167	0.265	0.0109	0.574	0.0338	0.957	0.0166	0.443	0.0648
	0.50	0.145	0.0166	0.234	0.0108	0.568	0.0334	0.947	0.0165	0.620	0.0756
	0.90	0.210	0.0178	0.185	0.0129	0.544	0.0337	0.938	0.0160	1.136	0.1260
All	0.10	0.114	0.0132	0.244	0.0106	0.605	0.0256	0.963	0.0130	0.467	0.0586
	0.25	0.136	0.0130	0.224	0.0090	0.601	0.0256	0.960	0.0131	0.606	0.0629
	0.50	0.156	0.0133	0.203	0.0086	0.598	0.0254	0.956	0.0133	0.767	0.0746
	0.90	0.233	0.0158	0.143	0.0098	0.567	0.0261	0.943	0.0134	1.629	0.1733

Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using LP.

Table 18: LP Coefficient Estimates and Standard Errors for US Manufacturing Firms

NAICS	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
31	0.167	0.0468	0.084	0.0161	0.708	0.0728	0.960	0.0386	1.983	0.8032
32	0.170	0.0234	0.181	0.0151	0.620	0.0396	0.971	0.0221	0.936	0.1643
33	0.134	0.0156	0.248	0.0120	0.572	0.0330	0.954	0.0160	0.541	0.0663
All	0.156	0.0128	0.205	0.0091	0.597	0.0253	0.958	0.0129	0.760	0.0702

Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using LP.

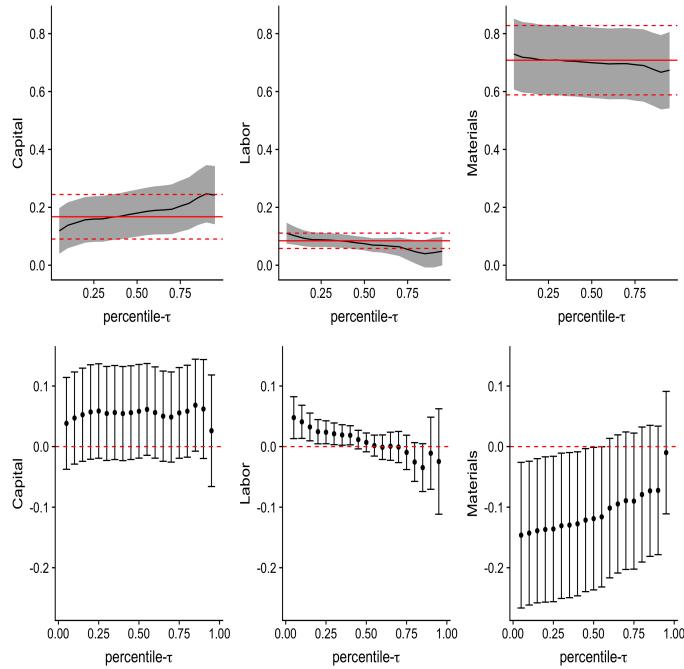


Figure 20: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

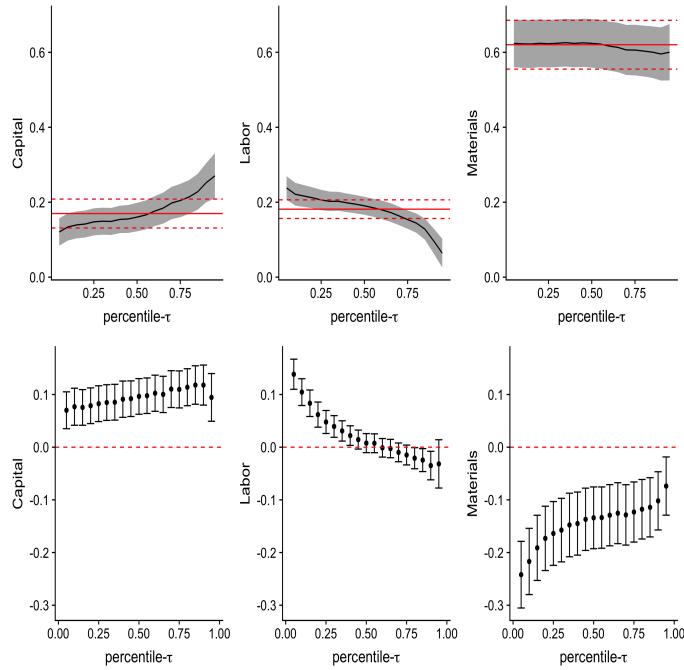


Figure 21: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

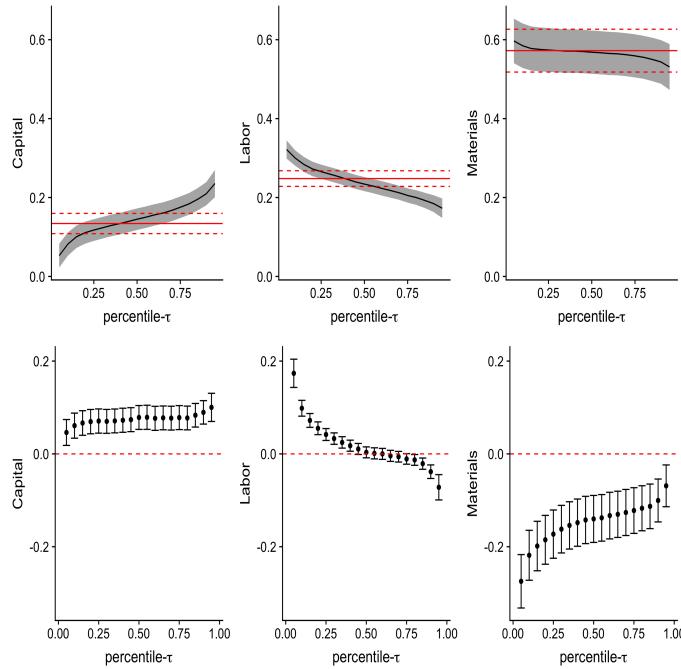


Figure 22: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

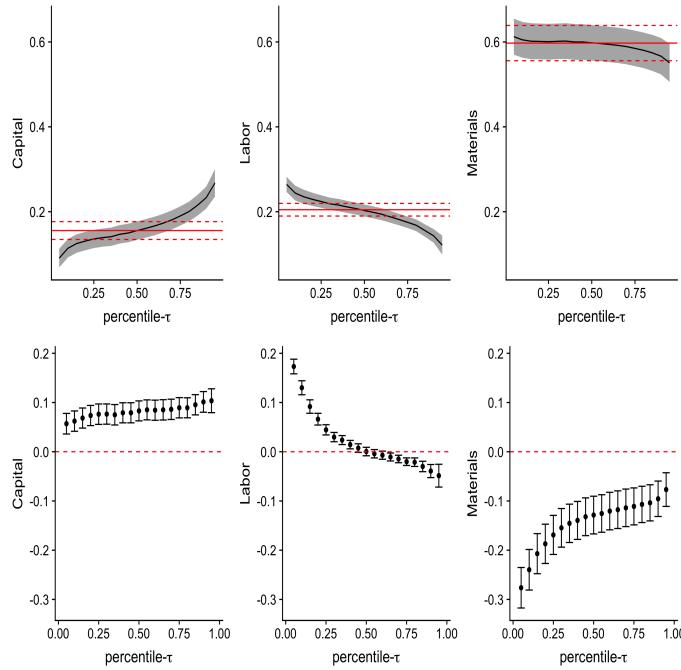


Figure 23: Top row: Estimated values of production function coefficients and their point-wise 90% confidence interval. Bottom row: Difference between QLP and quantile regression estimates and their 95% confidence intervals.

Figure 24: Mean and Median Estimates of Total Factor Productivity

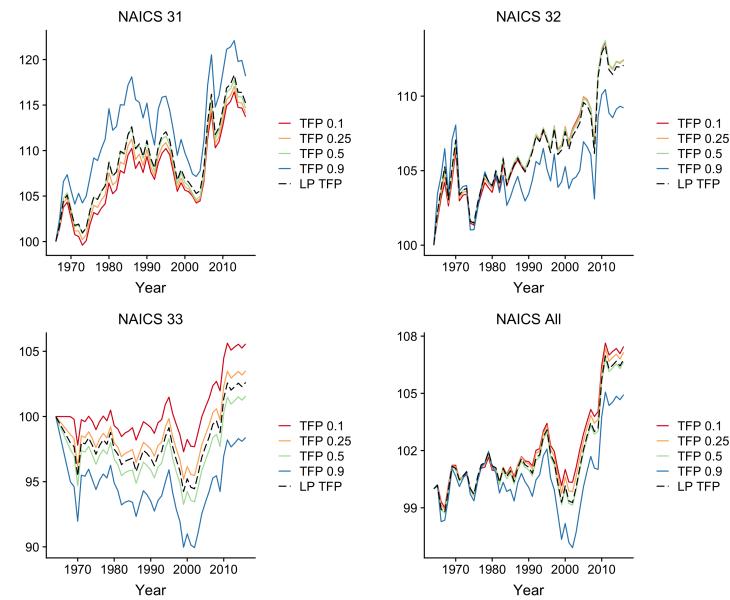
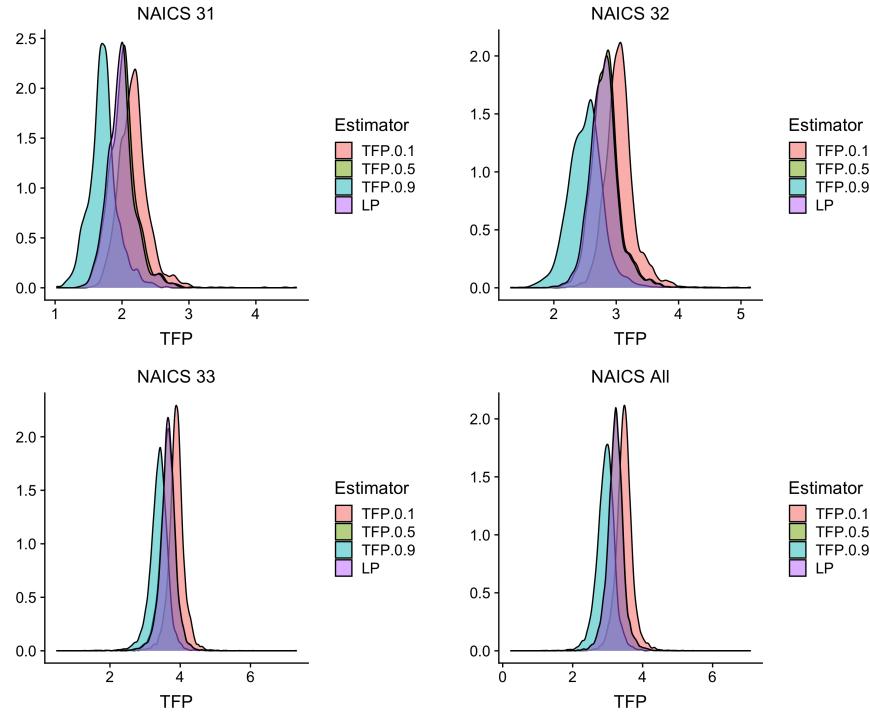


Figure 25: Estimated average productivity over time for the US. Base productivity in 1961 is set to 100.

Table 19: Productivity Differentials for US Manufacturing Firms using DS

NAICS	τ	R&D		Advertising	
		Coef.	s.e	Coef.	s.e
31	0.10	0.041	0.0236	0.049	0.0276
	0.25	0.040	0.0236	0.047	0.0277
	0.50	0.039	0.0238	0.046	0.0279
	0.90	0.030	0.0235	0.035	0.0274
32	0.10	0.023	0.0153	0.028	0.0158
	0.25	0.023	0.0152	0.027	0.0157
	0.50	0.023	0.0155	0.026	0.0159
	0.90	0.036	0.0161	0.035	0.0165
33	0.10	0.038	0.0121	0.044	0.0111
	0.25	0.041	0.0121	0.045	0.0111
	0.50	0.044	0.0121	0.046	0.0111
	0.90	0.045	0.0118	0.044	0.0108
All	0.10	0.034	0.0091	0.042	0.0089
	0.25	0.033	0.0092	0.041	0.0090
	0.50	0.034	0.0093	0.040	0.0091
	0.90	0.035	0.0094	0.039	0.0091

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using the DS-LP estimator

Table 20: Productivity Differentials for US Manufacturing Firms using LP

NAICS	R&D		Advertising	
	Coef.	s.e	Coef.	s.e
31	0.036	0.0234	0.043	0.0274
32	0.024	0.0154	0.027	0.0158
33	0.041	0.0118	0.044	0.0108
All	0.033	0.0091	0.039	0.0089

*Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using the LP estimator

6.2.2 Chile

Table 21: Coefficient Estimates and Standard Errors for Chilean Manufacturing Plants

ISIC	τ	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.026	0.0058	0.109	0.0086	0.762	0.0106	0.898	0.0109	0.237	0.0610
	0.25	0.045	0.0056	0.102	0.0065	0.761	0.0106	0.907	0.0106	0.437	0.0618
	0.50	0.061	0.0059	0.118	0.0063	0.754	0.0105	0.934	0.0107	0.517	0.0592
	0.90	0.081	0.0069	0.153	0.0111	0.756	0.0110	0.991	0.0118	0.528	0.0665
381	0.10	0.053	0.0137	0.326	0.0266	0.589	0.0183	0.968	0.0197	0.163	0.0502
	0.25	0.084	0.0129	0.295	0.0195	0.594	0.0166	0.973	0.0161	0.285	0.0530
	0.50	0.109	0.0131	0.274	0.0167	0.599	0.0158	0.982	0.0156	0.398	0.0611
	0.90	0.132	0.0159	0.285	0.0273	0.615	0.0165	1.032	0.0213	0.462	0.0845
321	0.10	0.022	0.0121	0.272	0.0224	0.620	0.0169	0.914	0.0178	0.082	0.0475
	0.25	0.048	0.0120	0.253	0.0168	0.621	0.0152	0.923	0.0154	0.191	0.0532
	0.50	0.072	0.0121	0.232	0.0152	0.625	0.0144	0.929	0.0150	0.310	0.0646
	0.90	0.108	0.0131	0.202	0.0204	0.639	0.0150	0.949	0.0178	0.536	0.1000
All	0.10	0.038	0.0043	0.187	0.0070	0.672	0.0062	0.898	0.0063	0.205	0.0255
	0.25	0.074	0.0043	0.165	0.0062	0.667	0.0061	0.906	0.0060	0.450	0.0338
	0.50	0.104	0.0043	0.155	0.0057	0.664	0.0059	0.923	0.0057	0.670	0.0409
	0.90	0.148	0.0055	0.174	0.0100	0.663	0.0064	0.984	0.0071	0.852	0.0708

Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using LP.

Table 22: LP Coefficient Estimates and Standard Errors for Chilean Manufacturing Plants

ISIC	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.057	0.0055	0.133	0.0069	0.755	0.0105	0.945	0.0105	0.427	0.0496
381	0.093	0.0130	0.288	0.0187	0.607	0.0154	0.988	0.0157	0.325	0.0574
321	0.067	0.0106	0.234	0.0146	0.629	0.0140	0.930	0.0150	0.286	0.0539
All	0.095	0.0039	0.171	0.0062	0.668	0.0059	0.934	0.0055	0.554	0.0343

*Standard errors are obtained using bootstrap with 500 replications

Figure 26: Estimated Coefficients of Capital and Labor for Chile: ISIC 311

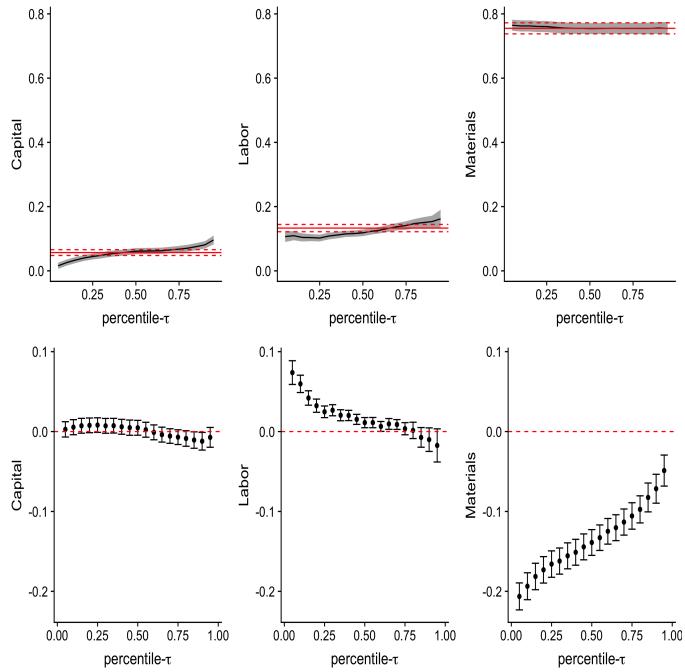


Figure 27: Estimated Coefficients of Capital and Labor for Chile: ISIC 381

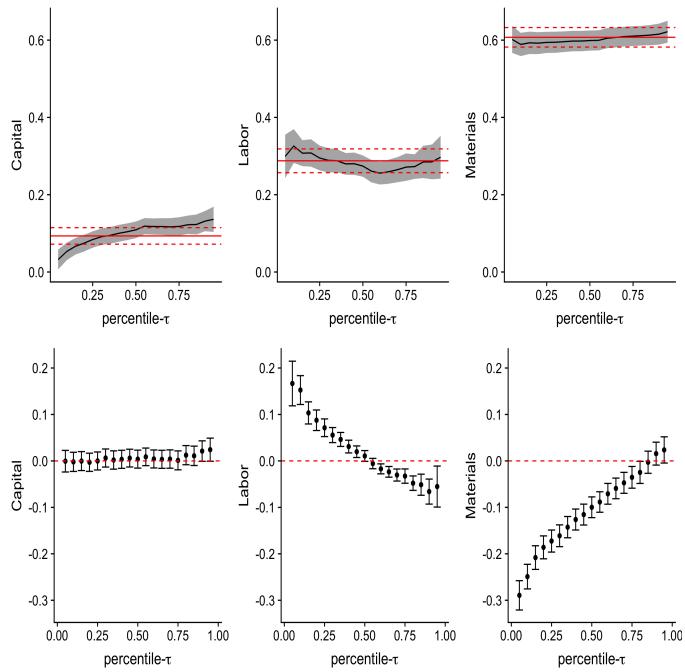


Figure 28: Estimated Coefficients of Capital and Labor for Chile: ISIC 321

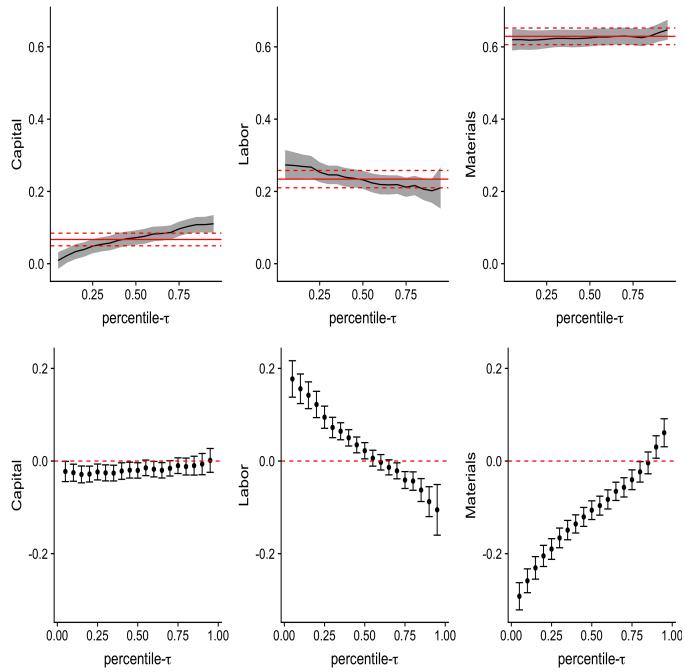


Figure 29: Estimated Coefficients of Capital and Labor for all Chilean Manufacturing Plants

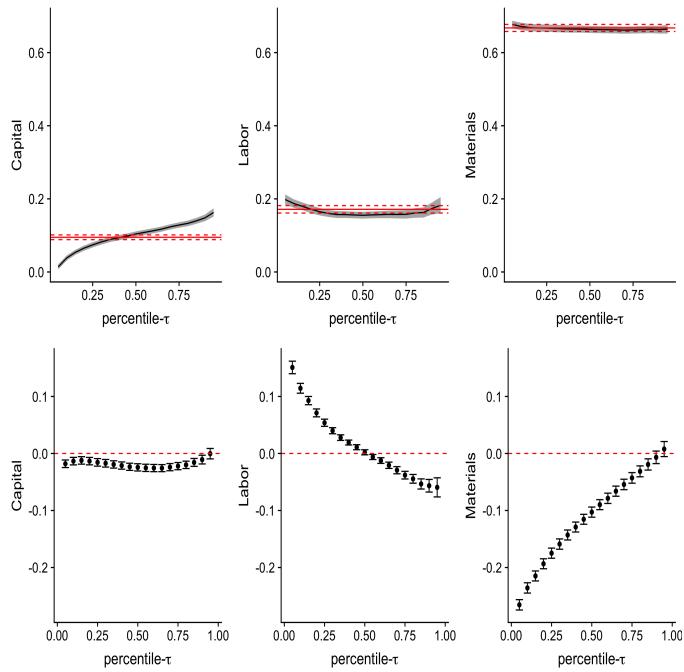


Figure 30: Mean and Median Estimates of Total Factor Productivity

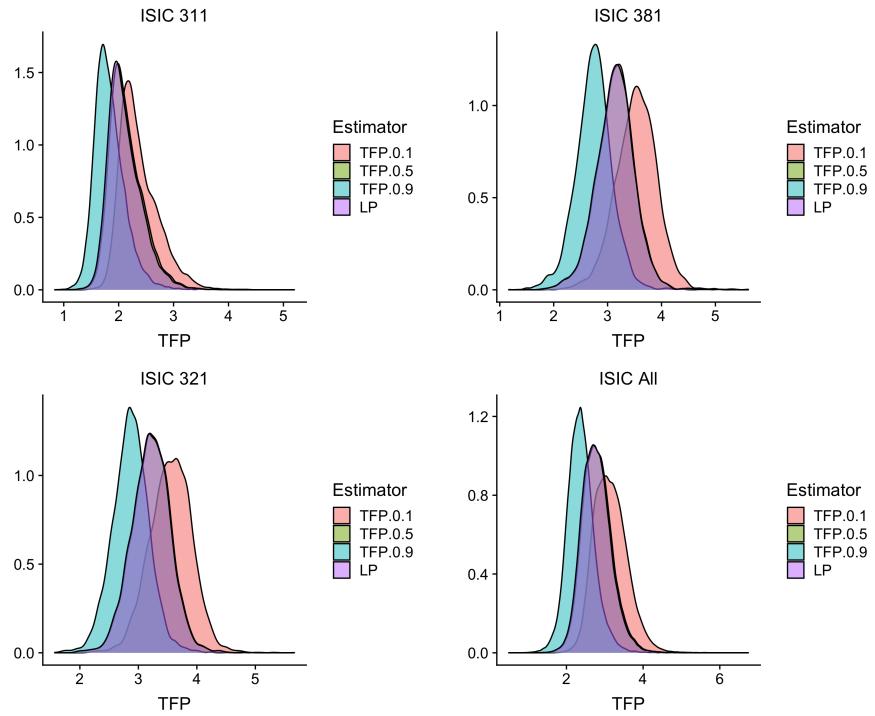
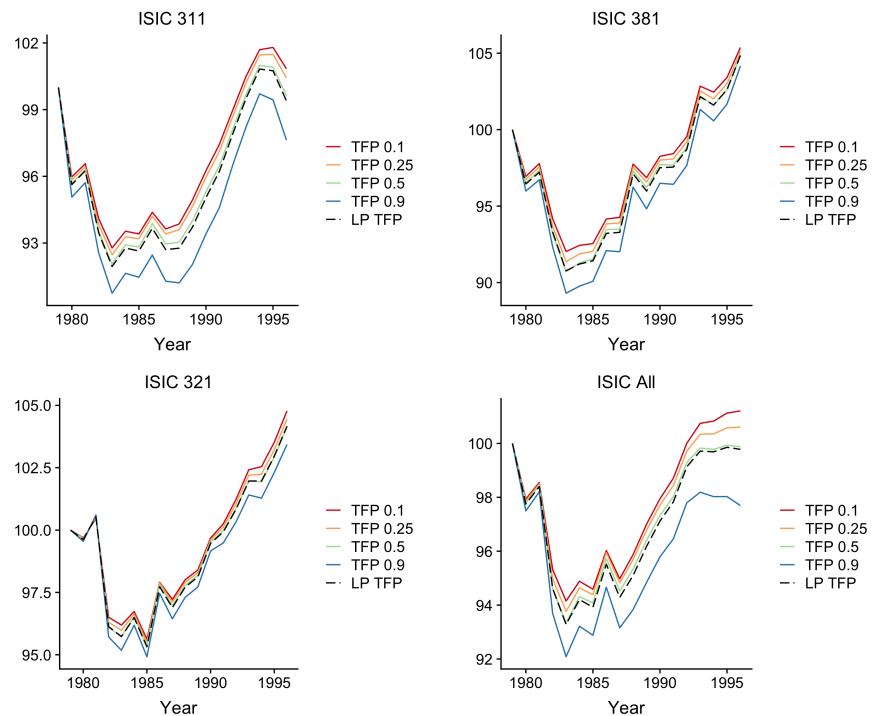


Figure 31: Chile Productivity Over Time



ISIC	τ	Exporter		Importer		Advertiser	
		Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.026	0.0110	0.039	0.0122	0.019	0.0087
	0.25	0.022	0.0108	0.037	0.0120	0.017	0.0087
	0.50	0.018	0.0107	0.028	0.0116	0.012	0.0087
	0.90	0.012	0.0111	0.009	0.0120	0.002	0.0088
381	0.10	0.024	0.0125	0.049	0.0176	0.037	0.0154
	0.25	0.016	0.0123	0.037	0.0179	0.030	0.0149
	0.50	0.008	0.0125	0.026	0.0185	0.024	0.0149
	0.90	-0.011	0.0136	-0.003	0.0196	0.006	0.0165
321	0.10	0.022	0.0111	0.031	0.0145	0.043	0.0133
	0.25	0.019	0.0107	0.027	0.0140	0.040	0.0122
	0.50	0.018	0.0107	0.023	0.0142	0.038	0.0122
	0.90	0.012	0.0112	0.012	0.0152	0.031	0.0130
All	0.10	0.038	0.0047	0.061	0.0070	0.053	0.0055
	0.25	0.031	0.0046	0.054	0.0069	0.051	0.0053
	0.50	0.023	0.0046	0.045	0.0068	0.046	0.0052
	0.90	0.002	0.0046	0.017	0.0069	0.030	0.0053

ISIC	Exporter		Importer		Advertiser	
	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.018	0.0108	0.024	0.0117	0.010	0.0086
381	0.008	0.0124	0.026	0.0182	0.023	0.0149
321	0.017	0.0106	0.022	0.0139	0.037	0.0122
All	0.021	0.0046	0.042	0.0068	0.043	0.0052

6.2.3 Colombia

Table 23: Coefficient Estimates and Standard Errors for Colombian Manufacturing Plants

ISIC	τ	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
		Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.049	0.0139	0.199	0.0174	0.648	0.0170	0.896	0.0167	0.244	0.0799
	0.25	0.083	0.0137	0.179	0.0103	0.642	0.0169	0.904	0.0142	0.466	0.0856
	0.50	0.110	0.0143	0.155	0.0078	0.642	0.0173	0.906	0.0139	0.707	0.1035
	0.90	0.198	0.0166	0.144	0.0151	0.612	0.0188	0.954	0.0156	1.373	0.1989
322	0.10	0.112	0.0157	0.394	0.0198	0.410	0.0165	0.916	0.0233	0.285	0.0451
	0.25	0.157	0.0155	0.353	0.0162	0.408	0.0161	0.917	0.0218	0.444	0.0522
	0.50	0.198	0.0167	0.302	0.0155	0.412	0.0163	0.912	0.0216	0.655	0.0705
	0.90	0.270	0.0231	0.256	0.0243	0.415	0.0185	0.942	0.0228	1.053	0.1670
381	0.10	0.070	0.0251	0.339	0.0244	0.378	0.0221	0.787	0.0320	0.206	0.0811
	0.25	0.113	0.0249	0.291	0.0176	0.380	0.0208	0.784	0.0308	0.388	0.0922
	0.50	0.153	0.0249	0.254	0.0170	0.377	0.0203	0.785	0.0312	0.602	0.1137
	0.90	0.261	0.0287	0.218	0.0294	0.348	0.0239	0.827	0.0333	1.197	0.2541
All	0.10	0.072	0.0060	0.309	0.0070	0.468	0.0070	0.849	0.0078	0.234	0.0210
	0.25	0.110	0.0060	0.280	0.0058	0.468	0.0066	0.859	0.0075	0.393	0.0247
	0.50	0.148	0.0061	0.256	0.0048	0.466	0.0066	0.870	0.0071	0.578	0.0285
	0.90	0.238	0.0072	0.257	0.0090	0.442	0.0073	0.937	0.0082	0.927	0.0492

Standard errors are obtained using bootstrap with 500 replications. Productivity is estimated using LP.

Table 24: LP Coefficient Estimates and Standard Errors for Colombian Manufacturing Plants

ISIC	Capital		Labor		Materials		Returns to Scale		Capital Intensity	
	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.121	0.0135	0.166	0.0109	0.633	0.0178	0.919	0.0142	0.730	0.0960
322	0.189	0.0166	0.305	0.0175	0.420	0.0162	0.914	0.0219	0.621	0.0717
381	0.162	0.0256	0.244	0.0212	0.378	0.0205	0.784	0.0308	0.663	0.1385
All	0.153	0.0059	0.266	0.0069	0.461	0.0068	0.881	0.0077	0.576	0.0294

*Standard errors are obtained using bootstrap with 500 replications

Figure 32: Estimated Coefficients of Capital and Labor for Colombia: ISIC 311

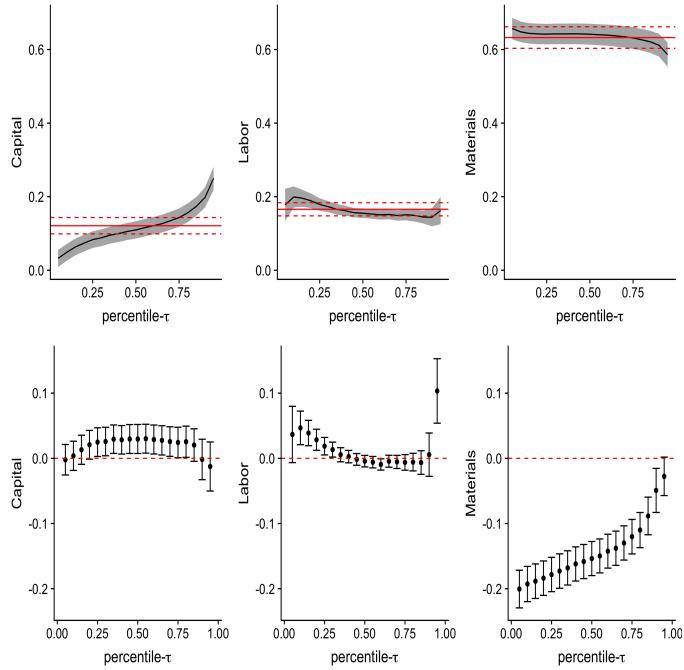


Figure 33: Estimated Coefficients of Capital and Labor for Colombia: ISIC 321

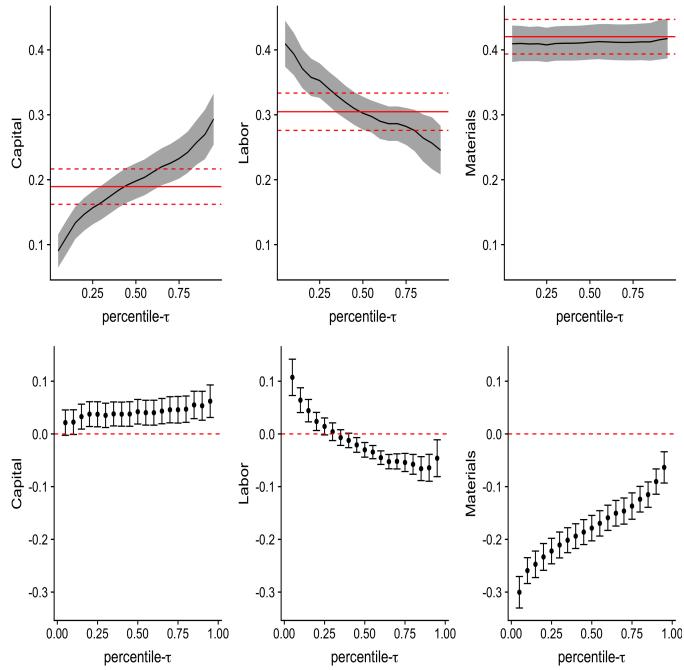


Figure 34: Estimated Coefficients of Capital and Labor for Colombia: ISIC 381

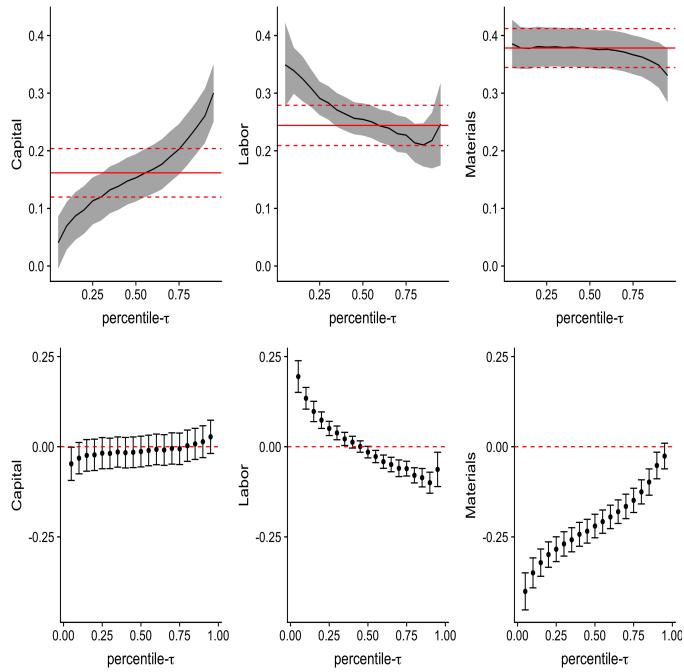


Figure 35: Estimated Coefficients of Capital and Labor for all Colombian Manufacturing Plants

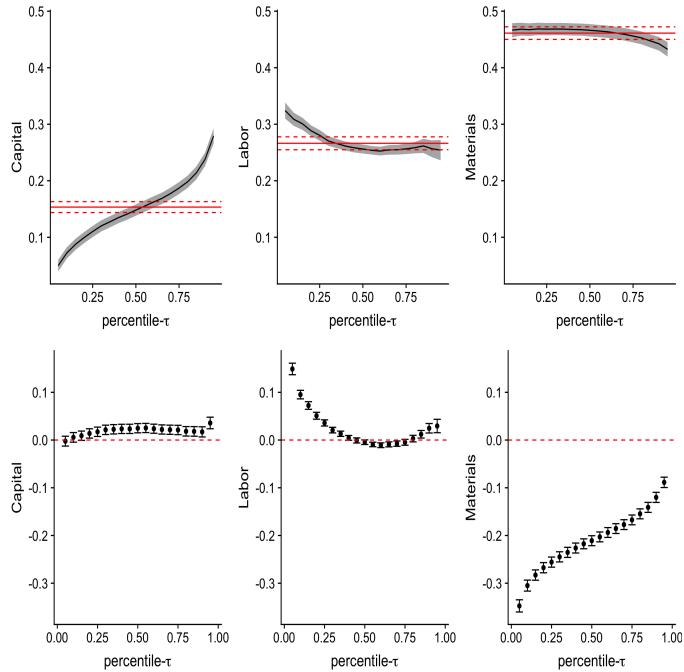


Figure 36: Mean and Median Estimates of Total Factor Productivity

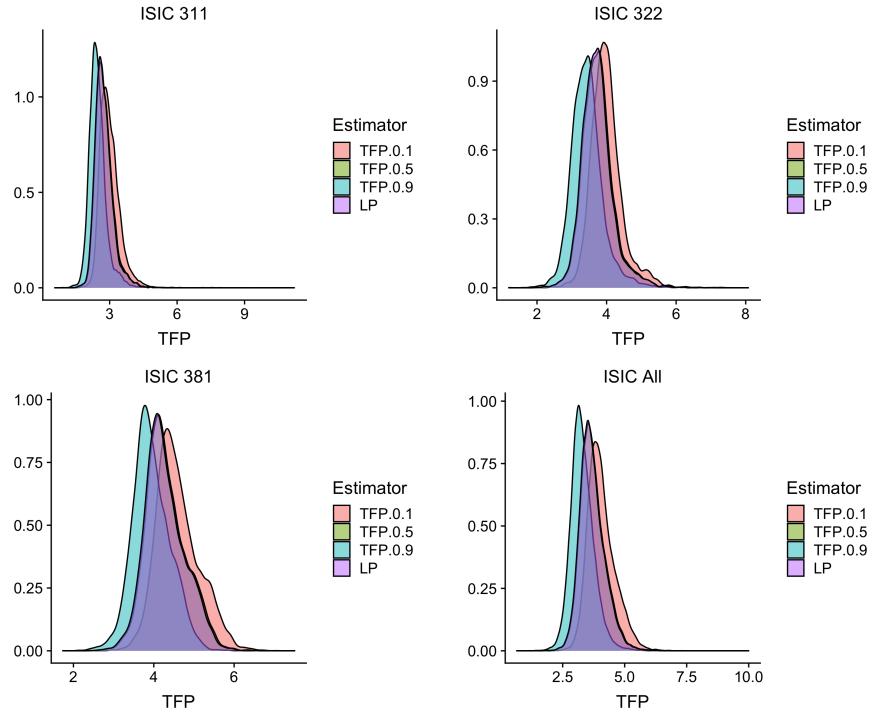
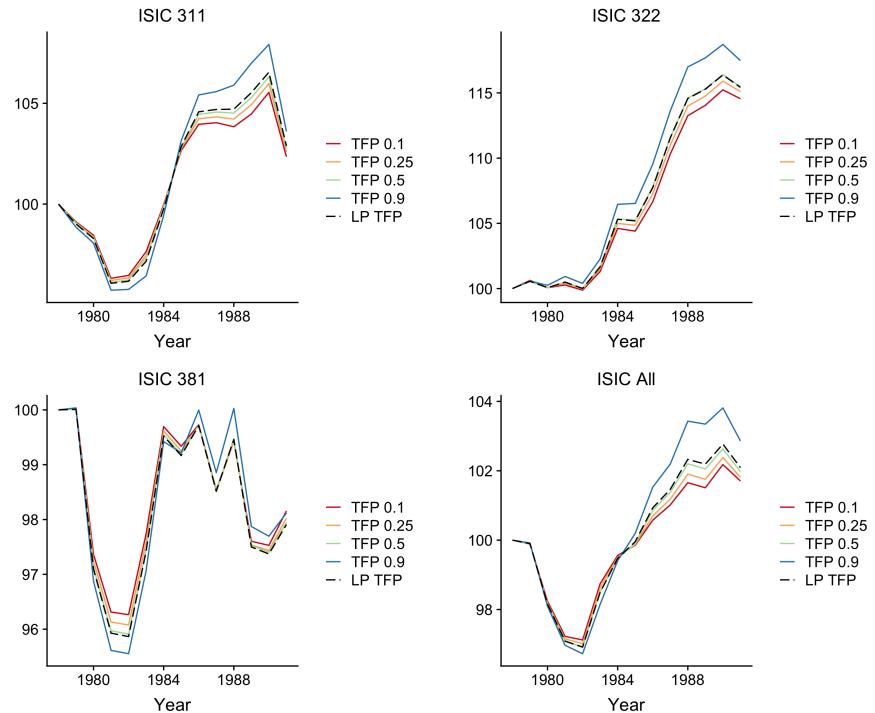


Figure 37: Chile Productivity Over Time



ISIC	τ	Exporter		Importer		Advertiser	
		Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.10	0.041	0.0144	0.087	0.0167	0.064	0.0144
	0.25	0.036	0.0139	0.084	0.0152	0.063	0.0134
	0.50	0.032	0.0137	0.082	0.0141	0.063	0.0126
	0.90	0.015	0.0135	0.069	0.0124	0.056	0.0113
322	0.10	0.049	0.0126	0.092	0.0147	0.073	0.0117
	0.25	0.046	0.0118	0.086	0.0141	0.066	0.0115
	0.50	0.044	0.0114	0.079	0.0140	0.060	0.0116
	0.90	0.034	0.0106	0.060	0.0142	0.043	0.0120
381	0.10	0.118	0.0226	0.155	0.0246	0.142	0.0227
	0.25	0.114	0.0221	0.151	0.0238	0.137	0.0220
	0.50	0.109	0.0219	0.146	0.0233	0.131	0.0216
	0.90	0.089	0.0210	0.126	0.0216	0.108	0.0202
All	0.10	0.074	0.0059	0.129	0.0058	0.101	0.0049
	0.25	0.068	0.0056	0.120	0.0056	0.094	0.0047
	0.50	0.061	0.0053	0.111	0.0055	0.088	0.0046
	0.90	0.042	0.0048	0.086	0.0055	0.068	0.0044

ISIC	Exporter		Importer		Advertiser	
	Coef.	s.e	Coef.	s.e	Coef.	s.e
311	0.029	0.0135	0.079	0.0139	0.061	0.0124
322	0.043	0.0114	0.078	0.0139	0.060	0.0115
381	0.108	0.0216	0.145	0.0230	0.130	0.0212
All	0.060	0.0053	0.109	0.0055	0.086	0.0046

6.3 Ex-Post Shocks and Measurement Error in Output

6.4 Identification of Gross-Output Production Functions

6.5 Proofs