Political Methodology III: Model Based Inference

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Model Based Inference

1) Likelihood inference
2) Logit/Probit
    a) DGP
    b) Optimization
    c) Quantities of Interest
    d) Perfect + Near Perfect Separation
3) Ordered Probit
    a) DGP
    b) Optimization
    c) Quantities of Interest

Civil Conflict and Political Instability (h/t: Yamamoto)

Fearon & Laitin (2003):

- Y_i: Civil conflict
- T_i: Political instability
- W_i: Geography (log % mountainous)

Estimated model:

$$\widehat{\Pr(Y_i = 1 \mid T_i, W_i)} = \text{logit}^{-1}(-2.84 + 0.91T_i + 0.35W_i)$$

Predicted probability:

$$\hat{\pi}(T_i = 1, W_i = 3.10) = 0.299$$

ATE:

$$\hat{\tau} = \frac{1}{n}\sum_{i=1}^{n}\{\hat{\pi}(1, W_i) - \hat{\pi}(0, W_i)\} = 0.127$$

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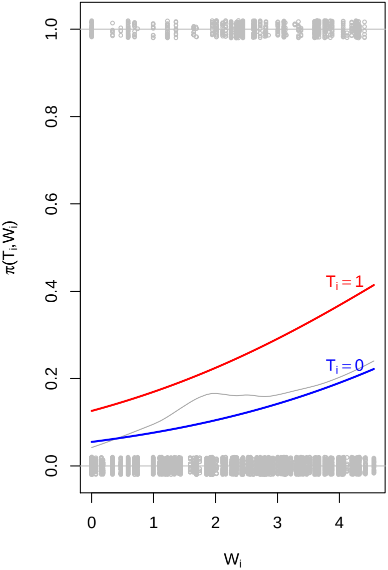

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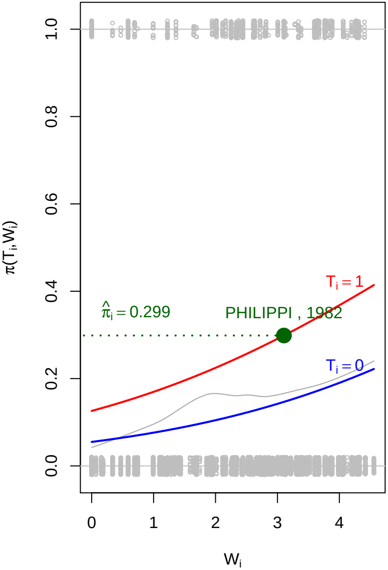

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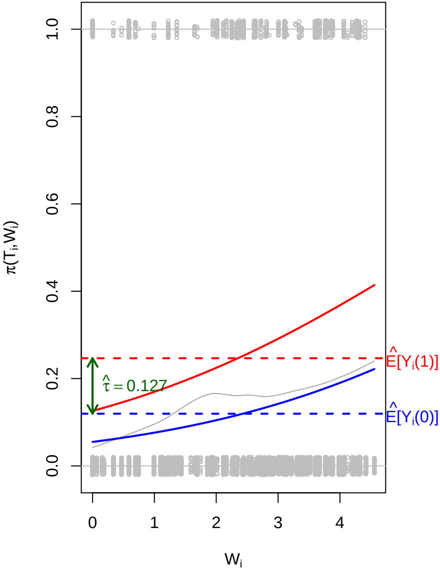

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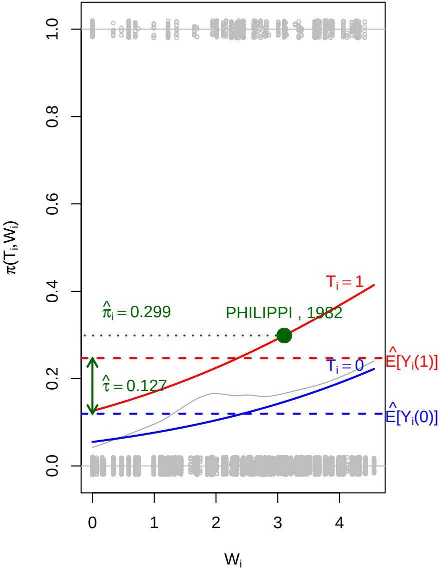

Final Passage Vote for the Affordable Care Act

$Y_i =$ Vote on ACA

$X_{i1} =$ Democrat

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	Nay	Yea
Republican	178	0
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$$\Pr(Y_i = 1 | X_{i1} = 0) \quad = \quad 0$$

Final Passage Vote for the Affordable Care Act

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> test_model<- glm(vote~dem + ideo,
data = as.data.frame(final_votes), family = binomial('logit'))
Warning message:
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$$E[Y_i|X_i = \widehat{(1,0,0.7)}, \boldsymbol{\beta}^*] \quad = \quad 0$$

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We have problems!

Is There Anything the Tea Party Doesn't Screw Up?

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$$L(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) = f(\boldsymbol{Y}|\boldsymbol{X}, \boldsymbol{\beta})$$

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L(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) &= f(\boldsymbol{Y}|\boldsymbol{X}, \boldsymbol{\beta}) \\
&= \sum_{i=1}^{N} Y_i \log\left(\frac{1}{1+\exp(-\boldsymbol{X}^{'}\boldsymbol{\beta})}\right) + {\color{red}(1-Y_i)\log\left(1 - \frac{1}{1+\exp(-\boldsymbol{X}^{'}\boldsymbol{\beta})}\right)}
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Remember:

Is There Anything the Tea Party Doesn't Screw Up?

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Remember:

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Remember:

$$
\Pr(Y_i = 1|X_{i1} = 0) = 0
$$

To fit data: set $\beta_0 \rightarrow -\infty$.

Perfect and Near-Perfect Separation

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Perfect separation: one covariate perfectly separates 0's and 1's

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Solution?:

Perfect and Near-Perfect Separation

Perfect separation: one covariate perfectly separates 0's and 1's
Near perfect separation: one covariate perfectly separates 0's or 1's
Solution?:
You need to make more assumptions

Add a Few Observations...

	Nay	Yea
Republican	178	0
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Add a Few Observations...

	Nay	Yea
Republican	178.5	0.5
Democrat	34.5	219.5

Penalized (Prior)-Logistic Regression

- Separation: causes coefficients to diverge
- Penalty (prior): force coefficients towards zero

Step 1: Standardize inputs (Gelman et al)

- Binary variables: mean 0, differ by 1.
    - Democrats: (30%). (0.3, -0.7)
- Other variables: mean 0, sd 0.5.

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Step 2: Penalize Likelihood

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$$L(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) = \prod_{i=1}^{N} \pi_i^{Y_i}(1-\pi_i)^{1-Y_i}|I(\boldsymbol{\beta})|^{1/2}$$

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where:

$$\begin{aligned}
\pi_i &= \frac{1}{1+\exp(-\boldsymbol{X}_i^{'}\boldsymbol{\beta})} \\
|I(\boldsymbol{\beta})| &= \text{Determinant of Fisher's information at } \boldsymbol{\beta} \\
I(\boldsymbol{\beta}) &= \boldsymbol{X}^{'}\boldsymbol{W}\boldsymbol{X} \\
\boldsymbol{W} &= \begin{pmatrix} \pi_1(1-\pi_1) & 0 & \ldots & 0 \\ 0 & \pi_2(1-\pi_2) & \ldots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \ldots & \pi_N(1-\pi_N) \end{pmatrix}
\end{aligned}$$

Penalized (Prior)-Logistic Regression

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L(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) &= \prod_{i=1}^{N} \pi_i^{Y_i}(1-\pi_i)^{1-Y_i}|I(\boldsymbol{\beta})|^{1/2} \\
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\end{aligned}
$$

Penalized (Prior)-Logistic Regression

```
jef_pri<- function(params, X, Y){
    beta<- params
    y.tilde<- X%*%beta
    y.prob<- plogis(y.tilde)
    temp<- matrix(0, nrow = length(Y), ncol=length(Y))
    part1<- Y%*%log(y.prob) + (1-Y)%*%log(1- y.prob)
    diag(temp)<- y.prob*(1-y.prob)
    part2<- 0.5*log(det(t(X)%*%temp%*%X))
    out<- part1 + part2
}
firth<- optim(rnorm(3), jef_pri, method = 'BFGS',
control=list(fnscale=-1), hessian=T,
X = cbind(1, dem, ideo), Y =clean[,3] )
```

Comparison

	ACA Vote (GLM)	Firth
Intercept	-14.00	-5.70
	(1670.439)	(24.68)
Democrat	11.67	3.30
	(1670.439)	(42.10)
Ideology	-16.86	-17.72
	(2.71)	(2.85)

Penalized (Prior) Logistic Regression

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- Suggestion Cauchy (DOF = 1) with scale 2.5.
  - Prioritize $\beta < 0.5$
  - Occasionally allows very large values (Cauchy)

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lst<- function(x, nu, mu, sigma2){
    part1<- lgamma( (nu + 1)/2)
    part2<- lgamma(nu/2)
    part3<- sqrt(pi *nu*sqrt(sigma2))
    part4<- 1 + (1/nu)*(( (x- mu)^2)/sigma2)
    part4<- ( - (nu + 1)/2)*log(part4)
    out<- part4
    return(out)
}
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```
log_t<- function(params, X, Y, nu, mu, sigma2){
beta<- params
prior<- 0
for(k in 2:ncol(X)){
prior<- prior + lst(beta[k], nu, mu, sigma2)
}
prior<- prior + lst(beta[1], 1, 0, 10)
y.tilde<- X%*%beta
y.prob<- plogis(y.tilde)
out<- Y%*%log(y.prob) + (1- Y)%*%log(1- y.prob)
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In practice: `bayesglm` in `library(arm)` is awesome!

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	ACA Vote (GLM)	Firth	Cauchy
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Democrat	11.67	3.30	-0.19
	(1670.439)	(42.10)	(1.12)
Ideology	-16.86	-17.72	-16.25
	(2.71)	(2.85)	(2.65)

Ordered Outcome Data

Modeling Ordered Outcomes (ht: Yamamoto)

- Suppose that the J choices are ordered in a substantively meaningful way
- Examples:
  - "Likert scale" in survey questions ("strongly agree", "agree", etc.)
  - Party positions (extreme left, center left, center, right, extreme right)
  - Levels of democracy (autocracy, anocracy, democracy)
  - Health status (healthy, sick, dying, dead)
- Why not use continuous outcome models?
  $\longrightarrow$ Don't want to assume equal distances between levels
- Why not use categorical outcome models? (on Monday)
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Ordered Logit and Probit Models (ht: Yamamoto)

- Again, the latent variable representation: $Y_i^* \; = \; X_i'\beta + \epsilon_i$
- Assume that Y_i^* gives rise to Y_i based on the following scheme:

$$Y_i \;=\; \begin{cases} 1 & \text{if} \quad -\infty(=\psi_0) \;<\; Y_i^* \;\le\; \psi_1, \\ 2 & \text{if} \quad \psi_1 \;<\; Y_i^* \;\le\; \psi_2, \\ \vdots & \qquad\qquad\qquad \vdots \\ J & \text{if} \quad \psi_{J-1} \;<\; Y_i^* \;\le\; \infty(=\psi_J) \end{cases}$$

  where $\psi_1, ..., \psi_{J-1}$ are the threshold parameters to be estimated
- If X_i contains an intercept, one of the ψ's must be fixed for identifiability (typically $\psi_1 = 0$)
- $\epsilon_j \sim_{\text{iid}}$ logistic $\Rightarrow$ the ordered logit model:

$$\Pr(Y_i \le j \mid X_i) \;=\; \frac{\exp(\psi_j - X_i'\beta)}{1 + \exp(\psi_j - X_i^{\top}\beta)}$$

- $\epsilon_j \sim_{\text{iid}}$ Normal$(0, 1) \Rightarrow$ the ordered probit model:

$$\Pr(Y_i \le j \mid X_i) \;=\; \Phi\left(\psi_j - X_i'\beta\right)$$

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$$\Pr(Y_i \leq j \mid X_i) = \Phi\left(\psi_j - X_i'\beta\right)$$

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- Again, the latent variable representation: $Y_i^* = X_i'\beta + \epsilon_i$
- Assume that Y_i^* gives rise to Y_i based on the following scheme:

$$Y_i = \begin{cases} 1 & \text{if} \quad -\infty(=\psi_0) < Y_i^* \leq \psi_1, \\ 2 & \text{if} \quad \psi_1 < Y_i^* \leq \psi_2, \\ \vdots & \qquad \vdots \\ J & \text{if} \quad \psi_{J-1} < Y_i^* \leq \infty(=\psi_J) \end{cases}$$

where $\psi_1, ..., \psi_{J-1}$ are the threshold parameters to be estimated

- If X_i contains an intercept, one of the ψ's must be fixed for identifiability (typically $\psi_1 = 0$)
- $\epsilon_j \sim_{\text{iid}}$ logistic $\Rightarrow$ the ordered logit model:

$$\Pr(Y_i \leq j \mid X_i) = \frac{\exp(\psi_j - X_i'\beta)}{1 + \exp(\psi_j - X_i^\top \beta)}$$

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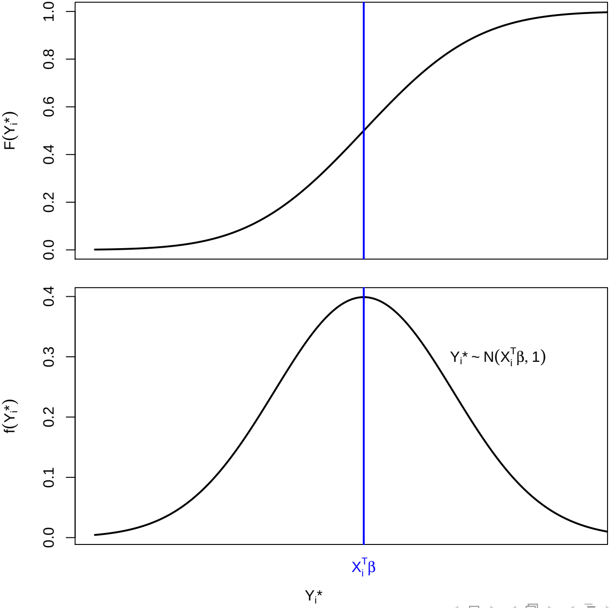

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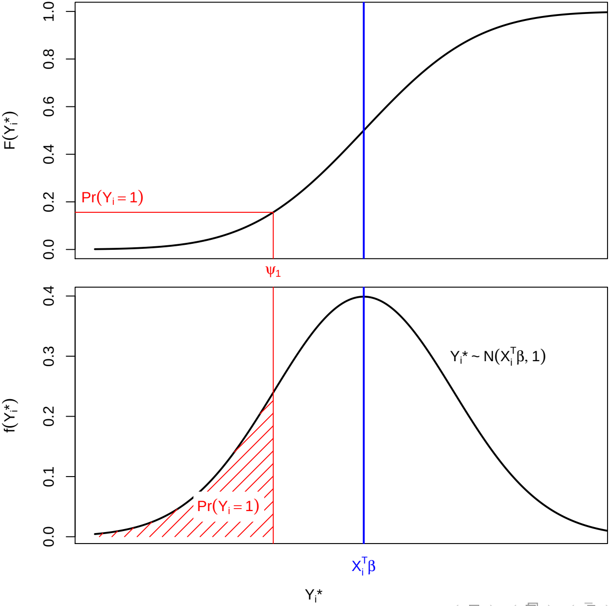

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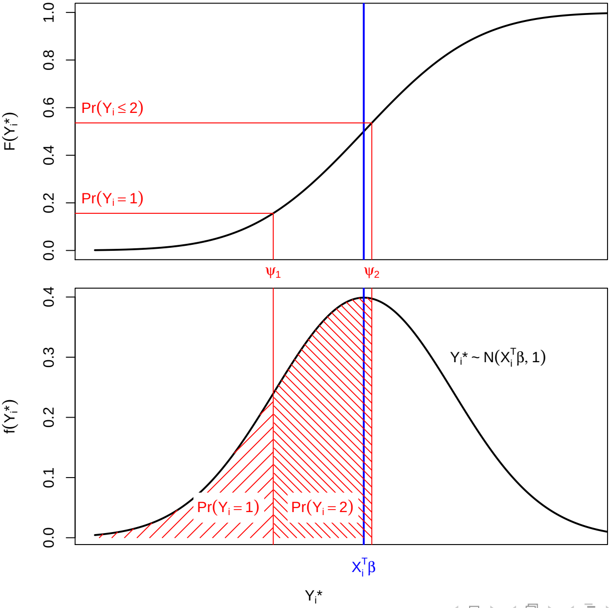

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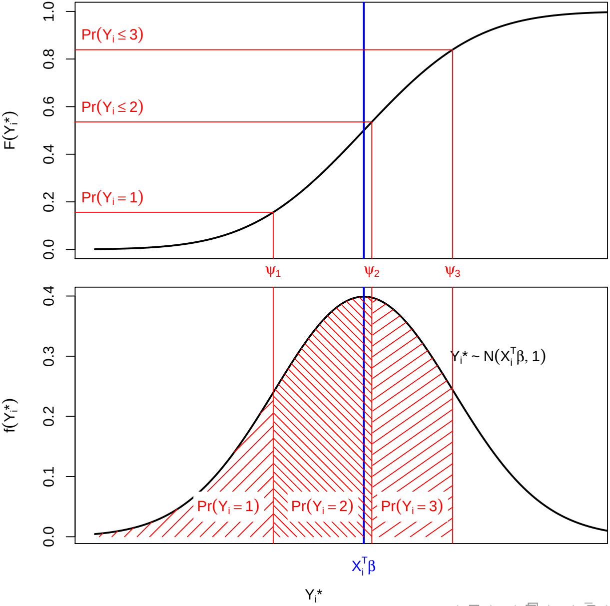

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P(Y_i = J) &= \int_{\psi_{j-1}}^{\psi_j} \phi(\tilde{y}|\boldsymbol{X}_i^{'}\boldsymbol{\beta})d\tilde{y} \\
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fit with `polr` package

Calculating Quantities of Interest (ht: Yamamoto)

- Predicted probability:

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- ATE (APE): $\tau_j = \mathrm{E}\left[\pi_j(T_i = 1, W_i) - \pi_j(T_i = 0, W_i)\right]$

- Estimate β and ψ via MLE, plug the estimates in, replace E with $\frac{1}{n}\sum$, and compute CI by delta or MC or bootstrap

- Note that $X_i'\beta$ appears both before and after the minus sign in π_{ij}
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Immigration and Media Priming(ht: Yamamoto)

Brader, Valentino and Suhay (2008):

- Y_i: Ordinary response to question about increasing immigration
- T_{1i}, T_{2i}: Media cues (immigrant ethnicity $\times$ story tone)
- W_i: Respondent age and income

Estimated coefficients:

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Coefficients:
          Value  s.e.   t
tone       0.27  0.32  0.85
eth       -0.33  0.32 -1.02
ppage      0.01  0.02  1.40
ppincimp   0.00  0.03  0.06
tone:eth   0.90  0.46  2.16

Intercepts:
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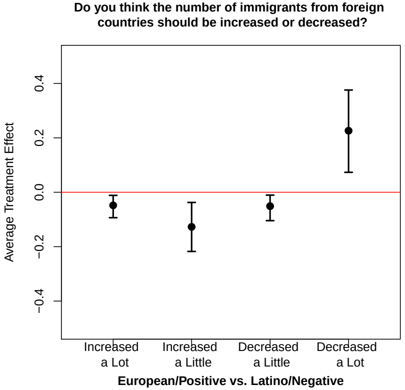
ATE:



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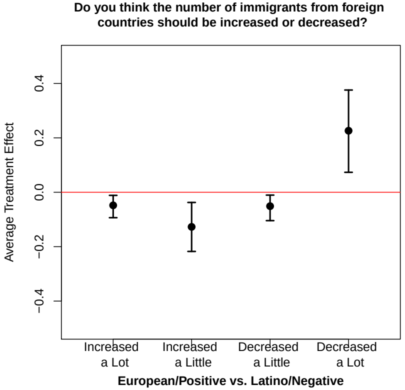

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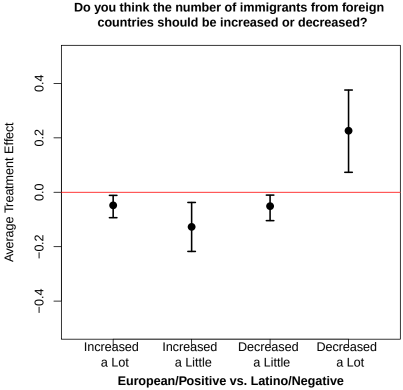

Model Based Inference

1) Likelihood inference
2) Logit/Probit
    a) DGP
    b) Optimization
    c) Quantities of Interest
    d) Perfect + Near Perfect Separation
3) Ordered Probit
    a) DGP
    b) Optimization
    c) Quantities of Interest
4) Choice Models: Multinomial Logit/Probit