Political Methodology III: Model Based Inference

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Going Public (Franco, Grimmer, and Whang 2017)

- Presidents interrupt prime time coverage → why? what effect?
- There are at least 6 explanations (Canes Wrone 2001, 2006) (others: credit claim, veto bargain, beauty contest, highlight obstruction, stupid)
- Gathered serendipitous surveys → happen to be in field when presidents go public (before → control; after → treatment) (we also have social media data and newspapers)
- Does going public increase probability respondents identify topic of president's speech as salient problem?
- Do respondents identify it as most important problem?

Modeling Bivariate Responses

$$Y_i \in \{0,1\}$$
 $X_{i1} = \text{Treatment status (0/1)}$ $X_{i2} = \text{Republican (0/1)}$

Infer effect of going public, condition on Republican as well

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Infer effect of going public, condition on Republican as well

	Control	Treat
\bar{Y}	0.373	0.367

Linear Probabilty Model

$$Y_i \sim \mathsf{Normal}(\mu_i, \sigma^2)$$

 $E[Y_i|X_i, \boldsymbol{\beta}] = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2}$

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Issues:

- 1) Predictions outside of 0 and 1
- 2) Constant effect
- 3) Efficiency loss

$$\tilde{Y}_i \sim \mathsf{Normal}(\mu_i, 1)$$

$$\begin{split} \tilde{Y}_i & \sim & \mathsf{Normal}(\mu_i, 1) \\ \tilde{Y}_i & = & \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i \end{split}$$

$$\begin{split} \tilde{Y}_i & \sim & \mathsf{Normal}(\mu_i, 1) \\ \tilde{Y}_i & = & \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i \\ \epsilon_i & \sim & \mathsf{Normal}(0, \textcolor{red}{\textbf{1}}) \end{split}$$

$$\begin{array}{lcl} \tilde{Y}_i & \sim & \mathsf{Normal}(\mu_i, 1) \\ \tilde{Y}_i & = & \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i \\ \epsilon_i & \sim & \mathsf{Normal}(0, \mathbf{1}) \\ Y_i & = & I(\tilde{Y}_i > 0) \end{array}$$

Suppose $Y_i \sim \mathsf{Bernoulli}(\pi_i)$ Assume:

$$\begin{array}{lcl} \tilde{Y}_i & \sim & \mathsf{Normal}(\mu_i, 1) \\ \tilde{Y}_i & = & \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i \\ \epsilon_i & \sim & \mathsf{Normal}(0, \mathbf{1}) \\ Y_i & = & I(\tilde{Y}_i > 0) \end{array}$$

We will write:

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We will write:

$$X_i = (1, X_{i1}, X_{i2})$$

 $\beta = (\beta_0, \beta_1, \beta_2)$
 $X_i\beta = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2}$

$$\pi_i = P(Y_i = 1) = P(\tilde{Y}_i > 0)$$

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= $P(\beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i > 0)$

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= $P(\beta_{0} + \beta_{1}X_{i1} + \beta_{2}X_{i2} + \epsilon_{i} > 0)$
= $P(\epsilon_{i} > -X'_{i}\beta)$

$$\pi_{i} = P(Y_{i} = 1) = P(\tilde{Y}_{i} > 0)$$

$$= P(\beta_{0} + \beta_{1}X_{i1} + \beta_{2}X_{i2} + \epsilon_{i} > 0)$$

$$= P(\epsilon_{i} > -\mathbf{X}_{i}'\boldsymbol{\beta})$$

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$$= P(\epsilon_{i} < \mathbf{X}_{i}'\boldsymbol{\beta})$$

$$= \int_{-\infty}^{\mathbf{X}_{i}'\boldsymbol{\beta}} \phi(\epsilon_{i})d\epsilon_{i}$$

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$$= \Phi(X_{i}'\beta)$$

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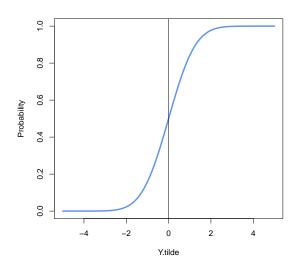
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$$Y_i \sim \mathsf{Bernoulli}(\pi_i)$$

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$$L(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = f(\boldsymbol{Y}|\boldsymbol{X},\boldsymbol{\beta})$$

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$$L(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = f(\boldsymbol{Y}|\boldsymbol{X},\boldsymbol{\beta})$$
$$= \prod_{i=1}^{N} f(Y_i|\boldsymbol{X}_i\boldsymbol{\beta})$$

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$$= \prod_{i=1}^{N} f(Y_i|\boldsymbol{X}_i\boldsymbol{\beta})$$

$$= \prod_{i=1}^{N} \pi_i^{Y_i} (1 - \pi_i)^{1 - Y_i}$$

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Maximize likelihood with respect to $oldsymbol{eta}$

$$l(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = \sum_{i=1}^{N} \left(Y_{i} \log \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta}) + (1 - Y_{i}) \log(1 - \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})) \right)$$

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Score Function for β_k :

$$\frac{\partial l(\boldsymbol{\beta}|\boldsymbol{X}_{i},Y_{i})}{\partial \beta_{k}} = \left(Y_{i} \frac{\phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})}{\Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})} - (1 - Y_{i}) \frac{\phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})}{1 - \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})}\right) X_{k}$$

$$l(\boldsymbol{\beta}|\boldsymbol{X}, \boldsymbol{Y}) = \sum_{i=1}^{N} \left(Y_{i} \log \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta}) + (1 - Y_{i}) \log(1 - \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})) \right)$$

Score Function for β_k :

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Hessian (typical entry h_{kj}):

$$\frac{\partial^{2}l(\boldsymbol{\beta}|\boldsymbol{X}_{i},Y_{i})}{\partial\beta_{k}\partial\beta_{j}} = \phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})[Y_{i}\frac{\phi(\boldsymbol{X}_{i}'\boldsymbol{\beta}) + \boldsymbol{X}_{i}\boldsymbol{\beta}\Phi(\boldsymbol{X}_{i}\boldsymbol{\beta})}{\Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta})^{2}} + (1 - Y_{i})\frac{\phi(\boldsymbol{X}_{i}\boldsymbol{\beta}) - \boldsymbol{X}_{i}'\boldsymbol{\beta}(1 - \Phi(\boldsymbol{X}_{i}'\boldsymbol{\beta}))}{(1 - \Phi(\boldsymbol{X}_{i}\boldsymbol{\beta}))^{2}}]X_{ik}X_{ij}$$

Maximum Likelihood Estimates

$$0 = \sum_{i=1}^{N} \left(Y_i \frac{\phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)}{\Phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)} - (1 - Y_i) \frac{\phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)}{1 - \Phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)} \right) X_k$$

Solve for β ?

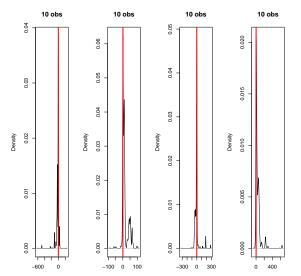
Probit Model

Maximum Likelihood Estimates

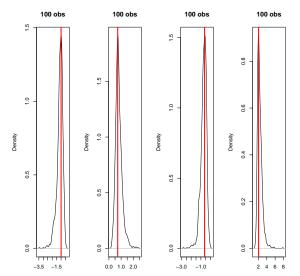
$$0 = \sum_{i=1}^{N} \left(Y_i \frac{\phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)}{\Phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)} - (1 - Y_i) \frac{\phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)}{1 - \Phi(\boldsymbol{X}_i' \boldsymbol{\beta}^*)} \right) X_k$$

Solve for β ? Computational Approaches

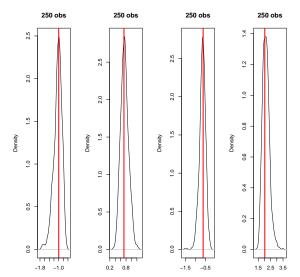
- Newton Raphson
- BFGS (Quasi-Newton. Requires only an approximate Hessian.)



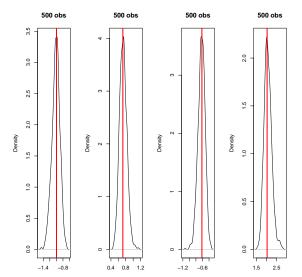
N = 1000 Bandwidth = 4.7: N = 1000 Bandwidth = 2.2! N = 1000 Bandwidth = 3.6: N = 1000 Bandwidth = 9.1



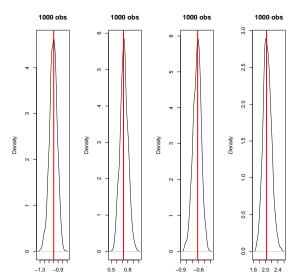
N = 1000 Bandwidth = 0.06: N = 1000 Bandwidth = 0.05: N = 1000 Bandwidth = 0.05: N = 1000 Bandwidth = 0.1



N = 1000 Bandwidth = 0.031 N = 1000 Bandwidth = 0.03; N = 1000 Bandwidth = 0.03; N = 1000 Bandwidth = 0.06;



N = 1000 Bandwidth = 0.02! N = 1000 Bandwidth = 0.02: N = 1000 Bandwidth = 0.02. N = 1000 Bandwidth = 0.04



 $N = 1000 \quad Bandwidth = 0.01i \quad N = 1000 \quad Bandwidth = 0.01! \quad N = 1000 \quad Bandwidth = 0.01! \quad N = 1000 \quad Bandwidth = 0.03! \quad N = 1000 \quad Bandwidth = 0.03! \quad N = 1000 \quad Bandwidth = 0.03! \quad N = 1000 \quad Bandwidth = 0.01! \quad N = 1000 \quad Bandwidth = 0.01!$

Probit Model: Maximum Likelihood Estimates

R Code

Probit Model: Most Important Problem

	Most Important
	•
	Problem
Intercept	-0.26
	(0.03)
Post-Speech	-0.01
	(0.05)
Republican	-0.20
	(0.05)

How do we interpret the coefficients?

Probit Model: Most Important Problem

	Most Important
	Problem .
Intercept	-0.26
	(0.03)
Post-Speech	-0.01
	(0.05)
Republican	-0.20
	(0.05)

How do we interpret the coefficients?

They are on the latent scale → Need to define all values when determining predicted probabilities

$$E[Y_i|X_i] = \Phi(\beta_0^* + \beta_1^*X_{i1} + \beta_2^*X_{i2})$$

$$E[Y_i|\mathbf{X}_i] = \Phi(\beta_0^* + \beta_1^* X_{i1} + \beta_2^* X_{i2})$$

$$E[Y_i|X_{i1} = 0, X_{i2} = 0] = \Phi(-0.26 - 0.01 \times 0 - 0.20 \times 0) = 0.40$$

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$$E[Y_i|X_{i1} = 0, X_{i2} = 1] = \Phi(-0.26 - 0.01 \times 0 - 0.20 \times 1) = 0.32$$

Expected Value:

$$E[Y_i|X_i] = \Phi(\beta_0^* + \beta_1^*X_{i1} + \beta_2^*X_{i2})$$

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First difference:

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First difference:

$$E[Y_i|X_{i1} = 1, X_{i2} = 0] - E[Y_i|X_{i1} = 0, X_{i2} = 0] = \Phi(-0.26 - 0.01) - \Phi(-0.26) = -0.0039$$

Expected Value:

$$E[Y_i|X_i] = \Phi(\beta_0^* + \beta_1^*X_{i1} + \beta_2^*X_{i2})$$

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$$E[Y_i|X_{i1} = 1, X_{i2} = 1] - E[Y_i|X_{i1} = 0, X_{i2} = 1] = \Phi(-0.26 - 0.01 - 0.2) - \Phi(-0.26 - 0.2) = -0.0036$$

Expected Value:

$$E[Y_i|X_i] = \Phi(\beta_0^* + \beta_1^*X_{i1} + \beta_2^*X_{i2})$$

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We will perform inference on these next lecture

$$\begin{array}{lcl} \frac{\partial \Phi(\boldsymbol{X}_i \boldsymbol{\beta})}{\partial X_{ij}} & = & \phi(\boldsymbol{X}_i \boldsymbol{\beta}) \beta_j \\ \text{Max Effect} & = & 0.4 \times \beta_j \end{array}$$

Define:

$$\begin{array}{rcl} \mathsf{Odds}(\pi) & = & \frac{\pi}{1-\pi} \\ \log \mathsf{Odds}(\pi) & = & \log\left(\frac{\pi}{1-\pi}\right) \\ \mathsf{logit}(\pi) & \equiv & \log\left(\frac{\pi}{1-\pi}\right) \end{array}$$

$$\log\left(\frac{\pi}{1-\pi}\right) = \alpha$$

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$$\pi = \exp(\alpha)(1-\pi)$$

$$\pi(1+\exp(\alpha)) = \exp(\alpha)$$

$$\log\left(\frac{\pi}{1-\pi}\right) = \alpha$$

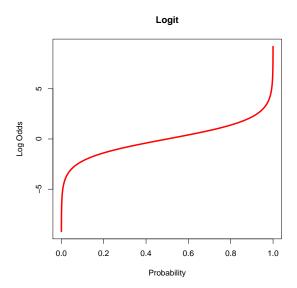
$$\left(\frac{\pi}{1-\pi}\right) = \exp(\alpha)$$

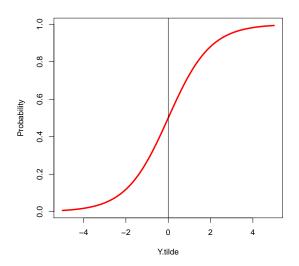
$$\pi = \exp(\alpha)(1-\pi)$$

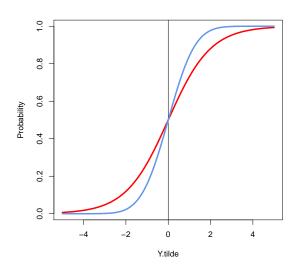
$$\pi(1+\exp(\alpha)) = \exp(\alpha)$$

$$\pi = \frac{\exp(\alpha)}{1+\exp(\alpha)} = \operatorname{Logit}^{-1}(\alpha)$$

$$\begin{split} \log\left(\frac{\pi}{1-\pi}\right) &= \alpha \\ \left(\frac{\pi}{1-\pi}\right) &= \exp(\alpha) \\ \pi &= \exp(\alpha)(1-\pi) \\ \pi(1+\exp(\alpha)) &= \exp(\alpha) \\ \pi &= \frac{\exp(\alpha)}{1+\exp(\alpha)} = \mathsf{Logit}^{-1}(\alpha) \\ \frac{\exp(\alpha)}{1+\exp(\alpha)} &= \frac{1}{1+\exp(-\alpha)} = \mathsf{Logistic}(\alpha) \end{split}$$







 $Z \sim \mathsf{Logistic}(\mu, s)$

$$\begin{array}{lcl} Z & \sim & \mathsf{Logistic}(\mu,s) \\ f(z) & = & \frac{\exp(\frac{z-\mu}{s})}{s(1+\exp(\frac{z-\mu}{s}))^2} \end{array}$$

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$$f(z) = \frac{\exp(\frac{z-\mu}{s})}{s(1 + \exp(\frac{z-\mu}{s}))^2}$$

$$Z \sim \operatorname{Logistic}(0, 1)$$
 Standard Logistic

$$\begin{array}{rcl} Z & \sim & \mathsf{Logistic}(\mu,s) \\ f(z) & = & \frac{\exp\left(\frac{z-\mu}{s}\right)}{s(1+\exp\left(\frac{z-\mu}{s}\right))^2} \\ Z & \sim & \underbrace{\mathsf{Logistic}(0,1)}_{\mathsf{Standard Logistic}} \\ f(z) & = & \frac{\exp(z)}{(1+\exp(z))^2} \end{array}$$

$$\begin{split} \tilde{Y}_i &= & \boldsymbol{X}_i' \boldsymbol{\beta} + \epsilon_i \\ \epsilon_i &\sim & \mathsf{Logistic}(0,1) \end{split}$$

$$\begin{split} \tilde{Y}_i &= & \boldsymbol{X}_i'\boldsymbol{\beta} + \epsilon_i \\ \epsilon_i &\sim & \mathsf{Logistic}(0,1) \\ P(Y_i = 1|\boldsymbol{X}_i,\boldsymbol{\beta}) &= & I(\tilde{Y}_i > 0) \end{split}$$

$$egin{array}{lcl} ilde{Y}_i &=& oldsymbol{X}_i'oldsymbol{eta} + \epsilon_i \ & \epsilon_i &\sim & \mathsf{Logistic}(0,1) \ P(Y_i = 1 | oldsymbol{X}_i, oldsymbol{eta}) &=& I(ilde{Y}_i > 0) \ &=& I(oldsymbol{X}_i'oldsymbol{eta} + \epsilon_i > 0) \end{array}$$

$$\begin{split} \tilde{Y}_i &= \boldsymbol{X}_i'\boldsymbol{\beta} + \epsilon_i \\ \epsilon_i &\sim \operatorname{Logistic}(0,1) \\ P(Y_i = 1 | \boldsymbol{X}_i, \boldsymbol{\beta}) &= I(\tilde{Y}_i > 0) \\ &= I(\boldsymbol{X}_i'\boldsymbol{\beta} + \epsilon_i > 0) \\ &= I(\epsilon_i > -\boldsymbol{X}_i'\boldsymbol{\beta}) \end{split}$$

$$\begin{split} \tilde{Y}_i &= \mathbf{X}_i' \boldsymbol{\beta} + \epsilon_i \\ \epsilon_i &\sim \mathsf{Logistic}(0,1) \\ P(Y_i = 1 | \mathbf{X}_i, \boldsymbol{\beta}) &= I(\tilde{Y}_i > 0) \\ &= I(\mathbf{X}_i' \boldsymbol{\beta} + \epsilon_i > 0) \\ &= I(\epsilon_i > -\mathbf{X}_i' \boldsymbol{\beta}) \\ &= I(\epsilon_i < \mathbf{X}_i' \boldsymbol{\beta}) \end{split}$$

$$\begin{split} \tilde{Y}_i &= \boldsymbol{X}_i'\boldsymbol{\beta} + \epsilon_i \\ \epsilon_i &\sim \operatorname{Logistic}(0,1) \\ P(Y_i = 1 | \boldsymbol{X}_i, \boldsymbol{\beta}) &= I(\tilde{Y}_i > 0) \\ &= I(\boldsymbol{X}_i'\boldsymbol{\beta} + \epsilon_i > 0) \\ &= I(\epsilon_i > -\boldsymbol{X}_i'\boldsymbol{\beta}) \\ &= I(\epsilon_i < \boldsymbol{X}_i'\boldsymbol{\beta}) \\ &= \operatorname{Logistic}(\boldsymbol{X}_i'\boldsymbol{\beta}) = F(\boldsymbol{X}_i'\boldsymbol{\beta}) \end{split}$$

$$Y_i \sim \operatorname{Bernoulli}(\pi_i)$$

 $\pi_i = F(\boldsymbol{X}_i'\boldsymbol{\beta})$

Likelihood function

$$L(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = f(\boldsymbol{Y}|\boldsymbol{X},\boldsymbol{\beta})$$

$$= \prod_{i=1}^{N} f(Y_i|\boldsymbol{X}_i,\boldsymbol{\beta})$$

$$= \prod_{i=1}^{N} F(\boldsymbol{X}_i'\boldsymbol{\beta})^{Y_i} (1 - F(\boldsymbol{X}_i'\boldsymbol{\beta}))^{1-Y_i}$$

Log-likelihood:

$$l(\boldsymbol{\beta}|\boldsymbol{X},\boldsymbol{Y}) = \sum_{i=1}^{N} Y_{i} \log F(\boldsymbol{X}_{i}'\boldsymbol{\beta}) + (1 - Y_{i}) \log(1 - F(\boldsymbol{X}_{i}'\boldsymbol{\beta}))$$

Homework: Derive Score + Hessian for Logistic Normal + Implement with Optim

R Code

Logit Model: Most Important Problem

Most Important Problem

	Probit	Logit
Intercept	-0.26	-0.41
	(0.03)	(0.05)
Post-Speech	-0.01	-0.02
	(0.05)	(0.09)
Republican	-0.20	-0.32
	(0.05)	(80.0)

$$E[Y_i|\mathbf{X}_i] = \frac{1}{1 + \exp(-\beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2})}$$

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$$E[Y_i|X_{i1} = 0, X_{i2} = 0] = \frac{1}{1 + \exp(0.41)} = 0.40$$

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$$E[Y_i|X_{i1} = 0, X_{i2} = 1] = \frac{1}{1 + \exp(0.41 + 0.32)} = 0.33$$

Expected Value:

$$E[Y_i|X_i] = \frac{1}{1 + \exp(-\beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2})}$$

$$E[Y_i|X_{i1} = 0, X_{i2} = 0] = \frac{1}{1 + \exp(0.41)} = 0.40$$

$$E[Y_i|X_{i1} = 0, X_{i2} = 1] = \frac{1}{1 + \exp(0.41 + 0.32)} = 0.33$$

First Differences

Expected Value:

$$E[Y_i|X_i] = \frac{1}{1 + \exp(-\beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2})}$$

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First Differences

$$E[Y_i|X_{i1} = 1, X_{i2} = 0] - E[Y_i|X_{i1} = 0, X_{i2} = 0] = 0.3941 - 0.3989 = -0.005$$

Expected Value:

$$E[Y_i|X_i] = \frac{1}{1 + \exp(-\beta_0 - \beta_1 X_{i1} - \beta_2 X_{i2})}$$

$$E[Y_i|X_{i1} = 0, X_{i2} = 0] = \frac{1}{1 + \exp(0.41)} = 0.40$$

$$E[Y_i|X_{i1} = 0, X_{i2} = 1] = \frac{1}{1 + \exp(0.41 + 0.32)} = 0.33$$

First Differences

$$E[Y_i|X_{i1} = 1, X_{i2} = 0] - E[Y_i|X_{i1} = 0, X_{i2} = 0] = 0.3941 - 0.3989 = -0.005$$

$$E[Y_i|X_{i1} = 1, X_{i2} = 1] - E[Y_i|X_{i1} = 0, X_{i2} = 1] = 0.3208 - 0.32519 = -0.0043$$

$$\begin{array}{ll} \frac{\partial F(\boldsymbol{X}_{i}'\boldsymbol{\beta})}{\partial X_{ij}} & = & f(\boldsymbol{X}_{i}'\boldsymbol{\beta})\beta_{j} \\ \text{Max Effect} & = & 0.25 \times \beta_{j} \end{array}$$

Most Important Problem: Linear Probability Model

	Probit	Logit	LPM
Intercept	-0.26	-0.41	0.40
	(0.03)	(0.05)	(0.13)
Post-Speech	-0.01	-0.02	-0.004
	(0.05)	(0.09)	(0.02)
Republican	-0.20	-0.32	-0.07
	(0.05)	(80.0)	(0.02)

LPM might (usually) be ok, but

- 1) Probit/Logit = easy to run
- 2) Base for more complicated models