350c

Lecture 8: Goodness of Fit Measures for Binary Outcome Models

Jens Hainmueller

Stanford

Hainmueller (Stanford) 1 / 4

Outline

- Binary Outcomes
- Calibration Plots
- 3 Discrimination
 - Classification accuracy
 - Sensitivity and specificity
- 4 Summary Accuracy Measures

Hainmueller (Stanford)

Outline

- Binary Outcomes
- Calibration Plots
- Oiscrimination
 - Classification accuracy
 - Sensitivity and specificity
- Summary Accuracy Measures

Regression Models for Binary Outcomes

- Often we are interested in binary outcomes Y that take on two values 1 or 0 depending on whether an event of interest occurred or not (vote or not, war or peace, etc).
- Three widely used regression models for binary outcomes:
 - Linear Probability Model (LPM)
 - Probit regression
 - Logistic regression
- Goals
 - Model the probability of the event occurring as a function of the independent variables
 - Estimate the probability of the event occurring given certain values of the independent variables
 - Predict how the probability of the event occurring changes for different values of the independent variables
 - Classify the observations into categories based on the probability of the event occurring

Hainmueller (Stanford) 4 / -

Assessing Goodness of Fit

- How do we validate that our fitted model is any good? How well does our model fit the data?
- Two core concepts to assess model fit for binary outcome models:

Calibration

- A model is well calibrated if the observed risk matches the predicted risk (e.g. for observations with a predicted probability of .2 there should be approximately 20% of Y = 1 observed outcomes)
- Well calibrated model ensures that in future samples the observed proportions of positives will be close to the estimated probabilities

Discrimination

- A model has high discrimination ability if it allows us to discriminate between low and high risk observations
- This means observations with Y = 1 should have high and observations with Y = 0 should have low predicted probabilities \hat{Y}
- But we don't want a lot of \hat{Y} close to .5 if outcomes are all Y = 1 or Y = 0 (sharpness)

Hainmueller (Stanford) 5 / 41

Outline

- Binary Outcomes
- Calibration Plots
- Oiscrimination
 - Classification accuracy
 - Sensitivity and specificity
- Summary Accuracy Measures

Calibration

- In a well calibrated model, the observed risk matches the predicted risk such that in future samples the observed proportion of positives will match the predicted probabilities
- Can check this for any model prediction using calibration plots:
 - Bin predicted probabilities \hat{Y} into 10 equal interval bins $g_k \in ([0,.1], (.1,.2], ..., (.9,1])$
 - In each bin g_k compute the average predicted probability $\hat{Y}_{\in g_k}$ and the fraction of observed positives $\bar{Y}_{\in g_k}$
 - Plot the two vectors of averages against each other
- If the model is well calibrated, the binned averages should trace the identity line

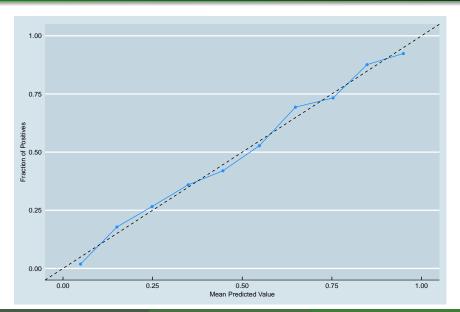
Hainmueller (Stanford) 7 / 4

Calibration Plot: Correct Model

```
> n <- 1000
> x <- rnorm(n)
> xh <- 2*x
> pr <- exp(xb)/(1+exp(xb))
> v <- 1*(runif(n) < pr)
> d <-data.frame(y,x)
> logit <- glm(v~x, data=d, family=binomial(link="logit"))
> summary(logit)
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) -0.05078 0.08164 -0.622
                                          0.534
            1 99622 0 12797 15 599 <2e-16 ***
>
> # cut into 10 bins
> d$predprobs <- logit$fitted
> d$pptbin <- cut(d$predprobs,seq(from=0,to=1,by=0.1),include.lowest = TRUE)
> # compute average predicted probs
> rel.logit <- (d %>% group_by(pptbin) %>% summarise(out = mean(y)))
> # compute observed averages
           <- (d %>% group_by(pptbin) %>% summarise(out = mean(predprobs)))
> mpv
> pdf <- data.frame(fop = rel.logit$out.
                    mpv = mpv sout)
> ggplot(pdf.aes(x=mpv.v=fop)) + geom point(color="dodgerblue") + geom line(color="dodgerblue") +
    theme economist(dkpanel = T) + geom abline(intercept = 0.slope=1.linetype=2) +
    xlim(0,1) + ylim(0,1) + xlab("Mean Predicted Value") + ylab("Fraction of Positives")
```

Hainmueller (Stanford) 8 / 4

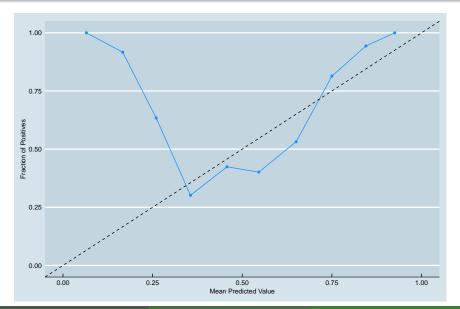
Calibration Plot: Correct Model



Calibration Plot: Miss-specified Model

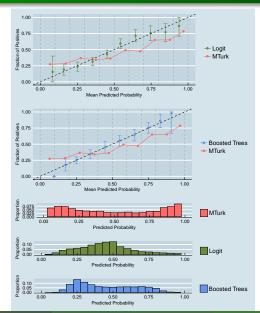
R Code > xb <- -2*x + 1.25*x^2 > pr <- exp(xb)/(1+exp(xb))> v <- 1*(runif(n) < pr) > d <-data.frame(v.x) > logit <- glm(y~x, data=d, family=binomial(link="logit")) > summarv(logit) Coefficients: Estimate Std. Error z value Pr(>|z|) 0.07215 8.475 <2e-16 *** (Intercept) 0.61150 -0.92570 0.08506 -10.883 <2e-16 *** > d\$predprobs <- logit\$fitted > head(d) x predprobs 0.07894658 0.6314440 2 0 -0 21551746 0 6923231 3 0 -0 25559249 0 7001684 0.77489539 0.4735684 5 0 2.62295669 0.1398448 6 0 0 50200285 0 5366323 > # cut predicted probabilitie sinto 10 bins > d\$pptbin <- cut(d\$predprobs.seq(from=0.to=1.bv=0.1).include.lowest = TRUE)</pre> > rel.logit <- (d %>% group bv(pptbin) %>% summarise(out = mean(v))) > mpv <- (d %>% group_by(pptbin) %>% summarise(out = mean(predprobs))) > pdf <- data.frame(fop = rel.logit\$out. mpv = mpv sout)> ggplot(pdf.aes(x=mpv.v=fop)) + geom point(color="dodgerblue") + geom line(color="dodgerblue") + theme economist(dkpanel = T) + geom abline(intercept = 0.slope=1.linetype=2) + xlim(0,1) + ylim(0,1) + xlab("Mean Predicted Value") + ylab("Fraction of Positives")

Example with Miss-specified Model



Hainmueller (Stanford) 11 /

Predicting Recidivism



Outline

- Binary Outcomes
- Calibration Plots
- 3 Discrimination
 - Classification accuracy
 - Sensitivity and specificity
- 4 Summary Accuracy Measures

Hainmueller (Stanford) 13 /

Discrimination

- Good calibration of a model is not sufficient for model validation
 - A silly model that predicts the same predicted probability for every observation will have good calibration (in future samples the observed proportion will be close to the predicted probability)
 - However, such a model would not be helpful since it does not help us to discriminate between high and low risk observations
- Need to check discrimination ability of a model
 - Discrimination is high if Y = 1 observations correspond with high \hat{Y} values and observations with Y = 0 correspond to low \hat{Y} values
- Discrimination can be captured in various ways including classification accuracy, sensitivity, specificity, precision, recall, etc.
- Many metrics are used and different fields have different preferences

Hainmueller (Stanford) 14 / 41

Outline

- Binary Outcomes
- Calibration Plots
- 3 Discrimination
 - Classification accuracy
 - Sensitivity and specificity
- Summary Accuracy Measures

Hainmueller (Stanford)

Classification accuracy

For classification we choose a cutoff value on the probability scale, say c=.5, and classify all predicted values above c as predicting a positive, and all below c as predicting a negative

The contingency table or confusion matrix for a binary classifier is given by:

| | Observed | |
|--------------------------------------|----------------------|---------------------|
| | Positive ($Y = 1$) | Negative $(Y = 0)$ |
| Predicted positive ($\hat{Y} > c$) | True Positive (TP) | False Positive (FP) |
| Predicted negative ($\hat{Y} < c$) | False Negative (FN) | True Negative (TN) |
| | | |
| | Total Positive (P) | Total Negative (N) |

- Accuracy: ACC = (TP + TN)/(P + N)
 - Percent correctly classified of all predictions

Hainmueller (Stanford) 16 / 41

Recidivism Example: Classification accuracy

```
> source("http://pcwww.liv.ac.uk/~william/R/crosstab.r")
> d <- read.csv("bail.csv")
> d$v <- d$two vear recid
> crosstab(d,row.vars="y")
      Count Total %
      524 0
               52 4
      476.0 47.6
  Sum 1000 0 100 0
> # model 1
> logit1 <- glm(y~sex+age+as.factor(race),data=d)
> summarv(logit1)
Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.653519 0.055209 11.837 < 2e-16 ***
             -0.162660 0.038255 -4.252 2.32e-05 ***
sex
               -0.006605 0.001296 -5.096 4.16e-07 ***
age
as.factor(race)2 0.162315 0.033086 4.906 1.09e-06 ***
as.factor(race)3 -0.012272 0.057799 -0.212
                                            0.832
as.factor(race)4 -0.195496   0.183733 -1.064
                                            0.288
as.factor(race)5 -0.475177 0.480880 -0.988
                                            0 323
> d$vhat1 <- logit1$fitted
> d$vpred1 <- as.numeric(d$vhat1>.5)
```

Hainmueller (Stanford) 17 / 41

Recidivism Example: Classification accuracy

```
R Code
> # model 2
> logit2 <- glm(v~sex+age+as.factor(race)+log(1+priors count).data=d)</pre>
> summary(logit2)
Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
                     0.553299
                                0.053324 10.376 < 2e-16 ***
(Intercept)
                     -0.104176
                                0.036774 -2.833 0.00471 **
SPX
                     -0.007815
                                0.001237 -6.318 3.99e-10 ***
age
                    0.089658 0.032197 2.785 0.00546 **
as.factor(race)2
                                0.054915 -0.384 0.70094
as.factor(race)3
                   -0.021096
                  -0.070564
as.factor(race)4
                                0.174957 -0.403 0.68680
as.factor(race)5
                   -0.573363 0.456928 -1.255 0.20984
log(1 + priors count) 0.166681
                                0 016016 10 407 < 2e-16 ***
> d$yhat2 <- logit2$fitted
> d$vpred2 <- as.numeric(d$vhat2>.5)
```

Hainmueller (Stanford) 18 / 41

Recidivism Example: Classification accuracy

```
R Code
> # model 1
> crosstab(d,row.vars="ypred1",col.vars="y",type="frequency",percentages=F)
                 1 Sum
ypred1
          352
                    556
         172 272
                    444
          524 476 1000
> ACCm1 <- mean(d$ypred1 == d$y)
> ACCm1
Γ17 0.624
> # model 2
> crosstab(d,row,vars="vpred2",col,vars="v",type="frequency",percentages=F)
                 1 Sum
ypred2
          378
              178
                    556
              298 444
         146
         524 476 1000
> ACCm2 <- mean(d$ypred2 == d$y)
> ACCm2
Γ17 0.676
> # all negative model (null model)
> ACCneg <- mean(0 == d$v)
> ACCneg
Γ17 0.524
> # percent increase in accuracy over null model
> ((ACCm2-ACCneg)/ACCneg)*100
Γ17 29.00763
```

Hainmueller (Stanford)

Classification accuracy (Cutpoint .85)

```
R Code
> # model 1
> d$ypred1 <- as.numeric(d$yhat1>.85)
> # model 2
> d$ypred2 <- as.numeric(d$yhat2>.85)
> # model 1
> crosstab(d,row.vars="ypred1",col.vars="y",type="frequency",percentages=F)
ypred1
          524 476 1000
          524 476 1000
> ACCm1 <- mean(d$vpred1 == d$v)
> ACCm1
Γ17 0.524
> # model 2
> crosstab(d,row.vars="ypred2",col.vars="y",type="frequency",percentages=F)
               1 Sum
ypred2
          521 455 976
                21
                     24
          524 476 1000
Sum
> ACCm2 <- mean(d$vpred2 == d$v)
> ACCm2
Γ17 0.542
> # all negative model (null model)
> ACCneg <- mean(0 == d$v)
> ACCneg
Γ17 0.524
```

Is classification accuracy sufficient?

Hainmueller (Stanford) 21 / 41

Accuracy Paradox

```
R Code
> logit1 <- glm(war~lpop+polity2+gdpen+ethfrac+lmtnest+0il.data=d.family= binomial(link = "logit"))</pre>
> summarv(logit1)
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) -6.38555 1.95858 -3.260 0.00111 **
          0.47944 0.20212 2.372 0.01769 *
1pop
          0.04639 0.04466 1.039 0.29896
polity2
gdpen
        -0.32338 0.14613 -2.213 0.02690 *
ethfrac 0.77218 0.97873 0.789 0.43014
Imtnest 0.37545 0.21782 1.724 0.08476 .
Oi1
           -0.23107 0.77083 -0.300 0.76435
> d$yhat1 <- logit1$fitted
> d$vpred1 <- as.numeric(d$vhat1>.5)
> crosstab(d,row.vars="ypred1",col.vars="y",type="frequency",percentages=F)
          0 1 Sum
      У
ypred1
        112 19 131
         3
             2
        115 21 136
> ACCm1 <- mean(d$ypred1 == d$y)
> ACCm1
Γ17 0.8382353
> # all negative model (null model)
> ACCneg <- mean(0 == d$v)
> ACCneg
[1] 0.8455882
```

Hainmueller (Stanford)

Outline

- Binary Outcomes
- Calibration Plots
- Oiscrimination
 - Classification accuracy
 - Sensitivity and specificity
- Summary Accuracy Measures

Hainmueller (Stanford)

Sensitivity and specificity

| | Observed | |
|--------------------------------------|---------------------|----------------------|
| | Positive $(Y = 1)$ | Negative ($Y = 0$) |
| Predicted positive ($\hat{Y} > c$) | True Positive (TP) | False Positive (FP) |
| Predicted negative ($\hat{Y} < c$) | False Negative (FN) | True Negative (TN) |
| | | |
| | Total Positive (P) | Total Negative (N) |

- Sensitivity: TPR = TP/P
 - Percent of positives that are correctly classified (also called recall or true positive rate)
 - For a screening test with higher sensitivity, fewer actual positives go undetected (lower type 2 error)
- Specificity: SPC = TN/N
 - Percent of negatives that are correctly classified (also called true negative rate)
 - For a screening test with higher specificity, fewer negatives are labeled as positive (lower type 1 error)
- In theory a classifier might achieve 100% on both, but in practice there are often tradeoffs and we do not attain 100%
 - The higher c the lower the sensitivity, but the higher the specificity

Hainmueller (Stanford) 24 / 41

Sensitivity and specificity

```
> # model 1
> crosstab(d,row.vars="ypred1",col.vars="y",type="frequency",percentages=F)
                1 Sum
ypred1
                   556
          172 272 444
          524 476 1000
Sum
> crosstab(d,row.vars="ypred1",col.vars="y",type="column.pct",percentages=F)
           0 1
ypred1
         0.67 0.43
        0.33 0.57
        1 00 1 00
Sum
> # model 2
> crosstab(d,row.vars="ypred2",col.vars="y",type="frequency",percentages=F)
                 1 Sum
            0
ypred2
         146 298 444
          524 476 1000
> crosstab(d,row.vars="ypred2",col.vars="y",type="column.pct",percentages=F)
ypred2
         0.72 0.37
         0.28 0.63
         1.00 1.00
Sum
```

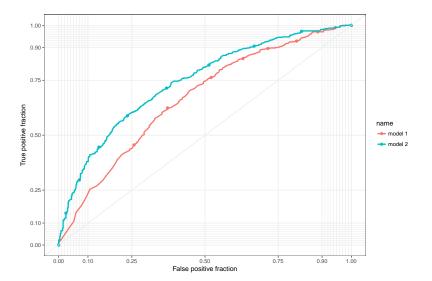
Hainmueller (Stanford)

- ROC curves plot the sensitivity against the specificity across all cutpoints c from 0 to 1
- High discrimination ability means high sensitivity and specificity simultaneously. For such models the ROC curve goes close to the top left corner
- Low discrimination ability means low sensitivity and specificity simultaneously. For such models the ROC curve goes close to the 45 degree diagonal line.
- Can plot ROC curves to compare discrimination ability across models

Hainmueller (Stanford) 26 / 41

```
R Code
> library(plotROC)
> dd <- data.frame(y=rep(d$y,2),
                  pred=c(d$yhat1,
                         d$vhat2).
                  name=rep(c("model 1"."model 2").each=nrow(d))
> head(dd)
         pred
                 name
   0.3030943 model 1
2 0 0 5318072 model 1
3 1 0.5119914 model 1
   0.4091240 model 1
5 0 0 3163049 model 1
6 1 0 6837284 model 1
> ggroc <- ggplot(dd, aes(d = y, m = pred, color = name)) + geom_roc(labels = FALSE) + style_roc()
> pdf("ROC.pdf",width=9,height=6)
> ggroc
> dev.off()
```

Hainmueller (Stanford) 27 / 41



Hainmueller (Stanford)

Area Under ROC Curve

- We can summarize the discrimination ability of a model by computing the area under the estimated ROC curve (AUC).
- Maximum discrimination ability means AUC = 1
- Minimum discrimination ability means AUC = 0.5
- Can show that AUC is the probability a random pair of observations, one with Y = 1 and one with Y = 0, the observation with Y = 1 is correctly ranked as having higher predicted probability than the other.

Hainmueller (Stanford) 29 / 41

```
> ggroc <- ggplot(dd, aes(d = y, m = pred, color = name)) + geom_roc(labels = FALSE) + style_roc()
> calc_auc(ggroc)
  PANEL group
                  AUC
        1 0.6654632
    1 2 0.7413461
> library(pROC)
> roccurve <- roc(d$v ~ d$vhat1)
> auc(roccurve)
Area under the curve: 0.6655
```

Hainmueller (Stanford) 30 / 41

Outline

- Binary Outcomes
- Calibration Plots
- 3 Discrimination
 - Classification accuracy
 - Sensitivity and specificity
- 4 Summary Accuracy Measures

Hainmueller (Stanford)

Brier Score

To summarize the accuracy of binary classifiers we use the Brier score:

$$BS = \frac{1}{N} \sum_{i}^{N} (\hat{Y}_i - Y_i)^2$$

- If $\hat{Y}_i = 1$ and $Y_1 = 1$, then BS = 0, the best score achievable.
- If $\hat{Y}_i = 1$ and $Y_1 = 0$, then BS = 1, the worst score achievable.
- Note that this measure is based on relative deviations ($\hat{Y}_i = .51$ pays a much lager cost than $\hat{Y}_i = .99$ for $Y_i = 1$), rather than binary deviations (such as correctly or incorrectly classified for a given cutoff c)
- Note that this is proper score function:
 - If \hat{Y} is the vector of predictions and you use a scoring function to give a reward of $S(\hat{Y}, i)$ if the i-th event occurs
 - The highest expected reward is obtained by reporting the true probability distribution so you encourages the forecaster to be honest to maximize the expected reward

Hainmueller (Stanford) 32 / 4

Brier Score

BS can be decomposed as the sum of a measure of calibration and refinement:

$$BS = \frac{1}{N} \sum_{i}^{N} (\hat{Y}_{i} - Y_{i})^{2} = \frac{1}{N} \sum_{k}^{K} (\hat{Y}_{k} - \bar{Y}_{k})^{2} + \frac{1}{N} \sum_{k}^{K} n_{k} (\bar{Y}_{k} (1 - \bar{Y}_{k}))$$

- K is the # of unique predicted probabilities, n_k is the # of predictions in predicted probability category k (s.t. $N = \sum_k n_k$), \bar{Y}_k is the observed frequency in k given the predicted probability \hat{Y}_k
- The first term captures calibration
 - How much do observed frequency of positives in k vary from predicted fraction \hat{Y}_k
- The second term refinement (related to the area under the ROC curve)
 - How uniform is occurrence of outcomes within each group k

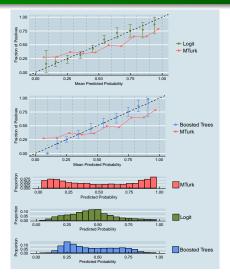
Hainmueller (Stanford) 33 / 41

Brier Score

```
> mean((d$ypred1-d$y)^2)
[1] 0.1617647
> mean((d-d$y)^2)
[1] 0.1544118
```

Hainmueller (Stanford) 34 / 41

Predicting Recidivism



Brier scores are: Mturk .24, Logit = .21 and Boosted Trees = .19 (all three model are similar on AUC)

- Often it is useful to validate model looking at out of sample fit instead of in-sample fit
- Most common method is k-fold Cross Validation
- Cut the dataset randomly into say 10 folds
- Start a loop, pick the first fold as the hold out, fit the model on the remaining 9 folds and predict outcomes for hold out fold
- Repeat while using next fold as hold out
- In the end you have an out of sample prediction for each observation and then examine model accuracy using this prediction
- Common is 10 folds, but can use other values

Hainmueller (Stanford) 36 / 41

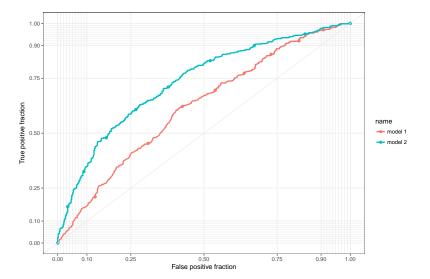
```
R Code
> ## bail example cross validation
> rm(list=ls())
> d <- read.csv("bail.csv")
> d$y <- d$two_year_recid
> Foldsize <- nrow(d)/10
> Foldsize
Γ17 100
> # fold var
> d$fold <- sample(rep(1:10,each=Foldsize),nrow(d),replace=F)
> d$yhat1 <- NA
> d$yhat2 <- NA
> for(i in 1:10){
+ # model 1
+ logit1
               <- glm(v~sex+age.data=d[d$fold!=i.])
+ d$yhat1[d$fold==i] <- predict(logit1,newdata=d[d$fold==i,])
+ # model 2
+ logit2 <- glm(y~sex+age+log(1+priors_count),data=d[d$fold!=i,])
+ d$yhat2[d$fold==i] <- predict(logit2,newdata=d[d$fold==i,])
+ }
```

Hainmueller (Stanford) 37 / 4

. R Code .

```
> dd <- data.frame(y=rep(d$y,2),
                   pred=c(d$yhat1,
                           d$vhat2).
                   name=rep(c("model 1", "model 2"), each=nrow(d))
> head(dd)
         pred
                 name
1 0 0.2103555 model 1
2 0 0.4504963 model 1
> dd$name <- factor(dd$name)
> ggroc <- ggplot(dd, aes(d = y, m = pred, color = name)) + geom_roc(labels = FALSE) + style_roc()
> pdf("ROC2.pdf",width=9,height=6)
> ggroc
> dev.off()
null device
> # AUC
> calc_auc(ggroc)
  PANEL group
                    ALIC
           1 0 6183747
            2 0.7311746
>
> # RS
> mean((d$yhat1-d$y)^2)
Γ17 0.2383642
> mean((d$yhat2-d$y)^2)
Γ17 0.2106819
```

Cross-Validation ROC Curve

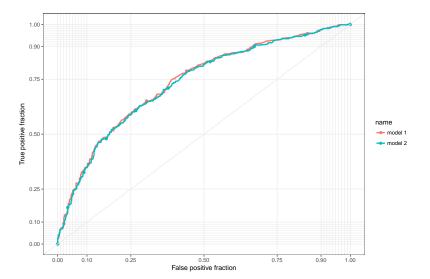


Hainmueller (Stanford) 39 / 4

```
R Code
> logit2
               <- glm(y~sex+age+log(1+priors_count),data=d)
> d$yhat2in
               <- predict(logit2,newdata=d)
> dd <- data.frame(y=rep(d$y,2),
                   pred=c(d$vhat2in.
                          d$yhat2),
                   name=rep(c("model 1", "model 2"), each=nrow(d))
> ggroc <- ggplot(dd, aes(d = y, m = pred, color = name)) + geom_roc(labels = FALSE) + style_roc()
> # ALIC
> calc_auc(ggroc)
  PANEL group
                    AUC.
      1 1 0 7369259
            2 0.7311746
> # RS
> mean((d$yhat2in-d$y)^2)
[1] 0.2087114
> mean((d$yhat2-d$y)^2)
[1] 0.2106819
```

Hainmueller (Stanford) 40 / 4

Cross-Validation ROC Curve



Hainmueller (Stanford) 41 / 4