

# Political Methodology III: Model Based Inference

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# Model Based Inference

- 1) Likelihood inference
- 2) Logit/Probit
  - a) DGP
  - b) Optimization
  - c) Quantities of Interest
  - d) Perfect + Near Perfect Separation
- 3) Ordered Probit
  - a) DGP
  - b) Optimization
  - c) Quantities of Interest

# Civil Conflict and Political Instability (h/t: Yamamoto)

Fearon & Laitin (2003):

- $Y_i$ : Civil conflict
- $T_i$ : Political instability
- $W_i$ : Geography (log % mountainous)

Estimated model:

$$\begin{aligned}\widehat{\Pr(Y_i = 1 \mid T_i, W_i)} \\ = \text{logit}^{-1}(-2.84 + 0.91T_i + 0.35W_i)\end{aligned}$$

Predicted probability:

$$\hat{\pi}(T_i = 1, W_i = 3.10) = 0.299$$

ATE:

$$\begin{aligned}\hat{\tau} &= \frac{1}{n} \sum_{i=1}^n \{\hat{\pi}(1, W_i) - \hat{\pi}(0, W_i)\} \\ &= 0.127\end{aligned}$$

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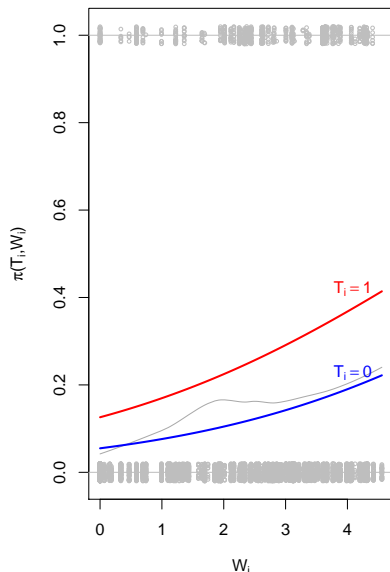
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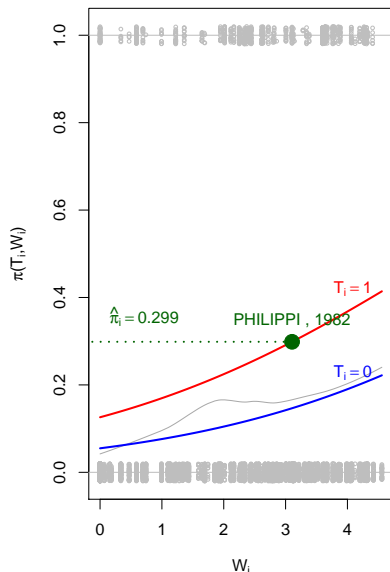
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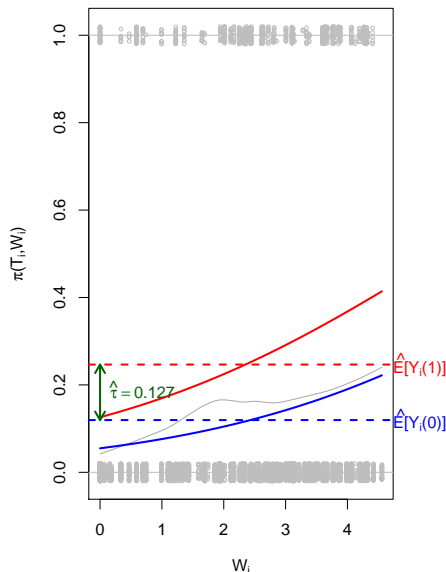
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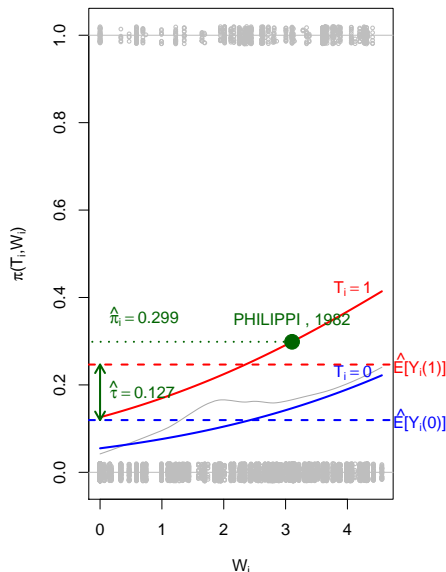
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$Y_i$  = Vote on ACA

$X_{i1}$  = Democrat

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> test_model<- glm(vote~dem + ideo,  
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We have problems!



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To fit data: set  $\beta_0 \rightarrow -\infty$ .

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**You need to make more assumptions**

# Add a Few Observations...

	Nay	Yea
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# Add a Few Observations...

	Nay	Yea
Republican	178.5	0.5
Democrat	34.5	219.5

# Penalized (Prior)-Logistic Regression

- Separation: causes coefficients to diverge
- Penalty (prior): force coefficients towards zero

Step 1: Standardize inputs (Gelman et al )

- Binary variables: mean 0, differ by 1.
  - Democrats: (30%). (0.3, -0.7)
- Other variables: mean 0, sd 0.5.

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where:

$$\pi_i = \frac{1}{1 + \exp(-\mathbf{X}'_i \boldsymbol{\beta})}$$

$$|I(\boldsymbol{\beta})| = \text{Determinant of Fisher's information at } \boldsymbol{\beta}$$

$$I(\boldsymbol{\beta}) = \mathbf{X}' \mathbf{W} \mathbf{X}$$

$$\mathbf{W} = \begin{pmatrix} \pi_1(1 - \pi_1) & 0 & \dots & 0 \\ 0 & \pi_2(1 - \pi_2) & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \pi_N(1 - \pi_N) \end{pmatrix}$$

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$$L(\boldsymbol{\beta}|\mathbf{X}, \mathbf{Y}) = \prod_{i=1}^N \pi_i^{Y_i} (1 - \pi_i)^{1-Y_i} |\mathbf{I}(\boldsymbol{\beta})|^{1/2}$$

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# Penalized (Prior)-Logistic Regression

```
jef_pri<- function(params, X, Y){  
  beta<- params  
  y.tilde<- X%*%beta  
  y.prob<- plogis(y.tilde)  
  temp<- matrix(0, nrow = length(Y), ncol=length(Y))  
  part1<- Y%*%log(y.prob) + (1-Y)%*%log(1- y.prob)  
  diag(temp)<- y.prob*(1-y.prob)  
  part2<- 0.5*log(det(t(X)%*%temp%*%X))  
  out<- part1 + part2  
}  
  
firth<- optim(rnorm(3), jef_pri, method = 'BFGS',  
  control=list(fnscale=-1), hessian=T,  
  X = cbind(1, dem, ideo), Y =clean[,3] )
```

# Comparison

	ACA Vote (GLM)	Firth
Intercept	-14.00 (1670.439)	-5.70 (24.68)
Democrat	11.67 (1670.439)	3.30 (42.10)
Ideology	-16.86 (2.71)	-17.72 (2.85)

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  - Occasionally allows very large values (Cauchy)

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lst<- function(x, nu, mu, sigma2){  
  part1<- lgamma( (nu + 1)/2)  
  part2<- lgamma(nu/2)  
  part3<- sqrt(pi *nu*sqrt(sigma2))  
  part4<- 1 + (1/nu)*((( x- mu)^2)/sigma2)  
  part4<- ( - (nu + 1)/2)*log(part4)  
  out<- part4  
  return(out)  
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log_t<- function(params, X, Y, nu, mu, sigma2){  
  beta<- params  
  prior<- 0  
  for(k in 2:ncol(X)){  
    prior<- prior + lst(beta[k], nu, mu, sigma2)  
  }  
  prior<- prior + lst(beta[1], 1, 0, 10)  
  y.tilde<- X%*%beta  
  y.prob<- plogis(y.tilde)  
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In practice: bayesglm in library(arm) is awesome!

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# Ordered Outcome Data

# Modeling Ordered Outcomes (ht: Yamamoto)

- Suppose that the  $J$  choices are **ordered** in a substantively meaningful way
- Examples:
  - “Likert scale” in survey questions (“strongly agree”, “agree”, etc.)
  - Party positions (extreme left, center left, center, right, extreme right)
  - Levels of democracy (autocracy, anocracy, democracy)
  - Health status (healthy, sick, dying, dead)
- Why not use continuous outcome models?
  - Don’t want to assume equal distances between levels
- Why not use categorical outcome models? (on Monday)
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# Ordered Logit and Probit Models (ht: Yamamoto)

- Again, the **latent variable** representation:  $Y_i^* = X_i' \beta + \epsilon_i$
- Assume that  $Y_i^*$  gives rise to  $Y_i$  based on the following scheme:

$$Y_i = \begin{cases} 1 & \text{if } -\infty(=\psi_0) < Y_i^* \leq \psi_1, \\ 2 & \text{if } \psi_1 < Y_i^* \leq \psi_2, \\ \vdots & \vdots \\ J & \text{if } \psi_{J-1} < Y_i^* \leq \infty(=\psi_J) \end{cases}$$

where  $\psi_1, \dots, \psi_{J-1}$  are the **threshold parameters** to be estimated

- If  $X_i$  contains an intercept, one of the  $\psi$ 's must be fixed for identifiability (typically  $\psi_1 = 0$ )
- $\epsilon_j \sim_{\text{iid}} \text{logistic} \Rightarrow$  the **ordered logit** model:

$$\Pr(Y_i \leq j \mid X_i) = \frac{\exp(\psi_j - X_i' \beta)}{1 + \exp(\psi_j - X_i' \beta)}$$

- $\epsilon_j \sim_{\text{iid}} \text{Normal}(0, 1) \Rightarrow$  the **ordered probit** model:

$$\Pr(Y_i \leq j \mid X_i) = \Phi(\psi_j - X_i' \beta)$$

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$$\Pr(Y_i \leq j \mid X_i) = \frac{\exp(\psi_j - X_i' \beta)}{1 + \exp(\psi_j - X_i' \beta)}$$

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# Ordered Logit and Probit Models (ht: Yamamoto)

- Again, the **latent variable** representation:  $Y_i^* = X_i' \beta + \epsilon_i$
- Assume that  $Y_i^*$  gives rise to  $Y_i$  based on the following scheme:

$$Y_i = \begin{cases} 1 & \text{if } -\infty(=\psi_0) < Y_i^* \leq \psi_1, \\ 2 & \text{if } \psi_1 < Y_i^* \leq \psi_2, \\ \vdots & \vdots \\ J & \text{if } \psi_{J-1} < Y_i^* \leq \infty(=\psi_J) \end{cases}$$

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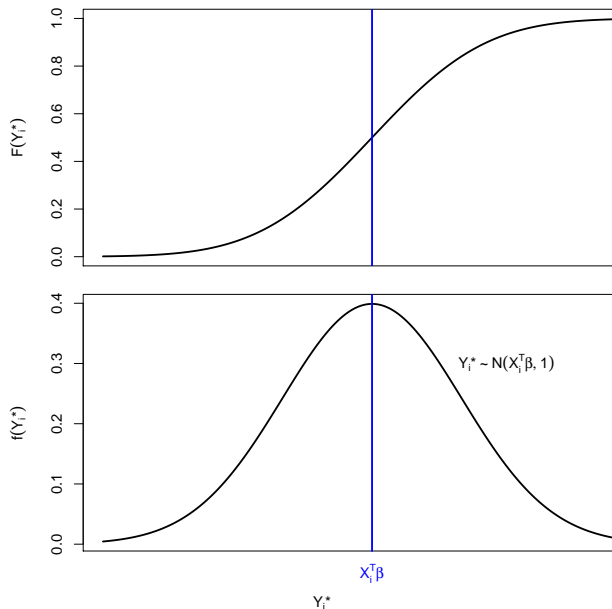
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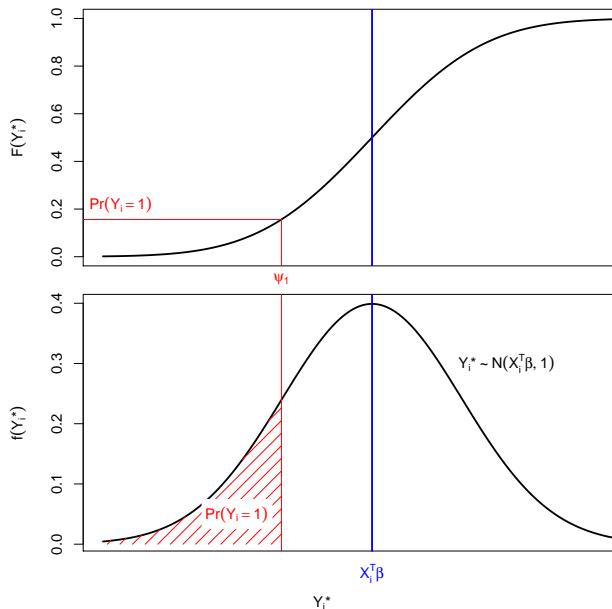
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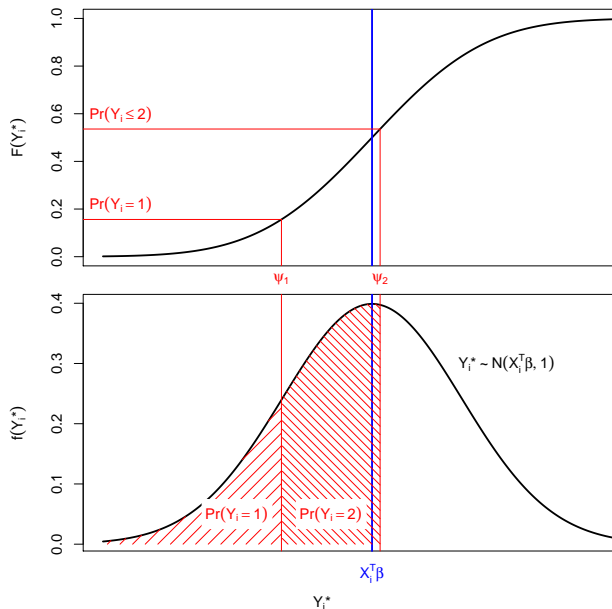
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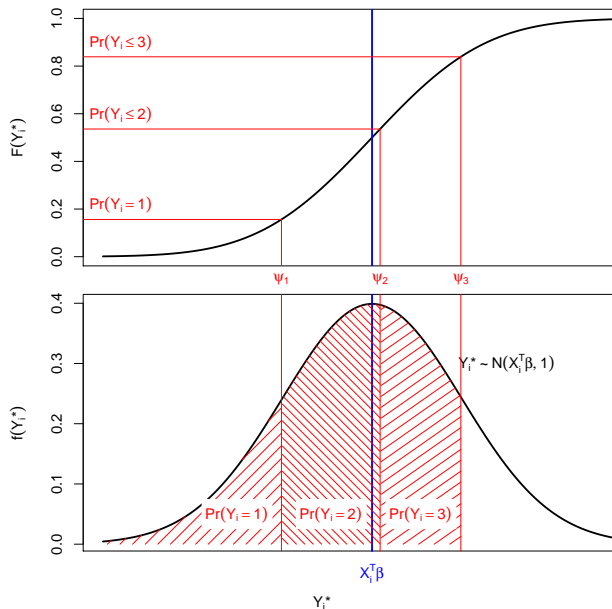


# Ordered Logit and Probit Models (ht: Yamamoto)





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$$L(\boldsymbol{\beta}, \boldsymbol{\Psi}|\mathbf{X}, \mathbf{Y}) = \prod_{i=1}^N \left[ \prod_{j=1}^J [\Phi(\psi_j|\mathbf{X}'_i\boldsymbol{\beta}) - \Phi(\psi_{j-1}|\mathbf{X}'_i\boldsymbol{\beta})]^{I(Y_i=j)} \right]$$

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fit with polr package

# Calculating Quantities of Interest (ht: Yamamoto)

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$$\begin{aligned}\pi_{ij}(X_i) &\equiv \Pr(Y_i = j \mid X_i) = \Pr(Y_i \leq j \mid X_i) - \Pr(Y_i \leq j-1 \mid X_i) \\ &= \begin{cases} \frac{\exp(\psi_j - X_i' \beta)}{1 + \exp(\psi_j - X_i' \beta)} - \frac{\exp(\psi_{j-1} - X_i' \beta)}{1 + \exp(\psi_{j-1} - X_i' \beta)} & \text{for logit} \\ \Phi(\psi_j - X_i' \beta) - \Phi(\psi_{j-1} - X_i' \beta) & \text{for probit} \end{cases}\end{aligned}$$

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- Estimate  $\beta$  and  $\psi$  via MLE, plug the estimates in, replace E with  $\frac{1}{n} \sum$ , and compute CI by delta or MC or bootstrap
- Note that  $X_i' \beta$  appears both before and after the minus sign in  $\pi_{ij}$ 
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Brader, Valentino and Suhay (2008):

- $Y_i$ : Ordinary response to question about increasing immigration
- $T_{1i}, T_{2i}$ : Media cues (immigrant ethnicity  $\times$  story tone)
- $W_i$ : Respondent age and income

Estimated coefficients:

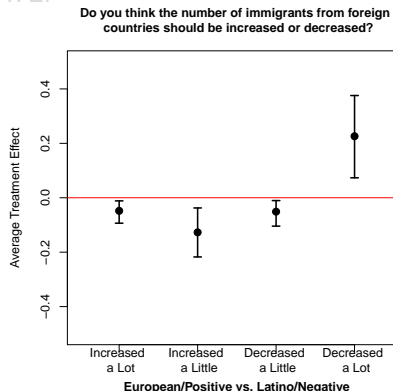
Coefficients:

	Value	s.e.	t
tone	0.27	0.32	0.85
eth	-0.33	0.32	-1.02
ppage	0.01	0.02	1.40
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tone:eth	0.90	0.46	2.16

Intercepts:

	Value	s.e.	t
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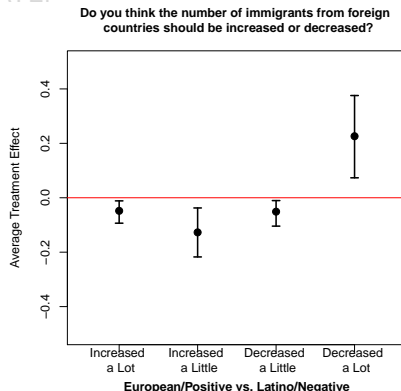
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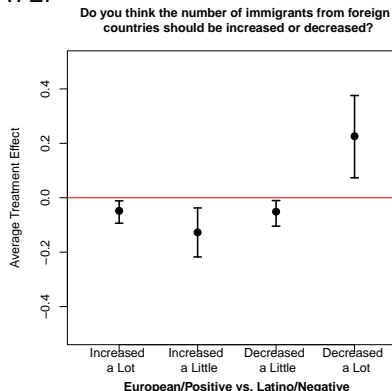
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# Model Based Inference

- 1) Likelihood inference
- 2) Logit/Probit
  - a) DGP
  - b) Optimization
  - c) Quantities of Interest
  - d) Perfect + Near Perfect Separation
- 3) Ordered Probit
  - a) DGP
  - b) Optimization
  - c) Quantities of Interest
- 4) Choice Models: Multinomial Logit/Probit