

Throughout, let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space.

## 1 Conditional expectation

**Theorem 1.1** (Existence and uniqueness of conditional expectation). *Let  $X \in L^1$ , and  $\mathcal{G} \subseteq \mathcal{F}$ . Then there exists a random variable  $Y$  such that*

- $Y$  is  $\mathcal{G}$ -measurable
- $Y \in L^1$ , and  $\mathbb{E}X\mathbf{1}_A = \mathbb{E}Y\mathbf{1}_A$  for all  $A \in \mathcal{G}$ .

Moreover, if  $Y'$  is another random variable satisfying these conditions, then  $Y' = Y$  almost surely.

We call  $Y$  a (version of) the conditional expectation given  $\mathcal{G}$ .

*Proof.* (Existence)

Case 1:  $X \in L^2$ .

Recall that  $L^2$  is a Hilbert space, and that the set of  $\mathcal{G}$ -measurable random variables is a closed subspace of  $L^2$  (it is closed because the space  $L^2(\Omega, \mathcal{G}, \mathbb{P})$  is complete). The projection theorem then gives us the existence and uniqueness of  $Y \in L^2 \subseteq L^1$ .

Case 2:  $X \geq 0 \in L^1$ .

Let  $X_n = X \wedge n \in L^2$ . Then by case 1, we can define  $Y_n = \mathbb{E}(X_n \mid \mathcal{G}) \in L^2$ . We make the following observation

**Lemma 1.0.1.** Suppose  $(X, Y)$  and  $(X', Y')$  are two pairs of random variables satisfying the conditions of the theorem, then  $X \geq X'$  implies  $Y \geq Y'$  almost surely.

*Proof.* Let  $A = \{Y < Y'\}$ . Then  $\mathbb{E}Y\mathbf{1}_A = \mathbb{E}X\mathbf{1}_A \geq \mathbb{E}X'\mathbf{1}_A = \mathbb{E}Y'\mathbf{1}_A$ , so  $\mathbb{E}(Y - Y')\mathbf{1}_A \geq 0$  and  $\mathbb{P}(A) = 0$ . □

It follows that there is some random variable  $Y$  such that  $Y_n \uparrow Y$ . Clearly  $Y$  is  $\mathcal{G}$ -measurable. For any  $A \in \mathcal{G}$ , we have

$$\begin{aligned} \mathbb{E}Y\mathbf{1}_A &= \lim_{n \rightarrow \infty} \mathbb{E}Y_n\mathbf{1}_A && \text{(MCV)} \\ &= \lim_{n \rightarrow \infty} \mathbb{E}X_n\mathbf{1}_A && \text{(MCV)} \\ &= \mathbb{E}X\mathbf{1}_A \end{aligned}$$

Case 3:  $X \in L^1$ .

Write  $X = X^+ - X^-$ , and apply case 2 to  $X^+$  and  $X^-$ .

(Uniqueness) Suppose  $Y$  and  $Y'$  are two random variables satisfying the conditions of the theorem. The  $\{Y > Y'\}$  is in  $\mathcal{G}$  so  $\mathbb{E}Y\mathbf{1}_{\{Y > Y'\}} = \mathbb{E}Y'\mathbf{1}_{\{Y > Y'\}} \implies \mathbb{E}(Y - Y')\mathbf{1}_{\{Y > Y'\}} = 0 \implies \mathbb{P}(Y > Y') = 0$ . Similarly,  $\mathbb{P}(Y' > Y) = 0$ .

**Remark.** The above can also be proved using the Radon-Nikodym theorem.

(Proof via Radon-Nikodym) First recall the Radon-Nikodym theorem

**Proposition 1.1** (Radon-Nikodym theorem). Let  $\mu, \nu$  be two  $\sigma$ -finite measures on  $(\Omega, \mathcal{F})$  such that  $\nu \ll \mu$ . Then there exists a unique (up to a.e. equivalence)  $f \in L^1(\Omega, \mathcal{F}, \mu)$  such that  $\nu(A) = \int_A f d\mu$  for all  $A \in \mathcal{F}$ .

Consider the measure on  $(\Omega, \mathcal{G})$  given by

$$\mu(A) = \mathbb{E}X\mathbf{1}_A, \quad A \in \mathcal{G}$$

so  $\mu \ll \mathbb{P}$ . By the Radon-Nikodym theorem, there exists a unique  $Y \in L^1(\Omega, \mathcal{G}, \mathbb{P})$  such that  $\mu(A) = \int_A Y d\mathbb{P}$  for all  $A \in \mathcal{G}$ .

For general  $X \in L^1$ , we can write  $X = X^+ - X^-$  and apply the above to  $X^+$  and  $X^-$ . □

**Proposition 1.2** (Equivalent definition for conditional expectation). Let  $X, \mathcal{G}$  be as above. Then there exists a random variable  $Y$  such that

- $Y$  is  $\mathcal{G}$ -measurable
- $Y \in L^1$  and  $\mathbb{E}XZ = \mathbb{E}YZ$  for all  $Z \in L^\infty(\mathcal{G})$

Moreover,  $Y = \mathbb{E}(X \mid \mathcal{G})$  almost surely.

*Proof.* (Existence) Set  $Y = \mathbb{E}(X \mid \mathcal{G})$ . It is straightforward to see that  $Y$  satisfies the conditions of the proposition for simple functions  $Z$ . Note that simple functions that are in  $L^p$  are dense in  $L^p$  for  $1 \leq p \leq \infty$ . Let  $Z_n \in L^\infty(\mathcal{G})$  be a sequence of simple functions such that  $Z_n \rightarrow Z$  in  $L^\infty$  (in particular, we have almost sure pointwise convergence). Then

$$\begin{aligned} \mathbb{E}XZ &= \lim_{n \rightarrow \infty} \mathbb{E}XZ_n && \text{(DCT)} \\ &= \lim_{n \rightarrow \infty} \mathbb{E}YZ_n \\ &= \mathbb{E}YZ && \text{(DCT)} \end{aligned}$$

□

(Uniqueness) Note that any two random variables satisfying the conditions of the proposition are versions of the conditional expectation given  $\mathcal{G}$ , which was shown to be unique.

**Lemma 1.2.1** (Conditional expectation as a function). Let  $X, Y : (\Omega, \mathcal{F}) \rightarrow (\mathbb{R}, \mathcal{B}(\mathbb{R}))$ . Then  $Y$  is measurable with respect to  $\sigma(X)$  if and only if there exists a Borel-measurable function  $f : \mathbb{R} \rightarrow \mathbb{R}$  such that  $Y(\omega) = f(X(\omega))$  for all  $\omega \in \Omega$ .

**Proposition 1.3** (Properties of conditional expectation). All (in)equality relations below hold almost surely.

1. If  $X \geq 0$  a.s., then  $\mathbb{E}(X | \mathcal{G}) \geq 0$
2. If  $X$  and  $\mathcal{G}$  are independent, then  $\mathbb{E}(X | \mathcal{G}) = \mathbb{E}[X]$
3. If  $\alpha, \beta \in \mathbb{R}$  and  $X_1, X_2 \in L^1$ , then

$$\mathbb{E}(\alpha X_1 + \beta X_2 | \mathcal{G}) = \alpha \mathbb{E}(X_1 | \mathcal{G}) + \beta \mathbb{E}(X_2 | \mathcal{G}).$$

4. *Tower property:* If  $\mathcal{H} \subseteq \mathcal{G}$ , then

$$\mathbb{E}(\mathbb{E}(X | \mathcal{G}) | \mathcal{H}) = \mathbb{E}(X | \mathcal{H}) = \mathbb{E}(\mathbb{E}(X | \mathcal{H}) | \mathcal{G})$$

5. If  $Z$  is bounded and  $\mathcal{G}$ -measurable, then

$$\mathbb{E}(ZX | \mathcal{G}) = Z\mathbb{E}(X | \mathcal{G}).$$

6. Let  $X \in L^1$  and  $\mathcal{G}, \mathcal{H} \subseteq \mathcal{F}$ . Assume that  $\sigma(X, \mathcal{G})$  is independent of  $\mathcal{H}$ . Then

$$\mathbb{E}(X | \mathcal{G}) = \mathbb{E}(X | \sigma(\mathcal{G}, \mathcal{H})).$$

*Proof.* 1. Follows from the proof of existence and uniqueness of conditional expectation, or just use monotonicity.

2. Let  $A \in \mathcal{G}$ . Then  $\mathbb{E}(\mathbb{E}(X) \mathbf{1}_A) = \mathbb{E}X \mathbb{E} \mathbf{1}_A = \mathbb{E}(X \mathbf{1}_A)$

3. Use linearity of conditional expectation.

4. For the left equality, let  $A \in \mathcal{H}$ . Then  $\mathbb{E}[\mathbb{E}(\mathbb{E}(X | \mathcal{G}) | \mathcal{H}) \mathbf{1}_A] = \mathbb{E}[\mathbb{E}(X | \mathcal{G}) \mathbf{1}_A] = \mathbb{E}(X \mathbf{1}_A)$ . For the right equality, note that  $\mathbb{E}(X | \mathcal{H})$  is  $\mathcal{H}$ -measurable

5. Easy if  $Z$  is an indicator function. Then use linearity and convergence theorems.

6. Note  $\mathbb{E}(X | \mathcal{G})$  is  $\sigma(\mathcal{G}, \mathcal{H})$ -measurable and  $\sigma(\mathcal{G}, \mathcal{H})$  is generated by the  $\pi$ -system  $\{A \cap B : A \in \mathcal{G}, B \in \mathcal{H}\}$ . We show that  $\mathbb{E}(X | \mathcal{G})$  satisfies the defining property of  $\mathbb{E}(X | \sigma(\mathcal{G}, \mathcal{H}))$ . Let  $A \in \mathcal{G}$  and  $B \in \mathcal{H}$ . Then for any element of the  $\pi$ -system, we have

$$\mathbb{E}(\mathbb{E}(X | \mathcal{G}) \mathbf{1}_{A \cap B}) = \mathbb{E}[\mathbb{E}(X | \mathcal{G}) \mathbf{1}_A \mathbf{1}_B] = \mathbb{E}[\mathbb{E}(X \mathbf{1}_A | \mathcal{G}) \mathbf{1}_B] = \mathbb{E}(\underbrace{X \mathbf{1}_A}_{\in \sigma(\mathcal{G}, X)}) \mathbb{E}(\mathbf{1}_B) = \mathbb{E}(X \mathbf{1}_{A \cap B})$$

Since finite measures extend uniquely from  $\pi$ -systems, the above holds if  $A \cap B$  is replaced by any element of  $\sigma(\mathcal{G}, \mathcal{H})$

□

**Proposition 1.4** (Properties of conditional expectation). All (in)equality relations below hold almost surely.

1. *Jensen's inequality:* If  $c : \mathbb{R} \rightarrow \mathbb{R}$  is convex, then

$$\mathbb{E}(c(X) | \mathcal{G}) \geq c(\mathbb{E}(X | \mathcal{G})).$$

2. *Conditional expectation is a contraction* For  $p \geq 1$ ,

$$\|\mathbb{E}(X | \mathcal{G})\|_p \leq \|X\|_p.$$

3. *Monotone convergence theorem* Suppose  $X_n \uparrow X$  is a sequence of non-negative random variables. Then

$$\mathbb{E}(X_n \mid \mathcal{G}) \uparrow \mathbb{E}(X \mid \mathcal{G}).$$

4. *Fatou's lemma*: If  $X_n$  are non-negative measurable, then

$$\mathbb{E}\left(\liminf_{n \rightarrow \infty} X_n \mid \mathcal{G}\right) \leq \liminf_{n \rightarrow \infty} \mathbb{E}(X_n \mid \mathcal{G}).$$

5. *Dominated convergence theorem*: If  $X_n \rightarrow X$  and  $Y \in L^1$  such that  $Y \geq |X_n|$  for all  $n$ , then

$$\mathbb{E}(X_n \mid \mathcal{G}) \rightarrow \mathbb{E}(X \mid \mathcal{G}).$$

*Proof.* 1. Note that a convex function is the supremum of countably many affine functions  $c(x) = \sup_{i \in I} a_i x + b_i$ . Then

$$\begin{aligned} \mathbb{E}(c(X) \mid \mathcal{G}) &= \mathbb{E}\left(\sup_{i \in I} (a_i X + b_i) \mid \mathcal{G}\right) \\ &\geq \mathbb{E}(a_i X + b_i \mid \mathcal{G}) \quad \forall i \in I \end{aligned} \quad (\text{monotonicity})$$

So  $\mathbb{E}(c(X) \mid \mathcal{G}) \geq \sup_{i \in I} \mathbb{E}(a_i X + b_i \mid \mathcal{G}) = c(\mathbb{E}(X \mid \mathcal{G}))$ .

2. Jensen

3. By monotonicity,  $\mathbb{E}(X_n \mid \mathcal{G}) \uparrow Y$  for some  $Y$ . By the usual monotone convergence theorem,  $\mathbb{E}\mathbb{E}(X_n \mid \mathcal{G}) = \mathbb{E}X_n \rightarrow \mathbb{E}Y \leq \mathbb{E}X$  so  $Y \in L^1$ . Since each of the  $\mathbb{E}(X_n \mid \mathcal{G})$  are  $\mathcal{G}$ -measurable, so is  $Y$ . Finally, for any  $A \in \mathcal{G}$ ,

$$\begin{aligned} \mathbb{E}Y \mathbf{1}_A &= \lim_{n \rightarrow \infty} \mathbb{E}\mathbb{E}(X_n \mid \mathcal{G}) \mathbf{1}_A && (\text{MCV}) \\ &= \lim_{n \rightarrow \infty} \mathbb{E}X_n \mathbf{1}_A \\ &= \mathbb{E}X \mathbf{1}_A && (\text{MCV}) \end{aligned}$$

4.

$$\begin{aligned} \mathbb{E}\left(\liminf_{n \rightarrow \infty} X_n \mid \mathcal{G}\right) &= \mathbb{E}\left(\lim_{n \rightarrow \infty} \underbrace{\inf_{m \geq n} X_m}_{\text{increasing}} \mid \mathcal{G}\right) \\ &= \lim_{n \rightarrow \infty} \mathbb{E}\left(\inf_{m \geq n} X_m \mid \mathcal{G}\right) && (\text{MCV}) \\ &= \liminf_{n \rightarrow \infty} \mathbb{E}\left(\underbrace{\inf_{m \geq n} X_m}_{\leq X_n} \mid \mathcal{G}\right) \\ &\leq \liminf_{n \rightarrow \infty} \mathbb{E}(X_n \mid \mathcal{G}) && (\text{monotonicity}) \end{aligned}$$

5. Use Fatou's lemma on  $Y + X_n$  and  $Y - X_n$ .

□

## 2 Martingales

### 2.1 Definition and Properties

**Definition** ((Discrete) stochastic process). A *stochastic process* (in discrete time) is a collection of random variables  $(X_n)_{n \in \mathbb{N}}$ . A stochastic process is *integrable* if  $X_n \in L^1$  for all  $n$ .

**Definition** (Filtration). A *filtration* is a sequence of  $\sigma$ -algebras  $\mathcal{F}_n \subseteq \mathcal{F}$  such that  $\mathcal{F}_n \subseteq \mathcal{F}_{n+1}$  for all  $n$ . We define  $\mathcal{F}_\infty = \sigma(\bigcup_{n=1}^\infty \mathcal{F}_n)$ . The *natural filtration* of a stochastic process  $X$  is the filtration  $\mathcal{F}_n = \sigma(X_1, \dots, X_n)$ . A stochastic process is *adapted* to a filtration  $\mathcal{F}_n$  if  $X_n$  is  $\mathcal{F}_n$ -measurable for all  $n$ .

**Definition** (Martingale). An integrable adapted process  $(X_n)_{n \geq 0}$  is a *martingale* if for all  $n \geq m$ , we have

$$\mathbb{E}(X_n \mid \mathcal{F}_m) = X_m.$$

We say it is a *super-martingale* if

$$\mathbb{E}(X_n \mid \mathcal{F}_m) \leq X_m,$$

and a *sub-martingale* if

$$\mathbb{E}(X_n \mid \mathcal{F}_m) \geq X_m,$$

By the tower property, it is sufficient to check the martingale property for  $n = m + 1$ .

**Remark.** The definition can be adapted for any *totally ordered* index set  $T$ , such as  $\mathbb{R}_+$  or  $\mathbb{N}_-$ .

**Theorem 2.1** (Doob decomposition, non-examinable). *Let  $X_n$  be an integrable adapted process. Then there exists a martingale  $M_n$  and an integrable predictable process  $A_n$  such that  $X_n = M_n + A_n$  and  $A_0 = 0$ , where predictable means that  $A_n$  is  $\mathcal{F}_{n-1}$ -measurable for all  $n \geq 1$ . Moreover,  $M_n$  and  $A_n$  are unique up to a.s. equivalence.*

*Proof.* (Existence) Add up the ‘known’ bits to get  $A$  and the ‘surprises’ to get  $M$ . Formally,

$$\begin{aligned} A_n &= A_{n-1} + \mathbb{E}(X_n \mid \mathcal{F}_{n-1}) - X_{n-1} \\ M_n &= M_{n-1} + \underbrace{X_n - \mathbb{E}(X_n \mid \mathcal{F}_{n-1})}_{\text{surprise}} \end{aligned}$$

(Uniqueness) Let  $X_n = M_n + A_n = M'_n + A'_n$ . Then  $M_n - M'_n = A'_n - A_n$  is  $\mathcal{F}_{n-1}$ -measurable. But  $M_n - M'_n$  is a martingale, so  $\mathbb{E}(M_n - M'_n \mid \mathcal{F}_{n-1}) = 0$  so  $M_n = M'_n$  almost surely. Similarly,  $A_n = A'_n$  almost surely.  $\square$

**Definition** (Stopping time). A random variable  $T : \Omega \rightarrow \mathbb{N} \cup \{\infty\}$  is a *stopping time* if  $\{T \leq n\} \in \mathcal{F}_n$  for all  $n$ .

In the discrete case, we can equivalently require that  $\{T = n\} \in \mathcal{F}_n$  for all  $n$ .

**Definition** ( $X_T$ ). Let  $X$  be a stochastic process and  $T$  a stopping time. Then  $X_T : \Omega \rightarrow \mathbb{R}$  is defined by cases

$$X_T(\omega) = \begin{cases} X_n(\omega) & T(\omega) = n \\ 0 & T(\omega) = \infty \end{cases}$$

**Definition** (Stopped  $\sigma$ -algebra). Let  $T$  be a stopping time. Then the *stopped  $\sigma$ -algebra* is

$$\mathcal{F}_T = \{A \in \mathcal{F} : A \cap \{T \leq n\} \in \mathcal{F}_n \text{ for all } n\}.$$

**Example.** Let  $N = \#$  of times a random walk hits -5 before it first hits 10 and  $T$  be the first time the random walk hits 10.  $N$  is  $\mathcal{F}_T$ -measurable

**Definition** (Stopped process). Let  $X$  be a stochastic process and  $T$  a stopping time. Then the *stopped process* is  $X_n^T = X_{T \wedge n}$

**Proposition 2.1.**

1. If  $T, S, (T_n)_{n \geq 0}$  are all stopping times, then

$$T \vee S, T \wedge S, \sup_n T_n, \inf_n T_n, \limsup T_n, \liminf T_n$$

are all stopping times.

2.  $\mathcal{F}_T$  is a  $\sigma$ -algebra
3. If  $S \leq T$ , then  $\mathcal{F}_S \subseteq \mathcal{F}_T$ .
4.  $X_T \mathbf{1}_{T < \infty}$  is  $\mathcal{F}_T$ -measurable.
5. If  $(X_n)$  is an adapted process, then so is  $(X_n^T)_{n \geq 0}$  for any stopping time  $T$ .
6. If  $(X_n)$  is an integrable process, then so is  $(X_n^T)_{n \geq 0}$  for any stopping time  $T$ .

*Proof.*

1. Elementary
2. Elementary
3. Let  $A \in \mathcal{F}_S$ . For any  $n$ , we have  $A \cap \{S \leq n\} \in \mathcal{F}_n$  and  $A \cap \{T \leq n\} = A \cap \{S \leq n\} \cap \{T \leq n\} \in \mathcal{F}_n$ .
4.  $X_T \mathbf{1}_{T < \infty} = \sum_{n=1}^{\infty} X_n \mathbf{1}_{\{T=n\}}$  where each of the terms is  $\mathcal{F}_T$ -measurable.
5.  $X_n^T = X_n \mathbf{1}_{\{T \geq n\}} + X_T \mathbf{1}_{\{T < n\}} \mathbf{1}_{\{T < \infty\}}$ .
6.  $X_n^T = X_n \mathbf{1}_{\{T \geq n\}} + \sum_{k=1}^{n-1} X_k \mathbf{1}_{\{T=k\}}$  so  $E|X_n^T| \leq E|X_n| + \sum_{k=1}^{n-1} E|X_k| < \infty$ .

□

**Theorem 2.2** (Equivalent definitions for super-martingales). *Let  $(X_n)_{n \geq 0}$  be an integrable and adapted process. Then the following are equivalent:*

1.  $(X_n)_{n \geq 0}$  is a super-martingale.

2. For any bounded stopping time  $T$  and any stopping time  $S$ ,

$$\mathbb{E}(X_T \mid \mathcal{F}_S) \leq X_{S \wedge T}.$$

3.  $(X_n^T)$  is a super-martingale for any stopping time  $T$ .

4. For bounded stopping times  $S, T$  such that  $S \leq T$ , we have

$$\mathbb{E}X_T \leq \mathbb{E}X_S.$$

*Proof.* – (2)  $\Rightarrow$  (1): Let  $n \geq m$  and set  $T = n, S = m$ .

– (2)  $\Rightarrow$  (4): Tower rule

– (2)  $\Rightarrow$  (3): Let  $n \geq m$

$$\mathbb{E}(X_n^T \mid \mathcal{F}_m) = \mathbb{E}(X_{T \wedge n} \mid \mathcal{F}_m) \leq X_{T \wedge m \wedge n} = X_m^T.$$

– (1)  $\Rightarrow$  (2) Let  $T \leq N$

$$X_T = X_{S \wedge T} + \sum_{k=0}^N (X_{k+1} - X_k) \mathbf{1}_{S \leq k < T} \quad (*)$$

Let  $A \in \mathcal{F}_S$ .

$$\begin{aligned} \mathbb{E}[(X_{k+1} - X_k) \mathbf{1}_{S \leq k < T} \mathbf{1}_A] &= \mathbb{E} \left[ \mathbb{E} \left[ (X_{k+1} - X_k) \underbrace{\mathbf{1}_{S \leq k < T} \mathbf{1}_A}_{\in \mathcal{F}_k} \mid \mathcal{F}_k \right] \right] \\ &= \mathbb{E} \left[ \mathbf{1}_{S \leq k < T} \mathbf{1}_A \underbrace{\mathbb{E}[(X_{k+1} - X_k) \mid \mathcal{F}_k]}_{\leq 0} \right] \\ &\leq 0 \end{aligned}$$

so  $\mathbb{E}X_T \mathbf{1}_A \leq \mathbb{E}X_{S \wedge T} \mathbf{1}_A$ . By Radon-Nikodym,  $\mathbb{E}(X_{S \wedge T} - X_T \mid \mathcal{F}_S) \geq 0$ . But  $X_{S \wedge T}$  is  $\mathcal{F}_S$ -measurable, so  $X_{S \wedge T} - X_T \geq 0$  almost surely.

– (4)  $\Rightarrow$  (2) Let  $n \geq m$  and  $A \in \mathcal{F}_m$ . One can check that  $T = m \mathbf{1}_A + n \mathbf{1}_{A^c} \leq n$  is a stopping time such that

$$\mathbb{E}((X_n - X_m) \mathbf{1}_A) = \mathbb{E}(X_n - X_T) \leq 0$$

By Radon-Nikodym,  $\mathbb{E}(X_m - X_n \mid \mathcal{F}_m) \geq 0$  so  $\mathbb{E}(X_n \mid \mathcal{F}_m) \leq X_m$ .

– (3)  $\Rightarrow$  (1) Let  $T = \infty$

□

**Proposition 2.2** (Convex transformations of martingales). Let  $(X_n)$  be a martingale and  $c : \mathbb{R} \rightarrow \mathbb{R}$  a convex function. Then  $(c(X_n))$  is a sub-martingale.

*Proof.* Let  $S \leq T$  be bounded stopping times. Then

$$\begin{aligned}\mathbb{E}(c(X_T) \mid \mathcal{F}_S) &\geq c(\mathbb{E}(X_T \mid \mathcal{F}_S)) \\ &= c(X_S)\end{aligned}\tag{Jensen}$$

□

**Theorem 2.3** (Optional stopping). *Let  $(X_n)_{n \geq 0}$  be a martingale and  $T$  a stopping time. Then  $E(X_T) = E(X_0)$  if any of the following conditions hold:*

1.  *$T$  is almost surely bounded, i.e. there is some  $N$  such that  $T \leq N$  almost surely.*
2.  *$X$  has bounded increments, i.e. there is some  $K$  such that  $|X_{n+1} - X_n| \leq K$  for all  $n$  almost surely and  $T$  is integrable*
3. *There exists an integrable random variable  $Y$  such that  $|X_n| \leq Y$  for all  $n$  almost surely and  $T$  is finite almost surely, i.e.  $\mathbb{P}(T < \infty) = 1$ .*

*Proof.* 1. Use (4) of the previous theorem with  $S = 0$ , or prove directly.

2.

$$\begin{aligned}X_{T \wedge n} &= \sum_{i=1}^{T \wedge n} (X_i - X_{i-1}) + X_0 \\ &= \sum_{i=1}^n (X_i - X_{i-1}) \mathbf{1}_{T \geq i} + X_0\end{aligned}$$

Note  $|X_{T \wedge n}| \leq |X_0| + \sum_{i=1}^n |X_i - X_{i-1}| \mathbf{1}_{T \geq i} \leq |X_0| + KT \in L^1$ . Note that  $T \wedge n \rightarrow T$  almost surely as  $T < \infty$  almost surely. Hence,  $X_{T \wedge n} \rightarrow X_T$  almost surely and  $\mathbb{E}X_{T \wedge n} \rightarrow \mathbb{E}X_T$  by dominated convergence theorem. By the previous case,  $\mathbb{E}X_{T \wedge n} = \mathbb{E}X_0$  so  $\mathbb{E}X_T = \mathbb{E}X_0$ .

3. Using the same reasoning,  $X_{T \wedge n} \rightarrow X_T$  almost surely and  $\mathbb{E}|X_{T \wedge n}| \leq \mathbb{E}Y$  so dominated convergence holds.

□

## 2.2 Convergence

**Definition** (Upcrossing). Let  $(x_n)$  be a sequence and  $(a, b)$  an interval. An *upcrossing* of  $(a, b)$  by  $(x_n)$  is a sequence  $j, j+1, \dots, k$  such that  $x_j \leq a$  and  $x_k \geq b$ . We define

$$\begin{aligned}U_n[a, b, (x_n)] &= \text{number of disjoint upcrossings contained in } \{1, \dots, n\} \\ U[a, b, (x_n)] &= \lim_{n \rightarrow \infty} U_n[a, b, (x_n)].\end{aligned}$$



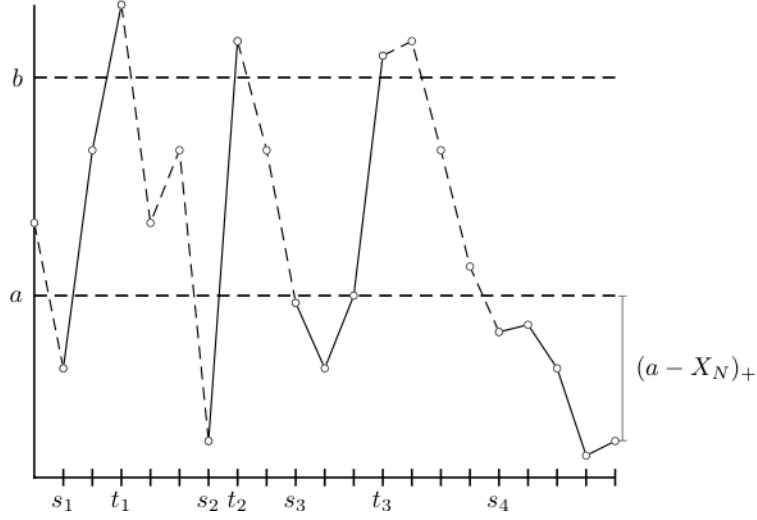


Figure 1: Three upcrossings. [Source](#)

The notion of upcrossings is related to the notion of convergence.

**Lemma 2.2.1** (Upcrossing and convergence). A sequence  $(x_n)$  converges to a limit in the extended real numbers if and only if  $U[a, b, (x_n)] < \infty$  for all rationals  $a < b$ .

**Proposition 2.3** (Doob's upcrossing inequality). Let  $X = (X_k)$  be a super-martingale and  $a < b$ . Then

$$(b - a)\mathbb{E}U_n[a, b, X] \leq \mathbb{E}[(X_n - a)^-] \leq \mathbb{E}(|X_n|) + |a|$$

*Proof.* Assume that  $X$  is a super-martingale. We define stopping times  $S_k, T_k$  as follows:

- $T_0 = 0$
- $S_{k+1} = \inf\{n : X_n \leq a, n \geq T_n\}$
- $T_{k+1} = \inf\{n : X_n \geq b, n \geq S_{k+1}\}.$

Note that the times alternate, i.e.  $S_k \leq T_k \leq S_{k+1} \leq T_{k+1}$

The idea is to sum up the increments of each upcrossing. Consider the set  $\mathcal{I} := \{k : S_k < T_k < n\}$  and the sum

$$\underbrace{\sum_{k \in \mathcal{I}} (X_{T_k} - X_{S_k})}_{\geq (b-a)U_n[a, b, X]} + \begin{cases} X_n - X_{S_{\max \mathcal{I}+1}} & \text{if } n > S_{\max \mathcal{I}+1} \\ 0 & \text{otherwise} \end{cases}$$

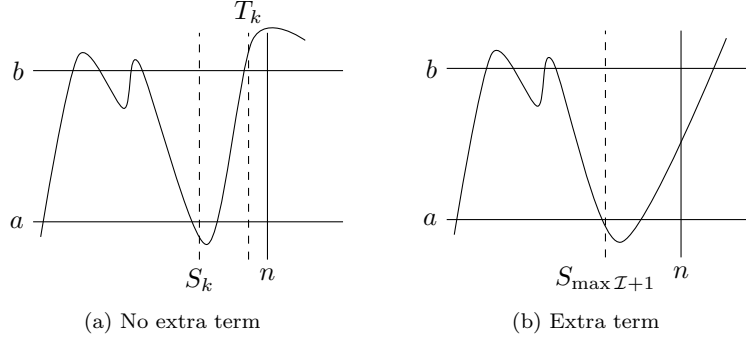


Figure 2: Graphs taken from [here](#)

Note that each term is of the form  $X_U - X_V$  for some *bounded* stopping times  $U \geq V$ . By the property of super-martingales,  $\mathbb{E}(X_U - X_V) \leq 0$  so the sum is negative in expectation.

Hence,

$$\begin{aligned}
 (b-a)\mathbb{E}U_n[a, b, X] + \underbrace{\mathbb{E}((X_n - X_{S_{\max T+1}}) \mathbf{1}_{n > S_{\max T+1}})}_{\substack{\geq X_n - a \geq -(X_n - a)^- \\ \geq -(X_n - a)^-}} &\leq 0 \\
 (b-a)\mathbb{E}U_n[a, b, X] &\leq \mathbb{E}((X_n - a)^-) \leq \mathbb{E}(|X_n|) + |a|
 \end{aligned}$$

□

**Corollary 1.** Under the same assumptions of the previous proposition, by the monotone convergence theorem and noting that  $U_n[a, b, X] \uparrow U[a, b, X]$ ,

$$(b-a)\mathbb{E}U[a, b, X] \leq \sup_n \mathbb{E}[(X_n - a)^-] \leq \sup_n \mathbb{E}(|X_n|) + |a|$$

**Proposition 2.4** (Almost sure martingale convergence theorem). Suppose  $X = (X_n)_{n \geq 0}$  is a super-martingale that is bounded in  $L^1$ , i.e.  $\sup_n \mathbb{E}|X_n| < \infty$ . Then for any  $a < b$ , we have  $U[a, b, X] < \infty$  almost surely. In particular, there exists an  $\mathcal{F}_\infty$ -measurable  $X_\infty \in L^1$  such that

$$X_n \rightarrow X_\infty \text{ a.s. as } n \rightarrow \infty.$$

**Remark.** The intuition is that you cannot make money by betting on a super-martingale (without shorting). For any  $a < b$ , you can devise a betting strategy where you buy at  $a$  and sell at  $b$ . If  $U[a, b, X] = \infty$ , then you make money almost surely.

*Proof.* Let  $a < b$  and  $n$  be arbitrary and  $\sup_m \mathbb{E}|X_m| = M$ . By the previous corollary, So  $U[a, b, X] < M + |a| < \infty$  almost surely.

Now consider the set of events where  $X_n$  converges

$$A := \cap_{a,b \in \mathbb{Q}} \{U[a, b, X] < \infty\}, \quad \mathbb{P}(A) = 1$$

and define

$$X_\infty = \begin{cases} \lim X_n & \text{on } A \\ 0 & \text{on } A^c \end{cases}$$

which is  $\mathcal{F}_\infty$ -measurable. Note that  $|X_\infty| = \liminf |X_n|$  almost surely and by Fatou's lemma,

$$\mathbb{E} \liminf |X_n| \leq \liminf \mathbb{E} |X_n| \leq M < \infty.$$

So  $X_\infty \in L^1$ . □

**Remark.** Some propositions below concern non-negative submartingales. Recall convex transformations, such as  $|\cdot|$ , turn a martingale into a sub-martingale.

**Proposition 2.5** (Doob's maximal inequality). Let  $X = (X_n)_{n \geq 0}$  be a non-negative sub-martingale. Then for any  $\lambda > 0$  and writing  $X_n^* = \max_{0 \leq k \leq n} X_k$ , we have

$$\lambda \mathbb{P}(X_n^* \geq \lambda) \leq \mathbb{E} X_n \mathbf{1}_{X_n^* \geq \lambda} \leq \mathbb{E} X_n.$$

*Proof.* Let  $T = \inf\{n : X_n \geq \lambda\}$

$$\begin{aligned} \mathbb{E} X_n &\geq \mathbb{E} X_{T \wedge n} = \mathbb{E} X_n \mathbf{1}_{T > n} + \mathbb{E} X_T \mathbf{1}_{T \leq n} \\ \mathbb{E} X_n \mathbf{1}_{T \leq n} &\geq \lambda \mathbb{P}(T \leq n) \\ \mathbb{E} X_n \mathbf{1}_{X_n^* \geq \lambda} &\geq \lambda \mathbb{P}(X_n^* \geq \lambda) \end{aligned}$$

□

**Corollary 2.** Under the same hypotheses, let  $X_n^* \uparrow X^*$ . Then

$$\lambda \mathbb{P}(X^* \geq \lambda) \leq \sup_{n \geq 0} \mathbb{E} X_n$$

**Proposition 2.6** (Doob's  $L^p$  inequality). Let  $X = (X_n)_{n \geq 0}$  be a martingale or a non-negative sub-martingale. Then for any  $p > 1$  and  $n \geq 1$ , we have

$$\|X_n^*\|_p \leq \frac{p}{p-1} \|X_n\|_p.$$

*Proof.* (magic)

Let  $k > 0$  and  $T = \inf\{n : X_n \geq k\}$ . Then

$$\begin{aligned}
\int |X_n^* \wedge k|^p \, d\mathbb{P} &= \int \left( \int p x^{p-1} \mathbf{1}_{\{0 \leq x \leq |X_n^* \wedge k|\}} \, dx \right) d\mathbb{P} \\
&= \int \left( \int p x^{p-1} \mathbf{1}_{\{|X_n^*| \geq x\}} \mathbf{1}_{\{0 \leq x \leq k\}} \, dx \right) d\mathbb{P} \\
&= \int p x^{p-1} \mathbb{P}(|X_n^*| \geq x) \mathbf{1}_{\{0 \leq x \leq k\}} \, dx \\
&\leq \int p x^{p-1} \frac{1}{x} \mathbb{E}|X_n| \mathbf{1}_{\{|X_n^*| \geq x\}} \mathbf{1}_{\{0 \leq x \leq k\}} \, dx && \text{(Doob's maximal inequality)} \\
&= \int \int p x^{p-2} |X_n| \mathbf{1}_{\{|X_n^*| \geq x\}} \mathbf{1}_{\{0 \leq x \leq k\}} \, dx \, d\mathbb{P} \\
&= \int \frac{p}{p-1} |X_n| (|X_n^*| \wedge k)^{p-1} \, d\mathbb{P} \\
&\leq \frac{p}{p-1} \|X_n\|_p \|X_n^* \wedge k\|_p^{p-1} && \text{(Hölder's inequality)}
\end{aligned}$$

By monotone convergence and taking  $k \rightarrow \infty$ , we get

$$\|X_n^*\|_p^p \leq \frac{p}{p-1} \|X_n\|_p \|X_n^*\|_p^{p-1}$$

and the result follows. □

**Corollary 3.** Under the same hypotheses, let  $X_n^* \uparrow X^*$ . Then

$$\|X^*\|_p \leq \frac{p}{p-1} \sup_{n \geq 0} \|X_n\|_p$$

**Proposition 2.7** (Equivalent conditions for  $L^p$  convergence,  $p > 1$ ). Let  $X = (X_n)_{n \geq 0}$  be a martingale, and  $p > 1$ . Then the following are equivalent:

1.  $(X_n)_{n \geq 0}$  is bounded in  $L^p$ , i.e.  $M = \sup_n \mathbb{E}|X_n|^p < \infty$ .
2.  $(X_n)_{n \geq 0}$  converges as  $n \rightarrow \infty$  to a random variable  $X_\infty \in L^p$  almost surely and in  $L^p$ .
3. There exists a random variable  $Z \in L^p$  such that

$$X_n = \mathbb{E}(Z \mid \mathcal{F}_n) \quad \lim_{n \rightarrow \infty} X_n = \mathbb{E}(Z \mid \mathcal{F}_\infty) \text{ a.s.}$$

This gives a bijection between martingales bounded in  $L^p$  and  $L^p(\mathcal{F}_\infty)$ , sending  $(X_n)_{n \geq 0} \mapsto X_\infty$ .

*Proof.* – (1)  $\Rightarrow$  (2) By Jensen,  $\mathbb{E}|X_n|^p \geq (\mathbb{E}|X_n|)^p$  so  $X$  is bounded in  $L^1$ . By the almost sure martingale convergence theorem, there exists an  $\mathcal{F}_\infty$ -measurable  $X_\infty \in L^1$  such that  $X_n \rightarrow X_\infty$  almost surely. Note also  $|X|^* \geq |X_n|$  for all  $n$  and  $X^* \in L^p$  by Corollary 3. Hence,  $X_n \rightarrow X_\infty$  in  $L^p$  by  $L^p$  dominated convergence

- (2)  $\Rightarrow$  (3) Let  $Z = X_\infty$ . Note  $X_n \xrightarrow{L^p} X_\infty$  implies that  $X$  is bounded in  $L^p$ . By Doob's maximal inequality,  $|X|^* \in L^p \subseteq L^1$ . By conditional dominated convergence,

$$\mathbb{E}(X_\infty | \mathcal{F}_n) = \lim_{m \rightarrow \infty} \mathbb{E}(X_m | \mathcal{F}_n) = X_n$$

- (3)  $\Rightarrow$  (1) Conditional expectation is a contraction (Proposition 1.4)

□

**Definition** (Closed martingale). A martingale in the form  $X_n = \mathbb{E}(Z | \mathcal{F}_n)$  for some  $Z \in L^p$  is called a *martingale closed in  $L^p$*

**Definition** (Non-examinable, uniform integrability for general measure space). Let  $(f_n)_n$  be a family of absolutely integrable functions on some measure space. The family is said to be *uniformly integrable (UI)* if all of the following hold

1. Uniform bound on  $L^1$  norm ( $\sup_n \int |f_n| < \infty$ )
2. No escape to vertical infinity ( $\sup_n \int_{\{|f_n| > \lambda\}} |f_n| \rightarrow 0$  as  $\lambda \rightarrow \infty$ )
3. No escape to horizontal infinity (for any  $\varepsilon > 0$ , there exists a finite measure subset  $A$  such that  $\sup_n \int_{A^c} |f_n| < \varepsilon$ )

**Example.** A single integrable function is uniformly integrable.

**Example.** A family of functions which is dominated by some integrable function, i.e. there is  $g \in L^1$  such that  $|f_i| \leq g$  for all  $i \in \mathcal{I}$ , is uniformly integrable.

**Definition** (Uniform integrability for random variables). A family of random variables  $(X_i)_{i \in \mathcal{I}}$  is *uniformly integrable* if

$$\sup_{i \in \mathcal{I}} \mathbb{E}(|X_i| \mathbf{1}_{|X_i| > \alpha}) \rightarrow 0 \text{ as } \alpha \rightarrow \infty.$$

**Remark.** In the finite measure case, functions cannot escape to horizontal infinity and the conditions simplifies just to no escape to vertical infinity.

**Proposition 2.8** (Equivalent definition of uniform integrability). A family of random variables  $(X_i)_{i \in \mathcal{I}}$  is *uniformly integrable* if and only if both of the following hold

1. It is bounded in  $L^1$  ( $\sup_{i \in \mathcal{I}} \mathbb{E}(|X_i|) < \infty$ )
2. It is equi-integrable (For any  $\varepsilon > 0$ , there exists a  $\delta > 0$  such that for any  $A \in \mathcal{F}$  with  $\mathbb{P}(A) < \delta$ , we have

$$\mathbb{E}(|X_i| \mathbf{1}_A) < \varepsilon.$$

for all  $i \in \mathcal{I}$ .)

**Proposition 2.9** (Uniform integrability by domination). Let  $\{Y_j : j \in J\}$  be uniformly integrable. Suppose the set  $X = \{X_i : i \in I\}$  satisfies for any  $i \in I$ , there exists  $j \in J$  such that  $|X_i| \leq Y_j$ . Then  $X$  is uniformly integrable.

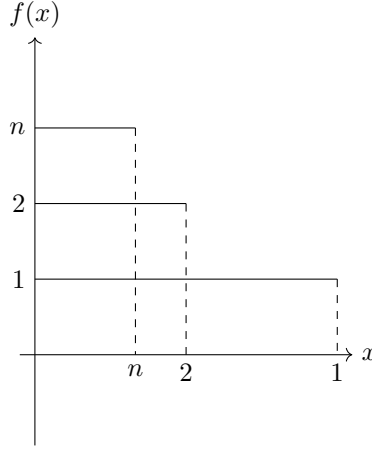


Figure 3: Non uniformly integrable sequence

**Theorem 2.4** (Non-examinable, Vitali convergence theorem). *Let  $f_1, f_2, \dots$  be a sequence of integrable functions on some measure space and  $1 \leq p < \infty$ . Then  $f_n \xrightarrow{L^p} f$  for some measurable  $f$  if and only if all of the following hold*

1.  $(f_n^p)$  is uniformly integrable
2.  $f_n \rightarrow f$  in measure
3. The sequence cannot escape to horizontal infinity, i.e. for any  $\varepsilon > 0$ , there exists a finite measure subset  $A$  such that  $\sup_n \int_{A^c} |f_n|^p < \varepsilon$ .

**Remark.** In the finite measure case, the third condition is trivially true and almost sure convergence implies convergence in measure, so this implies the  $L^p$  dominated convergence theorem.

**Proposition 2.10** (Conditional expectations are uniformly integrable). Let  $\mathcal{S}$  be a uniformly integrable family of random variables. Then the following set is uniformly integrable

$$\mathcal{S}^* = \{\mathbb{E}(X|\mathcal{G}) \mid X \in \mathcal{S}, \mathcal{G} \text{ is a sub } \sigma\text{-algebra of } \mathcal{F}\}.$$

*Proof.* Since  $\mathcal{S}$  is bounded in  $L^1$ ,  $\mathcal{S}^*$  is bounded in  $L^1$ . Let  $\varepsilon > 0$ . By uniform integrability of  $\mathcal{S}$ , there exists  $\delta > 0$  such that for any  $A \in \mathcal{F}$  with  $\mathbb{P}(A) < \delta$ , we have  $\mathbb{E}(|X|\mathbf{1}_A) < \varepsilon$  for all  $X \in \mathcal{S}$ . Note

$$\mathbb{E}(|\mathbb{E}(X|\mathcal{G})|\mathbf{1}_A) \underbrace{\leq}_{\text{Jensen}} \mathbb{E}(\mathbb{E}(|X||\mathcal{G})\mathbf{1}_A)$$

we wish to show that when  $A$  is of the form  $\{|X| > \alpha\}$ , the right hand side converges to zero as  $\alpha \rightarrow \infty$ . The right hand side becomes  $\mathbb{E}|X|\mathbf{1}_{\{|X|>\alpha\}}$

We want to choose  $\alpha$  such that  $\mathbb{P}(|X| > \alpha) < \delta$ . By Markov's inequality,

$$\mathbb{P}(|X| > \alpha) \leq \frac{\mathbb{E}|X|}{\alpha}$$

so picking any  $\alpha > \frac{\mathbb{E}|X|}{\delta}$  works.

We have shown that for any  $\varepsilon > 0$ , for any  $\alpha$  sufficiently large,  $\mathbb{E}(|\mathbb{E}(X | \mathcal{G})| \mathbf{1}_{\{|X| > \alpha\}}) \leq \varepsilon$   $\square$

**Proposition 2.11** (Equivalent conditions for  $L^1$  convergence). Let  $(X_n)_{n \geq 0}$  be a martingale. Then the following are equivalent:

1.  $(X_n)_{n \geq 0}$  is uniformly integrable.
2.  $(X_n)_{n \geq 0}$  converges to some  $X_\infty \in L^1$  almost surely and in  $L^1$ .
3. There exists  $Z \in L^1$  such that  $X_n = \mathbb{E}(Z | \mathcal{F}_n)$  almost surely.

Moreover,  $X_\infty = \mathbb{E}(Z | \mathcal{F}_\infty)$

*Proof.* – (1)  $\Rightarrow$  (2)  $X$  is  $L^1$  bounded and hence converges to some  $X_\infty$  almost surely by Proposition 2.4. By Vitali,  $X_n \rightarrow X_\infty$  in  $L^1$ .

– (2)  $\Rightarrow$  (3) Let  $Z = X_\infty$ . Then  $X_n = \mathbb{E}(Z | \mathcal{F}_n)$  almost surely by Proposition 1.4. (Same as previous proposition)  
Let  $Z = X_\infty$ .

$$\|X_n - \mathbb{E}(X_\infty | \mathcal{F}_n)\|_1 = \|\mathbb{E}(X_m - X_\infty | \mathcal{F}_n)\|_1 \leq \|X_m - X_\infty\|_1$$

for any  $m \geq n$  and the right hand side converges to 0 by  $L^1$  convergence.

– (3)  $\Rightarrow$  (1) Conditional expectation is uniformly integrable, see previous example.  $\square$

**Lemma 2.11.1** (Stopped UI process). Let  $X$  be a uniformly integrable martingale and  $T$  be any stopping time. Then the following statements about the stopped process  $X^T$  hold:

1.  $X_{T \wedge n} = \mathbb{E}(X_\infty | \mathcal{F}_{T \wedge n})$
2.  $X^T$  is uniformly integrable
3.  $X_n^T \rightarrow X_T$  in  $L^1$  and almost surely

*Proof.* Since  $X$  is UI, we use the fact that the martingale can be represented as a conditional expectation of some  $X_\infty \in L^1$ . Then

$$\begin{aligned} X_{T \wedge n} &= \mathbb{E}(X_n | \mathcal{F}_{T \wedge n}) \\ &= \mathbb{E}(\mathbb{E}(X_\infty | \mathcal{F}_n) | \mathcal{F}_{T \wedge n}) \\ &= \mathbb{E}(X_\infty | \mathcal{F}_{T \wedge n}) \end{aligned}$$

By Proposition 2.10, conditional expectations are uniformly integrable. By Proposition 2.11,  $X_n^T \rightarrow X_\infty^T$  in  $L^1$  and almost surely for some  $X_\infty^T \in L^1$ . By considering different values of  $T$ , one can see that  $X_\infty^T = X_T$  almost surely.  $\square$

**Proposition 2.12** (Optional stopping for arbitrary stopping times). If  $(X_n)_{n \geq 0}$  is a uniformly integrable martingale, and  $S, T$  are arbitrary stopping times, then  $\mathbb{E}(X_T \mid \mathcal{F}_S) = X_{S \wedge T}$ . In particular  $\mathbb{E}X_T = X_0$ .

Note that we are now allowing arbitrary stopping times, so  $T$  may be infinite with non-zero probability. Hence we define

$$X_T = \sum_{n=0}^{\infty} X_n \mathbf{1}_{T=n} + X_{\infty} \mathbf{1}_{T=\infty}.$$

*Proof.* We have proven the result for bounded stopping times. For the stopped process  $X^T = (X_{T \wedge n})_{n \geq 0}$ , we have  $\mathbb{E}(X_n^T \mid \mathcal{F}_S) = X_{S \wedge T \wedge n}$ . What we would like to do is take the limit as  $n \rightarrow \infty$ .

From Lemma 2.11.1,

$$\|\mathbb{E}(X_{T \wedge n} - X_T \mid \mathcal{F}_S)\|_1 \leq \|X_{T \wedge n} - X_T\|_1 \rightarrow 0$$

so  $X_{S \wedge T \wedge n} \xrightarrow{L^1} \mathbb{E}(X_T \mid \mathcal{F}_S)$ .

Lemma 2.11.1 also says  $X_{S \wedge T \wedge n} \xrightarrow{L^1} X_{S \wedge T}$  so  $X_{S \wedge T} = \mathbb{E}(X_T \mid \mathcal{F}_S)$  almost surely □

## 2.3 Applications of martingales

**Definition** (Backwards filtration). A *backwards filtration* on a measurable space  $(E, \mathcal{E})$  is a sequence of  $\sigma$ -algebras  $\hat{\mathcal{F}}_n \subseteq \mathcal{E}$  such that  $\hat{\mathcal{F}}_{n+1} \subseteq \hat{\mathcal{F}}_n$ . We define

$$\hat{\mathcal{F}}_{\infty} = \bigcap_{n \geq 0} \hat{\mathcal{F}}_n.$$

**Definition** (Backwards martingale). An adapted process  $(X_n)_{n \geq 0}$  is a *backwards martingale* with respect to a backwards filtration  $(\hat{\mathcal{F}}_n)_{n \geq 0}$  if all of the following hold

1.  $X_0 \in L^1$
2.  $\mathbb{E}(X_n \mid \hat{\mathcal{F}}_{n+1}) = X_{n+1}$

**Proposition 2.13** (Equivalent definition for backwards martingale).  $(X_n)_{n \geq 0}$  is a backwards martingale if and only if there exists some  $Y \in L^1$  such that  $X_n = \mathbb{E}(Y \mid \hat{\mathcal{F}}_n)$  almost surely.

*Proof.* – (  $\implies$  ) Let  $Y = X_0$ .

– (  $\impliedby$  ) Tower property and conditional expectation being a contraction □

**Remark.** Let  $I = \mathbb{N}$  or  $\mathbb{R}_+$  and  $(X_t)_{t \in I}$  be a backwards martingale with respect to a backwards filtration  $(\hat{\mathcal{F}}_t)_{t \in I}$  and  $s \in I$ . Then  $(X_{s-t})_{0 \leq t \leq s}$  is a martingale with respect to the filtration  $(\mathcal{F}_{s-t})_{0 \leq t \leq s}$ .



**Proposition 2.14** (Backwards martingale convergence). Let  $(X_n)_{n \geq 0}$  be a backwards martingale with  $X_0 \in L^p$  for some  $p \in [0, \infty)$  and  $X_\infty = \mathbb{E}(X_0 | \hat{\mathcal{F}}_\infty)$ . Then  $X_n \rightarrow X_\infty$  in  $L^p$  and almost surely.

*Proof.* Proof is basically the same as the forwards case. Note that the martingale  $(X_{s-t})_{0 \leq t \leq s}$  has the same number of upcrossings on  $[a, b]$  as  $(-X_{s-t})_{0 \leq t \leq s}$  on  $[-b, -a]$ . By noting that  $\|X_t\|_1 \leq \|X_0\|_1$  for all  $t$ , the martingale is bounded in  $L_1$  and as before we have  $X_t \rightarrow X_\infty$  almost surely to some  $X_\infty$  which is also in  $L^1$  by Fatou's lemma.

Using similar arguments, we see that  $X_\infty \in L^p$ .

Next, show  $X_\infty = \mathbb{E}(X_0 | \hat{\mathcal{F}}_\infty)$  almost surely. Pick any  $A \in \hat{\mathcal{F}}_\infty$ . Then  $A \in \hat{\mathcal{F}}_n$  for all  $n$ . So

$$\begin{aligned} \mathbb{E}(X_\infty \mathbf{1}_A) &= \lim_{n \rightarrow \infty} \mathbb{E}(X_n \mathbf{1}_A) && \text{(Vitali convergence theorem)} \\ &= \lim_{n \rightarrow \infty} \mathbb{E}(\mathbb{E}(X_0 | \hat{\mathcal{F}}_n) \mathbf{1}_A) \\ &= \lim_{n \rightarrow \infty} \mathbb{E}(X_0 \mathbf{1}_A) \\ &= \mathbb{E}(X_0 \mathbf{1}_A) \end{aligned}$$

Finally, prove  $X_n \rightarrow X_\infty$  in  $L^p$ . Note that  $(|X_n - X_\infty|^p)_{n \geq 0}$  is UI by 2.9 since

$$|X_n - X_\infty|^p = |\mathbb{E}(X_0 - X_\infty | \hat{\mathcal{F}}_n)|^p \leq \mathbb{E}(|X_0 - X_\infty|^p | \hat{\mathcal{F}}_n)$$

and the set of variables on the RHS is UI. By Vitali,  $X_n \rightarrow X_\infty$  in  $L^p$ . □

### 3 Continuous time stochastic processes

Uncountability bad.

**Definition** (Continuous time stochastic process). A *continuous time stochastic process* is a family of random variables  $(X_t)_{t \geq 0}$  (or  $(X_t)_{t \in [a, b]}$ ).

We can think of a continuous time stochastic process as a function

$$(\omega, t) \mapsto X_t(\omega)$$

In the discrete case, this function is  $(\mathcal{F} \otimes \mathcal{P}(\mathbb{N}))$ -measurable. This is not true in the continuous case, i.e. the function is not necessarily  $\mathcal{F} \otimes \mathcal{B}(\mathbb{R}_+)$ -measurable.

Worse still, the first hitting time of a measurable set  $A$  need not be a stopping time. Let  $T_A = \inf\{t \geq 0 : X_t \in A\}$  so

$$\{T_A \leq t\} = \bigcup_{0 \leq s \leq t} \{X_s \in A\} \notin \mathcal{F}_t$$

since the union is uncountable.

To remedy this, we reduce to the countable case by considering continuous or more generally cadlag processes, which are determined by any dense subset of  $\mathbb{R}_+$ , e.g.  $\mathbb{Q}_+$ .

**Definition (Cadlag).** Let  $(M, d)$  be a metric space and  $E \subseteq \mathbb{R}$ . A function  $f : E \rightarrow M$  is *cadlag* if it is right continuous and has left limits, i.e. for all  $t \geq 0$ ,

$$\lim_{s \rightarrow t^+} f(s) = f(t) \quad \lim_{s \rightarrow t^-} f(s) \text{ exists.}$$

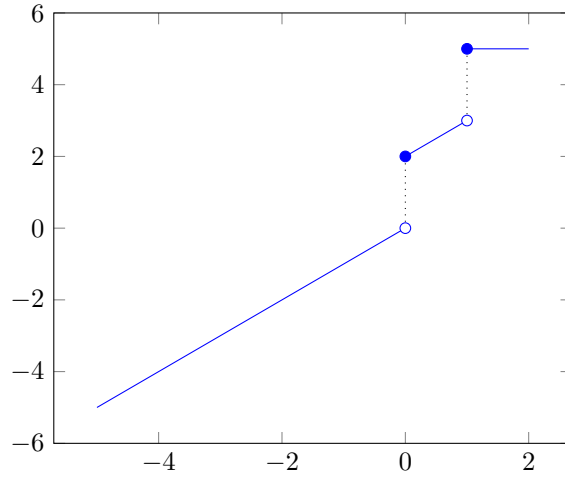


Figure 4: Cadlag function

**Definition (Continuous/cadlag processes).** We say a stochastic process is *continuous* (resp. *cadlag*) if the set of  $\omega \in \Omega$  such the map  $t \mapsto X_t(\omega)$  is continuous (resp. *cadlag*) has probability 1.

**Definition (Continuous time filtration).** A *continuous-time filtration* is a family of  $\sigma$ -algebras  $(\mathcal{F}_t)_{t \geq 0}$  such that  $\mathcal{F}_s \subseteq \mathcal{F}_t \subseteq \mathcal{F}$  if  $s \leq t$ . Define  $\mathcal{F}_\infty = \sigma(\mathcal{F}_t : t \geq 0)$ .

**Proposition 3.1.** Let  $(X_t)_{t \geq 0}$  be a cadlag adapted process and  $S, T$  stopping times. Then

1.  $S \wedge T$  is a stopping time.
2. If  $S \leq T$ , then  $\mathcal{F}_S \subseteq \mathcal{F}_T$ .
3.  $X_T \mathbf{1}_{T < \infty}$  is  $\mathcal{F}_T$ -measurable.
4.  $(X_t^T)_{t \geq 0} = (X_{T \wedge t})_{t \geq 0}$  is adapted.

To show 3., it is useful to have the following characterisation of  $\mathcal{F}_T$  measurability.

**Lemma 3.1.1 (Characterisation of  $\mathcal{F}_T$ -measurability).** A random variable  $Z$  is  $\mathcal{F}_T$ -measurable iff  $Z \mathbf{1}_{\{T \leq t\}}$  is  $\mathcal{F}_t$ -measurable for all  $t \geq 0$ .

*Proof.*  $\implies$  is just by definition. For  $\impliedby$ ,

1. It is true by working with the definition for  $Z = c\mathbf{1}_A$ ,  $A \in \mathcal{F}$
2. It is true for simple functions  $Z$  since measurability is preserved under linear combinations
3. It is true for any measurable  $Z$  since it is a pointwise limit of simple functions and measurability is preserved under limits

□

**Corollary 4.**  $X_T\mathbf{1}_{T=t}$  is  $\mathcal{F}_T$ -measurable.

*Proof.* Let  $s \geq 0$ . Then  $X_T\mathbf{1}_{T=t}\mathbf{1}_{T \leq s} = X_t\mathbf{1}_{s \geq t}$  which is  $\mathcal{F}_s$ -measurable by considering cases. □

**Proposition 3.2.** Let  $(X_t)_{t \geq 0}$  be a cadlag adapted process and  $S, T$  stopping times. Then

1.  $S \wedge T$  is a stopping time.
2. If  $S \leq T$ , then  $\mathcal{F}_S \subseteq \mathcal{F}_T$ .
3.  $X_T\mathbf{1}_{T < \infty}$  is  $\mathcal{F}_T$ -measurable.
4.  $(X_t^T)_{t \geq 0} = (X_{T \wedge t})_{t \geq 0}$  is adapted.

*Proof.* 1 and 2 are same as the discrete case. For 3, we have

$$X_T\mathbf{1}_{T \leq t} = X_T\mathbf{1}_{\{T < t\}} + X_T\mathbf{1}_{\{T=t\}}.$$

The term on the right is  $\mathcal{F}_t$ -measurable by Corollary 4. Let  $(T_n)_{n \geq 0} \subseteq \mathbb{Q}$  be your favourite rational sequence of stopping times with  $T_n \downarrow T$ . Since the process is right continuous,  $X_T$  can be approximated from above, i.e.  $X_T\mathbf{1}_{\{T < t\}} = \lim_{n \rightarrow \infty} X_{T_n \wedge t}\mathbf{1}_{T < t}$ . Moreover,

$$X_{T_n \wedge t} = \sum_{q \in \mathbb{Q}, q \leq t} \underbrace{X_{T_n}\mathbf{1}_{T_n=q}}_{\in \mathcal{F}_{T_n} \subseteq \mathcal{F}_t} + \underbrace{X_t\mathbf{1}_{T_n > t}}_{\in \mathcal{F}_t}$$

is  $\mathcal{F}_t$ -measurable. It follows that  $X_T\mathbf{1}_{\{T < t\}}$  is  $\mathcal{F}_t$ -measurable and the result follows by Lemma 3.1.1.

For 4,  $X_{T \wedge t} \in \mathcal{F}_{T \wedge t} \subseteq \mathcal{F}_t$ . □

**Definition** (Hitting time). Let  $A \in \mathcal{B}(\mathbb{R})$ . Then the hitting time of  $A$  is

$$T_A = \inf\{t \geq 0 : X_t \in A\}$$

**Lemma 3.2.1** (Process at hitting time of a closed set). Let  $A \in \mathcal{B}(\mathbb{R})$  be closed and  $X$  be a *continuous* adapted process. then  $X_{T_A} \in A$

*Proof.* There is a sequence  $t_n \downarrow T_A$  such that  $X_{t_n} \in A$  for all  $n$ . Since  $A$  is closed and  $X_{t_n} \rightarrow X_{T_A}$ , we have  $X_{T_A} \in A$ . □

**Proposition 3.3** (Hitting time of closed set is stopping time). Let  $A \in \mathcal{B}(\mathbb{R})$  be closed and  $X$  be a *continuous* adapted process. Then  $T_A$  is a stopping time if  $T_A < \infty$ .

*Proof.* Let  $t \geq 0$ . Then

$$\begin{aligned} \{T_A \leq t\} &= \{\text{at least one } 0 \leq s \leq t \text{ such that } X_s \in A\} && (X_{T_A} \in A) \\ &= \{\inf_{s \leq t} d(X_s, A) = 0\} && (s \rightarrow d(X_s, A) \text{ is a.s. continuous}) \\ &= \{\inf_{s \leq t, s \in \mathbb{Q}} d(X_s, A) = 0\} && (X \text{ is continuous}) \end{aligned}$$

□

**Definition** (Right continuous filtration). Given a continuous filtration  $(\mathcal{F}_t)_{t \geq 0}$ , we define

$$\mathcal{F}_t^+ = \bigcap_{s > t} \mathcal{F}_s \supseteq \mathcal{F}_t.$$

We say  $(\mathcal{F}_t)_{t \geq 0}$  is *right continuous* if  $\mathcal{F}_t = \mathcal{F}_t^+$ .

**Proposition 3.4** (Hitting time of open set is stopping time for right continuous filtrations). Let  $(X_t)_{t \geq 0}$  be an adapted process (to  $(\mathcal{F}_t)_{t \geq 0}$ ) that is cadlag, and let  $A$  be an open set. Then  $T_A$  is a stopping time with respect to  $\mathcal{F}_t^+$ .

*Proof.*

$$\begin{aligned} \{T_A < t\} &= \bigcup_{s < t, s \in \mathbb{Q}} \{X_s \in A\} \in \mathcal{F}_t \\ \{T_A \leq t\} &= \bigcap_{n \geq 1} \{T_A < t + \frac{1}{n}\} \in \mathcal{F}_t^+ \end{aligned}$$

□

Before proving continuous time analogues of previous theorems, recall the following properties of cadlag functions and processes.

**Proposition 3.5** (Properties of Cadlag/continuous functions). 1. Let  $D$  be a dense subset of  $\mathbb{R}$  and  $f : A \rightarrow \mathbb{R}$  be a cadlag function. Then for any  $x \in \mathbb{R} \cup \{\pm\infty\}$ ,

$$\lim_{y \rightarrow x} f(y) = z \iff \lim_{y \rightarrow x, y \in D} f(y) = z.$$

2. Let  $D$  be a dense subset of  $\mathbb{R}_+$  and  $(X_t)_{t \geq 0}, (Y_t)_{t \geq 0}$  be cadlag processes. Suppose almost surely

$$X_t = Y_t \text{ for all } t \in D.$$

then almost surely  $X_t = Y_t$  for all  $t \geq 0$ .

3. Let  $f : X \rightarrow Y$  be a continuous function on a metric space and  $E \subseteq X$  be dense. Then  $f(E)$  is dense in  $f(X)$ .  
 Analogously (I *think* the following is true), let  $f : X \rightarrow Y$  with  $X, Y \subseteq \mathbb{R}$  be a cadlag function and  $E \subseteq X$  with  $E$  being dense in  $X$  and  $X$  being open. Then  $f(E)$  is dense in  $f(X)$ .

**Proposition 3.6** (Almost sure martingale convergence theorem, continuous version). Suppose  $X = (X_t)_{t \geq 0}$  is a super-martingale that is bounded in  $L^1$ . Then there exists an  $\mathcal{F}_\infty$ -measurable  $X_\infty \in L^1$  such that

$$X_n \rightarrow X_\infty \text{ a.s. as } n \rightarrow \infty.$$

*Proof.* Note that  $X' = (X_t)'_{t \in \mathbb{Q}_+}$  is a super-martingale that is bounded in  $L^1$ . By the discrete time almost sure martingale convergence theorem, there exists an  $\mathcal{F}_\infty$ -measurable  $X'_\infty \in L^1$  such that  $X'_q \rightarrow X'_\infty$  almost surely as  $q \rightarrow \infty$ . By the previous proposition,  $X_t \rightarrow X_\infty$  almost surely as  $t \rightarrow \infty$ .  $\square$

**Proposition 3.7** (Doob's maximal inequality, continuous version). Let  $X = (X_t)_{t \geq 0}$  be a non-negative cadlag sub-martingale. Then for any  $\lambda > 0$  and writing  $X_t^* = \sup_{0 \leq s \leq t} X_s$ , we have

$$\lambda \mathbb{P}(X_t^* \geq \lambda) \leq \mathbb{E}X_t$$

*Proof.* Note that  $X_t^* = \sup_{0 \leq s \leq t, s \in \mathbb{Q}_+} X_s$ . the result follows from the discrete time analogue.  $\square$

**Proposition 3.8** (Doob's  $L^p$  inequality, continuous version). Let  $(X_t)_{t \geq 0}$  be as above. Then

$$\|X_t^*\|_p \leq \frac{p}{p-1} \|X_t\|_p.$$

*Proof.* The proof of the discrete time version can be applied directly after using the continuous version of Doob's maximal inequality.  $\square$

**Proposition 3.9** (Optional stopping theorem for UI martingales, continuous version). Let  $(X_t)_{t \geq 0}$  be a UI martingale and  $S, T$  be any stopping times. Then

$$\mathbb{E}(X_T \mid \mathcal{F}_S) = X_{S \wedge T} \text{ a.s.}$$

*Proof.* Let  $T_n = \frac{1}{2^n} \lceil 2^n T \rceil$  and  $S_n = \frac{1}{2^n} \lceil 2^n S \rceil$  so  $T_n \downarrow T$  and  $S_n \downarrow S$  as  $n \rightarrow \infty$ . It follows that  $X_{T_n} \rightarrow X_T$  and  $X_{S_n} \rightarrow X_S$  almost surely. By the discrete time optional stopping theorem,  $\mathbb{E}(X_{T_n} \mid \mathcal{F}_{S_n}) = X_{S_n \wedge T_n}$  almost surely. Let  $A \in \mathcal{F}_S \subseteq \mathcal{F}_{S_n}$ . Then

$$\mathbb{E}X_{T_n} \mathbf{1}_A = \mathbb{E}X_{S_n \wedge T_n} \mathbf{1}_A$$

By the discrete time version again,  $X_{T_n} = \mathbb{E}(X_\infty \mid \mathcal{F}_{T_n})$  almost surely so  $(X_{T_n})_{n \geq 0}$  is UI and similarly so is  $(X_{S_n \wedge T_n})_{n \geq 0}$ . By Vitali convergence, take  $n \rightarrow \infty$  in the above equation to get

$$\mathbb{E}(X_T \mid \mathcal{F}_S) = X_{S \wedge T} \text{ a.s.}$$

$\square$

The previous discussion focused on cadlag processes. The following proposition shows that this is not too harsh a restriction.

**Definition** (Version). Let  $X$  and  $Y$  be processes defined on the same probability space. We say  $X$  is a *version* of  $Y$  if for any  $t$  in the index set,  $X_t = Y_t$  almost surely, i.e.

$$\forall t, \mathbb{P}(X_t = Y_t) = 1.$$

**Definition** (Usual conditions). Let  $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbb{P})$  be a filtered probability space. Let  $\mathcal{N}$  be the set of all  $\mathcal{F}$ -measurable sets of measure zero. Define the filtration

$$\tilde{\mathcal{F}}_t = \sigma(\mathcal{F}_t \cup \mathcal{N}).$$

We say  $(\mathcal{F}_t)_{t \geq 0}$  satisfies the *usual conditions* if

$$\tilde{\mathcal{F}}_t = \mathcal{F}_t \text{ for all } t \geq 0.$$

**Lemma 3.9.1** (Upcrossing and convergence). Let  $f : \mathbb{Q}_+ \rightarrow \mathbb{R}$ ,  $a < b$ ,  $a, b \in \mathbb{Q}_+$ ,  $I \subseteq \mathbb{Q}_+$  be bounded. Suppose all of the following statements are true

1.  $f$  is bounded on  $I$
2. Number of upcrossings of the interval  $[a, b]$  by  $f$  on  $I$  is finite, i.e.  $U(a, b, I, f) < \infty$ , where

$$U(a, b, I, f) = \sup\{n \geq 0 : \exists 0 \leq s_1 < t_1 < \dots < s_n < t_n \in I \text{ such that } f(s_i) < a, f(t_i) > b, 1 \leq i \leq n\}$$

Then, for any  $t \in \mathbb{Q}_+$ ,

$$\lim_{s \downarrow t} f(s), \lim_{s \uparrow t} f(s) \text{ exist and are finite.}$$

**Remark.** If  $I$  satisfies the conditions of the lemma, then any  $J \subseteq I$  also satisfies the conditions of the lemma.

*Proof.* Consider any rational sequence  $t_n \downarrow t$ . By 2.2.1,  $f(t_n)$  converges. Since  $f$  is bounded on  $I$ , the limit is finite. Let  $s_n \downarrow t$  be any other rational sequence. Combine the two sequences to get a sequence  $r_n \downarrow t$ . Then  $f(r_n)$  converges to the same limit as  $f(t_n)$  and hence  $f(s_n)$  converges to the same limit. Hence,  $\lim_{s \downarrow t} f(s)$  exists and is finite. The proof for  $\lim_{s \uparrow t} f(s)$  is similar.  $\square$

**Proposition 3.10** (Martingale regularisation theorem). Let  $(X_t)_{t \geq 0}$  be a martingale wrt a filtration  $(\mathcal{F}_t)_{t \geq 0}$ . Then there exists a cadlag process  $\tilde{X}$  such that which is a martingale with respect to  $(\tilde{\mathcal{F}}_t)_{t \geq 0}$  and satisfies

$$\mathbb{P}(X_t = \mathbb{E}(\tilde{X}_t \mid \mathcal{F}_t)) = 1 \text{ for all } t \geq 0.$$

*Proof.* The candidate for  $\tilde{X}_t$  is  $\lim_{s \downarrow t, s \in \mathbb{Q}_+} X_s$  on a set of probability 1. The outline of the proof is as follows:

1. That the limit exists and is finite on a suitable set of probability 1 using the previous lemma.
2. For any  $t$ ,  $X_t = \mathbb{E}(\tilde{X}_t \mid \mathcal{F}_t)$  almost surely.
3.  $\tilde{X}$  is a martingale
4.  $\tilde{X}$  is cadlag

### Step 1

Let  $a, b, I$  be as in the previous lemma and  $\sup I < K < \infty$ . By the corollary to Doob's maximal inequality 2, for any  $\lambda > 0$ ,

$$\lambda \mathbb{P}(\sup_{s \in I} |X_s| \geq \lambda) \leq \sup_{s \in I} \mathbb{E}|X_s| \leq \mathbb{E}|X_K|$$

so  $\mathbb{P}(\sup_{s \in I} |X_s| \geq \lambda) \rightarrow 0$  as  $\lambda \rightarrow \infty$  so  $\mathbb{P}(\sup_{s \in I} |X_s| < \infty) = 1$ .

The following analogue to the corollary of the upcrossing inequality 1 can be proven

$$(b - a)\mathbb{E}U[a, b, I, X] \leq \sup_n \mathbb{E}(|X_n|) + |a| \leq \mathbb{E}|X_K| + |a| < \infty.$$

by taking finite subsets  $J \subseteq I$  and taking the limit as  $J \uparrow I$ . It follows that  $U[a, b, I, X] < \infty$  almost surely.

Hence, for any  $a, b, I$  as in the previous lemma, we can construct a set  $\Omega_{a,b,I}$  satisfying the conditions of the previous lemma. Define  $I_M = \mathbb{Q}_+ \cap [0, M]$  and  $\Omega_0 := \bigcap_{M \in \mathbb{N}} \bigcap_{a,b \in \mathbb{Q}, a < b} \Omega_{a,b,I_M}$ . Then  $\mathbb{P}(\Omega_0) = 1$ .

We show that on  $\Omega_0$ , the conditions of the lemma hold. For any  $a', b', I'$ , pick  $M'$  so  $I' \subseteq I_{M'}$ . Then  $\Omega_{a',b',I'} \subseteq \Omega_{a',b',I_{M'}} \subseteq \Omega_0$ .

Now, define

$$\tilde{X}_t = \begin{cases} \lim_{s \downarrow t, s \in \mathbb{Q}_+} X_s & \text{on } \Omega_0 \\ 0 & \text{on } \Omega_0^c \end{cases}$$

Observe that  $\tilde{X}_t$  is  $\tilde{\mathcal{F}}_t$ -measurable

### Step 2

Let  $t_n \downarrow t$ ,  $t_n \in \mathbb{Q}_+$ . Then  $(X_{t_n})_{n \geq 0}$  is a backwards martingale with respect to the backwards filtration  $(\mathcal{F}_{t_n})_{n \geq 0}$ . By the backwards martingale convergence theorem 2.14, we have  $X_{t_n} \xrightarrow{L^1} \tilde{X}_t$  and hence  $\underbrace{\mathbb{E}(X_{t_n} | \mathcal{F}_t)}_{=X_t} \xrightarrow{L^1} \mathbb{E}(\tilde{X}_t | \mathcal{F}_t)$  and hence

$X_t = \mathbb{E}(\tilde{X}_t | \mathcal{F}_t)$  almost surely.

### Step 3

Let  $s < t$ ,  $s_n \downarrow s$ ,  $s_n \in \mathbb{Q}_+$  and  $s_0 < t$ . By the backwards martingale convergence theorem again,  $\underbrace{\mathbb{E}(X_t | \mathcal{F}_{s_n})}_{=X_{s_n}} \rightarrow \mathbb{E}(X_t | \bigcap_{n \geq 0} \mathcal{F}_{s_n}) = \mathbb{E}(X_t | \mathcal{F}_s^+)$  almost surely. Recall that by definition  $X_{s_n} \rightarrow \tilde{X}_s$  almost surely. Hence  $\tilde{X}_s = \mathbb{E}(\tilde{X}_t | \mathcal{F}_s^+)$  almost surely.

To conclude that  $\tilde{X}$  is a martingale, we use the following result: for any random variable  $X$ ,  $\sigma$ -algebra  $\mathcal{G}$  and collection of measure zero sets  $\mathcal{N}$ ,

$$\mathbb{E}(X | \sigma(\mathcal{G} \cup \mathcal{N})) = \mathbb{E}(X | \mathcal{G}).$$

which can probably be proven using  $\pi$ - $\lambda$  theorem.

It follows that  $\tilde{X}$  is a martingale.

Step 4

Only show right continuity and leave left limit as an exercise. Suppose not. Then there exists sequence  $t_n \downarrow t$  and  $\varepsilon > 0$  such that  $|\tilde{X}_{t_n} - \tilde{X}_t| > \varepsilon$  for all  $n$ . We can pick a rational sequence  $t'_n \downarrow t$  with  $t'_n \geq t_n$  and approximate  $\tilde{X}_{t_n}$  by  $X_{t'_n}$  such that  $|X_{t'_n} - \tilde{X}_{t_n}| < \frac{\varepsilon}{2}$  for all  $n$ . Then  $|X_{t'_n} - \tilde{X}_t| \geq \frac{\varepsilon}{2}$  for all  $n$ . But  $X_{t'_n} \rightarrow \tilde{X}_t$  almost surely, which is a contradiction.  $\square$