

# 1 Martingales

**Definition 1.1.** Let probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  together with a filtration  $\{\mathcal{F}_t\}_{t \in \mathcal{T}}$ . Then the quadruple  $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \in \mathcal{T}}, \mathbb{P})$  is also called a **filtered probability space**.

In this course,  $\mathcal{T}$  will typically be the natural numbers  $\mathbb{N} = \{0, 1, 2, \dots\}$  or  $\mathbb{R}^+ = [0, \infty)$  the non-negative numbers. A martingale is a stochastic process defined with respect to a filtered probability space. Loosely speaking, it represents the total payout of a fair game. That is, the expected value in the future is equal to its current value.

**Definition 1.2.** Let  $X = \{X_t\}_{t \in \mathcal{T}}$  be a stochastic process satisfies the following two conditions.

- $X$  is **adapted** to  $\{\mathcal{F}_t\}_{t \in \mathcal{T}}$ , i.e.,  $X_t$  is  $\mathcal{F}_t$  measurable for all  $t \in \mathcal{T}$ .
- $\mathbb{E}[|X_t|] < \infty$  for all  $t \in \mathcal{T}$ .

$X$  is called a **martingale** (with respect to the filtration  $\{\mathcal{F}_t\}_{t \in \mathcal{T}}$ ) if

$$\mathbb{E}[X_t | \mathcal{F}_s] = X_s \quad \text{for all } s, t \in \mathcal{T} \text{ with } s \leq t. \quad (1)$$

If we say that  $X = \{X_t\}_{t \in \mathcal{T}}$  is a martingale without specifying the filtration, we mean that  $X = \{X_t\}_{t \in \mathcal{T}}$  is a martingale w.r.t. its natural filtration  $\mathcal{F}_t^X = \sigma(X_s | s \in \mathcal{T}, s \leq t)$ .

**Remark 1.3.** The condition (1) is equivalent to

$$\mathbb{E}[X_t - X_s | \mathcal{F}_s] = 0 \quad \text{for all } s, t \in \mathcal{T} \text{ with } s \leq t. \quad (2)$$

If we let  $X_t$  denote the total payouts of a game at time  $t$ , then  $X_t - X_s$  represents the gain (or loss) accumulated between times  $t$  and  $s$ . Condition (2) implies that based on all the information available at time  $s$ , the expected value of this gain (or loss) is zero. In this sense, a martingale can be understood the mathematical formalization of a fair game.

**Remark 1.4.** In discrete time,  $\mathcal{T} = \mathbb{N}$ , the condition (1) is equivalent to

$$\mathbb{E}[X_{n+1} | \mathcal{F}_n] = X_n \quad \text{for all } n \geq 0. \quad (3)$$

The proof is a direct application of the tower property.

**Example 1.5** (Simple Random Walk). Let  $Y_1, Y_2, \dots$  be i.i.d. Rademacher random variables, i.e.  $\mathbb{P}(Y = \pm 1) = \frac{1}{2}$ . Then  $\{X_n\}_{n \geq 0}$  defined through

$$X_0 = 0 \quad \text{and} \quad X_n = \sum_{k=1}^n Y_k \quad (4)$$

is a martingale in discrete time with respect to the natural filtration  $\mathcal{F}_n^X$ . Indeed, since  $Y_{n+1}$  is independent of  $\mathcal{F}_n^X$ ,

$$\mathbb{E}[X_{n+1} | \mathcal{F}_n] = \mathbb{E}[X_n + Y_{n+1} | \mathcal{F}_n] = \mathbb{E}[X_n | \mathcal{F}_n] + \mathbb{E}[Y_{n+1} | \mathcal{F}_n] = X_n + \mathbb{E}[Y_{n+1}] = X_n$$

which satisfies condition (3).

## 1.1 Properties

Naturally, the expected value of the earnings of a fair game is equal to zero.

**Proposition 1.6**

If  $\{X_t\}_{t \in \mathcal{T}}$  is a martingale, then

$$\mathbb{E}[X_t] = \mathbb{E}[X_0] \quad \text{for all } t \in \mathcal{T}.$$

We have the following formula for the second moment of the earnings between time  $s$  and  $t$ .

**Proposition 1.7**

Let  $\{X_t\}_{t \in \mathcal{T}}$  be a martingale with  $\mathbb{E}[(X_t)^2] < \infty$  for all  $t \in \mathcal{T}$ . Then, for  $s, t \in \mathcal{T}$  with  $s \leq t$ ,

$$\mathbb{E}[(X_t - X_s)^2 | \mathcal{F}_s] = \mathbb{E}[X_t^2 | \mathcal{F}_s] - X_s^2.$$

In particular,

$$\mathbb{E}[(X_t - X_s)^2] = \mathbb{E}[X_t^2] - \mathbb{E}[X_s^2].$$

**Example 1.8** (Martingale Betting Strategy). Let  $X_t$  be a simple random walk defined Example 4. We now take an *adapted* stochastic process  $\{\xi_n\}_{n \geq 0}$  where  $\xi_0 = 1$  and, for  $n \geq 1$ ,

$$\xi_n = \begin{cases} 2^n, & \text{if } Y_1 = \cdots = Y_n = -1, \\ 0, & \text{otherwise.} \end{cases}$$

This represents a betting strategy where we double our bet until we win. Then the gambler's total return at time  $n \geq 1$  is

$$\begin{aligned} V_n &= \sum_{k=0}^{n-1} \xi_k (X_{k+1} - X_k) \\ &= \xi_0 Y_1 + \cdots + \xi_{n-1} Y_n \\ &= \begin{cases} -1 - 2 - \cdots - 2^{n-1} = -(2^n - 1), & \text{if } Y_1 = \cdots = Y_n = -1 \\ +1, & \text{otherwise.} \end{cases} \end{aligned}$$

One can show that with probability one there will eventually be some (random) integer  $n$  such that  $Y_n = 1$ , in which case the gambler will have won \$1.

**Example 1.9** (General Betting Strategies). In general, let  $\{X_n\}_{n \geq 0}$  be a martingale denoting the outcomes of a fair game. We let the process  $\{\xi_n\}_{n \geq 0}$  be an adapted process denoting a betting strategy. This means that the  $\xi_n$  bet is a function of the information up to the  $n$ th game.

Suppose we are at game  $k$ , if we bet  $\xi_k$  on the  $k$ th game, then we earn  $\xi_k(X_{k+1} - X_k)$  on the  $k$ th game. Our earnings associated with this betting strategy is therefore

$$V_0 = 0, \quad V_n = \sum_{k=0}^{n-1} \xi_k (X_{k+1} - X_k). \quad (5)$$

A natural question is if one can come up with a smart betting strategy such that  $\mathbb{E}[V_n] > 0 = \mathbb{E}[V_0]$  for some  $n$ ? The answer to that question is no, and it is demonstrated in the following theorem. That is, no betting strategy that can turn a martingale into a favorable game.

**Theorem 1.10**

Suppose  $\{X_n\}_{n \geq 0}$  is an adapted process such that for every  $n$  there exists a constant  $C_n$  such that  $|\xi_n(\omega)| \leq C_n$  for all  $\omega \in \Omega$ . If  $\{X_n\}_{n \geq 0}$  is a martingale, then  $\{V_n\}_{n \geq 0}$  defined in (5) is again a martingale. In particular, we have  $\mathbb{E}[V_n] = 0$  for all  $n$ .

## 1.2 Example Problems

### 1.2.1 Proofs of Results

**Problem 1.1.** Prove Proposition 1.6.

**Solution 1.1.** It follows from (1.2) that

$$\mathbb{E}[X_t] = \mathbb{E}[\mathbb{E}[X_t | \mathcal{F}_0]] = \mathbb{E}[X_0].$$

**Problem 1.2.** Prove Proposition 1.7.

**Solution 1.2.** We have

$$\begin{aligned} \mathbb{E}[(X_t - X_s)^2 | \mathcal{F}_s] &= \mathbb{E}[X_t^2 - 2X_t X_s + X_s^2 | \mathcal{F}_s] \\ &= \mathbb{E}[X_t^2 | \mathcal{F}_s] - 2\mathbb{E}[X_t X_s | \mathcal{F}_s] + \mathbb{E}[X_s^2 | \mathcal{F}_s] \\ &= \mathbb{E}[X_t^2 | \mathcal{F}_s] - 2X_s \mathbb{E}[X_t | \mathcal{F}_s] + X_s^2 \\ &= \mathbb{E}[X_t^2 | \mathcal{F}_s] - 2X_s X_s + X_s^2 \\ &= \mathbb{E}[X_t^2 | \mathcal{F}_s] - X_s^2. \end{aligned}$$

The second identity follows from the first by taking expectations.

**Problem 1.3.** Prove Theorem 1.10.

**Solution 1.3.** We check the properties of a martingale.

- (i) Clearly,  $\{V_n\}_{n \geq 0}$  is adapted.
- (ii) Since  $|\xi_k| \leq C_k$ , we define  $C := \max\{C_1, \dots, C_{n-1}\}$  so that

$$\begin{aligned} \mathbb{E}[|V_n|] &= \mathbb{E}\left[\left|\sum_{k=1}^n \xi_{k-1}(X_k - X_{k-1})\right|\right] \leq \sum_{k=1}^n \mathbb{E}[|\xi_{k-1}(X_k - X_{k-1})|] \\ &\leq \sum_{k=1}^n C_{k-1} \mathbb{E}[|X_k - X_{k-1}|] \leq C \sum_{k=1}^n (\mathbb{E}[|X_k|] + \mathbb{E}[|X_{k-1}|]) < \infty. \end{aligned}$$

- (iii) Next, we have

$$\begin{aligned} \mathbb{E}[V_{n+1} - V_n | \mathcal{F}_n] &= \mathbb{E}[\xi_n(X_{n+1} - X_n) | \mathcal{F}_n] \\ &= \xi_n \mathbb{E}[X_{n+1} - X_n | \mathcal{F}_n] = 0 \end{aligned}$$

so  $\{V_n\}_{n \geq 0}$  is a martingale. Finally, the martingale property Proposition 1.6 implies that

$$\mathbb{E}[V_n] = \mathbb{E}[V_0] = 0, \quad \text{for all } n.$$

### 1.2.2 Definitions and Properties of Martingales

**Problem 1.4.** Let  $Y_1, Y_2, \dots$  be independent (though not necessarily identically distributed) random variables with common expectation  $\mathbb{E}[Y_k] = 0$  for all  $k$ . Show that  $\{X_n\}_{n=0,1,2,\dots}$  defined by

$$X_0 = 0 \quad \text{and} \quad X_n = \sum_{k=1}^n Y_k$$

is a martingale in discrete time with respect to its natural filtration  $\mathcal{F}_n^X$ .

**Solution 1.4.** Since  $Y_{n+1}$  is independent of  $\mathcal{F}_n^X$ ,

$$\mathbb{E}[X_{n+1} | \mathcal{F}_n] = \mathbb{E}[X_n + Y_{n+1} | \mathcal{F}_n] = \mathbb{E}[X_n | \mathcal{F}_n] + \mathbb{E}[Y_{n+1} | \mathcal{F}_n] = X_n + \mathbb{E}[Y_{n+1}] = X_n.$$

**Problem 1.5.** Let  $Y_1, Y_2, \dots$  be independent and nonnegative (though not necessarily identically distributed) random variables with common expectation  $\mathbb{E}[Y_k] = 1$  for all  $k$ .

1. Show that  $\{X_n\}_{n \geq 0}$  defined through

$$X_0 = 1 \quad \text{and} \quad X_n = \prod_{k=1}^n Y_k$$

is a martingale.

2. Let  $Y_k$  be of the form  $Y_k = e^{Z_k - c_k}$  for independent random variables  $Z_k$  with distribution  $N(0, \sigma_k^2)$  and certain constants  $c_k$ . That is, determine  $c_k$  such that  $\{X_n\}_{n \geq 0}$  is a martingale.

**Solution 1.5.**

**Part 1:** The fact that  $X$  is adapted and integrable is clear. To show (3), notice that by independence,

$$\mathbb{E}[X_{n+1} | \mathcal{F}_n] = \mathbb{E}\left[Y_{n+1} \prod_{k=1}^n Y_k \middle| \mathcal{F}_n\right] = \prod_{k=1}^n Y_k \mathbb{E}[Y_{n+1} | \mathcal{F}_n] = \prod_{k=1}^n Y_k \mathbb{E}[Y_{n+1}] = \prod_{k=1}^n Y_k = X_n.$$

**Part 2:** From part 1, it suffices to find a constant so that  $\mathbb{E}[Y_k] = 1$ . We have by the moment generating function formula for the Gaussian,

$$\mathbb{E}[Y_k] = e^{-c_k} \mathbb{E}[e^{Z_k}] = e^{-c_k} e^{\frac{\sigma^2}{2}} = 1 \iff c_k = \frac{\sigma^2}{2}.$$

**Problem 1.6.** Let  $X$  be a random variable such that  $\mathbb{E}[|X|] < \infty$  and  $\mathcal{T}$  either  $\{0, 1, 2, \dots\}$  or  $[0, \infty)$ . Show that

$$X_t := \mathbb{E}[X | \mathcal{F}_t], \quad t \in \mathcal{T},$$

is a martingale.

**Solution 1.6.** Clearly  $X_t$  is  $\mathcal{F}_t$  measurable because the conditional expected value. Furthermore, by Jensen's inequality and the law of total expectation

$$\mathbb{E}[|X_t|] = \mathbb{E}[|\mathbb{E}[X | \mathcal{F}_t]|] \leq \mathbb{E}[\mathbb{E}[|X| | \mathcal{F}_t]] = \mathbb{E}[|X|] < \infty.$$

Next, so show property (1) we have by the tower property that

$$\mathbb{E}[X_t | \mathcal{F}_s] = \mathbb{E}[\mathbb{E}[X | \mathcal{F}_t] | \mathcal{F}_s] = \mathbb{E}[X | \mathcal{F}_s] = X_s.$$

## 2 Stopping time

For this section, we focus on discrete time martingales, but similar statements can be made in continuous time. A stopping time is a random variable that depends on the historical information up to time  $n$ . We can think of a stopping time as a rule that tells use when to stop playing a game, which naturally can only depends on past historical information.

**Definition 2.1.** A random time  $\tau : \Omega \rightarrow \mathbb{N} \cup \{\infty\}$  is called a **stopping time** if  $\{\tau \leq n\} \in \mathcal{F}_n$  for all  $n \geq 0$ .

**Example 2.2.** Let  $\{Y_n\}_{n \geq 0}$  be any adapted process and define

$$\tau = \min\{n : Y_n \geq c\}.$$

Then  $\tau$  is a stopping time, which is sometimes called the **first passage time** of the level  $c$ .

Given a stopping time, we can define  $\sigma$ -algebra

$$\mathcal{F}_\tau = \{A \in \mathcal{F} : \{\tau \leq n\} \cap A \in \mathcal{F}_n \text{ for all } n\}$$

which consists of the events that depend on the information up to a random stopping time  $\tau$ . If  $X_n$  is  $\mathcal{F}_n$  measurable, then the random variable  $X_\tau$  is  $\mathcal{F}_\tau$  measurable.

A stopping time can be interpreted as a strategy to stop a game based only on current and historical information. The next theorem states that we cannot come up with a clever stopping strategy that can turn a martingale into a favorable game.

### Theorem 2.3 (*Optional stopping theorem*)

Let  $\{X_n\}_{n \geq 0}$  be a martingale and  $\tau$  be a stopping time. Suppose that

$$\mathbb{E}|X_\tau| < \infty, \quad \lim_{n \rightarrow \infty} \mathbb{E}|X_n| \mathbb{1}(n \leq \tau) = 0$$

then

$$\mathbb{E}[X_\tau] = \mathbb{E}[X_0].$$

**Remark 2.4.** Notice that this condition is satisfied if there exists a constant  $C$  such that  $\tau \leq C$  almost surely or  $|X_{n \wedge \tau}| \leq C$  for all  $n$  almost surely.

The integrability conditions are essential as demonstrated by the following example.

**Example 2.5.** Consider the situation of Example 1.8, where  $\{X_n\}_{n \geq 0}$  is a simple random walk and  $\{\xi_n\}_{n \geq 0}$  is the martingale betting strategy. We have seen that  $\{V_n\}_{n \geq 0}$  is a martingale with  $V_0 = 0$ . We let

$$\tau(\omega) = \min\{n : Y_n(\omega) = +1\},$$

where  $Y_n = X_n - X_{n-1}$ , denote the first time we win a game. Then  $\tau$  is a stopping time with  $\mathbb{P}(\tau < \infty) = 1$  and  $V_\tau = 1$ . Therefore,

$$\mathbb{E}[V_\tau] = 1 \neq 0 = \mathbb{E}[V_0].$$

This does not contradict Theorem 2.3 because

$$\mathbb{P}(\tau = n) = 2^{-n} \quad \mathbb{P}(\tau > n) = \sum_{k \geq n+1} 2^{-k} = 2^{-n}$$

so

$$\mathbb{E}|V_n| \mathbb{1}(n \leq \tau) = \mathbb{P}(\tau = n) + (2^n - 1) \mathbb{P}(\tau > n) = 1$$

which does not go to zero.

## 2.1 Example Problems

### 2.1.1 Proofs of Results

**Problem 2.1.** Prove the Optional Stopping Theorem (Theorem 2.3) under the assumption that  $\tau$  is a bounded stopping time.

**Solution 2.1.** We first show that  $X_{n \wedge \tau}$  is a martingale. We only consider the case  $\mathcal{T} = \{0, 1, 2, \dots\}$ . We have

$$X_{n \wedge \tau} - X_{(n-1) \wedge \tau} = \mathbb{1}_{\{\tau > n-1\}}(X_n - X_{n-1}).$$

Thus, stopping the process is the same as using the betting strategy  $\xi_n = \mathbb{1}_{\{\tau > n\}}$ , which is adapted since  $\tau$  is a stopping time. More precisely,

$$X_{n \wedge \tau} = X_0 + \sum_{k=1}^n \xi_{k-1}(X_k - X_{k-1})$$

Therefore Theorem 1.10 implies that  $X_{n \wedge \tau}$  is a martingale.

If  $\tau$  is almost surely bounded by some constant  $C$ , then  $X_{N \wedge \tau} = X_\tau$  for all  $N > C$ . Hence, by Proposition 1.6,

$$\mathbb{E}[X_\tau] = \mathbb{E}[X_{N \wedge \tau}] = \mathbb{E}[X_{0 \wedge \tau}] = \mathbb{E}[X_0].$$

**Remark 2.6.** The general proof of Theorem 2.3 uses the dominated convergence theorem to interchange the limit and expected value.

**Problem 2.2.** Show that for a random time  $\tau : \Omega \rightarrow \{0, 1, \dots\} \cup \{\infty\}$ , the following conditions are equivalent:

- (a)  $\tau$  is a stopping time.
- (b) For every  $n \geq 0$ , we have  $\{\tau \leq n\} \in \mathcal{F}_n$ .
- (c) For every  $n \geq 0$ , we have  $\{\tau > n\} \in \mathcal{F}_n$ .
- (d) For every  $n \geq 0$ , we have  $\{\tau = n\} \in \mathcal{F}_n$ .

**Solution 2.2.** The equivalence of (a) and (b) is immediate by the definition. To see that (b) and (c) are equivalent, recall that

$$A \in \mathcal{F}_n \iff A^c \in \mathcal{F}_n.$$

Since  $\{\tau \leq n\}^c = \{\tau > n\}$ , it follows that  $\{\tau \leq n\} \in \mathcal{F}_n \iff \{\tau > n\} \in \mathcal{F}_n$  so (b) and (c) are equivalent. To see that (b) and (c) are equivalent, notice that

$$\{\tau = n\} = \{\tau \leq n\} \setminus \{\tau \leq n-1\}$$

so  $\{\tau = n\}$  if (b) holds since the  $\sigma$ -algebra is closed under unions and complements.

### 2.1.2 Applications

**Problem 2.3.** Let  $\{X_n\}_{n \geq 0}$  is a simple random walk,  $a, b \in \mathbb{N}$ , and

$$\tau = \min\{n \mid X_n = -a \text{ or } X_n = b\}.$$

Find

$$\mathbb{P}[X_\tau = b].$$

**Solution 2.3.** Recall that  $\{X_n\}_{n \geq 0}$  is a martingale and  $\tau$  is a stopping time. We can interpret  $X_{n \wedge \tau}$  as the balance in a fair coin-tossing game between two players with respective capital  $a$  and  $b$ . We are interested in the probability that the player with capital  $b$  goes bankrupt before the other player, i.e.,  $\mathbb{P}(X_\tau = b)$ .

We have that the stopping time  $\tau$  satisfies  $\mathbb{P}(\tau < \infty) = 1$  and  $|X_{n \wedge \tau}| \leq a \vee b$  for all  $n$  almost surely. Therefore, using Theorem 2.3 with uniformly bounded stopped martingales implies that

$$0 = \mathbb{E}[X_0] = \mathbb{E}[X_\tau] = b\mathbb{P}(X_\tau = b) - a\mathbb{P}(X_\tau = -a) = b\mathbb{P}(X_\tau = b) - a(1 - \mathbb{P}(X_\tau = b))$$

which gives

$$\mathbb{P}[X_\tau = b] = \frac{a}{a + b}.$$