EC 655 Limited Dependent Variables 1

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Introduction

- ▶ Many economic outcomes are qualitative in nature
 - Grouped into categories
 - ▶ Work or not
 - Drive, take bus, cycle to work
 - Number of visits to Doctor
- ▶ People sometimes use alternative methods in this context
 - Though linear regression does still work in many cases
- ▶ In this section we cover models for these types of variables

Potential Outcomes and Linear Regression

- Start with the same potential outcomes setup
 - y₁ is the outcome with treatment
 - $ightharpoonup y_0$ is the outcome without treatment
 - w is a binary variable with 1 denoting treatment, and 0 no treatment
- We observe (y, w), where

$$y = y_0 + (y_1 - y_0)w$$

▶ The average difference between treated and control is

$$E(y|w=1) - E(y|w=0)$$
$$= [E(y_1|w=1) - E(y_0|w=1)] + E(y_0|w=1) - E(y_0|w=0)$$



▶ Key difference here is that y_1 and y_0 are binary

$$y_0 \in \{0, 1\}$$
$$y_1 \in \{0, 1\}$$

- ▶ You can get treatments effects in the same way we did before
 - ▶ Use linear regression
 - Estimate by OLS
- Only the interpretation changes
 - Since y is a Bernoulli random variable,

$$E(y|w) = Pr(y = 1|w)$$

So you can restate the difference in observed means as

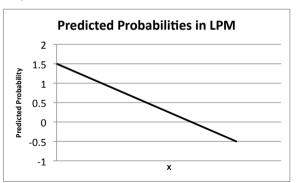
$$Pr(y = 1|w = 1) - Pr(y = 1|w = 0)$$

= $[Pr(y_1 = 1|w = 1) - Pr(y_0 = 1|w = 1)] + Pr(y_0 = 1|w = 1) - Pr(y_0 = 1|w = 0)$

- You can interpret as the difference in response probabilities
- ▶ Difference in observed response probabilities will estimate treatment effects if
 - Independence of potential outcomes
 - Mean independence of potential outcomes
 - Conditional mean independence of potential outcomes
- ► Can also use linear regression with binary or continuous treatment
- Called the Linear Probability Model

Issues with Linear Probability Model

- 1. Predicted probabilities are not restricted to [0,1] interval
 - ► Can have nonsense probabilities



2. Heteroskedasticity

$$Var[y|w] = E[y - E[y|w]|w]$$

= $E[y^2|w] - (E[y|w])^2$
= $Pr[y = 1|w](1 - Pr[y = 1|w])$

- As noted above Pr[y=1|w] depends on w
- lacktriangle So the variance of y will depend on w
- ▶ We therefore must use robust standard errors
- You could also use GLS

Nonlinear Models

- ▶ In some cases we may want to fix the predicted probability issue
 - ▶ If you are doing prediction, for example
- One way to do this is to feed the model through a CDF
- ▶ This is often motivated with an index model

$$y_0^* = \beta_0 - \eta$$
$$y_1^* = y_0^* + \beta_1$$
$$y^* = y_0^* + (y_1^* - y_0^*)w$$
$$y^* = \beta_0 + \beta_1 w - \eta$$

- ▶ In this setup, we do not observe y^*
 - ▶ It is some underlying continuous outcome driving our decisions
- ▶ Instead, we observe the binary *y* where

$$y = 1\{y^* > 0\}$$

► Plugging the model into this

$$y = 1\{\beta_0 + \beta_1 w - \eta > 0\}$$

▶ The probability that y = 1 is therefore

$$Pr(y = 1|w) = Pr(\beta_0 + \beta_1 w - \eta > 0|w)$$

▶ The random component here is η , so rearrange to isolate it

$$Pr(y = 1|w) = Pr(\beta_0 + \beta_1 w > \eta|w)$$
$$Pr(y = 1|w) = Pr(\eta < \beta_0 + \beta_1 w|w)$$

- ▶ The probability that *y* equals 1 depends on
 - ightharpoonup Treatment status w
 - ▶ The distribution of eta
- ightharpoonup Different choices for the η distribution lead to different models

Probit Model

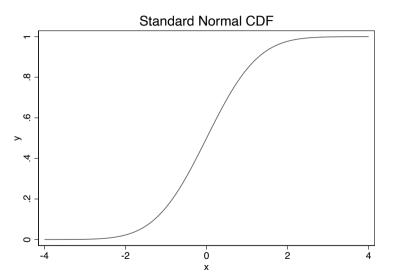
• Assuming $\eta \sim \mathcal{N}(0, \sigma_{\eta}^2)$ leads to the **Probit Model**

$$Pr(y=1|w) = \Phi\left(\frac{\beta_0 + \beta_1 w}{\sigma_\eta^2}\right)$$

where

$$\Phi(z) = \int_{-\infty}^{z} \frac{1}{\sqrt{2\pi}} e^{\frac{-v^2}{2}} dv$$

- ▶ Because $\Phi(.)$ is a CDF, Pr(y=1|w) is always between 0 and 1
 - ▶ This solves the predicted probability problem



▶ The observed difference in probabilities is

$$Pr(y = 1|w = 1) - Pr(y = 1|w = 0)$$
$$= \Phi\left(\frac{\beta_0 + \beta_1}{\sigma_\eta^2}\right) - \Phi\left(\frac{\beta_0}{\sigma_\eta^2}\right)$$

- ▶ Notice that β_1 does <u>not</u> equal the difference in response probabilities
 - They are the slopes in the index model
 - ▶ The index model parameters are not usually of interest
- ▶ To get the difference in response probabilities, feed parameters into the CDF first
- In nonlinear models, parameters are not "marginal effects"
 - ▶ You need to separately compute them after estimating the model

▶ In models with more variables and where they are continuous

$$Pr(y=1|\mathbf{x}) = \Phi\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)$$

▶ The marginal effect for continuous variable x_i

$$\frac{\partial Pr(y=1|\mathbf{x})}{\partial x_j} = \phi\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right) \frac{\beta_j}{\sigma_{\eta}^2}$$

- ▶ This is a function of the entire vector x
 - You need to specify their values to get the marginal effect
 - Normally people hold them at the mean
 - ▶ In theory you can get a distribution of marginal effects

Logit Model

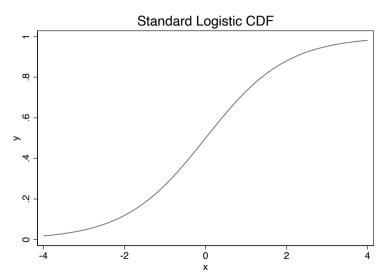
• Assuming $\eta \sim \operatorname{Logistic}(0, \sigma_{\eta}^2)$ leads to the **Logit Model**

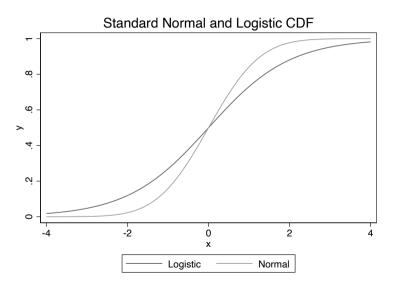
$$Pr(y=1|w) = \Lambda\left(\frac{\beta_0 + \beta_1 w}{\sigma_\eta^2}\right)$$

where

$$\Lambda(z) = \frac{e^z}{1 + e^z}$$

▶ Again, because $\Lambda(.)$ is a CDF, Pr(y=1|w) is always between 0 and 1





▶ The observed difference in probabilities is

$$Pr(y = 1|w = 1) - Pr(y = 1|w = 0)$$
$$= \Lambda \left(\frac{\beta_0 + \beta_1}{\sigma_\eta^2}\right) - \Lambda \left(\frac{\beta_0}{\sigma_\eta^2}\right)$$

In models with more variables and where they are continuous

$$Pr(y=1|\mathbf{x}) = \Lambda\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)$$

▶ The marginal effect for continuous variable x_i

$$\frac{\partial Pr(y=1|\mathbf{x})}{\partial x_j} = \Lambda\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right) \left(1 - \Lambda\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)\right) \frac{\beta_j}{\sigma_{\eta}^2}$$

Estimation of Probit and Logit

- Both models usually estimated by Maximum Likelihood
- ▶ Method maximizes the probability of getting our sample by choosing parameters
 - ▶ The**Likelihood Function** is the function of the parameters given the data
- ▶ The probability distribution of y_i is

$$f(y_i|\mathbf{x_i};\boldsymbol{\beta}) = F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)^{y_i} \left(1 - F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)\right)^{1 - y_i}$$

ightharpoonup The joint probability of observing the all the y_i in the data is

$$f(\mathbf{y}|\mathbf{X};\boldsymbol{\beta}) = \prod_{i=1}^{n} F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)^{y_i} \left(1 - F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)\right)^{1 - y_i}$$

▶ The Likelihood Function recasts as a function of the parameters given the data

$$\mathcal{L}(\boldsymbol{\beta}|\mathbf{y}, \mathbf{X}) = \prod_{i=1}^{n} F\left(\frac{\mathbf{x}_{i}\boldsymbol{\beta}}{\sigma_{n}^{2}}\right)^{y_{i}} \left(1 - F\left(\frac{\mathbf{x}_{i}\boldsymbol{\beta}}{\sigma_{n}^{2}}\right)\right)^{1 - y_{i}}$$

- Both models usually estimated by Maximum Likelihood
- \triangleright Method maximizes the joint probability of y values conditional on x
 - ► This is called the **Likelihood Function**
- \blacktriangleright In the case of a binary y, the likelihood function is

$$f[y|\mathbf{x};\boldsymbol{\beta}] = P[Y_1 = y_1, Y_2 = y_2, \dots, Y_n = y_n | \mathbf{x_i}; \boldsymbol{\beta}]$$

$$= \Pi_{y_i=1} F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right) \Pi_{y_i=0} \left[1 - F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)\right]$$

$$= \Pi_{i=1}^n F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)^{y_i} \left[1 - F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right)\right]^{1-y_i}$$

- Researchers usually focus on the log of the likelihood instead
 - ▶ It is a monotonic (increasing) transformation of the likelihood
 - ▶ The same parameter vector solves both versions
 - Log likelihoods are easier to work with
- Log Likelihood

$$ln\mathcal{L}(\boldsymbol{\beta}|\mathbf{y}, \mathbf{X}) = \sum_{i=1}^{N} \{y_i lnF\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right) + (1 - y_i) ln(1 - F\left(\frac{\mathbf{x_i}\boldsymbol{\beta}}{\sigma_{\eta}^2}\right))\}$$

- ▶ To solve this equation, you need numerical methods
 - ► A grid search algorithm that finds the maximum value

In ML environments, the estimated variance of $\hat{\beta}$ is estimated as the negative of the expected value of the Hessian (information matrix)

$$\hat{Var}(\hat{\boldsymbol{\beta}}) = -E[(\frac{\partial^{2} lnL}{\partial \hat{\boldsymbol{\beta}} \partial \hat{\boldsymbol{\beta}}'})^{-1}]$$
$$= (\sum_{i=1}^{n} \frac{f(\mathbf{x}_{i}'\hat{\boldsymbol{\beta}})^{2}}{F(\mathbf{x}_{i}'\hat{\boldsymbol{\beta}})(1 - F(\mathbf{x}_{i}'\hat{\boldsymbol{\beta}}))} \mathbf{x}_{i} \mathbf{x}_{i}')^{-1}$$

• where F(.) is either the Normal or Logistic CDF, and f(.) is the associated PDF

Hypothesis Testing in Probit and Logit

- ▶ Simple tests for coefficient significance is done by the usual *t*-test method
 - lacktriangle Assume \hat{eta} has normal distribution (asymptotically)
 - ▶ Test statistic $Z = \frac{\hat{\beta}_k}{\hat{SE}(\hat{\beta}_k)}$



- ▶ More complicated tests done using one of 3 methods:
 - 1. Likelihood Ratio (LR) Test
 - ▶ Test statistic $LR = 2[ln\hat{L}_U ln\hat{L}_R]$
 - $lacktriangleright ln\hat{L}_R$ is log likelihood evaluated at restricted parameter vector
 - 2. Wald (W) Test
 - ▶ Test statistic $W = \hat{g}^{'}[\hat{G}\hat{Var}(\hat{\pmb{\beta}})\hat{G}^{'}]\hat{g}$
 - $lackbox{} \hat{g}$ is a vector of restrictions evaluated at $\hat{oldsymbol{eta}}$
 - $lacksim \hat{G}$ is the derivative of a vector of restrictions evaluated at $\hat{oldsymbol{eta}}$
 - 3. Lagrange Multiplier (LM) Test
 - $lackbox{ Test statistic } LM = \hat{d}^{'}[\hat{Var}(\hat{oldsymbol{eta}})]\hat{d}$
 - $lackbox{}{\hat{d}}$ is the derivative of lnL evaluated at restricted $\hat{oldsymbol{eta}}$
- ▶ Restrictions can be linear or non-linear

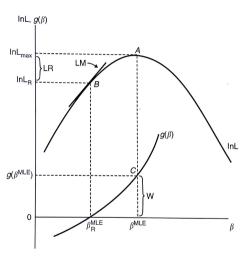


Figure: Source: Kennedy (2008)

Goodness of Fit in Probit and Logit

1. Confusion Matrix

		Actual	
Predicted	0	1	Total
0	# Correct 0	# Incorrect 1	# Pred 0
1	# Incorrect 0	$\#Correct\ 1$	# Pred 1
Total	# True 0	# True 1	

2. Pseudo- R^2

- ightharpoonup McFadden $ightarrow R^2 = 1 rac{ln\hat{L}_U}{ln\hat{L}_0}$
 - ▶ $ln\hat{L}_0$ is log-likelihood with no explanatory variables
- Others are possible, but goodness of fit is not incredibly important

Introduction

- Arises when a continuous dependent variable is limited in its range
 - Censoring
 - Income is top-coded at some level for privacy
 - Corner Solutions
 - Spending on consumer durables limited below by 0
 - Incidental Truncation
 - Wage is not observed for people who do not work
- You can sometimes use OLS depending on context
- ▶ We will cover models for censoring and corner solutions

Potential Outcomes

▶ We will again appeal to a latent variable model

$$y_0^* = \beta_0 - \eta$$
$$y_1^* = y_0^* + \beta_1$$
$$y^* = y_0^* + (y_1^* - y_0^*) * w$$
$$y^* = \beta_0 + \beta_1 * w - \eta$$

ightharpoonup In this case, the observed y is

$$y = max\{0, y^*\}$$



ightharpoonup Because it is either zero or positive, the expected value of y is

$$E[y|w] = E[y|y>0, w] Pr[y>0|w]$$

▶ Taking the difference in the observed *y*, we get

$$\begin{split} E[y|w=1] - E[y|w=0] \\ = E[y|y>0, w=1] Pr[y>0|w=1] - E[y|y>0, w=0] Pr[y>0|w=0] \\ = (Pr[y>0|w=1] - Pr[y>0|w=1]) E[y|y>0, w=1] \\ + (E[y|y>0, w=1] - E[y|y>0, w=0]) Pr[y>0|w=0] \end{split}$$

- ► There are two key pieces
 - Participation effect
 - Conditional on Positive effect
- ▶ In terms of the potential outcomes

$$E[y|w=1] - E[y|w=0]$$

$$= E[y_1|w=1] - E[y_0|w=0]$$

$$= E[y_1|w=1] - E[y_0|w=1] + E[y_0|w=1] - E[y_0|w=0]$$

- ▶ The limited dependent variable does not change causal interpretation
 - \blacktriangleright As long as potential outcome y_0 is mean independent of treatment

Conditional on Positive

- ▶ In some contexts people run regressions with just the positive outcomes
 - ▶ If you wanted to analyze participation decision separately
- lacktriangle Difference in observed y for this group is biased if under random assignment

$$E[y|y > 0, w = 1] - E[y|y > 0, w = 0]$$

$$= E[y_1|y_1 > 0] - E[y_0|y_0 > 0]$$

$$= E[y_1|y_1 > 0] - E[y_0|y_1 > 0] + E[y_0|y_1 > 0] - E[y_0|y_0 > 0]$$

- ▶ The treatment changes who has positive values of potential outcomes
 - ► More subtle form of bias
- You cannot interpret conditional on positive effects as causal



Tobit Model

- ▶ This model is used in the context of censoring and corner solutions
- ▶ We have data on a random sample
- ▶ The outcome is limited in its range
- ▶ There is a mass of observations at 1 or more values
 - Usually zeroes
 - Sometimes some upper amount, like income
- Using OLS may be a bad strategy depending on your goals
 - Can produce predicted values outside the limited range

▶ The Tobit model starts with the latent variable model

$$y^* = \mathbf{x} \boldsymbol{\beta} - \eta, \text{ where } \eta \sim N(0, \sigma^2)$$

$$y = \max(0, y^*)$$

- ▶ The conditional expectation of interest depends on context
 - Censored data
 - $ightharpoonup E[y^*|\mathbf{x}]$
 - $ightharpoonup y^*$ usually has meaning when data are censored
 - Corner Solutions
 - $E[y|\mathbf{x}]$ and $E[y|y>0,\mathbf{x}]$
 - $lackbox{ } y_i^*$ usually has no meaning for corner solutions
 - ▶ This is the most common situation

- Estimate this model by maximum likelihood
- ► The likelihood function has two pieces
 - When $y_i = 0$

$$Pr(y_i = 0|\mathbf{x}) = Pr(y_i^* < 0|\mathbf{x})$$

$$= Pr(\mathbf{x}\boldsymbol{\beta} - \eta < 0|\mathbf{x})$$

$$= Pr(\eta > \mathbf{x}\boldsymbol{\beta}|\mathbf{x})$$

$$= Pr(\frac{\eta}{\sigma_{\eta}} > \frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}|\mathbf{x})$$

$$= 1 - \Phi\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right)$$

- ▶ likelihood function continued...
 - When $y_i > 0$

$$f(y_i|y_i > 0, \mathbf{x}) = \frac{1}{\sigma_{\eta}} \phi\left(\frac{y_i - \mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right)$$

Combining terms, we can form the likelihood function

$$\mathcal{L}(\boldsymbol{\beta}|\mathbf{y}, \mathbf{X}) = \Pi_{y_i=0} \left(1 - \Phi\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right) \right) \Pi_{y_i>0} \left(\frac{1}{\sigma_{\eta}} \phi\left(\frac{y_i - \mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right) \right)$$

► The log likelihood is

$$ln\mathcal{L}(\boldsymbol{\beta}|\mathbf{y}, \mathbf{X}) = \sum_{y_i = 0} ln \left(1 - \Phi\left(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right) \right) + \sum_{y_i > 0} ln \left(\frac{1}{\sigma_{\eta}} \phi\left(\frac{y_i - \mathbf{x}\boldsymbol{\beta}}{\sigma_{\eta}}\right) \right)$$

lacktriangle Maximize this function by choosing the parameter vector $oldsymbol{eta}$ and σ_{η}



- Marginal effects will depend on the context of our estimation
 - Censored data

$$\frac{\partial E[y^*|\mathbf{x}]}{\partial x_k} = \beta_k$$

- ▶ In a Tobit with censored data, you can interpret the slope directly
- Corner Solutions
- $\qquad \qquad \qquad \bullet \frac{\partial E[y|y>0,\mathbf{x}]}{\partial x_k} = \{1 \lambda(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma})[\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma} + \lambda(\frac{\mathbf{x}\boldsymbol{\beta}}{\sigma})]\}\beta_k$
 - With corner solutions it depends on what you want
 - ightharpoonup You may want slope for random person, or conditional on y>0

Issues with Tobit Model

- 1. It must be possible for the dependent variable to take values near the limit
 - Example: not the case with consumer durables
 - You either spend zero or a large amount
- 2. Intensive and Extensive margins have same parameters
 - Means the model is relatively inflexible
 - Can be solved by modelling each separately
- 3. Normality assumption
- 4. Care must be taken in interpreting the coefficients
 - ▶ Do we care about the effect of x_k on y or y^* ?

