# IMA205 - Introduction Supervised Learning

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### 1 OLS

The OLS estimator is defined as  $\beta^* = (x^T x)^{-1} x^T y = Hy$ . Another linear unbiased estimator of  $\beta$  is defined as  $\tilde{\beta} = Cy$ , where C is a matrix dxn and C = H + D, D being a non-zero matrix.

- Calculating expected value and variance of  $\tilde{\beta}$  :
- $E[\tilde{\beta}] = E[Cy] = (I_d + Dx)\beta$ , as  $\tilde{\beta}$  is unbiased, it can be concluded that Dx must be equal to zero.
- $Var(\tilde{\beta}) = Var(Cy)$ , using the variance property:  $Var(Cy) = Cvar(y)C^T = \sigma^2 CC^T$ .

$$\begin{split} \sigma^2 C C^T &= \sigma^2 ((x^T x)^{-1} x^T + D) (x (x^T x)^{-1} + D^T) \\ \sigma^2 C C^T &= \sigma^2 (x^T x)^{-1} + \sigma^2 (x^T x)^{-1} (Dx)^T + \sigma^2 (Dx) (x^T x)^{-1} + \sigma^2 (DD^T) \end{split}$$

As proved in the expected value Dx must be equal to 0 so the estimator is unbiased, this means that:

$$Var(\tilde{\beta}) = \sigma^2 (x^T x)^{-1} + \sigma^2 (DD^T)$$

A matrix  $DD^T$  is always symmetric and semi positive, that means  $DD^T \ge 0$ , if it's equal to 0, is equivalent to the OLS, to conclude:

$$Var(\tilde{\beta}) = Var(\beta^*) + \sigma^2(DD^T)$$

As the second term is greater than 0, the variance of this new estimator is greater than the OLS.

Given the calculations above, the assumption that  $Var(\beta^* < Var(\tilde{\beta})$  holds. That is assuming that x is deterministic and  $E[\epsilon] = 0$  (normality assumption holds,  $\epsilon \sim \mathcal{N}(0, \sigma^2 I)$ 

## 2 Ridge regression

• The explicit Ridge solution can be written as follows:  $\beta^*_{ridge} = (x_c^T x_c + \lambda I)^{-1} x_c^T y_c$ . So the expected value is given by:  $E[\beta^*_{ridge}] = E[(x_c^T x_c + \lambda I)^{-1} x_c^T y_c] = [(x_c^T x_c + \lambda I)^{-1} x_c^T] E[y_c]$ 

$$E[\beta_{ridge}^*] = [(x_c^T x_c + \lambda I)^{-1} x_c^T x_c] \beta$$

Which is different from  $\beta$  unless lambda = 0 (the OLS case), meaning is a biased estimator.

• The SVD decomposition for the Ridge estimator can be written as:

$$\beta_{ridge}^* = (x_c^T x_c + \lambda I)^{-1} x_c^T y_c = ([UDV^T]^T [UDV^T] + \lambda I)^{-1} (UDV^T)^T y_c$$

$$= (VD^T U^T UDV^T + \lambda I)^{-1} VD^T U^T y_c = (VD^T DV^T + \lambda I)^{-1} VD^T U^T y_c = V(D^T D + \lambda I)^{-1} V^T VD^T U^T y_c$$

$$\beta_{ridge}^* = V(D^T D + \lambda I)^{-1} D^T U^T y_c$$

The manipulations to get the result above use that U and V are orthogonal matrix (the inverse is equal to the transpose). Using this transformation might be computationally useful because there is no need to invert a matrix, as  $(D^TD + \lambda I)^{-1}D^T$  is equal to a diagonal matrix, where each element is equal to  $\frac{eigenvalue}{(eigenvalue^2 + \lambda)}$ .

• The variance of the Ridge estimator can be calculated as:  $Var(\beta_{ridge}^*) = Var((x_c^T x_c + \lambda I)^{-1} x_c^T y_c)$ 

$$Var(\beta_{ridge}^*) = ((x_c^T x_c + \lambda I)^{-1} x_c^T) Var(y_c) ((x_c^T x_c + \lambda I)^{-1} x_c^T)^T$$
$$Var(\beta_{ridge}^*) = \sigma^2 (x_c^T x_c + \lambda I)^{-1} x_c^T x_c (x_c^T x_c + \lambda I)^{-1}$$

For a positive  $\lambda$ ,  $(x_c^T x_c + \lambda I)$  will always be greater than  $x_c^T x_c$ , as consequence  $(x_c^T x_c + \lambda I)^{-1} x_c^T x_c (x_c^T x_c + \lambda I)^{-1}$  will always be smaller than  $(x_c^T x_c)^{-1}$ , meaning that  $Var(\beta_{OLS}^*) \geq Var(\beta_{Bidge}^*)$ .

• The Ridge estimator promotes a trade-off between Bias and Variance. As  $\lambda$  increases, the Bias becomes bigger, and the variance becomes smaller.

This is logical given that if we take a  $\lambda$  really close to zero, the solution will tend to the OLS solution, with 0 bias and high variance, and if  $\lambda$  is close to infinity, the solution will be all parameters equal to zero, meaning zero variance, but high variance.

• As:  $\beta_{ridge}^* = (x_c^T x_c + \lambda I_d)^{-1} x_c^T y_c$ . If  $x_c^T x_c = I_d$ , Therefore  $\beta_{ridge}^* = (I_d + \lambda I_d)^{-1} x_c^T y_c = ((1 + \lambda) I_d)^{-1} x_c^T y_c$ . Remembering:  $\beta_{OLS}^* = (x_c^T x_c)^{-1} x_c^T y_c$ , where too  $x_c^T x_c = I_d$ , than  $\beta_{OLS}^* = x_c^T y_c$ Substituting, it's demonstrated that  $\beta_{ridge}^* = \frac{\beta_{OLS}^*}{1+\lambda}$ 

## 3 Elastic Net

Rewriting equation 2 from the exercise list:

$$\beta_{ElNet}^* = argmin_{\beta}(y_c - x_c\beta)^T (y_c - x_c\beta) + \lambda_2 ||\beta||_2^2 + \lambda_1 ||\beta||_1$$

As the function is strictly convex, the minimum can be obtained equaling the subgradient to zero  $(\lambda_1||\beta||_1)$  is not differentiable in 0.)

$$\frac{\partial f}{\partial \beta} = 2x_c^T (y_c - x_c \beta) + 2\lambda_2 \beta + \lambda_1 \begin{cases} \{-1\} & , \beta < 0 \\ \{1\} & , \beta > 0 \\ [-1, 1] & , \beta = 0 \end{cases}$$

$$2x_c^T(y_c - x_c\beta) + 2\lambda_2\beta \pm \lambda_1 = 0$$

$$2x_c^T y_c - 2x_c^T x_c \beta + 2\lambda_2 \beta \pm \lambda_1 = 0$$

Remembering that  $x_c^T x_c = I_d$ , so  $\beta_{OLS}^* = x_c^T y_c$ .

$$2\beta_{OLS}^* - 2\beta(1 - \lambda_2) \pm \lambda_1 = 0$$
$$\beta = \frac{\beta_{OLS}^* \pm \frac{\lambda_1}{2}}{(1 - \lambda_2)}$$

Giving the expected value proved by this demonstration.