

# ATE & Variance Estimators for the Case with Multiple Treatments and Linear Adjustments

Juri Trifonov

February 16, 2024

## 1 CAR without Clusters

### ATE Estimator

Let us recall the version for a binary treatment case:

$$\hat{\tau} \equiv \frac{1}{N} \sum_{i=1}^N \left[ \frac{A_i(Y_i - \hat{\mu}(1, S_i, X_i))}{\hat{\pi}(S_i)} - \frac{(1 - A_i)(Y_i - \hat{\mu}(0, S_i, X_i))}{1 - \hat{\pi}(S_i)} + \hat{\mu}(1, S_i, X_i) - \hat{\mu}(0, S_i, X_i) \right].$$

Since in the case with multiple treatments  $\hat{\pi}_a(s) + \hat{\pi}_0(s) \neq 1$ , we should make the following corrections to the expression. First, instead of  $1 - \hat{\pi}(s)$  we should directly use  $\hat{\pi}_0 = \frac{n_0(s)}{n(s)}$ . Second, in the original expression, the numerator of the first two terms implies that  $\hat{\pi}(s) + (1 - \hat{\pi}(s)) = 1$ . Since in our setting the last equality does not hold, we need to multiply the last two terms by  $\hat{\pi}_a(s) + \hat{\pi}_0(s)$ . Thus, the estimator for any treatment  $a \in \mathcal{A}$  (relative to control) becomes:

$$\hat{\tau}_a \equiv \frac{1}{n_a + n_0} \sum_{i \in I_{a,0}} \left[ (\hat{\pi}_0(S_i) + \hat{\pi}_a(S_i)) \times \left( \frac{\tilde{A}_i(Y_i - \hat{\mu}(a, S_i, X_i))}{\hat{\pi}_a(S_i)} - \frac{(1 - \tilde{A}_i)(Y_i - \hat{\mu}(0, S_i, X_i))}{\hat{\pi}_0(S_i)} \right) + \hat{\mu}(a, S_i, X_i) - \hat{\mu}(0, S_i, X_i) \right]$$

where

$$\tilde{A}_i = \begin{cases} 1, & \text{if } A_i = a, \forall a \in \mathcal{A} \\ 0, & \text{otherwise.} \end{cases}$$

### Variance Estimator

Recall that the expression of the variance estimator for the binary treatment case has the following form:

$$\hat{\sigma}^2 = \frac{1}{N} \sum_{i=1}^N \left[ A_i \hat{\Xi}_1^2(\mathcal{D}_i, S_i) + (1 - A_i) \hat{\Xi}_0^2(\mathcal{D}_i, S_i) + \hat{\Xi}_2^2(\mathcal{D}_i, S_i) \right],$$

where

$$\begin{aligned}
\hat{\Xi}_1(\mathcal{D}_i, s) &= \tilde{\Xi}_1(s) - \frac{1}{N_1(s)} \sum_{j \in I_1(s)} \tilde{\Xi}_{1,j}(\mathcal{D}_i, s), \\
\hat{\Xi}_0(\mathcal{D}_i, s) &= \tilde{\Xi}_0(s) - \frac{1}{N_0(s)} \sum_{j \in I_0(s)} \tilde{\Xi}_{0,j}(\mathcal{D}_i, s), \\
\hat{\Xi}_2 &= \left( \frac{1}{N_1(s)} \sum_{j \in I_1(s)} (Y_j - \hat{\tau} A_j) \right) - \left( \frac{1}{N_0(s)} \sum_{j \in I_0(s)} (Y_j - \hat{\tau} A_j) \right) \\
\tilde{\Xi}_1(\mathcal{D}_i, s) &= \left( 1 - \frac{1}{\hat{\pi}(s)} \right) \hat{\mu}(1, s, X_i) - \hat{\mu}(0, s, X_i) + \frac{Y_i}{\hat{\pi}(s)}, \\
\tilde{\Xi}_0(\mathcal{D}_i, s) &= \left( \frac{1}{1 - \hat{\pi}(s)} - 1 \right) \hat{\mu}(0, s, X_i) + \hat{\mu}(1, s, X_i) - \frac{Y_i}{1 - \hat{\pi}(s)}.
\end{aligned}$$

Following the same reasoning as for the ATE estimator, we need to make the following corrections to get the expression for the multiple treatments case.

First, notice that we can rewrite  $\tilde{\Xi}_1(\mathcal{D}_i, s)$  and  $\tilde{\Xi}_0(\mathcal{D}_i, s)$  as follows:

$$\begin{aligned}
\tilde{\Xi}_1(\mathcal{D}_i, s) &= \hat{\mu}(1, s, X_i) - \hat{\mu}(0, s, X_i) + \frac{Y_i - \hat{\mu}(1, s, X_i)}{\hat{\pi}(s)}, \\
\tilde{\Xi}_0(\mathcal{D}_i, s) &= \hat{\mu}(1, s, X_i) - \hat{\mu}(0, s, X_i) - \frac{Y_i - \hat{\mu}(0, s, X_i)}{1 - \hat{\pi}(s)}.
\end{aligned}$$

For the same reasoning as before, to generalize these terms for the multiple treatments scenario we need to substitute  $\hat{\pi}_a(s)$  and  $\hat{\pi}_0(s)$  for  $\hat{\pi}(s)$  and  $1 - \hat{\pi}(s)$  respectively.

Second, for the same reason as before, we need to multiply terms  $\frac{Y_i - \hat{\mu}(a, s, X_i)}{\hat{\pi}_a(s)}$  and  $\frac{Y_i - \hat{\mu}(0, s, X_i)}{\hat{\pi}_0(s)}$  in expressions for  $\tilde{\Xi}_a$  and  $\tilde{\Xi}_0$  by  $\hat{\pi}_0(s) + \hat{\pi}_a(s)$ .

Combining the corrections, we get the following expression of the variance estimator for every treatment  $a \in \mathcal{A}$ :

$$\hat{\sigma}_a^2 = \frac{1}{n_a + n_0} \sum_{i \in I_{a,0}} \left[ \tilde{A}_i \hat{\Xi}_a^2(\mathcal{D}_i, S_i) + (1 - \tilde{A}_i) \hat{\Xi}_0^2(\mathcal{D}_i, S_i) + \hat{\Xi}_2^2(\mathcal{D}_i, S_i) \right],$$

where

$$\begin{aligned}
\hat{\Xi}_a(\mathcal{D}_i, s) &= \tilde{\Xi}_a(s) - \frac{1}{N_a(s)} \sum_{j \in I_a(s)} \tilde{\Xi}_{a,j}(\mathcal{D}_i, s), \\
\hat{\Xi}_0(\mathcal{D}_i, s) &= \tilde{\Xi}_0(s) - \frac{1}{N_0(s)} \sum_{j \in I_0(s)} \tilde{\Xi}_{0,j}(\mathcal{D}_i, s), \\
\hat{\Xi}_2 &= \left( \frac{1}{N_a(s)} \sum_{j \in I_a(s)} (Y_j - \hat{\tau} \tilde{A}_j) \right) - \left( \frac{1}{N_0(s)} \sum_{j \in I_0(s)} (Y_j - \hat{\tau} \tilde{A}_j) \right), \\
\tilde{\Xi}_a(\mathcal{D}_i, s) &= \hat{\mu}(a, s, X_i) - \hat{\mu}(0, s, X_i) + [Y_i - \hat{\mu}(a, s, X_i)] \times \frac{\hat{\pi}_a(s) + \hat{\pi}_0(s)}{\hat{\pi}_a(s)}, \\
\tilde{\Xi}_0(\mathcal{D}_i, s) &= \hat{\mu}(a, s, X_i) - \hat{\mu}(0, s, X_i) - [Y_i - \hat{\mu}(0, s, X_i)] \times \frac{\hat{\pi}_a(s) + \hat{\pi}_0(s)}{\hat{\pi}_0(s)},
\end{aligned}$$

where

$$\tilde{A}_i = \begin{cases} 1, & \text{if } A_i = a, \forall a \in \mathcal{A} \\ 0, & \text{otherwise.} \end{cases}$$

## 2 CAR with Clusters

Using the same reasoning as before, we can generalize the ATE and variance estimators for the case with clusters and multiple treatments as follows.

### ATE Estimator

$$\hat{\tau}_a = \frac{1}{\sum_{j \in I_{\{a,0\}}} N_j} \sum_{g \in I_{\{a,0\}}} \hat{\Xi}_g,$$

$$\hat{\Xi}_g = (\hat{\pi}_0(S_i) + \hat{\pi}_a(S_i)) \times \left[ \frac{\tilde{A}_g (N_g \bar{Y}_g - \hat{\mu}(a, S_g, X_g, N_g))}{\hat{\pi}_a(S_g)} - \frac{(1 - \tilde{A}_g) (N_g \bar{Y}_g - \hat{\mu}(0, S_g, X_g, N_g))}{\hat{\pi}_0(S_g)} \right]$$

$$+ \hat{\mu}(a, S_g, X_g, N_g) - \hat{\mu}(0, S_g, X_g, N_g),$$

where

$$\tilde{A}_g = \begin{cases} 1, & \text{if } A_g = a, \forall a \in \mathcal{A} \\ 0, & \text{otherwise.} \end{cases}$$

### Variance Estimator

$$\hat{\sigma}_a^2 = \frac{\frac{1}{g_a + g_0} \sum_{g \in I_{\{a,0\}}} [\tilde{A}_g \hat{\Xi}_a^2(\mathcal{D}_g, S_g) + (1 - \tilde{A}_g) \hat{\Xi}_0^2(\mathcal{D}_g, S_g) + \hat{\Xi}_2^2(\mathcal{D}_g, S_g)]}{\left( \frac{1}{g_a + g_0} \sum_{g \in I_{\{a,0\}}} N_g \right)^2},$$

where

$$\hat{\Xi}_a(\mathcal{D}_g, s) = \tilde{\Xi}_a(s) - \frac{1}{G_1(s)} \sum_{j \in I_1(s)} \tilde{\Xi}_{1,j}(s) - \hat{\tau} \left( N_g - \frac{1}{G(s)} \sum_{j \in I(s)} N_j \right),$$

$$\hat{\Xi}_0(\mathcal{D}_g, s) = \tilde{\Xi}_0(s) - \frac{1}{G_0(s)} \sum_{j \in I_0(s)} \tilde{\Xi}_{0,j}(s) - \hat{\tau} \left( N_g - \frac{1}{G(s)} \sum_{j \in I(s)} N_j \right),$$

$$\hat{\Xi}_2(s) = \left( \frac{1}{G_1(s)} \sum_{j \in I_1(s)} N_j \bar{Y}_j \right) - \left( \frac{1}{G_0(s)} \sum_{j \in I_0(s)} N_j \bar{Y}_j \right) - \hat{\tau} \times \left( \frac{1}{G(s)} \sum_{j \in I(s)} N_j \right),$$

$$\tilde{\Xi}_a(\mathcal{D}_g, s) = \hat{\mu}(a, s, X_g, N_g) - \hat{\mu}(0, s, X_g, N_g) + [N_g \bar{Y}_g - \hat{\mu}(a, s, X_g, N_g)] \times \frac{\hat{\pi}_a(s) + \hat{\pi}_o(s)}{\hat{\pi}_a(s)},$$

$$\tilde{\Xi}_0(\mathcal{D}_g, s) = \hat{\mu}(a, s, X_g, N_g) - \hat{\mu}(0, s, X_g, N_g) - [N_g \bar{Y}_g - \hat{\mu}(0, s, X_g, N_g)] \times \frac{\hat{\pi}_a(s) + \hat{\pi}_o(s)}{\hat{\pi}_a(s)},$$

where

$$\tilde{A}_g = \begin{cases} 1, & \text{if } A_g = a, \forall a \in \mathcal{A} \\ 0, & \text{otherwise.} \end{cases}$$