

The Cyclicalities of Earnings Growth along the Distribution - Causes and Consequences

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November 16, 2021

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Abstract

Earnings growth is more procyclical at the bottom of the income distribution than at the top. Using high-quality administrative data from Germany, I show that the heterogeneity is chiefly driven by transitions between employment and non-employment, specifically job-finding. I build a heterogeneous agent business cycle model that can rationalize these empirical findings. Agents in the model endogenously choose where to search for work in a labor market that features directed search. The model reproduces the heterogeneous procyclicality of earnings growth, as well as the contribution of job-finding, along the income distribution. I use this model to evaluate two policies aimed at reducing business cycle risk: countercyclical hiring subsidies and universal basic income (UBI). The first policy proposal increases welfare relative to the baseline economy. Implementing UBI decreases the volatility of aggregate consumption but decreases welfare overall.

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1 Introduction

Recessions disproportionately affect the incomes of the poor. More generally, their earnings growth follows the business cycle more closely than does the earnings growth of the rich (Guvenen et al., 2017). I show that the heterogeneity in income growth over the business cycle is mainly driven by individuals transitioning from nonemployment to employment. To investigate the welfare consequences of this phenomenon, I propose a heterogeneous-agent business cycle model that generates heterogeneous procyclicality in earnings growth driven by job-finding. I use this model to evaluate two policies aimed at reducing business cycle risk: countercyclical hiring subsidies and universal basic income (UBI). I find that, in the model, hiring subsidies in recessions can increase welfare. A UBI decreases it.

Using data on US earnings, Guvenen et al. (2017) show that the yearly earnings growth at the bottom of the permanent income distribution comoves more with GDP growth than earnings growth at the top. Utilizing administrative microdata from Germany, I find a similar pattern: as aggregate earnings growth increases by 1 percentage point, earnings growth for individuals in the lowest decile rise by more than 3 percentage points. In the top decile, earnings growth rises by less than 1 percentage point. I extend previous research by forensically investigating the sources of the heterogeneity, since it has implications for the distribution of business cycle risk along the income distribution. My main empirical contribution to the literature is the decomposition of earnings growth procyclicality into the contributions of different labor-market transitions. I find that the heterogeneity in earnings growth over the business cycle is driven by people transitioning from nonemployment to employment. First, I show that earnings at the bottom of the income distribution are more procyclical due to changes in employment status, i.e., because of individuals transitions along the extensive margin of the labor market. Second, within these transitions, job finders, i.e., individuals finding work out of unemployment, are responsible for the larger procyclicality of earnings growth at the bottom of the income distribution. To my knowledge, I am the first to document this decomposition.

Looking beyond earnings and at the labor market transitions themselves, I uncover the driving forces behind this result. First, at the bottom of the permanent income distribution, the probability of finding a job is much lower than at the top: within one year, individuals in the first decile transition from nonemployment to employment with a probability of less than ten percent, while the same probability is 20 percent in the top decile. This leads to a higher incidence of nonemployment at the bottom of the distribution. The second factor is the fact that job-finding is moves with aggregate earnings: as aggregate earnings rise by one percentage point, the probability of finding a job rises by 0.75 percentage points. This estimate is constant for the bottom half of the distribution. Combined, these two forces lead earnings growth at the low end of the permanent income distribution to comove more with the business cycle.

The second part of the paper presents a macroeconomic model featuring idiosyncratic and aggregate risk, and a frictional labor market, aiming to evaluate labor market policies which address the heterogeneous comovement of earnings and the business cycle. The model features heterogeneous, risk-averse workers who search for jobs across submarkets which differ in offered wages and job-finding probabilities. Workers differ along three dimensions: productivity, wealth and wages. When deciding where to search for jobs, they endogenously trade off higher job-finding probabilities for higher wages. Hence, in the model, both worker choices and firm choices about vacancy posting influence earnings growth. Agents can smooth consumption and insure against earnings fluctuations by saving in a risk free asset.

During an employment spell, workers receive a wage which is fixed for the duration of the spell. In unemployment, they receive unemployment benefits and produce a value of home production. In this setup, highly productive workers have strong incentives to find work quickly, for two reasons. First, their potential market wages substantially exceed their income in unemployment, i.e. they are missing out on potential earnings. Second, workers can insure against a fall in their productivity by entering a match, as the wage is fixed for the duration of the match. Workers achieve this by searching in submarkets which offer relatively low wages, relative to productivity. On the other hand, low productivity workers have high reservation wages, relative to their productivity, because for them benefit income and the value of home production have a meaningful impact. This leads them to search in submarkets which offer low job-finding probabilities. These forces endogenously create heterogeneous job-finding probabilities along the income distribution, which is a key feature of the data. I calibrate the model to match moments of the German labor market, including heterogeneous job-market transition rates in steady state, but leave earnings growth and fluctuations over the business cycle untargeted.

I solve for the model's response to a sequence of aggregate productivity shocks. It successfully reproduces a strongly decreasing procyclicality of earnings growth along the income distribution, as well as the heterogeneous contribution of job-finding. The main drivers behind the result are the lower job-finding probabilities for low productivity workers, combined with procyclical vacancy posting by firms. In a recession, as vacancy posting decreases, this effect has a disproportional impact at the bottom of the income distribution, where non-employment is most concentrated. While the model matches the procyclicality of earnings growth well in the first two deciles, it produces a steeper slope along the income distribution, thus under predicting the earnings betas at the top of the distribution. Further, the contribution of job-finding implied by the model is larger than in the data, as it does not diminish as quickly in the higher permanent income quantiles.

To evaluate the welfare consequences of heterogeneous procyclicality in earnings growth, I introduce two labor market policies aimed at mitigating it into the model. Each of the two policies is aimed at one of the two drivers of the model results discussed above, i.e.

procyclical vacancy posting by firms, which is exogenous to workers, and heterogeneous job-finding probabilities in steady state, which is the result of optimal worker choices.

The first policy is a countercyclical hiring subsidy, aimed at reducing the risk of lower vacancy posting during recessions, disproportionately borne by low-productivity workers. It aims to lessen the impact of aggregate productivity fluctuations on the unemployed. In this counterfactual, firms receive a one-time lump-sum transfer from the government when a match is formed, with the size of the subsidy indexed to aggregate productivity. In case of a negative productivity shock of 1 percent, the subsidy offsets the vacancy posting cost. Over the business cycle, the hiring subsidy reduces the volatility of consumption inequality. The reason is that vacancy creation does not fall as much as in the baseline case, thus enabling workers to find matches at higher rates, especially as the economy transitions back to steady state. This policy experiment shows that reducing the procyclicality of vacancy creation has welfare benefits, because it reduces a risk factor which is exogenous to the worker.

The second policy experiment eliminates unemployment benefits and in its stead introduces a UBI into the economy. This policy has received considerable attention in the German media. I calibrate the unconditional payment such that it is financed with the same steady state tax liability as the unemployment benefits in the baseline economy, a value of close to 450 euros. I find that this policy strongly increases job-finding, as, at the lower end of the income distribution, it removes the disincentive to work caused by unemployment benefits. At the top of the distribution, job-finding rises because the payments received in unemployment are lower than those under the baseline system with benefits indexed to productivity. I find that UBI decreases the cyclicity of earnings and consumption along the business cycle. However, aggregate welfare is lower compared to the baseline economy, as the insurance value of the unconditional payments is low, forcing workers to self-insure more than they would in the baseline economy. Workers in the baseline economy are willing to forego 6 percent of their consumption in order to avoid living in the UBI economy. This policy experiment indicates that attempting to reduce the procyclicality of earnings growth at the bottom of the income distribution can have negative welfare consequences. In the baseline economy, agents optimally choose submarkets with low job-finding probabilities, because their potential wage earnings are not high enough to compensate them for leaving unemployment. A UBI incentivizes these workers to find work, but at a high welfare cost.

Related Literature The empirical part of this paper is related to a fast growing literature documenting how earnings growth rates change over the business cycle.¹ [Storesletten et al. \(2004\)](#) find that earnings risk is highly countercyclical in the US, as measured by

¹A related literature studies earnings growth heterogeneity while not focusing on cyclicity (see [De Nardi et al., 2021](#); [Guvenen et al., 2014](#); [Halvorsen et al., 2019](#)). Similarly, there is a growing literature showing heterogeneous incidence of monetary policy along the earnings distribution (see [Broer et al., 2020](#); [Holm et al., 2020](#))

the variance of earnings growth. Using administrative tax data [Guvenen et al. \(2014\)](#) find that, while the variance of earnings growth rates is acyclical, higher order moments such as skewness and kurtosis are cyclical, as well as heterogeneous across the income distribution. Similar results have been found for other countries, such as Germany, France and Sweden ([Busch et al., 2018](#)), Denmark ([Harmenberg and Sievertsen, 2017](#)), as well as Italy ([Hoffmann and Malacrino, 2019](#)). [Pruitt and Turner \(2020\)](#), find evidence that in the US, intra-family insurance and added worker effects lead to less procyclical earnings growth skewness within households, compared to singles. Relatedly, [Guvenen et al. \(2017\)](#) and [Dany-Knedlik et al. \(2021\)](#) show that in the US, "earnings betas", i.e. the correlation between individual and aggregate earnings growth, are higher at the bottom of the income distribution. I extend these analyses by separately studying the cyclicalities of earnings growth rates and labor market transitions in the context of Germany. Similar to previous research I find that earnings growth is more cyclical towards the bottom of the income distribution. Crucially, I show that a large part of heterogeneous earnings growth cyclicalities can be explained by heterogeneous extensive margin transition probabilities.

Further, this paper is related to a literature seeking to account for the empirical regularities in earnings processes outlined above using theoretical models. [Hubmer \(2018\)](#) shows that a job-ladder model with savings, risk aversion, skill depreciation in unemployment and endogenous search effort can account for the stylized facts documented by [Guvenen et al. \(2015\)](#). Similarly, [Karahan et al. \(2019\)](#) document heterogeneous labor market transition patterns across the lifetime earnings distribution. They also show that a model with ex-ante worker heterogeneity in learning ability and job ladder risk can account for them. [Harmenberg and Sievertsen \(2017\)](#) show that a job-ladder model can replicate the procyclical skewness in Danish earnings growth data. This paper features the first model which replicates heterogeneous earnings betas along the earnings distribution.

The model presented in this paper is a directed search model similar to the one presented in [Chaumont and Shi \(2018\)](#), but featuring idiosyncratic worker-level productivity shocks. The model allows for tractability, even in the presence of borrowing constraints, labor market search and unemployment risk due to a feature referred to as a block recursive equilibrium ([Menzio and Shi, 2010](#)). In these models, the policy functions of firms and consumers can be solved for separately from the distribution of workers in equilibrium. [Chaumont and Shi \(2018\)](#) use their model to show that directed search models that feature precautionary savings can generate more realistic wealth distributions than other models. Along similar lines, [Eeckhout et al. \(2020\)](#) show that the interplay between wealth and search can explain labor market sorting into differently productive jobs. [Birinci and See \(2019\)](#) use a directed search model to show that time-varying unemployment benefits can help smooth consumption in recessions. I add to this literature by showing that directed search models combined with idiosyncratic worker productivity can explain patterns of labor market transition rates found in the data, in steady state and along the transition

path.

Finally, this paper relates to a long literature investigating policies which stabilize the macroeconomy during a recession (see, e.g., [McKay and Reis, 2016](#)). I evaluate the effectiveness of hiring subsidies, similar to [Kitao et al. \(2011\)](#), who show that this instrument can be particularly powerful when labor markets are weak. [Luduvise \(2019\)](#) studies the macroeconomic effects of implementing a UBI system instead of a means-tested system in the US. He finds that an expenditure neutral reform results in a welfare loss, which is similar to my findings. My findings are also related to [Fabre et al. \(2014\)](#), who find that the insurance value of unemployment insurance is considerably higher than that of UBI. The same is true in the model I present.

The rest of the paper is organized as follows. Section 2 presents the empirical analysis of the paper. Section 3 outlines the model, Section 4 describes the calibration. Section 5 presents the model results. Section 6 discusses the policy experiments and section 7 concludes.

2 Empirical Analysis

The empirical analysis in this paper is inspired by [Guvenen et al. \(2017\)](#). Using administrative social security data, I estimate individuals' exposure to the business cycle by estimating what they term "earnings betas" across the earnings distribution. The approach, inspired by the Capital Asset Pricing Model (CAPM), regresses individual earnings growth on aggregate earnings growth within a section of the permanent income distribution:

$$\Delta y_{i,t+1} = \alpha_q + \beta_q \Delta Y_{t+1} + \epsilon_{i,t}$$

where $\Delta y_{i,t}$ represents individual i 's yearly log-growth in pre-tax earnings between years t and $t + 1$ and ΔY_{t+1} represents GDP-growth in the same period. They find that the regression coefficient β_q is about three times larger at the bottom of the income distribution, compared to the top, i.e. a 1% increase in aggregate earnings growth leads to a 3% increase in individual earnings growth at the bottom of the distribution, but only a 1% increase towards the top.

In this section, I extend their analysis by decomposing the earnings betas into the contributions by different labor market transitions in order to pinpoint the drivers of the heterogeneity. For the case of Germany, I first document that, like in the US, earnings growth towards the bottom of the income distribution is considerably more procyclical compared to the top. The magnitudes in the regression coefficients are very similar across the two countries. To investigate the causes of the heterogeneity, I investigate the earnings betas across multiple subsamples of labor market transitions. When restricting

the sample to job-stayers, i.e. excluding the extensive margin, procyclicality is close to homogeneous across the distribution and earnings growth moves less than one-for-one with the aggregate. Along the extensive margin, I then identify the contributions of job-finding and separations. Excluding separations reduces the procyclicality of earnings growth, but homogeneously so. Excluding job-finders from the sample, however, reduces the regression coefficients β_q at the bottom of the distribution considerably, implying that this transition is responsible for most of the observed heterogeneity.

2.1 Data

For the empirical exercise, I utilize the Sample of Integrated Labor Market Biographies (SIAB), constructed by the Research Data Center (FDZ) of the German Federal Employment Agency (BA). It contains administrative data on a representative two percent sample of German labor market histories between 1975 and 2014, with the exception of civil servants, students and self-employed individuals.²

The dataset is divided into labor market spells, which mark distinct episodes during an individuals labor market biography. For each such spell, the data provides information about its start- and end-date, as well as the average daily earnings throughout its duration. Since firms are required to notify the responsible social security agencies about their employees at least once per year (Ganzer et al., 2017), the maximum spell length is one year.

For employed individuals, the earnings measure is pre-tax earnings liable to social security contributions; for non-employed individuals, some unemployment benefit receipts are reported.³ The dataset does not report spells of periods of non-employment without benefit receipts.⁴ To construct complete labor market biographies, I declare any missing values at any point during an individuals life as non-employment episodes, except for those before their first or after their last non-missing episode in the data.⁵

I deflate all earnings using the quarterly consumer price index.⁶ For the subset of non-employed individuals who receive unemployment benefits, I also observe benefit

²The excluded individuals make up around 20% of the workforce (Busch et al., 2018).

³ Several forms of unemployment benefits are reported: i) Unemployment Benefits I (*Arbeitslosengeld I*), which represent an earnings replacement payment immediately after job-loss, for a limited time, (ii) Unemployment Help (*Arbeitslosenhilfe*) benefits claimable after the exhaustion of unemployment benefits, (iii) Unemployment Benefits II (*Arbeitslosengeld II*), the successor of the Unemployment Help program, implemented through the Agenda 2010 program in 2005. There is no data on Unemployment Benefits II between the years 2004 and 2007.

⁴For the latter years of the sample period, the data also contains information on job-search and marginal part-time employment. To keep the sample consistent over time, individuals in these categories are assigned to non-employment.

⁵For individuals younger than 60 years, I also declare as non-employed all episodes after their last employment observation if that observation falls into the years 2013 and 2014. The number of non-employed individuals would fall drastically towards the end of the sample otherwise.

⁶The CPI index (base year 2015) is obtained from the Federal Reserve Bank of Saint Louis' FRED database, and is not seasonally adjusted. The variable ID is DEUCPIALLQINMEI.

income. However, since the duration and generosity of unemployment benefits changed throughout the sample, the series is not consistent over time. Hence, in the baseline estimation, the earnings of all non-employed individuals are set to zero. In the appendix, I perform a robustness exercise including Unemployment Benefits I as an income measure for the unemployed.

I convert average daily earnings during labor market spells to quarterly earnings. To obtain a binary indicator for quarterly employment status, I assign individuals who are employed for more than half of a quarter into employment and all others into non-employment. Employed individuals are assigned the average earnings value of the longest employment spell during the quarter.⁷

Because social security contributions are capped at the assessment ceiling for pension insurance, the earnings data are censored from above, affecting around 6 % of the total observations. In most of the empirical analysis that follows, I focus on the lower 90% of the earnings distribution but impute top-earnings following [Card et al. \(2013\)](#) and [Dustmann et al. \(2009\)](#). Their approach relies on predicting the censored observations using a Tobit estimation.⁸ Note that because the imputed earnings are randomly drawn from estimated distributions, the earnings growth they produce for each individual is not informative and the earnings growth rates among the top 10 percent should be interpreted with this in mind.

To make the data consistent over time, I first exclude individuals employed in the states of the former German Democratic Republic. Second, marginal part-time workers are declared as non-employed since they are only registered in the dataset after 1999 and are previously unobserved. Third, due to large fluctuations in the number of individuals whose employment status is coded as "Other employment status" ([Ganzer et al., 2017](#)), I label all of these as non-employed in all years of the sample. For my analysis, I restrict the sample period to 1980-2014 and to individuals older than 25 years and younger than 60, leaving close to 70 million person-quarter observations.

2.2 Descriptive Statistics

For each quarter, I generate an income distribution by sorting individuals according to their recent earnings, following [Guvenen et al. \(2014\)](#). Recent earnings are constructed by averaging individual quarterly labor earnings (including zeros) over five years prior

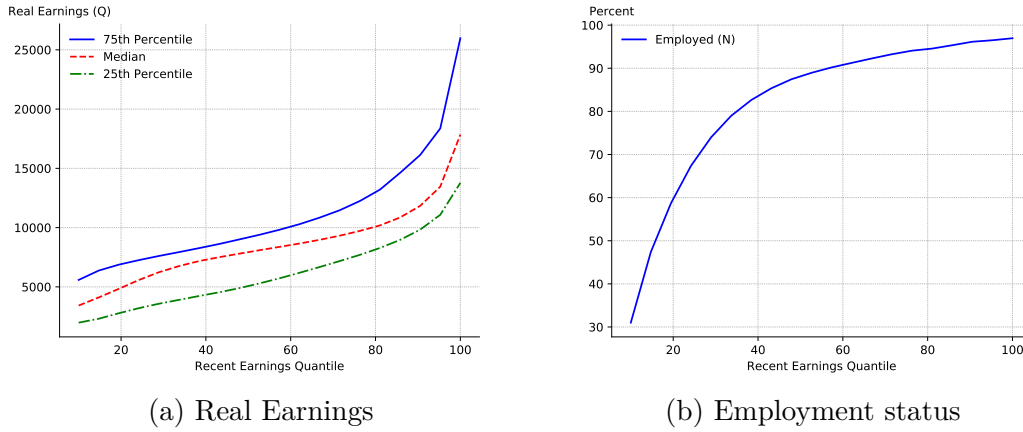
⁷Note that individuals who are non-employed for less than a full quarter and employed for the rest are declared to be employed. Employment spells that are longer than one quarter are split into multiple quarters. The earnings for each quarter are the earnings of the overarching spell.

⁸A second approach fits a Pareto distribution to the upper tail of the earnings distribution and draws replacements for the censored earnings observations using this method. Drawing on a separate dataset, [Dustmann et al. \(2009\)](#) conclude, that the Tobit approach is preferable. In the context of my exercise, results are very similar.

to quarter t .⁹ This measure is intended as a proxy for permanent income (the results are similar using lifetime earnings as a sorting variable) and individuals are split into 20 quantiles along its distribution, trading off noise reduction (larger quantiles) and granularity (smaller quantiles). I exclude individuals who do not receive any labor earnings during the five-year period. Furthermore, because age and gender are influential predictors of recent earnings, I construct the quantiles outlined above conditional on these two characteristics, by assigning individuals to 5-year age brackets.

The left-hand panel of Figure 1 shows the median of real quarterly earnings in quarter t , as well as the 25th and 75th percentile, conditional on employment. Median labor earnings are about three times larger at the top, compared to the bottom. Interestingly, the slope of the earnings curve starts out relatively steep, but flattens out around the median of the recent earnings distribution, steepening again towards the top.

Figure 1: Earnings and Employment across the distribution



Note: The *Left Panel* shows the median, 25th and 75th percentile of real gross quarterly labor earnings by quantile, deflated using the CPI with base year 2005. The *Right Panel* shows the average employment share by quantile. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

The right-hand panel of Figure 1 plots the share of employed individuals across the distribution. While only close to 30 % of individuals in the first quantile are employed according to the definition outlined above, the fraction rises to 80 % at the 4th decile and to above 90 % beyond the median.

Next, I compute average earnings growth by quantile. Here, I make a crucial modification to the approach used by [Guvenen et al. \(2017\)](#): I include the extensive margin of labor market transitions. I average real earnings by quantile (including zeros) and

⁹As a robustness exercise, I construct a recent earnings distribution excluding zeros in the appendix and perform the same estimations as in the baseline. The results are very similar.

compute the growth of this measure over time (Krueger et al., 2016):

$$\Delta \bar{y}_{t,k}^q = \log(\bar{y}_{t+k}^q) - \log(\bar{y}_t^q) \quad (1)$$

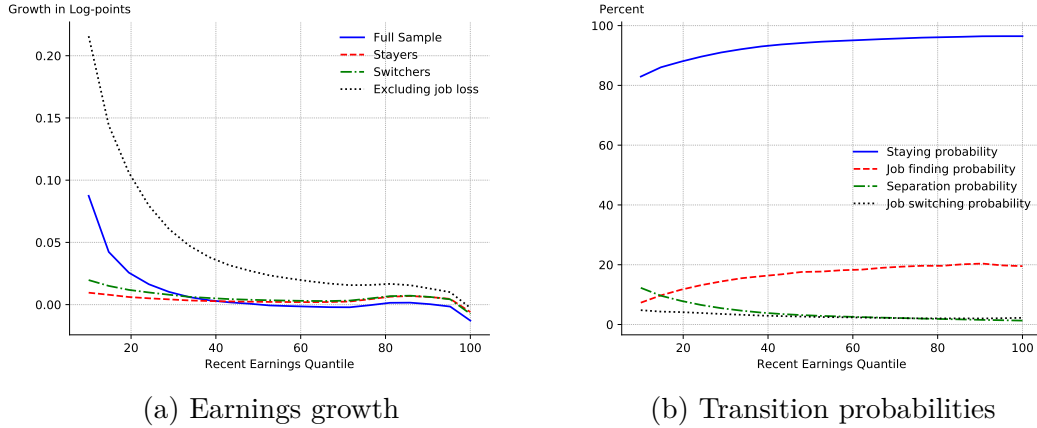
$$\text{where } \bar{y}_{t+k}^q = \frac{1}{N_q} \sum_i y_{i,t+k} \quad \forall i \in q \text{ at } t$$

where \bar{y}_{t+k}^q represents average earnings in quarter $t+k$, for all individuals who are in quantile q in period t and consequently, $\Delta \bar{y}_t^q$ is the log-growth of average earnings. This approach has the advantage of being able to include zeros in a log-growth rate, as they are subsumed in the aggregate. When using individual growth rates, this is not possible.

Furthermore, my dataset allows me to compute earnings growth rates separately by labor market transitions. I consider five such transitions: job-stayers are individuals who remain with the same employer between periods t and $t+k$, job-switchers are individuals who are employed in periods t and $t+k$, but with different employers, job losers are individuals who are employed in quarter t but non-employed in period $t+k$, job finders are non-employed in period t and employed in period $t+k$ and finally, all other individuals are non-employed in both periods t and $t+k$. In what follows, I decompose average earnings growth over time along these dimensions.

The red line in the left-hand panel of Figure 2 displays quarterly earnings growth ($k=1$) across the earnings distribution. While earnings growth is as high as 10 % in the first quantile, it turns negative around the median, indicating mean reversion over time. First, I exclude all individuals who transition from employment to non-employment between quarters t and $t+1$ (black, dotted line). This removes all downside risk for the employed, but retains the upside potential for the non-employed. Unsurprisingly, earnings growth for this restricted sample is considerably higher, especially at the bottom, thus increasing growth by more than 0.1 log-points. At the top of the distribution, growth only increases by around 0.02 log-points when separators are excluded.

Figure 2: Earnings growth and Transitions across the distribution



Note: The *Left Panel* shows average earnings growth between t and $t + 1$ for various sub-samples, calculated according to Equation (1), by quantile. The blue line represents average earnings growth for the full sample, the black line excludes job-losers, the green line limits the sample to those who are employed in periods t and $t + 1$ and the red line only includes job-stayers. The *Right Panel* shows the average transition probabilities between employment states. The blue line represents the probability of staying employed, the red line represents the job-finding probability, the green line the probability of separation and the black line the probability of switching employers. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

Next, I remove all extensive-margin transitions, by restricting the sample to those who are employed in periods t and $t + 1$. This almost perfectly equalizes average earnings growth across the distribution, with now only slightly higher growth towards the bottom end of the distribution. Restricting the sample even further, to only include individuals who stay with the same employer, shows that the slightly higher growth rate at the bottom seems to be mainly driven by job-switchers.

Focusing more explicitly on labor market transitions, the right-hand panel of Figure 2 shows quarterly transition probabilities across the distribution. Focusing on the initially employed, there is a striking difference in separation probabilities between the bottom and the top of the earnings distribution: while the employed in the first quantiles face a 30 % chance of transitioning into non-employment, the probability is almost zero at the very top. Furthermore, while generally low, the job-switching probability is higher at the bottom as well. Shifting to the non-employed, the figure shows a strong upward slope in quarterly job-finding probabilities across the distribution, implying longer unemployment duration at the bottom. These forces cause the strong heterogeneity in employment status documented in the right-hand panel of Figure 1.

2.3 Earnings procyclicality across the distribution

Next, I document heterogeneity in the procyclicality of earnings by quantile, similar to [Guvenen et al. \(2017\)](#). I regress the measure of earnings growth introduced in Equation (1) on aggregate earnings growth, by quantile. As a measure for aggregate growth, I choose growth in average earnings constructed analogously to the quantile-specific growth measure:

$$\Delta Y_{t,k} = \log(\bar{y}_{t+k}) - \log(\bar{y}_t)$$

$$\text{where } \bar{y}_{t+k} = \frac{1}{N} \sum_i y_{i,t+k}$$

Using the growth rates constructed in this way, the procyclicality of earnings along the income distribution can be estimated using the following regression:

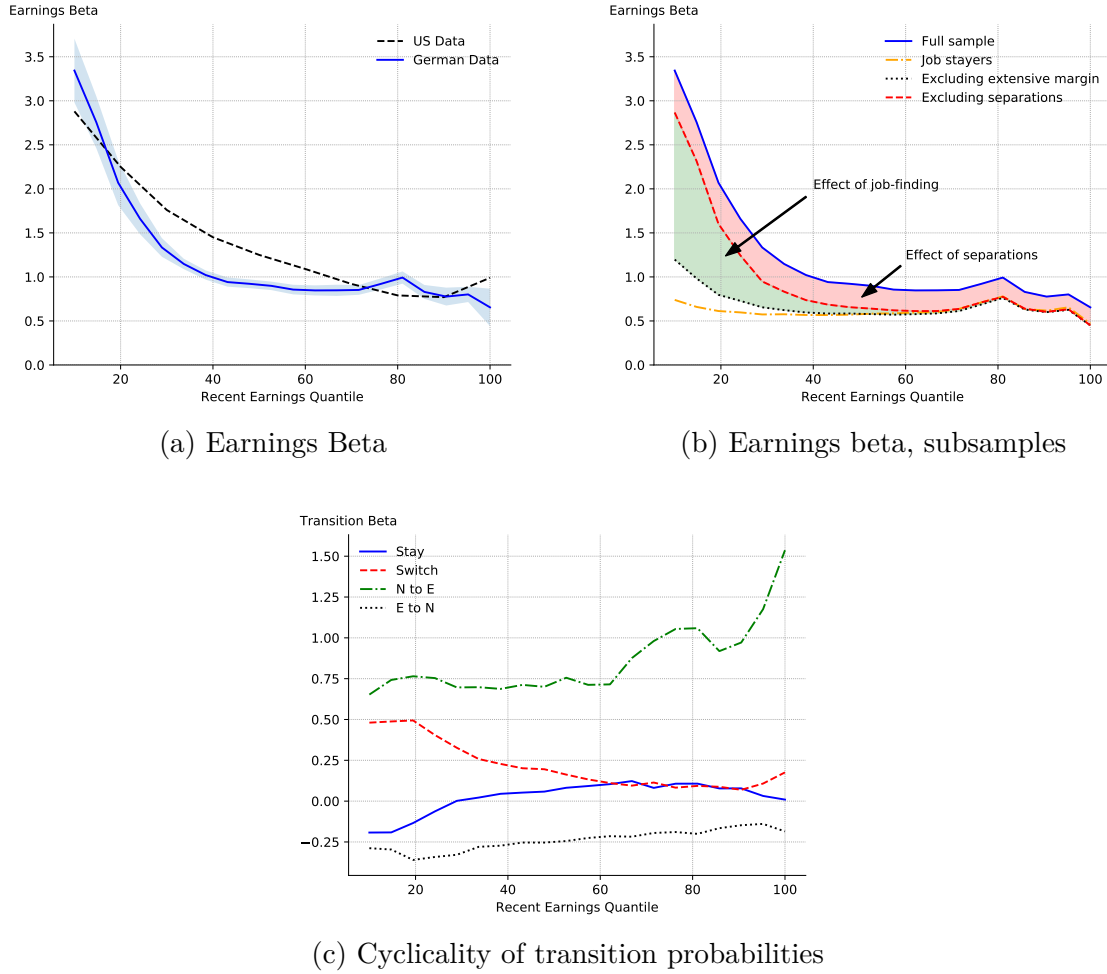
$$\Delta \bar{y}_{t,k}^q = \alpha + \beta_{Y,k}^q \Delta Y_{t,k} + X_t + \epsilon_{q,t} \quad (2)$$

where β_Y^q represents the change in quantile-specific growth in response to a 1 percentage point increase in aggregate growth and X_t contains calendar-quarter dummies to account for seasonality. For this analysis, I focus on yearly quarter-on-quarter growth rates, implying $k = 4$.

The left-hand panel of Figure 3 plots the values of $\beta_{Y,4}^q$ obtained from Equation (2) across the earnings distribution. Individuals at the bottom of the distribution see their earnings growth rise by more than 3 percentage points, on average, for every additional percentage point in aggregate earnings growth. The same relationship is less than 1-for-1 beyond the 40th percentile, but increasing slightly again between deciles seven and eight. Beyond the 80th percentile, the graph shows a decrease in the earnings betas, likely driven by the imputation of top incomes.

These results are strikingly similar to the ones documented by [Guvenen et al. \(2015\)](#) for the case of the US. The black dashed line in the top-left panel of Figure 3 reports the earnings betas for males, aged 36-45, as reported in the online appendix Table B1 in [Guvenen et al. \(2015\)](#). In both countries, earnings growth is about three times as procyclical at the bottom of the income distribution as it is at the top. The slope is steeper in Germany, implying that around the median, earnings growth is more procyclical in the US.

Figure 3: Cyclicity of Earnings Growth



Note: The *Left Panel* plots the coefficients $\beta_{Y,k}^q$ from Equation (2) (blue), by quantile. The shaded area represents 95 % confidence intervals. The dashed black line reproduces the GDP earnings beta for men aged 36-45, as reported by [Guvenen et al. \(2017\)](#). The *Right Panel* shows the estimates for the coefficient $\beta_{Y,k}^q$ from Equation (2) for two additional subsamples. The blue line utilizes the full sample, the red, dashed line excludes individuals who are employed in quarter t and non-employed in quarter $t+4$, the black, dotted line restricts the sample to those employed in both periods t and $t+4$. The area between the solid and dashed lines represents the change in the beta-coefficients upon the exclusion of individuals who separate. The shaded region marks the contribution of job-finding to the beta-coefficients. The *Bottom Panel* plots the coefficients $\beta_{Y,k,trans}^q$ in Equation (3). The blue line restricts the sample to job-stayers, the red line to job switchers, the green line to job-finders and the black line to separators. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

In order to investigate possible reasons for the stark heterogeneity in the procyclicality of earnings growth rates across the distribution, I proceed similarly to Figure 2, by restricting the sample to subgroups and investigating the effect on the regression-coefficients in Equation (2). The right-hand panel in Figure 3 shows the result of this exercise. When restricting the sample to individuals who stay in the same job between quarters t and $t+4$, the coefficients are close to homogeneous along the earnings distribution. Earnings are procyclical for job-stayers, but their earnings grow only half as much as aggregate earnings.

I interpret the difference between the beta coefficients estimated for the full sample on the one hand and the job-stayers on the other as the contribution of the extensive margin. The graph suggests that this contribution is considerably larger towards the bottom of the distribution than it is at the top.

Next, I reintroduce individuals who change employers or who are non-employed in periods t and $t + 4$. This exercise allows some labor market transitions, but excludes the extensive margin, i.e. individuals who are employed in one period and non-employed in the other. This raises the coefficients slightly towards the bottom of the income distribution. From this, I conclude that job-switching is driving some of the earnings procyclicality there, while the effect is almost zero at the top.

Decomposing the contribution of the extensive margin into the contributions by job-finders and job-separators presents a challenge. It is not possible to compute earnings growth rates for either group, as all job-finders, by definition, have zero earnings in period t , and all separators earn zero in period $t + 4$. In order to still pin down which margin is responsible for the strong heterogeneity in earnings growth rates across the distribution, I restrict the sample in steps. First, I focus on all individuals who do **not** transition from employment in quarter t to non-employment in quarter $t + 4$ (separations) and re-estimate the regression in Equation (2). The resulting coefficients are plotted as the red, dashed line in the right-hand panel of Figure 3b. Across the distribution, the earnings betas for this subsample are lower than they are for the full sample. The shift is almost homogeneous, with coefficients decreasing by between 0.25 and 0.5 points. Therefore, the contribution of separations to the overall beta-estimates, appears to be fairly similar between the bottom and the top of the recent earnings distribution. Importantly, the earnings betas estimated for the described subsample are still steeply decreasing from close to 3 in the first decile to 0.5 around the median of the income distribution.

Crucially, the beta-coefficients estimated for the above subsample also identify the contribution of job-finders to the earnings beta. In the right-hand panel of Figure 3b, this is the shaded green area between the dotted black line, the beta-coefficients for the subsample of job-stayers, and the dashed red line, the coefficients when the estimation sample includes job-stayers and job-finders (but excludes job-separators). The difference between the two lines is due to the addition of job-finders. In the first decile of the recent earnings distribution, the addition of job-finders increases the beta coefficient by more than 2 percentage points, implying that job-finding drives a large share of the procyclicality of earnings. The effect decreases towards the median and is close to zero approaching the seventh decile.

From this decomposition, I conclude that the main reason for the observed heterogeneity in the procyclicality of earnings growth rates along the income distribution is job-finding. This fact is not unexpected, as Figure 1 shows that non-employment is considerably more prevalent towards the lower end of the income distribution. It remains to be shown,

however, that job-finding is procyclical at the bottom of the recent earnings distribution. To this end, similarly to Equation (2), I estimate a regression of transition probabilities on aggregate earnings:

$$\Delta \overline{trans}_{t,k}^q = \gamma + \beta_{Y,k,trans}^q \Delta Y_t + X_t + \varepsilon_t \quad (3)$$

where $\overline{trans}_{t,k}^q = \frac{1}{N_q} \sum_i trans_{i,t+k} \forall i \in q$

where \overline{trans}_t^q is the fraction of all individuals in quantile q at quarter t who transition along *trans*. I focus on four transition paths (*trans*): (i) staying with the same employer, (ii) switching to a new employer, (iii) employed to non-employed and (iv) non-employed to employed.

The bottom panel of Figure 3 plots the values of $\beta_{Y,4,trans}^q$ from Equation (3) across the recent earnings distribution. Both staying and job-switching are moderately more procyclical at the low end of the distribution, compared to the top. Job-separations are countercyclical everywhere, but slightly more so at the bottom. Crucially, in the probability of transitioning from non-employment to employment (green line), a one percentage point increase in aggregate earnings growth leads to a 0.75 percentage point increase from the bottom of the distribution up to the 6th decile. Beyond that, the cyclicity is increasing.

Together, these results explain the heterogeneous incidence of job-finding in driving earnings procyclicality. Non-employment is considerably more common at the low end of the recent earnings distribution. This is driven by higher separation rates and lower job-finding (right-hand panel in Figure 2). In times of a rise in aggregate earnings, job-finding increases across the distribution, driving large income gains at the bottom, but almost none at the top, as there are very few non-employed workers there.

The previous section outlines the fact that earnings are heterogeneously procyclical, an effect driven by a heterogeneous impact of job-finding along the income distribution. In order to assess the consequences of these findings for welfare and policy, I build a model that can reproduce the effects observed in the data, while accounting for important margins of insurance like savings and unemployment benefits. The next section provides such a model.

3 Model

In order to conduct welfare analyses of policy proposals, I provide a model that mirrors the empirical earnings beta and its decomposition. Time is discrete and infinite. The model is a directed search model in the spirit of Chaumont and Shi (2018), bringing together a unit mass of workers and a continuum of firms. The model features heterogeneous agents, who

differ along three dimensions: productivity, wealth and wages. In the model, job-finding is an endogenous outcome of agents' decisions and can therefore be affected by labor market policies. Due to search frictions, the model features involuntary unemployment, endogenous to the aggregate state of the economy; and due to borrowing constraints, households are unable to perfectly insure against this risk.

3.1 Environment

The consumers in the model are risk averse, receive utility from consumption and discount the future at rate β . Hence, their expected utility is

$$U_t = \sum_{s=t}^{\infty} \beta^{s-t} u(c_s)$$

The utility function $u : \mathbb{R}_+ \rightarrow \mathbb{R}$ is twice differentiable, strictly increasing, strictly concave and $u'(0) = \infty$. All consumers can save in a risk-free asset a , which pays interest rate r . Borrowing is not possible.

Households can either be employed or unemployed, with the employed supplying one unit of labor inelastically and earning $(1 - \tau)w$. The variable w represents the constant wage associated with their firm-worker match and τ is a labor income tax imposed by the government. In equilibrium, wages differ across workers, giving rise to an earnings distribution as in the data. The unemployed receive unemployment benefits $b(z)$, indexed by individual productivity, and produce h of home production. Consequently, workers differ according to their asset holdings, their employment status, their productivity and their wages.

Each individual is endowed with productivity z_t , which evolves according to the Markov process

$$\begin{aligned} \log(z_t) &= \rho_z \log(z_{t-1}) + \varepsilon_t \\ \varepsilon_t &\sim N(\mu_z, \sigma_z) \end{aligned}$$

Labor market search in the model is directed. Households, employed and unemployed, can direct their search at wage submarkets indexed by productivity and wealth. In each submarket, workers only meet firms willing to employ them at the submarket-specific wage rate. Consequently, if a worker meets a (new) firm, there is no need for wage bargaining, and they start their relationship at the beginning of the subsequent period. Wages are fixed for the duration of the match. The unemployed can search every period, while the employed can only search with probability $0 < \Lambda < 1$, capturing the fact that most of their time is spent working.

Firms are owned by risk-neutral investors, maximize profits and discount the future at

the risk-free interest rate. They post vacancies across wage submarkets, taking wealth and productivity of the searching workers as given. The cost of posting a vacancy is κ and there is free entry. A firm-worker match produces according to the worker's productivity z and aggregate productivity A . Existing matches separate for two possible reasons: (i) exogenously with probability $\delta(z)$, or (ii) endogenously, if the worker finds a new job.

Submarkets are indexed by wealth, in order for firms to be able to make predictions about the workers' job-finding probabilities on-the-job.¹⁰ These affect expected match profits, which in turn influence which wage-submarket a firm enters. If submarkets were not indexed by worker wealth, firms would need to form expectations over the worker's wealth level given her productivity, requiring them to know the distribution of wealth conditional on productivity. Furthermore, upon meeting, firms and workers would need to bargain over the wage, since match surpluses would differ depending on the workers' assets.¹¹ This would make the problem less tractable and more costly to solve computationally.

If wealth was unobservable to the firm, workers in the model would have an incentive to signal their wealth to the firm (Chaumont and Shi, 2018). In equilibrium, asset-rich workers have an incentive to search for jobs in submarkets that offer higher wages at lower job-finding probabilities. Hence, they are less likely to leave their current match for a new one than wealth-poor workers. This, all else equal, is attractive to the firm, since it applies a longer match duration. Along these lines, Chaumont and Shi (2018) argue that workers who can signal to the firm any level of wealth less than or equal to their actual holdings, will choose to signal the true value. This, in turn, implies that submarkets will be indexed by the worker's wealth level.

Within each submarket, searchers and firms meet according to a matching function $M(S, V)$, where S and V represent the mass of searchers and vacancies within each submarket, respectively. The function is strictly increasing, concave in both arguments and exhibits constant returns to scale. If a vacant firm meets a worker, they form a match. The probability of a vacancy being filled can be expressed as $q(\theta) = M(S, V)/V = M(\frac{1}{\theta}, 1)$ where $\theta = \frac{V}{S}$ represents the submarket tightness. Likewise, the probability of a worker finding a match can be written as $\eta(\theta) = M(1, \theta)$. The vacancy filling probability $q(\theta)$ is decreasing in θ , while the job finding probability $\eta(\theta)$ is increasing in θ . If a vacant firm meets more than one worker, it randomly decides which one to hire. Submarkets that are not active because no worker or no firm visits them, have $\theta = 0$.

The government finances unemployment benefits and bond issuance through a tax on labor income τ . It balances its budget every period, following a budget rule, trading off fluctuations in bond issuance with tax rate changes. Further, the government taxes all

¹⁰If there was no on-the-job search, implying that the probability of an endogenous quit was zero, workers' assets would not enter the firm's problem. I choose to allow workers to search in order for the model to be able to match the labor market flows I observe in the data.

¹¹This problem is similar to Krusell et al. (2010), who solve a random search model with workers who can hold assets.

profits lump sum.¹²

A worker's state vector can be written as (L, z, x, w) , where $L \in \{U, E\}$ represents her employment status, $x \in \mathcal{X} \equiv [\underline{x}, \bar{x}] \subseteq \mathbb{R}_+$ represents her beginning of period wealth, $z \in Z \subseteq \mathbb{R}_+$ represents her idiosyncratic productivity realisation, and $w \in \mathcal{W} \subseteq [0, 1]$ is her piece-rate wage, if she is employed.

The aggregate state of the economy can be expressed as $\psi = (A, \Omega)$, with $A \in \mathbb{R}_+$ representing the aggregate productivity state of the economy and $\Omega : \{U, E\} \times \mathcal{X} \times Z \times \mathcal{W} \rightarrow [0, 1]$ denoting the distribution of agents across states. The aggregate state follows the law of motion $\psi' = Q(\psi)$.

Model timing: At the beginning of each period, the aggregate and individual productivity innovations are revealed. Subsequently, individuals who matched with a vacant firm in the previous period start their new jobs. In the next phase, matches produce and households consume and make their savings choices. In the last stage, firms post vacancies across submarkets and consumers choose the submarket in which to search. Lastly, separations and job-matching take place. In the notation that follows, the value functions of the workers are evaluated after production has taken place and wages were paid, but before the consumption/savings decision.

3.2 Firms

An unmatched firm can post a vacancy into a submarket indexed by the worker's productivity z , her cash-on-hand x and the posted wage w' . The value of a vacancy is

$$V_F(z, x, w'; \psi) = -\kappa + \frac{1}{1+r} \mathbb{E} [q(z, x, w'; \psi) J(z', x', w'; \psi') + (1 - q(z, x, w'; \psi)) V_F'] \quad (4)$$

The vacancy posting cost is κ , $q(z, x, w'; \psi)$ represents the probability that the firm will meet a worker and r is the interest rate for investments between period t and $t + 1$. If the vacancy is not filled, the firm can post a new vacancy in the subsequent period. Because the firm's owners are risk neutral, they discount the future at the risk-free interest rate.

A firm that meets a worker in a submarket starts producing output in the subsequent period. Its value is

$$\begin{aligned} J(z, x, w) = & Az - w + \frac{1}{1+r'} \mathbb{E} [\delta(z') V_{new} \\ & + (1 - \delta(z')) \eta(z, x, w) V_{new} \\ & + (1 - \delta(z')) (1 - \eta(z, x, w)) J(z', x', w)] \end{aligned} \quad (5)$$

¹²An alternative specification follows [Kaplan et al. \(2014\)](#), where each agent receives a fraction of total profits each period, in relation to their share of total risk-free investment. In my model, using such an approach would require guessing a path for aggregate profits along the transition path to aggregate productivity shocks, which would make the problem computationally much more costly to solve.

The firm produces according to aggregate productivity A and the worker's productivity z . With probability $\delta(z')$, the match dissolves exogenously before the production phase in the next period. If this is not the case, the worker finds a new match on-the-job with probability $\eta(z, x, w)$. In both cases, the firm has the opportunity to post a vacancy V_{new} again in the following period.

Because entry into all submarkets is free, in equilibrium, the value of posting a vacancy is driven to zero, $V_F = 0$. Utilizing this, it is possible to rewrite Equation (4) as

$$\frac{1}{1+r} \mathbb{E} [J(z', x', w'; \psi')] = \frac{\kappa}{q(\theta)}$$

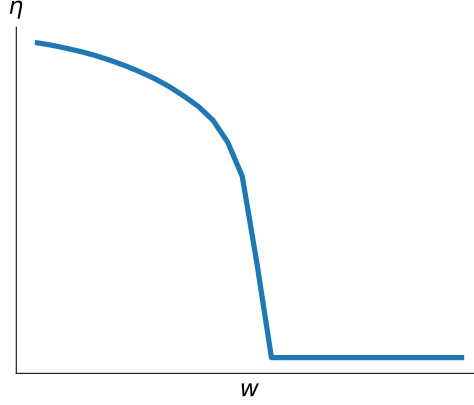
Intuitively, the expected value of a match must be equal to the vacancy posting cost, adjusted for the probability of meeting a worker, in each submarket. This equation pins down the relationship between the wage and tightness for each submarket, conditioning on the worker's current productivity z and the worker's savings choice:

$$\theta(z, x, w'; \psi) = \begin{cases} q^{-1} \left(\frac{\kappa(1+r)}{\mathbb{E}[J(z', x', w'; \psi')]} \right) & \text{if } \mathbb{E} [J(z', x', w'; \psi')] \geq \kappa(1+r) \\ 0 & \text{otherwise} \end{cases} \quad (6)$$

The inequality constraint implies that firms will not post vacancies in submarkets where the present discounted value of forming a match is too low to cover the vacancy posting cost, even if a match was guaranteed ($q(\theta) = 1$). In these markets, tightness is zero and no matches are formed. To save space, in what follows, I will suppress the dependencies of θ on current productivity z , current worker wealth x and the posted wage w' .

conditioning on productivity and cash-on-hand, equation (14) implies a downwards sloping relationship between the worker's job-finding probability and the wage across submarkets. Figure 4 illustrates this dependence. The x-axis displays wages offered across submarkets w , and the y-axis displays the job-finding probability implied by Equation (14). The latter decreases until it hits the wage value at which the expected discounted value of a potential match falls below the vacancy posting cost. At this point, submarket tightness is zero and no firm visits these submarkets.

Figure 4: Relationship between wage and job-finding



Note: This figure illustrates the relationship between the posted submarket wage w and the resulting job-finding probability η , as implied by Equation (14). The graph conditions on worker productivity z and cash-on-hand x .

3.3 Worker Problems

3.3.1 Unemployed

The unemployed enter each period with state variables productivity z and cash-on-hand x , and maximize their expected discounted utility. In recursive notation, their problem is

$$\begin{aligned}
 V^u(z, x; \psi) = \max_{a'} \quad & u(c) + \beta \max_{w'} \mathbb{E} \left[\underbrace{\eta(\theta) V^e(z', x'_n, w'; \psi')}_{\text{Find Work}} \right. \\
 & \left. + \underbrace{(1 - \eta(\theta)) V^u(z', x'_u; \psi')}_{\text{Stay Unemployed}} \right] \\
 \text{subject to} \quad & x'_u = (1 + r)a' + b(z') + h \\
 & x'_n = (1 + r)a' + (1 - \tau)w' \\
 & c + a' \leq x
 \end{aligned} \tag{7}$$

Using backward induction, households first solve which submarket w' to search in, taking as given the choice for consumption c and savings a' . As discussed in section 3.2, the free entry condition implies that, conditional on worker productivity z and wealth x , each submarket wage w implies a unique submarket tightness, which pins down a submarket specific job-finding probability $\eta(\theta)$. When choosing in which submarket to search, workers trade off higher wages w' , which promise higher continuation values V_{t+1}^w in case a match is formed, with higher job-finding probabilities. Note that unemployed households who successfully match to a firm cannot immediately separate exogenously. If an unemployed worker matches with a vacant firm, their cash-on-hand in the following period is x'_n .

Otherwise, they remain unemployed, receiving benefits $b(z)$ and producing h at home, leading to cash-on-hand x'_u .

In the first stage, knowing their choice $w'(a')$, workers choose the optimal consumption in each period, subject to their beginning of period wealth x , their productivity z , and the borrowing constraint.

3.3.2 Employed

Like the unemployed, the employed's problem consists of a consumption-savings choice and a choice of in which wage submarket to search. An employed worker enters the period with her individual productivity level z , cash-on-hand x and wage w . Recursively their problem can be formulated as

$$\begin{aligned}
V^e(z, x, w; \psi) = \max_{a', c} \quad & u(c) + \\
& \beta \max_{w'} \mathbb{E} \left[\underbrace{(1 - \delta(z'))(1 - \Lambda\eta(\theta))V^e(z', x'_e, w; \psi')}_{\text{Stay in same job}} \right. \\
& + \underbrace{(1 - \delta(z'))\Lambda\eta(\theta)V^e(z', x'_n, w'; \psi')}_{\text{Switch jobs}} \\
& \left. + \underbrace{\delta(z')V^u(z', x'_u; \psi')}_{\text{Separations}} \right] \\
\text{subject to} \quad & x'_u = (1 + r)a' + b(z') + h \\
& x'_e = (1 + r)a' + (1 - \tau)w \\
& x'_n = (1 + r)a' + (1 - \tau)w' \\
& c + a' \leq x
\end{aligned} \tag{8}$$

Analogous to unemployed's problem, an employed worker's problem can be solved by backward induction. The employed first choose the wage submarket in which they want to search on-the-job, taking savings and consumption decision as given. Each period, the employed are only able to search with probability Λ . Further, with exogenous probability $\delta(z')$, an employed worker becomes unemployed at the beginning of the next period, irrespective of their search outcome.

The employed search in the same submarkets as the unemployed, conditional on their respective savings decisions and productivity. Hence, with probability $\eta(\theta)$, their search is successful and they move to a new wage w' starting in the next period (unless they get separated exogenously). If their job-search is unsuccessful, or they are unable to search, they continue working with their current employer.

In equilibrium, workers who enjoy high levels of idiosyncratic productivity search

in submarkets with higher job-finding probabilities and lower wages, relative to their productivities. The reasons are two-fold. First, relative to their productivity level, low-productivity households are more productive at home than their high-productivity counterparts. For the former, this acts as an insurance mechanism, making them demand higher wages (and incur lower job-finding probabilities). However, the level of h is too low to insure highly productive workers, as their potential wages are multiples higher. Second, because individual level productivity z changes over time, highly productive workers want to insure themselves by matching themselves to an employer. Since wages are fixed for the duration of the match, they stand to gain more from being employed at relatively low wages than from waiting, with the risk of a low productivity realization in the next period.

Due to risk aversion, all households will try to smooth their consumption. Unemployed households with little wealth, whose income is likely to increase, due to a positive job-finding probability, would like to borrow, but are prevented from doing so by the borrowing constraint. To avoid large jumps in their earnings, they search in submarkets that offer low wages and high job-finding probabilities. At the same time, rich unemployed households search in markets with lower job-finding probabilities but higher wages. Through their asset holdings, they are insured against the resulting fluctuations in earnings and the risk associated with them.

This result is similar to those obtained in models with search effort ([Baily, 1978](#); [Chetty, 2006](#)), but for slightly different reasons. In those model, search effort is costly, and the wage is constant. Then, richer individuals exert less search effort to find jobs, because the difference in utilities between unemployment and employment is too small to justify paying high search costs. In my model, search is costless and there is no search effort, but agents trade off higher job-finding probabilities with higher wages. All else equal, asset rich individuals search in submarkets that are more risky (lower job-finding probabilities), because they are insured against the earnings fluctuations caused by finding high-paying jobs.

3.4 Government budget

The government collects income taxes and all firm profits, with which it finances a stock of risk free bonds, balancing its budget each period:

$$\begin{aligned}
 B(1+r) + UI &= B' + \tau W + \Pi \\
 \text{where } W &= \int_{i \in \Omega} w_i \, d\Omega \\
 UI &= \int_{i \in \Omega} b_i I_i^U \, d\Omega
 \end{aligned} \tag{9}$$

where B represents the supply stock of bonds at the beginning of period t , B' represents newly issued bonds, W are total earnings in the economy and UI is the total outlays of the unemployment insurance system.

The government follows the fiscal rule

$$\tau - \tau^* = \phi_\tau (B' - B^*) \quad (10)$$

where τ^* and B^* are the steady state tax rate and bond issuance, respectively. The parameter $\phi_\tau > 0$ governs how elastic bond issuance is relative to tax changes. For $\phi_\tau = 0$, taxes are constant and the budget is balanced through bond issuance. As the parameter rises, taxes become more elastic.

3.5 Equilibrium

Recursive Equilibrium definition: A recursive equilibrium in this economy is, given a path for the interest rate r and the labor income tax τ , a set of household savings policy functions $g_a^E(z, x, w; \psi)$ and $g_a^U(z, x; \psi)$, a set of household wage choice functions $g_w^E(z, x, w; \psi)$ and $g_w^U(z, x; \psi)$ and submarket tightnesses $\theta(z, x, w', \psi)$, such that

- the household's asset and wage policy functions are consistent with the problems in Equation (7) and (8),
- the government budget in Equation (9) is balanced,
- the free entry condition implies submarket tightness according to Equation (14),
- asset markets clear, such that

$$B' = \int_{i \in \Omega} g_a^W(z, x, w; \psi) + g_x^U(z, x; \psi) \quad d\Omega$$

- the aggregate state of the economy ψ evolves according to $\psi' = Q(\psi)$

The equilibrium, formulated in this way, is difficult to compute, because the distribution of agents Ω is an infinite-dimensional object. However, the model permits a block recursive formulation, as discussed in [Menzio and Shi \(2010\)](#) and [Menzio and Shi \(2011\)](#), which considerably simplifies the problem by decoupling the decision rules and submarket tightnesses from the distribution. The proposition and the corresponding proof follow [Karahan and Rhee \(2019\)](#), [Herkenhoff \(2019\)](#) and [Birinci and See \(2019\)](#).

Block Recursive Equilibrium (BRE) definition: A block recursive equilibrium is an equilibrium in which, given a path for the interest rate r and the labor income tax τ , the households' policy functions and submarket tightnesses only depend on the aggregate productivity state A , but not on the distribution of agents Ω .

Proposition *If i) utility function $u(\cdot)$ is strictly increasing, strictly concave, and satisfies the Inada conditions; ii) choice sets \mathcal{W} and \mathcal{A} , and sets of exogenous productivity processes z and A are bounded; iii) matching function M exhibits constant returns to scale; and iv) All policies are restricted to depend on the aggregate state only through aggregate match productivity, then there exists a unique BRE for this economy*

Proof See Appendix C

4 Calibration

I calibrate the steady state of the model to match a set of labor market moments I observe in my dataset. For this exercise, I set aggregate productivity to $A = 1$, and hence, there are no aggregate fluctuations. I use the Generalized Method of Moments to minimize the distance between the target moments in the model and the data. Tables 1 to 2 and Figure 5 summarize the model calibration.

The period length in the model is set to one quarter. I set the annual risk-free interest rate to 4%, resulting in a quarterly rate of $r = (1.04)^{\frac{1}{4}} - 1 = 0.0985$. To calibrate the bond supply B , I target the average liquid-asset to income ratio in Germany, which I obtain from the Household Finance and Consumption Survey. I utilize the survey's second wave, which was mainly conducted in 2014. Household asset holdings are categorized into liquid and illiquid following the approach by Kaplan et al. (2014), with slight modifications to the definitions used in their exercise. As household income, I classify the sum of self-reported labor earnings and social transfers, as this comes closest to the income measure I observe in the SIAB panel.¹³ The resulting average liquid-asset to income ratio for Germany is 4.5.

The utility function for workers is given by

$$u(c) = \frac{c^{1-\sigma}}{1-\sigma}$$

where σ represents the coefficient of relative risk aversion. I set this parameter to 2, in line with standard calibrations in the literature.

Labor Market Following Den Haan et al. (2000), labor market matches in the model, from a mass of searchers S and vacancies V , are formed according to the matching function

$$M(S, V) = \frac{SV}{(S^x + V^x)^{\frac{1}{x}}}.$$

This function guarantees that the job-finding and vacancy-filling probabilities lie within

¹³To establish comparability between the HFCS and the SIAB panel, and reliability of the income variable in the HFCS, I calculate the median quarterly income in 2014 in both datasets. In the SIAB, the individual median income is €5700, while in the HFCS, quarterly household income is €4160. Note that the SIAB values are gross earnings, while the HFCS are net of tax.

the interval $[0, 1]$. I set $x = 0.3$ to match the average job-finding probability in the data. I set the vacancy cost $\kappa = 0.04$ from [Hagedorn and Manovskii \(2008\)](#), diving by 12, as theirs is a weekly model. The exogenous quarterly separation probabilities, $\delta(z)$, are set to match the separation probability in the data, across the distribution. Unemployed households produce h of home production in each quarter. In the model, this parameter is closely related to the job-finding probability of the lower quantiles; hence, I calibrate it to match the job-finding probability at the low end of the income distribution.

Government In Germany, the replacement rate is 60 % of the last net-wage. In order to save on state variables (past wages), I model this benefit by indexing unemployment benefits to productivity such that $b(z) = \phi_B z$ with $\phi_B = 0.6$, similar to [McKay and Reis \(2016\)](#). I set the parameter governing the elasticity of the tax rate $\psi_{\tau} = 0.1$.

Productivity process The household's productivity z is a combination of aggregate productivity in the economy and the consumer's individual productivity. Both evolve according to AR(1) processes. Individual productivity follows

$$\log(z_{t+1}) = \rho_z \log(z_t) + \epsilon_{t+1} \quad (11)$$

with persistence ρ_z and innovations $\epsilon \sim N(0, \sigma_\epsilon)$. In the model, I discretize the process to a grid $\{z_1, \dots, z_s\}$ with $s = 7$ using the Rouwenhorst method. To calibrate the parameter ρ_z , I estimate the following earnings process for *the continuously employed* in both the model and in the data:

$$\log(y_{i,t+1}) = \rho_y \log(y_{i,t}) + e_{t+1}. \quad (12)$$

in order to exclude labor market transitions, which will be an endogenous outcome of the model. Then, I choose the parameters ρ_z in the productivity process such that the two match. In the data, I residualize earnings in the spirit of [Heathcote et al. \(2010\)](#) by regressing labor earnings on observable characteristics:

$$\log(earn_{i,t}) = \alpha + \gamma_1 X_{i,t} + \gamma_2 T_t + y_{i,t}$$

where $earn_{i,t}$ is an individual's quarterly earnings and $X_{i,t}$ contains an age polynomial, dummies for gender and education. I remove these aspects of the earnings process, because the model does not allow for heterogeneity along these dimensions. Using the variance/covariance structure between earnings observations of periods t and $t + s$, I back-out the parameter $\rho_z = (\text{Cov}(y_t, y_{t+s}) / \sigma_\epsilon^2)^{(1/(s-t))}$. As $s \rightarrow \infty$ (around $s = 50$), the latter equation converges to $\rho_{y,data} = 0.988$ (around $s = 20$) in the data. To calibrate the volatility of the earnings process, I target the ratio of the average gross earnings between the 25th and the 75th percentile, which is 2.402 in the data. In the model, I set $\rho_z = 0.98$ $\sigma_\epsilon^2 = 0.07$.

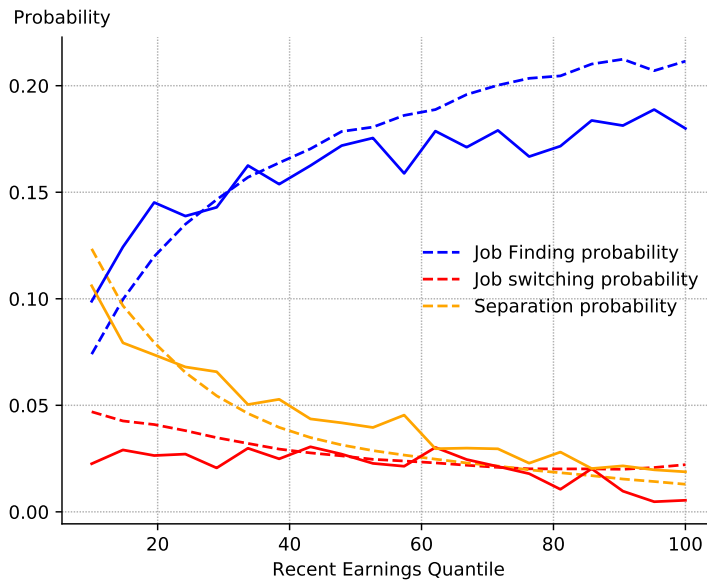
Table 1: Externally Calibrated Parameters

Parameter	Description	Value
r	Interest Rate	0.0985
σ	Coefficient of relative risk aversion	2
ϕ_B	Replacement Rate	0.60
κ	Vacancy posting cost	0.04
ϕ_τ	Persistence of tax rate	0.10

Table 2: Internally Calibrated Parameters

Parameter	Description	Value	Target Moment	Data	Model
B	Bond supply	2.552	Liq. wealth to income	2.670	2.663
λ	OTJ search Probability	0.300	Average switching prob.	0.028	0.022
x	Matching funct parameter	0.301	Average job finding	0.172	0.175
ρ_z	Persistence of prod. process	0.980	Employed earnings	0.988	0.983
σ_ϵ	Distribution of prod. innovations	0.070	Relative wage $P75/P25$	2.402	2.454
$\delta(z)$	Exogenous separation probability		see text		
h	Home production		see text		

Figure 5: Transition Probabilities - Model and Data



Note: This figure compares the transition probabilities in the data to their analog in the model. Dashed lines represent probabilities obtained in the data, solid lines are obtained from the model. The blue lines represent job-finding probability, the yellow line represents separation probability and the red line represents job-switching probability. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

The calibrated model is able to replicate the transition probabilities in the data fairly well.¹⁴ The separation probabilities in the data were targeted through the exogenous

¹⁴Appendix B shows steady state earnings growth rates in the data and the model.

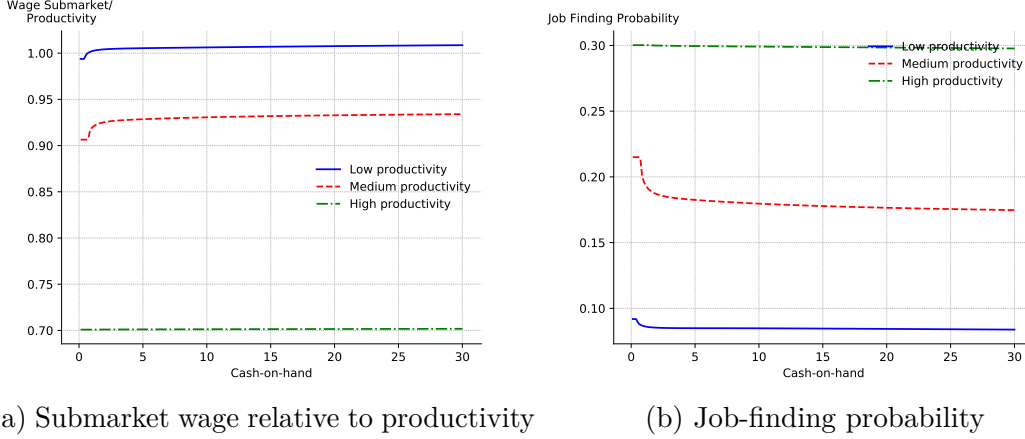
separation probabilities along the productivity grid in the model. However, job switching was not targeted, and still the model matches the observed switching probabilities in the data fairly well, especially around the median. Furthermore, the job-finding probabilities along the distribution were only targeted using two parameters: home production h and the parameter in the matching function x . Still, the model produces job-finding rates which are sloping upwards close to the data. In the next section I describe how the model achieves this.

4.1 Policy functions

The job-finding probability along the income distribution is strongly increasing in the model, as in the data. The reason for this is illustrated in the left-hand panel of Figure 6, which shows the search policy functions of the unemployed for different levels of cash-on-hand (x-axis) as well as idiosyncratic productivity (colors) *relative to their productivity realization*. The right panel shows the corresponding job-finding probabilities. Unemployed individuals with higher productivity realizations search in submarkets that offer lower wages, relative to their productivities. The right-hand panel shows that this leads to higher job-finding probabilities for high productivity workers. These workers have a strong incentive to become matched: if they find employment, their wages are fixed, even if their productivity realizations fall in the future.

At the lower end of the productivity spectrum, on the other hand, the unemployed have little incentive to find employment. Their outside option (home-production) is high, relative to their productivity. Furthermore, their productivity may rise in the future, in which case they do not want to be tied to a job, anticipating that matched individuals can only search for new jobs with probability $\Lambda < 1$. Figure 6 also shows search behavior along the wealth (cash-on-hand) dimension. Richer individuals search in submarkets that pay higher wages and offer lower job-finding probabilities, conditional on productivity. This is because individuals with relatively little wealth, while wishing to find work so they can increase their consumption, prefer small fluctuations in their earnings. As they get richer, the unemployed are more able to insure against the large earnings jumps that come from finding a job which pays a high wage. Hence the wage-policy function is upwards-sloping in wealth. Wealth effects, however, are trumped by productivity effects. Only for intermediate values of productivity do they play a role, and the wage changes associated with them are small.

Figure 6: Submarket choice of the unemployed



Note: The *Left Panel* shows the wage submarket choices for workers with high, medium and low-productivity realizations. The x-axis represents their cash-on-hand state and the y-axis represents their wage choice relative to their current productivity realization. The *Right Panel* shows the associated job-finding probabilities.

The mechanism for employed workers who search on the job is analogous. Higher productivity realizations make workers search for jobs at relatively low wages that offer high job-finding probabilities, while wealth makes individuals search for jobs that pay higher piece-rates but offer lower job-finding probabilities. Additionally, employed workers take their current wages into account, never searching in wage submarkets that offer lower pay.

5 Results

5.1 Earnings procyclicality in the model

In this section, I show that the model presented in Section 3 can reproduce the heterogeneity in earnings dynamics observed in the data. The earnings betas presented in section 2 are computed as the correlation between individual income growth and aggregate earnings growth. In the steady state of the model, aggregate earnings are constant, however. To estimate earnings betas in the model, I introduce aggregate shocks and solve for the economy's response along a series of such shocks over time.

An economy with aggregate shocks requires rational, forward looking agents to take the aggregate state of the economy into account when making their consumption, savings, and wage choices. This state can be summarized by the vector $\psi = (\Omega, A)$, where Ω is the distribution of agents in the economy and A is the level of aggregate productivity (Krusell and Smith, 1998). Once the model is solved in this general way, the researcher can simulate its response to a series of shocks over time and estimate the earnings betas

as the correlation, over time, between aggregate and quantile-specific earnings growth.

The dynamic programming problem in an economy with aggregate risk is difficult to solve, as the distribution Ω is an infinite dimensional object. In principle, the state space is infinitely large. I rely on the solution method proposed by [Boppart et al. \(2018\)](#) (BKM) to solve this problem. Their approach employs the impulse response to a single unexpected shock in aggregate productivity as a first-order approximation to the global solution of the model. In order for this method to provide accurate results, the model's policy functions, and hence aggregate outcomes, need to be approximately linear in the aggregate productivity shock. Below, I show that this is the case in my model.

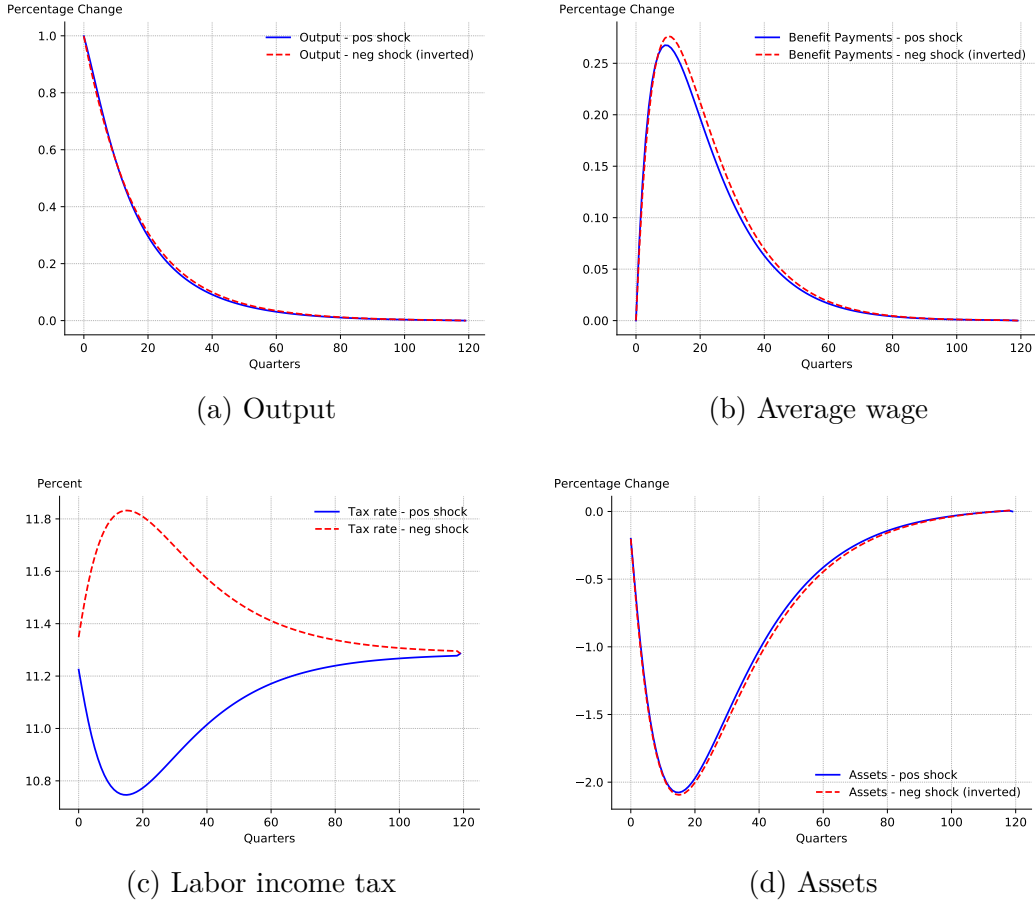
I estimate two impulse responses to aggregate productivity shocks of 1% and persistence $\rho = 0.9$. Using these impulse responses as a linear approximation of the model to any aggregate productivity shock of the same sign, I simulate the economy for 600 periods in response to a series of aggregate shocks drawn from a normal distribution with $\sigma_{agg} = 1\%$ and persistence $\rho_{agg} = 0.9$. Along the transition path, I assume that the government balances its budget according to the fiscal rule in Equation (10) by varying both the labor-income tax rate τ_t and the bond issuance B_t . Each period, the real interest rate r_t adjusts to clear the asset market. I solve for the path of the tax rate τ_t and the interest rate r_t using the sequence space Jacobian method proposed by [Auclert et al. \(2021\)](#).¹⁵

Figure 7 shows the responses of several aggregate variables to a 1% increase and decrease in aggregate match productivity. Where feasible, the responses for the negative shock were inverted to facilitate the comparison. The impulse responses to the negative productivity shocks are almost perfectly symmetrical to those after a positive shock. In response to the positive shock, output increases by close to 1 percent upon impact and then converges back to its steady state equilibrium. Wages also increase, but with a lag. This is related to the structure of the labor market. Upon impact, job finding increases as firms want to hire workers who will temporarily be more productive. Workers trade off some of the increase in the job-finding probability for higher wages. As aggregate productivity converges back to its steady state value, these incentives wane away. The aggregate converges back to steady state as the workers who were hired at higher wages separate from their matches.

The labor income tax falls in response to a positive shock for two reasons. Unemployment decreases, which leads to less government expenditures. Additionally, wages rise, implying a larger tax base. In period $t = 0$, as agents in the model are confronted with the new, decreasing path for productivity, they want to save for the worse times ahead and consume less today. To clear the bond market, the interest rate falls to dampen the demand for savings.

¹⁵For a discussion of the computation, see appendix D.

Figure 7: Impulse response to Productivity shock



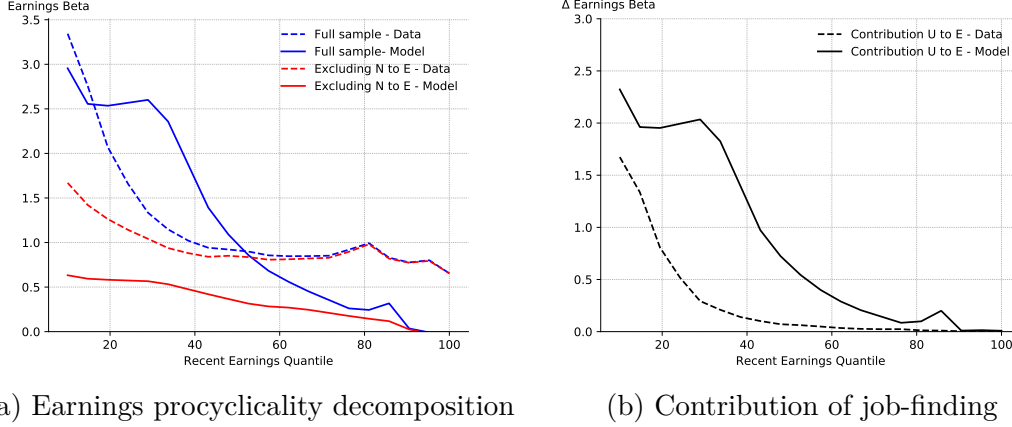
Note: The *Top Left Panel* shows the response of output, in percent, relative to steady state, to a positive (solid) and negative (dashed) productivity shock which reduces aggregate match productivity by 1%, with persistence 0.9. The response to the negative shock is inverted. The *Top Right Panel* shows the responses of benefit payments, in percent, relative to the steady state; the negative response is inverted. The *Bottom Left Panel* shows the responses of the labor income tax rate to the two shocks. The *Bottom Right Panel* shows the response of asset demand, in percent, relative to the steady state, with the response to the negative shock inverted. The horizontal axis shows quarters since the shock.

I employ the algorithm proposed by BKM to simulate an economy with aggregate shocks over time. Then, I construct the model-generated analog to the dataset discussed in section 2 and use it to estimate Equation (2), the gross-earnings betas, in the model.

The left-hand panel of Figure 8 shows the result of this exercise. The solid and dashed blue lines represent the model generated beta and the betas estimated in Section 2.3, respectively. The model generated beta coefficients imply that earnings at the bottom of the recent earnings distribution are about three times as procyclical as those around the median, which is in line with the data. At the bottom of the income distribution, the model replicates the steeply decreasing earnings betas, but then briefly flattens out between the second and fourth deciles. Moreover, the model does not replicate the flattening out of the empirical earnings betas beyond the median. Instead, the model implies estimates

which decrease towards zero. This is intuitive, as in the model, job-stayers' earnings are fixed for the duration of a match. As job-staying is the most common transition path for individuals at the top of the earnings distribution, this drives the correlation between aggregate and quantile specific earnings growth to zero.

Figure 8: Model implied earnings betas



Note: The *Left Panel* shows the results for the regression coefficients in Equation (2) in the model (solid lines) and the data (dashed lines). The blue lines represent results for the full sample, the red lines represent results for the subsample excluding job-finders. The *Right Panel* shows the difference between the regression coefficients obtained from Equation (2) in the full sample and the restricted sample. The solid line represents the model, the dashed lines represent the data. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. In the data, the sample period is 1980-2014.

Next, I restrict the sample by excluding individuals who are non-employed in quarter t and employed in quarter $t + 4$. This results in the red dashed (data) and solid (model) lines in the left-hand panel of Figure 8. The difference between the two can be interpreted as the contribution of job-finding to the procyclicality of earnings growth. The difference between the data and the model is more pronounced, with the model implied coefficients falling from 0.5 to zero, while in the data, they fall from close to 1.5 to slightly below one. The level difference, again, is due to the fact that the model does not generate any gross-earnings fluctuations for job-stayers, leading the model implied betas to be closer to zero.

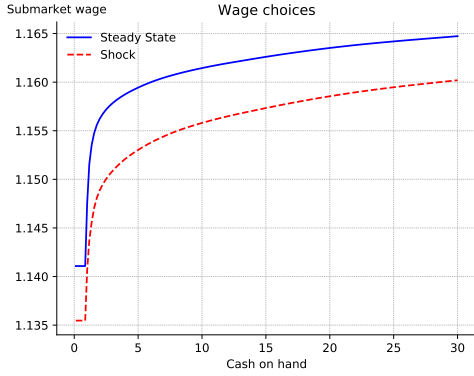
An additional comparison between the model and the data can be conducted by comparing the contribution of job-finding to the overall beta coefficients, i.e. the change in the earnings betas when moving from the full sample to a restricted one which excludes job-finding. The result is plotted in the right-hand panel of Figure 8, with the dotted line representing the size of the shaded area in Figure 3b and the solid line representing its model analog. The model, as the data, implies a decreasing contribution of job-finding to the overall beta coefficient. However, it generates a contribution that is considerably larger than that implied by the data. Still, the model qualitatively matches the empirical

patters.

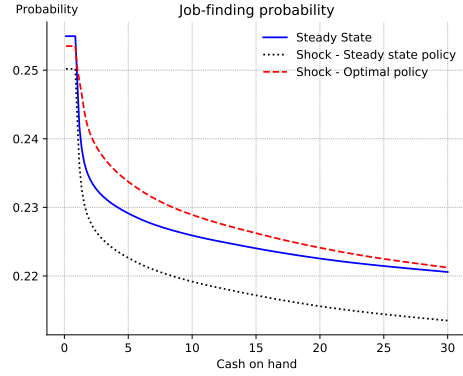
The model is able to generate the patterns observed through the following mechanism. In response to a negative productivity shock, firm value decreases, all else equal. This decrease in match-profits leads to fewer firms entering into all submarkets. Consequently, workers face lower job-finding probabilities if they follow their steady-state wage submarket choices. To counteract some of this decrease, workers can choose to search in lower-wage submarkets, which offer higher job-finding probabilities. The attractiveness of this tradeoff, however, varies across the productivity distribution (and, by extension, the recent earnings distribution). Furthermore, it is ex-ante unclear how much of the fall in vacancy creation workers can compensate for by switching to different submarkets.

Figure 9 summarizes this channel for median productivity individuals (top row) and low productivity individuals (bottom row). The top left panel shows the submarket choices of median productivity individuals in steady state (blue line) and in the first period after the realization of a one-percent decrease in aggregate match productivity. In response to the shock, these agents choose to search in lower-wage submarkets. The top right panel investigates how this affects their job-finding prospects. The blue line shows the job-finding probabilities in steady state; after the realization of the shock, the job finding probabilities drop for all wealth levels, holding submarket choices fixed (black, dotted line). Due to individuals adjusting their submarket choices, however, the job-finding probabilities shift upwards beyond their steady state values.

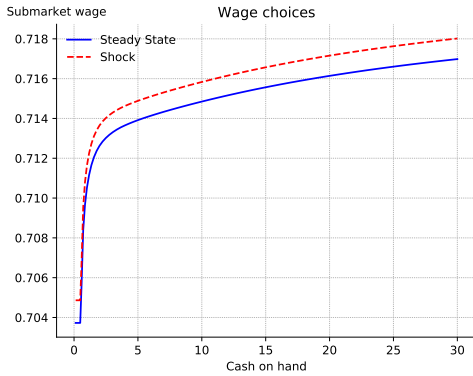
Figure 9: Model mechanism



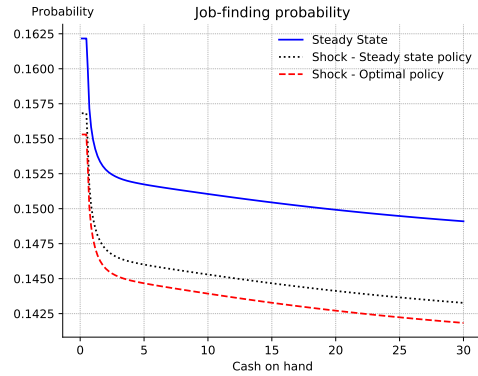
(a) Wage policy, high productivity



(b) Job-finding, high productivity



(c) Wage policy, low productivity



(d) Job-finding, low productivity

Note: The *Top Panels* show the wage choices (left) and the job-finding probabilities (right) of high-productivity workers. The solid lines represent steady state outcomes, the dashed lines represent outcomes in the first period after the realization of a negative productivity shock, the dotted line represents the job-finding probabilities after a negative productivity shock when worker choices are held fixed at their steady state values. The *Bottom Panels* contain the same information for low-productivity individuals.

Analogous to the top left panel, the bottom left panel shows the submarket choices of low productivity individuals. In response to a negative productivity shock, they choose to search in wage-submarkets that offer *higher* wages. This leads their job finding probabilities (bottom right panel) to drop even lower than the hypothetical scenario that keeps policy functions fixed to their steady state values. This is due to the fact that their income from home production (together with unemployment benefits) is large, relative to their potential wage earnings. In response to a negative productivity shock, their potential wage earnings fall even further, making non-employment even more attractive. Furthermore, as tax rates increase modestly after a negative productivity shock, low-productivity individuals ask for higher wages to be compensated for this loss, leading to even lower job-finding probabilities.

The mechanism discussed here implies that wages of new matches are procyclical at the top of the income distribution, but countercyclical at the bottom. In the data, the

cyclicalities of new wages at the low end of the distribution is not countercyclical, but still considerably less cyclical than new wages at the top.

6 Policy Experiments

Using the calibrated model outlined in section 3, I conduct two policy experiments motivated by the observation that earnings are more procyclical at the bottom of the income distribution than at the top. As described above, the heterogeneous procyclicalities are due to two factors: first, firm entry and vacancy posting are procyclical in the data (and the model) and second, non-employment is concentrated towards the lower end of the permanent income distribution. The two policy experiments discussed in this section address each of these observations in turn.

First, I introduce a countercyclical hiring subsidy for firms. This policy incentivizes vacancy creation in recessions and allows job-finding to remain closer to its steady state level over the business cycle. Intuitively, this policy weakens the risk to workers caused by a decrease in vacancies posting, which they cannot counteract by adjusting their search decisions.¹⁶

The second experiment considers a Universal Basic Income (UBI), which substitutes for and eliminates unemployment benefits. This policy has received much attention in Germany and around the world.¹⁷ In 2020, the German Institute for Economic Research started a pilot project on the subject, aiming to assess its impact on well-being and values.¹⁸ In the context of the documented heterogeneity in earnings risk, the concept of a UBI is interesting as it removes the disincentive on working introduced by unemployment insurance. Hence, it has the potential to encourage workers to seek out vacancies with higher job-finding probabilities, reducing non-employment at the lower end of the income distribution. In what follows, I introduce both policies in more detail and compare their welfare effects in the model.

I evaluate the policies according to two measures. First, I calculate the welfare effects of each policy by estimating the percentage increase in consumption necessary in order to make agents indifferent between the two economies, following Krusell et al. (2010):

$$E_0 \left[\sum_{t=0}^{\infty} \beta^t u((1 + \lambda)c_{i,t}) \right] = E_0 \left[\sum_{t=0}^{\infty} \beta^t u(\bar{c}_{i,t}) \right] \quad (13)$$

¹⁶In the appendix, I discuss an additional policy experiment, which introduces countercyclical unemployment benefits. This measure offers insurance beyond the steady state unemployment benefit payments in times of negative productivity shocks and decreases insurance when it is not as needed, i.e. in times of positive productivity shocks.

¹⁷For coverage in a major German newspaper, see, e.g.: <https://www.zeit.de/wirtschaft/2021-02/bedingungsloses-grundeinkommen-corona-arbeit-existenzangst-finanzierung>

¹⁸see <https://www.pilotprojekt-grundeinkommen.de/english>

where $c_{i,t}$ is each agent's consumption in the baseline economy and $\bar{c}_{i,t}$ is their consumption in the counterfactual economy. Unfortunately, I cannot calculate this measure in the full aggregate risk version of the model, as I solve it using first-order approximation. For welfare statements, however, a second-order approximation would be needed. Instead, I evaluate Equation (13) in period 0 after an unexpected, one-time productivity shock. The value of λ then indicates how much consumption each agent would be willing to give up in order to switch to the counterfactual economy, as opposed to the baseline economy. Averaging the values of λ over all agents in the model is referred to as the utilitarian consumption equivalent variation of the counterfactual economy.

The second measure, which I can compute along the model's transition path including aggregate shocks, is the volatility of aggregate consumption. A higher cyclical volatility points to more risk in the model, which is linked to lower welfare.

6.1 Countercyclical hiring subsidies

As outlined above, the goal of the countercyclical hiring subsidy is to dampen the cyclical volatility of firm vacancy creation. This, in turn, has the potential to mitigate the countercyclical job-finding risk from the workers' perspectives.

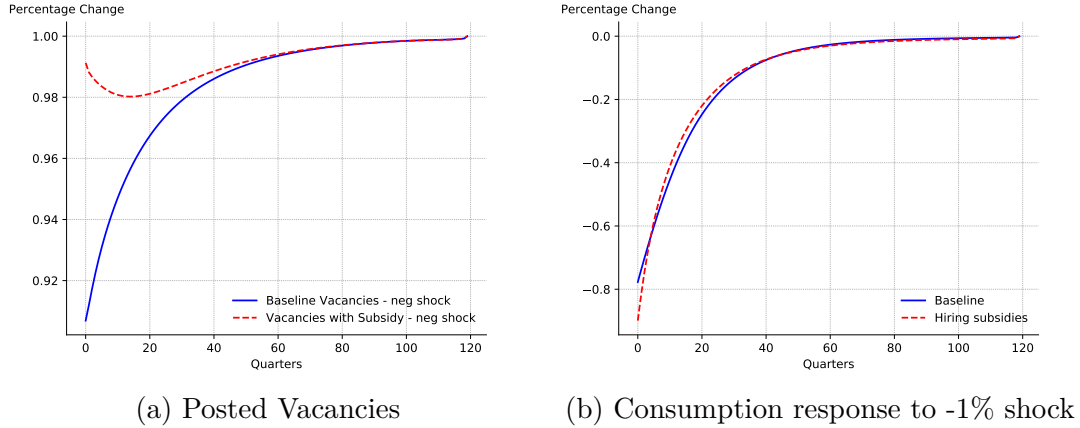
I implement this policy by letting the government award (or tax) firms the lump-sum transfer $v(A)$, indexed by aggregate productivity A , with the following function form:

$$v(A) = \vartheta(A - \bar{A})$$

where the multiplier ϑ is calibrated such that, in response to a negative productivity shock of 1 percent, hiring firms receive a one time transfer which equals the hiring cost: $v(0.99) = \kappa$. As aggregate match productivity returns to its steady state value, the hiring subsidy decreases back to zero. The process is analogous upon the arrival of positive productivity shocks.

Figure 10 shows the results of this exercise. The left-hand panel displays the percentage change in the number of vacancies created in a low-productivity submarket. In the baseline economy, vacancy creation contracts by more than 8 percent. With the subsidy, this fall is reduced considerably: vacancy creation falls by at most 2 percent, and the impulse response is more hump-shaped. Due to the small size of the hiring subsidy, the effect on vacancy creation in high productivity submarkets is small.

Figure 10: Effect of hiring subsidies



Note: The *Left Panel* shows the response of vacancy posting from firms entering low-productivity submarkets to a negative aggregate productivity shock of 1 percent in the baseline economy (solid) and the economy with hiring subsidies (dashed). The *Right Panel* shows the response of aggregate consumption to a negative productivity shock in the two economies.

The right-hand panel of Figure 10 shows the response of aggregate consumption. The initial impact of a negative productivity shock is slightly larger than in the baseline economy. Afterwards, the economy with hiring subsidies rebounds somewhat faster, however.

Welfare, as measured by Equation 13, immediately after the realization of a negative, unexpected, transitory shock to aggregate match productivity, increases. To reach the same level of aggregate welfare in the baseline economy, each worker would need to be given an increase of 0.02% in their consumption, forever. To put this number into perspective, I calculate the consumption equivalent variation between the steady state economy and the first period after a negative productivity shock in the same economy. If given a choice, workers would, on average, give up 0.19% of their consumption each period in order to avoid negative productivity shocks. A hiring subsidy attenuates this impact by about ten percent.

The largest welfare gains accrue around the median of the recent earnings distribution. While positive, the effect is small at the bottom, because the hiring subsidy does not raise wages enough to compensate workers for the loss of home production and unemployment benefits. At the top of the distribution, hiring subsidies decrease welfare in recessions. For high productivity individuals, there is little marginal benefit in raising job-finding probabilities. However, they are negatively affected by the higher tax-burden, which lowers their welfare.

6.2 Universal Basic Income

The second policy experiment introduces a UBI into the economy while eliminating unemployment benefits. This experiment does not only affect the business cycle behavior of the economy, but also its steady state. To make meaningful comparisons between the baseline economy and the UBI economy, I calibrate the unconditional payment such that the income tax rate τ remains constant between the two, implying the same fiscal cost to the government. Without this constraint, some welfare changes might be driven by a higher tax rate leading to more redistribution by construction. The resulting UBI payment is $ubi = 0.147$, which equates to about 450 euros per quarter for every inhabitant of the economy.¹⁹

I keep the asset supply in the two economies the same, implying that the risk free interest rate in the UBI economy must adjust to clear the market. In equilibrium, the quarterly interest rate falls to -0.0009 , which implies that agents demand considerably more safe assets than in the baseline economy. This makes intuitive sense, as the removal of unemployment benefits implies that there is less government provided insurance for workers who become non-employed. In the baseline economy, highly productive workers receive 60% of their productivity value when separating from their firm, which implies a monthly income of up to 4000 euros. With the introduction of UBI, this value drops by almost 90%. Incomes in unemployment also decrease for low-productivity workers. While the drop is much smaller, from around 800 euros, the effect is large, as the separation probabilities are much higher and unemployment is higher.

The effect on the labor market is shown in the top-left panel of Figure 11. The solid lines represent the transition probabilities in the UBI economy and the dashed lines reproduce the same probabilities in the baseline economy. Separation probabilities are almost equivalent between the two economies, as they are exogenous and explicitly targeted to the data. The probability of job-finding rises by about 10 percentage points across the distribution. As explained above, the prime reason for this shift is the unattractiveness of non-employment without unemployment benefits. Workers choose to search for jobs in submarkets that offer lower wages, but higher job-finding probabilities, as their reservation utility is only influenced by their home-production value.

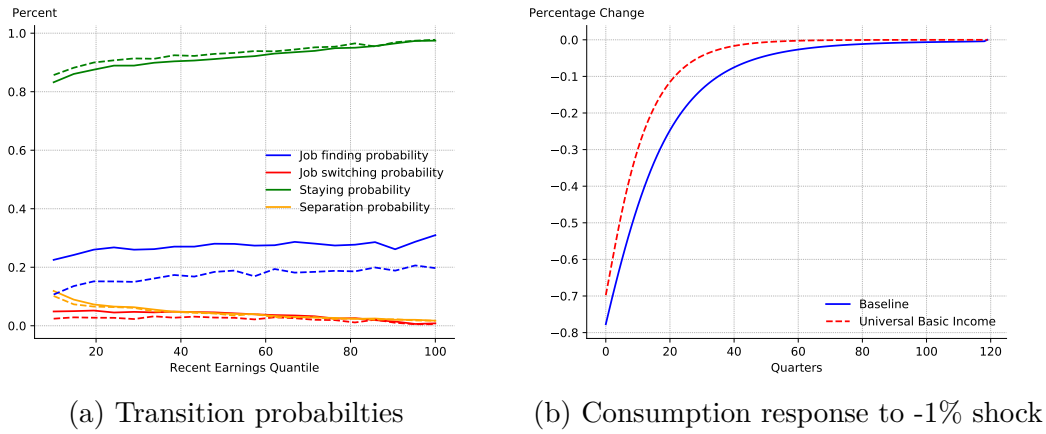
In the aggregate, higher job-finding across the distribution leads to higher output and higher aggregate consumption. Relative to the baseline economy, output rises by five percent in the UBI economy, while consumption rises by close to one percent. The difference is due to the fact that the UBI economy features more firm-entry, leading to more resources spent on vacancy posting. Aggregate, ex-post profits in the new economy are higher by a factor of two, compared to the baseline economy, since workers, relative to

¹⁹ Average quarterly pre-tax earnings in the economy are 0.8438, while average real pre-tax earnings in the data equal 7816 euros per quarter. Hence a UBI benefit of 0.147 in the model implies a payment of 448 euros.

their productivities, search in lower wage submarkets. This leads to higher profit shares for firms.

The top right panel of Figure 11 shows the consumption response of the economy to a negative 1% shock to aggregate productivity. The percentage change in aggregate consumption is smaller in the UBI economy and the series converges back to steady state considerably faster. The reason for this is the stronger incentive to find work, even in times of low aggregate productivity. In the baseline economy, only highly productive workers adjusted their wage demands such that their job-finding probabilities stayed close to their steady state values. In the UBI economy, when aggregate productivity falls, all workers search in lower wage submarkets and thus offset the initial fall in job-finding probabilities.

Figure 11: Effect of Universal Basic Income



Note: The *Left Panel* shows the steady state transition probabilities in the model with UBI (solid) and the baseline economy (dashed). The *Right Panel* shows the response of aggregate consumption to a negative productivity shock of 1 percent *relative to steady state* in the baseline economy (solid) and the economy with a UBI (dashed).

As in the previous policy experiment, I calculate the welfare impact of the introduction of a UBI into the economy. While previously, the steady state was unaffected, and therefore no welfare comparison was warranted, here, I first estimate the utilitarian consumption equivalence variation of introducing a UBI into the baseline economy, following Equation (13). I find that $\lambda = 0.0598$, implying that workers, on average, in the baseline economy are willing to give up close to 6% of their consumption each period in order to avoid living in the economy that provides them with a UBI. This considerable reduction in aggregate welfare is mainly driven by the reduction of targeted insurance already mentioned above. Low productivity workers are willing to give up close to 16% of their consumption each period in order to stay in the baseline economy. They are most affected by the exogenously higher separation probabilities and hence stay non-employed much longer than workers at high productivities.

In order to assess the welfare impact of business cycles, I conduct a “difference in differences” analysis between the baseline and the UBI economy. I expose both to a negative 1% shock to aggregate match productivity and calculate the average consumption equivalence loss in both economies. Then I compare these values across economies. In the baseline economy, workers are willing to give up on average 0.2% of their consumption in each future period to avoid a single negative productivity shock. In the UBI economy, this number is only 0.1%. From this I conclude that productivity shocks and, by extension, business cycles, are less costly in the UBI economy *relative to the steady state*. This smaller exposure, however, comes at a high cost, as explained above.

7 Conclusion

In this paper I show that earnings growth is more procyclical at the low end of the income distribution due to extensive margin transitions over the business cycle. Using administrative microdata from Germany, I find that as aggregate earnings growth rises by 1 percentage point, earnings growth in the first decile of the income distribution rises by three times as much. At the top of the distribution, the comovement is less than one-for-one. Decomposing the effect, I show that at the bottom of the distribution, transitions from unemployment to employment account for half of the procyclicality, whereas in the last decile, they account for almost nothing.

In order to speak to the welfare consequences of these findings, I build a heterogeneous agent macroeconomic model with idiosyncratic and aggregate risk. The model features a frictional labor market in which workers direct their search at different submarkets, based on their characteristics. They endogenously trade off higher job-finding probabilities with higher wages. Highly productive workers find work faster by choosing submarkets which offer low wages, relative to productivity. The opposite is true for less productive workers. In this way, the model reproduces heterogeneous job-finding probabilities along the income distribution. I calibrate the model to match key features of the German labor market, such as heterogeneous labor market transition rates. Importantly, I leave earnings growth rates untargeted.

The model produces heterogeneity in the procyclicality of earnings growth similar to the data. The reason for these results are the lower job-finding probabilities for low productivity workers in steady state, combined with procyclical vacancy posting by firms. The slope of procyclicality along the income distribution produced by the model is steeper than what I find in the data, which leads to an underprediction of procyclicality at the top of the distribution.

I introduce two labor market policies into the model. Countercyclical hiring subsidies, which rise in recessions and fall in booms, reduce the cyclicity of vacancy posting by firms. This reduces the impact of aggregate productivity fluctuations on the unemployed. After

the realization of a negative productivity shock, workers in the baseline economy would require an increase of 0.02 percent in their consumption in order for aggregate welfare to be the same as in an economy with such hiring subsidies, implying an increase in overall welfare. To assess the magnitude of this increase, I calculate the welfare impact of a single productivity shock in the baseline economy, asking workers how much consumption they would be willing to forego each period to avoid the shock. According to this measure, a negative productivity shock reduces welfare by 0.2 percent, implying that hiring subsidies lessen the welfare impact of recessions by 10 percent.

The second labor market policy I introduce is a UBI. This policy has received considerable attention in the media in recent years. In the context of my framework, it serves as a substitute to unemployment benefits, which disincentivize job-finding at the bottom of the income distribution. I calibrate the unconditional payment to be financed with the same fiscal capacity as the unemployment benefit regime in the baseline economy, which leads to a value of close to 450 euros. The introduction of a UBI decreases the volatility of aggregate output and consumption over the business cycle. Negative productivity shocks reduce welfare relative to steady state, but by a smaller amount compared to the baseline economy: workers are willing to give up, on average, 0.1 percent of their consumption each period to avoid the shock, in the UBI economy, which is a reduction of 50 percent. However, despite the reduction in economic volatility in response to shocks, aggregate welfare is lower. Comparing steady states workers in the baseline economy are willing to give up 6 percent of their consumption in all future periods, on average, in order to avoid moving to the UBI economy.

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A Robustness

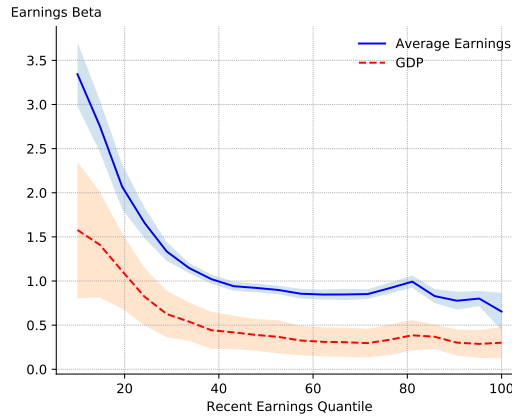
A.1 GDP

For robustness, I construct a second aggregate earnings growth measure using GDP. Growth is measured as the quarter-on-quarter log difference in real GDP between periods t and $t + k$ ²⁰:

$$\Delta Y_{t,k}^{GDP} = \log(GDP_{t+k}) - \log(GDP_t)$$

Figure 12: Earnings beta with GDP

I then substitute $\Delta Y_{t,k}^{GDP}$ into equation (2) and estimate the beta coefficients. Figure 12 reports the results from this exercise. While the beta-coefficients are shifted downwards, implying a lower cyclicity of quantile-specific earnings relative to GDP, the shape of the curve remains similar. Earnings at the bottom of the income distribution are still about three-times as procyclical as those at the top.



Note: This figure shows something

A.2 Asymmetry

The heterogeneity in the cyclicity of earnings growth outlined in section 2.3 raises the question whether the effect is symmetric between expansions and contractions of the business cycle. To investigate this issue, I obtain an OECD based recession indicator²¹, which identifies the time-period between the peak and trough values of recessions. Using

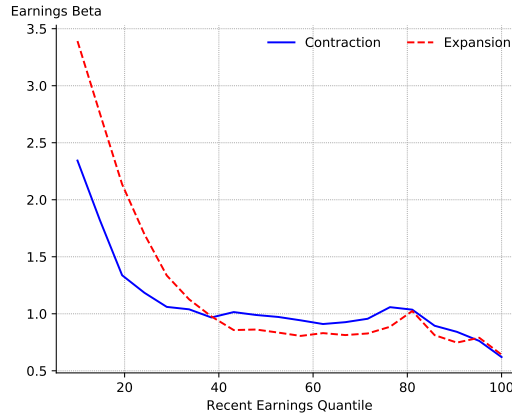
²⁰Deflated, seasonally and calendar adjusted GDP is obtained from the German Statistical Office. Due to German reunification in 1990, there only exist two separate time series for (i) West Germany and (ii) today's Germany, but both contain values for 1991. Consequently, I normalize both series by the GDP values for the first quarter of 1991 and append them.

²¹The indicator is constructed by the Federal Reserve Bank of Saint Louis, the variable name is DEUREC.

this indicator, I reestimate Equation (2), separating the sample into expansions and contractions. Figure 13 presents the results for β_Y^q along the recent earnings distribution, employed average earnings as the measure of aggregate earnings.

Similar to the results in Figure 3, earnings are more procyclical at the bottom of the earnings distribution. The procyclicality is fairly symmetrical, between expansions and contractions, with betas estimated the former slightly larger than those of the latter. One possible reason for this difference are employment protection schemes like short-time work, by which employers are subsidised in order to keep employees on their payroll. The asymmetry decreases as one moves up along the distribution, almost vanishing beyond the seventh quantile. These results indicate that the cyclicity in earnings observed especially at the bottom of the distribution is not driven solely by either expansions or contractions. Conversely, earnings at the bottom are highly correlated with the aggregate during both periods.

Figure 13: Cyclicity of earnings – Asymmetry

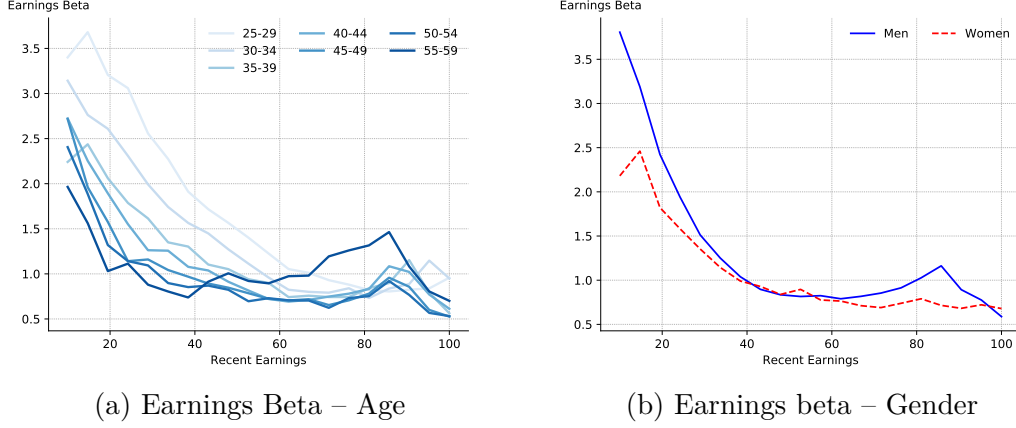


Note: This figure shows the coefficient $\beta_{Y,k}^q$ from Equation (2) by quantile, estimated during periods of contraction (blue) and expansion (red), respectively. Contractionary periods are identified based on the OECD’s recession indicator. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

A.3 Age and Gender

In the graphs presented so far, all quantiles are constructed within age-gender bins. However, it is of course of interest how the earnings betas differ along those dimensions, as well. Figure 14, therefore, shows the estimates of β_Y^q for quantiles based on recent earnings separately for each age group (left panel) and each gender (right panel).

Figure 14: Cyclicalities of earnings – Age and Gender



Note: The *Left Panel* shows the coefficient $\beta_{Y,k}^q$ from Equation (2) by quantile, estimated for subsamples of different five-year age groups. The *Right Panel* plots the coefficient $\beta_{Y,k}^q$ from Equation (2) by quantile, for men (blue) and women (red). Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

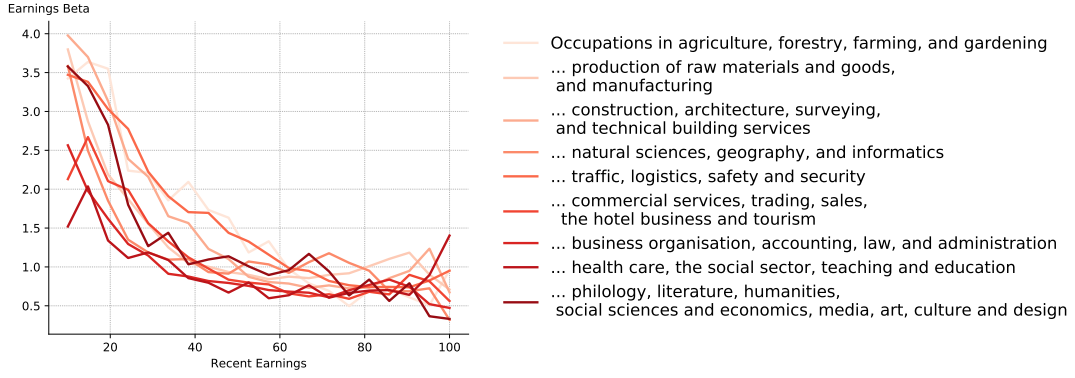
Similar to results presented in [Guvenen et al. \(2017\)](#), the size of the β_Y^q , as well as their heterogeneity, diminish with age (with the exception of the top earners of the last age group, where growth rates are imputed). While earnings growth of individuals between the ages of 25 and 34, at the bottom of the earnings distribution, moves with aggregate earnings by a factor of 3, this coefficient diminishes to 2 along the life-cycle. Towards to top of the distribution, procyclicality is weaker at all ages.

Earnings growth at the bottom is considerably more procyclical for men than it is for women as their β_Y^q is almost 50% larger. However, towards the median of the recent earnings distribution, the two graphs converge. At the very top, the two groups diverge again, potentially due to the fact that considerably more male earnings observations are censored and hence imputed. Hence, the stark difference is to be taken with a grain of salt.

A.4 Occupation and Industry

Another potential driver of the heterogeneous cyclicalities in earnings growth rates are an individuals occupation or industry. Hence, again, I reestimate Equation (2), splitting the sample into ten separate occupation bins. Figure 15 presents the results. The familiar picture of a higher correlation between individual and aggregate earnings growth at the bottom of the distribution is present in all occupations, but to varying degrees. Among the most cyclical occupations related to construction and manufacturing, moving 3:1 with average earnings in the first quantile. In the same quantile, occupations related to health care and science only co-move slightly 2:1 with the aggregate. Towards the median quantile, all graphs converge towards a value between 0.5 and 1.

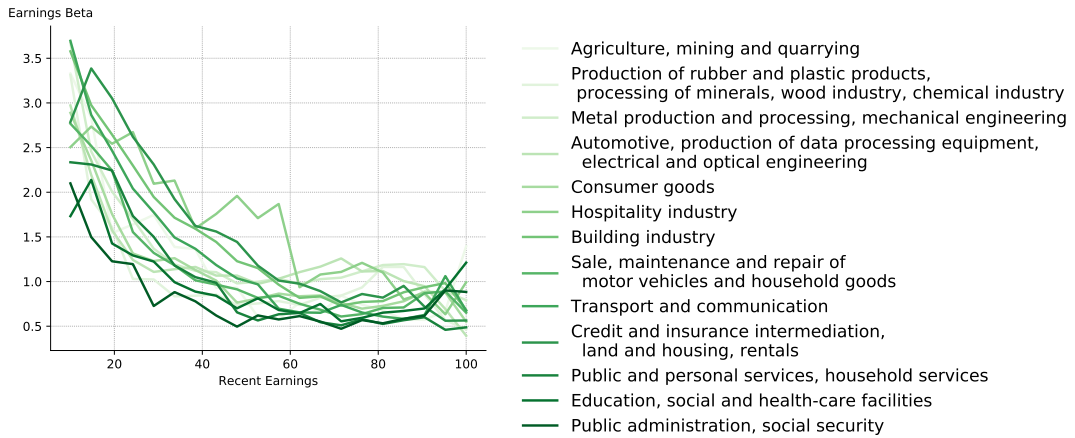
Figure 15: Cyclicity of earnings – Occupation



Note: The figure shows the coefficient $\beta_{Y,k}^q$ from Equation (2) by quantile, estimated for subsamples of different occupational groups. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

Figure 16 shows the results of performing a similar analysis separately by industry. For all industries, there is a strong downward trend in procyclicality as one moves up across quantiles. The least cyclical industry, for most quantiles, is public administration. Among the most cyclical are the automotive industry and agriculture.

Figure 16: Cyclicity of earnings – Industry



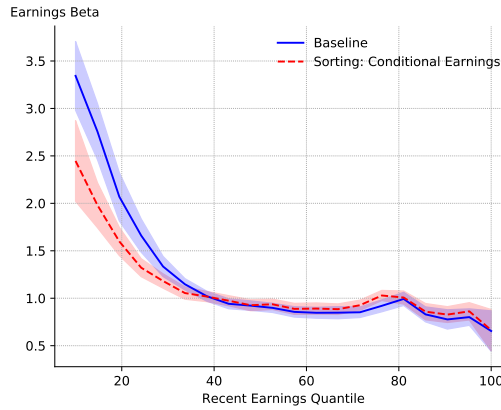
Note: The figure shows the coefficient $\beta_{Y,k}^q$ from Equation (2) by quantile, estimated for subsamples of different industry groups. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The sample period is 1980-2014.

A.5 Quantile ordering with Conditional earnings

In the baseline estimation of the earnings betas, I sort individuals into quantiles based on their earnings history over the previous five years, including zero-earnings spells, i.e., unconditional earnings. Alternatively, individuals can be sorted according to their conditional earnings history, i.e., excluding zeros. This measure is likely closer to individual level productivity, holding constant age and gender.

Figure 17 shows the result of this alternative sorting method on the earnings beta estimations (blue) and the baseline estimation discussed above (red). Sorting individuals based on conditional earnings considerably lowers the earnings beta for the bottom quantiles, but does not alter the general shape of the graph, which still indicates that those at the lower end of the (now conditional) earnings distribution receive earnings which are much more correlated with the aggregate business cycle than those further up in the same distribution.

Figure 17: Alternative Sorting – Conditional Earnings



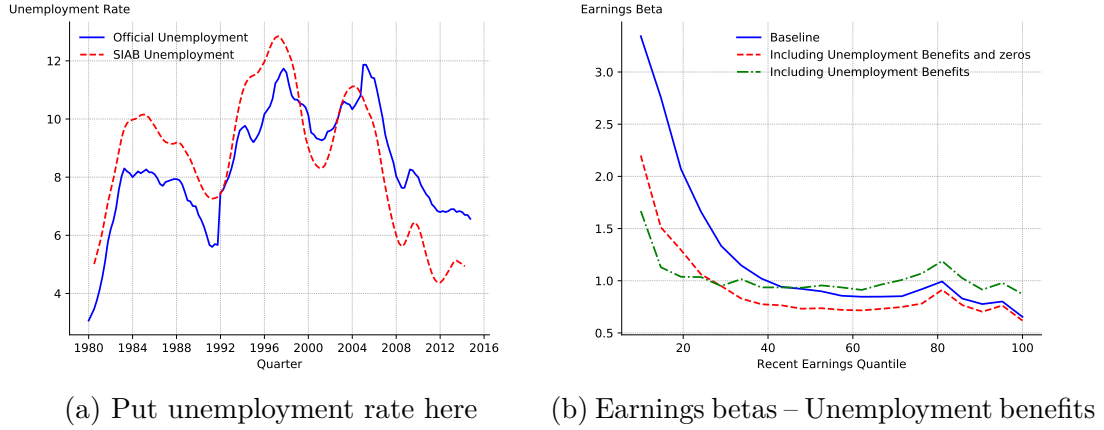
Note: This figure shows the coefficients $\beta_{Y,k}^q$, from equation (2), by quantile. Individuals are sorted into 20 quantiles each quarter based on their most recent five-year earnings history. The blue line plots the coefficient when zero-earnings observations are included in this earnings history, the red line conditions on employment and excludes zeros from the earnings history. The sample period is 1980-2014.

A.6 Unemployment benefits

As outlined in footnote 3 the only consistent measure of unemployment benefits in the data are Unemployment Benefits I. They are paid out for a limited period of time after job-loss with a replacement rate of 65%. The left panel of Figure 18 compares the official German Unemployment rate²² (red) to the one obtained using the SIAB sample (blue). To make the series comparable, I seasonally adjust the SIAB unemployment rate. From 1980 to 1998, the SIAB sample overestimates the unemployment rate by close to two percentage points. After 2000, the pattern reverses and it underestimates unemployment, especially towards the end of the sample. However, the dynamics of the two unemployment rates are very similar over time. Note that the duration of unemployment benefit receipts was shortened to 12 months in 2005, potentially leading to fewer observed unemployed individuals after that period.

²²I obtain the Quarterly Registered Seasonally Adjusted Unemployment Rate for Germany from Fred (LMUNRRTTDEQ156S) .

Figure 18: Cyclicalitity of earnings – Unemployment benefits



Note: The *Left Panel* shows the official German unemployment rate (blue) and the unemployment rate implied by the SIAB sample (red). The official unemployment rate is constructed for West-Germany during the years before 1990. The *Right Panel* shows the coefficients $\beta_{Y,k}^q$, from equation (2), by quantile. The blue line represents the baseline model, which sets earnings to zero during non-employment spells. The red line shows the coefficients when unemployment benefits are included as earnings for the non-employed. The green line restricts the sample of the non-employed only to those individuals who receive unemployment benefits.

Here, I investigate the consequences of including non-employed earnings in the baseline estimation by reestimating Equation (2) while including the observed benefits as earnings for the unemployed. I perform two robustness tests: (i) keeping earnings at zero for all non-employed individuals who are not unemployed (ii) excluding all non-employed individuals who are not unemployed. The former approach uses the same sample as the baseline, only changing earnings for some non-employed, while the second approach produces a smaller sample. Both approaches require new quantiles to be calculated.

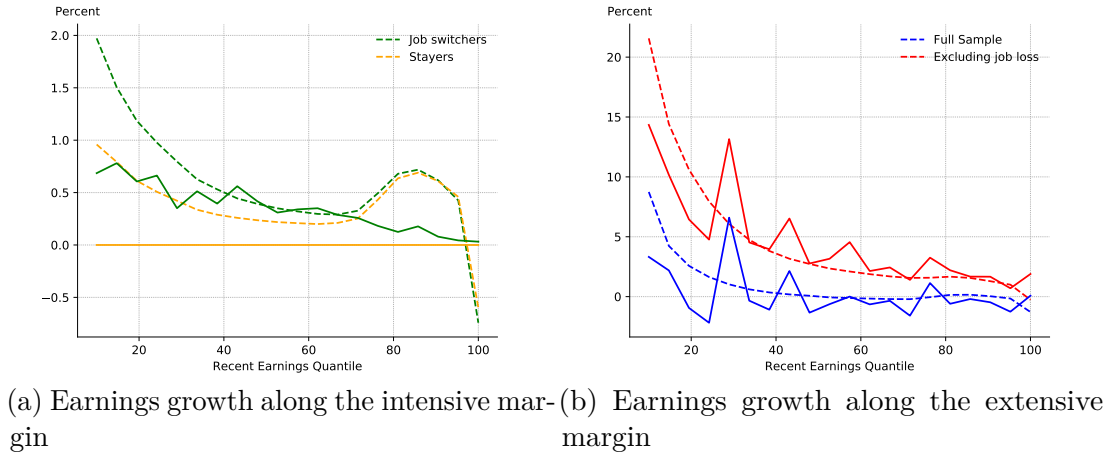
The left panel of Figure 18 compares the baseline estimation, previously reported in Figure 3 (blue) to the two alternative approaches, including zero earnings for the non-unemployed non-employed (yellow) and excluding the non-unemployed non-employed (red). The Figure shows that the earnings betas at the bottom get progressively smaller with each step. When moving from the baseline to including non-employment earnings, this is intuitive, as unemployment benefits introduce an insurance mechanism that smooths the strongest earnings fluctuations. Not also that there is almost no difference between the two lines at the very top, implying that unemployment transitions are less important here.

Excluding zeros for the non-unemployed non-employed leads the betas at the bottom to fall further. This is likely due to a reshuffling of individuals across percentiles, dampening the effects seen in the other two graphs. Beyond the 6th quantile, the quantile specific earnings betas are slightly higher than using the other two approaches.

B Additional calibration results

As described in the section 4, the calibration explicitly targets labor market transition probabilities. However, it leaves earnings growth rates across the distribution untargeted. Figure 19 shows the earnings growth rates that result from the model. The left panel compares the results along the intensive margin. In the model, job-stayers earnings are flat, by definition, as wages are fixed for the duration of a match. Job switchers' earnings growth, however, decreases in both the model and the data. The model underestimates the slope, especially at the bottom of the distribution. Furthermore, it cannot produce the increase in switchers' earnings growth at the very top of the distribution.

Figure 19: Steady State earnings growth rates



Note: The *Left Panel* shows a comparison between quarterly growth-rates implied by the model and the same statistic measured in the data. Dashed lines represent the data, dashed lines represent the model's output. The yellow lines plot earnings growth for job-stayers, the yellow line for job-switchers. The *Right Panel* shows earnings growth rates along the extensive margin. The blue line represents earnings growth for the full all workers, the red line excludes job-loss. The dashed lines represent the data, the solid lines represent the model's output.

The right panel compares earnings growth along the extensive margin in the model and the data. Both produce downwards sloping earnings growth rates. The model's output exhibits pronounced humps, however. This is due to the fact that agents are not smoothly distributed over different levels of recent earnings. A specific level of recent earnings, say R , is reached by agents who make very different decisions from those who reach level $R + 1$. This can lead to the non-smoothness of earnings growth rates along the distribution.

B.1 Countercyclical unemployment benefits

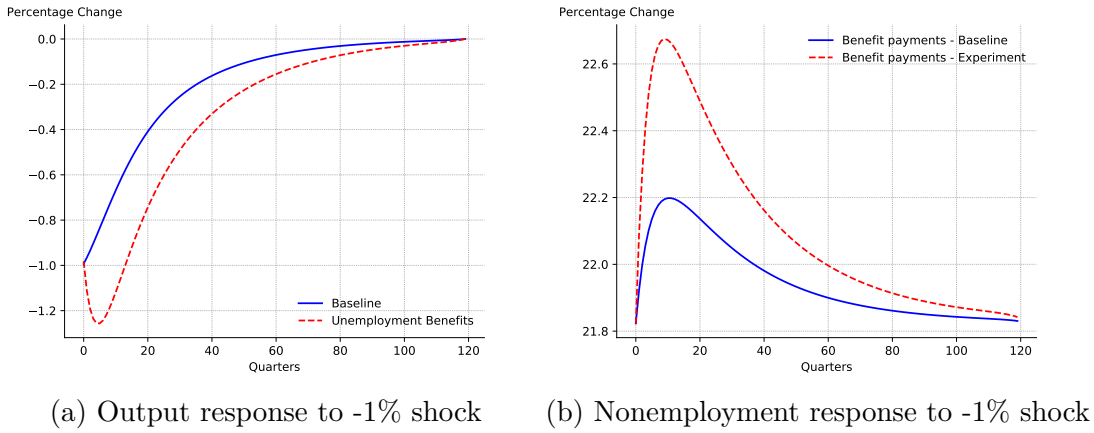
Next, I introduce countercyclical unemployment benefits into the economy. In response to a negative productivity shock, unemployment benefits are supplemented by a lump sum payment of

$$b(A) = \omega(A - \bar{A})$$

where A is the current level of aggregate match productivity and \bar{A} is its steady state value. I calibrate the size of the benefits to have similar fiscal impact as the hiring subsidies, which implies $\omega = 1.5$.

Figure 20 shows the effect of introducing this policy into the baseline economy. The left panel shows that the output response to a negative productivity shock becomes hump-shaped. While in the baseline, output immediately converges back to its steady state level, output in the counterfactual economy drops by 1.2 percent before converging back to steady state. The reason for this is shown in the right panel of the same figure. Non-employment rises by much more when unemployment benefits are countercyclical.

Figure 20: Effect of countercyclical unemployment benefits



Note: This Figure shows something

The reason for these findings are related to the incentive effects of unemployment benefits in the model. Higher benefits lead individuals to search in submarkets which offer lower job-finding probabilities, decreasing job-finding. In the context of this model, this is the moral hazard effect of unemployment insurance provision.

The welfare effect of providing unemployment benefits, according to equation 13 is positive. After the realization of a negative productivity shock, individuals in the baseline economy are willing to give up 0.004% of their consumption every quarter in order to live in an economy with the unemployment benefit supplement described here. Hence, relative to the positive insurance effect of higher benefits, However, the variance of logarithmic

aggregate consumption increases almost four-fold, relative to the baseline economy.

To investigate the moral hazard effect further, I introduce a lump sum increase of unemployment benefits into the baseline economy, but do not allow workers to reoptimize their submarket choices. This turns off the moral hazard effect of unemployment benefit provision. In spirit this exercise is similar to [Baily \(1978\)](#) and [Chetty \(2006\)](#), calculating a statistic akin to the partial derivative of benefit provision, while holding search constant. In the model, $\lambda_{nochoice} = 1.3\%$, implying a welfare increase. However, allowing workers to adjust their choices leads to $\lambda_{GE} = -1\%$.²³ I conclude that in the model presented here, the moral hazard effect of unemployment provision strongly dominates the insurance effect it provides.

²³The effect sizes are larger than those from the cyclical benefit experiment, because here, I increase benefits forever, not just in response to a productivity shock.

C Block recursive equilibrium

Block recursive equilibrium (BRE) definition: A block recursive equilibrium is an equilibrium in which, given a path for the interest rate r and the labor income tax τ , the households' policy functions and submarket tightnesses only depend on the aggregate productivity state A , but not on the distribution of agents Ω .

Proposition *If i) utility function $u(\cdot)$ is strictly increasing, strictly concave, and satisfies the Inada conditions; ii) choice sets \mathcal{W} and \mathcal{A} , and sets of exogenous productivity processes z and A are bounded; iii) matching function M exhibits constant returns to scale; and iv) All policies are restricted to depend on the aggregate state only through aggregate match productivity, then there exists a unique BRE for this economy*

Proof As mentioned in the body of the paper, the proposition and the proof closely follow Karahan and Rhee (2019), Herkenhoff (2019) and Birinci and See (2019).

I prove the existence of a block recursive equilibrium in two steps. The first step is showing that the firms' value functions and the resulting market tightnesses only depend on the aggregate state ψ through aggregate productivity A , **given a tax rate τ and an interest rate r** . The second step shows that the households' policy and value functions are similarly independent of the distribution of agents γ , **given τ and r** . Consequently, there is a solution to the households' problem which, together with the solution to the firms' problems and the resulting market tightnesses constitutes a BRE, as long as τ and r are given. Note that I condition on the tax rate τ and the interest rate r , two objects which, in order to clear the government budget and the asset market, will depend on the aggregate distribution of workers across states.

Let $\mathcal{J}(\mathcal{Z}, \mathcal{X}, \mathcal{W}, \mathcal{A} | \tau, r)$ be the set of continuous and bounded functions which map $\mathcal{J} : \mathcal{W} \times \mathcal{X} \times \mathcal{W} \times \mathcal{A} \rightarrow \mathbb{R}$, given a tax rate τ and an interest rate r . Further, let $\mathbf{T}_{\mathcal{J}}$ be the operator associated with the firm's value function, Equation (5). One can verify, using Blackwell's sufficiency conditions, that $\mathbf{T}_{\mathcal{J}} : \mathcal{J} \rightarrow \mathcal{J}$ is a contraction, the unique fixed point of which I denote as $J^* \in \mathcal{J}$. From this, it follows that the firm's value function only depends on the aggregate state ψ through the state of aggregate productivity A . In turn, this implies that the wage posting choices by firms, conditional on worker wealth \mathcal{X} and worker productivity \mathcal{Z} , are also only affected by the aggregate state ψ through aggregate productivity A . Upon substituting J^* into Equation 14, I obtain

$$\theta^*(z, x, w'; A) = \begin{cases} q^{-1} \left(\frac{\kappa(1+r)}{\mathbb{E}[J^*(z', x', w'; A')] \right) & \text{if } w \in \mathcal{W}(z, x, w; A) \\ 0 & \text{otherwise} \end{cases} \quad (14)$$

This condition shows that market tightness does not depend on the distribution of agents across states, Γ , apart from the interest rate r , which I am conditioning on.

Next, I move the the workers' problems. I combine all value functions into a single functional equation and show that this equation is a contraction. It can be shown that this

function maps the set of functions which depend on the aggregate state ψ only through aggregate productivity A . This function V is of the form $V : \{U, E\} \times \mathcal{X} \times \mathcal{Z} \times \mathcal{W} \times \Omega \rightarrow \mathbb{V}$, where Ω defines all possible realizations of the aggregate state. In this formulation,

$$\begin{aligned} V(U, z, x; \psi) &= V^U(z, x; \psi | \tau, r) \\ V(E, z, x, w; \psi) &= V^E(z, x, w; \psi | \tau, r) \end{aligned}$$

I now define a set of functions $\mathcal{V} : \{U, E\} \times \mathcal{X} \times \mathcal{Z} \times \mathcal{W} \times \mathcal{A} \rightarrow \mathbb{V}$ and let \mathbf{T}_V be an operator such that

$$\begin{aligned} (T_V V) = & \mathbb{I}_U \left\{ \max_{c_U} u(c_U) + \beta \left[\max_{w'} \mathbb{E} [\eta(\theta(z, x, w'; A | \tau, r)) V(E, z', x'_n, w'; A' | \tau, r) \right. \right. \\ & \left. \left. + (1 - \eta(\theta(z, x, w'; A | \tau, r))) V(U, z', x'_u; A' | \tau, r) \right] \right\} \\ & \mathbb{I}_E \left\{ \max_{a', c} u(c) + \beta \max_{w'} \mathbb{E} [\right. \\ & (1 - \delta(z')) (1 - \Lambda \eta(\theta(z, x, w'; A | \tau, r))) V(E, z', x'_e, w; A' | \tau, r) \\ & + (1 - \delta(z')) \Lambda \eta(\theta(z, x, w'; A | \tau, r)) V(E, z', x'_n, w'; A' | \tau, r) \\ & \left. + \delta(z') V(U, z', x'_u; A' | \tau, r) \right\} \end{aligned}$$

subject to

$$\begin{aligned} c + a' &\leq x \\ x'_u &= (1 + r)a' + b(z') + h \\ x'_e &= (1 + r)a' + (1 - \tau)w \\ x'_n &= (1 + r)a' + (1 - \tau)w' \\ A' &= F_A(A) \\ z' &= F_Z(z). \end{aligned}$$

In this formula, I use the fact that submarket tightness θ does not depend on the aggregate distribution of workers across states. Further, x_u is the cash-on-hand of unemployed workers, x_n is the cash-on-hand value for workers who find a new job at w' and x_e is the cash-on-hand value for workers who stay employed at wage w . The first two lines represent the problem of an unemployed worker, the last three lines represent the problem of an employed worker.

If we assume that the utility function is bounded and continuous, then \mathcal{V} is a set of bounded and continuous functions. It can be shown that the operator T_V maps $\mathcal{V} \rightarrow \mathcal{V}$. Using Blackwell's sufficiency conditions for a contraction and the assumptions on the boundedness of the sets defining the exogenous processes \mathcal{Z} and \mathcal{A} , as well as the choice sets \mathcal{W} and \mathcal{X} , one can show that T_R is a contraction with a fixed point $V^* \in \mathcal{V}$. Thus, the solution to the workers' problem does not depend on the distribution of workers across states Ω . This, in combination with the firm's problem above, constitutes a block recursive

equilibrium.

D Computational appendix

To solve the model in steady state, I employ the following algorithm:

- Guess values for τ and r
- Solve the firms' and the workers' problems
 - Guess wage and consumption policy functions for the workers
 - Given these policies together with the free entry condition, solve the firms' problem
 - Given the resulting job-finding probabilities and taking wage choices as given, update the consumption policy function using an endogenous grid point method
 - Given the job-finding probabilities and the consumption policy functions, update the wage policy functions
- Using the policy functions computed in the previous step, compute the distribution of agents across states, Ω .
- Check that asset markets clear and the government budget holds
- If not, update the guess and repeat

Along the transition path, when solving for the economy's response to a single unexpected aggregate productivity shock, I employ the sequence space Jacobian method proposed by [Auclert et al. \(2021\)](#). Using this approach speeds up the solving process by a factor of 20, compared to traditional methods of gradient descent.

I calculate the earnings procyclicality measures using a non-stochastic simulation approach. I