# Pooled Analysis

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### Preamble

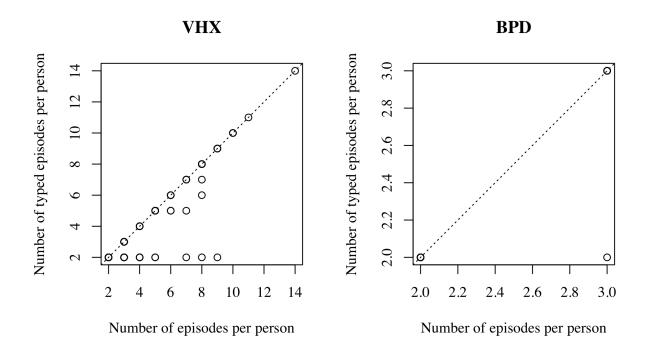
Load R packages, functions and data.

Summary of the data and the whole of the VHX data set versus the subset typed (in terms of number of episodes):

```
## Number of individuals in VHX and BPD trials: 644 and 655, respectively
## Number of individuals with at least one episode typed: 217
## Number of episodes typed: 710
## Number of recurrences typed: 494
## Overall in the dataset: breakdown by treatment group (individuals typed):
##
   AS CHQ PMQ
##
  13 90 114
##
## Within VHX: breakdown by treatment group (VHX individuals typed):
##
##
  AS CHQ PMQ
##
  13 90 34
##
## From BPD trial there are 80 individuals with total of 167 episodes typed (enrollment: 79; recurrent
## From VHX trial there are 137 individuals with total of 543 episodes typed (enrollment: 137; recurren
# Were all episodes typed if a person was selected for genotyping?
#-----
MS_pooled_summary = MS_pooled[!duplicated(MS_pooled$Episode_Identifier),] # Collapse rows due to COI >
# No. of typed episodes per person with one or more typed episodes in VHX and BPD
no_of_typed_epi_per_person_typed_VHX = table(MS_pooled_summary$ID[grep1('VHX',MS_pooled_summary$ID)])
no_of_typed_epi_per_person_typed_BPD = table(MS_pooled_summary$ID[grep1('BPD',MS_pooled_summary$ID)])
# No. of total episodes per person with one or more typed episodes in VHX and BPD
no_of_epi_per_person_typed_VHX = All_VHX_epi_count[names(All_VHX_epi_count) %in% names(no_of_typed_epi_
no_of_epi_per_person_typed_BPD = All_BPD_epi_count[names(All_BPD_epi_count) %in% names(no_of_typed_epi_
# VHX data set summary: breif because genotyping VHX was not exhaustive
XO = length(no_of_typed_epi_per_person_typed_VHX) # Number of people typed
ind_untyped = no_of_epi_per_person_typed_VHX != no_of_typed_epi_per_person_typed_VHX
X1 = sum(ind_untyped) # Number of people selected for genotyping but some episodes untyped
# How many untyped per person with incomplete set of episodes typed:
```

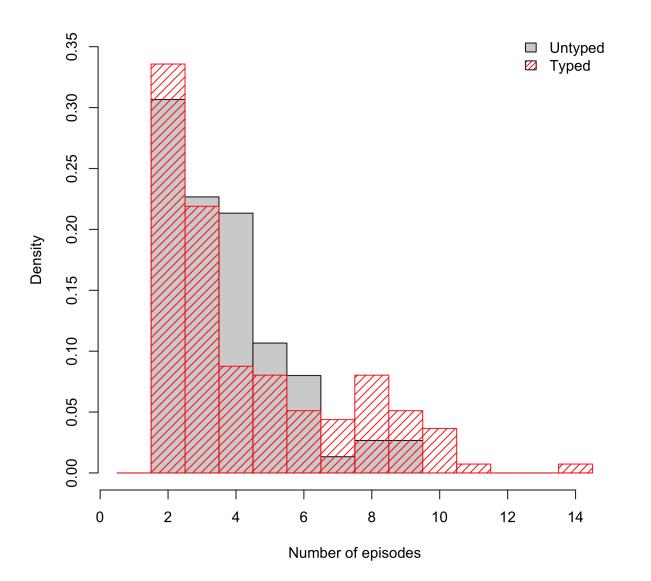
```
X2 = range(no_of_epi_per_person_typed_VHX[ind_untyped] - no_of_typed_epi_per_person_typed_VHX[ind_untyp
X3 = sum(no_of_epi_per_person_typed_VHX - no_of_typed_epi_per_person_typed_VHX) # Total number untyped
writeLines(sprintf('VHX: for %s of %s VHX individual/s selected for genotyping: %s to %s of their episo
## VHX: for 27 of 137 VHX individual/s selected for genotyping: 1 to 7 of their episodes were not typed
#-----
# BPD data set: comprehensive because genotyping BPD was exhaustive
#-----
# How many people who experience one or more recurrences had one or more episodes genotyped?
recurrences = All_BPD_epi_count[All_BPD_epi_count > 1]-1
indivs who recurred = names(recurrences)
indivs_who_were_typed = names(no_of_typed_epi_per_person_typed_BPD)
indivs_who_were_not_typed = indivs_who_recurred[!indivs_who_recurred %in% indivs_who_were_typed]
# Summary over individuals typed
X0 = length(indivs_who_were_typed) # Number of people typed
ind_untyped = no_of_epi_per_person_typed_BPD != no_of_typed_epi_per_person_typed_BPD
X1 = sum(ind_untyped) # Number of episodes untyped
# How many untyped per person with incomplete set of episodes typed:
X2 = range(no_of_epi_per_person_typed_BPD[ind_untyped] - no_of_typed_epi_per_person_typed_BPD[ind_untyp
X3 = sum(no_of_epi_per_person_typed_BPD - no_of_typed_epi_per_person_typed_BPD) # Total number untyped
# Individuals with not all episodes typed
ind_missing_typed_epi <- names(which(no_of_epi_per_person_typed_BPD != no_of_typed_epi_per_person_typed
# All episodes of the BPD individuals missing one or episodes
X4 = lapply(ind_missing_typed_epi, function(x){
 ind = grepl(x, uncensored_patientids)
 Combined_Time_Data$episode[!censored_ind][ind]
})
# Typed episodes of the BPD individuals missing one or episodes
X5 = lapply(ind_missing_typed_epi, function(x){
 ind = grepl(x, MS_pooled_summary$ID)
 MS_pooled_summary$Episode[ind]
})
X6 = lapply(1:length(X5), function(i){setdiff(X4[[i]], X5[[i]])}) # Not typed episodes
X7 = sum(sapply(X6, function(x)sum(x>1))) # Not typed recurrence
writeLines(paste(sprintf('BPD: of %s of the people who recurred: %s person/people with %s recurrence/s
                        length(unique(indivs_who_recurred)),
                        length(unique(indivs_who_were_not_typed)),
                        recurrences[indivs_who_were_not_typed]),
                 sprintf('Of %s of %s BPD individual/s selected for genotyping: %s to %s of their episo
                 sprintf('Of the %s episodes not typed %s were recurrences.',X3, X7),
                sprintf('In total there were %s recurrences: %s untyped.', sum(recurrences), recurrence
```

## BPD: of 81 of the people who recurred: 1 person/people with 1 recurrence/s was not selected for geno



```
x2 = All_VHX_rec_count[!names(All_VHX_rec_count) %in% names(x1)] # No of epi per person untyped
# setequal(names(x1), names(x2)) # Check mutually exclusive
# setequal(names(x2), unique(names(MS_pooled_summary$ID))) # Further check
max_rec = max(All_VHX_rec_count)
```

### VHX subset



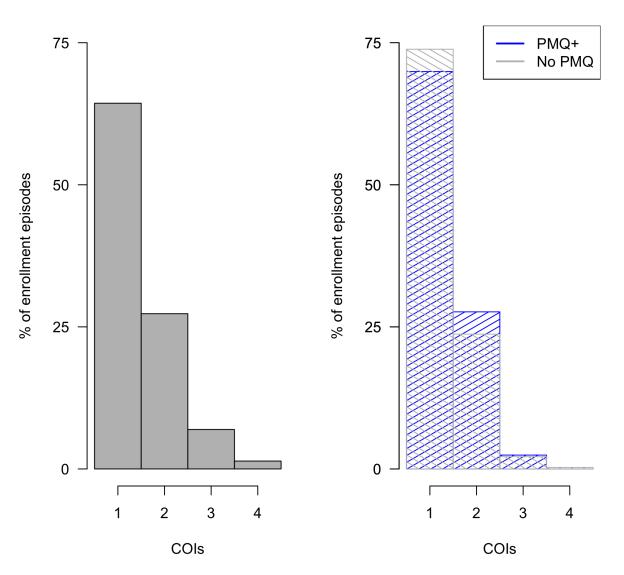
Summary of complexity of infection based on numbers of alleles observed. This is broken down by enrollment episodes (this is independent of drug given) and subsequent recurrences which could be drug dependent.

```
COIs = data.frame(t(sapply(unique(MS_pooled$Episode_Identifier), function(x){
  ind = which(MS_pooled$Episode_Identifier == x)
  c(MOI=max(MS_pooled$MOI_id[ind]),
     Enrollment = MS_pooled$Episode[ind[1]] == 1,
     Drug = MS_pooled$Treatment[ind[1]])
})))
```

```
COIs$MOI = as.numeric(COIs$MOI)
COIs$Enrollment = COIs$Enrollment=='TRUE'
COIs$PMQ = 0
COIs$PMQ[!COIs$Enrollment & COIs$Drug=='PMQ']=1
```

## **Enrollment episodes**

# **Recurrent episodes**



```
##
   glm(formula = MOI ~ enrollment + drug, family = "poisson", data = data.frame(MOI = COIs$MOI -
##
       1, enrollment = as.numeric(COIs$Enrollment), drug = COIs$PMQ))
##
## Deviance Residuals:
##
       Min
                      Median
                                   3Q
                 1Q
                                           Max
## -0.9526 -0.8065 -0.7595
                                        2.9375
                               0.6986
##
```

```
## Coefficients:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.24337
                          0.09667 -12.862 < 2e-16 ***
## enrollment
               0.45306
                          0.13982
                                    3.240 0.00119 **
## drug
               0.12007
                          0.18533
                                    0.648 0.51707
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##
      Null deviance: 619.82 on 709 degrees of freedom
## Residual deviance: 609.30 on 707 degrees of freedom
## AIC: 1057.2
##
## Number of Fisher Scoring iterations: 6
## Mean complexity of recurrent episodes is 1.29, and mean complexity of enrollment episodes is 1.45
## Median COI in VHX and BPD: 1 and 1, respectively
## 30 of 710 episodes (4 percent) with COI greater than or equal to 3
```

From this Poisson regression, there appears to be evidence that enrollment episodes have higher complexities of infection than recurrences. This implies that relapses are more likely to be single hypnozoite activated infections?

### Allele frequencies

First we define the set of microsatellite markers used in this analysis:

```
MSs_all = c("PV.3.502","PV.3.27","PV.ms8",

"PV.1.501","PV.ms1","PV.ms5",

"PV.ms6","PV.ms7","PV.ms16")
```

We use a multinomial-dirichlet model with subjective weight  $\omega$ .  $\omega = 0$  recovers the empirical allele frequencies.

## Number of episodes used to compute frequencies: 216

Calculate the effective marker cardinality for each microsatellite marker using a simulation approach.

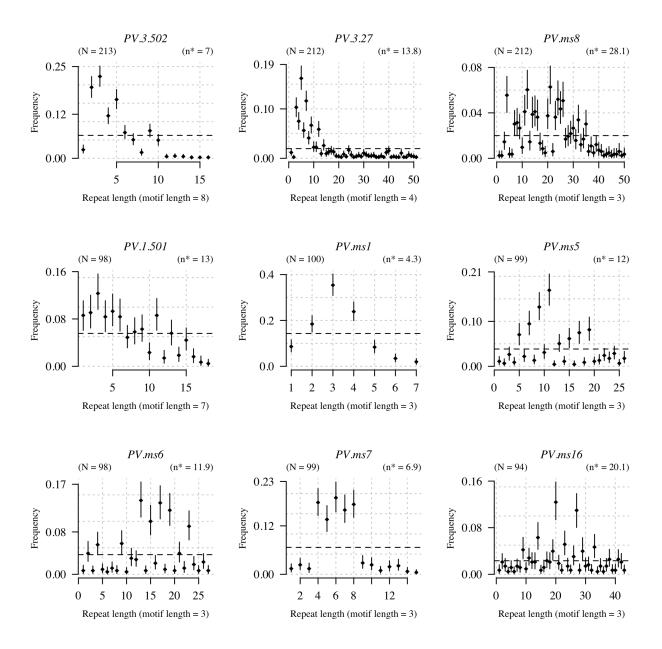
```
N = 10^6
Effective_Allele_size = list()
for(ms in MSs_all){
    n_obs_alleles = length(table(MS_pooled[Ind_Primary,ms]))
    draw1 = sample(x = names(Fs_Combined[[ms]]), replace = T, size = N, prob = Fs_Combined[[ms]])
    draw2 = sample(x = names(Fs_Combined[[ms]]), replace = T, size = N, prob = Fs_Combined[[ms]])
    x = mean(draw1 == draw2)
    n = 1/x
    writeLines(sprintf('The effective cardinality for %s with %s observed alleles is %s', ms, n_obs_allel
```

```
Effective_Allele_size[[ms]] = round(n,1)
}
## The effective cardinality for PV.3.502 with 13 observed alleles is 6.99
## The effective cardinality for PV.3.27 with 33 observed alleles is 13.76
## The effective cardinality for PV.ms8 with 46 observed alleles is 28.1
## The effective cardinality for PV.1.501 with 17 observed alleles is 13.03
## The effective cardinality for PV.ms1 with 7 observed alleles is 4.32
## The effective cardinality for PV.ms5 with 24 observed alleles is 12.04
## The effective cardinality for PV.ms6 with 25 observed alleles is 11.93
## The effective cardinality for PV.ms7 with 14 observed alleles is 6.91
## The effective cardinality for PV.ms16 with 39 observed alleles is 20.06
# The mean and range in our data set
writeLines(sprintf('The mean effective marker cardinality is %s, range: %s to %s',
                   round(mean(unlist(Effective_Allele_size)),2),
                   round(min(unlist(Effective_Allele_size)),2),
                   round(max(unlist(Effective_Allele_size)),2)))
```

## The mean effective marker cardinality is 13.01, range: 4.3 to 28.1

### Plotting allele frequencies

These are the mean posterior allele frequencies (dots) and 95% credible intervals (bars) given pooled enrollment data and  $\omega = D_{\text{weight\_Prior}}$ .



## Computing the probability of relatedness across infections

The approach is Bayesian and consists of the following:

- A prior probability vector for the recurrence state from the time-to-event model
- An allele frequency estimate from the posterior distribution of allele frequencies
- A likelihood based on the genetic data of being a *relapse*, a *recrudescence*, or a *reinfection* given the observed microsatellite data.

The following iterates through each individual and computes the probability of relatedness states.

```
# First we remove MS data for which there are no recurrent data
N_episodes_typed = table(MS_pooled$ID[!duplicated(MS_pooled$Episode_Identifier)])
MS_pooled = filter(MS_pooled, ID %in% names(N_episodes_typed[N_episodes_typed>1]))
```

#### Load the time-to-event priors

### Computation using full dataset

We use all 9MS markers (when available).

Full posterior computation

### Plot results

```
# Output of time-to-event model (sorted by episode number s.t. columns correspond)
Time_Estimates_1 = filter(Mod2_ThetaEstimates, Episode_Identifier %in% thetas_9MS$Episode_Identifier)
Time_Estimates_1 = arrange(Time_Estimates_1, Episode_Identifier)
# Outputs of genetic model w/wo time prior
# sorted by episode number s.t. columns correspond and drug added
thetas_9MS = arrange(thetas_9MS, Episode_Identifier)
```

```
thetas_9MS_Tagnostic = arrange(thetas_9MS_Tagnostic, Episode_Identifier)
thetas_9MS$drug = Time_Estimates_1$arm_num # Add drug
thetas_9MS_Tagnostic$drug = Time_Estimates_1$arm_num # Add drug

# Extract BPD only for BPD only plots
BPD_data = Thetas_full_post[grep('BPD',rownames(Thetas_full_post)),]
Thetas_BPD = thetas_9MS[grep('BPD', thetas_9MS$Episode_Identifier),]

# Extract prior used in absence of time-to-event
Time_agnostic_p = as.list(formals(post_prob_CLI)$p)
```

### Going from time-to-event prior to posterior

Plotted by radical cure versus no radical cure, as that is the most informative distinction here.

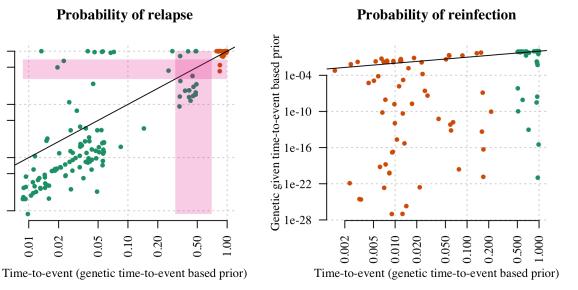
# Qenetic given time-to-event pased briout 0.500 0.500 0.050 0.005 0.005 0.001 0.05 0.10 0.20 0.50 1.00 Genetic given time-agnostic prior

Probability of relapse

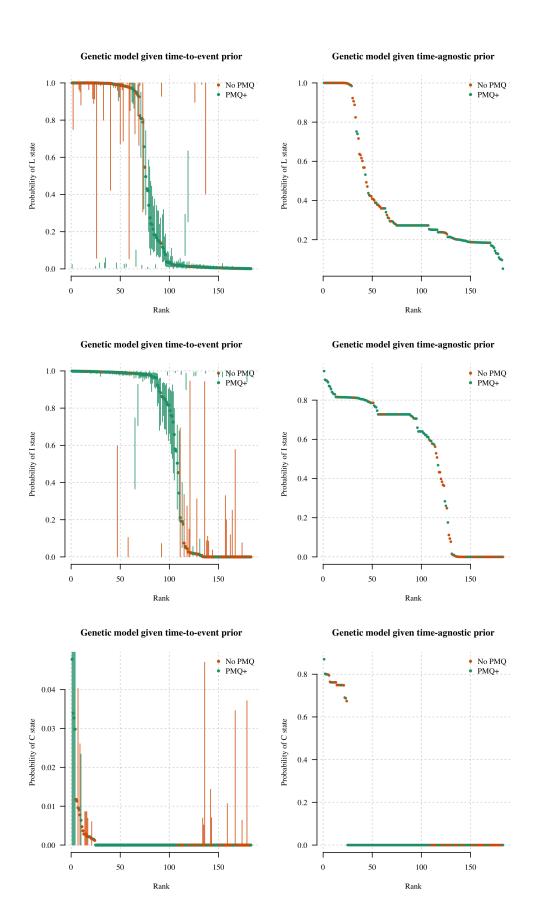
# Genetic given time-to-event based prior 1e-10 1e-16 1e-28 No PMO PMQ+ 1e-15 1e-03 1e-091e-27 1e-21 Genetic given time-agnostic prior

Probability of reinfection

# Probability of relapse O.000 description of the control of 0.50 0.02 0.20 0.01 0.05 0.10 1.00

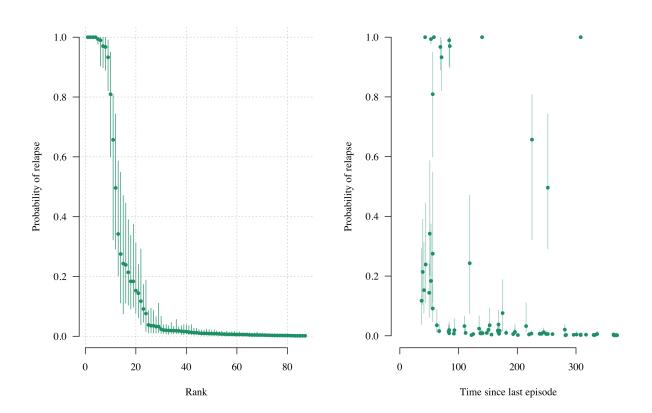


## Based on time-to-event alone, 66 of 66 No PMQ classified as relapse ## Based on genetic alone, 23 of 66 No PMQ classified as relapse Based on all available data, 61 of 66 No PMQ classified as relapse ## Based on time-to-event alone, 0 of 120 PMQ+ classified as relapse ## Based on genetic alone, 13 of 120 PMQ+ classified as relapse ## Based on all available data, 12 of 120 PMQ+ classified as relapse Probability of states, ordered from most to least likely:



### **BPD** Final plot

```
## The weighted average of recurrences which are estimated to be failures is 15.92%
## The weighted average of recurrences which are estimated to be relapses is 15.78%
## Number of BPD recurrences analysed: 87
```



## Extra computations for VHX: too complex episodes

We remove the IDs that can be straightforwardly calculated:

```
ind_calculated = which(MS_pooled_summary$Episode_Identifier %in% thetas_9MS$Episode_Identifier)
IDs_calculated = unique(MS_pooled_summary$ID[ind_calculated])
IDs_remaining = unique(MS_pooled_summary$ID[! MS_pooled_summary$ID %in% IDs_calculated])
writeLines(sprintf('individuals with more than two recurrences: %s',length(IDs_remaining)))
```

## individuals with more than two recurrences: 54

We blow up the pooled analysis into all pairs within individuals:

Construct adjacency graphs and compute probabilities of relapse and reinfection.

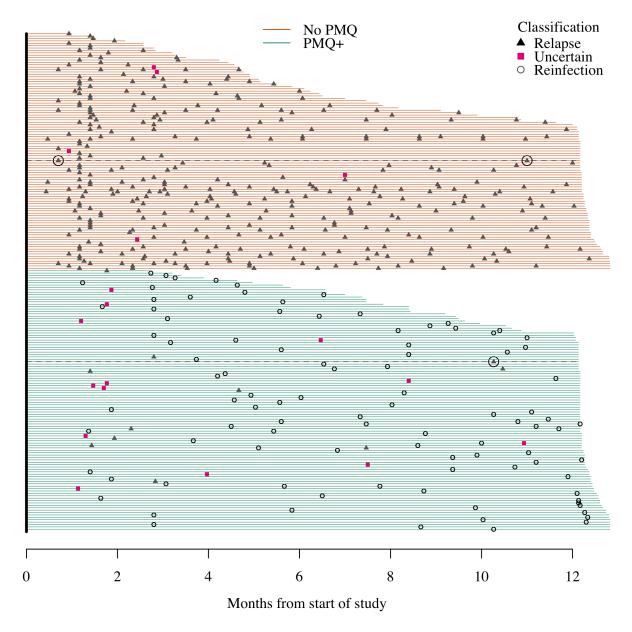
```
MS_pooled_summary$L_or_C_state = MS_pooled_summary$TotalEpisodes = NA
MS_pooled_summary$L_lower = MS_pooled_summary$L_upper = MS_pooled_summary$L_median = NA
MS_pooled_summary$C_lower = MS_pooled_summary$C_upper = MS_pooled_summary$C_median = NA
MS_pooled_summary$I_lower = MS_pooled_summary$I_upper = MS_pooled_summary$I_median = NA
# Arrange by complexity
```

```
# Get single rows per episode (throw away the extra MOI information)
MS_inflated_summary = MS_inflated[!duplicated(MS_inflated$Episode_Identifier) &
                                     MS_inflated$Episode==2,]
Results_Inflated$Episode_Identifier = as.character(Results_Inflated$Episode_Identifier)
for(i in 1:nrow(MS_inflated_summary)){
  if(!length(which(MS_inflated_summary$Episode_Identifier[i] ==
                   Results_Inflated$Episode_Identifier))>0){
    MS_inflated_summary = MS_inflated_summary[-i,]
    print('removing')
}
## [1] "removing"
## [1] "removing"
Results_Inflated$ID_True = NA
Results_Inflated$First_EpNumber = NA
Results_Inflated$Second_EpNumber = NA
# The ordering has changed so need to be careful about naming
for(i in 1:nrow(Results_Inflated)){
  ind_MS\_inflated = \\ which(MS\_inflated\_summary\\ \\ *Episode\_Identifier\\ \\ = \\ Results\_Inflated\\ \\ *Episode\_Identifier\\ \\ [inflated]
  Results_Inflated$ID_True[i] =
    MS_inflated_summary$ID_True[ind_MS_inflated]
  Results_Inflated$First_EpNumber[i] =
    MS_inflated_summary $First_EpNumber[ind_MS_inflated]
  Results_Inflated$Second_EpNumber[i] =
    MS_inflated_summary$Second_EpNumber[ind_MS_inflated]
}
# Iterate through the ones we can calculate in one go
episodes_full_model = unique(Thetas_full_post$Episode_Identifier)
cols_remove = grep('Episode_Identifier', colnames(Thetas_full_post))
Thetas_full_post = Thetas_full_post[, -cols_remove]
for(ep in episodes_full_model){
  ind1 = (MS_pooled_summary$Episode_Identifier==ep)
  ind2 = rownames(Thetas_full_post)==ep
  ## Summaries for relapse
  L_cols = grep('L',colnames(Thetas_full_post))
  MS_pooled_summary$L_upper[ind1] = quantile(unlist(Thetas_full_post[ind2,L_cols]),
                                              probs=upperCI, na.rm = T)
  MS_pooled_summary$L_lower[ind1] = quantile(unlist(Thetas_full_post[ind2,L_cols]),
                                              probs=lowerCI, na.rm = T)
  MS_pooled_summary$L_median[ind1] = quantile(unlist(Thetas_full_post[ind2,L_cols]),
                                               probs=0.5, na.rm = T)
  ## Summaries for recrudescence
  C_cols = grep('C',colnames(Thetas_full_post))
  MS_pooled_summary C_upper[ind1] = quantile(unlist(Thetas_full_post[ind2,C_cols]),
                                              probs=upperCI, na.rm = T)
  MS_pooled_summary$C_lower[ind1] = quantile(unlist(Thetas_full_post[ind2,C_cols]),
                                              probs=lowerCI, na.rm = T)
```

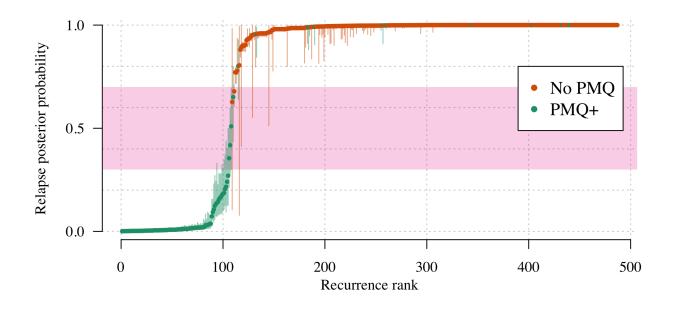
```
MS_pooled_summary$C_median[ind1] = quantile(unlist(Thetas_full_post[ind2,C_cols]),
                                             probs=0.5, na.rm = T)
 ## Summaries for reinfection
 I_cols = grep('I',colnames(Thetas_full_post))
 MS_pooled_summary$I_upper[ind1] = quantile(unlist(Thetas_full_post[ind2,I_cols]),
                                            probs=upperCI, na.rm = T)
 MS_pooled_summary$I_lower[ind1] = quantile(unlist(Thetas_full_post[ind2,I_cols]),
                                            probs=lowerCI, na.rm = T)
 MS_pooled_summary$I_median[ind1] = quantile(unlist(Thetas_full_post[ind2,I_cols]),
                                             probs=0.5, na.rm = T)
 # Just going to classify on relapse versus reinfection
 if(!is.na(MS_pooled_summary$L_upper[ind1])){
   if(MS_pooled_summary$L_upper[ind1]+MS_pooled_summary$C_upper[ind1] < Epsilon_lower){</pre>
     MS_pooled_summary$L_or_C_state[ind1] = 'I'
   } else if(MS_pooled_summary$L_lower[ind1]+MS_pooled_summary$C_lower[ind1] > Epsilon_upper){
     MS_pooled_summary$L_or_C_state[ind1] = 'L'
   } else {
     MS_pooled_summary$L_or_C_state[ind1] = 'Uncertain'
 } else {
   MS_pooled_summary$L_or_C_state[ind1] = NA
# Now iterate through the complex ones
for(i in 1:length(IDs_remaining)){
 id = IDs_remaining[i]
 Doubles_Thetas = filter(Results_Inflated, ID_True==id)
 for(ep in unique(Doubles_Thetas$Second_EpNumber)){
    # indices on the MS pooled summary
   ind1 = which(MS_pooled_summary$ID==id & MS_pooled_summary$Episode==ep)
    # indices on DOubles thetas: looking for relapse evidence
   ind2 = which(Doubles_Thetas$Second_EpNumber == ep)
    # index for recrudescence evidence
   ind3 = which(Doubles Thetas$Second EpNumber == ep &
                  Doubles_Thetas$First_EpNumber == (ep-1))
   best_match_relapse = which.max(Doubles_Thetas$L_median[ind2])
   if(length(best match relapse)>0){
      # Relapse probability
     MS_pooled_summary$L_lower[ind1] = Doubles_Thetas$L_min[ind2[best_match_relapse]]
     MS_pooled_summary$L_upper[ind1] = Doubles_Thetas$L_max[ind2[best_match_relapse]]
     MS_pooled_summary$L_median[ind1] = Doubles_Thetas$L_median[ind2[best_match_relapse]]
      # Reinfection probability
     MS_pooled_summary$I_lower[ind1] = Doubles_Thetas$I_min[ind2[best_match_relapse]]
     MS_pooled_summary$I_upper[ind1] = Doubles_Thetas$I_max[ind2[best_match_relapse]]
     MS_pooled_summary$I_median[ind1] = Doubles_Thetas$I_median[ind2[best_match_relapse]]
      # Recrudescence probability
```

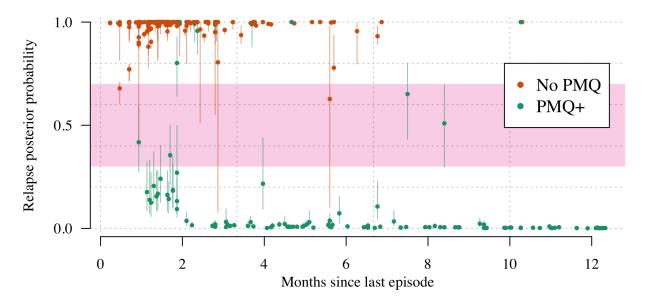
```
if(length(ind3)>0){
        # we can compute using previous episode
        MS_pooled_summary$C_lower[ind1] = Doubles_Thetas$C_min[ind3]
        MS_pooled_summary$C_upper[ind1] = Doubles_Thetas$C_max[ind3]
        MS_pooled_summary$C_median[ind1] = Doubles_Thetas$C_median[ind3]
      if(is.na(MS_pooled_summary$C_median[ind1])){
        MS pooled summary $C lower[ind1] =
          1-MS_pooled_summary$L_lower[ind1]+MS_pooled_summary$I_lower[ind1]
        MS pooled summary $C upper[ind1] =
          1-MS_pooled_summary$L_upper[ind1]+MS_pooled_summary$I_upper[ind1]
        MS_pooled_summary$C_median[ind1] =
          1-MS pooled summary L median[ind1]+MS pooled summary I median[ind1]
      }
    }
    if(!is.na(MS_pooled_summary$C_median[ind1])){
      if(MS_pooled_summary$L_upper[ind1] < MS_pooled_summary$L_lower[ind1]){</pre>
        writeLines(sprintf('Problem with ID %s',id))
        stop()
      }
      if(MS_pooled_summary$L_upper[ind1]+MS_pooled_summary$C_upper[ind1] < Epsilon_lower){</pre>
        MS_pooled_summary$L_or_C_state[ind1] = 'I'
      } else if(MS_pooled_summary$L_lower[ind1] +MS_pooled_summary$C_lower[ind1] > Epsilon_upper){
        MS pooled summary$L or C state[ind1] = 'L'
      } else {
        MS_pooled_summary$L_or_C_state[ind1] = 'Uncertain'
      }
    }
  }
}
MS_pooled_summary$Drug = MS_pooled_summary$FU = NA
for(id in MS_pooled_summary$ID){
  ind = MS_pooled_summary$ID==id
  MS_pooled_summary$TotalEpisodes[ind] = max(MS_pooled_summary$Episode[ind])
  MS_pooled_summary$Drug[ind] = as.numeric(
    Combined_Time_Data$arm_num[Combined_Time_Data$patientid==id][1] == 'CHQ/PMQ') + 2
  MS_pooled_summary$FU[ind] = Combined_Time_Data$FU_time[Combined_Time_Data$patientid==id][1]
MS_pooled_summary$Plotting_pch_Values =
  as.numeric(mapvalues(MS_pooled_summary$L_or_C_state, from = c('L', 'Uncertain', 'I'), to = c(17,15,1)))
MS pooled summary $Plotting col Values =
  as.numeric(mapvalues(MS_pooled_summary$L_or_C_state, from = c('L', 'Uncertain', 'I'), to = 1:3))
# How many too complex to generate estimate for?
ind_recur = MS_pooled_summary$Episode > 1 # Filter out enrollment
ind_complex_recur = is.na(MS_pooled_summary$L_median[ind_recur])
no_complex_recur = sum(ind_complex_recur) # Recurrences with NAs
# How many of the complex infections result in the loss of an individual
no_indiv_removed = sum(N_episodes_typed[MS_pooled_summary$ID[ind_recur][ind_complex_recur]] <= 2)
# Final number of people with recurrences analysed total and by trial
indiv_recur_analysed = length(unique(MS_pooled_summary$ID[ind_recur][!ind_complex_recur]))
```

## Of 493 recurrences analysed, 6 were too complex to estimate recurrence state probabilities, resultin



## The Coatney style plot is showing 487 recurrences in 208 individuals





Individuals who appear to relapse very late (more than 300 days after last episode):

BPD\_27\_1

## 60

BPD\_27

1

PMQ

```
## 61
         BPD_27
                        2
                                      BPD_27_2
                                                        PMQ
                                                                   1
         BPD 27
                        2
                                      BPD_27_2
                                                                   2
## 62
                                                        PMQ
   355 VHX 235
                                     VHX 235 1
                        1
                                                        CHQ
                                                                   1
   356 VHX_235
                                     VHX_235_1
                                                        CHQ
                                                                   2
##
                        1
   357 VHX_235
##
                        2
                                     VHX_235_2
                                                        CHQ
                                                                   1
                        3
                                     VHX 235 3
                                                                   1
   358 VHX 235
                                                        CHQ
##
        timeSinceLastEpisode timeSinceEnrolment PV.1.501 PV.3.27 PV.3.502
##
## 60
                              0
                                                    0
                                                               3
                                                                       33
                                                                                   7
## 61
                            308
                                                  308
                                                               3
                                                                       33
                                                                                   7
                            308
                                                  308
                                                               3
                                                                       35
                                                                                   7
## 62
## 355
                              0
                                                    0
                                                               1
                                                                        5
                                                                                   2
                                                                                   2
                              0
                                                    0
                                                                         5
## 356
                                                               1
                             21
## 357
                                                   21
                                                               1
                                                                         5
                                                                                   3
   358
                            309
                                                                         5
##
                                                  330
                                                               1
                                                                                   3
        PV.ms1 PV.ms16 PV.ms5 PV.ms6 PV.ms7 PV.ms8
##
## 60
              4
                      27
                              24
                                      15
                                                5
                                                       17
              4
                      27
                                      15
                                                5
## 61
                              24
                                                       17
   62
              4
                      27
                              24
                                      15
                                                5
                                                       17
##
              3
                      23
                                       9
##
   355
                              13
                                               10
                                                       12
##
   356
              3
                      23
                              13
                                      15
                                               10
                                                       33
## 357
              4
                      20
                              13
                                        9
                                               10
                                                       12
## 358
              4
                      23
                                      15
                                               10
                                                       12
                              11
```

The summaries of the final dataset. Results for all those genotyped who did not receive primaquine (artesunate or chloroquine monotherapy):

```
## ## AS CHQ PMQ
## 11 88 109
```

- ## In no-primaquine individuals, the weighted average of relapse is 0 (97-99.3), for 366 recurrences
- ## In no-primaquine individuals, the weighted average of recrudescences is 0 (1.4-2.2), for 366 recurred
- ## In no-primaquine individuals, the weighted average of reinfections is 0 (0.1-2), for 366 recurrences Results for all those genotyped who did receive primaquine (VHX and BPD studies combined):
- ## In primaquine treated individuals, the weighted average of relapses is 0 (11.7-17.7), for 121 recurr
- ## In primaquine treated individuals, the weighted average of recrudescences is 0 (0-0.4), for 121 recu
- ## In primaquine treated individuals, the weighted average of reinfections is 0 (82.3-88), for 121 recu
- Results for all those genotyped who did receive primaquine in the VHX study (unknown denominator)
- ## In primaquine treated individuals (VHX), the weighted average of relapses is 0 (8-14.1), for 34 recu
- ## In primaquine treated individuals (VHX), the weighted average of recrudescences is 0 (0-0.3), for 34
- ## In primaquine treated individuals (VHX), the weighted average of reinfections is 0 (85.9-91.7), for
- Results for all those genotyped who did receive primaquine, only in the BPD study (known denominator)
- ## In primaquine treated individuals (BPD), the weighted average of relapses is 0 (13.1-19.2), for 87 r
- ## In primaquine treated individuals (BPD), the weighted average of recrudescences is 0 (0-0.4), for 87
- ## In primaquine treated individuals (BPD), the weighted average of reinfections is 0 (80.8-86.5), for

### False positive rate of relapse

We want to know how often our model estimates evidence of relapse across pairs of episodes when the episodes are in different people (i.e. have not possibility of being a relapse)

```
if(RUN MODELS FALSE POSITIVE){
  # check if the massive pairwise dataset has been made, if not make it
  # (takes a long time ~20hours)
  if(!"APC_MSdata.bigRData"%in%list.files(path = '../RData/LargeFiles/')){
    # The pooled MS data from BPD and VHX
   load('../RData/GeneticModel/MS data PooledAnalysis.RData')
   APC_MSdata = Make_All_Pairwise_Comparisons(MS_data = MS_pooled, ncores=42)
   save(APC_MSdata, file = '../RData/LargeFiles/APC_MSdata.bigRData')
  }
  load('../RData/LargeFiles/APC MSdata.bigRData')
  print('The inflated pairwise dataset is available, now running the analysis...')
  # Run the genetic model on the pairwise data
  tic()
  Inflated_Results = post_prob_CLI(MSdata = APC_MSdata,
                                   Fs = Fs_Combined,
                                   UpperComplexity = 10<sup>6</sup>,
                                   verbose = F,
                                    cores = 42)
 toc()
  save(Inflated_Results, file = '../RData/LargeFiles/Inflated_Results.bigRData')
} else {
  load('../RData/LargeFiles/Inflated Results.bigRData')
  Inflated_Results = Inflated_Results[!is.na(Inflated_Results$L),]
 load('../RData/LargeFiles/APC_MSdata.bigRData')
}
## The false-positive discovery rate of the genetic model is estimated as 2.15 percent.
## This is based on 90194 pairwise comparisons
```

## Analysis of radical cure efficacy in BPD

Almost all episodes in BPD were typed. Therefore we can estimate the true efficacy comparing with historical controls (VHX).

```
Mod2_ThetaEstimates$Failure_Identifier =
  apply(Mod2_ThetaEstimates, 1,
        function(x) paste(x['patientid'], as.integer(x['episode']),sep='_'))
sss=0
for(i in 1:nrow(Combined Time Data)){
  ep_id = Combined_Time_Data$Episode_Identifier[i]
  # we look one ahead
 MS_id = paste(Combined_Time_Data$patientid[i],
                as.integer(Combined Time Data$episode[i])+1, sep=' ')
  # If in MS_final then use full probability
  if(MS_id %in% MS_final$Episode_Identifier){
   Combined_Time_Data$Reinfection_Probability[i] =
      MS_final$I_median[MS_final$Episode_Identifier==MS_id]
   Combined_Time_Data$Reinfection_Probability_UL[i] =
      MS_final$I_upper[MS_final$Episode_Identifier==MS_id]
    Combined_Time_Data$Reinfection_Probability_LL[i] =
      MS_final$I_lower[MS_final$Episode_Identifier==MS_id]
  } else { # use the time to event model
    ind = which(Mod2 ThetaEstimates$Failure Identifier==MS id)
    if(length(ind)>0){
      Combined_Time_Data$Reinfection_Probability[i] =
        Mod2_ThetaEstimates$ReInfection_mean_theta[ind]
      Combined Time Data$Reinfection Probability UL[i] =
        Mod2 ThetaEstimates $ReInfection 975 theta[ind]
      Combined Time Data$Reinfection Probability LL[i] =
        Mod2_ThetaEstimates$ReInfection_025_theta[ind]
      sss=sss+1
   }
  }
}
```

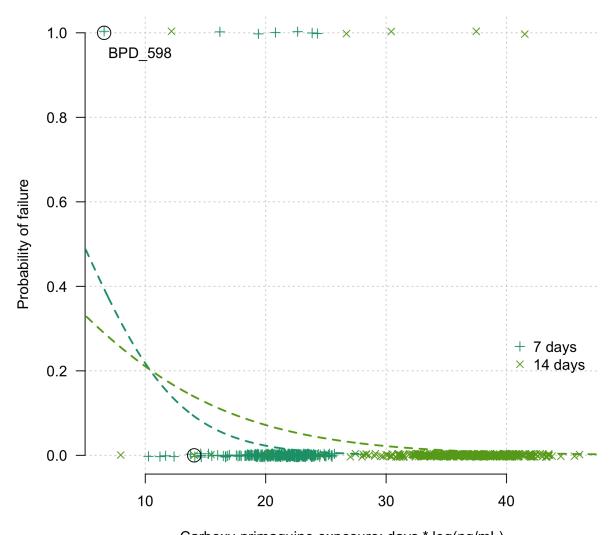
Now we look at whether the PK (carboxy-primaquine) can predict failure: First we add the carboxy to the dataset:

```
Combined_Time_Data = arrange(Combined_Time_Data, patientid, episode)
load('../RData/PK_data/BPD_pk.RData')
BPD_pk = filter(BPD_pk, !is.na(Episode))
Combined_Time_Data$log10_carboxyPMQ = NA
Combined_Time_Data$log10_PMQ = NA
Combined_Time_Data$NumberDaysPMQ = NA
# The default is 14 days
Combined_Time_Data$NumberDaysPMQ[Combined_Time_Data$Censored != 1 &
                                   Combined_Time_Data$arm_num=='CHQ/PMQ'] = 14
for(i in 1:nrow(Combined_Time_Data)){
  id = Combined_Time_Data$patientid[i]
  ep_i = Combined_Time_Data$episode[i]
  all id eps = Combined Time Data episode [Combined Time Data patientid == id]
  pk_ind = which(BPD_pk$ID == id & BPD_pk$Episode==ep_i)
  if(length(pk ind)>0){
    if(length(pk_ind)>1) print(id)
    Combined_Time_Data$log10_carboxyPMQ[i] = mean(BPD_pk$log10_carboxyPQ_PK[pk_ind])
   Combined_Time_Data$log10_PMQ[i] = mean(BPD_pk$log10_PQ_PK[pk_ind])
```

```
Combined_Time_Data$NumberDaysPMQ[i] = BPD_pk$NumberofPKDays[pk_ind[1]]
 }
}
## [1] "BPD 34"
We exclude the two recurrences seen in patient BPD 444 who was G6PD deficient and received the 8 weekly
regimen (not daily dosing).
# These are two outliers - have discussed with Cindy
BPD444_recurrences = Combined_Time_Data$patientid=='BPD_44' & Combined_Time_Data$episode>1
BPD_598 = which(Combined_Time_Data$patientid=='BPD_598')
ind_keep = !BPD444_recurrences #6 !BPD_598
Combined Time Data$Failure YN = Combined Time Data$Reinfection Probability < 0.5
mod = glmer(Failure_YN ~ log10_carboxyPMQ + NumberDaysPMQ +
              (1 | patientid),
            family = 'binomial', data=Combined_Time_Data[ind_keep,])
summary(mod)
## Generalized linear mixed model fit by maximum likelihood (Laplace
     Approximation) [glmerMod]
## Family: binomial (logit)
## Formula: Failure_YN ~ log10_carboxyPMQ + NumberDaysPMQ + (1 | patientid)
     Data: Combined_Time_Data[ind_keep, ]
##
##
##
        AIC
                 BIC
                       logLik deviance df.resid
##
      119.8
               138.1
                        -55.9
                                 111.8
##
## Scaled residuals:
##
      Min
               1Q Median
                                3Q
## -0.5862 -0.1303 -0.1095 -0.0912 13.6453
##
## Random effects:
             Name
## Groups
                          Variance Std.Dev.
## patientid (Intercept) 1.624e-14 1.274e-07
## Number of obs: 721, groups: patientid, 639
##
## Fixed effects:
##
                    Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                    2.48402
                                1.77436 1.400 0.161529
## log10_carboxyPMQ -1.73527
                                0.49252 -3.523 0.000426 ***
## NumberDaysPMQ
                   -0.18324
                                0.09035 -2.028 0.042554 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##
               (Intr) 110 PM
## lg10_crbPMQ -0.860
## NumbrDysPMQ -0.728 0.315
# Plot the data and model
xs = seq(0,4,by=.01)
par(las = 1, bty='n')
regimen_colors = brewer.pal(8, 'Dark2')[c(1,5)]
plot(Combined Time Data$log10 carboxyPMQ[ind keep] *Combined Time Data$NumberDaysPMQ[ind keep],
```

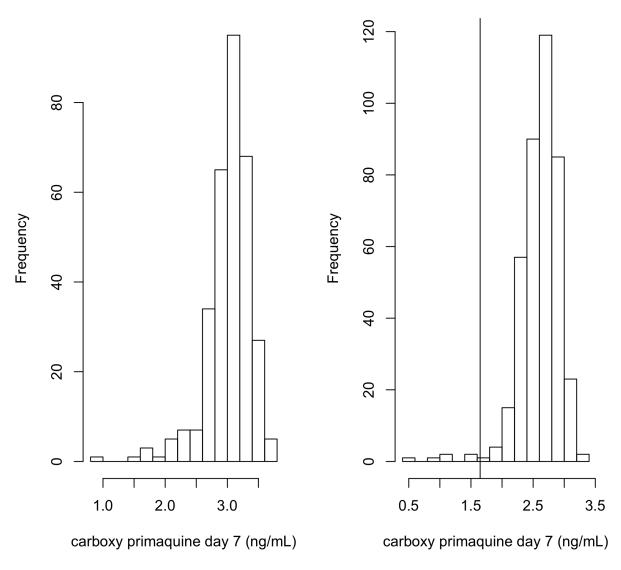
jitter(as.numeric(Combined\_Time\_Data\$Failure\_YN[ind\_keep]), factor = 0.02),

```
col = regimen_colors[mapvalues(Combined_Time_Data$NumberDaysPMQ[ind_keep],c(7,14),c(1,2))],
     pch = mapvalues(Combined_Time_Data$NumberDaysPMQ[ind_keep],c(7,14),c(3,4)),
     xlab = 'Carboxy-primaquine exposure: days * log(ng/mL)',
     panel.first = grid(), ylab = 'Probability of failure')
legend(x = 40, y = 0.3, bty='n', col =regimen_colors,
      pch=c(3,4), legend = c('7 days','14 days'))
lines(xs*7, predict(mod, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=7,
                                            patientid='new'),allow.new.levels=T,
                    type='response'), lwd=2, col= regimen_colors[1], lty=2)
lines(xs*14, predict(mod, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=14,
                                             patientid='new'),allow.new.levels=T,
                     type='response'), lwd=2, col= regimen_colors[2], lty = 2)
points (Combined Time Data$log10 carboxyPMQ[BPD 598] *Combined Time Data$NumberDaysPMQ[BPD 598],
       Combined_Time_Data$Failure_YN[BPD_598], cex=2)
text(Combined_Time_Data$log10_carboxyPMQ[BPD_598[1]]*
       Combined_Time_Data$NumberDaysPMQ[BPD_598[1]]+3,
     Combined_Time_Data$Failure_YN[BPD_598[1]]-0.05, labels = 'BPD_598')
```



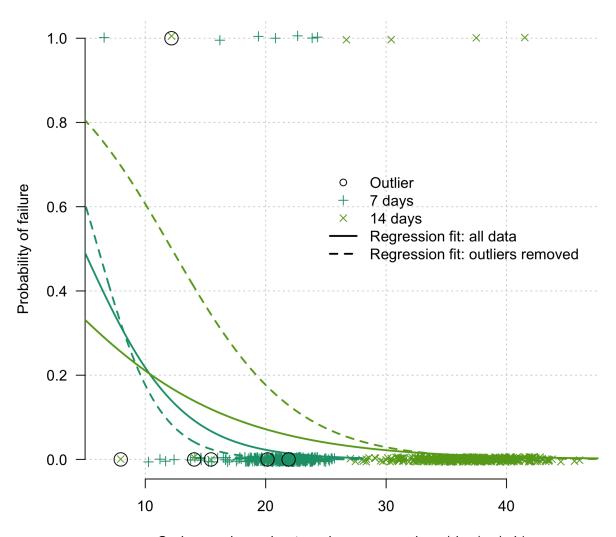
Carboxy-primaquine exposure: days \* log(ng/mL)

Now we remove outliers and fit the same model (CPMQ outliers)



## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]

```
## Family: binomial (logit)
## Formula: Failure YN ~ log10 carboxyPMQ + NumberDaysPMQ + (1 | patientid)
      Data: Combined Time Data[ind keep & !outliers14 & !outliers7, ]
##
##
        AIC
                 BIC
                       logLik deviance df.resid
       48 5
                60.5
                        -21.3
                                  42.5
##
##
## Scaled residuals:
##
       Min
                1Q Median
                                30
                                       Max
## -0.2931 -0.1136 -0.0856 -0.0675 17.9880
## Random effects:
## Groups
              Name
                          Variance Std.Dev.
## patientid (Intercept) 2.102e-14 1.45e-07
## Number of obs: 396, groups: patientid, 352
##
## Fixed effects:
##
                    Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                       2.404
                                  4.488 0.536
                                                   0.592
                                  1.832 -1.507
## log10 carboxyPMQ
                      -2.760
                                                    0.132
##
## Correlation of Fixed Effects:
##
               (Intr)
## lg10 crbPMQ -0.994
## fit warnings:
## fixed-effect model matrix is rank deficient so dropping 1 column / coefficient
Compare results with and without outliers:
par(las = 1, bty='n')
plot(Combined_Time_Data$log10_carboxyPMQ[ind_keep] *Combined_Time_Data$NumberDaysPMQ[ind_keep],
     jitter(as.numeric(Combined Time Data$Failure YN[ind keep]),factor = 0.03),
     col = regimen_colors[mapvalues(Combined_Time_Data$NumberDaysPMQ[ind_keep],c(7,14),1:2)],
     pch = mapvalues(Combined_Time_Data$NumberDaysPMQ[ind_keep],c(7,14),c(3,4)),
     xlab = 'Carboxy-primaquine trough exposure: days * log(ng/mL)',
     ylab = 'Probability of failure', panel.first = grid())
legend(x = 25, y = 0.7, bty='n', col = c(1, regimen_colors, 1,1),
       pch=c(1,3,4,NA,NA), lty = c(NA,NA,NA,1,2), lwd=c(NA,NA,NA,2,2),
       legend = c('Outlier','7 days','14 days',
                  'Regression fit: all data', 'Regression fit: outliers removed'))
lines(xs*7, predict(mod, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=7,
                                            patientid='new'),allow.new.levels=T,
                    type='response'), lwd=2, col= regimen_colors[1], lty=1)
lines(xs*14, predict(mod, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=14,
                                             patientid='new'),allow.new.levels=T,
                     type='response'), lwd=2, col= regimen_colors[2], lty = 1)
lines(xs*7, predict(mod_No_Outliers, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=7,
                                                         patientid='new'), allow.new.levels=T,
                    type='response'), lwd=2, col= regimen_colors[1], lty=2)
lines(xs*14, predict(mod_No_Outliers, newdata=data.frame(log10_carboxyPMQ=xs, NumberDaysPMQ=14,
                                                         patientid='new'),allow.new.levels=T,
                     type='response'), lwd=2, col= regimen_colors[2], lty = 2)
# outline the outliers
```



Carboxy-primaguine trough exposure: days \* log(ng/mL)

Now we calculate a compressed dataset and failure for each individual

```
# now we calculate the primaquine failure rate
# For individuals with two episodes: P(failure) = 1 - P(Rec 1 = I)*P(Rec 2 = I)
Summary_data = Combined_Time_Data[!duplicated(Combined_Time_Data$patientid),]
Summary_data$Failure_UL = Summary_data$Failure_LL =
    Summary_data$Failure = Summary_data$CPMQ =
    Summary_data$CPMQ = NA
for(i in 1:nrow(Summary_data)){
    ind = which(Combined_Time_Data$patientid==Summary_data$patientid[i])
    Summary_data$Failure[i] = 1-prod(Combined_Time_Data$Reinfection_Probability[ind],na.rm=T)
```

```
Summary_data$Failure_UL[i] = 1-prod(Combined_Time_Data$Reinfection_Probability_UL[ind],na.rm=T)
Summary_data$Failure_LL[i] = 1-prod(Combined_Time_Data$Reinfection_Probability_LL[ind],na.rm=T)
Summary_data$CPMQ[i] = median(Combined_Time_Data$log10_carboxyPMQ[ind],na.rm=T)
}
BPD_data = Summary_data[grep('BPD', Summary_data$patientid),]

P_Failure=100*sum(BPD_data$Failure)/nrow(BPD_data)
# invert the intervals here - optimistic for not failure = pessimistic for failure
P_Failure_UL = 100*sum(BPD_data$Failure_LL)/nrow(BPD_data)
P_Failure_LL = 100*sum(BPD_data$Failure_UL)/nrow(BPD_data)

writeLines(sprintf('The primaquine failure rate in the %s individuals is %s%% (%s-%s) over the course of nrow(BPD_data), round(P_Failure,2), round(P_Failure_LL,2), round(Sum(BPD_data$FU_time)/365)))
```

## The primaquine failure rate in the 655 individuals is 2.59% (1.96-3.58) over the course of 522 years

### Extra Analyses

### Looking at the effect of inbreeding coefficient

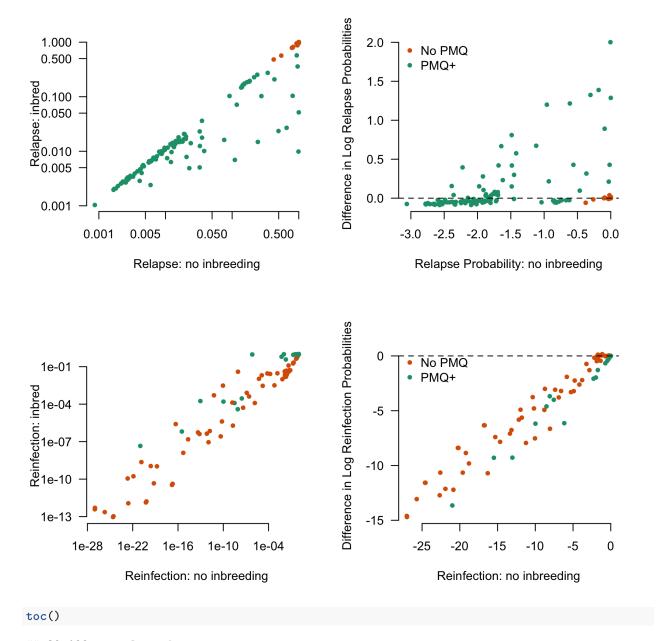
Our model has a parameter  $\alpha$  which defines the level of inbreeding within the population. Taylor is developing methods for the estimation of  $\alpha$  from genetic data (in preparation).

We look at the sensitivity of the results (all the above is with  $\alpha = 0$ ) for a reasonable upper bound of  $\alpha = 0.175$ .

We rerun the analysis on the single run isolates (low computational complexity):

```
alphaUpper = 0.175
if(RUN_MODELS_SINGLE_SIMPLE){
 # Run (with time-to-event)
 # Approx 100 secs per full model run
 tic()
 thetas 9MS alphaUpper = post prob CLI(MSdata = MS pooled, Fs = Fs Combined,
                                   p = p, cores = 6, verbose = F, alpha = alphaUpper)
 thetas_9MS_alphaUpper$Episode_Identifier = rownames(thetas_9MS_alphaUpper)
 save(thetas_9MS_alphaUpper, file = '../RData/GeneticModel/thetas_9MS_alphaUpper.RData')
 toc()
} else {
 load('../RData/GeneticModel/thetas_9MS_alphaUpper.RData')
par(las=1, bty='n', mfrow=c(2,2))
plot(thetas_9MS$L,
    thetas 9MS alphaUpper$L,
    log = 'xy',
    ylab = 'Relapse: inbred',
    xlab = 'Relapse: no inbreeding',
    col= drug_cols2[thetas_9MS$drug],pch=20)
lines(c(-10,10),c(-10,10),lty=2)
```

```
plot(log10(thetas_9MS$L), log10(thetas_9MS$L)-log10(thetas_9MS_alphaUpper$L),
     ylab = 'Difference in Log Relapse Probabilities',
     xlab = 'Relapse Probability: no inbreeding',
     col= drug_cols2[thetas_9MS$drug],pch=20)
abline(h=0,lty=2)
legend('topleft', legend = c('No PMQ', 'PMQ+'), col = drug_cols2[2:3], pch = 20, bty = 'n')
###**** Reinfection : comparison ****####
par(las=1, bty='n')
plot(thetas_9MS$I, thetas_9MS_alphaUpper$I,
     log = 'xy',
     ylab = 'Reinfection: inbred',
     xlab = 'Reinfection: no inbreeding',
     col= drug_cols2[thetas_9MS$drug],pch=20)
lines(c(-10,10),c(-10,10),lty=2)
plot(log10(thetas_9MS$I), log10(thetas_9MS$I)-log10(thetas_9MS_alphaUpper$I),
     ylab = 'Difference in Log Reinfection Probabilities',
     xlab = 'Reinfection: no inbreeding',
     col= drug_cols2[thetas_9MS$drug],pch=20)
legend('topleft', legend = c('No PMQ', 'PMQ+'), col = drug_cols2[2:3], pch = 20, bty = 'n')
abline(h=0,lty=2)
```



### ## 32.022 sec elapsed

Interpretation: Adding the inbreeding coefficient slightly changes some of the probabilities of relapse for some primaquine treated individuals (only green dots are being shifted). This means that inbreeding would imply that fewer of the primaquine treated episodes are relapses, implying higher efficacy of the drug.

For the non-primaquine group, it is just tempering the very low probabilities of reinfection seen for some episodes.

In conclusion, this isn't changing the results significantly and would imply a greater primaquine efficacy than reported in the paper.