

# Errata for Proof of Eq. (S2.42)

Jiaxin Qiu, Zeng Li, Jianfeng Yao

(LATEST UPDATE: JULY 30, 2025)

## Abstract

This document provides the errata for the proof of Eq. (S2.42) in [Qiu et al. \[2023\]](#): for any  $\varepsilon > 0$  and any  $\ell > 0$ ,

$$\Pr(\|\mathbf{B}_n\| \geq \eta + \varepsilon) = o(n^{-\ell}), \quad (1)$$

where  $\eta = 2 \limsup \|\Sigma\|/\sqrt{\theta}$ ,  $\theta = \lim_p p^{-1} \text{tr}(\Sigma^2)$ . To correct the proof, we impose an additional assumption that  $p = \Theta(n^t)$  for some  $t > 1$ .

## 1 Counterexample for Eq. (S2.45)

In the proof of Eq. (S2.42) in [Qiu et al. \[2023\]](#), we used an intermediate inequality (S2.45) as follows:

$$\max_{\|z\|=1} \sum_{i=1}^n \left\{ \sum_{j \neq i} (\mathbf{x}'_i \Sigma \mathbf{x}_j) z_j \right\}^2 \leq \|\Sigma\|^2 \max_{\|z\|=1} \sum_{i=1}^n \left\{ \sum_{j \neq i} (\mathbf{x}'_i \mathbf{x}_j) z_j \right\}^2.$$

Unfortunately, this inequality turns out to be incorrect. The following provides a counterexample:

$$p = n = 2, \quad \mathbf{x}_1 = \begin{pmatrix} 1 \\ 0 \end{pmatrix}, \quad \mathbf{x}_2 = \begin{pmatrix} 0 \\ 1 \end{pmatrix}, \quad \Sigma = \begin{pmatrix} 1 & \varepsilon \\ \varepsilon & 1 \end{pmatrix}, \quad 0 < \varepsilon < 1, \quad \implies \quad \text{LHS} = \varepsilon^2 > \text{RHS} = 0.$$

## 2 Corrected Proof of Eq. (S2.42)

**Sketch of the proof:**

- **Step 1:** Establish the result for the Gaussian case. Under the Gaussian assumption,  $\Sigma$  can be diagonalized, then we apply the method of [Chen and Pan \[2012\]](#) to prove the result.
- **Step 2:** Use the Lindeberg replacement method and the mathematical induction to generalize the result to the non-Gaussian case.
  - To apply the Lindeberg replacement method, we impose an additional assumption that  $p = \Theta(n^t)$  for some  $t > 1$ . This growth condition ensures that the total error introduced by the replacement steps remains negligible in the asymptotic regime.

### 2.1 Step 1: Proof under Gaussian case

Recall that

$$\mathbf{B}_n = \frac{1}{\sqrt{npb_p}} \underbrace{\{\mathbf{X}'\Sigma\mathbf{X} - \text{diag}(\mathbf{X}'\Sigma\mathbf{X})\}}_{=: \mathbf{B}_n^\circ}.$$

We take  $\delta = \delta_n$  even  $k = k_n$  and satisfying, as  $n, p \rightarrow \infty$ ,  $\delta \rightarrow 0$ ,  $\delta(np)^{1/4} \rightarrow 0$ ,

$$\frac{k}{\log n} \rightarrow \infty, \quad \frac{\delta^{1/3}k}{\log n} \rightarrow 0, \quad \frac{\delta^2 \sqrt[4]{n}}{k^3} \geq 1, \quad \frac{\delta^2 \sqrt{p/n}}{k} > 1, \quad \frac{k}{\log(\delta^2 \sqrt{p/n/k})} > 1, \quad \frac{\delta^2 k^3}{\log(2\delta^2 \sqrt{p/n/k})} \leq 1. \quad (2)$$

*Remark 1.* • We assume that  $k$  is even such that the inequality holds:  $\|\mathbf{B}_n\|^k \leq \sum_{i=1}^n \lambda_i^k(\mathbf{B}_n) = \text{tr}(\mathbf{B}_n^k)$ .

- The first three conditions for  $k$  in (2) are taken from [Chen and Pan \[2012\]](#) to obtain (3) below. The 4th and 5th conditions are used in (8), and the last condition is used in (11).
- Recall that  $p = \Theta(n^t)$  for some  $t > 1$ . Let  $\delta = n^{-\alpha}$  and  $k = k_n = (\log n)^2$ . For any  $t > 1$ , we can choose  $\alpha = \min\{\frac{1}{16}, \frac{t-1}{8}\}$  to make all the conditions in (2) hold.

We assume that the elements of  $\mathbf{X}$  are i.i.d. standard Gaussian, and denote  $x_{ij}$  by  $z_{ij}$ . Moreover, we denote the corresponding  $\mathbf{B}_n$  as  $\mathbf{B}_g$ . Without loss of generality, we assume that  $\Sigma$  is diagonal with eigenvalues  $\sigma_1 \geq \dots \geq \sigma_p > 0$ . Using the bound of the case where  $\Sigma = \mathbf{I}$  shown in [Chen and Pan \[2012, Page 1418\]](#), we have

$$\begin{aligned} 0 \leq \mathbb{E} \text{tr}(\mathbf{B}_g^k) &= \frac{1}{(\sqrt{npb_p})^k} \sum \mathbb{E} \left\{ \left( \mathbf{Z}' \Sigma \mathbf{Z} \right)_{j_1 j_2} \left( \mathbf{Z}' \Sigma \mathbf{Z} \right)_{j_2 j_3} \cdots \left( \mathbf{Z}' \Sigma \mathbf{Z} \right)_{j_k j_1} \right\} \\ &\leq \frac{1}{(\sqrt{npb_p})^k} \sum \mathbb{E} \left| \left( z_{i_1 j_1} \sigma_{i_1} z_{i_1 j_2} \right) \left( z_{i_2 j_2} \sigma_{i_2} z_{i_2 j_3} \right) \cdots \left( z_{i_k j_k} \sigma_{i_k} z_{i_k j_1} \right) \right| \\ &= \left( \frac{\|\Sigma\|}{\sqrt{b_p}} \right)^k \cdot \underbrace{\frac{1}{(np)^{k/2}} \sum \mathbb{E} |z_{i_1 j_1} z_{i_1 j_2} z_{i_2 j_2} z_{i_2 j_3} \cdots z_{i_k j_k} z_{i_k j_1}|}_{\leq (2+\zeta)^k \text{ by } \text{Chen and Pan [2012, Page 1418]}} \leq \left\{ \frac{\|\Sigma\|(2+\zeta)}{\sqrt{b_p}} \right\}^k, \end{aligned} \quad (3)$$

where  $0 < \zeta < \varepsilon/\eta$ , and the summation is over all  $i_1, \dots, i_k \in [p]$  and  $j_1, \dots, j_k \in [n]$  such that  $j_1, \dots, j_k$  are distinct. Thus, we have

$$\mathbb{P}(\|\mathbf{B}_g\| \geq \eta + \varepsilon) \leq \frac{\mathbb{E} \|\mathbf{B}_g\|^k}{(\eta + \varepsilon)^k} \leq \frac{\mathbb{E} \text{tr}(\mathbf{B}_g^k)}{(\eta + \varepsilon)^k} \leq \left( \frac{2+\zeta}{\eta + \varepsilon} \cdot \frac{\|\Sigma\|}{\sqrt{b_p}} \right)^k = o(n^{-\ell}),$$

since  $k/\log n \rightarrow \infty$  and  $\frac{2+\zeta}{\eta + \varepsilon} \cdot \frac{\|\Sigma\|}{\sqrt{b_p}} < 1$ .

## 2.2 Step 2: Extension to non-Gaussian case

As shown in the last section, the key input is the upper bound of  $\mathbb{E} \text{tr}(\mathbf{B}_n^k)$  as in (3). We can apply the Lindeberg replacement trick to extend the result to the non-Gaussian case.

We define  $\mathbf{B}_{i,j}$  as the matrix obtained by replacing all entries  $z_{k\ell}$  in  $\mathbf{B}_g$  with  $x_{k\ell}$  for all  $(k, \ell)$  such that  $k < i$  or  $k = i$  and  $\ell \leq j$ . Then, we have  $\mathbf{B}_n = \mathbf{B}_{p,n}$  and  $\mathbf{B}_g = \mathbf{B}_{1,0}$ , and we can write

$$\mathbb{E} \text{tr}(\mathbf{B}_n^k) - \mathbb{E} \text{tr}(\mathbf{B}_g^k) = \sum_{i=1}^p \sum_{j=1}^n \left\{ \mathbb{E} \text{tr}(\mathbf{B}_{i,j}^k) - \mathbb{E} \text{tr}(\mathbf{B}_{i,j-1}^k) \right\}.$$

We claim that (will be proved later, see [Page 3](#)) for any  $i$  and  $j$ ,

$$\mathbb{E} \text{tr}(\mathbf{B}_{i,j}^k) - \mathbb{E} \text{tr}(\mathbf{B}_{i,j-1}^k) \lesssim \left\{ \frac{\|\Sigma\|(2+2\zeta)}{\sqrt{b_p}} \right\}^k. \quad (4)$$

Starting from  $\mathbf{B}_g = \mathbf{B}_{1,0}$ , after  $pn$  replacements we obtain  $\mathbf{B}_n = \mathbf{B}_{p,n}$ , and we have

$$\mathbb{E} \text{tr}(\mathbf{B}_n^k) \lesssim (pn) \left\{ \frac{\|\Sigma\|(2+2\zeta)}{\sqrt{b_p}} \right\}^k,$$

and therefore

$$\mathbb{P}(\|\mathbf{B}_n\| \geq \eta + \varepsilon) \leq \frac{\mathbb{E} \text{tr}(\mathbf{B}_n^k)}{(\eta + \varepsilon)^k} \lesssim (pn) \left( \frac{2+2\zeta}{\eta + \varepsilon} \cdot \frac{\|\Sigma\|}{\sqrt{b_p}} \right)^k = o(n^{-\ell}), \quad (5)$$

the last equality follows from (2),  $p = O(n^\alpha)$ , and the fact that  $\frac{2+2\zeta}{\eta + \varepsilon} \cdot \frac{\|\Sigma\|}{\sqrt{b_p}} < 1$ .

*Remark 2.* The additional assumption  $p = \Theta(n^t)$  is necessary to derive the last equality in (3). Let  $\gamma = \frac{2+2\zeta}{\eta+\varepsilon} \cdot \frac{\|\Sigma\|}{\sqrt{b_p}} < 1$ . Then, we have

$$\frac{\log(pn\gamma^k n^\ell)}{\log n} \lesssim \frac{k}{\log n} \log \gamma + (t + \ell + 1) \xrightarrow{(2)} -\infty \implies pn\gamma^k = o(n^{-\ell}).$$

Now, we prove (4). We only prove the case  $i = j = 1$  for simplicity, and the other cases can be proved similarly. Let  $\mathbf{Z}_{(1)} = \mathbf{Z} - \mathbf{z}_1 \mathbf{e}_1'$ , we have

$$\begin{aligned} \mathbf{B}_{1,0}^\circ &= \mathbf{B}_g^\circ = \{\mathbf{Z}_{(1)} + \mathbf{z}_1 \mathbf{e}_1'\}' \Sigma \{\mathbf{Z}_{(1)} + \mathbf{z}_1 \mathbf{e}_1'\} - \text{diag}(\mathbf{Z}' \Sigma \mathbf{Z}) \\ &= \mathbf{Z}_{(1)}' \Sigma \mathbf{z}_1 \mathbf{e}_1' + \mathbf{e}_1 \mathbf{z}_1' \Sigma \mathbf{Z}_{(1)} + \mathbf{Z}_{(1)}' \Sigma \mathbf{Z}_{(1)} + \mathbf{e}_1 \mathbf{z}_1' \Sigma \mathbf{z}_1 \mathbf{e}_1' - \text{diag}(\mathbf{Z}' \Sigma \mathbf{Z}) \\ &= z_{11} (\mathbf{Z}_{(1)}' \sigma_1 \mathbf{e}_1' + \mathbf{e}_1 \sigma_1' \mathbf{Z}_{(1)}) + \Delta_{1,1}, \end{aligned}$$

where  $\sigma_1$  is the first column of  $\Sigma$ , and  $\Delta_{1,1}$  is obtained by replacing  $z_{11}$  in  $\mathbf{B}_{1,0}^\circ$  with zero. Similarly, we have

$$\mathbf{B}_{1,1}^\circ = x_{11} (\mathbf{Z}_{(1)}' \sigma_1 \mathbf{e}_1' + \mathbf{e}_1 \sigma_1' \mathbf{Z}_{(1)}) + \Delta_{1,1}.$$

Using these notations, we have

$$\begin{aligned} & \left| \mathbb{E} \text{tr}(\mathbf{B}_{1,1}^k) - \mathbb{E} \text{tr}(\mathbf{B}_{1,0}^k) \right| \\ &= \frac{1}{(npb_p)^{k/2}} \left| \mathbb{E} \text{tr} \left\{ x_{11} (\mathbf{Z}_{(1)}' \sigma_1 \mathbf{e}_1' + \mathbf{e}_1 \sigma_1' \mathbf{Z}_{(1)}) + \Delta_{1,1} \right\}^k - \mathbb{E} \text{tr} \left\{ z_{11} (\mathbf{Z}_{(1)}' \sigma_1 \mathbf{e}_1' + \mathbf{e}_1 \sigma_1' \mathbf{Z}_{(1)}) + \Delta_{1,1} \right\}^k \right| \\ &\lesssim \frac{1}{(npb_p)^{k/2}} \left[ \mathbb{E} \sum_{\ell=1}^k \binom{k}{\ell} \left\{ 2|x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| \right\}^\ell \|\Delta_{1,1}\|^{k-\ell} + \mathbb{E} \sum_{\ell=1}^k \binom{k}{\ell} \left\{ 2|z_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| \right\}^\ell \|\Delta_{1,1}\|^{k-\ell} \right] \\ &\leq \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left( 2|x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| + \|\Delta_{1,1}\| \right)^k + \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left( 2|z_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| + \|\Delta_{1,1}\| \right)^k, \end{aligned}$$

where in the first inequality, we use the fact that  $\mathbf{Z}_{(1)}' \sigma_1 \mathbf{e}_1' + \mathbf{e}_1 \sigma_1' \mathbf{Z}_{(1)}$  is of rank 2, and  $\text{tr}(\mathbf{A}) \leq r(\mathbf{A}) \|\mathbf{A}\|$ . It suffices to estimate the first term on the RHS, since the second term have the same bound. We have

$$\begin{aligned} & \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left( 2|x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| + \|\Delta_{1,1}\| \right)^k \\ &\leq \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left( 2|x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| + \|\Delta_{1,1}\| \right)^k \mathbb{1} \left\{ \|\Delta_{1,1}\| \geq 2\delta^{-1} |x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| \right\} \\ &\quad + \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left( 2|x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| + \|\Delta_{1,1}\| \right)^k \mathbb{1} \left\{ \|\Delta_{1,1}\| < 2\delta^{-1} |x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| \right\} \\ &\leq \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left\{ (1 + \delta) \|\Delta_{1,1}\| \right\}^k + \frac{1}{(npb_p)^{k/2}} \mathbb{E} \left\{ 2(1 + \delta^{-1}) |x_{11}| \cdot \|\mathbf{Z}_{(1)}' \sigma_1\| \right\}^k \\ &=: \mathcal{I}_1 + \mathcal{I}_2. \end{aligned}$$

We claim that (will be proved later, see Page 5) for any even  $k$ ,

$$\mathbb{E} \text{tr}(\Delta_{1,1}^k) \leq \mathbb{E} \text{tr}(\mathbf{B}_{1,0}^k). \quad (6)$$

Using this fact, we have

$$\frac{\mathbb{E} \|\Delta_{1,1}\|^k}{(npb_p)^{k/2}} \leq \frac{\mathbb{E} \text{tr}(\Delta_{1,1}^k)}{(npb_p)^{k/2}} \leq \text{tr}(\mathbf{B}_{1,0}^k) \leq \left\{ \frac{\|\Sigma\|(2 + \zeta)}{\sqrt{b_p}} \right\}^k, \quad (7)$$

and thus

$$\mathcal{I}_1 \leq \left\{ \frac{\|\Sigma\|(1 + \delta)(2 + \zeta)}{\sqrt{b_p}} \right\}^k \leq \left\{ \frac{\|\Sigma\|(2 + 2\zeta)}{\sqrt{b_p}} \right\}^k.$$

Using multinomial expansion, we have

$$\begin{aligned}
\mathbb{E}\|\mathbf{Z}'_{(1)}\sigma_1\|^{2k} &= \mathbb{E}\left\{\sum_{j=2}^n (\mathbf{z}'_j\sigma_1)^2\right\}^k \\
&= \sum_{\ell=1}^k \sum_{2 \leq j_1 < \dots < j_\ell \leq n} \sum_{\substack{k_1 + \dots + k_\ell = k \\ k_1, \dots, k_\ell \geq 1}} \binom{k}{k_1, \dots, k_\ell} \mathbb{E}(\mathbf{z}'_{j_1}\sigma_1)^{2k_1} \dots \mathbb{E}(\mathbf{z}'_{j_\ell}\sigma_1)^{2k_\ell} \\
&\stackrel{(9)}{\leq} \sum_{\ell=1}^k n^\ell \ell^k \frac{\{k\delta(np)^{1/4}\|\Sigma\|\}^{2k} k^\ell}{\{\delta^2(np)^{1/2}\}^\ell} \\
&= \{k\delta(np)^{1/4}\|\Sigma\|\}^{2k} \sum_{\ell=1}^k \frac{\ell^k}{\{\delta^2(p/n)^{1/2}/k\}^\ell} \\
&\stackrel{(2),(10)}{\leq} \{k\delta(np)^{1/4}\|\Sigma\|\}^{2k} k \left( \frac{k}{\log\{\delta^2(p/n)^{1/2}/k\}} \right)^k,
\end{aligned} \tag{8}$$

where

- the first inequality follows from, for any  $d \leq 2k$ ,

$$\begin{aligned}
\mathbb{E}|\mathbf{z}'_2\sigma_1|^d &= \mathbb{E}\left|\sum_{i=1}^p \sigma_{i1}z_{i2}\right|^d \\
&= \sum_{\ell=1}^{d/2} \sum_{1 \leq i_1 < \dots < i_\ell \leq p} \sum_{\substack{d_1 + \dots + d_\ell = d \\ d_1, \dots, d_\ell \geq 2}} \binom{d}{d_1, \dots, d_\ell} \mathbb{E}|\sigma_{i_1 1}z_{i_1 2}|^{d_1} \dots \mathbb{E}|\sigma_{i_\ell 1}z_{i_\ell 2}|^{d_\ell} \\
&\leq \sum_{\ell=1}^{d/2} \sum_{1 \leq i_1 < \dots < i_\ell \leq p} |\sigma_{i_1 1}|^2 \dots |\sigma_{i_\ell 1}|^2 \sum_{\substack{d_1 + \dots + d_\ell = d \\ d_1, \dots, d_\ell \geq 2}} \binom{d}{d_1, \dots, d_\ell} \|\Sigma\|^{d-2\ell} \{\delta(np)^{1/4}\}^{d-2\ell} \\
&\leq \sum_{\ell=1}^{d/2} \|\sigma_1\|^{2\ell} \ell^d \|\Sigma\|^{d-2\ell} \{\delta(np)^{1/4}\}^{d-2\ell} \\
&\leq \frac{\{\delta(np)^{1/4}\|\Sigma\|k\}^d k}{\delta^2(np)^{1/2}}.
\end{aligned} \tag{9}$$

- the last inequality follows from the (2) and the inequality [Yin et al., 1988, Page 519]

$$a^{-t}t^b \leq \left(\frac{b}{\log a}\right)^b, \quad a > 1, b > 0, t \geq 1, \frac{b}{\log a} > 1. \tag{10}$$

The above estimates yield

$$\begin{aligned}
\mathcal{I}_2 &\leq \frac{1}{(npb_p)^{k/2}} \{2(1 + \delta^{-1})\}^k \cdot \{\delta(np)^{1/4}\}^k \cdot \left\{\frac{k}{2}\delta(np)^{1/4}\|\Sigma\|\right\}^k \frac{k}{2} \left(\frac{k/2}{\log\{2\delta^2(p/n)^{1/2}/k\}}\right)^{k/2} \\
&\leq \underbrace{\left(\frac{\|\Sigma\|}{\sqrt{b_p}}\right)^k}_{\rightarrow 1} \underbrace{\{(k/2)^{1/k}\}^k \left(\frac{2\delta^2 k^3}{\log\{2\delta^2(p/n)^{1/2}/k\}}\right)^{k/2}}_{< 2} \\
&\stackrel{(2)}{\leq} \left\{\frac{\|\Sigma\|(2 + \varepsilon_0)}{\sqrt{b_p}}\right\}^k,
\end{aligned} \tag{11}$$

for some small constant  $0 < \varepsilon_0 < 1$ . The above estimates complete the proof of (4) for the case  $(i, j) = (1, 1)$ .

Assume that (4) holds for  $(i, j) = (i_0, j_0)$ , and we consider the case  $(i, j) = (i_0, j_0 + 1)$ . Similar to (6) – (7), we have

$$\frac{\mathbb{E}\text{tr}(\Delta_{i_0, j_0+1}^k)}{(npb_p)^{k/2}} \leq \mathbb{E}\text{tr}(\mathbf{B}_{i_0, j_0}^k) \lesssim \left\{ \frac{\|\Sigma\|(2+2\zeta)}{\sqrt{b_p}} \right\}^k,$$

where the last inequality is proven in the previous case  $(i, j) = (i_0, j_0)$ . All the other estimates are similar to the proof of the case  $(i, j) = (1, 1)$ . Thus, we complete the proof of (4) by using the mathematical induction.

*Proof of (6).* Recall that  $\mathbf{B}_{1,0}^\circ = \Delta_{1,1} + z_{11}\mathbf{H}$ , where  $\mathbf{H} = \mathbf{Z}'_{(1)}\boldsymbol{\sigma}_1\mathbf{e}'_1 + \mathbf{e}_1\boldsymbol{\sigma}'_1\mathbf{Z}_{(1)}$  is a rank-2 symmetry matrix. For notation simplicity, we denote  $\mathbf{B}_{1,0}^\circ$  as  $\mathbf{B}(z_{11})$  and let  $\psi(z_{11}) = \text{tr}(\mathbf{B}(z_{11})^k)$ , where  $k$  is any even integer.

We claim that  $\psi(z_{11})$  is convex in  $z_{11}$  (will be proved later). Using the conditional Jensen's inequality, we have

$$\mathbb{E}\{\psi(z_{11}) \mid z_{ij}, (i, j) \neq (1, 1)\} \geq \psi(\mathbb{E}[z_{11}]) = \psi(0) = \text{tr}(\Delta_{1,1}^k),$$

that is,

$$\mathbb{E}\{\text{tr}(\mathbf{B}_{1,0}^{\circ k}) - \text{tr}(\Delta_{1,1}^k) \mid z_{ij}, (i, j) \neq (1, 1)\} \geq 0.$$

Taking the outer expectation, we obtain (6).

Finally, we prove the convexity of  $\psi(z_{11})$ . Using the non-commutative chain rule for matrix powers, we have  $\partial \mathbf{B}(z_{11})^k / \partial z_{11} = \sum_{j=0}^{k-1} \mathbf{B}(z_{11})^j \mathbf{H} \mathbf{B}(z_{11})^{k-1-j}$ , so that  $\psi'(z_{11}) = k \text{tr}(\mathbf{H} \mathbf{B}(z_{11})^{k-1})$  and

$$\psi''(z_{11}) = k \sum_{j=0}^{k-2} \text{tr}(\mathbf{H} \mathbf{B}(z_{11})^j \mathbf{H} \mathbf{B}(z_{11})^{k-2-j}).$$

Diagonalizing the symmetric matrix  $\mathbf{B}(z_{11}) = \mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^\top$  and writing  $\tilde{\mathbf{H}} = \mathbf{U}^\top \mathbf{H} \mathbf{U}$ , we obtain

$$\text{tr}(\mathbf{H} \mathbf{B}(z_{11})^j \mathbf{H} \mathbf{B}(z_{11})^{k-2-j}) = \text{tr}(\tilde{\mathbf{H}} \boldsymbol{\Lambda}^j \tilde{\mathbf{H}} \boldsymbol{\Lambda}^{k-2-j}) = \sum_{i=1}^n \sum_{\ell=1}^n \Lambda_{\ell\ell}^j \Lambda_{ii}^{k-2-j} \tilde{H}_{i\ell}^2,$$

and thus

$$\psi''(z_{11}) = k \sum_{i,\ell} \tilde{H}_{i\ell}^2 \sum_{j=0}^{k-2} \Lambda_{\ell\ell}^j \Lambda_{ii}^{k-2-j} = k \left( \sum_{i \neq \ell} \frac{\Lambda_{\ell\ell}^{k-1} - \Lambda_{ii}^{k-1}}{\Lambda_{\ell\ell} - \Lambda_{ii}} \tilde{H}_{i\ell}^2 + \sum_{i=1}^n \Lambda_{ii}^{k-2} \tilde{H}_{ii}^2 \right) \geq 0 \quad (k \text{ is even}).$$

Hence,  $\psi$  is convex in  $z_{11}$ , and we complete the proof of (6). □

## References

- Binbin Chen and Guangming Pan. Convergence of the largest eigenvalue of normalized sample covariance matrices when  $p$  and  $n$  both tend to infinity with their ratio converging to zero. *Bernoulli*, 18(4), 2012. doi: [10.3150/11-bej381](https://doi.org/10.3150/11-bej381). (cited on pages 1 and 2)
- Jiixin Qiu, Zeng Li, and Jianfeng Yao. Asymptotic normality for eigenvalue statistics of a general sample covariance matrix when  $p/n \rightarrow \infty$  and applications. *The Annals of Statistics*, 51(3):1427–1451, 2023. doi: [10.1214/23-AOS2300](https://doi.org/10.1214/23-AOS2300). (cited on page 1)
- Yongquan Yin, Zhidong Bai, and Pathak R Krishnaiah. On the limit of the largest eigenvalue of the large dimensional sample covariance matrix. *Probability Theory and Related Fields*, 78:509–521, 1988. doi: [10.1007/BF00353874](https://doi.org/10.1007/BF00353874). (cited on page 4)