#### 1B Markov Chains

Learn the theorems. Do the 2 example sheets. Weber's notes and cross-reference Norris' book.

**Definition 1.2** We say that  $(X_n)_{n\geq 0}$  is a **Markov chain** with initial distribution  $\lambda$  and transition matrix P if for all  $n\geq 0$  and  $i_0,\ldots,i_{n+1}\in I$ ,

1. 
$$P(X_0 = i_0) = \lambda_{i_0}$$

2. 
$$P(X_{n+1} = i_{n+1} | X_0 = i_0, \dots, X_n = i_n) = P(X_{n+1} = i_{n+1} | X_n = i_n) = p_{i_n i_{n+1}}$$

**Theorem 1.3**  $(X_n)_{n\geq 0}$  is  $Markov(\lambda, P)$  iff for all  $n\geq 0$  and  $i_0,\ldots,i_n\in I$ ,

$$P(X_0 = i_0, \dots, X_n = i_n) = \lambda_{i_0} p_{i_0 i_1} \dots p_{i_{n-1} i_n}$$
(1.1)

## Example Sheet 1 Michaelmas 2020 (Bauerschmidt)

 $= P(X_{m+n+1} = i_{n+1} | X_{m+n} = i_n) = p_{i_n i_{n+1}} = P(Z_{n+1} = i_{n+1} | Z_n = i_n)$ 

Q1 Let  $X = (X_n)_{n\geq 0}$  be a Markov chain. Show that, conditioned on  $X_m = i$ ,  $Z = (Z_n)_{n\geq 0}$  given by  $Z_n = X_{n+m}$  is a Markov chain with starting state i.

For each  $j \in I$ ,

$$P(Z_0 = j) = P(X_m = j) = \delta_{ij} = \begin{cases} 1 & \text{if } j = i \\ 0 & \text{else} \end{cases}$$

For each  $n \geq 0$  and  $i_0, \ldots, i_{n+1} \in I$ ,

$$\begin{split} &P(Z_{n+1}=i_{n+1}|Z_0=i_0,\ldots,Z_n=i_n)\\ &=P(X_{m+n+1}=i_{n+1}|X_m=i_0,\ldots,X_{m+n}=i_n)\\ &=\frac{P(X_m=i_0,\ldots,X_{m+n}=i_n,X_{m+n+1}=i_{n+1})}{P(X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_{m+n+1}=i_{n+1}|X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_{m+n+1}=i_{n+1}|X_{m+n}=i_n)P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=P(X_{m+n+1}=i_{n+1}|X_{m+n}=i_n)\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=P(X_{m+n+1}=i_{n+1}|X_{m+n}=i_n)\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=P(X_{m+n+1}=i_{n+1}|X_{m+n}=i_n)\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=j_0,\ldots,X_{m-1}=j_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=P(X_{m+n+1}=i_{n+1}|X_{m+n}=i_n)\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{k_0,\ldots,k_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}\\ &=\frac{\sum_{j_0,\ldots,j_{m-1}}P(X_0=k_0,\ldots,X_{m-1}=k_{m-1},X_m=i_0,\ldots,X_{m+n}=i_n)}{\sum_{j_0,\ldots$$

**Q2** Let  $X = (X_n)_{n \ge 0}$  be a sequence of independent random variables. Show that X is a Markov chain. Under what condition is this chain homogeneous?

Since  $X_0, X_1, \ldots, X_n$  are independent, for all  $i_0, \ldots, i_{n+1} \in I$ ,

$$P(X_{n+1} = i_{n+1} | X_0 = i_0, \dots, X_n = i_n) = \frac{P(X_0 = i_0, \dots, X_n = i_n, X_{n+1} = i_{n+1})}{P(X_0 = i_0, \dots, X_n = i_n)}$$

$$= \frac{P(X_0 = i_0) \dots P(X_n = i_n) P(X_{n+1} = i_{n+1})}{P(X_0 = i_0) \dots P(X_n = i_n)}$$

$$= \frac{P(X_n = i_n) P(X_{n+1} = i_{n+1})}{P(X_n = i_n)}$$

$$= \frac{P(X_n = i_n, X_{n+1} = i_{n+1})}{P(X_n = i_n)}$$

$$= P(X_{n+1} = i_{n+1} | X_n = i_n) = P(X_{n+1} = i_{n+1})$$

so X is a Markov chain. It is homogeneous iff the  $X_i$  are identically distributed.

Q3 Let  $X = (X_n)_{n\geq 0}$  be a sequence of fair coin tosses (with the two possible outcomes interpreted as 0 and 1) and set  $M_n = \max_{k\leq n} X_k$ . Show that  $(M_n)_{n\geq 0}$  is a Markov chain and find the transition probabilities.

$$P(M_{n+1} = 0 | M_0 = i_0, \dots, M_n = i_n) = \begin{cases} 1/2 & \text{if } i_0 = \dots = i_n = 0 \\ 0 & \text{else} \end{cases}$$

and

$$P(M_{n+1} = 0 | M_n = i_n) = \begin{cases} 1/2 & \text{if } i_n = 0 \\ 0 & \text{else} \end{cases}$$

but  $i_0 = \cdots = i_n = 0$  iff  $i_n = 0$  so the two (conditional) probabilities are the same.

$$P(M_{n+1} = 1 | M_0 = i_0, \dots, M_n = i_n) = \begin{cases} 1/2 & \text{if } i_0 = \dots = i_n = 0\\ 1 & \text{else} \end{cases}$$

and

$$P(M_{n+1} = 1 | M_n = i_n) = \begin{cases} 1/2 & \text{if } i_n = 0\\ 1 & \text{else} \end{cases}$$

but  $i_0 = \cdots = i_n = 0$  iff  $i_n = 0$  so the two (conditional) probabilities are the same.

In particular,  $p_{01} = p_{00} = 1/2$  and  $p_{11} = 1$  and  $p_{10} = 0$ .

#### Q4

#### 1B Statistics

Do the 3 example sheets. Weber's notes.

Statistics is a collection of procedures and principles for gaining and processing information in order to make decisions when faced with uncertainty. Two of its principal concerns are **parameter estimation** and **hypothesis testing**.

**Data** refers to a collection of numbers or other pieces of information to which meaning has been attached. The numbers 1.1, 3.0, 6.5 are not necessarily data. They become so when we are told that they are the muscle weight gains in kg of three athletes who have been trying a new diet.

In statistics, our data are modelled by a vector of random variables

$$X = (X_1, X_2, \dots, X_n)$$

where  $X_i$  takes values in  $\mathbb{Z}$  or  $\mathbb{R}$ .

A statistic T(x) is any function of the data. An estimator of a parameter  $\theta$  is a function T = T(X) which we use to estimate  $\theta$  from an observation of X. T is said to be **unbiased** if

$$\mathbb{E}(T) = \theta$$

The expectation above is taken over X. Once the actual data x is observed, t = T(x) is the **estimate** of  $\theta$  obtained via the estimator T.

Suppose that the random variable X has probability density function  $f(x|\theta)$ . Given the observed value x of X, the **likelihood** of  $\theta$  is defined by

$$lik(\theta) = f(x|\theta)$$

Thus we are considering the density as a function of  $\theta$ , for a fixed x. In the case of multiple observations, i.e. when  $x = (x_1, \ldots, x_n)$  is a vector of observed values of  $X_1, \ldots, X_n$ , we assume unless otherwise stated that  $X_1, \ldots, X_n$  are IID; in this case  $f(x_1, \ldots, x_n | \theta)$  is the product of the marginals,

$$lik(\theta) = f(x_1, \dots, x_n | \theta) = \prod_{i=1}^n f(x_i | \theta)$$

It makes intuitive sense to estimate  $\theta$  by whatever value gives the greatest likelihood to the observed data. Thus the **maximum likelihood estimate**  $\hat{\theta}(x)$  of  $\theta$  is defined as the value of  $\theta$  that maximises the likelihood. Then  $\hat{\theta}(X)$  is called the **maximum likelihood estimator (MLE)** of  $\theta$ .

Of course the maximum likelihood estimator need not exist, but in many examples it does. In practice, we usually find the MLE by maximising  $\log f(x|\theta)$ , which is known as the **loglikelihood**.

#### 1B Optimisation

Learn the simplex algorithm. Do the 2 example sheets. Weber's notes.

#### Part 2 Probability & Measure

The relationships between various modes of convergence: almost sure, in probability, in  $L^p$ . Strong law of large numbers and central limit theorem.

# Part 2 Applied Probability

Review statements and proofs of the basic theorems on queues and the starting point of the story for Poisson processes. Bauerschmidt's notes.

## Part 3 Advanced Probability

Would really like to get hold of an interesting almost current concrete research problem to give direction to the theory. Martingale convergence theorems and basic secret examples of stochastic integrals and Feynman-Kac.

#### Part 3 Stochastic Calculus

Would really like to get hold of an interesting almost current concrete research problem to give direction to the theory.

## **Number Theory**

Greg is taking this in the spring semester though I probably can't fit it into my schedule and the registration is already full last I checked anyways. My main interest in this is to find motivation for commutative algebra from a source other than algebraic geometry, and also to see what the algebraic number theorists have to say about solving diophantine equations, which seems to be a basic and interesting problem, as introduced by Shafarevich.

Not sure what books or notes to start with. Could look into the Cambridge notes and the standard textbook for Greg's course.

# Algebraic Geometry

Want to understand (moduli spaces of) J-holomorphic curves in symplectic geometry. This was invented by Gromov and apparently inspired by pre-existing theory of holomorphic curves in algebraic geometry. Sounds to me it is basically about counting how many lines or curves of a particular degree there are in an algebraic variety. Like "27 lines on a cubic hypersurface" kind of thing, which is of course the classical starting point of the mirror symmetry business.

So I should learn about (moduli spaces of) algebraic curves. And probably Riemann surfaces and compare and contrast. "Curves and their Jacobians" seems to be a thing. Riemann surfaces version is in the Rick Miranda book.

Prerequisite for reading about curves and their Jacobians is to review and strengthen general basic algebraic geometry e.g. from Shafarevich, Kempf, Mumford. Would like to also take the opportunity to learn about sheaves, sheaf cohomology, and schemes. Mumford's Red Book is probably the desired "endpoint", though I am not sure how it compares vs Hartshorne. Shafarevich is very nice and concrete but does not talk about sheaves or schemes. Kempf would be a useful stepping stone.

So probably review Shafarevich for concrete elementary examples, read Kempf for the sheaves, read Mumford for the schemes, then read about moduli spaces of curves and Jacobians of curves. I suddenly recall something about Hilbert polynomial or Hilbert scheme which I should also look into. Of course on the Riemann surfaces side there is also the thing about period integrals and Hodge theory that I have wanted to look into for the longest time ever.

# Symplectic Geometry

Want to understand framework and applications and computations of Morse theory and Floer homology, in particular J-holomorphic curves for the Floer homology and also more generally in symplectic geometry other than Floer homology. As always, Audin & Damian and the original papers of Floer, but also Abouzaid for the symplectic homology and cohomology and isomorphisms with string or loop homology for the noncompact manifolds.

Other basic concepts: Liouville manifolds, contact manifolds, Legendrian submanifolds, conical ends, Lagrangian torus fibrations over an affine base, almost complex structures which of course reminds one of Kähler manifolds and Calabi-Yau manifolds, the canonical and anticanonical bundles.

#### 3-Manifolds

The basic state of knowledge of 3-manifolds. Required for writing the paper with Nathan when the time comes again. Classification of curves on surfaces and the proof that the algorithm works via flip moves or arcslides.

## Differential Geometry

Origins of the subject leading to theories of Riemannian metrics, curvature, and connections. Spivak of course, but also PMH Wilson, which might remind me of some questions regarding Lie groups and hence topological and algebraic groups.

# Algebraic Topology

Fundamental class, characteristic classes, computations of homology and cohomology using simplicial and other theories, isomorphisms between these theories with a view towards computations. Spin manifolds.

#### General brainteasers and problem solving exercises

Literally to prepare for interviews and online assessments. Would like to solve and create my own collection. Start from quant interview guide books and 1A Numbers & Sets but also look into olympiad problems and Putnam and other competitions.

## Day 1, 27th December 2024, Friday

## $\mathbf{Q}\mathbf{1}$

If  $p \in Z(I(X))$  then for every polynomial  $f \in k[x_1, \ldots, x_n]$  such that f = 0 on X, have  $p \in Z(f)$ . So for every Zariski closed subset of  $\mathbb{A}^n$  containing X, say defined by (the vanishing of) polynomials  $f_1, \ldots, f_m$ , have  $p \in Z(f_1, \ldots, f_m)$ , i.e. p is in every Zariski closed subset of  $\mathbb{A}^n$  containing X.

Conversely, suppose  $p \in \mathbb{A}^n$  is contained in every Zariski closed subset of  $\mathbb{A}^n$  containing X. Then for any  $f \in I(X)$ , since f = 0 on X, have that X is contained in the Zariski closed subset Z(f). Hence  $p \in Z(f)$ . This shows that f(p) = 0 for every  $f \in I(X)$  so  $p \in Z(I(X))$ .

# $\mathbf{Q2}$

Let  $X \subseteq \mathbb{A}^n$  be an irreducible affine variety and  $U \subseteq X$  a non-empty open set. Let Y be a closed subset of  $\mathbb{A}^n$  containing U, say  $Y = Z(f_1, \ldots, f_m)$  for some polynomials  $f_1, \ldots, f_m$ . Suppose there is some  $p \in X$  that is not contained in Y, so  $f_i(p) \neq 0$  for some  $i \in \{1, \ldots, m\}$ , wlog say  $f_1(p) \neq 0$ . In particular,  $p \notin U$  since  $U \subseteq Y$ . Write  $X_1 = X - U$  and  $X_2 = X \cap Z(f_1)$ . Clearly these are both closed subsets of X, and since  $U \subseteq Z(f_1)$ , have that  $X = X_1 \cup X_2$ . But  $X_1$  is not the whole of X since U is nonempty, and U is not the whole of U is nonempty, and U is irreducibility of U. Thus it is impossible for there to be a closed subset U but with U but with U and U is dense in U. That is, every closed subset of U containing U must contain all of U, i.e. U is dense in U.

Suppose U is not irreducible, so we can write  $U = U_1 \cup U_2$  where  $U_1, U_2$  are proper closed subsets of U. Let  $X_1, X_2 \subseteq X$  be closed subsets of X s.t.  $U_1 = U \cap X_1$  and  $U_2 = U \cap X_2$ . Clearly  $X_1 \neq X$  since  $U_1 \neq U$ , and similarly  $X_2 \neq X$ . But  $U = U_1 \cup U_2 = (U \cap X_1) \cup (U \cap X_2) = U \cap (X_1 \cup X_2)$  so  $X_1 \cup X_2$  is a closed subset of X containing all of U. If  $X_1 \cup X_2 = X$ , then we would have obtained X as a union of two proper closed subsets, contradicting the irreducibility of X. But if  $X_1 \cup X_2$  is not all of X, then writing  $X_3 = X - U$  we definitely have that  $X = (X_1 \cup X_2) \cup X_3$  as a union of two proper closed subsets, again contradicting the irreducibility of X. Hence it is impossible to write U as a union of proper closed subsets, i.e. U is irreducible.

Suppose X is an irreducible affine variety with at least 2 distinct points p and q. If X is Hausdorff, then there are disjoint open subsets U and V of X containing p and q respectively. Write Y = X - U and Z = X - V. These are proper closed subsets of X since  $p \notin Y$  and  $q \notin Z$ . Furthermore, since U and V are disjoint, have  $V \subseteq Y$  and  $U \subseteq Z$  so  $Y \cup Z = X$ . This gives X as a union of proper closed subsets, contradicting the irreducibility of X. Hence it is impossible for X to have two distinct points p and q and be Hausdorff.

#### Q3

Let  $X \supseteq X_1 \supseteq X_2 \supseteq X_3 \supseteq \ldots$  be a descending chain of closed subsets in an affine variety  $X \subseteq \mathbb{A}^n$ . Then  $I(X_1) \subseteq I(X_2) \subseteq I(X_3) \subseteq \ldots$  is an increasing chain of ideals in  $k[x_1, \ldots, x_n]$ , which is a Noetherian ring, so there is some N s.t.  $I(X_n) = I(X_N)$  for all  $n \ge N$ . Since the  $X_i$  are closed subsets,  $Z(I(X_i)) = X_i$  by Q1, so  $X_n = Z(I(X_n)) = Z(I(X_N)) = X_N$  for all  $n \ge N$ , i.e. the descending chain of closed subsets is eventually constant.

# $\mathbf{Q4}$

The coordinate ring of Y is k[x,y]/(xy-1) but the coordinate ring of  $\mathbb{A}^1$  is k[z]. An isomorphism between Y and  $\mathbb{A}^1$  induces an isomorphism between their coordinate rings. In particular,  $x,y \in k[x,y]/(xy-1)$  will be sent under this isomorphism to polynomials  $p(z), q(z) \in k[z]$  such that p(z)q(z) = 1. But this can only happen if p(z) and q(z) are constants, i.e. if  $\mathbb{A}^1$  is mapped to a single point in Y by the isomorphism, which is impossible (because k is algebraically closed so both Y and  $\mathbb{A}^1$  have infinitely many points).

As argued above, any morphism  $\mathbb{A}^1 \to Y$  induces a pullback of the coordinate functions x and y on Y, sending them to polynomial functions p(z) and q(z) on X, which must satisfy p(z)q(z)=1, which is only possible if x and y are constants. So the only morphism  $\mathbb{A}^1 \to Y$  is the constant map.

Regarding morphisms  $Y \to \mathbb{A}^1$ , there are the two obvious projections onto the coordinate axes  $(x,y) \mapsto x$  and  $(x,y) \mapsto y$ . More generally, we can project onto any straight line through the origin,  $(x,y) \mapsto ax + by$  for some constants  $a, b \in k$ , and then compose with any morphism  $\mathbb{A}^1 \to \mathbb{A}^1$ , which is simply a polynomial function in one variable, to obtain  $(x,y) \mapsto p(ax + by)$  for any  $p(z) \in k[z]$ .

## Day 2, 29th December 2024, Sunday

#### 2.1.c Existence of Pseudo-Gradient Fields

Pseudo-gradient fields exist for all Morse functions on manifolds. This is, for example, a consequence of the existence of Riemannian metrics, and more exactly, of the existence of Riemannian metrics with a prescribed form on a given subset of the manifold. In any case, it is a simple consequence of the existence of partitions of unity.

**Proposition 2.1.5** The stable and unstable manifolds of the critical point a are submanifolds of V that are diffeomorphic to open disks. Moreover, we have

$$\dim W^u(a) = \operatorname{codim} W^s(a) = \operatorname{Ind}(a)$$

**Proposition 2.1.6** We suppose that the manifold V is compact. Let  $\gamma : \mathbb{R} \to V$  be a trajectory of the pseudo-gradient field X. Then there exist critical points c and d of f such that

$$\lim_{s \to -\infty} \gamma(s) = c \quad \text{ and } \quad \lim_{s \to +\infty} \gamma(s) = d$$

Theorem 2.2.5 (Smale Theorem) Let V be a manifold with boundary and let f be a Morse function on V with distinct critical values. We fix Morse charts in the neighbourhood of each critical point of f. Let  $\Omega$  be the union of these charts and let X be a pseudo-gradient field on V that is transversal to the boundary. Then there exists a pseudo-gradient field X' that is close to X (in the  $C^1$  sense), equals X on  $\Omega$  and for which we have

$$W_{X'}^s(a) \cap W_{X'}^u(b)$$

for all critical points a, b of f.

**Theorem 3.2.2** The space  $\overline{\mathcal{L}}(a,b)$  is compact.

**Theorem 3.2.7** If  $\operatorname{Ind}(a) = \operatorname{Ind}(b) + 2$  then  $\overline{\mathcal{L}}(a,b)$  is a compact manifold of dimension 1 with boundary.

**Proposition 3.2.8** Let V be a compact manifold, let  $f:V\to\mathbb{R}$  be a Morse function ad let X be a pseudo-gradient for f satisfying the Smale property. Let a,c and b be three critical points of indices k+1, k and k-1 respectively. Let  $\lambda_1\in\mathcal{L}(a,c)$  and  $\lambda_2\in\mathcal{L}(c,b)$ . There exists a continuous embedding  $\psi$  from an interval  $[0,\delta)$  onto a neighbourhood of  $(\lambda_1,\lambda_2)$  in  $\overline{\mathcal{L}}(a,b)$  that is differentiable on  $(0,\delta)$  and satisfies

$$\psi(0) = (\lambda_1, \lambda_2) \in \overline{\mathcal{L}}(a, b), \qquad \psi(s) \in \mathcal{L}(a, b) \quad \text{ for } \quad s \neq 0$$

Moreover, if  $(l_n)$  is a sequence in  $\mathcal{L}(a,b)$  that tends to  $(\lambda_1,\lambda_2)$ , then  $l_n$  is contained in the image of  $\psi$  for n sufficiently large.

**Lemma** Let k b an arbitrary field,  $f \in k[x, y]$  an irreducible polynomial, and  $g \in k[x, y]$  an arbitrary polynomial. If g is not divisible by f then the system of equations f(x, y) = g(x, y) = 0 has only a finite number of solutions.

# Day 3, 30th December 2024, Monday

# 2.3 $L^2$ bounded martingales

**Definition** Let

$$M^2 = \{X: \Omega \times [0,\infty) \to \mathbb{R}: \ X \text{ is a cadlag martingale with } \sup_{t \geq 0} \mathbb{E} X_t^2 < \infty\} / \sim M_c^2 = \{X \in M^2: X(\omega,\cdot) \text{ is continuous for every } \omega \in \Omega\} / \sim$$

where  $\sim$  means that indistinguishable processes are identified. Moreover, set

$$\|X\|_{M^2} = \left(\sup_t \mathbb{E}X_t^2\right)^{1/2} = \left(\mathbb{E}X_\infty^2\right)^{1/2}$$

**Proposition**  $M^2$  is a Hilbert space and  $M_c^2$  is a closed subspace.

## 2.4 Quadratic variation

**Definition** For a sequence of processes  $(X^n)$  and a process X, we say  $X^n \to X$  ucp (uniformly on compact sets in probability) if for all t > 0 and  $\epsilon > 0$ ,

$$\mathbb{P}(\sup_{s\in[0,t]}|X_s^n - X_s| > \epsilon) \to 0$$

as  $n \to \infty$ .

## Day 4, 31st December 2024, Tuesday

It is 11.30pm on 31st December 2024, Tuesday. The last day of 2024. The new year is approaching and I feel mostly fear for what the future brings. Fear and stress. The past year seems to have gone by quickly and unproductively. I have accomplished little in the past 4 to 5 years but somehow it feels as if I accomplished the least in the past year. Somehow it feels I will accomplish even less in the coming year, perhaps even degenerate and regress.

I have found it hard to focus in the past few days since coming back to Singapore. Many reasons for my depression and unproductivity have occurred to me. The first reason that came to mind was loneliness. Christmas day was one of the most depressing days in recent memory, almost on par with that worst day of the fall 2024 semester. I spent it alone, aimlessly wandering from the Starbucks at Clarke Quay Central to the tables at Guoco tower at Tanjong Pagar, to the Starbucks at Marina Bay Link Mall. Revisiting all the usual old places. These places used to bring joy to me but now they only bring sadness. I saw so many old lonely people eating alone, sitting alone, aimlessly. It feels as though I am destined to become one of them, or perhaps I am already one of them. Lonely and poor with no purpose in life.

But it is also perhaps that I have come back to Singaproe too often in the past year. I was only back here in July-August earlier this year, barely 4 months ago. I have accomplished nothing =in the meantime while I was away. It feels as if the city has moved on and left me behind. My peers from Raffles and Cambridge certainly have. I am once again studying at the library. It used to be such a joyful nostalgic thing, learning. But now it seems I am destined, doomed to do this for the rest of my life. Aimlessly, hopelessly studying into oblivion with nothing to show for it. Poor and lonely. Unaccomplished.

Studying math used to bring solace. It was pure, clean, escapism into a beautiful world where I could be myself. But not anymore. I find myself reading the same things again that I was reading 4 months ago, 1 year ago, 2 years ago. It feels like I am going in circles, or perhaps marching on the spot. Going nowhere. And I do have somewhere I need to be, somewhere I want to be. Somewhere else. There here and now do not spark joy. I long for a better life, a better place. But is that even possible at all now. It all feels so hopelessly impossibly out of reach. I am destined to fail. Forever.

Or perhaps it is just the same greed and lack of focus that has always plagued me. I think I have reached my limit. I am being torn apart. My answer to not knowing what I really want to do, all these years, has been to (try to) do everything. It is exhausting. I am exhausted.

It is the new year tomorrow. 1st of January. Just 15 minutes away now. The library won't be open. Sadly? Or perhaps thankfully. I think I will go cycling. I hope the weather will be good. I just want to run away from all of it. But where can I run to? I am stuck. Hopelessly stuck. And alone.

Youth is a wonderful thing. Too bad it is wasted on the young. I wish I could be a secondary school kid at Raffles again. Here I am at 30, turning 31, wishing I was 20 or 15, when one day I shall be 40 wishing I was 30.

 $(X_n)_{n\geq 0}$  is  $Markov(\lambda, P)$  if and only if for all  $n\geq 0$  and  $i_0,\ldots,i_n\in I$ ,

$$P(X_0 = i_0, \dots, X_n = i_n) = \lambda_{i_0} p_{i_0 i_1} \dots p_{i_{n-1} i_n}$$

Let  $\mathcal{A}$  be a  $\pi$ -system. Then any d-system containing  $\mathcal{A}$  contains also the  $\sigma$ -algebra generated by  $\mathcal{A}$ .

Let X be an integrable random variable and let  $\mathcal{G} \subseteq \mathcal{F}$  be a  $\sigma$ -algebra. Then there exists a random variable Y such that

- (a) Y is  $\mathcal{G}$ -measurable
- (b) Y is integrable and  $\mathbb{E}(X1_A) = \mathbb{E}(Y1_A)$  for all  $A \in \mathcal{G}$ .

Let X be a continuous local martingale with  $X_0 = 0$ . If X is also a finite variation process then  $X_t = 0 \ \forall t$  a.s.

# 1st January 2025, Wednesday

Today was better. Went cycling.

## 2nd January 2025, Thursday

## **ALTAM Chapter 2 Survival Models**

- 2.1 Summary
- 2.2 The future lifetime random variable
- 2.3 The force of mortality
- 2.4 Actuarial notation
- 2.5 Mean and standard deviation of  $T_x$
- 2.6 Curtate future lifetime

#### 15 regular exercises, 3 excel based exercises

Let (x) denote a life aged x, where  $x \ge 0$ . The death of (x) can occur at any age greater than x, and we model the future lifetime of (x) by a continuous random variable which we denote by  $T_x$ . This means that  $x + T_x$  represents the age-at-death random variable for (x). We write

$$F_x(t) = \mathbb{P}(T_x \le t)$$

for the lifetime distribution from age x. The survival function

$$S_x(t) = 1 - F_x(t) = \mathbb{P}(T_x > t)$$

is the probability that (x) survives for at least t years.

The fundamental assumption: the events  $\{T_x \leq t\}$  and  $\{T_0 \leq x + t\}$  are equivalent in the sense that

$$\mathbb{P}(T_x \le t) = \mathbb{P}(T_0 \le x + t | T_0 > x)$$

i.e. given that a newborn individual survives to age x, the probability that they die by age x + t is the same as the probability that an individual currently aged x dies within t years. Equivalently,

$$F_x(t) = \mathbb{P}(T_x \le t) = \mathbb{P}(T_0 \le x + t | T_0 > x) = \frac{\mathbb{P}(x < T_0 \le x + t)}{\mathbb{P}(T_0 > x)} = \frac{F_0(x + t) - F_0(x)}{1 - F_0(x)}$$

$$S_x(t) = 1 - F_x(t) = \frac{1 - F_0(x+t)}{1 - F_0(x)} = \frac{S_0(x+t)}{S_0(x)}$$

$$S_0(x+t) = S_0(x)S_x(t)$$

i.e. the probability of a newborn surviving to age x + t is the probability that they survive to x and the probability that a person aged x survives another t years. The events of a newborn surviving two disjoint contiguous time periods are independent. More generally,

$$S_0(x+t+u) = S_0(x+t)S_{x+t}(u) = S_0(x)S_x(t+u) \quad \Rightarrow \quad S_x(t+u) = \frac{S_0(x+t)}{S_0(x)}S_{x+t}(u) = S_x(t)S_{x+t}(u)$$

The events of a life of any age surviving two disjoint contiguous time periods are independent.

We denote the force of mortality at age x by  $\mu_x$  and define it as

$$\mu_{x} = \lim_{dx \to 0^{+}} \frac{\mathbb{P}(T_{x} \le x + dx | T_{0} > x)}{dx}$$

$$= \lim_{dx \to 0^{+}} \frac{\mathbb{P}(T_{x} \le dx)}{dx}$$

$$= \lim_{dx \to 0^{+}} \frac{\mathbb{P}(T_{x} \le dx) - \mathbb{P}(T_{x} \le 0)}{dx}$$

$$= \lim_{dx \to 0^{+}} \frac{F_{x}(dx) - F_{x}(0)}{dx}$$

$$= -\lim_{dx \to 0^{+}} \frac{S_{x}(dx) - S_{x}(0)}{dx}$$

$$= -\lim_{dx \to 0^{+}} \frac{S_{x}(dx) - 1}{dx}$$

$$= -\frac{1}{S_{0}(x)} \lim_{dx \to 0^{+}} \frac{S_{0}(x + dx) - S_{0}(x)}{dx} = -\frac{1}{S_{0}(x)} S'_{0}(x)$$

or equivalently since  $S_0(0) = 1$ ,

$$S_0(x) = e^{-\int_0^x \mu_y dy}, \qquad x \ge 0$$

More generally,

$$S_x(t) = \frac{S_0(x+t)}{S_0(x)} = e^{-\int_x^{x+t} \mu_y dy}$$

In any case, the survival function  $S_x(t)$  is kind of exponentially decreasing over time and the force of mortality  $\mu_x$  contains all of the information about the rate at which people die off at any age so the fraction of people at age x surviving for at least t years is (one over) the proportion beating the (exponential of) the total force of mortality over those years.

Two of the most useful mortality laws are Gompertz's law

$$\mu_x = Bc^x$$
 for some constants  $B > 0$  and  $c > 1$ 

and Makeham's law, which is a generalisation of Gompertz's law, given by

$$\mu_x = A + Bc^x$$
 for some constants  $A, B > 0$  and  $c > 1$ 

One of the earliest mortality laws proposed was De Moivre's law

$$\mu_x = \frac{1}{\omega - x}, \qquad 0 \le x < \omega$$

This is unrealistic since it implies that  $T_x$  is uniformly distributed on the interval  $(0, \omega - x)$ , which means that he probability that a life currently aged x dies between ages x + t and x + t + dt is the same for all t > 0 as long as  $x + t > \omega$ .

Another simple mortality law that is very unrealistic for modelling human mortality is the constant force of mortality assumption which states that  $\mu_x = \mu$  for all  $x \ge 0$ . Under this model,  $T_x$  has an exponential distribution.

Although the De Moivre and constant force of mortality models are not useful for overall mortality for humans, they may be used in other contexts, such as modelling the failure time of machine components. We also use these models in a limited sense, to model mortality between integer ages.

The actuarial notation for survival and mortality probabilities is

$$_{t}p_{x} = S_{x}(t), \qquad _{t}q_{x} = F_{x}(t), \qquad _{u}|_{t}q_{x} = S_{x}(u) - S_{x}(u+t)$$

the last being the probability that a life currently aged x survives for u years but then dies within the t years after that. We drop the subscript t if its value is 1 so  $p_x$  is the probability that (x) survives to at least age x + 1 and  $q_x$  is the probability that (x) dies before age x + 1. We call  $q_x$  the mortality rate at age x and  $u|_tq_x$  the deferred mortality probability.