

Exam I : next week

$$f(x) = ax^2 + bx + c = 0$$

$$x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}$$

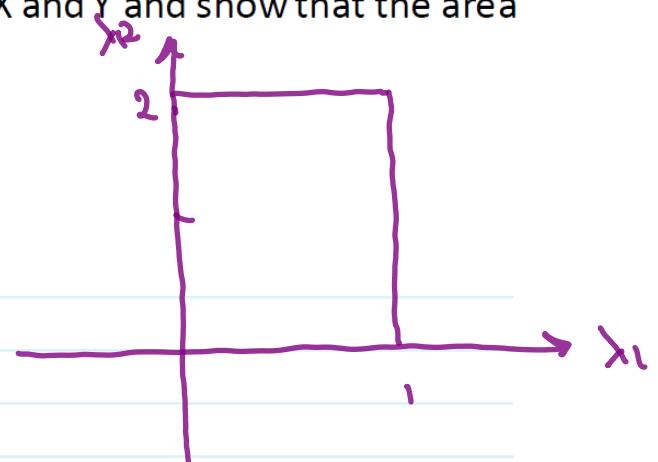
Final term project

3. Consider the following joint density function for the random variable X and Y uniformly distributed as:

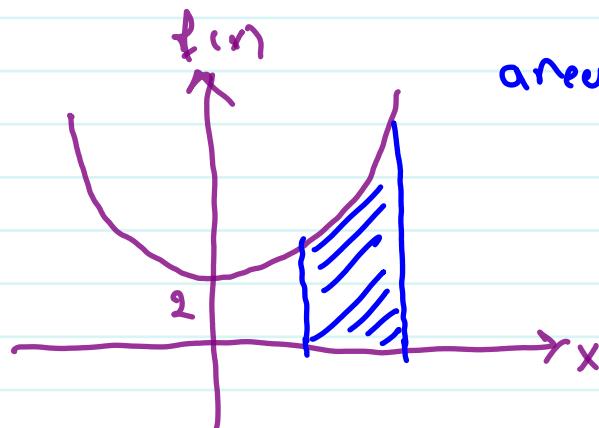
$$f_{X,Y}(\lambda_1, \lambda_2) = \begin{cases} c\lambda_1^2 + \frac{\lambda_1\lambda_2}{3} & ; 0 < \lambda_1 < 1 \text{ and } 0 < \lambda_2 < 2 \\ 0 & ; \text{ Else} \end{cases}$$

- a. Find the constant c.
- b. Find the marginal density $f_X(\lambda_1)$ and $f_Y(\lambda_2)$.
- c. Graph the marginal density function for random variable X and Y and show that the area under each curve is unity.
- d. Find $P(\lambda_1 + \lambda_2 > 1)$
- e. Are random variable X and Y independent?

$$f(x) = x^2 + 2$$



area $f(x) \quad 1 < x < 2$



$$\text{area} = \int_1^2 f(x) dx$$

independency check

$$f_{x,y}(\lambda_1, \lambda_2) \stackrel{?}{=} f_x(\lambda_1) \times f_y(\lambda_2)$$

$$c\lambda_1^2 + \frac{\lambda_1 \lambda_2}{3} \stackrel{?}{=}$$

$\forall \lambda_1, \lambda_2$

$$f_{x,y}(\lambda_1, \lambda_2) = \begin{cases} \frac{1}{\lambda_1 \lambda_2} & : \\ 0 & : \end{cases}$$

$$\frac{1}{\lambda_1 \lambda_2} = \frac{1}{\lambda_1} \cdot \frac{1}{\lambda_2}$$

x, y are independent

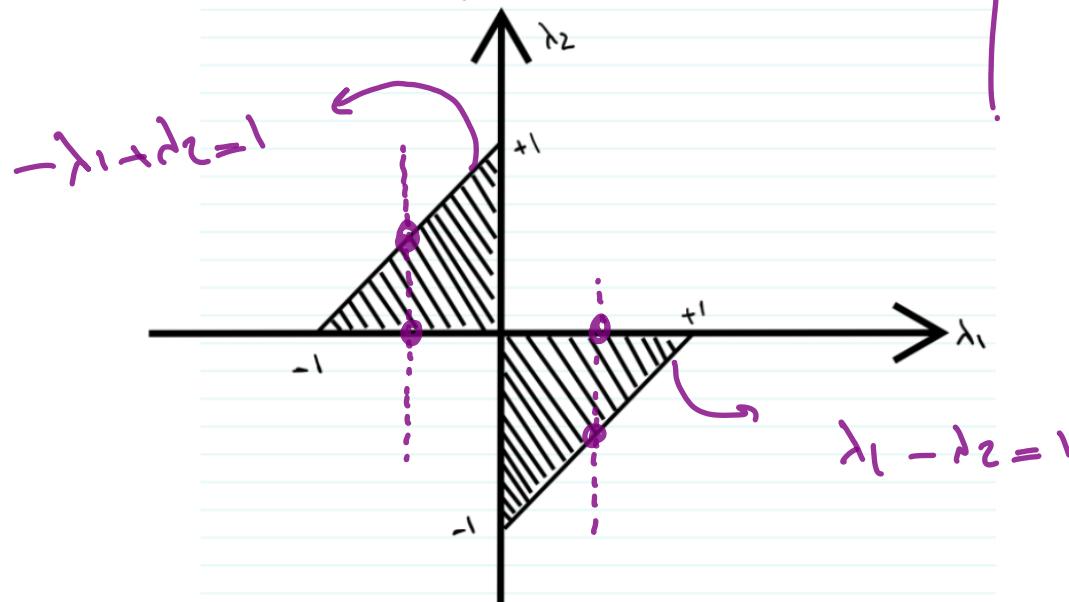
$$f_x(\lambda_1) = \frac{1}{\lambda_1}$$

$$f_y(\lambda_2) = \frac{1}{\lambda_2}$$

1. Consider the following joint density function for the random variable X and Y with the following distribution:

$$f_{X,Y}(\lambda_1, \lambda_2) = \begin{cases} c & ; -1 < \lambda_1 < 1 \text{ and } -1 < \lambda_2 < 1 \\ 0 & ; \text{ Else} \end{cases}$$

- a. Find the constant c.
- b. Find the marginal density $f_X(\lambda_1)$ and $f_Y(\lambda_2)$
- c. Graph the marginal density function for random variable X and Y and show that the area under each curve is unity.
- d. What is $E[X]$? what is $E[Y]$?
- e. Find and graph $f_{X|Y}(\lambda_1|\lambda_2)$. What is $f_{X|Y}(\lambda_1|\lambda_2)$ when $\lambda_2 = 0.5$.
- f. Find and graph $f_{Y|X}(\lambda_2|\lambda_1)$. What is $f_{Y|X}(\lambda_2|\lambda_1)$. when $\lambda_1 = 0.5$.
- g. Are random variable X and Y independent?



Alternative :

Volume = area * height

$$1 = \left(\frac{1}{2} + \frac{1}{2} \right) * c$$

$$c = 1$$

$$\int_{-\infty}^{\infty} \int f_{x,y}(\lambda_1, \lambda_2) d\lambda_2 d\lambda_1 = 1$$

$$\int_{-1}^0 \int_0^{1+\lambda_1} c d\lambda_2 d\lambda_1 + \int_0^1 \int_{\lambda_1-1}^0 c d\lambda_2 d\lambda_1 = 1$$

$$\int_{-1}^0 c(1+\lambda_1) d\lambda_1 + \int_0^1 c(1-\lambda_1) d\lambda_1 = 1$$

$$c\left(\lambda_1 + \frac{\lambda_1^2}{2}\right) \Big|_{-1}^0 + c\left(\lambda_1 - \frac{\lambda_1^2}{2}\right) \Big|_0^1 = 1$$

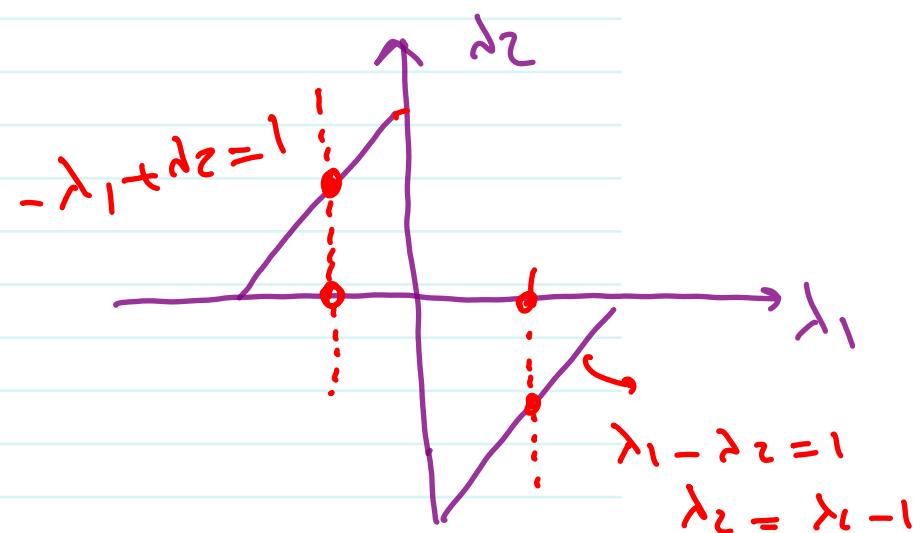
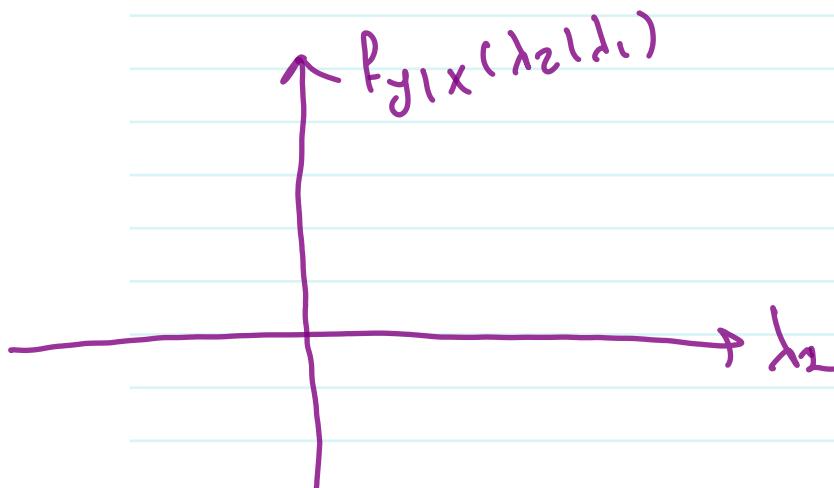
$$c\left(0 - \underbrace{(-1 + \frac{1}{2})}_{-1/2}\right) + c\left(\underbrace{1 - \frac{1}{2}}_{1/2}\right) = 1$$

$$\frac{c}{2} + \frac{c}{2} = 1$$

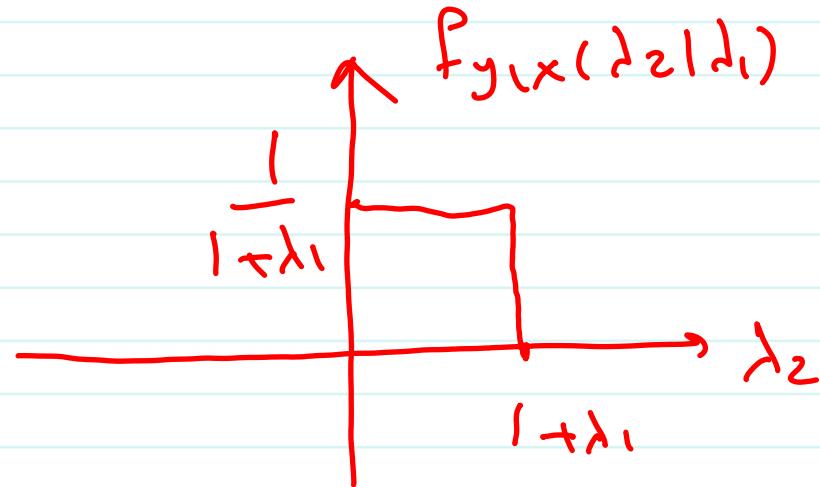
$$c = 1$$

$$f_x(\lambda_1) = \begin{cases} 1 + \lambda_1 & : -1 < \lambda_1 < 0 \\ 1 - \lambda_1 & : 0 < \lambda_1 < 1 \\ 0 & : \text{Else} \end{cases}$$

$$\underline{f_{y|x}(\lambda_2|\lambda_1)} = \frac{f_{xy}(\lambda_1, \lambda_2)}{f_x(\lambda_1)} = \begin{cases} \frac{1}{1+\lambda_1} & -1 < \lambda_1 < 0 \\ \frac{1}{1-\lambda_1} & 0 < \lambda_1 < 1 \\ 0 & : \text{Else} \end{cases}$$

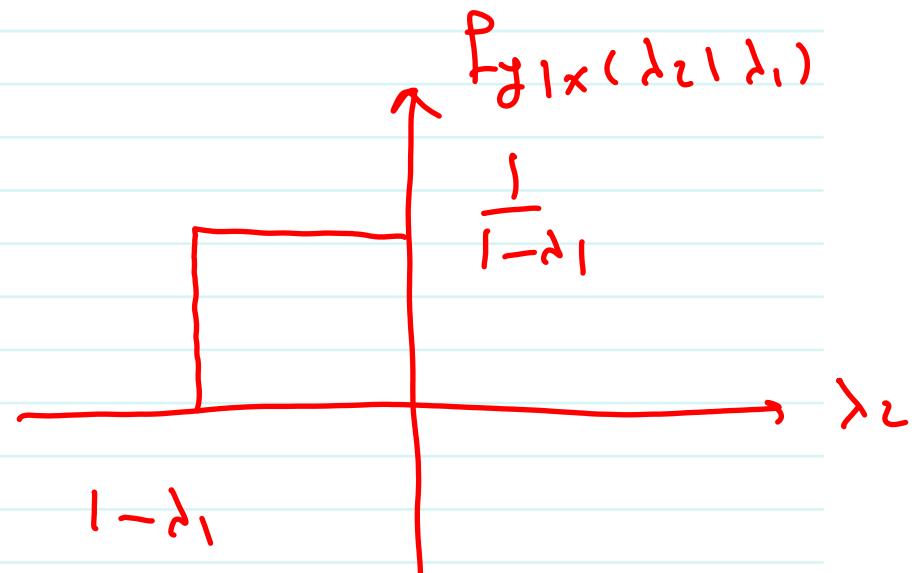


Case I : $[-1 < \lambda_1 < 0]$

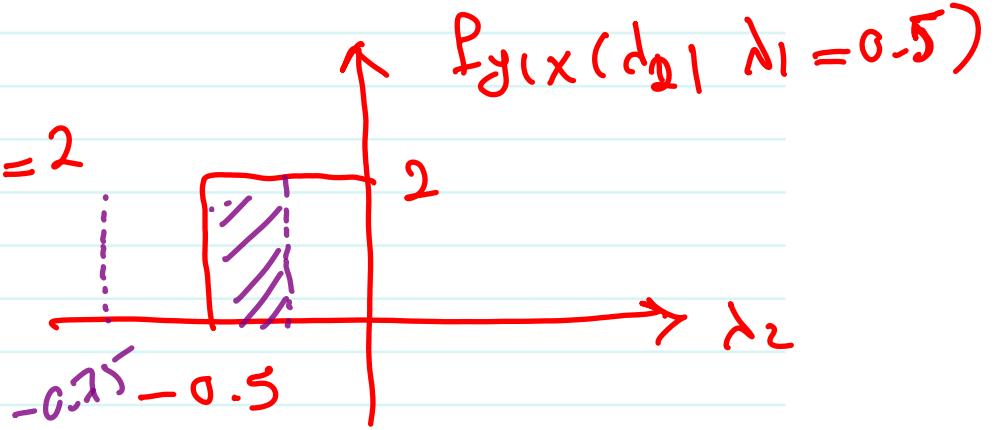


Case II: $[0 < \lambda_1 < 1]$

$$f_{Y|X}(\lambda_2 | \lambda_1) = \frac{1}{1 - \lambda_1}$$



$$f_{Y_1|X}(\lambda_2 | \lambda_1 = 0.5) = \frac{1}{\frac{1}{2}} = 2$$



$$\Pr(-0.5 < \lambda_2 < -0.75 \mid \lambda_1 = 0.5) = 0$$

$$\Pr(-0.5 < \lambda_2 < -0.25 \mid \lambda_1 = 0.5) = 50\%$$

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p + \epsilon$$

$$\left\{ \begin{array}{l} y_1 = \beta_0 + \beta_1 x_{11} + \dots + \beta_p x_{1p} + \epsilon_1 \\ y_2 = \beta_0 + \beta_1 x_{21} + \dots + \beta_p x_{2p} + \epsilon_2 \\ \vdots \\ y_n = \beta_0 + \beta_1 x_{n1} + \dots + \beta_p x_{np} + \epsilon_n \end{array} \right.$$

β_i are
unknown

$$\underline{Y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}_{n \times 1}$$

$$\underline{X} = \begin{bmatrix} 1 & x_{11} & \dots & x_{1p} \\ 1 & \vdots & \ddots & \vdots \\ \vdots & & & \vdots \\ 1 & x_{n1} & \dots & x_{np} \end{bmatrix}_{n \times (p+1)}$$

$$\underline{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}_{(p+1) \times 1}$$

$$\epsilon = \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix}_{n \times 1}$$

$n \times (p+1)$

$$Y = X\beta + \epsilon$$

$n \times 1$ $n \times (p+1)$ $(p+1) \times 1$
 \uparrow \nearrow \nearrow
 $\hat{Y} = X\hat{\beta}$ $+ \hat{\epsilon}$ $n \times 1$
 $\hat{Y} = X\hat{\beta}$

Find $\underline{\beta}$: $SSE = e_1^2 + \dots + e_n^2$

$$\begin{aligned}
 \underline{e} &= \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix} \\
 SSE &= \underline{e}^T \cdot \underline{e} \\
 &= (\underline{Y} - \hat{\underline{Y}})^T (\underline{Y} - \hat{\underline{Y}}) \\
 &= (\underline{Y} - X\beta)^T (\underline{Y} - X\beta) \\
 &= (\underline{Y}^T - \beta^T X^T)(\underline{Y} - X\beta)
 \end{aligned}$$

$(AB)^T = B^T \cdot A^T$

$$\underline{Y}^T X \beta$$

$$\text{SSE} = \underbrace{\mathbf{y}^T \cdot \mathbf{y}} - \underbrace{\mathbf{y}^T \mathbf{x} \beta^*}_{\text{Residual sum of squares}} - \underbrace{\beta^T \mathbf{x}^T \mathbf{y}}_{\text{Sum of products}} + \beta^T \mathbf{x}^T \mathbf{x} \beta$$

$$\frac{\partial \text{SSE}}{\partial \beta} = 0 - \mathbf{x}^T \mathbf{y} - \mathbf{x}^T \mathbf{y} + 2 \mathbf{x}^T \mathbf{x} \beta = 0$$

$$0 = -2 \mathbf{x}^T \mathbf{y} + 2 \mathbf{x}^T \mathbf{x} \beta$$

$$\mathbf{x}^T \mathbf{x} \beta = \mathbf{x}^T \mathbf{y}$$

$$\beta^*_{\text{LSE}} = (\mathbf{x}^T \mathbf{x})^{-1} \mathbf{x}^T \mathbf{y}$$

$(P_{+1} \times 1)$

Normal
Eq.
Least square
estimator.

$$f(\mathbf{x}) = h_1 x_1 + h_2 x_2 + \dots + h_n x_n \\ = \mathbf{h}^T \mathbf{x}$$

$$\mathbf{h} = \begin{bmatrix} h_1 \\ h_2 \\ \vdots \\ h_n \end{bmatrix} \quad \mathbf{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix}$$

$$\begin{aligned} \frac{\partial f}{\partial x_1} &= h_1 \\ \frac{\partial f}{\partial x_2} &= h_2 \\ &\vdots \\ \frac{\partial f}{\partial x_n} &= h_n \end{aligned}$$

$$\frac{\partial f}{\partial \underline{x}} = \underline{h}$$

$$f(x) = \underline{x}^T A \underline{x}$$

$$\frac{\partial f}{\partial x} = 2 A x$$

erval for this estimate.

$x \leftarrow$	Bill(\$)	y Tip(\$)	\hat{y}	c	e^2
	1	2	2.8	-0.8	0.64
	2	4	3.4	0.6	0.36
	3	5	4	1	1
	4	4	4.6	-0.6	0.36
	5	5	5.2	-0.2	0.04

$$SSE = 2.4$$

$$\hat{y} = B_0 + B_1 x$$

$$\hat{y} = 2.2 + 0.6 x$$

$$Y = X\beta \rightarrow \beta^* = (X^T X)^{-1} X^T Y$$

$$Y = \begin{bmatrix} 2 \\ 4 \\ 5 \\ 4 \\ 5 \end{bmatrix}$$

$$X = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \\ 1 & 5 \end{bmatrix} \quad \begin{array}{l} x_1^* \\ x_2^* \\ x_3^* \\ x_4^* \\ x_5^* \end{array}$$

$$\beta^* = \left(\begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 1 & 2 & 3 & 4 & 5 \end{bmatrix} \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 1 & 3 \\ 1 & 4 \\ 1 & 5 \end{bmatrix} \right)^{-1} \begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 1 & 2 & 3 & 4 & 5 \end{bmatrix} \begin{bmatrix} 2 \\ 4 \\ 5 \\ 4 \\ 5 \end{bmatrix}$$

$$\begin{aligned} \beta^* &= \left(\begin{bmatrix} 5 & 15 \\ 15 & 55 \end{bmatrix} \right)^{-1} \begin{bmatrix} 20 \\ 66 \end{bmatrix} \\ &= \frac{\begin{bmatrix} 55 & -15 \\ -15 & 5 \end{bmatrix} \begin{bmatrix} 20 \\ 66 \end{bmatrix}}{5 * 55 - 15 * 15} \\ &\quad \boxed{\begin{bmatrix} a & b \\ c & d \end{bmatrix}^{-1} = \frac{\begin{bmatrix} d & -b \\ -c & a \end{bmatrix}}{ad - bc}} \end{aligned}$$

$$\begin{aligned} &= \frac{\begin{bmatrix} 11 & -3 \\ -3 & 1 \end{bmatrix} \begin{bmatrix} 20 \\ 66 \end{bmatrix}}{55 - 45} = \frac{\begin{bmatrix} 11 & -3 \\ -3 & 1 \end{bmatrix} \begin{bmatrix} 2 \\ 6.6 \end{bmatrix}}{6.6} = \begin{bmatrix} 2.2 \\ 0.6 \end{bmatrix} \end{aligned}$$

$$\beta_0 = 2.2$$

$$\beta_1 = 0.6$$

$$\hat{y} = 2.2 + 0.6x$$

$$\hat{y}_i \pm 1.96 \sigma_e \sqrt{1 + x^* (X^T X)^{-1} x^*}$$

$$\sigma_e^2 = \frac{2.4}{n-p-1} = \frac{2.4}{5-1-1} = 0.6$$

$$\hat{y}_{1,1} = \sqrt{1 + [1 \ 1] \begin{bmatrix} 11 & -3 \\ -3 & 1 \end{bmatrix}^{-1} \begin{bmatrix} 1 \\ 1 \end{bmatrix}} = \sqrt{1 + \frac{6}{10}} = \sqrt{\frac{16}{10}}$$

$$\hat{y}_{2,2} = \sqrt{1 + [1 \ 2] \begin{bmatrix} 11 & -3 \\ -3 & 1 \end{bmatrix}^{-1} \begin{bmatrix} 1 \\ 2 \end{bmatrix}}$$

\hat{y}_1 interval

$$2.8 \pm 1.96 \sqrt{\frac{8}{10}} \sqrt{\frac{16}{10}} =$$

$$2.8 \pm 1.96 \sqrt{\frac{8+16}{100}} = 2.8 \pm \frac{1.96}{10} \frac{4\sqrt{2}}{5}$$

$$2.8 \pm 1.96 \frac{4\sqrt{2}}{5}$$

1.2

$$[2.6, 3.2]$$

poly fit

