

# Math 741 Assignment 16 (Hand-In)

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9.3.6.(H) solution: Let  $\sigma_1^2$  and  $\sigma_2^2$  are the variances of American League teams fans changes and National League teams fans changes, then the test can be formulated as following

$$H_0 : \sigma_1^2 = \sigma_2^2$$

$$H_1 : \sigma_1^2 \neq \sigma_2^2$$

let  $\alpha = 0.05$ . Enter the data into calculator and use 2-sample F Test, then

$$n_1 = 12, n_2 = 14$$

$$\bar{x}_1 = -12.833, \bar{x}_2 = -15.143$$

$$s_1 = 16.5685, s_2 = 19.9687$$

$$F_0 = \frac{s_1^2}{s_2^2} = 0.688444, P - value = 0.541978$$

since  $P - value = 0.541978 > \alpha = 0.05 \implies$  fail to reject  $H_0$ . Therefore, they have the same variance. As a result, we can use the pooled two-sample t test.

9.3.10. (H) solution: Let  $X_1, \dots, X_n$  and  $Y_1, \dots, Y_n$  be independent random samples from normal distributions with means  $\mu_1$  and  $\mu_2$  and standard deviations  $\sigma_1$  and  $\sigma_2$  and unbiased standard deviations estimator  $S_1$  and  $S_2$ . We will derive the  $\lambda$  of  $\sigma_X^2 = \sigma_Y^2$  against  $\sigma_X^2 \neq \sigma_Y^2$ , since for one-sided it is just to change the subscript notation from  $\alpha/2$  to  $\alpha$ .

If  $\sigma_X^2 = \sigma_Y^2 = \sigma^2$ , then

$$\hat{\sigma}^2 = \frac{1}{n+m} \left( \sum_{i=1}^n (x_i - \bar{x})^2 + \sum_{i=1}^m (y_i - \bar{y})^2 \right) = \frac{n}{n+m}$$

$$L(\Omega_0) = \prod_{i=1}^{n+m} \frac{1}{\sqrt{2\pi\hat{\sigma}}} e^{-\frac{1}{2\hat{\sigma}^2} \left( \sum_{i=1}^n (x_i - \bar{x})^2 + \sum_{i=1}^m (y_i - \bar{y})^2 \right)}$$

$$L(\Omega_0) = \left( \frac{1}{2\pi\hat{\sigma}^2} \right)^{(n+m)/2} e^{-\frac{n+m}{2}}$$

If  $\sigma_X^2 \neq \sigma_Y^2$ , then

$$\sigma_X^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2, \sigma_Y^2 = \frac{1}{m} \sum_{i=1}^m (y_i - \bar{y})^2$$

$$L(\Omega) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma_X}} e^{-\frac{1}{2\sigma_X^2} \sum_{i=1}^n (x_i - \bar{x})^2} \cdot \prod_{i=1}^m \frac{1}{\sqrt{2\pi\sigma_Y}} e^{-\frac{1}{2\sigma_Y^2} \sum_{i=1}^m (y_i - \bar{y})^2}$$

$$L(\Omega) = \left( \frac{1}{2\pi\sigma_X} \right)^{n/2} e^{-n/2} \cdot \left( \frac{1}{2\pi\sigma_Y} \right)^{m/2} e^{-m/2}$$

Moreover,

$$\lambda = \frac{L(\Omega_0)}{L(\Omega)} = \frac{\left( \frac{1}{2\pi\hat{\sigma}^2} \right)^{(n+m)/2} e^{-\frac{n+m}{2}}}{\left( \frac{1}{2\pi\sigma_X} \right)^{n/2} e^{-n/2} \cdot \left( \frac{1}{2\pi\sigma_Y} \right)^{m/2} e^{-m/2}} = \frac{\sigma_X^{n/2} \sigma_Y^{m/2}}{\hat{\sigma}^{(n+m)/2}}$$

Therefore,

$$\lambda = \frac{\left( \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \right)^{n/2} \left( \frac{1}{m} \sum_{i=1}^m (y_i - \bar{y})^2 \right)^{m/2}}{\left( \frac{1}{n+m} \right)^{(n+m)/2} \left( \sum_{i=1}^n (x_i - \bar{x})^2 + \sum_{i=1}^m (y_i - \bar{y})^2 \right)^{(n+m)/2}}$$

$$\lambda = \frac{(n+m)^{(n+m)/2}}{n^{n/2} m^{m/2}} \frac{\left( \sum_{i=1}^n (x_i - \bar{x})^2 \right)^{n/2} \left( \sum_{i=1}^m (y_i - \bar{y})^2 \right)^{m/2}}{\left( \sum_{i=1}^n (x_i - \bar{x})^2 + \sum_{i=1}^m (y_i - \bar{y})^2 \right)^{(n+m)/2}}$$

It is known that under  $H_0$ , the statistic follows:

$$F_0 = \frac{S_Y^2}{S_X^2}$$

with  $(m-1, n-1)$  df.

Then the GLRT for  $H_0 : \sigma_X^2 = \sigma_Y^2$  versus  $H_1 : \sigma_X^2 > \sigma_Y^2$  if  $\frac{S_Y^2}{S_X^2} \leq F_{\alpha, m-1, n-1}$ .