

PROBABILITY AND STATISTICS I

HOMEWORK III

Jianyu Dong 2019511017

March, 13 2021

1

Let $B_1 = \{\text{The original ball in the box is black}\}$

Let $B_2 = \{\text{The original ball in the box is white}\}$

Information tells us that

$$P(B_1) = p, \quad P(B_2) = 1 - p$$

Let $A = \{\text{The ball we picked is white}\}$

We can easily get that

$$P(A | B_1) = 0.5, \quad P(A | B_2) = 1$$

Putting this into Baye's theorem implies that

$$P(B_1 | A) = \frac{P(B_1)P(A | B_1)}{\sum_{j=1}^2 P(B_j)P(A | B_j)} = \frac{p}{2-p}$$

So now the probability that the original ball in the box is black is $\frac{p}{2-p}$

2

Let $B_1 = \{\text{The person have cancer}\}$

Let $B_2 = \{\text{The person doesn't have cancer}\}$

Let $A = \{\text{The test show positive}\}$

Information from the question tells that

$$P(B_1) = 0.04, \quad P(B_2) = 0.96, \quad P(A | B_1) = 0.79, \quad P(A | B_2) = 1 - 0.95 = 0.05$$

Putting this into Baye's theorem implies that

$$P(B_1 | A) = \frac{P(B_1)P(A | B_1)}{\sum_{j=1}^2 P(B_j)P(A | B_j)} = \frac{79}{199}$$

3

The discrete random variable X follows the binomial distribution with parameters 10 and $\frac{1}{2}$, so that the probability mass function is

$$p_X(x) = P(X = x) = \binom{10}{x} \left(\frac{1}{2}\right)^x \left(1 - \frac{1}{2}\right)^{10-x} = \binom{10}{x} \frac{1}{1024} \quad \text{for } x = 0, 1, \dots, 10$$

4

The discrete random variable X follows the geometric distribution with parameter $\frac{1}{2}$, so that the probability mass function is

$$p_X(x) = P(X = x) = \left(1 - \frac{1}{2}\right)^{x-1} \frac{1}{2} = \left(\frac{1}{2}\right)^x$$

5

Since X have the p.m.f $p_X(x) = \left(\frac{1}{2}\right)^{|x|}$, for $x = -1, -2, -3, \dots$

If $Y = X^4$, we have $y = 1, 2^4, 3^4, \dots$, so the p.m.f of Y is

$$p_Y(y) = \left(\frac{1}{2}\right)^{\sqrt[4]{y}}$$

6

6.1

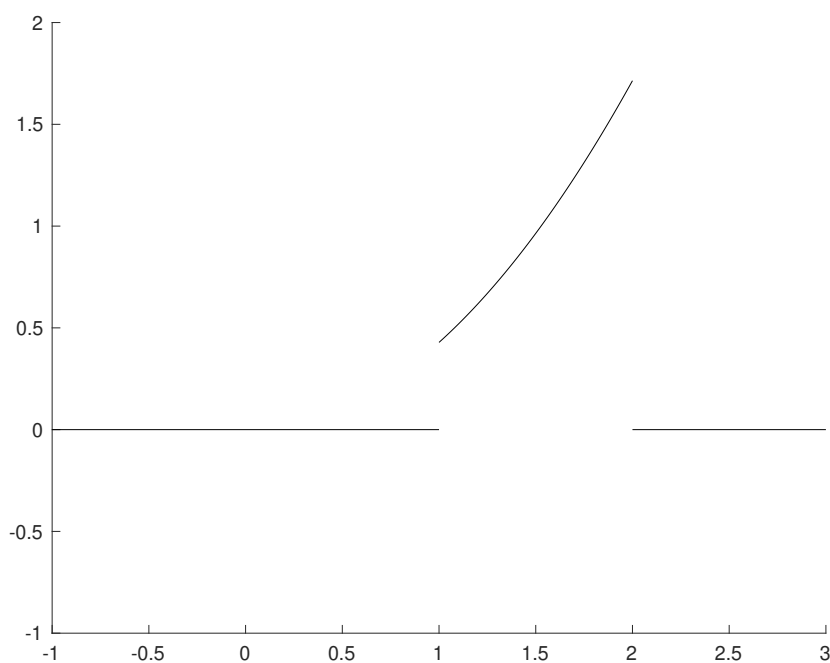
Since the probability density function of X is

$$f(x) = \begin{cases} cx^2, & \text{for } 1 \leq x \leq 2, \\ 0, & \text{otherwise.} \end{cases}$$

We have that

$$\int_{-\infty}^{+\infty} f(x) dx = \int_1^2 cx^2 dx = \frac{7}{3}c = 1$$

So we get the value of c is $\frac{3}{7}$



6.2

We could calculate that

$$P(X > \frac{3}{2}) = \int_{\frac{3}{2}}^{+\infty} f(x) dx = \int_{\frac{3}{2}}^2 \frac{3}{7}x^2 dx = \frac{37}{56}$$

7

Assume there exists a number c satisfying the function given is a probability density function.

If $c = 0$, $f(x) = 0$ for all x is a real number. So we have

$$\int_{-\infty}^{+\infty} f(x) dx = 0$$

Which shows $f(x)$ is not a probability density function, so $c \neq 0$.

For $c \neq 0$, we have

$$\int_{-\infty}^{+\infty} f(x) dx = \int_0^1 \frac{c}{x} dx = -c \lim_{x \rightarrow 0} \ln x$$

For all $c \neq 0$, we know that

$$\lim_{x \rightarrow 0} |c \ln x| = \infty$$

So the assumption is false. So there doesn't exist any number c for the given function is a p.d.f.

8

Problem information tells us that the p.d.f of X is

$$f(x) = \begin{cases} \frac{1}{x^2}, & \text{for } x > 1 \\ 0, & \text{otherwise.} \end{cases}$$

If $C_1 = \{x : 1 < x < 2\}$, $C_2 = \{x : 4 < x < 5\}$, we know that $C_1 \cup C_2 = \{x : 1 < x < 2 \text{ or } 4 < x < 5\}$ and $C_1 \cap C_2 = \emptyset$, so there are

$$P(C_1 \cup C_2) = \int_1^2 \frac{1}{x^2} dx + \int_4^5 \frac{1}{x^2} dx = \frac{11}{20}, \quad P(C_1 \cap C_2) = 0$$

9

Problem information tells us that the p.d.f of X is

$$f(x) = \begin{cases} \frac{x+2}{18} & \text{for } -2 < x < 4, \\ 0 & \text{otherwise.} \end{cases}$$

We can calculate that

$$P(|X| < 1) = P(-1 < X < 1) = \int_{-1}^1 f(x) dx = \int_{-1}^1 \frac{x+2}{18} dx = \frac{2}{9}$$

$$P(X^2 < 9) = P(-3 < X < 3) = \int_{-3}^3 f(x) dx = \int_{-2}^3 \frac{x+2}{18} dx = \frac{25}{36}$$

10

If $f(x) = ce^{-x^2}$ is a p.d.f for $x \in \mathbb{R}$, we have

$$\int_{-\infty}^{+\infty} f(x) dx = \int_{-\infty}^{+\infty} ce^{-x^2} dx = 1$$

So $c = \left(\int_{-\infty}^{+\infty} ce^{-x^2} dx \right)^{-1}$

In order to calculate above-mentioned integral, firstly we need to calculate

$$\int_{-R}^{+R} e^{-x^2} dx$$

Then we calculate

$$\left(\int_{-R}^R e^{-x^2} dx \right)^2 = \int_{-R}^R e^{-x^2} dx \int_{-R}^R e^{-y^2} dy = \iint_{\substack{-R \leq x \leq R \\ -R \leq y \leq R}} e^{-x^2-y^2} dx dy \quad (1)$$

Since $e^{-x^2-y^2} > 0$ and containment relationship of integration area, we know that

$$\iint_{x^2+y^2 \leq R^2} e^{-x^2-y^2} dx dy \leq \left(\int_{-R}^R e^{-x^2} dx \right)^2 \leq \iint_{x^2+y^2 \leq 2R^2} e^{-x^2-y^2} dx dy$$

With polar coordinate transformation, we get

$$\begin{aligned} \iint_{x^2+y^2 \leq R^2} e^{-x^2-y^2} dx dy &= \int_0^{2\pi} d\varphi \int_0^R e^{-r^2} r dr = \pi (1 - e^{-R^2}) \\ \iint_{x^2+y^2 \leq 2R^2} e^{-x^2-y^2} dx dy &= \int_0^{2\pi} d\varphi \int_0^{\sqrt{2}R} e^{-r^2} r dr = \pi (1 - e^{-2R^2}) \end{aligned}$$

So we get an inequality

$$\pi (1 - e^{-R^2}) \leq \left(\int_{-R}^R e^{-x^2} dx \right)^2 \leq \pi (1 - e^{-2R^2})$$

Let $R \rightarrow +\infty$, we get

$$\int_{-R}^R e^{-x^2} dx = \sqrt{\pi}$$

So if $f(x) = ce^{-x^2}$ is a p.d.f for $x \in \mathbb{R}$,

$$c = \frac{1}{\sqrt{\pi}}$$

11

Problem information tells us that the p.d.f of X is

$$f(x) = \begin{cases} 4x^3 & \text{for } 0 < x < 1, \\ 0 & \text{otherwise.} \end{cases}$$

So its c.d.f is

$$F(x) = \int_{-\infty}^x f(t) dt = \begin{cases} 0 & x < 0, \\ x^4 & 0 \leq x < 1, \\ 1 & x \geq 1. \end{cases}$$

Assume the 20th percentile is x_0 , since the c.d.f is continuous for $x \in \mathbb{R}$, so we have $F(x_0) = 0.2$. We could calculate that

$$x_0 = \frac{1}{\sqrt[4]{5}}$$

12

Problem information tells us that the c.d.f of X is

$$F(x) = \begin{cases} 0 & \text{for } x < 0, \\ 1 - e^{-x} - xe^{-x} & \text{for } x \geq 0. \end{cases}$$

So we could calculate the p.d.f is

$$f(x) = \frac{dF(x)}{dx} = \begin{cases} 0 & \text{for } x < 0, \\ xe^{-x} & \text{for } x \geq 0. \end{cases}$$

Assume the median is x_0 , so we have $F(x_0) = 1 - e^{-x_0} - x_0e^{-x_0} = 0.5$, then we could calculate that $x_0 \approx 1.678$

13

Since the probability set function of X is

$$P(C) = \int_C e^{-x} dx$$

where support is $C = (0, +\infty)$. If $C_k = \{x : 2 - \frac{1}{k} < x \leq 3\}$, $k = 1, 2, 3, \dots$

We can easily get that $\lim_{k \rightarrow \infty} C_k = \{x : 2 < x \leq 3\}$, so we could calculate that

$$P(\lim_{x \rightarrow \infty} C_k) = \int_2^3 e^{-x} dx = \frac{e-1}{e^3}$$

Then

$$P(C_k) = \int_{2-\frac{1}{k}}^3 e^{-x} dx = \sqrt[k]{e} e^{-2} - e^{-3}$$

We know that $\lim_{x \rightarrow \infty} \sqrt[x]{e} = 1$ So we have

$$\lim_{k \rightarrow \infty} P(C_k) = \lim_{k \rightarrow \infty} \sqrt[k]{e} e^{-2} - e^{-3} = \frac{e-1}{e^3}$$

Which shows that $\lim_{k \rightarrow \infty} P(C_k) = P(\lim_{x \rightarrow \infty} C_k)$

14

Since X follows the binomial distribution with parameters n and $p = 1/2$, we have

$$P(X = x) = \binom{n}{x} \left(\frac{1}{2}\right)^x \left(1 - \frac{1}{2}\right)^{n-x} = \binom{n}{x} \left(\frac{1}{2}\right)^n$$

When $k = 0$, we have

$$P(X \leq 0) = P(X = 0) = \frac{1}{2^n} \leq \frac{1}{2^n} \text{ for all sufficiently large } n$$

Assume when $k = t \geq 1$, there still is

$$P(X \leq t) \leq \frac{1}{2^n} \binom{n}{t} \frac{n - (t-1)}{n - (2t-1)} \text{ for all sufficiently large } n$$

Then when $k = t + 1$, we could calculate that

$$\begin{aligned} P(X \leq t+1) &= P(X \leq t) + P(X = t+1) \leq \frac{1}{2^n} \binom{n}{t} \frac{n - (t-1)}{n - (2t-1)} + \frac{1}{2^n} \binom{n}{t+1} \\ &= \frac{1}{2^n} \binom{n}{t+1} \left(1 + \frac{(t+1)(n-t+1)}{(n-t)(n-2t+1)}\right) \end{aligned}$$

To prove while $k = t + 1$, $P(k)$ still satisfy the inequality, we just need to prove that

$$1 + \frac{(t+1)(n-t+1)}{(n-t)(n-2t+1)} \leq \frac{n-t}{n-2t-1} = 1 + \frac{t+1}{n-2t-1}$$

So we need to prove

$$\frac{(t+1)(n-t+1)}{(n-t)(n-2t+1)} \leq \frac{t+1}{n-2t-1}$$

Since problem says $n \geq 2k$, we have that $n \geq 2(t+1)$, so that $n-t > 0, n-2t+1 > 0, n-2t-1 > 0, t+1 > 0$. Then we just need to prove that

$$(n-t)(n-2t+1) - (n-t+1)(n-2t-1) = (n^2 - 3nt + n + 2t^2 - t) - (n^2 - 3nt + 2t^2 - t - 1) = n + 1 > 0$$

It is obvious that $n + 1 > 0$, so we have proven that

$$1 + \frac{(t+1)(n-t+1)}{(n-t)(n-2t+1)} \leq \frac{n-t}{n-2t-1}$$

Which shows that

$$P(X \leq t+1) \leq 2^{-n} \binom{n}{t+1} \frac{n - ((t+1) - 1)}{n - (2(t+1) - 1)}$$

In summary, we have proven that for a fixed k and all $n \leq 2k$, there exists

$$P(X \leq k) \leq 2^{-n} \binom{n}{k} \frac{n - (k - 1)}{n - (2k - 1)}$$

a