

PROBABILITY AND STATISTICS I

HOMEWORK II

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Let $A_0 = \{ \text{A shopper chooses brand A on his first and second purchase and brand B on his third and fourth purchase} \}$

Let $A_1 = \{ \text{The shopper choose brand A on his first purchase} \}$

Let $A_2 = \{ \text{The shopper choose brand A on his second purchase} \}$

Let $A_3 = \{ \text{The shopper choose brand B on his third purchase} \}$

Let $A_4 = \{ \text{The shopper choose brand B on his fourth purchase} \}$

Then we have :

$$A_0 = A_1 \cap A_2 \cap A_3 \cap A_4$$

$$P(A_1) = \frac{1}{2}, P(A_2|A_1) = \frac{1}{3}, P(A_3|A_1 \cap A_2) = \frac{2}{3}, P(A_4|A_1 \cap A_2 \cap A_3) = \frac{1}{3}$$

$$P(A_0) = P(A_1) \times P(A_2|A_1) \times P(A_3|A_1 \cap A_2) \times P(A_4|A_1 \cap A_2 \cap A_3) = \frac{1}{27}$$

So the probability of both his first and second purchases will be brand A and both his third and fourth purchases will be brand B is $\frac{1}{27}$

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If after swapping n times the black ball is still in box A, there must be even times we pick the black ball and swap it with a white ball.

Let $C = \{ \text{After swapping n times, the black ball is still in box A} \}$

(1).If n is odd, there could be 0,2,4,6...(n-1) times we pick the black ball and swap it with a white ball. Then the probability of C is :

$$P(C) = \sum_{k=0}^{\frac{n-1}{2}} \binom{n}{2k} \left(\frac{1}{3}\right)^{2k} \times \left(\frac{2}{3}\right)^{n-2k}$$

(2).If n is even, there could be 0,2,4,6...n times we pick the black ball and swap it with a white ball. Then the probability of C is :

$$P(C) = \sum_{k=0}^{\frac{n}{2}} \binom{n}{2k} \left(\frac{1}{3}\right)^{2k} \times \left(\frac{2}{3}\right)^{n-2k}$$

We can calculate that:

$$P(C) = \frac{1 + 3^n}{2 \times 3^n}$$

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4

Let $A_0 = \{ \text{The first and third cards are spades, the second and fourth cards are red} \}$ Let $A_1 = \{ \text{The first card is spade.} \}$

Let $A_2 = \{ \text{The second card is red} \}$

Let $A_3 = \{ \text{The third card is spade} \}$

Let $A_4 = \{ \text{The fourth card is red} \}$

We have that:

$$P(A_1) = \frac{13}{52}, P(A_2 | A_1) = \frac{P(A_1 \cap A_2)}{P(A_1)} = \frac{26}{51},$$

$$P(A_3 | A_1 \cap A_2) = \frac{P(A_3 \cap A_1 \cap A_2)}{P(A_1 \cap A_2)} = \frac{12}{50}, \quad P(A_4 | A_1 \cap A_2 \cap A_3) = \frac{P(A_4 \cap A_1 \cap A_2 \cap A_3)}{P(A_1 \cap A_2 \cap A_3)} = \frac{25}{49}$$

Then we could get that:

$$P(A_0) = P(A_1 \cap A_2 \cap A_3 \cap A_4) = P(A_1) \times P(A_2 | A_1) \times P(A_3 | A_1 \cap A_2) \times P(A_4 | A_1 \cap A_2 \cap A_3) = \frac{13}{833}$$

Let $B_0 = \{ \text{The first and third cards are red, the second and fourth cards are spades} \}$ Let $B_1 = \{ \text{The first card is red.} \}$

Let $B_2 = \{ \text{The second card is spade} \}$

Let $B_3 = \{ \text{The third card is red} \}$

Let $B_4 = \{ \text{The fourth card is spade} \}$

We have that:

$$P(B_1) = \frac{26}{52}, \quad P(B_2 | B_1) = \frac{P(B_1 \cap B_2)}{P(B_1)} = \frac{13}{51},$$

$$P(B_3 | B_1 \cap B_2) = \frac{P(B_3 \cap B_1 \cap B_2)}{P(B_1 \cap B_2)} = \frac{25}{50}, \quad P(B_4 | B_1 \cap B_2 \cap B_3) = \frac{P(B_4 \cap B_1 \cap B_2 \cap B_3)}{P(B_1 \cap B_2 \cap B_3)} = \frac{12}{49}$$

Then we could get that:

$$P(B_0) = P(B_1 \cap B_2 \cap B_3 \cap B_4) = P(B_1) \times P(B_2 | B_1) \times P(B_3 | B_1 \cap B_2) \times P(B_4 | B_1 \cap B_2 \cap B_3) = \frac{13}{833}$$

Since A_0 and B_0 are disjoint, we can get the probability of the spades and red cards alternate is

$$P(A_0) + P(B_0) = \frac{26}{833}$$

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Since C_1 and C_2 are independent events, according to the definition of independent events we can get:

5.1

$$P(C_1 \cap C_2) = P(C_1) \times P(C_2) = 0.6 \times 0.3 = 0.18$$

5.2

$$P(C_1 \cup C_2) = P(C_1) + P(C_2) - P(C_1 \cap C_2) = 0.6 + 0.3 - 0.18 = 0.72$$

5.3

Since C_1 and C_2 are independent events and $P(C_2) + P(C_2^c) = 1$, we have C_1 and C_2^c are independent events and $P(C_2^c) = 0.7$.

Then we have:

$$P(C_1 \cap C_2^c) = P(C_1) \times P(C_2^c) = 0.6 \times 0.7 = 0.42$$

$$P(C_1 \cup C_2^c) = P(C_1) + P(C_2^c) - P(C_1 \cap C_2^c) = 0.6 + 0.7 - 0.42 = 0.88$$

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Let $A_1 = \{\text{Player A win in his first throw}\}$
 Let $A_2 = \{\text{Player A win in his second throw}\}$
 Let $A_3 = \{\text{Player A win in his third throw}\}$
 Let $B_1 = \{\text{Player B win in his first throw}\}$
 Let $B_2 = \{\text{Player B win in his second throw}\}$
 Let $B_3 = \{\text{Player B win in his third throw}\}$

Since player A and B play a sequence of independent games, according to the question stem, we can easily get that:

$$P(A \text{ win the game}) = P(A_1) + P(A_1^c \cap B_1^c \cap A_2) + P(A_1^c \cap B_1^c \cap A_2^c \cap B_2^c \cap A_3) = \frac{1}{6} + \frac{5}{6} \times \frac{4}{6} \times \frac{3}{6} + \frac{5}{6} \times \frac{4}{6} \times \frac{3}{6} \times \frac{2}{6} \times \frac{5}{6} = \frac{169}{324}$$

Since either A or B will win the game, we have:

$$P(B \text{ win the game}) = 1 - P(A \text{ win the game}) = 1 - \frac{169}{324} = \frac{155}{324}$$

In a word, the probability of player A win the game is $\frac{169}{324}$, the probability of player B win the game is $\frac{155}{324}$

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Let $A = \{\text{At least one six in four independent casts of a six-sided die}\}$, then $A^c = \{\text{There is no six in four independent casts of a six-sided die}\}$.

We can easily get that:

$$P(A^c) = \left(\frac{5}{6}\right)^4 = \frac{625}{1296}$$

$$P(A) = 1 - P(A^c) = \frac{671}{1296} \approx 0.5177$$

Let $B = \{\text{At least a pair of sixs in 24 independent casts of a pair of dice}\}$, then $B^c = \{\text{There is no pair of sixs in 24 independent casts of a pair of dice}\}$

We could caculate that:

$$P(B^c) = \left(\frac{35}{36}\right)^{24}$$

$$P(B) = 1 - P(B^c) = 1 - \left(\frac{35}{36}\right)^{24} \approx 0.4914$$

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Let $\Omega = \{\text{The person draw two cards from an ordinary deck of cards without replacement}\}$, and Ω is a simple sample space. We can get that:

$$|\Omega| = \binom{52}{2} = 1326$$

Let $A = \{\text{The person get two cards in same suit}\}$. We have:

$$A \subset \Omega \text{ and } |A| = \binom{4}{1} \times \binom{13}{2} = 312$$

Then, we have that:

$$P(A) = \frac{|A|}{|\Omega|} = \frac{4}{17}$$

If let the bet is fair, there must be that:

$$b \times P(A) = 1$$

So $b = 4.25$

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According to the definition of conditional probability, we have that:

$$P(B|A) = \frac{P(A \cap B)}{P(A)}$$

We know that $P(A \cap B) = P(A) + P(B) - P(A \cup B) = P(A) + 1 - P(A \cup B) - P(B^c) \geq P(A) - P(B^c)$, which uses $1 - P(A \cup B) \geq 0$

We get that:

$$P(B | A) = \frac{P(A \cap B)}{P(A)} \geq \frac{P(A) - P(B^c)}{P(A)} \geq 1 - \frac{P(B^c)}{P(A)}$$

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Since A and B are mutually exclusive, we get that $A \cap B^c = A$.

According to the definition of conditional probability, we have:

$$P(A | B^c) = \frac{P(A \cap B^c)}{P(B^c)} = \frac{P(A)}{1 - P(B)}$$

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If $A \subset B$ and $P(B) > 0$, we can get that $A \cap B = A$.

According to the multiplication rule, we have:

$$P(A) = P(A \cap B) = P(A | B) \times P(B) \leq P(A | B)$$

Using $P(B) \leq 1$, we prove that $P(A) \leq P(A | B)$

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According to the definition of conditional probability, we have:

$$P(A | B) = \frac{P(A \cap B)}{P(B)}$$

We know that $P(A \cap B) = P(A) + P(B) - P(A \cup B) \leq P(A) = p$. Which says that:

$$P(A | B) = \frac{P(A \cap B)}{P(B)} \leq \frac{p}{1 - \varepsilon}$$

We know that $P(A \cap B) = P(A) + P(B) - P(A \cup B) \geq P(A) + P(B) - 1 = p - \varepsilon$. Which says that:

$$P(A | B) = \frac{P(A \cap B)}{P(B)} \geq \frac{p - \varepsilon}{1 - \varepsilon}$$

In summary, there is:

$$\frac{p - \varepsilon}{1 - \varepsilon} \leq P(A | B) \leq \frac{p}{1 - \varepsilon}$$

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For any event B, if $P(A) = 0$, we can easily get that $P(A \cap B) = 0$, which says that $P(A \cap B) = P(A) \times P(B) = 0$. Then A is independent with any event B has been proven.

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According to the definition of conditional probability, we have:

$$P(A | C) = \frac{P(A \cap C)}{P(C)}$$

If $\{B_j\}_{j=1}^k$ is a partition of Ω , we have that:

$$P(A \cap C) = \sum_{j=1}^k P(A \cap C \cap B_j)$$

Then

$$P(A | C) = \frac{P(A \cap C)}{P(C)} = \sum_{j=1}^k \frac{P(B_j \cap C)}{P(C)} \times \frac{P(A \cap C \cap B_j)}{P(B_j \cap C)} = \sum_{j=1}^k P(B_j | C) \times P(A | B_j \cap C)$$

Which proves that if $\{B_j\}_{j=1}^k$ is a partition of Ω , for any A and C:

$$P(A | C) = \sum_{j=1}^k P(B_j | C) \times P(A | B_j \cap C)$$

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Since $0 < P(A) < 1$, $0 < P(B) < 1$, according to the definition of conditional probability, we have:

$$P(A | B) + P(A^c | B^c) = \frac{P(A \cap B)}{P(B)} + \frac{P(A^c \cap B^c)}{P(B^c)} = 1 \quad (*)$$

We know that

$$P(A^c \cap B^c) = P((A \cup B)^c) = 1 - P(A \cup B) = 1 - P(A) - P(B) + P(A \cap B)$$

$$P(B^c) = 1 - P(B)$$

Then (*) change into:

$$P(A \cap B) = P(A) \times P(B)$$

Which says that A and B are independent.

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If $P(A) > 0$ and $P(B) > 0$, A and B are independent, we have $P(A \cap B) = P(A) \times P(B) > 0$. While $P(\emptyset) = 0$, so $A \cap B \neq \emptyset$.