

(DO NOT WRITE YOUR ANSWER IN THIS AREA)

Seal line

**WARNING: MISBEHAVIOR AT EXAM TIME WILL LEAD TO SERIOUS CONSEQUENCE.**

# SCUT Final Exam

## Probability and Statistics Exam paper B (2021–2022-2)

- Notice:**
1. Make sure that you have filled the form on the left side of seal line.
  2. Write your answers on the exam paper.
  3. This is a close-book exam.
  4. The exam with full score of 100 points lasts 120 minutes.

Question No.	I	II	III	IV	V	VI	VII	VIII	IX	X	Sum
Score											

I. (10 points) We roll a die  $n$  times. Let  $A_{ij}$  be the event that the  $i$  th and  $j$  th rolls produce the same number. Show that the events  $\{A_{ij} : 1 \leq i < j \leq n\}$  are pairwise independent but not independent.

Score

Suppose  $i < j$  and  $m < n$ . If  $j < m$ , then  $A_{ij}$  and  $A_{mn}$  are determined by distinct independent rolls, and are therefore independent. For the case  $j = m$  we have that  $\mathbb{P}(A_{ij} \cap A_{jn}) = \mathbb{P}(i \text{ th, } j \text{ th, and } n \text{ th rolls show same number})$

$$= \sum_{r=1}^6 \frac{1}{6} \mathbb{P}(j \text{ th and } n \text{ th rolls both show } r \mid i \text{ th shows } r) = \frac{1}{36} = \mathbb{P}(A_{ij}) \mathbb{P}(A_{jn})$$

as required. However, if  $i \neq j \neq k$ ,

$$\mathbb{P}(A_{ij} \cap A_{jk} \cap A_{ik}) = \frac{1}{36} \neq \frac{1}{216} = \mathbb{P}(A_{ij}) \mathbb{P}(A_{jk}) \mathbb{P}(A_{ik})$$

II. (10 points) Two fair dice are rolled. Show that the event that their sum is 7 is independent of the score shown by the first die.

Score

$$\mathbb{P}(1 \text{ st shows } r \text{ and sum is } 7) = \frac{1}{36} = \frac{1}{6} \cdot \frac{1}{6} = \mathbb{P}(1 \text{ st shows } r) \mathbb{P}(\text{sum is } 7)$$

III. (10 points) Individuals **A** and **B** begin to play a sequence of chess games.

Score

Let  $S = \{\text{A wins a game}\}$ , and suppose that outcomes of successive games are independent with  $P(S) = p$  and  $P(F) = 1 - p$  (they never draw). They will play until one of them wins ten games. Let  $X =$  the number of games played (with possible values  $10, 11, \dots, 19$ ).

a. For  $x = 10, 11, \dots, 19$ , obtain an expression for  $p(x) = P(X = x)$

b. If a draw is possible, with  $p = P(S), q = P(F), 1 - p - q = P(\text{draw})$ , what is  $P(20 \leq X)$ ?

a.  $P(X = x) = P(\text{A wins in } x \text{ games}) + P(\text{B wins in } x \text{ games})$

$= P(9S' \text{ s in } 1^{\text{st}} x - 1 \cap S \text{ on the } x^{\text{th}}) + P(9F' \text{ s in } 1^{\text{st}} x - 1 \cap F \text{ on the } x^{\text{th}})$

$$= \binom{x-1}{9} p^9 (1-p)^{(x-1)-9} \cdot p + \binom{x-1}{9} (1-p)^9 p^{(x-1)-9} \cdot (1-p)$$

$$= \binom{x-1}{9} [p^{10} (1-p)^{x-10} + (1-p)^{10} p^{x-10}]$$

win the last

b. Possible values of  $X$  are now all positive integers  $\geq 10 : 10, 11, 12, \dots$

Similar to case a),

$P(X = x) = P(\text{A wins in } x \text{ games}) + P(\text{B wins in } x \text{ games})$

$= P(9S' \text{ s in } 1^{\text{st}} x - 1 \cap S \text{ on the } x^{\text{th}}) + P(9F' \text{ s in } 1^{\text{st}} x - 1 \cap F \text{ on the } x^{\text{th}})$

$$= \binom{x-1}{9} p^9 (1-p)^{(x-1)-9} \cdot p + \binom{x-1}{9} q^9 (1-q)^{(x-1)-9} \cdot q$$

$$= \binom{x-1}{9} [p^{10} (1-p)^{x-10} + q^{10} (1-q)^{x-10}] \text{ . Finally,}$$

$$P(X \geq 20) = 1 - P(X < 20) = \sum_{x=10}^{19} \binom{x-1}{9} [p^{10} (1-p)^{x-10} + q^{10} (1-q)^{x-10}]$$

IV. (10 points) A 12-in. bar that is clamped at both ends is to be subjected to an increasing amount of stress until it snaps. Let  $Y =$  the distance from the left end at which the break occurs. Suppose  $Y$  has pdf

Score

$$f(y) = \begin{cases} \left(\frac{1}{24}\right) y \left(1 - \frac{y}{12}\right) & 0 \leq y \leq 12 \\ 0 & \text{otherwise} \end{cases}$$

Compute the following:

- The cdf of  $Y$ .
- The expected length of the shorter segment  $\min(Y, 12 - Y)$  when the break occurs.

a. For  $0 \leq y \leq 12$ ,  $F(y) = \frac{1}{24} \int_0^y \left(u - \frac{u^2}{12}\right) du = \frac{1}{24} \left(\frac{u^2}{2} - \frac{u^3}{36}\right) \Big|_0^y = \frac{y^2}{48} - \frac{y^3}{864}$ .

b. The shorter segment has length equal to  $\min(Y, 12 - Y)$ , and

$$E[\min(Y, 12 - Y)] = \int_0^{12} \min(y, 12 - y) \cdot f(y) dy = \int_0^6 \min(y, 12 - y) \cdot f(y) dy + \int_6^{12} \min(y, 12 - y) \cdot f(y) dy = \int_0^6 y \cdot f(y) dy + \int_6^{12} (12 - y) \cdot f(y) dy = \frac{90}{24} = 3.75 \text{ inches.}$$

V. (10 points) Let  $X$  and  $Y$  be two continuous rv's with joint pdf

$$f(x, y) = \begin{cases} K(x + y^2) & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

Score

- What is the value of  $K$ ?
- What is the conditional probability that  $X$  is at most 0.5 given that  $Y = 0.5$ ?
- Compute the covariance between  $X$  and  $Y$ .

1.  $1 = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_0^1 \int_0^1 K(x + y^2) dx dy = K \int_0^1 \int_0^1 x dx dy + K \int_0^1 \int_0^1 y^2 dx dy = \frac{K}{2} + \frac{K}{3} = \frac{5K}{6} \Rightarrow K = \frac{6}{5}$

2. The marginal pdf of  $Y$  at 0.5 is  $f_Y(0.5) = \int_{-\infty}^{\infty} f(x, 0.5) dx = \frac{6}{5} \int_0^1 (x + \frac{1}{4}) dx = \frac{9}{10}$  and the conditional pdf of  $X$  given that  $Y = 0.5$  is

$$f_{X|Y}(x|0.5) = \frac{f(x, 0.5)}{f_Y(0.5)} = \begin{cases} \frac{4}{3} \left(x + \frac{1}{4}\right), & 0 \leq x \leq 1 \\ 0, & \text{otherwise.} \end{cases}$$

And thus  $P(X \leq 0.5 | Y = 0.5) = \int_0^{0.5} \frac{4}{3} \left(x + \frac{1}{4}\right) dx = \frac{1}{3}$ .

3.

$$E(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f(x, y) dx dy = \frac{7}{20};$$

$$E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dx dy = \frac{3}{5};$$

$$E(Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f(x, y) dx dy = \frac{3}{5};$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{7}{20} - \frac{9}{25} = -0.01.$$

VI. (10 points) Suppose that  $X$  and  $Y$  are two independent rv's, both of which has uniform distribution in  $(0, 2)$ .

Score

1. Determine the joint pdf of  $X$  and  $Y$ .
2. Compute the probability  $P(X + Y \leq 1)$
3. Compute the probability  $P(X \leq Y)$
4. Compute  $V(X - Y)$ .

1. Since  $X$  and  $Y$  are independent, their joint pdf is

$$f(x, y) = \begin{cases} f_X(x)f_Y(y) = \frac{1}{4}, & 0 \leq x \leq 2, 0 \leq y \leq 2; \\ 0, & \text{otherwise.} \end{cases}$$

$$2. P(X + Y \leq 1) = \int_0^1 \left[ \int_0^{1-x} \frac{1}{4} dy \right] dx = \frac{1}{4} \times \frac{1}{2} = \frac{1}{8}.$$

$$3. P(X \leq Y) = \int_0^2 \left[ \int_x^2 \frac{1}{4} dy \right] dx = \frac{1}{4} \times 2 = \frac{1}{2}.$$

$$4. V(X - Y) = V(X) + V(Y) = 2V(X) = 2 \times \frac{2^2}{12} = \frac{2}{3}.$$

VII. (10 points) Suppose the expected tensile strength of type-A steel is 100ksi and the standard deviation of tensile strength is 8ksi. For type-B steel, suppose the expected tensile strength is 95ksi and the standard deviation of tensile strength is 7ksi, respectively. Let  $\bar{X}$  = the sample average tensile strength of a random sample of 40 type-A specimens, and let  $\bar{Y}$  = the sample average tensile strength of a random sample of 35 type-B specimen. (Using the Central Limit Theorem to answer the following questions) ( $\Phi(1.65) = 0.95$ ,  $\Phi(2.89) = 0.998$ ,  $\Phi(1.96) = 0.975$ ) (a) What is the approximate distribution of  $\bar{X}$  ? of  $\bar{Y}$  ? (b) Calculate  $P(\bar{X} - \bar{Y} \geq 10)$ .

Score

a. According to the CLT,  $\bar{X}$  has approximately a normal distribution  $N\left(100, \frac{8^2}{40}\right)$ , i.e.  $N(100, 1.6)$ ,  $\bar{Y}$  has approximately a normal distribution  $N\left(95, \frac{7^2}{35}\right)$ , i.e.  $N(95, 1.4)$ .

b. According to the CLT,  $\bar{X} - \bar{Y}$  has approximately a normal distribution  $N(5, 3)$ .  $P(\bar{X} - \bar{Y} \geq 10) = 1 - P(\bar{X} - \bar{Y} < 10) \approx 1 - \Phi\left(\frac{10-5}{\sqrt{3}}\right) = 1 - \Phi\left(\frac{5}{\sqrt{3}}\right) = 1 - \Phi(2.886) = 1 - 0.998 = 0.002$ .

VIII. (10 points) Let  $X_1, X_2, \dots, X_n$  be a random sample of size  $n$  from the pdf

Score

$$f(x; \theta) = \begin{cases} \frac{x}{\theta} e^{-\frac{x^2}{2\theta}}, & x > 0 \\ 0, & \text{otherwise} \end{cases}$$

(a) Use the method of moments to find an estimator for  $\theta$ . (b) Find the maximum likelihood estimator for  $\theta$ .

(a)

$$\begin{aligned} E(x) &= \int_{-\infty}^{+\infty} x f(x, \theta) dx = \int_0^{+\infty} \frac{x^2}{\theta} e^{-\frac{x^2}{2\theta}} dx \\ &= \int_0^{+\infty} x e^{-\frac{x^2}{2\theta}} d\frac{x^2}{2\theta} = x \left( -e^{-\frac{x^2}{2\theta}} \right) \Big|_0^{+\infty} + \int_0^{+\infty} e^{-\frac{x^2}{2\theta}} dx \\ &= 0 + \frac{1}{2} \int_{-\infty}^{+\infty} e^{-\frac{x^2}{2\theta}} dx = \sqrt{\frac{\theta \cdot \pi}{2}} \\ \therefore \theta &= \frac{2E^2(x)}{\pi} \end{aligned}$$

The moment estimator is  $\hat{\theta}_M = \frac{2}{\pi} \bar{X}^2$  (b) The likelihood function is

$$f(x_1, \dots, x_n, \theta) = \prod_{i=1}^n \frac{x_i}{\theta} \cdot e^{-\sum_{i=1}^n \frac{x_i^2}{2\theta}} \quad (x_i > 0, i = 1, \dots, n)$$

The  $\ln(\text{likelihood})$  is

$$\begin{aligned} \ln f(x_1, \dots, x_n; \theta) &= \sum_{i=1}^n \ln \left( \frac{x_i}{\theta} \right) - \sum_{i=1}^n \frac{x_i^2}{2\theta} \\ &= \sum_{i=1}^n \ln x_i - n \ln \theta - \sum_{i=1}^n \frac{x_i^2}{2\theta} \\ \frac{d \ln f(x_1, \dots, x_n, \theta)}{d\theta} &= \frac{-n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n \frac{x_i^2}{2} = 0 \\ \Rightarrow \theta &= \frac{1}{2n} \sum_{i=1}^n x_i^2 \end{aligned}$$

So the maximum likelihood estimator of  $\theta$  is

$$\hat{\theta}_l = \frac{1}{2n} \sum_{i=1}^n X_i^2.$$

IX. (10 points) It is reported that for a sample of 49 kitchens with gas cooking appliances monitored during a one-week period, the sample mean CO<sub>2</sub> level (ppm) was 654. Suppose that the population of CO<sub>2</sub> level of all homes is normal.

Score

- (a) Calculate a 95% confidence interval for true average CO<sub>2</sub> level with the sample standard deviation  $s = 168$ . ( $t_{0.05,48} = 1.68$ ,  $t_{0.025,48} = 2.0$ .)
- (b) Suppose that  $\sigma = 175$ . What sample size would be necessary to obtain an interval width of at most 50ppm for a confidence level of 95%? ( $z_{0.05} = 1.65$ ,  $z_{0.025} = 1.96$ .)

(a) With  $n = 49$ ,  $\bar{x} = 654$  and  $t_{\alpha/2, n-1} = t_{0.025, 48} = 2.0$ , the 95% confidence interval for  $\mu$  is

$$\bar{x} \pm t_{\alpha/2, n-1} \frac{s}{\sqrt{n}} = 654 \pm 2.0 \frac{168}{7} = 654 \pm 48 = (606, 702).$$

(b) With  $\sigma = 175$  and  $\mu \in \bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$ , the width of CI is  $w = 2z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$ . So

$$n = \left( \frac{2z_{\alpha/2}\sigma}{w} \right)^2 = \left( \frac{2(1.96)(175)}{50} \right)^2 = (13.72)^2 = 188.24,$$

which rounds up to 189.

X. (10 points) The desired percentage of SiO<sub>2</sub> in a certain type of aluminous cement is 5.5. In a test 16 independently obtained samples are analyzed. Suppose that the percentage of SiO<sub>2</sub> is normally distributed with  $\sigma = 0.3$  and that  $\bar{x} = 5.25$ .

Score

- (a) Does this indicate conclusively that the true average percentage less than 5.5? Consider a significance level of  $\alpha = 0.01$ . ( $z_{0.01} = 2.33$ ,  $z_{0.005} = 2.58$ .)
- (b) If the true average percentage is  $\mu = 5.3$  and a level  $\alpha = 0.01$  test based on  $n = 16$  is used, what is the probability of rejecting  $H_0$ ? ( $\Phi(0.34) = 0.63$ ,  $\Phi(0.64) = 0.74$ .)
- (c) What value of  $n$  is required to satisfy  $\alpha = 0.01$  and  $\beta(5.3) = 0.01$ ? Note that  $\beta(5.3)$  is the probability of making a type II error when  $\mu = 5.3$ .

The hypotheses are  $H_0 : \mu = 5.5$  vs  $H_a : \mu < 5.5$ . The sample mean is  $\bar{x} = 5.25$ .

(a)  $z = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} = \frac{5.25 - 5.5}{0.3/\sqrt{16}} = -3.33 \leq -z_{\alpha} = -z_{0.01} = -2.33$ . Reject  $H_0$ .

(b) The probability of making a type II error when  $\mu = 5.3$  is

$$\beta(5.3) = 1 - \Phi \left( -z_{\alpha} + \frac{\mu_0 - \mu'}{\sigma/\sqrt{n}} \right) = 1 - \Phi \left( -2.33 + \frac{5.5 - 5.3}{0.3/\sqrt{16}} \right) = 1 - \Phi(0.34).$$

So the probability of rejecting  $H_0$  is  $1 - \beta(5.3) = \Phi(0.34) = 0.63$ .

(c) Since  $z_{0.01} = 2.33$ , the requirement that the level 0.01 test also have  $\beta(5.3) = 0.01$  necessitates

$$1 - \Phi \left( -2.33 + \frac{5.5 - 5.3}{0.3/\sqrt{n}} \right) = 0.01, \quad -2.33 + \frac{5.5 - 5.3}{0.3/\sqrt{n}} = 2.33, \quad n = \left[ \frac{0.3(2.33 + 2.33)}{5.5 - 5.3} \right]^2 = 48.9.$$

So  $n = 49$  should be used.