# The three paths of hierarchical modeling

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# Agenda

· Hierarchical GLM Notation

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- $\cdot$  Likelihood and model fit activity

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- · Hierarchical GLM Notation
- $\boldsymbol{\cdot}$  Likelihood and model fit activity
- · Radon!

# clustered data

A re-introduction to Generalized

Linear Models (GLMs) for

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- $\cdot x_i \in [0,1]$  is an indicator of exposure.
- ·  $\alpha$  is expected outcome when  $x_i$  = 0
- $\cdot$   $\epsilon_i$  are independently and identically distributed (i.i.d.) errors

# Independent errors

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In plain-ish English:

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Another way of writing it:

$$\cdot y_i \sim N(\alpha + \beta x_i, \sigma^2)$$

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### Three Approaches to Modeling Clustered Data



Which door will you choose?

# Door #1: Ignore clustering and fit a normal GLM

- · Pool data across all units, i.e. ignore clustering.
- · i.e. fit model  $y_{ij} = \alpha + \beta x_i + \epsilon_i$

Is this a good idea? Why or why not?

### NO!



Complete pooling ignores potential sources of *observed* and *unobserved*. unit-level confounding.

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# Pooling clustered data violates assumption of independent errors

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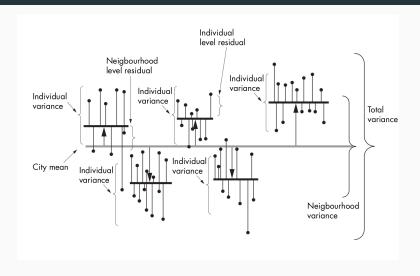
A pooled model:

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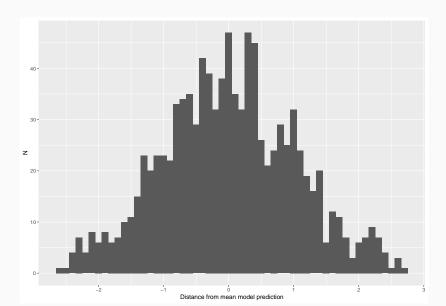
•  $y_i$  is a combination of systematic variation  $(\alpha + \beta x)$  and uncorrelated random noise  $(\epsilon_i)$  where:

$$i.i.d. \ \epsilon \sim Normal(0, \sigma^2)$$
 (2)

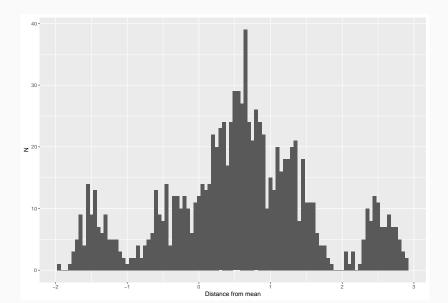
# Clustering may result in correlation between average differences from mean



### Your residuals should look like this



# When you ignore clustering you may see something like:



### Door #2: Fit a different model to each cluster

Fit  $unpooled \mod 1$  to each unit (j), assuming outcomes in each unit are independent:

$$\cdot \ y_{ij} = \alpha_j + \beta_j x_i + \epsilon_{ij}$$

$$\cdot \ \epsilon_{ij} \sim N(0,\sigma_j^2)$$

# More danger!



Totally unpooled models run the risk of overfitting the data, particularly in small samples.

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### What else could go wrong here?

- Some units (e.g. counties) may have few observations, making unpooled models impractical
- We may want to allow some effect of exposure (e.g. having a basement) to be consistent across counties.

# Door #3: Partial Pooling!

 Allow effects to vary across clusters, but constrain them with a prior distribution.

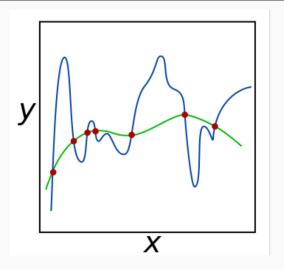
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- This approach accommodates variation across units without assuming they have no similarity.
- More likely to make accurate out-of-sample predictions than the fully-pooled or unpooled examples.

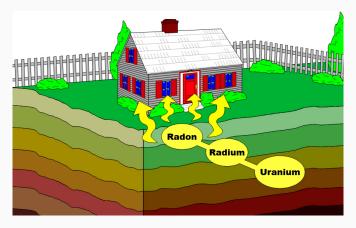
# Partial pooling = Regularization



Both functions fit the data perfectly...which one should you prefer?

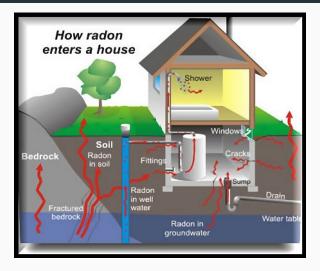
# Radon Example

# Radon is a carcinogenic gas



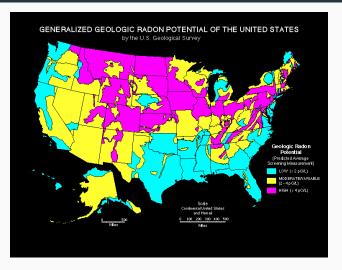
Radon is a byproduct of decaying soil uranium.

# Radon enters a house more easily when it is built into the ground



Ann Arbor is a radon hotspot!

# Considerable geographic variation in radon potential



Ann Arbor is a radon hotspot!

### Trust me on this one...



My very own radon mitigation system.

What should a model that accounts for important sources of variation in household radon potential include?

# What should a model that accounts for important sources of variation in household radon potential include?

- · County-level variation in soil uranium.
- Whether or not the radon measurement was taken in a basement.

#### Random intercepts account for county-level variation

Gelman [@Gelman2006] proposes a multi-level model to measure household radon in household i in county j,  $y_{ij}$ :

$$\cdot \ y_{ij} \sim N(\alpha_j + \beta x_{ij}, \sigma_y^2)$$
 , for  $i=1,\dots,n_j, j=1,\dots,J$ 

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#### Where:

- $\cdot \ \alpha_j$  is average, non-basement radon measure at county level
- $\beta$  is fixed effect measuring average change in radon level in houses with a basement.
- $\cdot$   $\sigma_y^2$  represents within-county variation in risk

#### Include predictors of county-level variation in second level

County-level random intercept is a function of county soil uranium measure,  $u_i$ :

$$\cdot \ \alpha_j \sim N(\gamma_0 + \gamma_1 u_j, \sigma_\alpha^2)$$
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#### Where:

- ·  $\gamma_0$  is expected household radon measure when  $u_i=0$
- ·  $\gamma_1$  scales expected county-level uranium with  $u_i$
- $\sigma_{\alpha}^2$  is between-county variation in radon risk not measured by  $u_{j}$ .

## Putting it all together

County-level intercept is a function of county soil uranium measure,  $\boldsymbol{u}_j$ :

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### Putting it all together

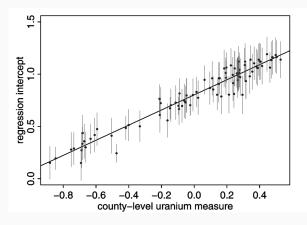
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Household-level radon measure is a function of having a basement and county-level intercept:

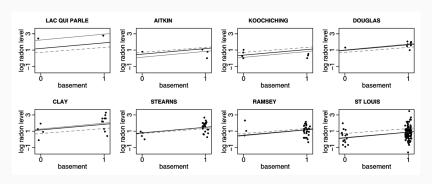
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#### County-level radon levels vary with soil uranium measures



County-level intercept,  $\alpha_j$ , ( $\pm 1$  standard error) as a function of county-level uranium.

### Model predictions vs. radon measures by county



Multi-level regression line,  $y=\alpha_j+\beta x$ , from 8 Minnesota counties. Unpooled estimates = light grey line; Totally pooled estimates = dashed grey line.

#### Next Time

• Hands-on with the Radon example

## References

#### References i