

# **Section: Module 5**

## **Instrumental Variable**

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# Agenda

- Review of Instrumental Variable
- Estimating Complier for multi-valued treatment
  - Point Estimate / Standard Error (Question 3)
- Weak Instrument

# Review of Instrumental Variable

- Setup
  - $Z_i$ : Instrument / Encouragement (randomized)
  - $T_i$ : Treatment (not randomized!)
  - $Y_i$ : Outcome
- Assumptions
  - Randomization of instrument
  - Exclusion restriction ( $Z_i$  influences outcome only through  $T_i$ )
  - Monotonicity (there is no defiers)
- Check review section's slide for identification

## Complier Type (Binary Treatment)

- When treatment is binary, we have four types
  - Compliers:  $T_i(Z_i = 1) = 1$  and  $T_i(Z_i = 0) = 0$
  - Always-takers:  $T_i(Z_i = 1) = T_i(Z_i = 0) = 1$
  - Never-takers:  $T_i(Z_i = 1) = T_i(Z_i = 0) = 0$
  - Defiers:  $T_i(Z_i = 1) = 0$  and  $T_i(Z_i = 0) = 1$

	$Z_i = 1$	$Z_i = 0$
$T_i = 1$	Complier / Always-taker	Defier / Always-taker
$T_i = 0$	Defier / Never-taker	Complier / Never-taker

- We exclude defier by monotonicity
  - As a result, we can identify each principal strata

## Estimating Complier for multi-valued treatment (1)

- What if we have multi-valued treatment
  - This is the setting of Question 3
- Consider the case where treatment is three category
  - I.e.,  $T_i \in \{0, 1, 2\}$
  - We keep instrument binary:  $Z_i \in \{0, 1\}$
- How many principal strata do we have?

## Estimating Complier for multi-valued treatment (2)

- We have 9 principal strata
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 2\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 2\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 2\}$
- We need to remove some of principal strata to identify the probability

## Estimating Complier for multi-valued treatment (2)

- Monotonicity  $\rightarrow$  we can remove strata  $T_i(Z_i = 0) > T_i(Z_i = 1)$
- We have 6 principal strata (remove 3 strata)
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 2\}$
  - $\{\cancel{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 0}\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 2\}$
  - $\{\cancel{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 0}\}$
  - $\{\cancel{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 1}\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 2\}$

## Estimating Complier for multi-valued treatment (3)

- Another Example:  $T_i(1) - T_i(0) \in \{-1, 0\}$ 
  - i.e., difference is at most one
- We have 5 principal strata (remove 4 strata)
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 0, T_i(Z_i = 1) = 2\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 1, T_i(Z_i = 1) = 2\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 0\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 1\}$
  - $\{T_i(Z_i = 0) = 2, T_i(Z_i = 1) = 2\}$

## Estimating Complier for multi-valued treatment (4)

- Can we estimate the proportion of each strata with standard errors?
- Because  $Z_i$  is randomized, for  $k \in \{0, 1, 2\}$

$$\mathbb{P}(T_i(1) = k) = \mathbb{P}(T_i(1) = k | Z_i = 1) = \mathbb{P}(T_i = k | Z_i = 1)$$

$$\mathbb{P}(T_i(0) = k) = \mathbb{P}(T_i(0) = k | Z_i = 0) = \mathbb{P}(T_i = k | Z_i = 0)$$

- Moreover, because sum of probability is 1,

$$\mathbb{P}(T_i(1) = 2 | T_i(0) = 2) + \mathbb{P}(T_i(1) = 1 | T_i(0) = 2) = 1$$

$$\mathbb{P}(T_i(1) = 1 | T_i(0) = 1) + \mathbb{P}(T_i(1) = 0 | T_i(0) = 1) = 1$$

$$\mathbb{P}(T_i(1) = 0 | T_i(0) = 0) = 1$$

## Estimating Complier for multi-valued treatment (5)

- Now, notice that by the definition of conditional probability,

$$1 = \mathbb{P}(T_i(1) = 0 \mid T_i(0) = 0) = \frac{\mathbb{P}(T_i(1) = 0, T_i(0) = 0)}{\mathbb{P}(T_i(0) = 0)}$$

- This means that

$$\mathbb{P}(T_i(0) = 0) = \mathbb{P}(T_i(1) = 0, T_i(0) = 0)$$

- Hence,

$$\mathbb{P}(T = 0 \mid Z = 0) = \mathbb{P}(T_i(1) = 0, T_i(0) = 0)$$

## Estimating Complier for multi-valued treatment (6)

- Now, notice that

$$\mathbb{P}(T_i(1) = 0) = \mathbb{P}(T_i(0) = 1, T_i(1) = 0) + \mathbb{P}(T_i(0) = 0, T_i(1) = 0)$$

- We know  $\mathbb{P}(T_i(1) = 0) = \mathbb{P}(T_i = 0 \mid Z_i = 1)$
- We also identify  $\mathbb{P}(T_i(0) = 0, T_i(1) = 0) = \mathbb{P}(T = 0 \mid Z = 0)$  from previous step
- Therefore,

$$\mathbb{P}(T_i(0) = 1, T_i(1) = 0) = \mathbb{P}(T_i = 0 \mid Z_i = 1) - \mathbb{P}(T_i = 0 \mid Z_i = 0)$$

## Estimating Complier for multi-valued treatment (7)

- What about  $\mathbb{P}(T_i(0) = 1, T_i(1) = 1)$ ?

$$\mathbb{P}(T_i(0) = 1, T_i(1) = 1) = \mathbb{P}(T_i(0) = 1) - \mathbb{P}(T_i(0) = 1, T_i(1) = 0)$$

- As we identify  $\mathbb{P}(T_i(0) = 1, T_i(1) = 0)$  already,

$$\begin{aligned}\mathbb{P}(T_i(0) = 1, T_i(1) = 1) \\ = \mathbb{P}(T_i = 1 \mid Z_i = 0) - (\mathbb{P}(T_i = 0 \mid Z_i = 1) - \mathbb{P}(T_i = 0 \mid Z_i = 0))\end{aligned}$$

- We simply repeat the same thing to identify all the strata

## Estimating Complier for multi-valued treatment (8)

- Can you estimate the standard error?

$$\mathbb{P}(\widehat{T_i(0)} = 1, \widehat{T_i(1)} = 0) = \mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 1) - \mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 0)$$

- Recall that  $\mathbb{V}[X - Y] = \mathbb{V}[X] + \mathbb{V}[Y] - 2\text{Cov}(X, Y)$
- Hence,

$$\begin{aligned}\mathbb{V}(\mathbb{P}(\widehat{T_i(0)} = 1, \widehat{T_i(1)} = 0)) \\ &= \mathbb{V}(\mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 1) - \mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 0)) \\ &= \mathbb{V}(\mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 1)) + \mathbb{V}(\mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 0)) \\ &\quad - 2 \underbrace{\text{Cov}(\mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 1), \mathbb{P}(\widehat{T_i = 0} \mid \widehat{Z_i} = 0))}_{=0(\because \text{Independence})}\end{aligned}$$

## Estimating Complier for multi-valued treatment (9)

- What about  $\mathbb{P}(\widehat{T}_i(0) = 1, \widehat{T}_i(1) = 1)$ ?

$$\begin{aligned}\mathbb{P}(\widehat{T}_i(0) = 1, \widehat{T}_i(1) = 1) \\ = \mathbb{P}(\widehat{T}_i = 1 \mid \widehat{Z}_i = 0) - (\mathbb{P}(\widehat{T}_i = 0 \mid \widehat{Z}_i = 1) - \mathbb{P}(\widehat{T}_i = 0 \mid \widehat{Z}_i = 0))\end{aligned}$$

- Hence,

$$\begin{aligned}\mathbb{V}\left(\mathbb{P}(\widehat{T}_i(0) = 1, \widehat{T}_i(1) = 1)\right) \\ = \mathbb{V}\left(\mathbb{P}(\widehat{T}_i = 1 \mid \widehat{Z}_i = 0) - (\mathbb{P}(\widehat{T}_i = 0 \mid \widehat{Z}_i = 1) - \mathbb{P}(\widehat{T}_i = 0 \mid \widehat{Z}_i = 0))\right)\end{aligned}$$

- **Caution:** Notice that  $\mathbb{P}(\widehat{T}_i = 1 \mid \widehat{Z}_i = 0)$  and  $\mathbb{P}(\widehat{T}_i = 0 \mid \widehat{Z}_i = 0)$  are not independent!
  - They are on the same sample ( $Z_i = 0$ )
  - Thus, you need to take into account **covariance**

## Estimating Complier for multi-valued treatment (9)

$$\begin{aligned} & \mathbb{V}\left(\mathbb{P}(T_i = \widehat{1} \mid Z_i = 0) - (\mathbb{P}(T_i = \widehat{0} \mid Z_i = 1) - \mathbb{P}(T_i = \widehat{0} \mid Z_i = 0))\right) \\ &= \mathbb{V}(\mathbb{P}(T_i = \widehat{1} \mid Z_i = 0)) \\ &\quad + \mathbb{V}(\mathbb{P}(T_i = \widehat{0} \mid Z_i = 1)) \\ &\quad + \mathbb{V}(\mathbb{P}(T_i = \widehat{0} \mid Z_i = 0))) \\ &\quad - 2 \underbrace{\text{Cov}(\mathbb{P}(T_i = \widehat{1} \mid Z_i = 0), \mathbb{P}(T_i = \widehat{0} \mid Z_i = 1))}_{=0} \\ &\quad + 2 \underbrace{\text{Cov}(\mathbb{P}(T_i = \widehat{1} \mid Z_i = 0), \mathbb{P}(T_i = \widehat{0} \mid Z_i = 0))}_{\text{Non-Zero!}} \\ &\quad - 2 \underbrace{\text{Cov}(\mathbb{P}(T_i = \widehat{0} \mid Z_i = 1), \mathbb{P}(T_i = \widehat{0} \mid Z_i = 0))}_{=0} \end{aligned}$$

## Estimating Complier for multi-valued treatment (10)

- How can we calculate covariance term?

$$\text{Cov}(\widehat{\mathbb{P}(T_i = 1 | Z_i = 0)}, \widehat{\mathbb{P}(T_i = 0 | Z_i = 0)})$$

- Why is it non-zero?  $\rightarrow$  If  $\widehat{\mathbb{P}(T_i = 1 | Z_i = 0)}$  becomes larger,  $\widehat{\mathbb{P}(T_i = 0 | Z_i = 0)}$  should be smaller
- From theory of multinomial distribution:

$$\begin{aligned}\text{Cov}(\widehat{\mathbb{P}(T_i = 1 | Z_i = 0)}, \widehat{\mathbb{P}(T_i = 0 | Z_i = 0)}) \\ = -\frac{\mathbb{P}(T_i = 1 | Z_i = 0)\mathbb{P}(T_i = 0 | Z_i = 0)}{\text{Number of } Z_i = 0}\end{aligned}$$

- Same for other covariance terms

## Added: What is multinomial distribution?

- Suppose there are  $K$  categories
- Think about how many times the trial where each  $T_i$  falls into one of  $K$  categories with probability

$$\mathbb{P}(T_i = k) = p_k \quad \text{with} \quad \sum_{k=1}^K p_k = 1$$

- We care about the counts of each category: i.e.,  
 $X_k = \sum_{i=1}^n \mathbb{1}\{T_i = k\}$
- The joint distribution of counts  $(X_1, \dots, X_K)$  follows multinomial distribution with  $\sum_{k=1}^K X_k = n$ 
  - $\mathbb{E}[X_k] = np_k$
  - $\mathbb{V}[X_k] = np_k(1 - p_k)$
  - $\text{Cov}(X_k, X_l) = -np_k p_l$
- As a result, the covariance of each probability estimates is

$$\text{Cov}(\hat{p}_k, \hat{p}_l) = \text{Cov}\left(\frac{X_k}{n}, \frac{X_l}{n}\right) = \frac{1}{n^2} \text{Cov}(X_k, X_l) = -\frac{p_k p_l}{n}$$

## Two Stage Least Squares

- Consider the following models:

$$Y_i = \alpha + \beta T_i + \epsilon_i$$

$$T_i = \gamma Z_i + \eta_i$$

where  $\mathbb{E}[\epsilon_i | Z_i] = \mathbb{E}[\eta_i | Z_i] = 0$

- Wald Estimator:**

$$\hat{\beta}_{\text{IV}} := \frac{\widehat{\text{Cov}(Y_i, Z_i)}}{\widehat{\text{Cov}(T_i, Z_i)}} = \frac{\text{Effect of } Z \text{ on } Y}{\text{Effect of } Z \text{ on } T}$$

## Two Stage Least Squares: Why it works?

$$\begin{aligned}\beta_{\text{IV}} &= \frac{\text{Cov}(Y_i, Z_i)}{\text{Cov}(T_i, Z_i)} \\&= \frac{\text{Cov}(\alpha + \beta T_i + \epsilon, Z_i)}{\text{Cov}(\gamma Z_i + \eta_i, Z_i)} \\&= \frac{\text{Cov}(\beta T_i, Z_i)}{\text{Cov}(\gamma Z_i, Z_i)} \quad (\because \text{Exogeneity}) \\&= \frac{\text{Cov}(\beta(\gamma Z_i + \eta_i), Z_i)}{\text{Cov}(\gamma Z_i, Z_i)} \\&= \frac{\beta \gamma \mathbb{V}[Z_i]}{\gamma \mathbb{V}[Z_i]} = \beta\end{aligned}$$

## Weak Instrument (1)

- IV is unstable when instrument weakly affects treatment  $\gamma \approx 0$ 
  - Let's see how bias appears
  - For the sake of simplicity, assume  $\bar{Z} = 0$
- Then, Wald estimator is written as

$$\hat{\beta}_{\text{IV}} := \frac{\widehat{\text{Cov}}(Y_i, Z_i)}{\widehat{\text{Cov}}(T_i, Z_i)} = \frac{\frac{1}{n} \sum_{i=1}^n Y_i Z_i}{\frac{1}{n} \sum_{i=1}^n T_i Z_i}$$

## Weak Instrument (2)

$$\begin{aligned}\hat{\beta}_{\text{IV}} - \beta &= \frac{\frac{1}{n} \sum_{i=1}^n Y_i Z_i}{\frac{1}{n} \sum_{i=1}^n T_i Z_i} - \beta \\&= \frac{\frac{1}{n} \sum_{i=1}^n (\alpha + \beta T_i + \epsilon_i) Z_i}{\frac{1}{n} \sum_{i=1}^n T_i Z_i} - \beta \\&= \frac{\frac{1}{n} \sum_{i=1}^n (\beta T_i Z_i + \epsilon_i Z_i)}{\frac{1}{n} \sum_{i=1}^n T_i Z_i} - \beta \\&= \beta + \frac{\frac{1}{n} \sum_{i=1}^n \epsilon_i Z_i}{\frac{1}{n} \sum_{i=1}^n T_i Z_i} - \beta = \frac{\frac{1}{n} \sum_{i=1}^n \epsilon_i Z_i}{\frac{1}{n} \sum_{i=1}^n T_i Z_i}\end{aligned}$$

- If  $\gamma = 0$ , then  $T_i = \gamma Z_i + \eta_i = \eta_i$ . So,

$$\hat{\beta}_{\text{IV}} - \beta = \frac{\frac{1}{n} \sum_{i=1}^n \epsilon_i Z_i}{\frac{1}{n} \sum_{i=1}^n \eta_i Z_i}$$