Regression and SVD

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Problem 1.1. Show that $\operatorname{rank}(X) = \operatorname{rank}(X^{\top}X)$ for any matrix X with real entries.

Solution:

Method 1: Suppose $X \in \mathbb{R}^{m \times n}$. Then, by SVD, we have $X = U\Sigma V^{\top}$, where $U \in \mathbb{R}^{m \times m}$, $V \in \mathbb{R}^{n \times n}$ are orthogonal matrices, and $\Sigma \in \mathbb{R}^{m \times n}$ having non-negative singular values on its diagonal and zeroes elsewhere.

Since U, V are orthogonal, they are invertible. Because multiplying an invertible matrix does not change the rank of a matrix, we get $\operatorname{rank}(X) = \operatorname{rank}(U\Sigma V^\top) = \operatorname{rank}(\Sigma)$.

On the other hand, we have

$$\operatorname{rank}(X^\top X) = \operatorname{rank}(V \Sigma^\top U^\top U \Sigma V^\top) = \operatorname{rank}(V \Sigma^\top \Sigma V^\top) = \operatorname{rank}(\Sigma^\top \Sigma) = \operatorname{rank}(\Sigma)$$

Hence, we have $\operatorname{rank}(X) = \operatorname{rank}(X^{\top}X)$.

Method 2: We will prove $Im(X^{\top}) = Im(X^{\top}X)$. First, we have

$$Im(X^{\top}X) = \{X^{\top}Xv \mid v \in \mathbb{R}^n\} = \{X^{\top}u \mid u \in Im(X)\}\$$

By the rank-nullity theorem, we get $dim(Im(X)) + dim(Ker(X^{\top})) = m$. Moreover, Im(X) and $Ker(X^{\top})$ are orthogonal. Therefore Im(X) and $Ker(X^{\top})$ together span the whole space \mathbb{R}^m .

It follows that $\forall u \in \mathbb{R}^m$, we can uniquely decompose it as the sum $u = u_r + u_n$ where $u_r \in Im(X)$ and $u_n \in Ker(X^\top)$.

Therefore, we obtain

$$Im(X^{\top}) = \{X^{\top}u, u \in \mathbb{R}^m\} = \{X^{\top}(u_r + u_n) \mid u_r \in Im(X), u_n \in Ker(X^{\top})\} = \{X^{\top}u_r \mid u_r \in Im(X)\} = Im(X^{\top}X)$$

Hence, we conclude

$$\operatorname{rank}(X) = \dim(\operatorname{Im}(X)) = \dim(\operatorname{Im}(X^\top)) = \dim(\operatorname{Im}(X^\top X)) = \operatorname{rank}(X^\top X)$$

Problem 1.2. Write a numpy function to compute the pseudo-inverse of a real matrix, building upon the numpy SVD function discussed in class. Test that your function works by generating a sequence of random invertible matrices, and check that for each one, your pseudo-inverse equals the actual inverse up to numerical precision.

Solution: WLOG, suppose $X \in \mathbb{R}^{m \times n}$, $m \le n$. Then, by SVD, we have $X = U\Sigma V^{\top}$, where $U \in \mathbb{R}^{m \times m}$, $V \in \mathbb{R}^{n \times n}$ are orthogonal matrices, and $\Sigma \in \mathbb{R}^{m \times n}$ has non-negative singular values $\sigma_1 \ge \cdots \ge \sigma_p \ge \sigma_{p+1} = \cdots = \sigma_m = 0$, $p \le \min\{m, n\}$, on its diagonal and zeroes elsewhere.

Then, we have

$$X^\dagger = \left(X^\top X\right)^{-1} X^\top = \left(V \Sigma^\top U^\top U \Sigma V^\top\right)^{-1} V \Sigma^\top U^\top = \left(V (\Sigma^\top \Sigma) \, V^\top\right)^{-1} V \Sigma^\top U^\top = V (\Sigma^\top \Sigma)^{-1} \, \Sigma^\top U^\top = V \Sigma^\dagger U^\top$$

That is, the pseudo-inverse given by the SVD is $X^{\dagger} = V \Sigma^{\dagger} U^{\top}$.

For the code, see the attached jupyter notebook.

Problem 1.3. Prove that the Moore-Penrose pseudo-inverse is given as a one-sided limit of ridge regression problems, i.e. prove

$$X^\dagger y = \lim_{\lambda \to 0^+} \left(X^\top X + \lambda I \right)^{-1} X^\top y$$

Hint: use the SVD.

Solution: Recall that an alternate definition of X^{\dagger} , see [1] Albert (1972), is

Definition 1.4. The Moore-Penrose pseudo-inverse X^{\dagger} is defined so that $X^{\dagger}y$ is the minimum-norm vector among all minimizers of $\min_{\beta} \|y - X\beta\|^2$.

Therefore, we want to show the limit we get from the RHS is a minimizer of $\min_{\beta} \|y - X\beta\|^2$, and it's the one with the smallest norm among all minimizers. We apply the SVD to help us show this.

WLOG, suppose $X \in \mathbb{R}^{m \times n}$, $m \le n$. Then, by SVD, we have $X = U \Sigma V^{\top}$, where $U \in \mathbb{R}^{m \times m}$, $V \in \mathbb{R}^{n \times n}$ are orthogonal matrices, and $\Sigma \in \mathbb{R}^{m \times n}$ has non-negative singular values $\sigma_1 \ge \cdots \ge \sigma_p \ge \sigma_{p+1} = \cdots = \sigma_m = 0$, $p \le \min\{m, n\}$, on its diagonal and zeroes elsewhere.

Then, we have

$$(X^{\top}X + \lambda I)^{-1} X^{\top}y = (V\Sigma^{\top}U^{\top}U\Sigma V^{\top} + \lambda I)^{-1} V\Sigma^{\top}U^{\top}y$$

$$= (V(\Sigma^{\top}\Sigma + \lambda I) V^{\top})^{-1} V\Sigma^{\top}U^{\top}y$$

$$= V(\Sigma^{\top}\Sigma + \lambda I)^{-1} \Sigma^{\top}U^{\top}y$$

$$= V\begin{bmatrix} \frac{\sigma_{1}}{\lambda + \sigma_{1}^{2}} & & \\ & \ddots & \\ & & \frac{\sigma_{p}}{\lambda + \sigma_{p}^{2}} & \end{bmatrix} U^{\top}y$$

We then take the limit and obtain

$$X^{\dagger}y = \lim_{\lambda \to 0^{+}} \left(X^{\top}X + \lambda I\right)^{-1} X^{\top}y = \lim_{\lambda \to 0^{+}} V \begin{bmatrix} \frac{\sigma_{1}}{\lambda + \sigma_{1}^{2}} & & \\ & \ddots & \\ & & \frac{\sigma_{p}}{\lambda + \sigma_{p}^{2}} \end{bmatrix} U^{\top}y = V \begin{bmatrix} \frac{1}{\sigma_{1}} & & \\ & \ddots & \\ & & \frac{1}{\sigma_{p}} & \end{bmatrix} U^{\top}y := VSU^{\top}y, \quad S \in \mathbb{R}^{n \times m}$$

To be a minimizer of $\min_{\beta} \|y - X\beta\|^2$, we need $\beta^* = X^{\dagger}y = VSU^{\top}y$ to satisfy the first-order necessary condition

$$\frac{\partial L(\beta)}{\partial \beta} = \frac{\partial}{\partial \beta} \frac{1}{2} \|y - X\beta\|^2 = \frac{\partial}{\partial \beta} \frac{1}{2} (y - X\beta)^\top (y - X\beta) = X^\top X\beta - X^\top y = 0 \tag{1}$$

Plug $\beta^* = VSU^{\top}y$ in Equation (1) and apply SVD on X, we get

$$\begin{split} X^\top X \beta^* - X^\top y &= V \Sigma^\top U^\top U \Sigma V^\top V S U^\top y - V \Sigma^\top U^\top y \\ U, \text{Vorthogonal} &\Longrightarrow = V \Sigma^\top \Sigma S U^\top y - V \Sigma^\top U^\top y \\ &= V \begin{bmatrix} \sigma_1 & & & \\ & \ddots & & \\ & & \sigma_p & & \\ & & & \ddots & \\ & & & \sigma_p & & \\ & & & & \end{bmatrix} \begin{bmatrix} \frac{1}{\sigma_1} & & & \\ & \ddots & & \\ & & \frac{1}{\sigma_p} & & \\ & & & \end{bmatrix} U^\top y - V \Sigma^\top U^\top y \\ &= V \Sigma^\top U^\top y - V \Sigma^\top U^\top y \\ &= 0 \end{split}$$

Therefore, $\beta^* = X^{\dagger}y = VSU^{\top}y$ is a minimizer of $\min_{\beta} \|y - X\beta\|^2$.

We now assume that there exists another minimizer $\beta_1 \in \mathbb{R}^n$ other than β^* and that $y = X\beta_1 + \varepsilon$ where $\varepsilon \in \mathbb{R}^m$.

Then, by Equation (1), we have

$$0 = X^{\top}X\beta_1 - X^{\top}y = X^{\top}X\beta_1 - X^{\top}(X\beta_1 + \epsilon) \Longrightarrow 0 = X^{\top}\varepsilon = V\Sigma^{\top}U^{\top}\varepsilon \Longrightarrow 0 = \Sigma^{\top}U^{\top}\varepsilon = \begin{bmatrix} \sigma_1u_1^{\top}\varepsilon \\ \vdots \\ \sigma_pu_p^{\top}\varepsilon \\ 0 \\ \vdots \\ 0 \end{bmatrix} \Longrightarrow u_i^{\top}\varepsilon = 0, \text{ for } i = 1, \dots, p$$

It follows that

$$SU^{\top} \varepsilon = \begin{bmatrix} \frac{1}{\sigma_1} & & & \\ & \ddots & & \\ & & \frac{1}{\sigma_p} & \\ & & & \end{bmatrix} U^{\top} \varepsilon = \begin{bmatrix} \frac{1}{\sigma_1} u_1^{\top} \varepsilon \\ \vdots \\ \frac{1}{\sigma_p} u_p^{\top} \varepsilon \\ 0 \\ \vdots \\ 0 \end{bmatrix} = 0$$

Finally, we consider $\beta^{*^{\top}}\beta^{*}$ and obtain

$$\begin{split} \|\beta^*\|^2 &= \beta^{*^\top}\beta^* = (VSU^\top y)^\top VSU^\top y \\ &= (VSU^\top X\beta_1 + VSU^\top \varepsilon)^\top (VSU^\top X\beta_1 + VSU^\top \varepsilon) \\ SU^\top \varepsilon &= 0 \Longrightarrow = (VS\Sigma V^\top \beta_1)^\top (VS\Sigma V^\top \beta_1) \\ &= \beta_1^\top V\Sigma^\top S^\top V^\top VS\Sigma V^\top \beta_1 \\ V \text{ orthogonal } &\Longrightarrow = \beta_1^\top V\Sigma^\top S^\top S\Sigma V^\top \beta_1 \\ &= \beta_1^\top V \begin{bmatrix} 1 & \ddots & & & \\ & 1 & & & \\ & & 0 & & \ddots & \\ & & & & 0 \end{bmatrix}^\top \begin{bmatrix} 1 & \ddots & & \\ & & 1 & & \\ & & & 0 & \\ & & & \ddots & \\ & & & & 0 \end{bmatrix} V^\top \beta_1 \\ &= \beta_1^\top \left(\sum_{i=1}^p v_i v_i^\top \right) \beta_1 \\ &= \beta_1^\top \left(\sum_{i=1}^n v_i v_i^\top \right) \beta_1 - \beta_1^\top \left(\sum_{i=p+1}^n v_i v_i^\top \right) \beta_1 \\ &= \beta_1^\top VV^\top \beta_1 - \sum_{i=p+1}^n \beta_1^\top v_i v_i^\top \beta_1 \\ &= \|\beta_1\|^2 - \sum_{i=p+1}^n \|v_i^\top \beta_1\|^2 \\ &< \|\beta\|^2 \end{split}$$

where the equality is attained when p = n, i.e., X must be full column ranked.

Since β_1 is chosen arbitrarily, we can now conclude that $\beta^* = X^{\dagger}y = VSU^{\top}y$ is a minimizer of $\min_{\beta} \|y - X\beta\|^2$, and it's the one with the smallest norm among all minimizers.

Hence, we conclude that the limit given by $X^{\dagger}y = \lim_{\lambda \to 0^+} \left(X^{\top}X + \lambda I\right)^{-1} X^{\top}y$ is, by definition, the Moore-Penrose pseudo-inverse.

References

[1] Albert, Arthur (1972). Regression and the Moore-Penrose pseudo-inverse. Academic Press.