

5.3

Solution:

(a) The n samples here should be independent.

$$Y = \sum_{i=1}^n x_i \sim \text{Gamma}(2n, \theta)$$

$$(b) L(\theta) = \prod_{i=1}^n \left(\frac{\theta^2}{\Gamma(2)} x_i e^{-\theta x_i} \right) = \frac{\theta^{2n}}{(\Gamma(2))^n} \left(\prod_{i=1}^n x_i \right) e^{-\theta \sum_{i=1}^n x_i}$$

$$\text{let } \frac{L(\theta_1)}{L(\theta_0)} = \frac{\theta_1^{2n}}{\theta_0^{2n}} e^{-\theta_1 \sum_{i=1}^n x_i + \theta_0 \sum_{i=1}^n x_i} = R. \quad (\neq)$$

Since $\theta_1 > 1 = \theta_0$, $\theta_1 - \theta_0 > 0$.

$$(\neq) \Leftrightarrow \sum_{i=1}^n x_i \leq \frac{\log\left(\frac{\theta_1}{\theta_0}\right) k}{\theta_1 - \theta_0} \triangleq C.$$

When H_0 is true we have $\sum_{i=1}^n x_i \sim \text{Gamma}(2n, \theta)$.

$$\begin{aligned} \alpha &= \Pr(\vec{x} \in C \mid \theta = \theta_0 = 1) = \Pr\left(\sum_{i=1}^n x_i \leq C \mid \theta = \theta_0 = 1\right) \\ &= \Pr(\text{Gamma}(2n, 1) \leq C) = \Pr(\chi^2(4n) \leq 2C) \\ &= \Pr(\chi^2(4n) \leq \chi^2(1-\alpha, 4n)) \\ \Rightarrow C &= \frac{1}{2} \chi^2(1-\alpha, 4n) \end{aligned}$$

\Rightarrow The test with critical region $C = \{\vec{x} : \sum_{i=1}^n x_i \leq \frac{1}{2} \chi^2(1-\alpha, 4n)\}$ is the most powerful test of size α .

(c)

$$P(\theta) = \Pr(\vec{x} \in C \mid \theta)$$

$$\begin{aligned} &= \Pr\left(\sum_{i=1}^n x_i \leq \text{Gamma}(1-\alpha, 2n, 1) \mid \theta\right) \\ &= \int_{-\infty}^{\chi^2(1-\alpha, 4n)} \frac{\theta^{2n}}{\Gamma(2n)} x^{2n-1} e^{-\theta x} dx. \end{aligned}$$

Solution:

$$(a) L(\theta) = \theta^n \prod_{i=1}^n (1-x_i)^{\theta-1} \mathbb{1}_{(0 < x_i < 1)}$$

$$\frac{\partial L(\theta_0)}{\partial L(\theta_1)} = \frac{\theta_0^n}{\theta_1^n} \prod_{i=1}^n (1-x_i)^{\theta_0 - \theta_1} \leq k$$

Since $\theta_1 > \theta_0 = 1$,

$$(\theta_0 - \theta_1) \sum_{i=1}^n \log(1-x_i) \leq \log\left(\frac{k \theta_1^n}{\theta_0^n}\right)$$

$$\sum_{i=1}^n \log(1-x_i) \geq \frac{1}{\theta_0 - \theta_1} \log\left(\frac{k \theta_1^n}{\theta_0^n}\right) = c.$$

Note that for $y_i \equiv -\log(1-x_i)$

$$y_i \sim \text{Exponential}(\theta)$$

$$\Rightarrow -\sum_{i=1}^n \log(1-x_i) = \sum_{i=1}^n y_i \sim \text{Gamma}(n, \theta)$$

$$\alpha = \Pr(\vec{x} \in C \mid \theta = \theta_0 = 1) = \Pr\left(\sum_{i=1}^n y_i \leq -c \mid \theta = \theta_0 = 1\right)$$

$$= \Pr(\text{Gamma}(n, 1) \leq -c) = \Pr(\chi^2(2n) \leq -2c) = \Pr(\chi^2(2n) \leq \chi^2(1-\alpha, 2n))$$

$$\Rightarrow -2c = \chi^2(1-\alpha, 2n)$$

$$c = -\frac{1}{2} \chi^2(1-\alpha, 2n)$$

\Rightarrow The test with critical region $\{\vec{x}: \sum_{i=1}^n \log(1-x_i) \geq -\frac{1}{2} \chi^2(1-\alpha, 2n)\}$.

is the most powerful test of size α .

$$(b). \Theta_0 = \{1\}, \Theta_1 = (0, 1) \cup (1, +\infty) \quad \Theta = (0, +\infty) \subseteq \mathbb{R} = \Theta^*$$

Step 1: Calculate $\lambda(\vec{x})$:

$$L(\theta) = \prod_{i=1}^n \theta(1-x_i)^{\theta-1} \mathbb{1}_{(0 < x_i < 1)}$$

$$= \theta^n \exp\left\{(\theta-1) \sum_{i=1}^n \log(1-x_i)\right\} \Rightarrow \hat{\theta} = \frac{-n}{\sum_{i=1}^n \log(1-x_i)} \quad \left(\sum_{i=1}^n \log(1-x_i) < 0\right)$$

$$\log L(\theta) = n \log \theta + (\theta-1) \sum_{i=1}^n \log(1-x_i)$$

$$\frac{d \log L(\theta)}{d\theta} = \frac{n}{\theta} + \sum_{i=1}^n \log(1-x_i)$$

$$\frac{\sup_{\theta \in \Theta_0} L(\theta)}{\sup_{\theta \in \Theta} L(\theta)} = \frac{1}{\hat{\theta}^n \left(\prod_{i=1}^n (1+x_i) \right)^{\hat{\theta}-1}}$$

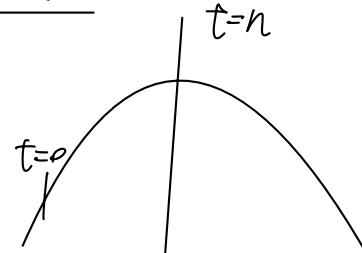
Step 2: Find the critical region C .

$$C = \left\{ \frac{1}{\hat{\theta}^n \left(\prod_{i=1}^n (1+x_i) \right)^{\hat{\theta}-1}} \leq \lambda_\alpha \right\}. \quad T(\vec{x}) = -\sum_{i=1}^n \log(1+x_i) \quad \hat{\theta} = \frac{n}{t}$$

$$= \left\{ \frac{1}{\frac{n^n}{t^n} e^{-(\frac{n}{t}-1)t}} \leq \lambda_\alpha \right\} = \left\{ \frac{1}{n^n t^n} e^{n-t} \leq \lambda_\alpha \right\}.$$

$$= \left\{ \vec{x}: t > 0, h(t) \leq \lambda_\alpha \right\}.$$

Step 2(a): Check if or not $f(t)$ is log-concave.



$$f(t) \triangleq \log f(t) = n \log t + t - n - n \log n$$

$$f'(t) = \frac{n}{t} - 1 \quad f''(t) = -\frac{n}{t^2} < 0$$

$\Rightarrow f(t)$ is strictly log-concave (attain its maximum at $t = n$).

Step 2(b): Find equivalent C .

$$t > 0, f(t) \leq \lambda_\alpha \iff 0 < t \leq C_1 \text{ or } t \geq C_2.$$

$$C = \left\{ \vec{x}: 0 < T \leq C_1 \text{ or } T \geq C_2 \right\}.$$

Step 2(c): Determine C_1 & C_2

$$\alpha = \Pr \{ T \leq C_1 \mid \theta = \theta_0 \} + \Pr \{ T \geq C_2 \mid \theta = \theta_0 \}$$

We use equal-tail approach.

$$\text{Let } \bar{\alpha} = \Pr \{ T \leq C_1 \mid \theta = \theta_0 \} = \Pr \{ T \geq C_2 \mid \theta = \theta_0 \}$$

Since $T \sim \text{Gamma}(n, \theta)$ by (a)

Under H_0 , $T \sim \text{Gamma}(n, 1)$

$$2T \sim \text{Gamma}(n, \frac{1}{2}) = \chi^2(2n)$$

$$\Rightarrow \frac{\alpha}{2} = \Pr \left\{ \chi^2(2n) \leq C_1 \right\} = \Pr \left\{ \chi^2(2n) \leq \chi^2(1 - \frac{\alpha}{2}, 2n) \right\}$$

$$= \Pr \left\{ \chi^2(2n) \geq C_2 \right\} = \Pr \left\{ \chi^2(2n) \geq \chi^2(\frac{\alpha}{2}, 2n) \right\}$$

$$\Rightarrow C_1 = \frac{1}{2} \chi^2(1 - \frac{\alpha}{2}, 2n), \quad C_2 = \frac{1}{2} \chi^2(\frac{\alpha}{2}, 2n).$$

\Rightarrow The LRT with size α is the test with critical region

$$C = \left\{ \vec{x} : 0 < -\sum_{i=1}^n \log(Tx_i) \leq \frac{1}{2} \chi^2(1 - \frac{\alpha}{2}, 2n) \right.$$

$$\left. \text{or} \quad -\sum_{i=1}^n \log(Tx_i) \geq \frac{1}{2} \chi^2(\frac{\alpha}{2}, 2n) \right\}.$$

5.5

Solution: Step 1: Consider test

$$H_{0S}: \theta = \theta_0 \quad \text{against} \quad H_{1S}: \theta = \theta_1 < \theta_0. \quad (*)$$

$$L(\theta) = (2\pi)^{-\frac{n}{2}} \exp \left\{ -\frac{1}{2} \sum_{i=1}^n (x_i - \theta)^2 \right\}.$$

$$\frac{L(\theta_0)}{L(\theta_1)} = \exp \left\{ -\frac{1}{2} \sum_{i=1}^n (x_i - \theta_0)^2 + \frac{1}{2} \sum_{i=1}^n (x_i - \theta_1)^2 \right\}.$$

$$= \exp \left\{ \frac{1}{2} \sum_{i=1}^n (2x_i - \theta_0 - \theta_1)(\theta_0 - \theta_1) \right\}$$

$$= \exp \left\{ (\theta_0 - \theta_1) \sum_{i=1}^n x_i - \frac{n}{2} (\theta_0 - \theta_1)^2 \right\} \leq k.$$

$$\Leftrightarrow \sum_{i=1}^n x_i = \frac{\log k + \frac{n}{2}(\theta_0 - \theta_1)^2}{(\theta_0 - \theta_1)} \triangleq C.$$

Since $\sum_{i=1}^n x_i \sim N(n\theta, n)$, $\frac{\sum_{i=1}^n x_i - n\theta}{\sqrt{n}} \sim N(0, 1)$

$$\begin{aligned}\alpha &= \Pr(\vec{x} \in C \mid \theta = \theta_0) = \Pr\left(\frac{\sum_{i=1}^n x_i - n\theta_0}{\sqrt{n}} \leq \frac{C - n\theta_0}{\sqrt{n}}\right) \\ &= \Pr\left\{\frac{\sum_{i=1}^n x_i - n\theta_0}{\sqrt{n}} \leq z_{1-\alpha}\right\}\end{aligned}$$

$$\Rightarrow \frac{C - n\theta_0}{\sqrt{n}} = z_{1-\alpha} = -z_\alpha \quad \Rightarrow \quad C = n\theta_0 - \sqrt{n} z_\alpha.$$

\Rightarrow the test γ with critical region

$C = \{\vec{x} : \sum_{i=1}^n x_i \leq n\theta_0 - \sqrt{n} z_\alpha\}$ is a MPT with size α .

Step 2: Consider the test

$$H_{0S}: \theta = \theta_0 \text{ against } H_1: \theta < \theta_0.$$

Since the C above depends only on n, θ_0, α and the fact that $\theta_1 < \theta_0$, but not on the value of θ_1 .

\Rightarrow The test γ is also the UMP test of size α for testing

$$H_{0S}: \theta = \theta_0 \text{ against } H_1: \theta < \theta_0$$

Step 3:

$$\begin{aligned}\sup_{\theta \geq \theta_0} P_\gamma(\theta) &= \sup_{\theta \geq \theta_0} \Pr\left\{\sum_{i=1}^n x_i \geq C \mid \theta\right\} \\ &= \sup_{\theta \geq \theta_0} \Pr\left\{\frac{\sum_{i=1}^n x_i - n\theta}{\sqrt{n}} \leq \frac{C - n\theta}{\sqrt{n}}\right\} \\ &= \sup_{\theta \geq \theta_0} \Pr\left\{N(0, 1) \leq \frac{C - n\theta}{\sqrt{n}}\right\} = \Pr\left\{N(0, 1) \leq \frac{C - n\theta_0}{\sqrt{n}}\right\} \\ &= \Pr\left\{N(0, 1) \leq z_{-\alpha}\right\} = \alpha\end{aligned}$$

Then, the test φ is also the UMP test of size α for testing

$$H_0: \theta \geq \theta_0, \quad H_1: \theta < \theta_0.$$

5.6

Solution:

(a) We first consider $H_0: \sigma^2 = \sigma_0^2$ against $H_1: \sigma^2 \neq \sigma_0^2$.

$$L(\mu, \sigma^2) = (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2}$$

$$\Theta_0 = \{(\mu, \sigma^2) : -\infty < \mu < +\infty, \sigma^2 = \sigma_0^2\}, \quad \Theta = \{(\mu, \sigma^2) : -\infty < \mu < +\infty, \sigma^2 \geq 0\}.$$

$$l(\mu, \sigma^2) = -\frac{n}{2} \log 2\pi - \frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2.$$

$$\text{Let } \begin{cases} \frac{\partial l}{\partial \mu} = \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu) = 0 \\ \frac{\partial l}{\partial \sigma^2} = -\frac{n}{2} \frac{1}{\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (x_i - \mu)^2 = 0 \end{cases} \Rightarrow \begin{cases} \mu = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x} \\ \sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \end{cases}$$

$$\lambda(\bar{x}) = \frac{(2\pi\sigma_0^2)^{-\frac{n}{2}} \exp(-\frac{1}{2\sigma_0^2} \sum_{i=1}^n (x_i - \bar{x})^2)}{\left[2\pi \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2\right]^{-\frac{n}{2}} \exp\left(-\frac{n}{2 \sum_{i=1}^n (x_i - \bar{x})^2} \sum_{i=1}^n (x_i - \bar{x})^2\right)}$$

$$= \left[\frac{(n-1) S^2}{n \sigma_0^2} \right]^{\frac{n}{2}} \exp\left(-\frac{1}{2\sigma_0^2} (n-1) S^2 + \frac{n}{2}\right) \triangleq f(S^2)$$

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2.$$

The critical region that H_0 is rejected is

$$\begin{aligned} C &= \{\bar{x} : \lambda(\bar{x}) \leq \lambda_\alpha\} \\ &= \{\bar{x} : \left[\frac{(n-1) S^2}{n \sigma_0^2} \right]^{\frac{n}{2}} \exp\left(-\frac{1}{2\sigma_0^2} (n-1) S^2 + \frac{n}{2}\right) \leq \lambda_\alpha\} \\ &= \left\{ \bar{x} : \frac{n}{2} \log \frac{n-1}{n \sigma_0^2} + \frac{n}{2} \log S^2 - \frac{1}{2\sigma_0^2} (n-1) S^2 + \frac{n}{2} \leq \log \lambda_\alpha \right\}. \end{aligned}$$

$$\text{Let } g(S^2) = \log f(S^2). \quad C = \left\{ \vec{x} : g(S^2) \leq \log \lambda_\alpha \right\}.$$

$$\frac{d g(S^2)}{d S^2} = \frac{n}{2} \frac{1}{S^2} - \frac{1}{2\alpha_0^2} (n-1), \quad \frac{d^2 g(S^2)}{d(S^2)^2} = -\frac{n}{2} \frac{1}{(S^2)^2} < 0.$$

$\Rightarrow f(S^2)$ is log-concave and attains its maximum at $\frac{n}{n-1}\alpha_0^2$.

$$\Rightarrow f(S^2) \leq \lambda_\alpha \Leftrightarrow 0 < S^2 \leq C_1 \text{ or } S^2 \geq C_2,$$

We select equal-tail approach.

$$\frac{\alpha}{2} = \Pr(0 < S^2 \leq C_1 \mid \alpha^2 = \alpha_0^2)$$

$$= \Pr\left(0 < \frac{(n-1)S^2}{\alpha_0^2} \leq \frac{(n-1)C_1}{\alpha_0^2}\right)$$

$$= \Pr\left(0 < \chi^2(n-1) \leq \frac{(n-1)C_1}{\alpha_0^2}\right).$$

$$= \Pr\left(0 < \chi^2(n-1) \leq \chi^2(1 - \frac{\alpha}{2}, n-1)\right).$$

$$\Rightarrow \frac{(n-1)C_1}{\alpha_0^2} = \chi^2(1 - \frac{\alpha}{2}, n-1). \quad C_1 = \frac{\alpha_0^2}{n-1} \chi^2(1 - \frac{\alpha}{2}, n-1).$$

$$\frac{\alpha}{2} = \Pr(S^2 \geq C_2 \mid \alpha^2 = \alpha_0^2)$$

$$= \Pr\left(\chi^2(n-1) \geq \frac{(n-1)C_2}{\alpha_0^2}\right)$$

$$= \Pr\left(\chi^2(n-1) \geq \chi^2(\frac{\alpha}{2}, n-1)\right) \Rightarrow C_2 = \frac{\alpha_0^2}{n-1} \chi^2(\frac{\alpha}{2}, n-1).$$

\Rightarrow The critical region is

$$C = \left\{ \vec{x} : 0 < S^2 \leq \frac{\alpha_0^2}{n-1} \chi^2(1 - \frac{\alpha}{2}, n-1) \text{ or } S^2 \geq \frac{\alpha_0^2}{n-1} \chi^2(\frac{\alpha}{2}, n-1) \right\}.$$

(b). Then we consider

$$H_0: \alpha^2 = \alpha_0^2 \text{ against } H_1: \alpha^2 > \alpha_0^2.$$

The critical region is $C = \left\{ \vec{x} : S^2 \geq \frac{\alpha_0^2}{n-1} \chi^2(\alpha, n-1) \right\}$.

(c) Finally, consider

$$H_0: \alpha^2 = \alpha_0^2 \text{ against } H_1: \alpha^2 < \alpha_0^2.$$

The critical region is $\mathbb{C} = \left\{ \vec{x} : 0 \leq S^2 \leq \frac{\alpha_0^2}{n-1} \chi^2(n-1, n-1) \right\}$.

5.7

Solution:

$$\vec{\theta} = (M_1, M_2, \sigma^2)^T, \Theta_0 = \{(0, 0, \sigma^2)^T, \sigma^2 > 0\}, \Theta = \{(M_1, M_2, \sigma^2)^T, M_1, M_2 \in \mathbb{R}, \sigma^2 > 0\}.$$

$H_0: M_1 = M_2 = 0$ against $H_1: M_1 \neq M_2$ or $M_1 = M_2 \neq 0$

$$L(M_1, M_2, \sigma^2) = (2\pi\sigma^2)^{-n} \exp \left[-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - M_1)^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - M_2)^2 \right].$$

$$l(M_1, M_2, \sigma^2) = \log L(M_1, M_2, \sigma^2)$$

$$= -n \log 2\pi - n \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{i=1}^n [(x_i - M_1)^2 + (y_i - M_2)^2].$$

$$\text{let } \begin{cases} \frac{\partial l}{\partial M_1} = \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - M_1) = 0 \\ \frac{\partial l}{\partial M_2} = \frac{1}{\sigma^2} \sum_{i=1}^n (y_i - M_2) = 0 \\ \frac{\partial l}{\partial \sigma^2} = -\frac{n}{\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n [(x_i - M_1)^2 + (y_i - M_2)^2] = 0. \end{cases}$$

$$\Rightarrow \hat{M}_1 = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}, \quad \hat{M}_2 = \bar{y}, \quad \hat{\sigma}^2 = \frac{1}{2} \frac{n-1}{n} (S_1^2 + S_2^2).$$

$$\begin{aligned} \text{Under } H_0: \hat{\sigma}^2 R &= \frac{1}{2n} \sum_{i=1}^n (x_i^2 + y_i^2) = \frac{1}{2n} \left[\sum_{i=1}^n (x_i - \bar{x})^2 + n\bar{x}^2 + \sum_{i=1}^n (y_i - \bar{y})^2 + n\bar{y}^2 \right] \\ &= \frac{(n-1)S_1^2 + (n-1)S_2^2 + n(\bar{x}^2 + \bar{y}^2)}{2n}. \end{aligned}$$

$$\begin{aligned} \lambda(\bar{x}) &= \frac{L(0, 0, \hat{\sigma}^2 R)}{L(\hat{M}_1, \hat{M}_2, \hat{\sigma}^2)} = \frac{(2\pi \hat{\sigma}^2 R)^{-n} \exp \left[-\frac{1}{2\hat{\sigma}^2 R} \sum_{i=1}^n (x_i^2 + y_i^2) \right]}{(2\pi \hat{\sigma}^2)^{-n} \exp \left[-\frac{n}{(n-1)(S_1^2 + S_2^2)} (n-1)(S_1^2 + S_2^2) \right]} \\ &= \left(\frac{\hat{\sigma}^2}{\hat{\sigma}^2 R} \right)^n = \left[\frac{(n-1)(S_1^2 + S_2^2)}{(n-1)(S_1^2 + S_2^2) + n(\bar{x}^2 + \bar{y}^2)} \right]^n \end{aligned}$$

$$= \left(\frac{1}{1+F} \right)^n, \quad F = \frac{n(\bar{x}^2 + \bar{y}^2)}{(n-1)(S_1^2 + S_2^2)}$$

$$= \frac{\frac{n}{\alpha^2} \bar{x}^2 + \frac{n}{\alpha^2} \bar{y}^2}{\frac{(n-1)S_1^2}{\alpha^2} + \frac{(n-1)S_2^2}{\alpha^2}} \underset{\text{Under } H_0}{\sim} \frac{\chi^2(2)}{\chi^2(2n-2)}$$

$$\Rightarrow (n-1)F \sim \frac{\chi^2(2)/2}{\chi^2(2n-2)/(2n-2)} = F(2, 2n-2)$$

under H_0 .

$$\alpha = \Pr \{ \lambda(\vec{x}) \leq \lambda_\alpha \mid H_0 \} = \Pr \{ F \geq c \mid H_0 \}$$

$$= \Pr \{ (n-1)F \geq (n-1)c \mid H_0 \} = \Pr \{ F(2, 2n-2) \geq (n-1)c \}$$

$$= \Pr \{ F(2, 2n-2) \geq f(\alpha, 2, 2n-2) \}.$$

$$\Rightarrow c = \frac{1}{n-1} f(\alpha, 2, 2n-2).$$

\Rightarrow The critical region is

$$\left\{ \vec{x}: \frac{n(\bar{x}^2 + \bar{y}^2)}{(n-1)(S_1^2 + S_2^2)} \geq \frac{1}{n-1} f(\alpha, 2, 2n-2) \right\}.$$

5.8

Solution:

(a) We first consider

$$H_0: \alpha_1^2 = \alpha_2^2 \quad \text{against} \quad H_1: \alpha_1^2 \neq \alpha_2^2.$$

$$\vec{\theta} = (\mu_1, \mu_2, \alpha_1^2, \alpha_2^2) \quad \Theta_0 = \{ \vec{\theta}: \alpha_1^2 = \alpha_2^2 \}, \quad \Theta_1 = \Theta \setminus \Theta_0 = \Theta^* \setminus \Theta_0.$$

$$L(\vec{\theta}) = (2\pi\alpha_1^2)^{-\frac{m_1}{2}} \exp \left(-\frac{1}{2\alpha_1^2} \sum_{i=1}^{m_1} (x_i - \mu_1)^2 \right)$$

$$\times (2\pi\alpha_2^2)^{-\frac{m_2}{2}} \exp \left(-\frac{1}{2\alpha_2^2} \sum_{i=1}^{m_2} (y_i - \mu_2)^2 \right).$$

$$l(\vec{\theta}) = \log L(\vec{\theta}) = -\frac{m_1}{2} \log 2\pi - \frac{m_1}{2} \log \alpha_1^2 - \frac{1}{2\alpha_1^2} \sum_{i=1}^{m_1} (x_i - \mu_1)^2$$

$$-\frac{n_2}{2} \log 2\pi - \frac{n_2}{2} \log \alpha_2^2 - \frac{1}{2\alpha_2^2} \sum_{i=2}^n (\gamma_i - \mu_2)^2.$$

$$\Rightarrow \hat{\mu}_1 = \bar{x}, \quad \hat{\mu}_2 = \bar{y}, \quad \hat{\sigma}_k^2 = \frac{(n_k-1) S_k^2}{n_k}. \quad k=1,2.$$

Under H_0 , the restricted mles are

$$\hat{\mu}_1^R = \bar{x}, \quad \hat{\mu}_2^R = \bar{y}, \quad \hat{\sigma}_2^R = \frac{(n_1-1) S_1^2 + (n_2-1) S_2^2}{n_1+n_2}$$

$$\begin{aligned} \lambda(\bar{x}) &= \frac{(\hat{\sigma}_1^2)^{\frac{n_1}{2}} (\hat{\sigma}_2^2)^{\frac{n_2}{2}}}{(\hat{\sigma}_2^R)^{\frac{n_1+n_2}{2}}} \\ &= \frac{[(n_1+n_2)(n_1-1)/n_1]^{\frac{n_1}{2}} [(n_1+n_2)(n_2-1)(S_2^2/S_1^2)/n_2]^{\frac{n_2}{2}}}{((n_1-1)+(n_2-1)\frac{S_2^2}{S_1^2})^{\frac{n_1+n_2}{2}}}. \end{aligned}$$

$$\text{let } t = \frac{S_1^2}{S_2^2}.$$

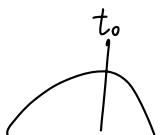
$$\Rightarrow \lambda(\bar{x}) \propto f(t) = \frac{t^{\frac{n_2}{2}}}{[n_1-1+(n_2-1)t]^{\frac{n_1+n_2}{2}}}.$$

$$\Rightarrow \log f(t) = \frac{n_2}{2} \log t - \frac{n_1+n_2}{2} \log [n_1-1+(n_2-1)t].$$

$$\frac{d \log f}{dt} = \frac{n_2}{2} \frac{1}{t} - \frac{(n_1+n_2)}{2} \frac{(n_2-1)}{[n_1-1+(n_2-1)t]}.$$

$$\frac{d^2 \log f}{dt^2} < 0 \quad \text{f}(t) \text{ is log-concave.}$$

$$t_0 = \frac{n_2(n_1-1)}{n_1(n_2-1)}$$



$$\Rightarrow C = \left\{ \bar{x} : t \leq C_1 \text{ or } t \geq C_2 \right\}.$$

We use equal-tail approach.

$$\frac{\alpha}{2} = \Pr \left\{ \frac{S_2^2}{S_1^2} \leq C_1 \mid H_0 \right\}$$

$$\text{Under } H_0, \quad \alpha_1^2 = \alpha_2^2, \quad \frac{S_2^2}{S_1^2} = \frac{\frac{(n_1-1) S_1^2}{\alpha_2^2} / (n_1-1)}{\frac{(n_1-1) S_1^2}{\alpha_1^2} / (n_1-1)} \sim F(n_2-1, n_1-1).$$

$$\Rightarrow \frac{\alpha}{2} = \Pr \{ F(n_2-1, n_1-1) \leq c_1 \}$$

$$= \Pr \{ F(n_2-1, n_1-1) \leq f(1 - \frac{\alpha}{2}, n_2-1, n_1-1) \}.$$

$$\Rightarrow C_1 = f(1 - \frac{\alpha}{2}, n_2-1, n_1-1).$$

$$\frac{\alpha}{2} = \Pr \left\{ \frac{s_2^2}{s_1^2} \geq c_2 \mid H_0 \right\}, \text{ Similarly, } C_2 = f(\frac{\alpha}{2}, n_2-1, n_1-1).$$

(b). Consider $H_0: \sigma_1^2 = \sigma_2^2$ against $H_1: \sigma_1^2 > \sigma_2^2$.

$$C = \left\{ \vec{x}: \frac{s_2^2}{s_1^2} \leq f(1-\alpha, n_2-1, n_1-1) \right\}.$$

(c). Consider $H_0: \sigma_1^2 = \sigma_2^2$ against $H_1: \sigma_1^2 < \sigma_2^2$.

$$C = \left\{ \vec{x}: \frac{s_2^2}{s_1^2} \geq f(\alpha, n_2-1, n_1-1) \right\}.$$

J.9.

Solution:

(a) Consider $H_0: \theta = \frac{1}{2}$ against $H_1: \theta \neq \frac{1}{2}$.

$$\Theta_0 = \left\{ \frac{1}{2} \right\}, \quad \Theta_1 = \Theta \setminus \Theta_0, \quad \Theta = \Theta^* = \mathbb{R}.$$

$$L(\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}.$$

$$l(\theta) = \log L(\theta) = \log \binom{n}{x} + x \log \theta + (n-x) \log (1-\theta).$$

$$\frac{dl}{d\theta} = \frac{x}{\theta} + \frac{x-n}{1-\theta} = 0 \Leftrightarrow x - x\theta = (x+n)\theta \Leftrightarrow \theta = \frac{x}{n}.$$

$$\lambda(x) = \frac{\binom{n}{x} \left(\frac{1}{2}\right)^x \left(\frac{1}{2}\right)^{n-x}}{\binom{n}{x} \left(\frac{x}{n}\right)^x \left(\frac{n-x}{n}\right)^{n-x}}$$

$$= \frac{\left(\frac{1}{2}n\right)^n}{x^x (n-x)^{n-x}}.$$

$$\Rightarrow \text{The LR statistic} \rightsquigarrow \lambda(X) = \frac{\left(\frac{1}{2}n\right)^n}{X^X (n-X)^{n-X}}$$

$$(b) \lambda(x) \leq \lambda_\alpha$$

$$\Leftrightarrow \frac{(\frac{1}{2}n)^n}{x^x(n-x)^{n-x}} \leq \lambda_\alpha$$

$$\Leftrightarrow n \log(\frac{1}{2}n) - x \log x - (n-x) \log(n-x) \leq \log \lambda_\alpha.$$

$$\Leftrightarrow x \log x + (n-x) \log(n-x) \geq n \log \frac{n}{2} - \log \lambda_\alpha \triangleq c.$$

(c).

$$f(x) = \log x + 1 - \log(n-x) - 1 = \log \frac{x}{n-x}$$

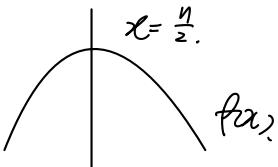
$$f(x) = f(n-x) = x \log x + (n-x) \log(n-x)$$

$\Rightarrow f(x)$ has symmetric axis $x = \frac{n}{2}$.

$$f(x) < 0, \quad x < \frac{n}{2}$$

$$f(x) > 0, \quad x > \frac{n}{2}$$

$$\Rightarrow$$



$$\Rightarrow f(x) \geq c_0 \Leftrightarrow |x - \frac{n}{2}| \geq c$$

5.11

Solution: $H_0: p_1 = \dots = p_6 = \frac{1}{6}$. against $H_1: \rightarrow (p_1 = \dots = p_6 = \frac{1}{6})$

$$\begin{aligned} Q_n &= \sum_{i=1}^6 \frac{(N_i - np_{i0})^2}{np_{i0}} \\ &= \frac{(43-50)^2}{50} + \frac{(49-50)^2}{50} + \frac{(56-50)^2}{50} \\ &\quad + \frac{(45-50)^2}{50} + \frac{(66-50)^2}{50} + \frac{(41-50)^2}{50} \end{aligned}$$

$$= 8.96 < \chi^2(0.05, 5) = 11.07$$

\Rightarrow we CANNOT reject H_0 .