
Appendix A:

Basic Statistical Distributions

A.1 Discrete Distributions

A.1.1 Finite discrete distribution

Notation: $X \sim \text{FDiscrete}_n(\mathbf{x}, \mathbf{p})$, $\mathbf{x} = (x_1, \dots, x_n)^\top$, $\mathbf{p} = (p_1, \dots, p_n)^\top \in \mathbb{T}_n \triangleq \{(p_1, \dots, p_n): p_i > 0, \sum_{i=1}^n p_i = 1\}$.

Density: $\Pr(X = x_i) = p_i$, $i = 1, \dots, n$.

Moments: $E(X) = \sum_{i=1}^n x_i p_i$ and $\text{Var}(X) = \sum_{i=1}^n x_i^2 p_i - (\sum_{i=1}^n x_i p_i)^2$.

Note: The *uniform discrete* distribution is a special case of the finite discrete distribution with $p_i = 1/n$ for all i .

Sampling: `sample(x, size, replace = FALSE, prob = NULL)` takes a sample of the specified size from the elements of `x` using either with or without replacement.

Examples:

```
> sample(c(0,1), 100, replace= T, prob=c(0.8, 0.2))
> sample(1:20, 4)      # the default: replace= F
```

A.1.2 Hypergeometric distribution

Notation: $X \sim \text{Hgeometric}(m, n, k)$, m, n, k are positive integers.

Density: $\text{Hgeometric}(x|m, n, k) = \binom{m}{x} \binom{n}{k-x} / \binom{m+n}{k}$,
where $x = \max(0, k - n), \dots, \min(m, k)$.

Moments: $E(X) = km/N'$ and $\text{Var}(X) = kmn(N' - k)/[N'^2(N' - 1)]$,
where $N' \triangleq m + n$.

Computing:

```
> prod(5:1) = 5!
> prod(20:16) = 20 × 19 × 18 × 17 × 16
> choose(40,5) =  $\binom{40}{5}$ 
```

Functions: `dhyper(x, m, n, k)`
`phyper(q, m, n, k)`
`qhyper(p, m, n, k)`
`rhyper(nn, m, n, k)`

A.1.3 Poisson distribution

Notation: $X \sim \text{Poisson}(\lambda)$, $\lambda > 0$

Density: $\text{Poisson}(x|\lambda) = \lambda^x e^{-\lambda}/x!$, $x = 0, 1, \dots, \infty$.

Moments: $E(X) = \lambda$ and $\text{Var}(X) = \lambda$.

Properties: • If $\{X_i\}_{i=1}^n \stackrel{\text{ind}}{\sim} \text{Poisson}(\lambda_i)$, then

$$\sum_{i=1}^n X_i \sim \text{Poisson}(\sum_{i=1}^n \lambda_i), \quad \text{and} \\ (X_1, \dots, X_n) | (\sum_{i=1}^n X_i = m) \sim \text{Multinomial}_n(m, \mathbf{p}),$$

where $\mathbf{p} = (\lambda_1, \dots, \lambda_n)^\top / \sum_{i=1}^n \lambda_i$;

• The Poisson and gamma distribution have relationship:

$$\sum_{x=k}^{\infty} \text{Poisson}(x|\lambda) = \int_0^\lambda \text{Gamma}(y|k, 1) \, dy.$$

Functions: `dpois(x, lambda)`
`ppois(q, lambda)`
`qpois(p, lambda)`
`rpois(n, lambda)`

```
=====
> x <- 0:20
> plot(x, dpois(x, 4), type="h")           # histogram-like
                                           # Figure A.1
*****
```

A.1.4 Binomial distribution

Notation: $X \sim \text{Binomial}(n, p)$, n is a positive integer, $p \in (0, 1)$.

Density: $\text{Binomial}(x|n, p) = \binom{n}{x} p^x (1-p)^{n-x}$, $x = 0, 1, \dots, n$.

Moments: $E(X) = np$ and $\text{Var}(X) = np(1-p)$.

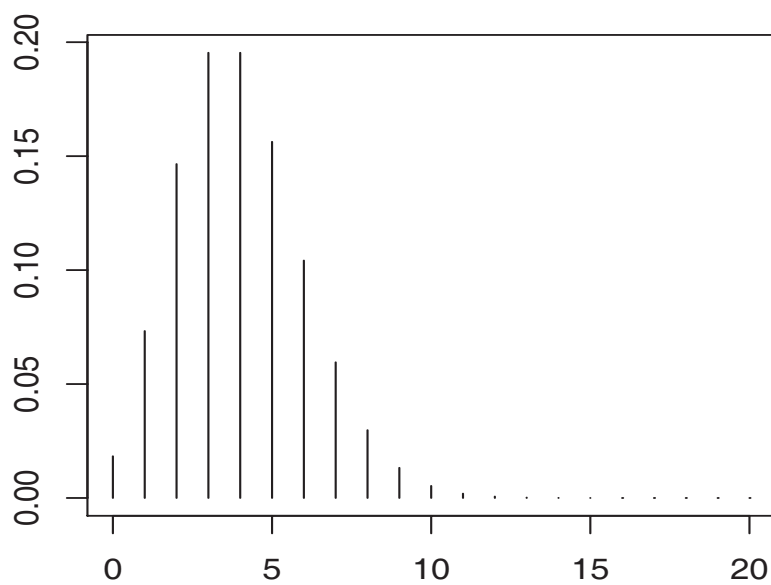


Figure A.1 Point probabilities of $\text{Poisson}(4)$.

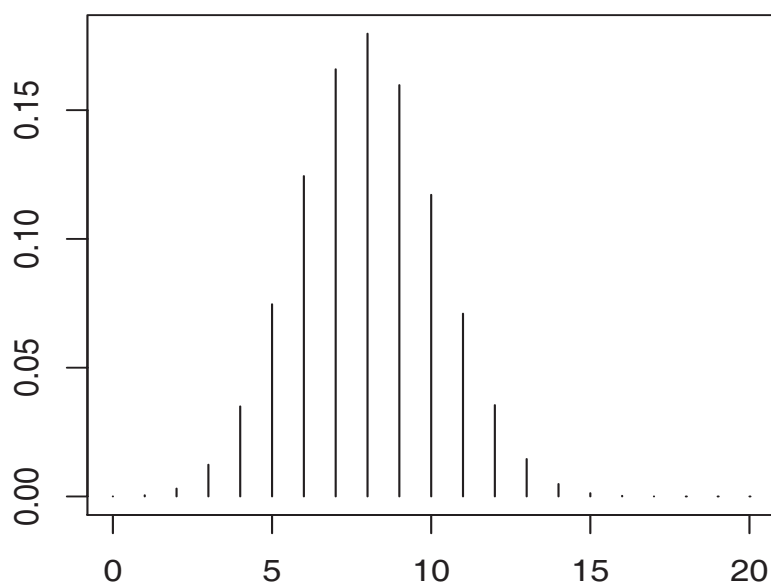


Figure A.2 Point probabilities of $\text{Binomial}(20, 0.4)$.

Properties: • If $\{X_i\}_{i=1}^d \stackrel{\text{ind}}{\sim} \text{Binomial}(n_i, p)$, then

$$\sum_{i=1}^d X_i \sim \text{Binomial}(\sum_{i=1}^d n_i, p);$$

• The binomial and beta distribution have relationship:

$$\sum_{x=0}^k \text{Binomial}(x|n, p) = \int_0^{1-p} \text{Beta}(x|n-k, k+1) dx,$$

where $0 \leq k \leq n$.

Note: When $n = 1$, binomial distribution is called *Bernoulli* distribution.

Functions: `dbinom(x, size, prob)` # size= n, prob= p
`pbinom(q, size, prob)`
`qbinom(p, size, prob)`
`rbinom(nn, size, prob)`

```
=====
> x <- 0:20
> plot(x, dbinom(x, size=20, prob=0.4), type="h")
# Figure A.2
*****
```

A.1.5 Multinomial distribution

Notation: $\mathbf{x} = (X_1, \dots, X_d)^\top \sim \text{Multinomial}(n; p_1, \dots, p_d)$ or
 $\mathbf{x} = (X_1, \dots, X_d)^\top \sim \text{Multinomial}_d(n, \mathbf{p})$,
 n is a positive integer, $\mathbf{p} = (p_1, \dots, p_d)^\top \in \mathbb{T}_d$,

Density: $\text{Multinomial}_d(\mathbf{x}|n, \mathbf{p}) = \binom{n}{x_1, \dots, x_d} \prod_{i=1}^d p_i^{x_i}$,
 $\mathbf{x} = (x_1, \dots, x_d)^\top$, $x_i \geq 0$, $\sum_{i=1}^d x_i = n$.

Moments: $E(X_i) = np_i$, $\text{Var}(X_i) = np_i(1-p_i)$ and $\text{Cov}(X_i, X_j) = -np_i p_j$.

Note: The binomial distribution is a special case of the multinomial with $d = 2$.

Functions: `dmultinom(x, size = NULL, prob)` # size= n, prob= \mathbf{p}
`rmultinom(nn, size, prob)`

A.2 Continuous Distributions

A.2.1 Uniform distribution

Notation: $X \sim U(a, b)$ or $X \sim U[a, b]$, $a < b$

Density: $U(x|a, b) = 1/(b - a)$, $x \in (a, b)$ or $x \in [a, b]$.

Moments: $E(X) = (a + b)/2$ and $\text{Var}(X) = (b - a)^2/12$.

Properties: If $Y \sim U(0, 1)$, then $X = a + (b - a)Y \sim U(a, b)$.

Functions: `dunif(x, min= 0, max= 1)` # min= a, max= b
 `punif(q, min= 0, max= 1)`
 `qunif(p, min= 0, max= 1)`
 `runif(n, min= 0, max= 1)`

A.2.2 Beta distribution

Notation: $X \sim \text{Beta}(a, b)$, $a > 0, b > 0$.

Density: $\text{Beta}(x|a, b) = x^{a-1}(1 - x)^{b-1}/B(a, b)$, $0 < x < 1$.

Moments: $E(X) = a/(a + b)$, $E(X^2) = a(a + 1)/[(a + b)(a + b + 1)]$ and $\text{Var}(X) = ab/[(a + b)^2(a + b + 1)]$.

Properties: If $Y_1 \sim \text{Gamma}(a, 1)$, $Y_2 \sim \text{Gamma}(b, 1)$, and $Y_1 \perp\!\!\!\perp Y_2$, then $Y_1/(Y_1 + Y_2) \sim \text{Beta}(a, b)$.

Note: When $a = b = 1$, $\text{Beta}(1, 1) = U(0, 1)$.

Functions: `dbeta(x, shape1, shape2)` # shape1= a, shape2= b
 `pbeta(q, shape1, shape2)`
 `qbeta(p, shape1, shape2)`
 `rbeta(n, shape1, shape2)`

A.2.3 Exponential distribution

Notation: $X \sim \text{Exponential}(\beta)$, rate parameter $\beta > 0$.

Density: $\text{Exponential}(x|\beta) = \beta e^{-\beta x}$, $x \geq 0$.

Moments: $E(X) = 1/\beta$ and $\text{Var}(X) = 1/\beta^2$.

Properties: • If $U \sim U(0, 1)$, then $-\log(U)/\beta \sim \text{Exponential}(\beta)$;

• If $\{X_i\}_{i=1}^n \stackrel{\text{iid}}{\sim} \text{Exponential}(\beta)$, then $\sum_{i=1}^n X_i \sim \text{Gamma}(n, \beta)$.

Functions: `dexp(x, rate= 1)` `# rate= β`
`pexp(q, rate= 1)`
`qexp(p, rate= 1)`
`rexp(n, rate= 1)`

A.2.4 Gamma distribution

Notation: $X \sim \text{Gamma}(\alpha, \beta)$, shape parameter $\alpha > 0$, rate parameter $\beta > 0$.

Density: $\text{Gamma}(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x}$, $x > 0$.

Moments: $E(X) = \alpha/\beta$ and $\text{Var}(X) = \alpha/\beta^2$.

Properties: • If $X \sim \text{Gamma}(\alpha, \beta)$ and $c > 0$, then $cX \sim \text{Gamma}(\alpha, \beta/c)$;

• If $\{X_i\}_{i=1}^n \stackrel{\text{iid}}{\sim} \text{Gamma}(\alpha_i, \beta)$, then $\sum X_i \sim \text{Gamma}(\sum \alpha_i, \beta)$;

• $\Gamma(\alpha + 1) = \alpha\Gamma(\alpha)$, $\Gamma(1) = 1$ and $\Gamma(1/2) = \sqrt{\pi}$.

Note: $\text{Gamma}(1, \beta) = \text{Exponential}(\beta)$. $\text{Gamma}(\nu/2, 1/2) = \chi^2(\nu)$.

Functions: `dgamma(x, shape, rate= 1)` `# shape= α , rate= β`
`pgamma(q, shape, rate= 1)`
`qgamma(p, shape, rate= 1)`
`rgamma(n, shape, rate= 1)`

A.2.5 Chi-square distribution

Notation: $X \sim \chi^2(\nu) \equiv \text{Gamma}(\frac{\nu}{2}, \frac{1}{2})$, degrees of freedom $\nu > 0$.

Density: $\chi^2(x|\nu) = \frac{2^{-\nu/2}}{\Gamma(\nu/2)} x^{\nu/2-1} e^{-x/2}$, $x > 0$.

Moments: $E(X) = \nu$ and $\text{Var}(X) = 2\nu$.

Properties: • If $Y \sim N(0, 1)$, then $X = Y^2 \sim \chi^2(1)$;

• If $\{X_j\}_{j=1}^m \stackrel{\text{iid}}{\sim} \chi^2(\nu_j)$, then $\sum_{j=1}^m X_j \sim \chi^2(\sum_{j=1}^m \nu_j)$.

Functions: dchisq(x, df) # df = nu
 pchisq(q, df)
 qchisq(p, df)
 rchisq(nn, df)

```
=====
> x <- seq(0.01, 25, 0.1)
> par(mfrow=c(2, 2))                    # Figure A.3
> curve(dchisq(x, df= 1), from=0.1, to = 25)
> curve(dchisq(x, df= 2), from=0.1, to = 25)
> curve(dchisq(x, df= 3), from=0.1, to = 25)
> curve(dchisq(x, df= 4), from=0.1, to = 25)
*****
```

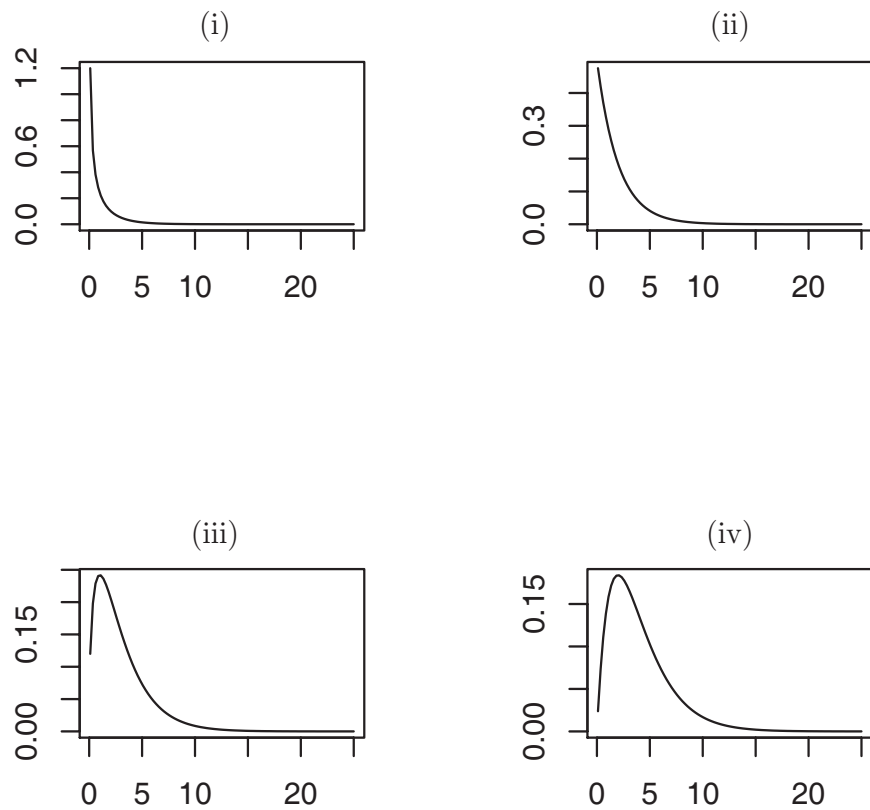


Figure A.3 Density functions of $\chi^2(\nu)$ for various ν . (i) $\nu = 1$; (ii) $\nu = 2$; (iii) $\nu = 3$; (iv) $\nu = 4$.

A.2.6 t - or Student's t -distribution

Notation: $X \sim t(\nu)$, $\nu > 0$ is a positive real number.

Density: $t(x|\nu) = \frac{\Gamma(\frac{\nu+1}{2})}{\sqrt{\pi\nu}\Gamma(\frac{\nu}{2})} \left(1 + \frac{x^2}{\nu}\right)^{-\frac{\nu+1}{2}}, \quad -\infty < x < \infty.$

Moments: $E(X) = 0$ (if $\nu > 1$) and $\text{Var}(X) = \frac{\nu}{\nu-2}$ (if $\nu > 2$).

Properties: Let $Z \sim N(0, 1)$, $Y \sim \chi^2(\nu)$, and $Z \perp\!\!\!\perp Y$, then

$$\frac{Z}{\sqrt{Y/\nu}} \sim t(\nu).$$

Note: When $\nu = 1$, $t(\nu) = t(1)$ is called *standard Cauchy distribution*, whose mean and variance are undefined.

Functions: `dt(x, df)` `# df = nu`
 `pt(q, df)`
 `qt(p, df)`
 `rt(nn, df)`

A.2.7 F or Fisher's F -distribution

Notation: $X \sim F(n_1, n_2)$, n_1, n_2 are positive integers.

Density: $F(x|n_1, n_2) = \frac{(n_1/n_2)^{n_1/2}}{B(\frac{n_1}{2}, \frac{n_2}{2})} x^{\frac{n_1}{2}-1} (1 + \frac{n_1 x}{n_2})^{-\frac{n_1+n_2}{2}}, \quad x > 0.$

Moments: $E(X) = \frac{n_2}{n_2-2}$ (if $n_2 > 2$), $\text{Var}(X) = \frac{2n_2^2(n_1+n_2-2)}{n_1(n_2-4)(n_2-2)^2}$ (if $n_2 > 4$).

Properties: Let $Y_i \sim \chi^2(n_i)$, $i = 1, 2$, and $Y_1 \perp\!\!\!\perp Y_2$, then

$$\frac{Y_1/n_1}{Y_2/n_2} \sim F(n_1, n_2).$$

Functions: `df(x, df1, df2)` `# df1= n1, df2= n2`
 `pf(q, df1, df2)`
 `qf(p, df1, df2)`
 `rf(n, df1, df2)`

A.2.8 Normal or Gaussian distribution

Notation: $X \sim N(\mu, \sigma^2)$, $-\infty < \mu < \infty$, $\sigma^2 > 0$.

Density: $N(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp[-\frac{(x-\mu)^2}{2\sigma^2}]$, $-\infty < x < \infty$.

Moments: $E(X) = \mu$ and $\text{Var}(X) = \sigma^2$.

Properties: • If $\{X_i\} \stackrel{\text{ind}}{\sim} N(\mu_i, \sigma_i^2)$, then $\sum a_i X_i \sim N(\sum a_i \mu_i, \sum a_i^2 \sigma_i^2)$;

• If $X_1|X_2 \sim N(X_2, \sigma_1^2)$ and $X_2 \sim N(\mu_2, \sigma_2^2)$, then

$$X_1 \sim N(\mu_2, \sigma_1^2 + \sigma_2^2).$$

Functions: `dnorm(x, mean=0, sd= 1) # mean= μ , sd= σ`
`pnorm(q, mean=0, sd= 1)`
`qnorm(p, mean=0, sd= 1)`
`rnorm(n, mean=0, sd= 1)`

```
=====
> x <- seq(-4, 4, 0.1)
> plot(x, dnorm(x), type="l",
      ylab="Density function of N(0,1)")
# Note that this is the letter "l", not the digit "1"
# Figure A.4
-----

# An alternative way of creating the plot is

> curve(dnorm(x), from=-4, to = 4,
      ylab="Density function of N(0,1)")
*****
```

A.2.9 Multivariate normal or Gaussian distribution

Notation: $\mathbf{x} = (X_1, \dots, X_d)^\top \sim N_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ or $N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, $\boldsymbol{\mu} \in \mathbb{R}^d$, $\boldsymbol{\Sigma} > 0$.

Density: $N_d(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(\sqrt{2\pi})^d |\boldsymbol{\Sigma}|^{\frac{1}{2}}} \exp\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\}$, $\mathbf{x} \in \mathbb{R}^d$.

Moments: $E(\mathbf{x}) = \boldsymbol{\mu}$ and $\text{Var}(\mathbf{x}) = \boldsymbol{\Sigma}$.

Functions: Producing one or more samples from the specified multivariate normal distribution

```
mvrnorm(n= 1, mu, Sigma, tol= 1e-6, empirical= F)
```

```
rmvn(n, mu, V)
```

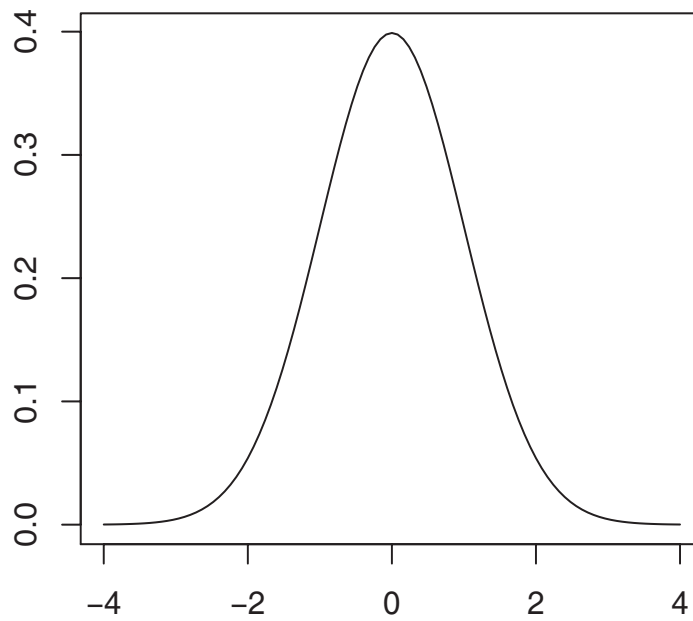


Figure A.4 Density functions of $N(0,1)$.