MAT271E - HOMEWORK 3

Name: Ö. Malik KALEMBAŞI

ID: 150180112

CRN: 13285

1. The hypothesis that mean length of battery charge time of iPhone 13 is under same circumstances is 50 minutes ($\mu_x = 50$) will be tested with alternate hypothesis that is length of battery charge time different from 50 minutes.

The length of battery charge of iPhone 13 is assumed to be normally distributed, with a standart deviation of 6 minutes ($\hat{s}_x = 6$).

Suppose the mean length of time estimated from measurements over N=11 times is: $\overline{x}=48$ minutes, at the level of significance of 0.05 ($\alpha=0.05$).

For small samples the sampling distribution of the mean is t distribution with degrees of freedom N-1=10.

For two-sided hypothesis test, 0.05/2=0.025 and $t_{10/0.025}=2.228$ (according to t-distribution table). The boundaries of the acceptence region:

$$b_1 = \mu_\chi - rac{t \cdot \hat{s}_\chi}{\sqrt{N}}$$
 and $b_2 = \mu_\chi + rac{t \cdot \hat{s}_\chi}{\sqrt{N}}$

$$b_1 = 50 - \frac{(2.228) \cdot 6}{\sqrt{11}} = 45.96939635$$

$$b_2 = 50 + \frac{(2.228)\cdot 6}{\sqrt{11}} = 54.03060365$$

As a result, $\overline{x} = 48$ is in the interval (45.96939635-54.03060365). So, hypothesis is **ACCEPTED**.

2. Changing alternate hypothesis to $\mu_x < 50$, for the same significance $\alpha = 0.05$, same sample size N = 11, same standard deviation $\hat{s}_x = 6$ and same mean of the sample $\bar{x} = 48$.

For the one-sided hypothesis test, **0.05** and $t_{10/0.05} = 1.812$. The lower boundary of the acceptance region:

$$50 - \frac{(1.812) \cdot 6}{\sqrt{11}} = 46.72196866$$

As a result, $\overline{x} = 48 > 46.72196866$. So, hypothesis is **ACCEPTED**.

3. A large group of a students took a final exam in Statistics and its claimed to be standard deviation of the grades different from $\sigma_{\chi} = 18$, at the level of significance of $\alpha = 0.10$, for a sample size N = 15. Samples are shown below:

$$H_0: \sigma_x = 18$$

$$H_1: \sigma_x \neq 18$$

For left-side,
$$\chi^2_{(\frac{N-1}{2},\frac{1-\alpha}{2})} = \chi^2_{(7,0.45)} \cong 6.346$$

For right-side,
$$\chi^2_{(\frac{N-1}{2},\frac{\alpha}{2})} = \chi^2_{(7,0.05)} = 14.067$$

$$\chi^2(test) = \frac{(N-1)\cdot s^2}{\sigma^2}$$

$$\mu_{x} = 51$$

$$S^2 = \sum \frac{(x_i - \mu_x)^2}{N-1} = 614$$

$$\chi^2(test) = \frac{(N-1)\cdot s^2}{\sigma^2} = \frac{14\cdot 614}{324} = 26.5308642$$

As a result, **26.530842** is not in the interval (6.346-14.067). So, hypothesis H_0 is rejected. Hypothesis H_1 , Standard deviation of the grades are different from $\sigma_{\chi}=18$ is **ACCEPTED** at the level of significance of $\alpha=0.10$.

4. A large group of a students took a final exam in Statistics and its claimed to be standard deviation of the grades higher than $\sigma_x > 18$, at the level of significance of $\alpha = 0.10$, for a sample size N = 15. Samples are shown below:

$$H_0: \sigma_x = 18$$

$$H_1: \sigma_x > 18$$

$$\chi^2_{(N-1,\alpha)} = \chi^2_{(14,0.10)} = 21.064$$

$$\chi^2(test) = \frac{(N-1)\cdot s^2}{\sigma^2}$$

$$\mu_{x} = 51$$

$$S^2 = \sum \frac{(x_i - \mu_x)^2}{N-1} = 614$$

$$\chi^2(test) = \frac{(N-1)\cdot s^2}{\sigma^2} = \frac{14\cdot 614}{324} = 26.5308642$$

As a result, **26.530842 > 21.064** and $\chi^2(test) > \chi^2(critical\ value)$. So, hypothesis H_0 is rejected. Hypothesis H_1 , Standard deviation of the grades are higher than $\sigma_{\chi} > 18$ is **ACCEPTED** at the level of significance of $\alpha = 0.10$.

| t Table | | | | | | | | | | | |
|-----------|------------------|-------|-------|-------|-------|-------|-------|-------|-------|--------|---------|
| cum. prob | t.50 | t.75 | t.80 | t .85 | t .90 | t.95 | t.975 | t.99 | t.995 | t.999 | t .9995 |
| one-tail | 0.50 | 0.25 | 0.20 | 0.15 | 0.10 | 0.05 | 0.025 | 0.01 | 0.005 | 0.001 | 0.0005 |
| two-tails | 1.00 | 0.50 | 0.40 | 0.30 | 0.20 | 0.10 | 0.05 | 0.02 | 0.01 | 0.002 | 0.001 |
| df | | | | | | | | | | | |
| 1 | 0.000 | 1.000 | 1.376 | 1.963 | 3.078 | 6.314 | 12.71 | 31.82 | 63.66 | 318.31 | 636.62 |
| 2 | 0.000 | 0.816 | 1.061 | 1.386 | 1.886 | 2.920 | 4.303 | 6.965 | 9.925 | 22.327 | 31.599 |
| 3 | 0.000 | 0.765 | 0.978 | 1.250 | 1.638 | 2.353 | 3.182 | 4.541 | 5.841 | 10.215 | 12.924 |
| 4 | 0.000 | 0.741 | 0.941 | 1.190 | 1.533 | 2.132 | 2.776 | 3.747 | 4.604 | 7.173 | 8.610 |
| 5 | 0.000 | 0.727 | 0.920 | 1.156 | 1.476 | 2.015 | 2.571 | 3.365 | 4.032 | 5.893 | 6.869 |
| 6 | 0.000 | 0.718 | 0.906 | 1.134 | 1.440 | 1.943 | 2.447 | 3.143 | 3.707 | 5.208 | 5.959 |
| 7 | 0.000 | 0.711 | 0.896 | 1.119 | 1.415 | 1.895 | 2.365 | 2.998 | 3.499 | 4.785 | 5.408 |
| 8 | 0.000 | 0.706 | 0.889 | 1.108 | 1.397 | 1.860 | 2.306 | 2.896 | 3.355 | 4.501 | 5.041 |
| 9 | 0.000 | 0.703 | 0.883 | 1.100 | 1.383 | 1.833 | 2.262 | 2.821 | 3.250 | 4.297 | 4.781 |
| 10 | 0.000 | 0.700 | 0.879 | 1.093 | 1.372 | 1.812 | 2.228 | 2.764 | 3.169 | 4.144 | 4.587 |
| 11 | 0.000 | 0.697 | 0.876 | 1.088 | 1.363 | 1.796 | 2.201 | 2.718 | 3.106 | 4.025 | 4.437 |
| 12 | 0.000 | 0.695 | 0.873 | 1.083 | 1.356 | 1.782 | 2.179 | 2.681 | 3.055 | 3.930 | 4.318 |
| 13 | 0.000 | 0.694 | 0.870 | 1.079 | 1.350 | 1.771 | 2.160 | 2.650 | 3.012 | 3.852 | 4.221 |
| 14 | 0.000 | 0.692 | 0.868 | 1.076 | 1.345 | 1.761 | 2.145 | 2.624 | 2.977 | 3.787 | 4.140 |
| 15 | 0.000 | 0.691 | 0.866 | 1.074 | 1.341 | 1.753 | 2.131 | 2.602 | 2.947 | 3.733 | 4.073 |
| 16 | 0.000 | 0.690 | 0.865 | 1.071 | 1.337 | 1.746 | 2.120 | 2.583 | 2.921 | 3.686 | 4.015 |
| 17 | 0.000 | 0.689 | 0.863 | 1.069 | 1.333 | 1.740 | 2.110 | 2.567 | 2.898 | 3.646 | 3.965 |
| 18 | 0.000 | 0.688 | 0.862 | 1.067 | 1.330 | 1.734 | 2.101 | 2.552 | 2.878 | 3.610 | 3.922 |
| 19 | 0.000 | 0.688 | 0.861 | 1.066 | 1.328 | 1.729 | 2.093 | 2.539 | 2.861 | 3.579 | 3.883 |
| 20 | 0.000 | 0.687 | 0.860 | 1.064 | 1.325 | 1.725 | 2.086 | 2.528 | 2.845 | 3.552 | 3.850 |
| 21 | 0.000 | 0.686 | 0.859 | 1.063 | 1.323 | 1.721 | 2.080 | 2.518 | 2.831 | 3.527 | 3.819 |
| 22 | 0.000 | 0.686 | 0.858 | 1.061 | 1.321 | 1.717 | 2.074 | 2.508 | 2.819 | 3.505 | 3.792 |
| 23 | 0.000 | 0.685 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069 | 2.500 | 2.807 | 3.485 | 3.768 |
| 24 | 0.000 | 0.685 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064 | 2.492 | 2.797 | 3.467 | 3.745 |
| 25 | 0.000 | 0.684 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060 | 2.485 | 2.787 | 3.450 | 3.725 |
| 26 | 0.000 | 0.684 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056 | 2.479 | 2.779 | 3.435 | 3.707 |
| 27 | 0.000 | 0.684 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052 | 2.473 | 2.771 | 3.421 | 3.690 |
| 28 | 0.000 | 0.683 | 0.855 | 1.056 | 1.313 | 1.701 | 2.048 | 2.467 | 2.763 | 3.408 | 3.674 |
| 29 | 0.000 | 0.683 | 0.854 | 1.055 | 1.311 | 1.699 | 2.045 | 2.462 | 2.756 | 3.396 | 3.659 |
| 30 | 0.000 | 0.683 | 0.854 | 1.055 | 1.310 | 1.697 | 2.042 | 2.457 | 2.750 | 3.385 | 3.646 |
| 40 | 0.000 | 0.681 | 0.851 | 1.050 | 1.303 | 1.684 | 2.021 | 2.423 | 2.704 | 3.307 | 3.551 |
| 60 | 0.000 | 0.679 | 0.848 | 1.045 | 1.296 | 1.671 | 2.000 | 2.390 | 2.660 | 3.232 | 3.460 |
| 80 | 0.000 | 0.678 | 0.846 | 1.043 | 1.292 | 1.664 | 1.990 | 2.374 | 2.639 | 3.195 | 3.416 |
| 100 | 0.000 | 0.677 | 0.845 | 1.042 | 1.290 | 1.660 | 1.984 | 2.364 | 2.626 | 3.174 | 3.390 |
| 1000 | 0.000 | 0.675 | 0.842 | 1.037 | 1.282 | 1.646 | 1.962 | 2.330 | 2.581 | 3.098 | 3.300 |
| Z | 0.000 | 0.674 | 0.842 | 1.036 | 1.282 | 1.645 | 1.960 | 2.326 | 2.576 | 3.090 | 3.291 |
| | 0% | 50% | 60% | 70% | 80% | 90% | 95% | 98% | 99% | 99.8% | 99.9% |
| Г | Confidence Level | | | | | | | | | | |

Chi-Square Table

Table 5-2 ${\rm Critical\ Values\ of\ the\ } \chi^{\rm \, 2} {\rm \, Distribution}$

| \ p | | | | | | | | | | |
|------|-------|-------|-------|--------|--------|--------|--------|--------|--------|----|
| df \ | 0.995 | 0.975 | 0.9 | 0.5 | 0.1 | 0.05 | 0.025 | 0.01 | 0.005 | df |
| 1 | .000 | .000 | 0.016 | 0.455 | 2.706 | 3.841 | 5.024 | 6.635 | 7.879 | 1 |
| 2 | 0.010 | 0.051 | 0.211 | 1.386 | 4.605 | 5.991 | 7.378 | 9.210 | 10.597 | 2 |
| 3 | 0.072 | 0.216 | 0.584 | 2.366 | 6.251 | 7.815 | 9.348 | 11.345 | 12.838 | 3 |
| 4 | 0.207 | 0.484 | 1.064 | 3.357 | 7.779 | 9.488 | 11.143 | 13.277 | 14.860 | 4 |
| 5 | 0.412 | 0.831 | 1.610 | 4.351 | 9.236 | 11.070 | 12.832 | 15.086 | 16.750 | 5 |
| 6 | 0.676 | 1.237 | 2.204 | 5.348 | 10.645 | 12.592 | 14.449 | 16.812 | 18.548 | 6 |
| 7 | 0.989 | 1.690 | 2.833 | 6.346 | 12.017 | 14.067 | 16.013 | 18.475 | 20.278 | 7 |
| 8 | 1.344 | 2.180 | 3.490 | 7.344 | 13.362 | 15.507 | 17.535 | 20.090 | 21.955 | 8 |
| 9 | 1.735 | 2.700 | 4.168 | 8.343 | 14.684 | 16.919 | 19.023 | 21.666 | 23.589 | 9 |
| 10 | 2.156 | 3.247 | 4.865 | 9.342 | 15.987 | 18.307 | 20.483 | 23.209 | 25.188 | 10 |
| 11 | 2.603 | 3.816 | 5.578 | 10.341 | 17.275 | 19.675 | 21.920 | 24.725 | 26.757 | 11 |
| 12 | 3.074 | 4.404 | 6.304 | 11.340 | 18.549 | 21.026 | 23.337 | 26.217 | 28.300 | 12 |
| 13 | 3.565 | 5.009 | 7.042 | 12.340 | 19.812 | 22.362 | 24.736 | 27.688 | 29.819 | 13 |
| 14 | 4.075 | 5.629 | 7.790 | 13.339 | 21.064 | 23.685 | 26.119 | 29.141 | 31.319 | 14 |
| 15 | 4.601 | 6.262 | 8.547 | 14.339 | 22.307 | 24.996 | 27.488 | 30.578 | 32.801 | 15 |